HW07q2

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Question 2

Data

```
dat = expand.grid(factory=c("East", "West"), accident=c("No", "Yes"))
dat$y = c(645, 1275, 28, 31)
tab = matrix(dat$y, nrow=2,
dimnames=list(factory=c("East", "West"), accident=c("No", "Yes")))
```

2(a)

```
fit_2a = glm(y \sim factory + accident, poisson, dat)
fit_2a
##
## Call: glm(formula = y \sim factory + accident, family = poisson, data = dat)
##
## Coefficients:
## (Intercept) factoryWest accidentYes
##
         6.481
                      0.663
                                   -3.483
```

Degrees of Freedom: 3 Total (i.e. Null); 1 Residual ## Null Deviance: 2423 ## Residual Deviance: 4.678 AIC: 38.43 2(b)

predict(fit_2a, newdata = data.frame(factory = factor("West"),

```
##
             1
 ## 38.93583
To attain the result manually, we can use the equation bellow:
```

 $e^{(6.481+0.663+-3.483)} = 38.9$

2(c)

factoryWest

0.6815

fit_2c = glm(y ~ factory*accident, poisson, dat) fit_2c

predict(fit_2c, newdata = data.frame(factory = factor("West"),

$glm(formula = y \sim factory * accident, family = poisson, data = dat)$

accident = factor("Yes")), type = "response")

Call: $glm(formula = y \sim factory * accident, family = poisson, data = dat)$

accident = factor("Yes")), type = "response")

##

##

##

##

Coefficients:

(Intercept)

6.4693

```
## factoryWest:accidentYes
##
                   -0.5797
##
## Degrees of Freedom: 3 Total (i.e. Null); 0 Residual
## Null Deviance:
                        2423
## Residual Deviance: 1.061e-13
                                     AIC: 35.75
```

accidentYes

-3.1370

```
##
 ## 31
To attain the result manually, we can use the equation bellow:
                                                    e^{(6.4693+0.6815-3.1370-0.5797)} = 31.0
```

perfectly. Thus, when we subtract the log-likelihood of the saturated model from the log-likelihood of the predicted model, we will get zero.

We get a residual deviance of 0 because our predicted model is the saturated model. Our model can adjust to fit any row in the data frame

Call:

2(d)

2(e)

summary(fit_2c)

```
## Deviance Residuals:
## [1] 0 0 0 0
##
## Coefficients:
```

```
##
                              Estimate Std. Error z value Pr(>|z|)
                               6.46925 0.03937 164.299 <2e-16 ***
 ## (Intercept)
 ## factoryWest
                                           0.04832 14.103
                                                              <2e-16 ***
                               0.68145
                                           0.19304 -16.251 <2e-16 ***
 ## accidentYes
                              -3.13705
 ## factoryWest:accidentYes -0.57967
                                        0.26515 -2.186 0.0288 *
 ## ---
 ## Signif. codes: 0 '***' 0.001 '**' 0.01 '*' 0.05 '.' 0.1 ' ' 1
 ##
 ## (Dispersion parameter for poisson family taken to be 1)
 ##
        Null deviance: 2.4235e+03 on 3 degrees of freedom
 ##
 ## Residual deviance: 1.0614e-13 on 0 degrees of freedom
 ## AIC: 35.749
 ## Number of Fisher Scoring iterations: 3
The z-value for the interaction term is -2.186. This is less than -1.645. We can reject the null that eta_{factory*accident}=0 and conclude
\beta_{factory*accident} \neq 0.
2(f)
The result from question 2 part e tells us the west factory on average is less likely to have accidents.
```

d = drop1(fit_2c, test="Chisq")

y ~ factory * accident

2(h)

ype = "link")

type = "link")

d: -0.5796684

summary(fit_2j)

##

Model:

<none>

Single term deletions

Df Deviance

log_odds_east_2h = east_accident_1 - east_accident_0 log_odds_west_2h = west_accident_1 - west_accident_0

fit_2j = glm(accident ~ factory, binomial, dat, weights = y)

2(g)

```
## factory:accident 1 4.678 38.427 4.678 0.03055 *
## ---
## Signif. codes: 0 '***' 0.001 '**' 0.01 '*' 0.05 '.' 0.1 ' ' 1
p_val_2g = d^{r(>Chi)}[2]
chi_sq_2g = qchisq(p_val_2g, 1, lower.tail = FALSE)
```

LRT Pr(>Chi)

Using the likelihood ratio test to evaluate the interaction, we get a chi-squared value of 4.6779804 and a p-value of 0.0305516.

AIC

0.000 35.749

```
east_accident_0 = predict(fit_2c, newdata = data.frame(factory = factor("East"), accident = factor("No")), t
ype = "link")
east_accident_1 = predict(fit_2c, newdata = data.frame(factory = factor("East"), accident = factor("Yes")),
type = "link")
west_accident_0 = predict(fit_2c, newdata = data.frame(factory = factor("West"), accident = factor("No")), t
```

west_accident_1 = predict(fit_2c, newdata = data.frame(factory = factor("West"), accident = factor("Yes")),

```
• The log odds of an accident in the east is -3.1370458

    The log odds of an accident in the west is -3.7167143

2(i)
 c = log_odds_east_2h
 d = log_odds_west_2h - log_odds_east_2h
 cat("c: ", c, "\n", "d: ", d)
        -3.137046
 ## C:
```

```
=log(m_{i0})+log(rac{m_{i1}}{m_{i0}})*west
                                                      = -3.1370 - 0.5797 * west
2(j)
```

 $log(rac{\pi_{1|i}}{1-\pi_{1|i}}) = log(rac{\pi_{1|i}}{\pi_{0|i}})$

```
##
## glm(formula = accident ~ factory, family = binomial, data = dat,
      weights = y)
##
## Deviance Residuals:
                        3
```

```
-7.404 -7.827 13.344 15.229
##
## Coefficients:
##
              Estimate Std. Error z value Pr(>|z|)
## (Intercept) -3.1370 0.1930 -16.251 <2e-16 ***
## factoryWest -0.5797
                         0.2651 -2.186 0.0288 *
## Signif. codes: 0 '***' 0.001 '**' 0.01 '*' 0.05 '.' 0.1 ' ' 1
## (Dispersion parameter for binomial family taken to be 1)
##
##
      Null deviance: 530.73 on 3 degrees of freedom
## Residual deviance: 526.06 on 2 degrees of freedom
## AIC: 530.06
## Number of Fisher Scoring iterations: 6
```

2(k)

I would expect the estimated logistic regression to be just the intercept: $log(rac{\pi_{1|i}}{1-\pi_{1|i}})=-3.1370$