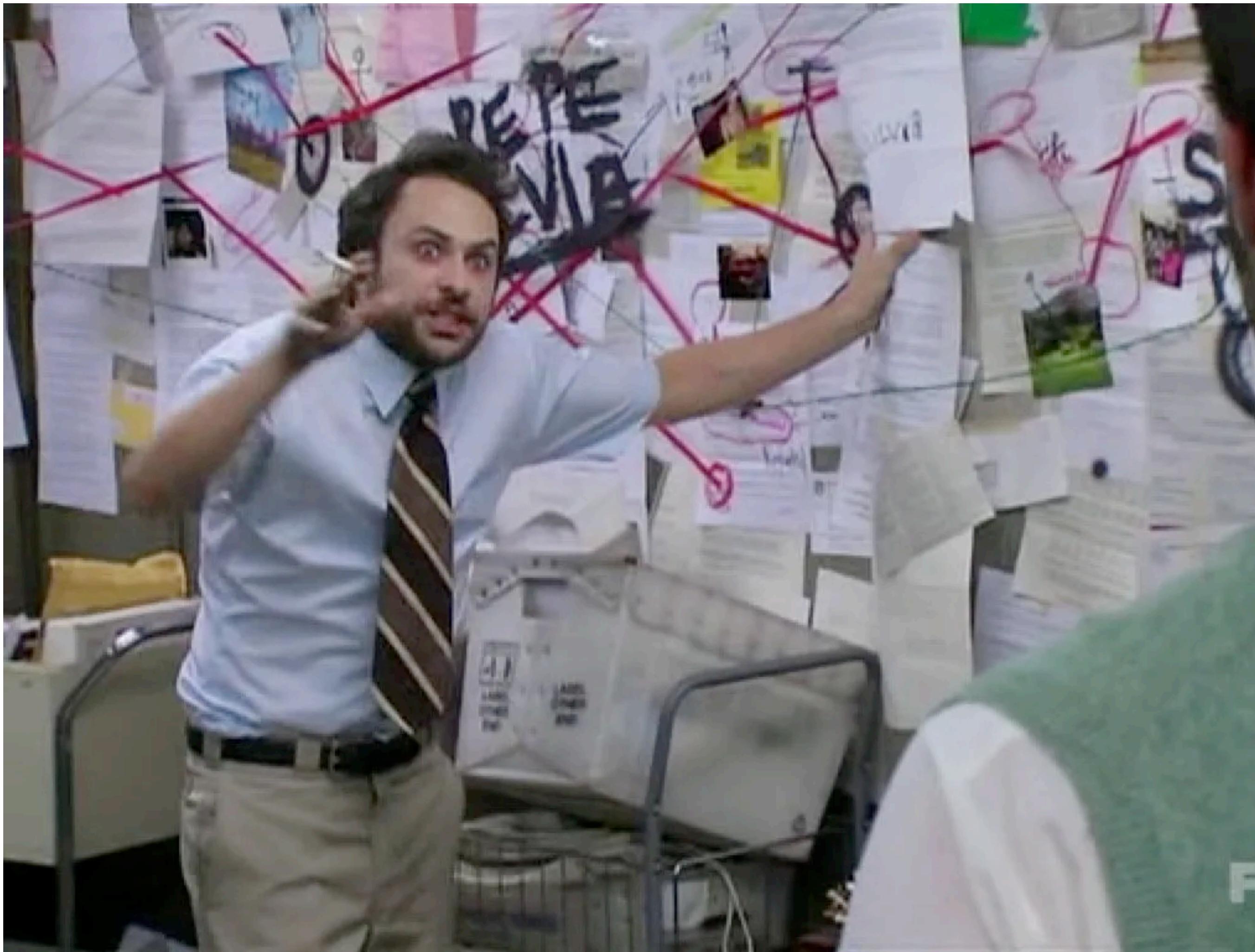


# Math for ML

Finale: Course Overview

By: Samuel Deng

# Lesson Overview



# Week 1.1

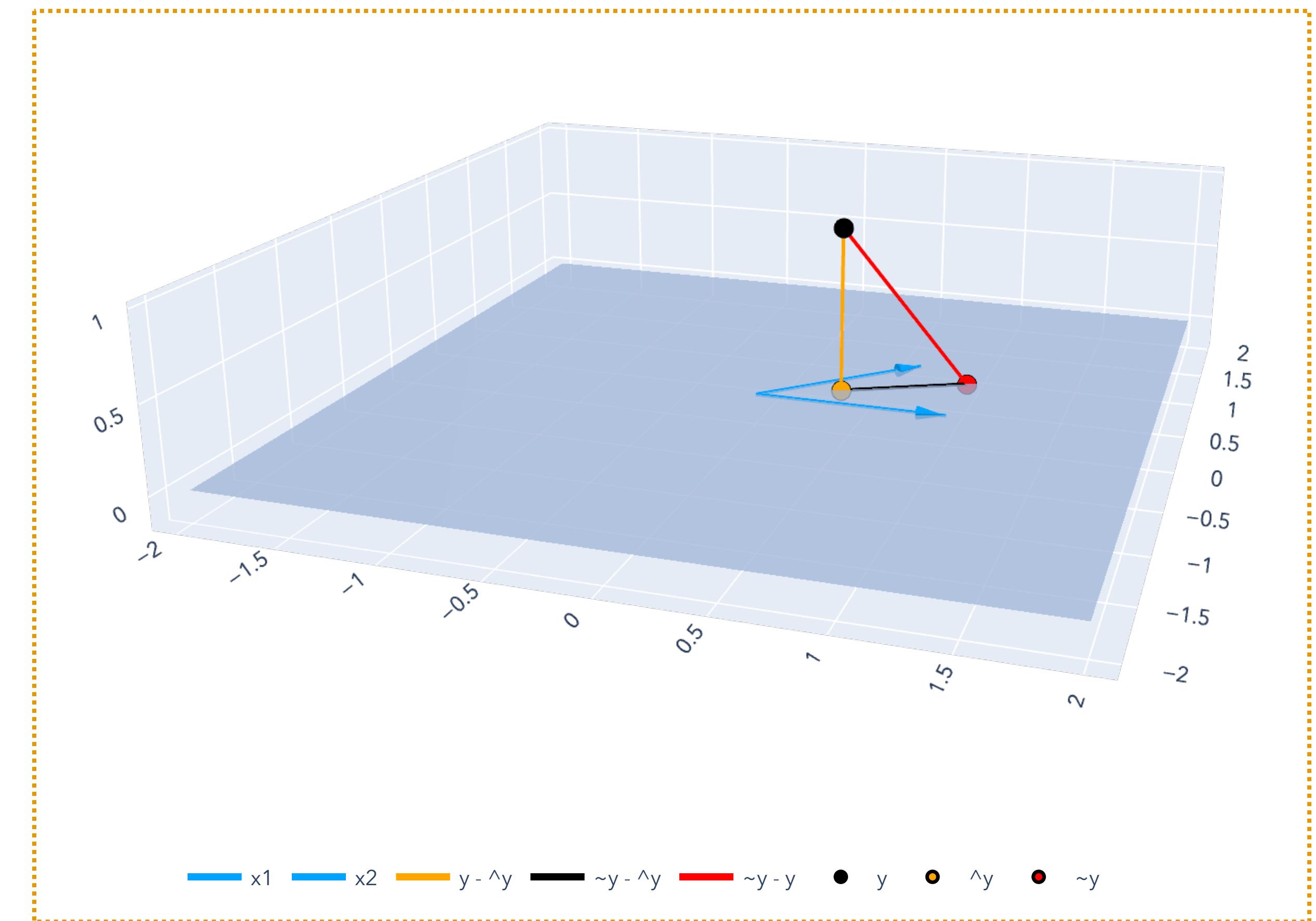
## Vectors, matrices, and least squares regression

# Vectors, matrices, and least squares regression

## Big Picture: Least Squares

Linear independence, span, and rank allowed us to get  $(\mathbf{X}^T \mathbf{X})^{-1}$  from  $\text{rank}(\mathbf{X}^T \mathbf{X}) = \text{rank}(\mathbf{X})$  sketching our first OLS theorem:

Theorem (OLS solution). If  $n \geq d$  and  $\text{rank}(\mathbf{X}) = d$ , then:  $\hat{\mathbf{w}} = (\mathbf{X}^T \mathbf{X})^{-1} \mathbf{X}^T \mathbf{y}$ .



# Vectors, matrices, and least squares regression

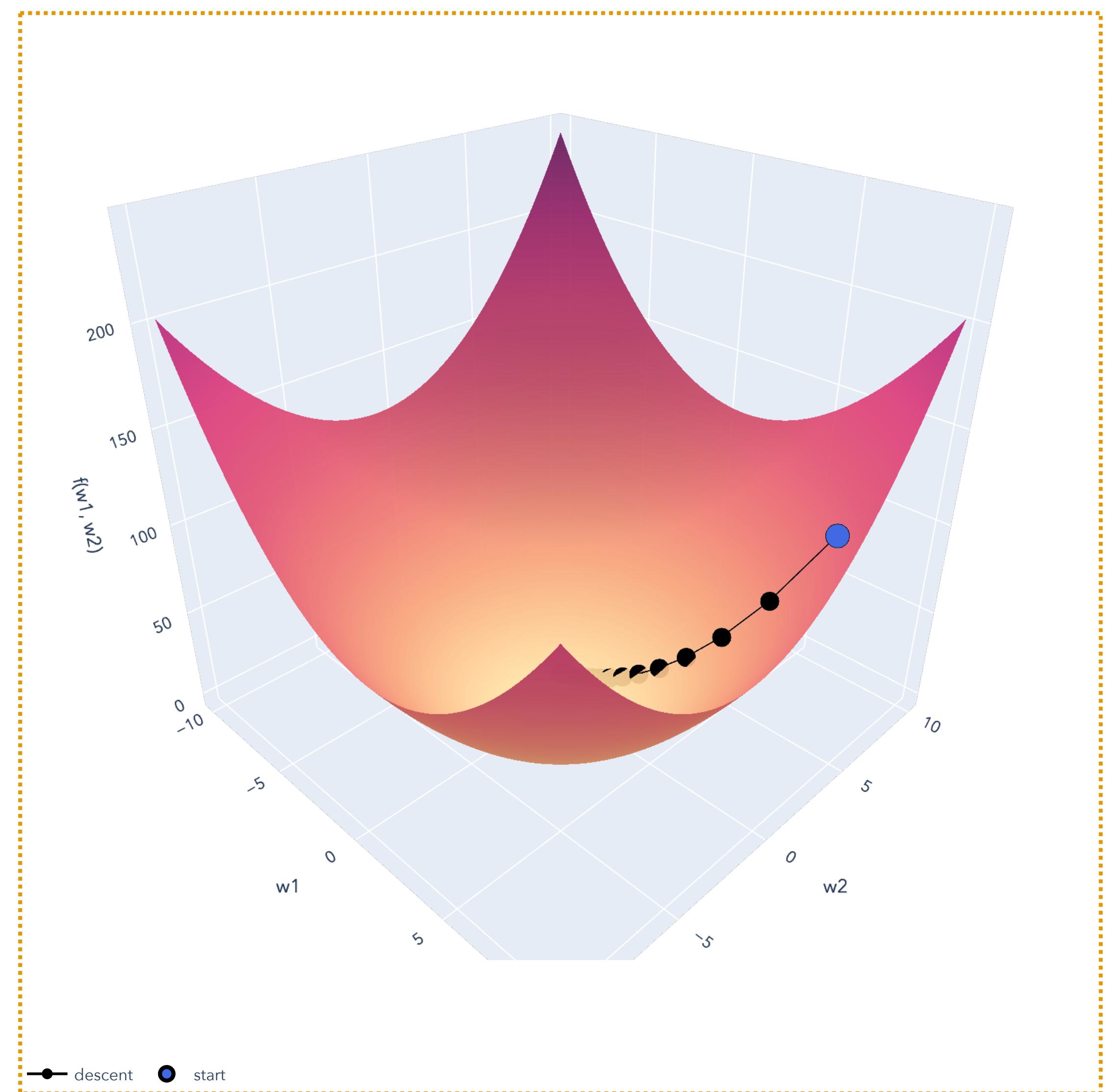
## Big Picture: Gradient Descent

Using [norm](#) to rewrite the sum of squared residual errors,

$$f(\mathbf{w}) = \sum_{i=1}^n (\mathbf{w}^\top \mathbf{x}_i - y_i)^2$$

we got a function measuring how “badly” each  $\mathbf{w}$  does:

$$f(\mathbf{w}) = \|\mathbf{X}\mathbf{w} - \mathbf{y}\|^2.$$



# Week 1.2

## Bases, subspaces, and orthogonality

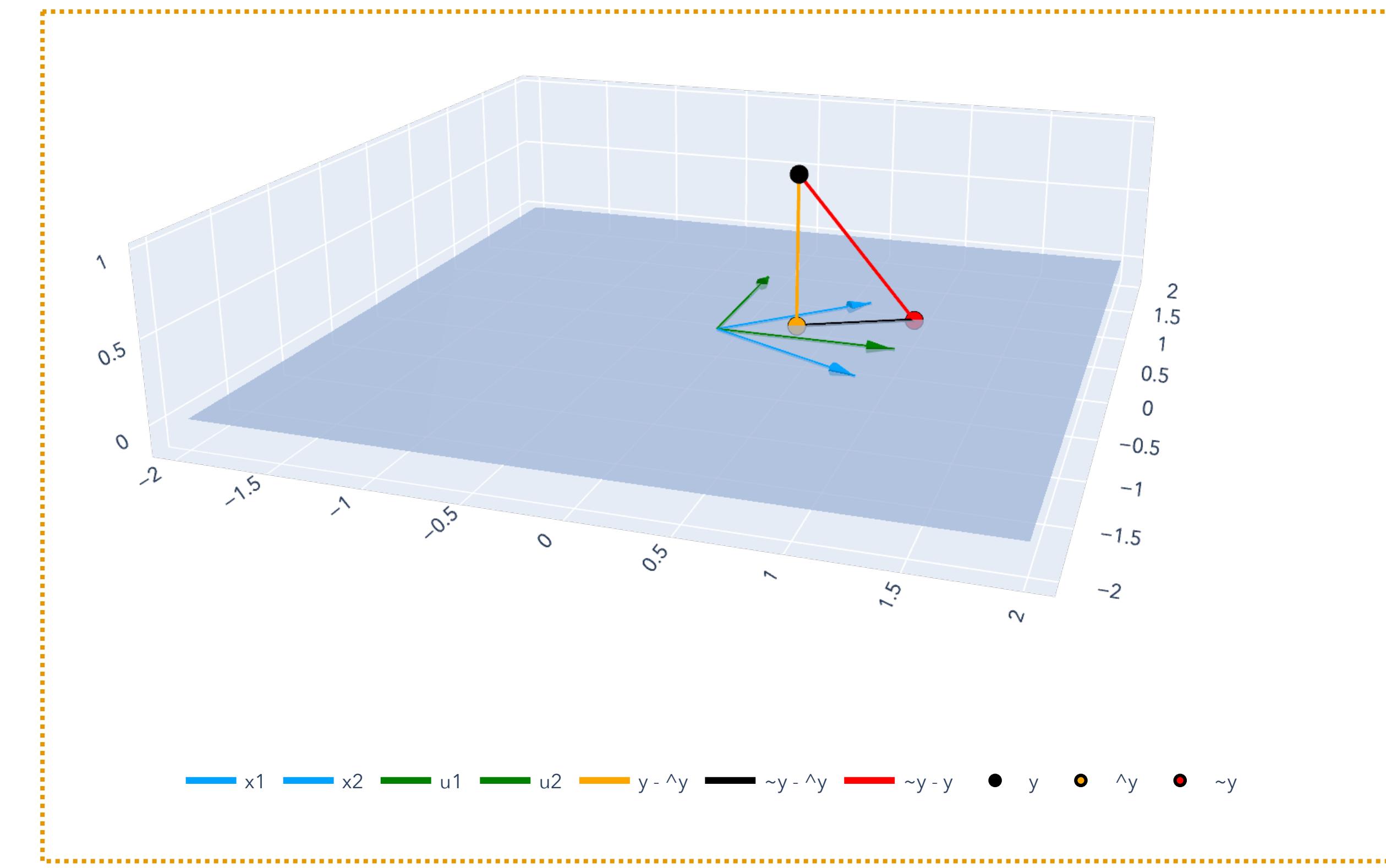
# Bases, subspaces, and orthogonality

## Big Picture: Least Squares

We formally defined subspace, a basis, the columnspace, and orthogonal basis. This filled in the gaps to get Theorem (invertibility of  $\mathbf{X}^T \mathbf{X}$ ) and Theorem (Pythagorean Theorem).

Using our new notion of orthogonality, we could simplify the OLS solution if we had an ONB.

**Theorem (OLS solution with ONB).** If  $n \geq d$  and  $\text{rank}(\mathbf{X}) = d$  and  $\mathbf{U} \in \mathbb{R}^{d \times d}$  an ONB:  $\hat{\mathbf{w}} = \mathbf{U}^T \mathbf{y}$ .



# Bases, subspaces, and orthogonality

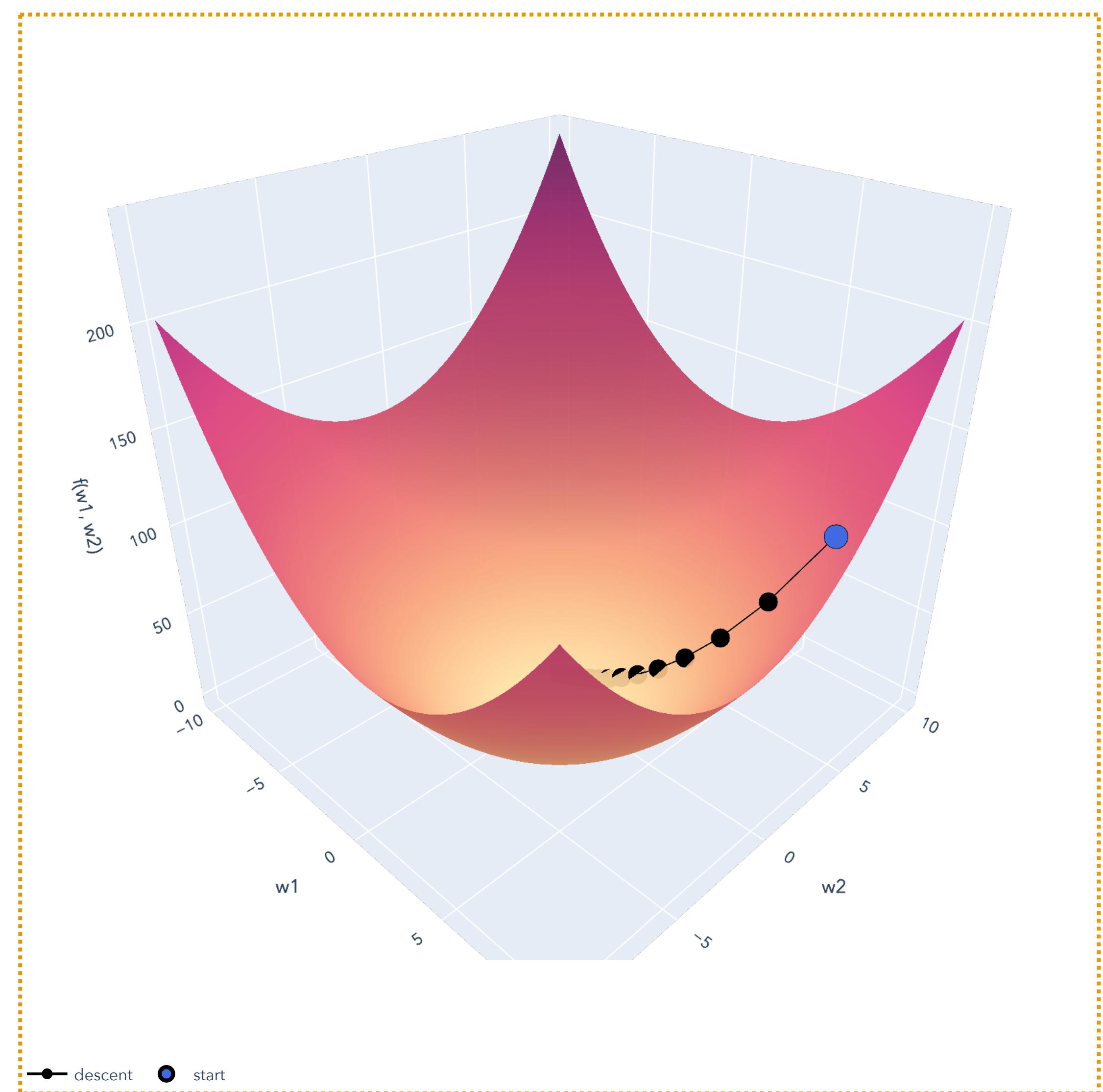
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# Week 2.1

## Singular Value Decomposition

# Singular Value Decomposition

## Big Picture: Least Squares

We defined orthogonal complements and projection matrices to solve the best-fitting 1D subspace problem, leading to SVD:

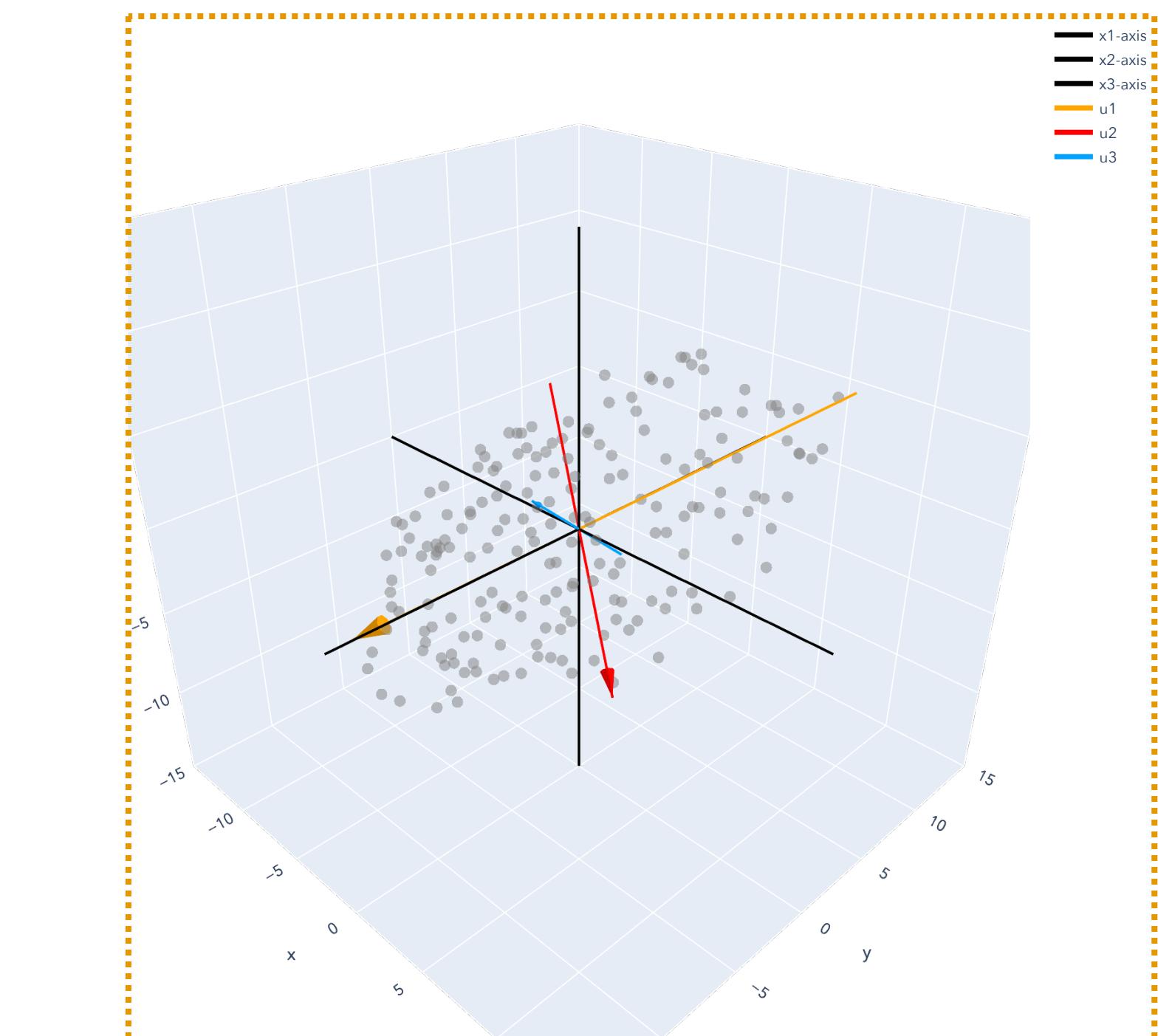
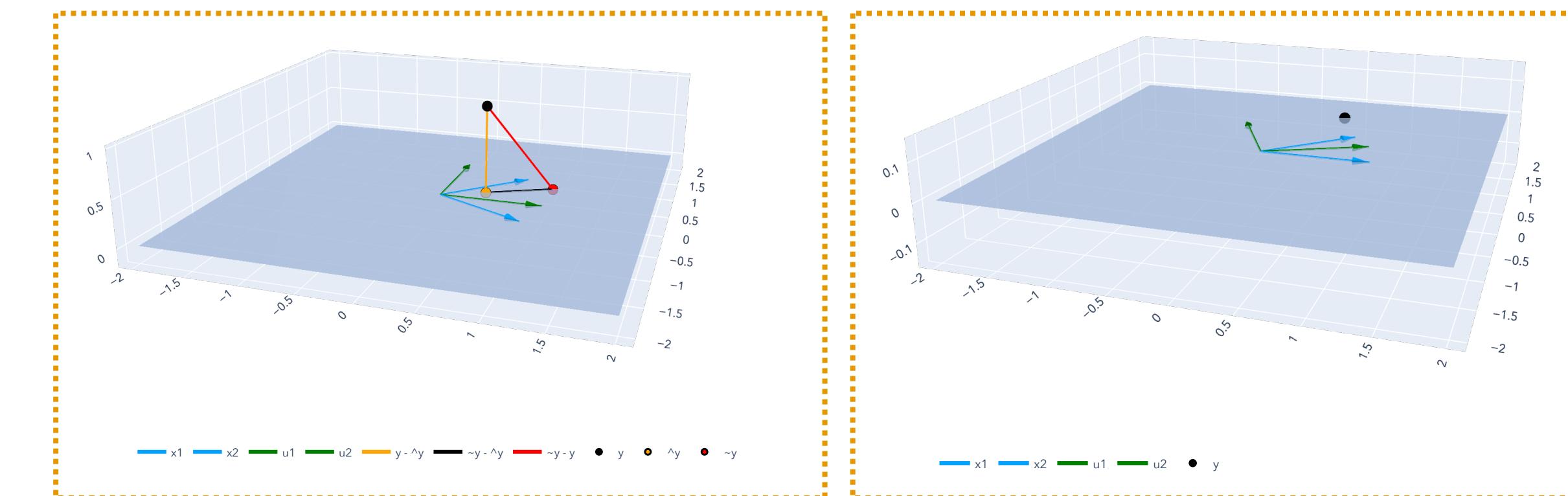
$$\mathbf{X} = \mathbf{U}\Sigma\mathbf{V}^\top$$

The SVD defined the pseudoinverse which unified OLS:

**Theorem (OLS solution with pseudoinverse).** Let  $\mathbf{X} \in \mathbb{R}^{n \times d}$  have pseudoinverse  $\mathbf{X}^+ \in \mathbb{R}^{d \times n}$ . Then:  $\hat{\mathbf{w}} = \mathbf{X}^+\mathbf{y}$ .

If  $n \geq d$ , then  $\hat{\mathbf{w}}$  minimizes  $\|\mathbf{X}\mathbf{w} - \mathbf{y}\|^2$ .

If  $d > n$ , then  $\hat{\mathbf{w}}$  is the exact solution  $\mathbf{X}\hat{\mathbf{w}} = \mathbf{y}$  with min. norm.



# Singular Value Decomposition

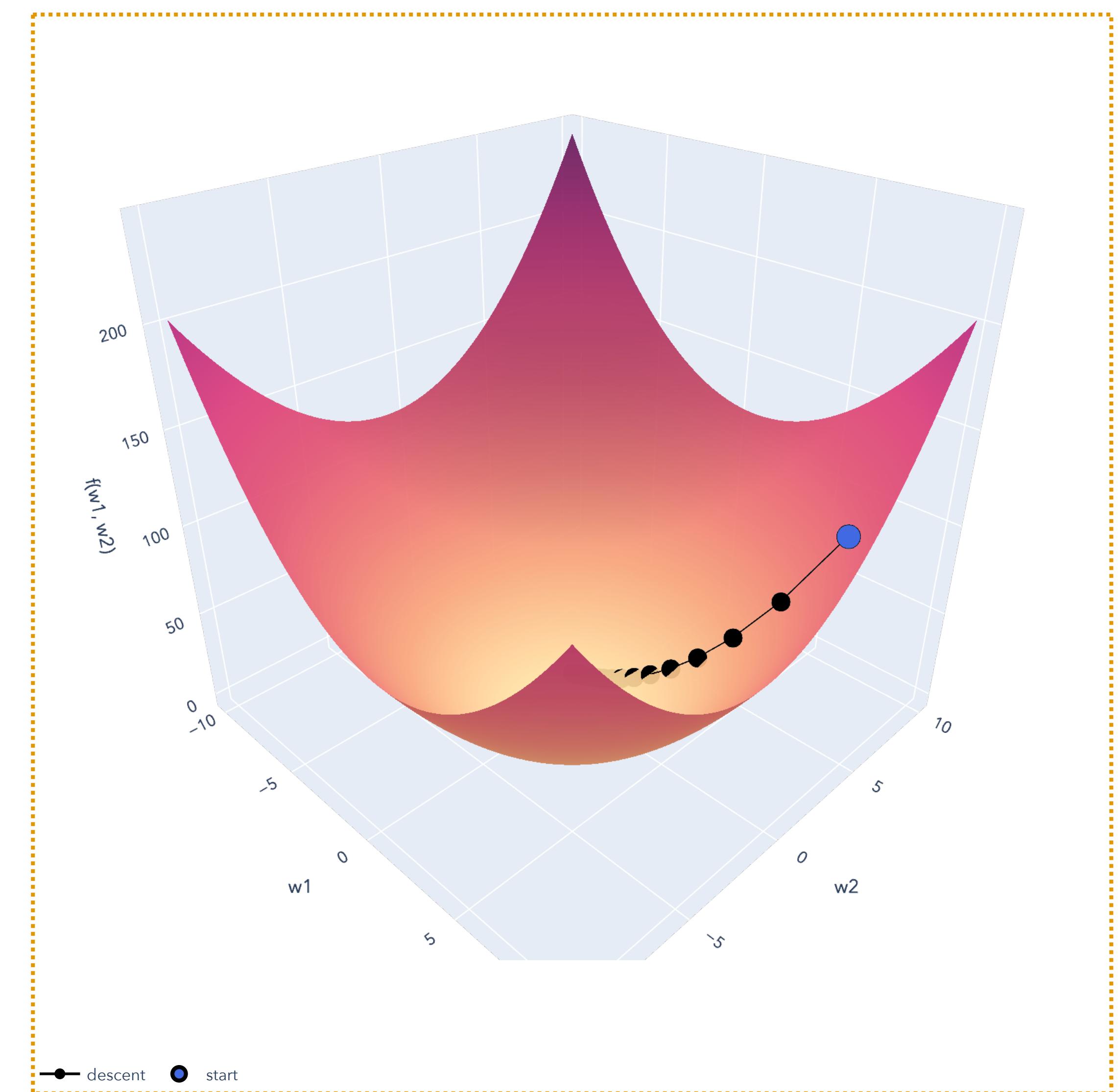
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we got a function measuring how “badly” each  $\mathbf{w}$  does:

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# Week 2.2

## Eigendecomposition and PSD Matrices

# Eigendecomposition and PSD Matrices

## Big Picture: Least Squares

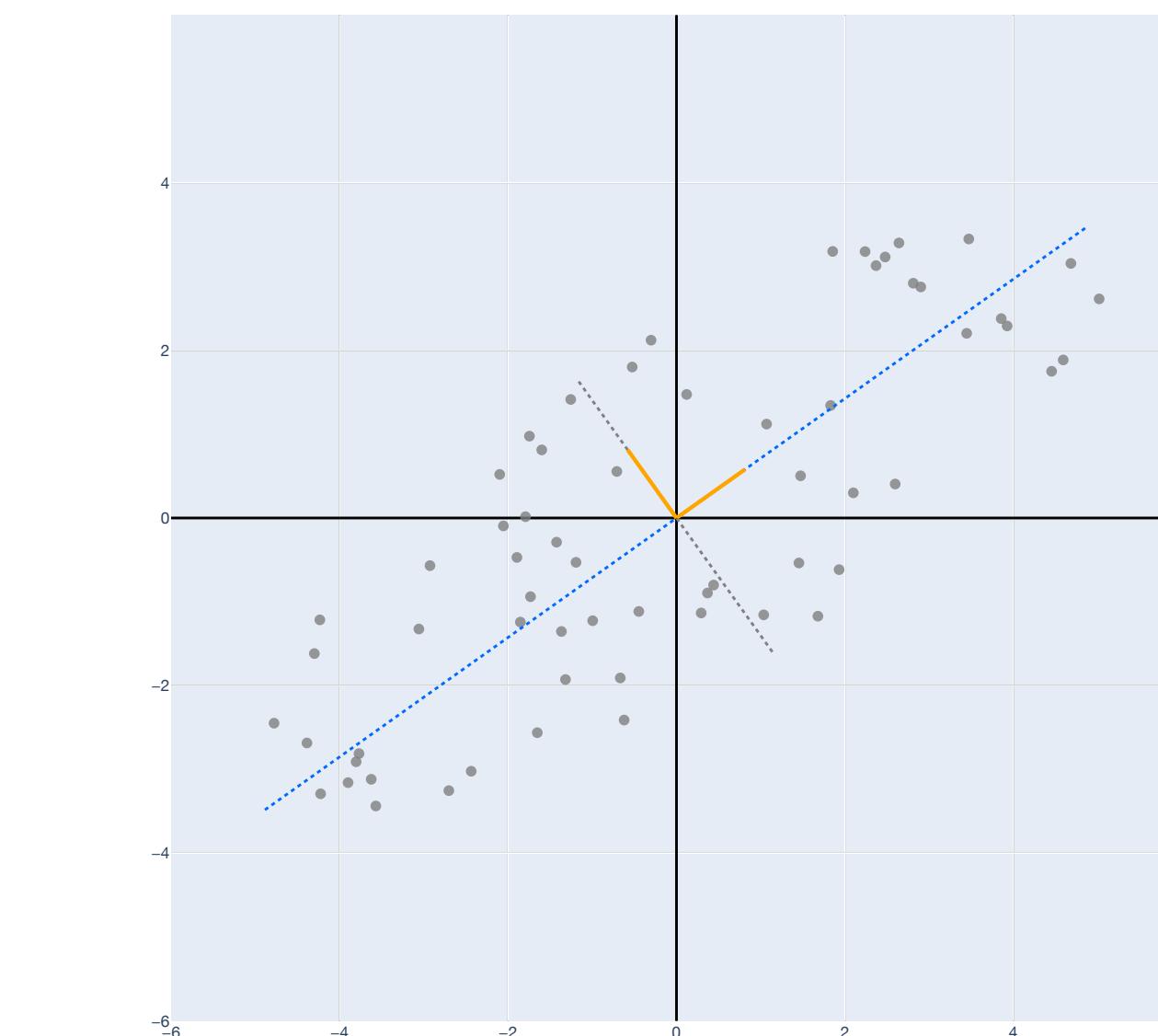
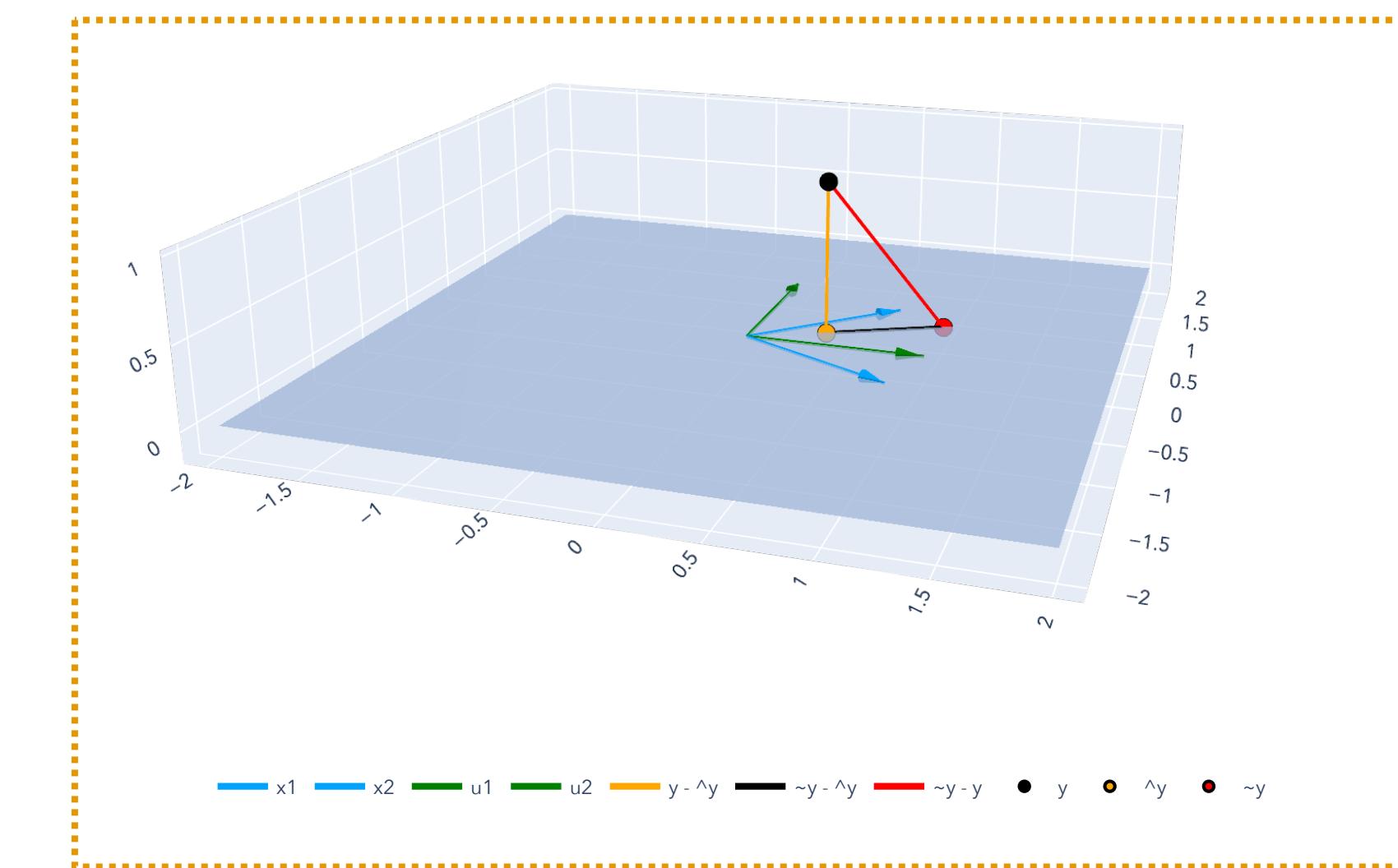
We defined eigenvectors and eigenvalues of square matrices. When a square matrix is diagonalizable:

$$\mathbf{X} = \mathbf{V}\Lambda\mathbf{V}^T$$

The spectral theorem tells us that symmetric matrices are diagonalizable.

One example of a symmetric matrix is  $\mathbf{X}^T\mathbf{X}$ , so we did a rudimentary eigenvector/eigenvalue analysis of  $(\mathbf{X}^T\mathbf{X})^{-1}\mathbf{X}^T\mathbf{y}$  in the error model:

$$\mathbf{y} = \mathbf{X}\mathbf{w}^* + \epsilon.$$



# Eigendecomposition and PSD Matrices

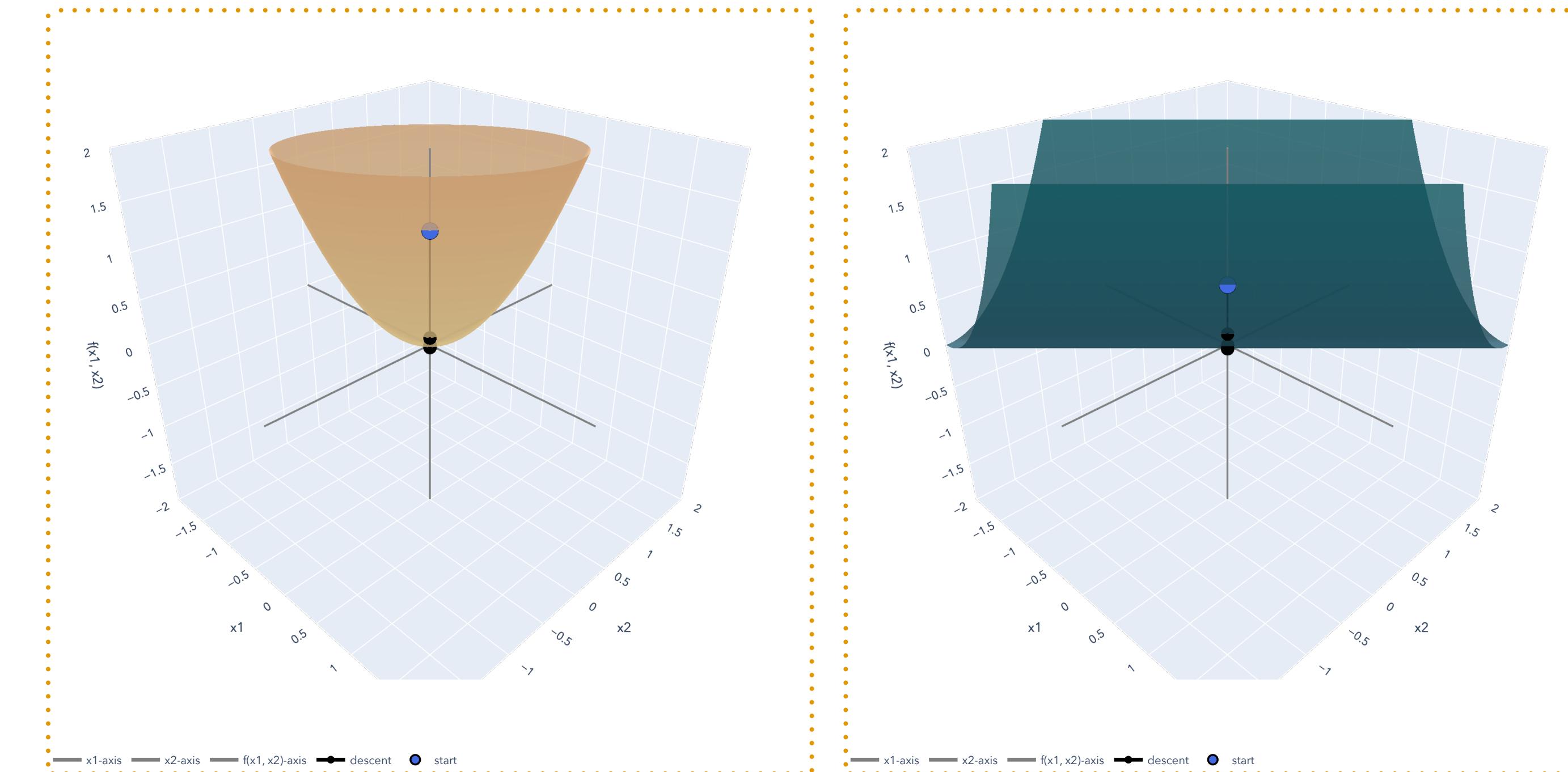
## Big Picture: Gradient Descent

Defined an important class of square, symmetric matrices, positive semidefinite (PSD) matrices.

PSD matrices are always associated with functions called quadratic forms

$$f(\mathbf{x}) := \mathbf{x}^T \mathbf{A} \mathbf{x},$$

which look “bowl” or “envelope” shaped.



# Week 3.1

## Differentiation and vector calculus

# Differentiation and vector calculus

## Big Picture: Least Squares

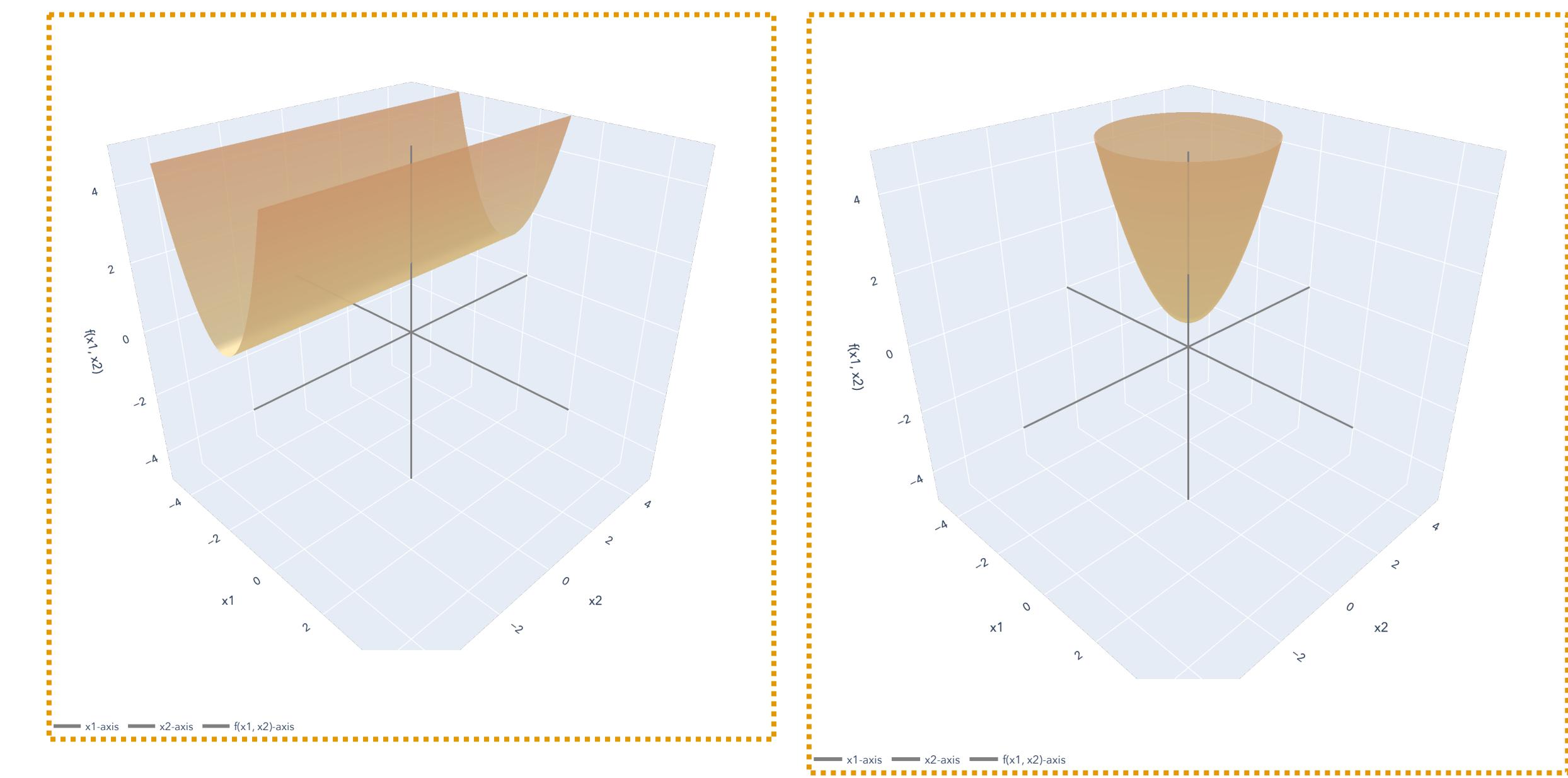
The directional, partial, and total derivatives are summarized with the gradient and Jacobian.

Using analogy to single variable calculus optimization, we treated

$$f(\mathbf{w}) = \|\mathbf{X}\mathbf{w} - \mathbf{y}\|^2$$

as a function to optimize and proved the same theorem, from a calculus/optimization perspective.

**Theorem (OLS solution).** If  $n \geq d$  and  $\text{rank}(\mathbf{X}) = d$ ,  
then:  $\hat{\mathbf{w}} = (\mathbf{X}^T \mathbf{X})^{-1} \mathbf{X}^T \mathbf{y}$ .

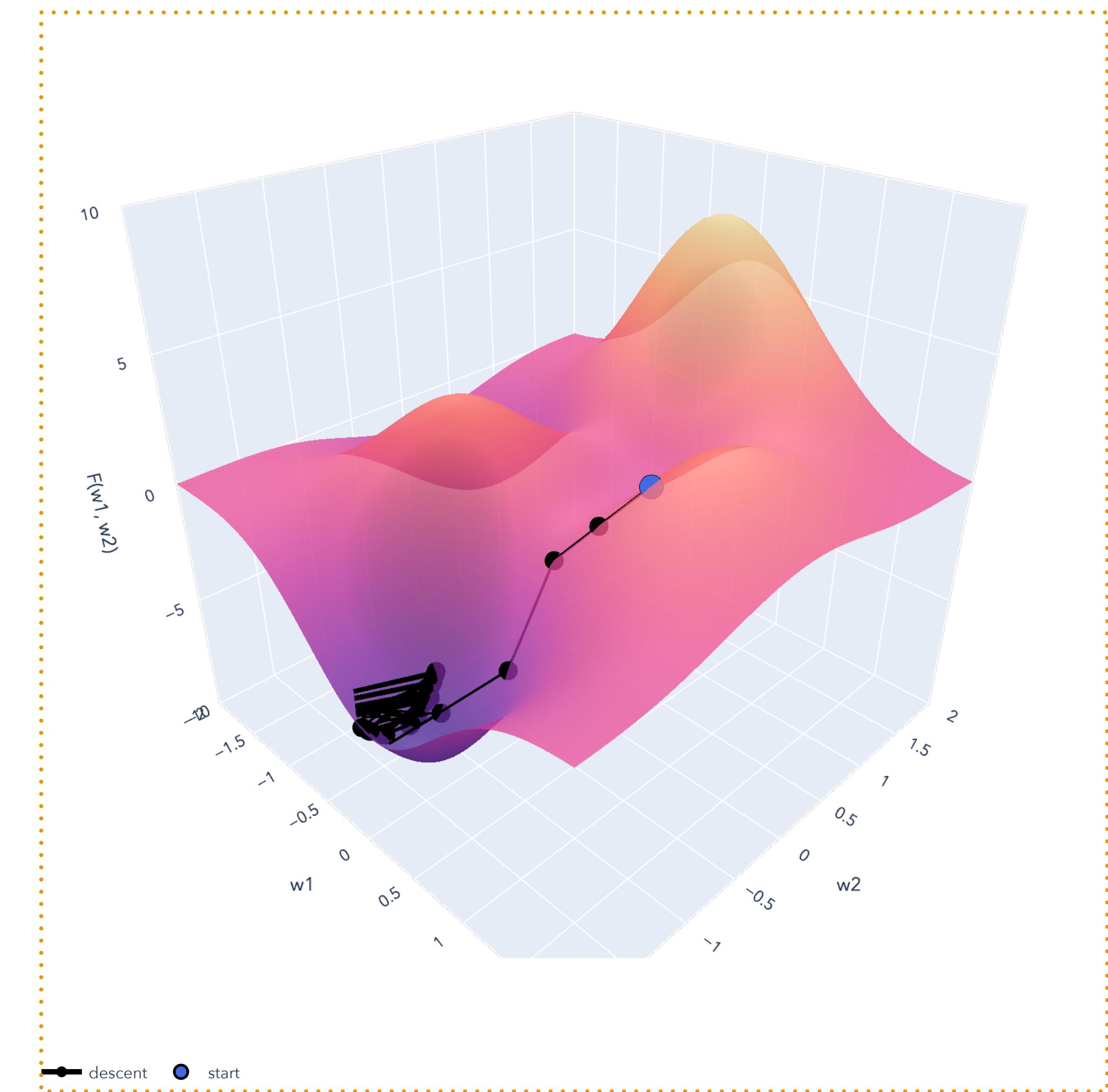


# Differentiation and vector calculus

## Big Picture: Gradient Descent

The gradient points in the direction of steepest ascent. This lets us write out the algorithm for gradient descent:

$$\mathbf{w}_t \leftarrow \mathbf{w}_{t-1} - \eta \nabla f(\mathbf{w}_{t-1}).$$



# Week 3.2

## Linearization and Taylor series

# Linearization and Taylor series

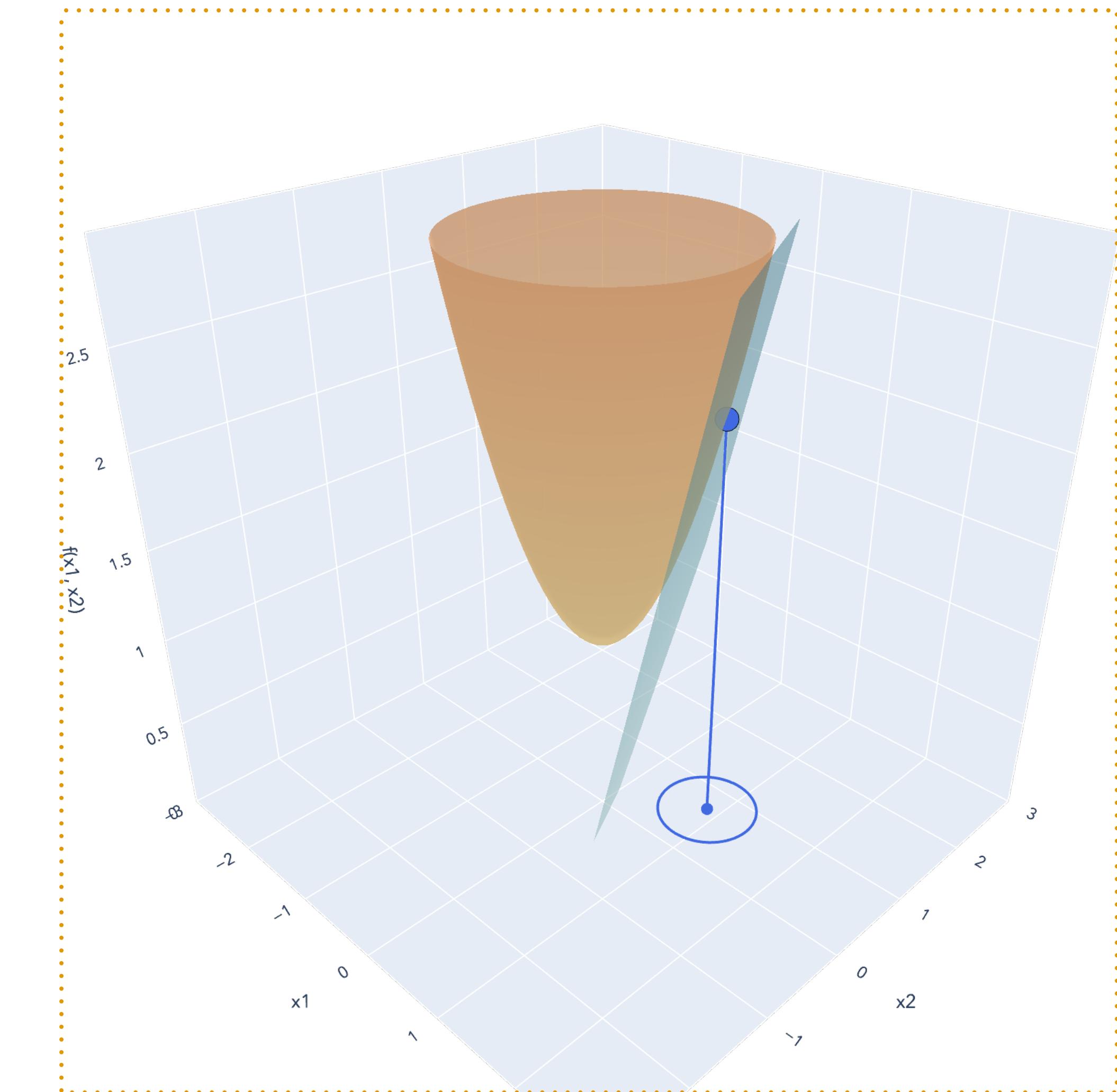
## Big Picture: Least Squares

We discussed [linearization](#), a main motivation for the techniques of multivariable calculus:

$$f(\mathbf{x}) \approx f(\mathbf{x}_0) + \nabla f(\mathbf{x}_0)^T (\mathbf{x} - \mathbf{x}_0)$$

This is a “part” of the [Taylor series](#) of a function.  
We quantified the approximation error of a Taylor series through [Taylor’s Theorem](#).

The error term in the first-order Taylor expansion was a function of the [Hessian](#), which is always a symmetric matrix for  $\mathcal{C}^2$  functions.



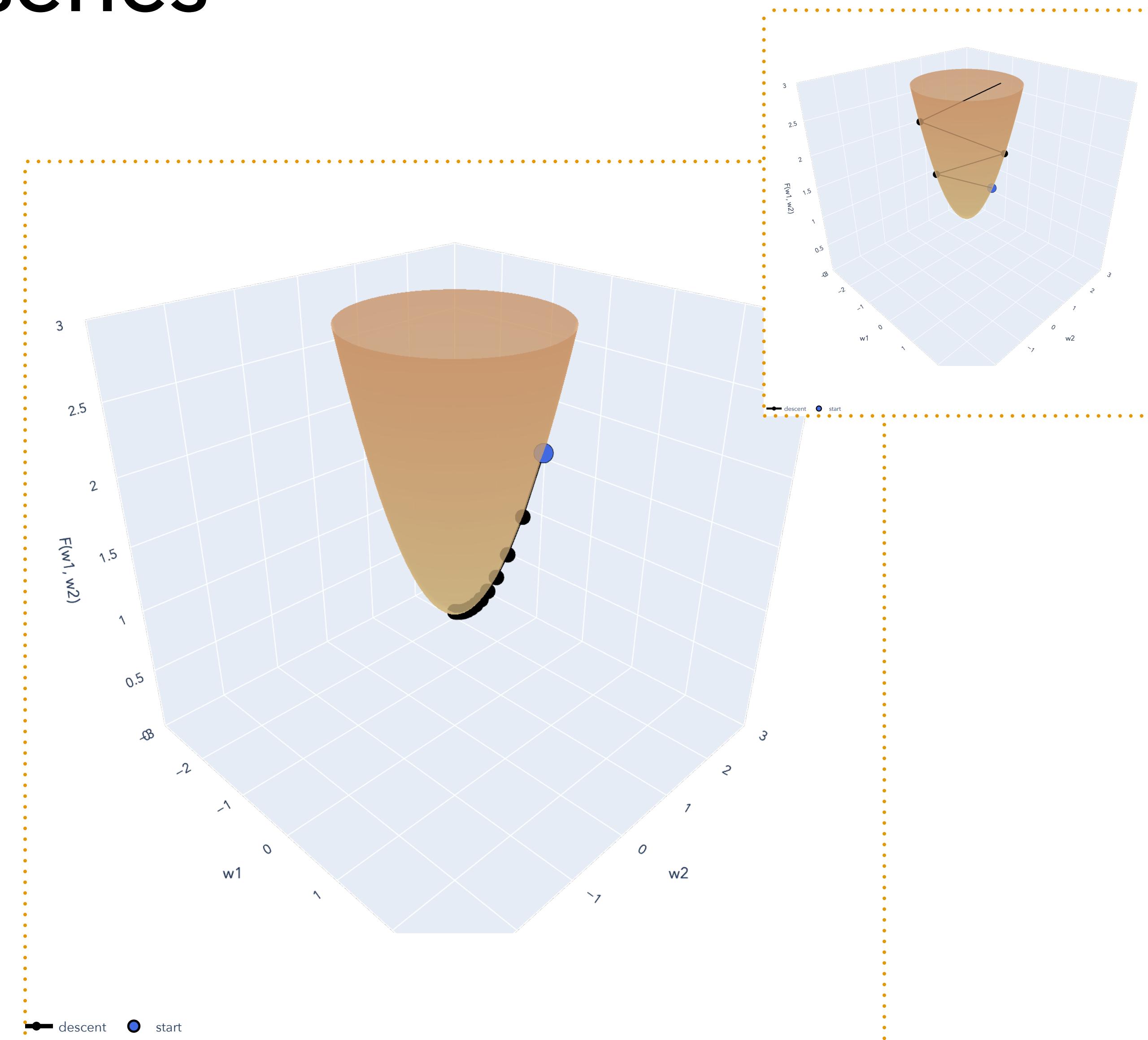
# Linearization and Taylor series

## Big Picture: Gradient Descent

Taylor's Theorem and smoothness of the Hessian allowed us to analyze the first-order Taylor approximation to get our first GD theorem:

**Theorem (Descent Lemma).** If  $f \in \mathcal{C}^2$  and is  $\beta$ -smooth, then with  $\eta = 1/\beta$ , for any  $\mathbf{w} \in \mathbb{R}^d$ ,

$$f(\mathbf{w} - \eta \nabla f(\mathbf{w})) \leq f(\mathbf{w}) - \frac{1}{2\beta} \|\nabla f(\mathbf{w})\|^2.$$



# Week 4.1

## Optimization and the Lagrangian

# Optimization and the Lagrangian

## Big Picture: Least Squares

minimize  $f(\mathbf{x})$   
 $\mathbf{x} \in \mathbb{R}^d$

subject to  $\mathbf{x} \in \mathcal{C}$

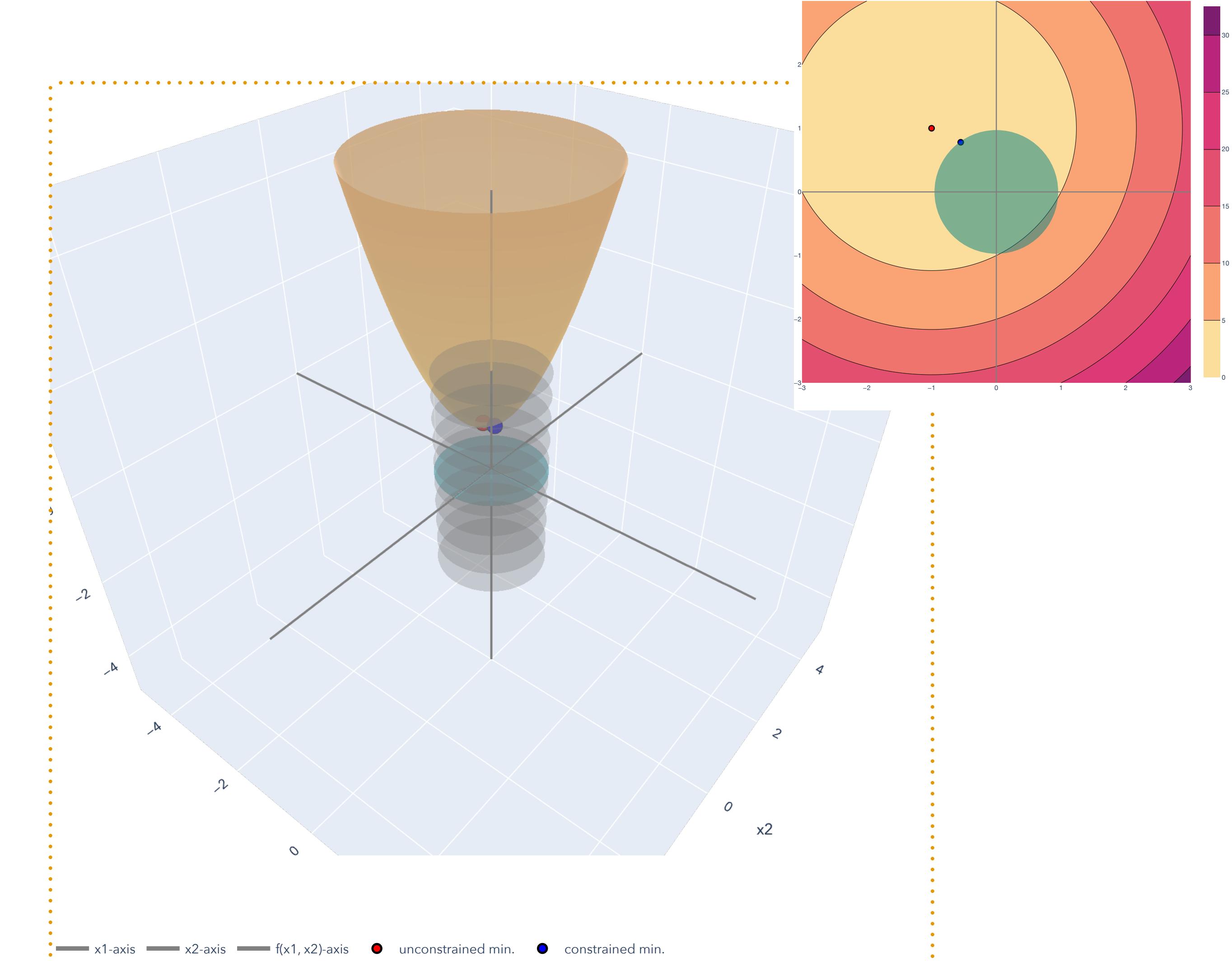
Gave the necessary conditions for unconstrained local minima, filled in gaps in OLS proof.

Defined the Lagrangian  $L(\mathbf{x}, \lambda)$ , which helped us solve constrained optimization problems by “unconstraining.”

Two constrained problems related to OLS:

1. Least norm solution.  $\hat{\mathbf{w}} = \mathbf{X}^+ \mathbf{y}$ .

2. Ridge regression.  $\hat{\mathbf{w}} = (\mathbf{X}^\top \mathbf{X} + \gamma \mathbf{I})^{-1} \mathbf{X}^\top \mathbf{y}$



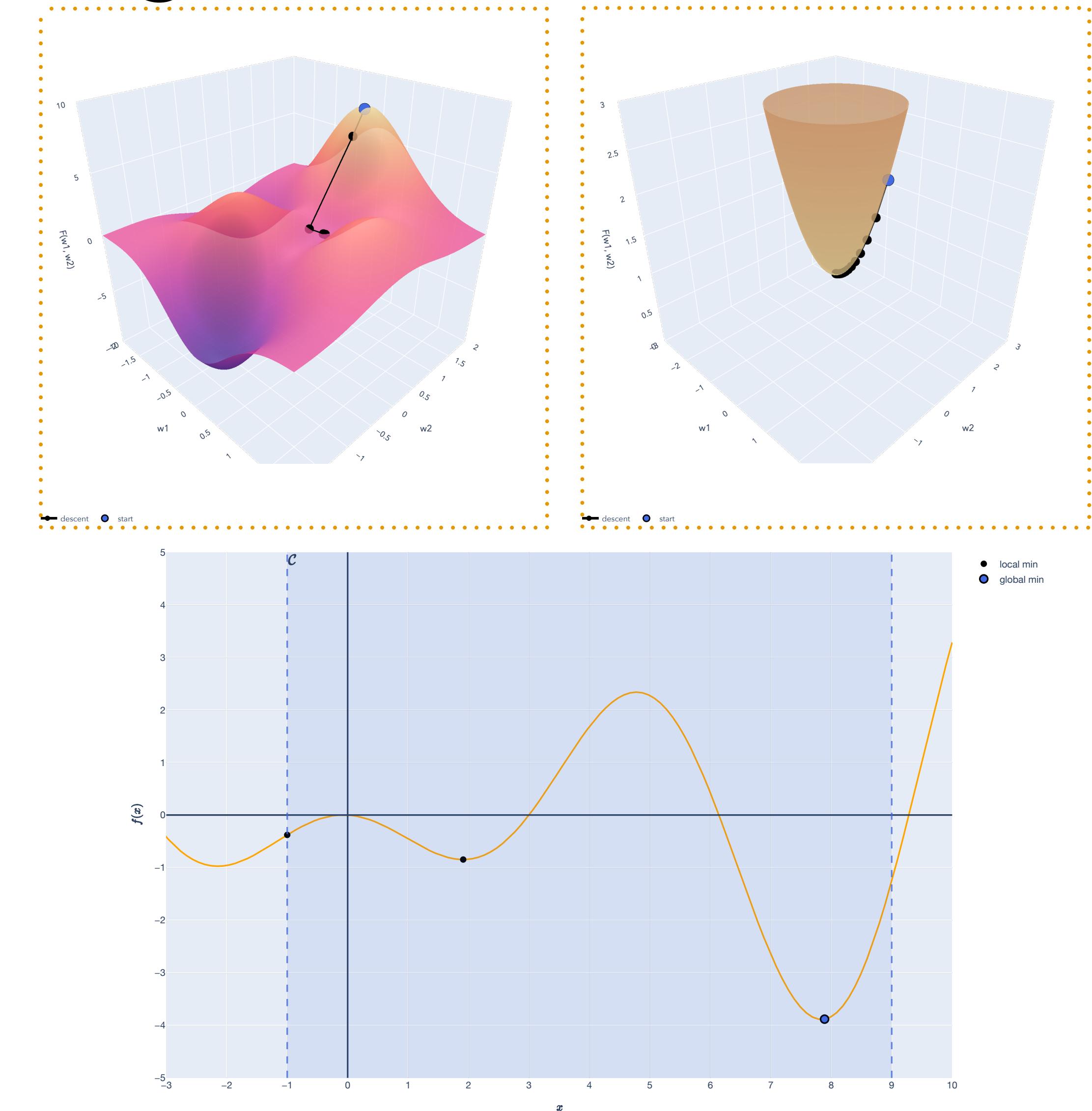
# Optimization and the Lagrangian

## Big Picture: Gradient Descent

Classified the types of minima we can hope for in an optimization problem: unconstrained local minima, constrained local minima, and global minima.

We want global minima but GD and the descent lemma only says something about getting to the local minima.

$$f(\mathbf{w} - \eta \nabla f(\mathbf{w})) \leq f(\mathbf{w}) - \frac{1}{2\beta} \|\nabla f(\mathbf{w})\|^2$$



# Week 4.2

## Basics of convex optimization

# Basics of convex optimization

## Big Picture: Least Squares

Convexity of functions and sets. Convex functions satisfy:

$$f(\alpha \mathbf{x} + (1 - \alpha) \mathbf{y}) \leq \alpha f(\mathbf{x}) + (1 - \alpha) f(\mathbf{y}).$$

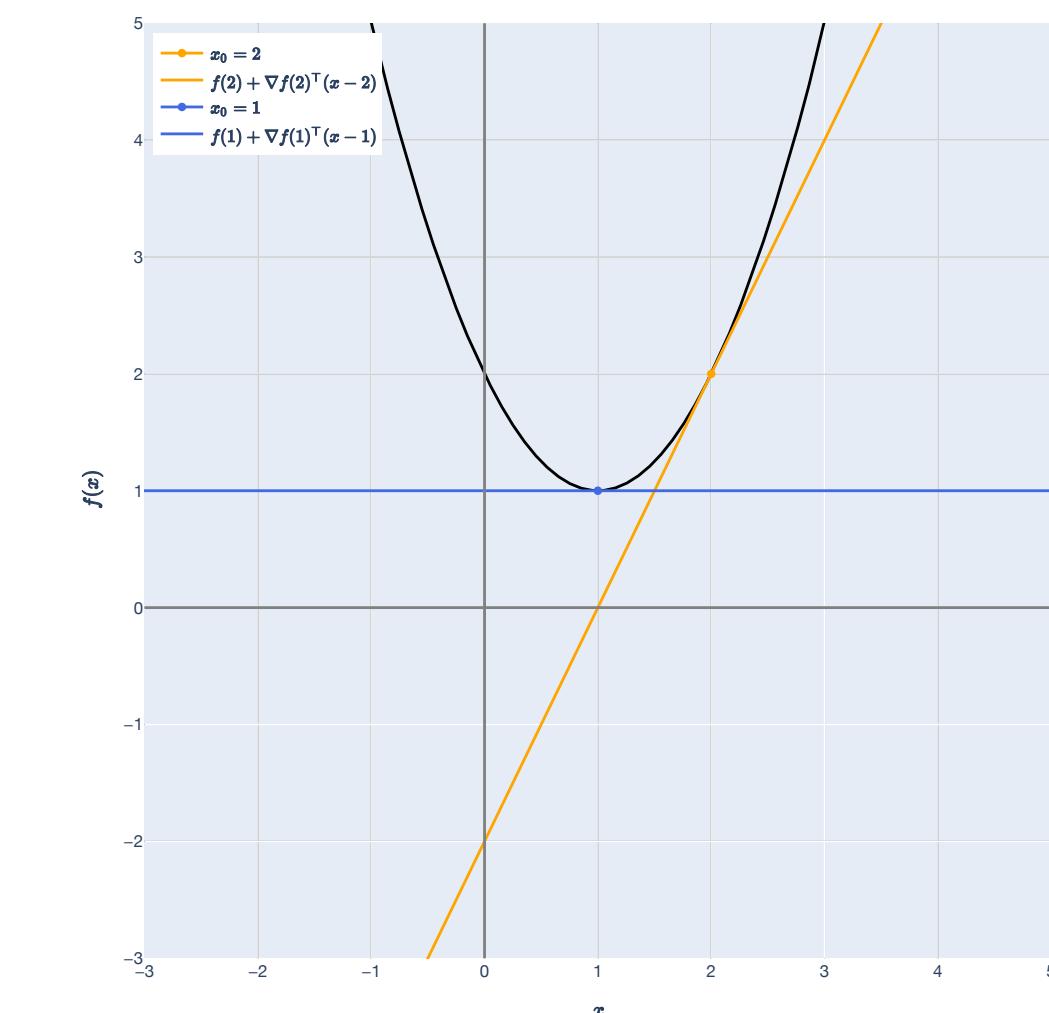
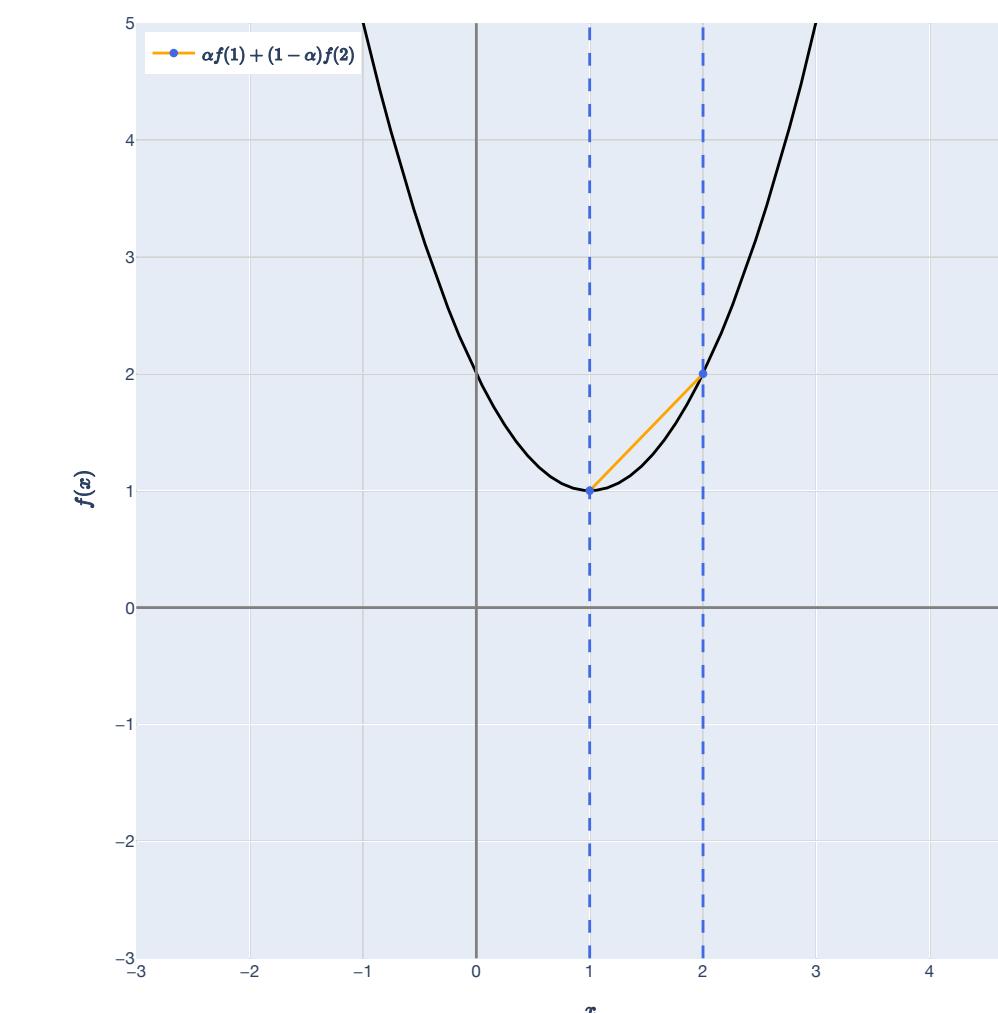
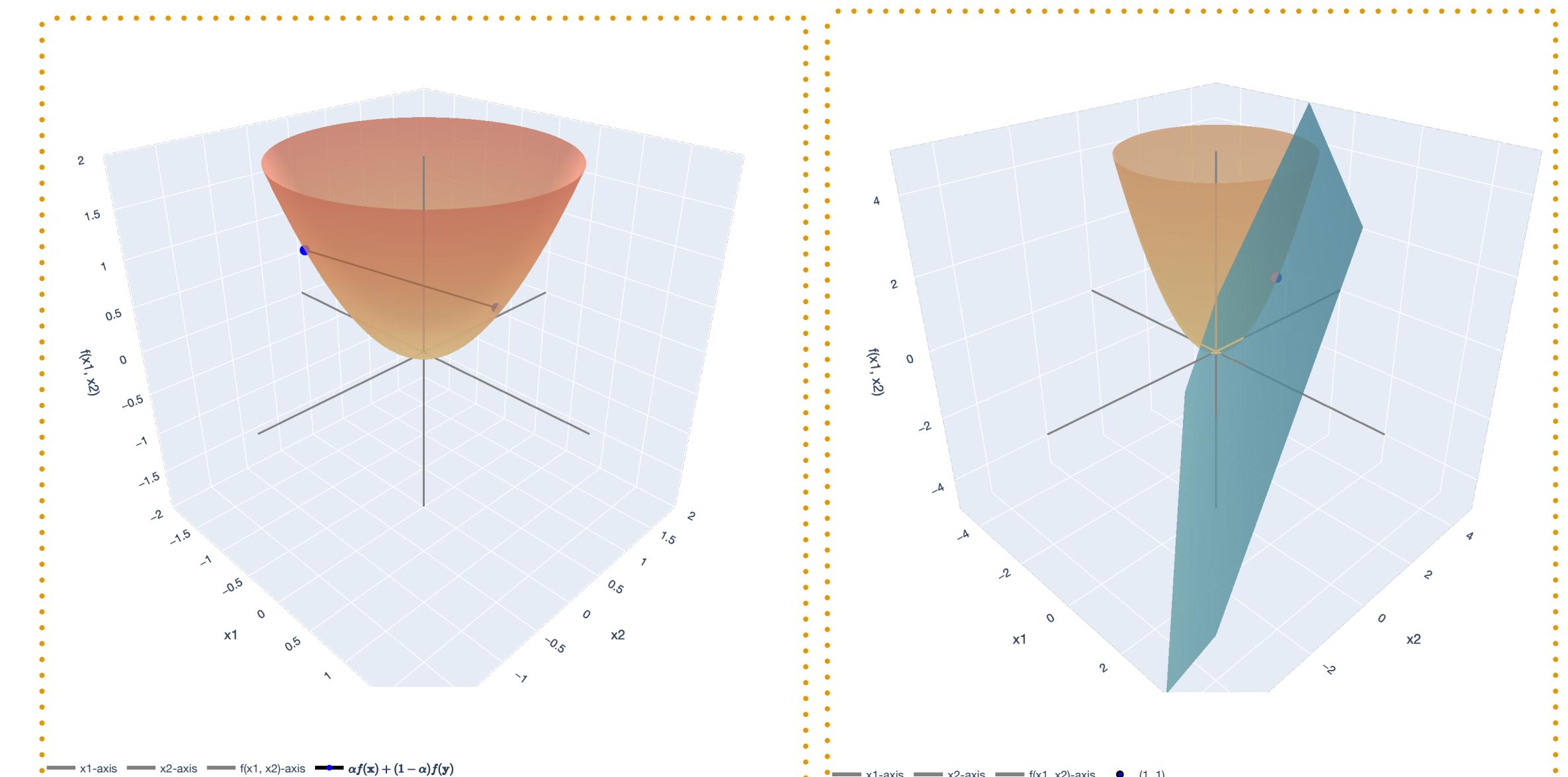
$$f(\mathbf{y}) \geq f(\mathbf{x}) + \nabla_{\mathbf{x}} f(\mathbf{x})^T (\mathbf{y} - \mathbf{x}).$$

$\nabla^2 f(\mathbf{x})$  is positive semidefinite.

The key property we proved is that for **convex functions, all local minima are global minima**.

We verified that the OLS objective is convex:

$$f(\mathbf{w}) = \|\mathbf{X}\mathbf{w} - \mathbf{y}\|^2 \text{ is convex.}$$



# Basics of convex optimization

## Big Picture: Gradient Descent

Assured that for **convex** functions, all local minima are global minima, we proved *global* convergence for GD:

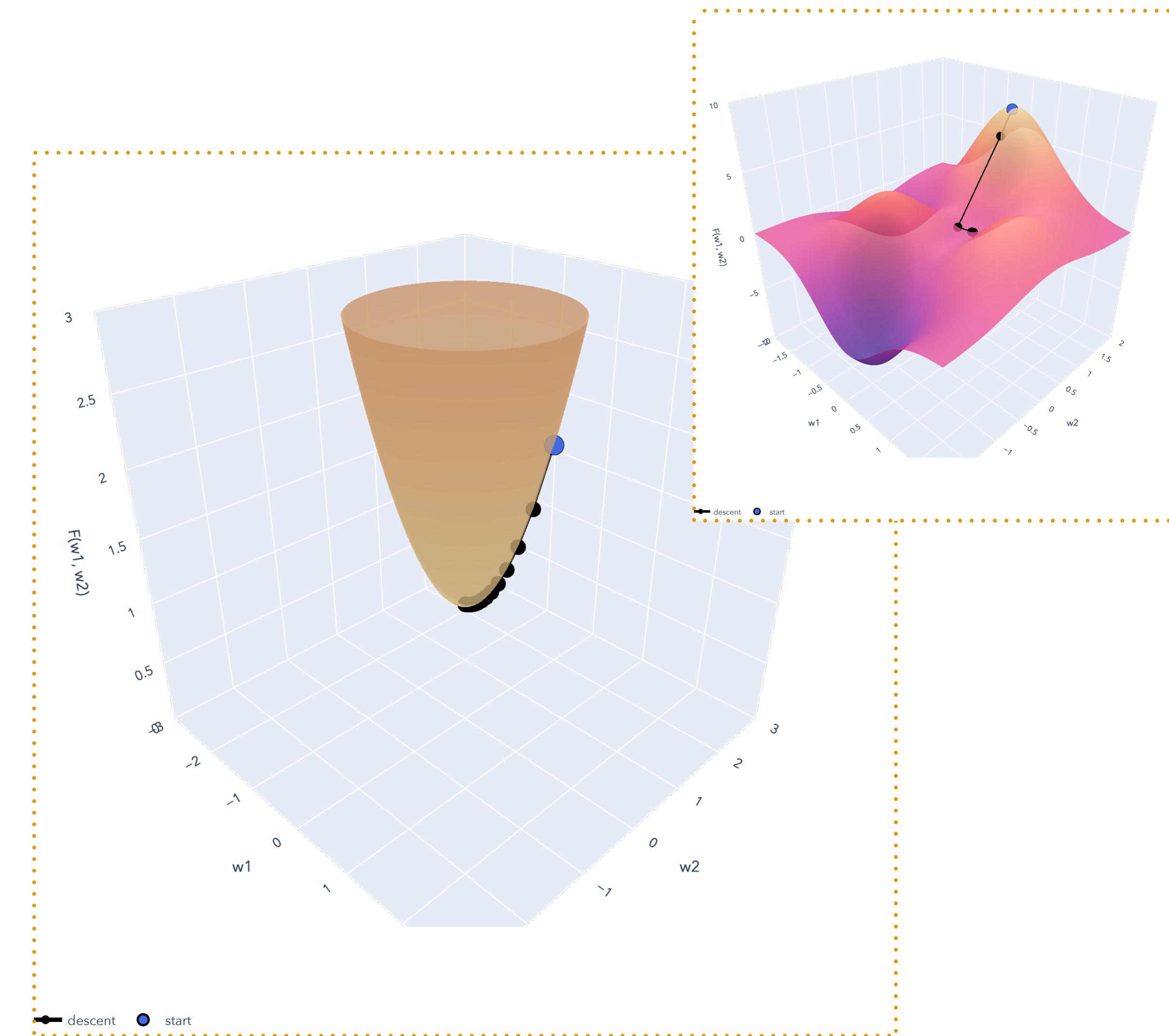
**Theorem (Convergence of GD for smooth, convex functions).**

$$f(\mathbf{x}_T) - f(\mathbf{x}^*) \leq \frac{\beta}{2T} (\|\mathbf{x}_0 - \mathbf{x}^*\|^2 - \|\mathbf{x}_T - \mathbf{x}^*\|^2),$$

after  $T$  iterations of our algorithm.

As a corollary, we were able to unite the two stories of our course and apply GD to OLS to get:

$$\|\mathbf{X}\mathbf{w}_T - \mathbf{y}\|^2 - \|\mathbf{X}\mathbf{w}^* - \mathbf{y}\|^2 \leq \frac{\beta}{2T} (\|\mathbf{w}_0 - \mathbf{w}^*\|^2 - \|\mathbf{w}_T - \mathbf{w}^*\|^2).$$



# Week 5.1

## Probability Theory, Models, and Data

# Probability Theory, Models, and Data

## Big Picture: Least Squares

Defined probability spaces and random variables.

Random variables come with a CDF and a PMF/PDF.

Most important stats: expectation and variance.

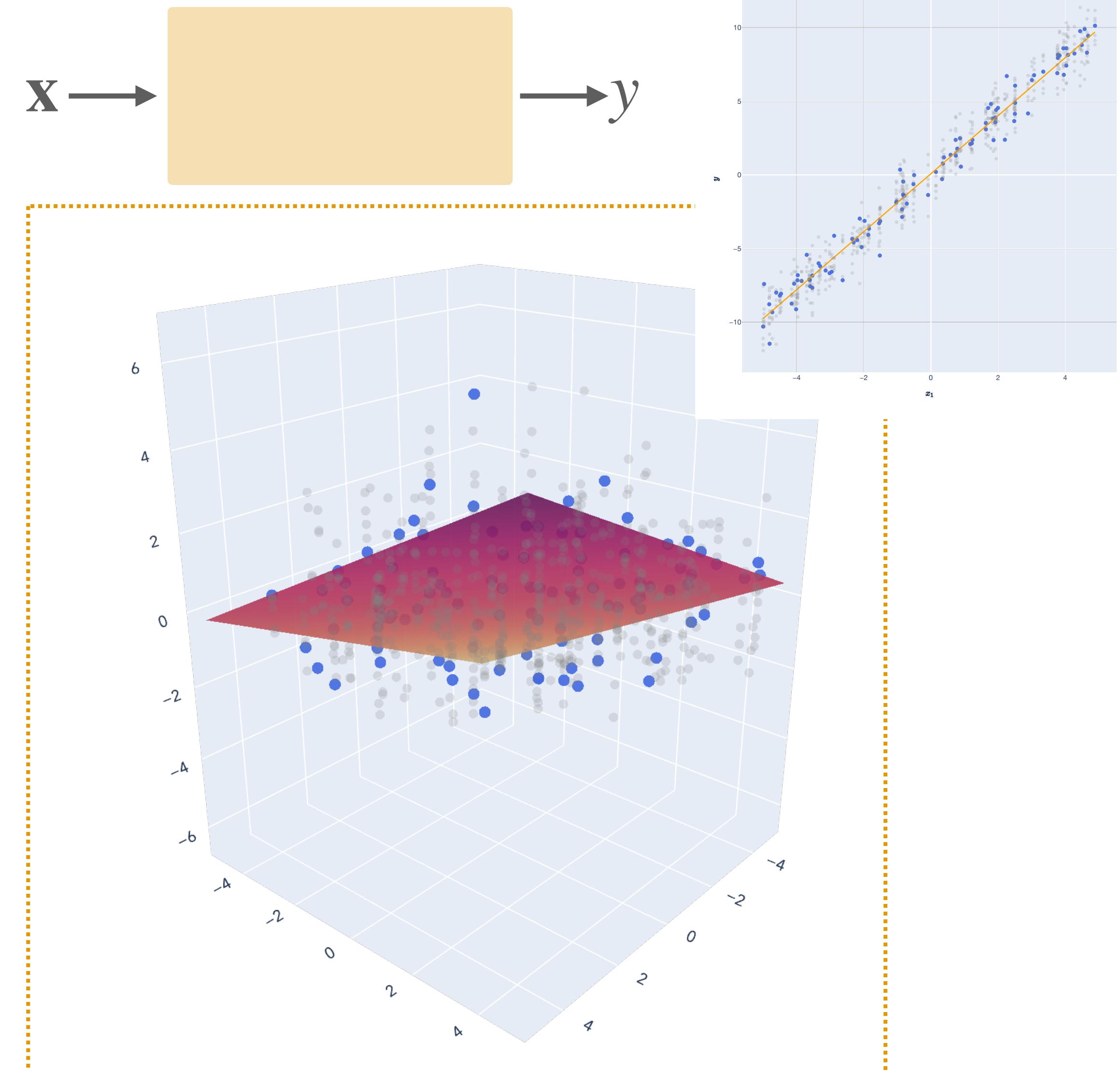
Random vector variances are given in a covariance matrix. This framework allowed us to define the random error model:

$\mathbf{y} = \mathbf{X}\mathbf{w}^* + \boldsymbol{\epsilon}$ , where  $\mathbb{E}[\boldsymbol{\epsilon}] = 0$  and  $\epsilon_i$  i.i.d. and indep. of  $\mathbf{X}$ .

Now,  $\hat{\mathbf{w}} = (\mathbf{X}^\top \mathbf{X})^{-1} \mathbf{X}^\top \mathbf{y}$  is a random variable with

Expectation:  $\mathbb{E}[\hat{\mathbf{w}}] = \mathbf{w}^*$ .

Variance:  $\text{Var}[\hat{\mathbf{w}}] = \sigma^2 \mathbb{E}[(\mathbf{X}^\top \mathbf{X})^{-1}]$ .



# Probability Theory, Models, and Data

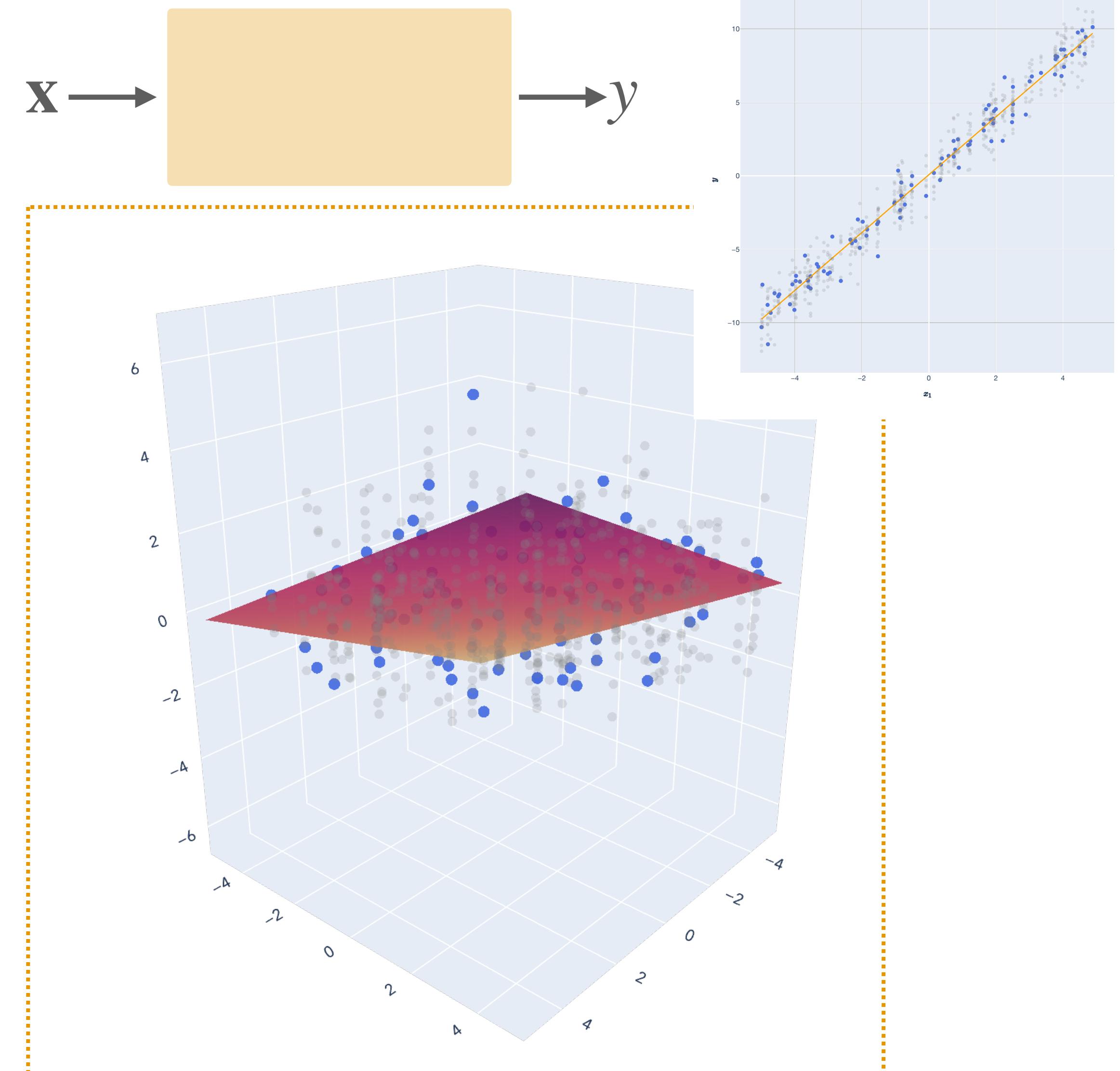
## Big Picture: Least Squares

Proof of OLS expectation and variance relied heavily on conditioning.

The conditional expectation of a random variable can be thought of as a “best guess” at a random variable given the information of an event or *another random variable*.

$$\mathbb{E}[X | A], \text{ for } A \subseteq \Omega.$$

$$\mathbb{E}[X | Y], \text{ for } Y : \Omega \rightarrow \mathbb{R}.$$



# Week 5.2

## Law of large numbers and statistical estimators

# Law of large numbers and statistical estimators

## Big Picture: Least Squares

The sample average of i.i.d. random variables:

$$\bar{X}_n := \frac{1}{n} \sum_{i=1}^n X_i.$$

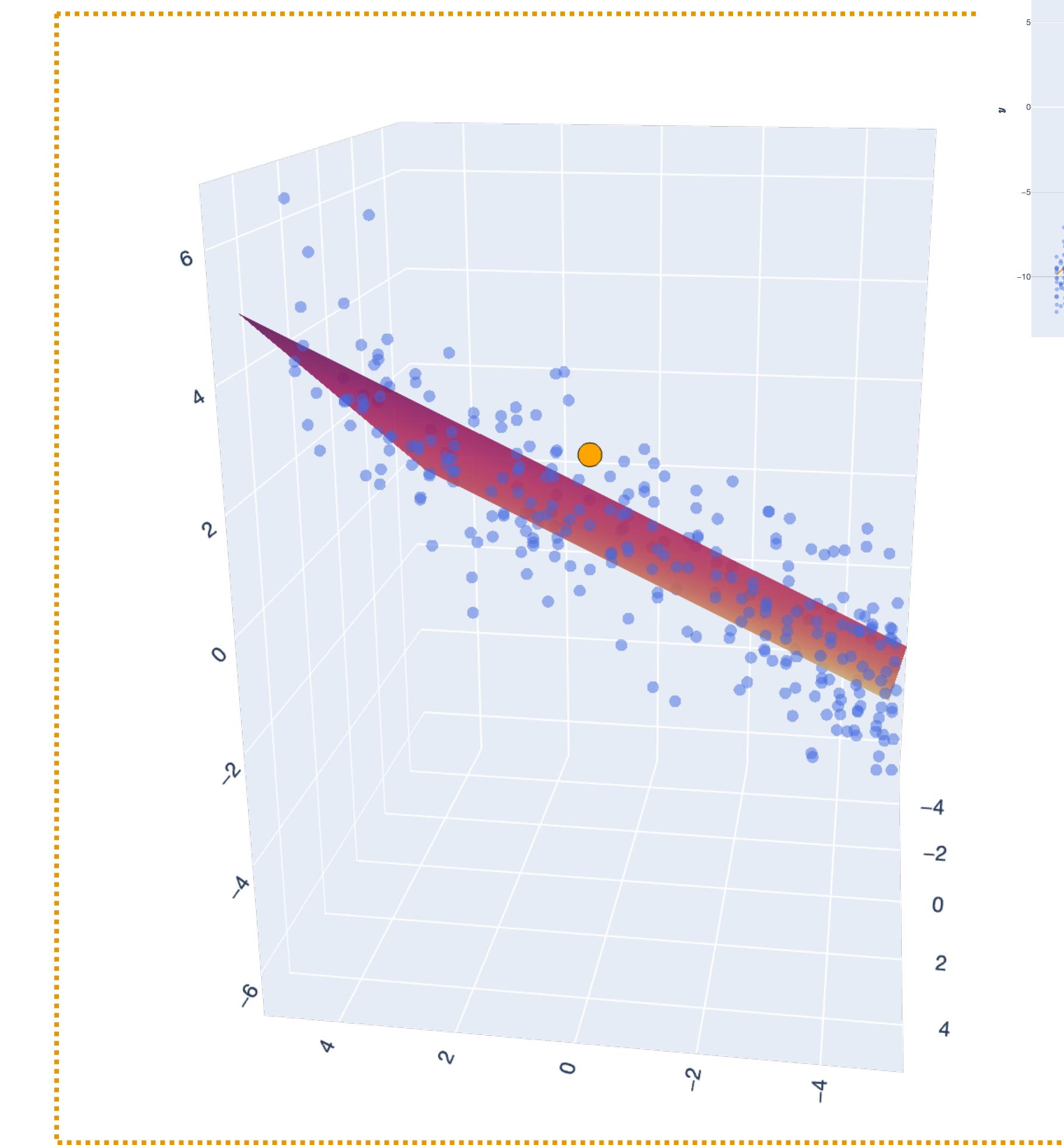
Chebyshev's inequality proved the (Weak) Law of Large Numbers: sample averages approach true means.

Sample average is a statistical estimator of the mean.

Estimators have bias and variance connected through the bias-variance decomposition of mean-squared error.

We found that OLS, as a random variable and estimator of  $\mathbf{w}^*$  is unbiased, has variance  $\text{Var}[\hat{\mathbf{w}}] = \sigma^2 \mathbb{E}[(\mathbf{X}^\top \mathbf{X})^{-1}]$ , and risk

$$R(\hat{\mathbf{w}}) = \mathbb{E}[(\hat{\mathbf{w}}^\top \mathbf{x}_0 - y_0)^2] \approx \sigma^2 + \frac{\sigma^2 d}{n}.$$



# Law of large numbers and statistical estimators

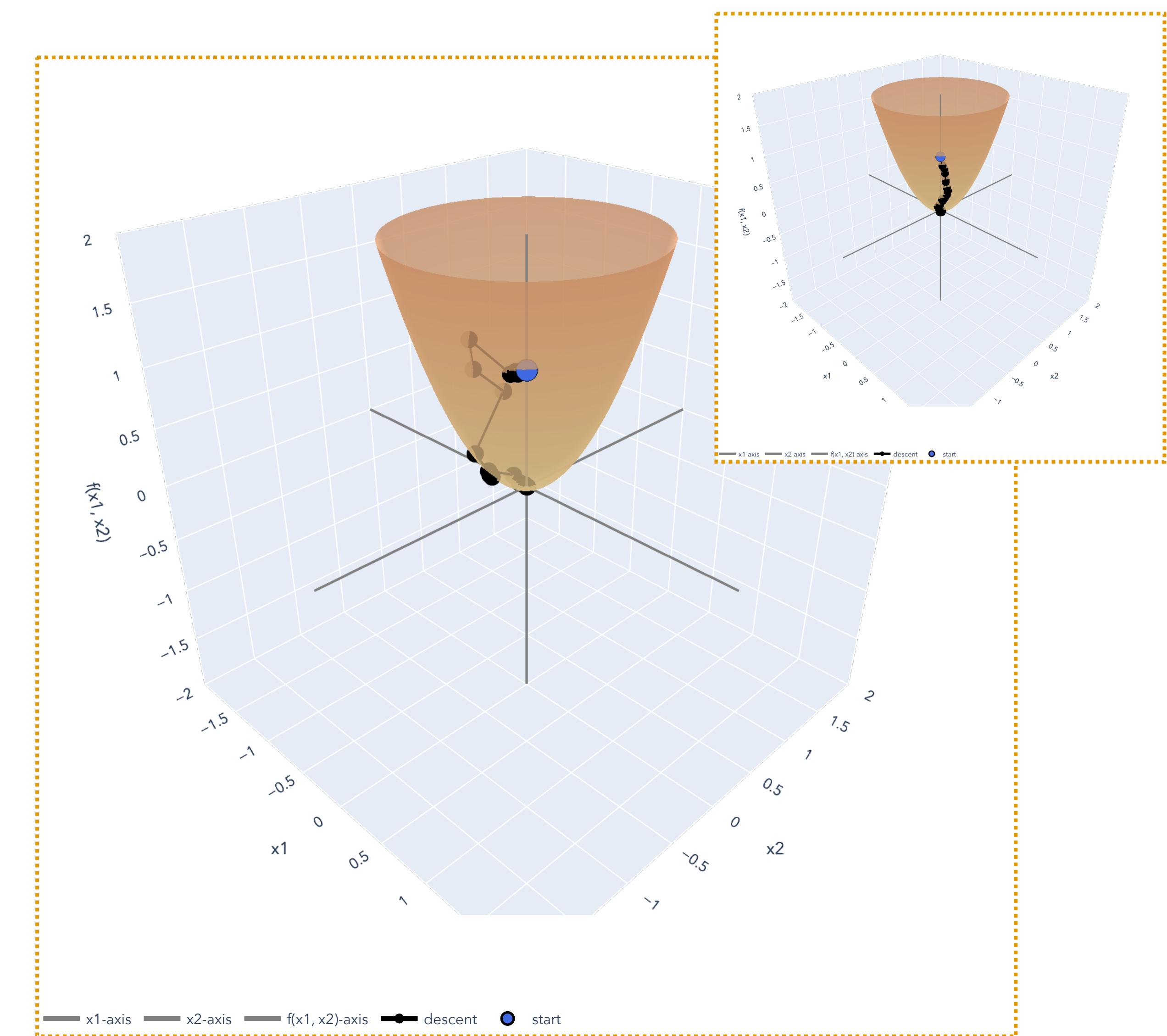
## Big Picture: Gradient Descent

We closed the story of gradient descent with stochastic gradient descent (SGD): instead of taking the gradient over *all* the samples  $(\mathbf{x}_1, y_1), \dots, (\mathbf{x}_n, y_n)$ , we used an unbiased estimator of the gradient:

$$\text{Estimand: } \nabla f(\mathbf{w}) = \frac{1}{n} \sum_{i=1}^n \nabla_{\mathbf{w}} (\mathbf{w}^\top \mathbf{x}_i - y_i)^2.$$

**Estimator:** Sample a single example  $i$  uniformly from  $1, \dots, n$  and take the gradient:

$$\widehat{\nabla f(\mathbf{w})} = \nabla_{\mathbf{w}} (\mathbf{w}^\top \mathbf{x}_i - y_i)^2.$$



# Week 6.1

## Central Limit Theorem, Distributions, and MLE

# Central Limit Theorem, Distributions, and MLE

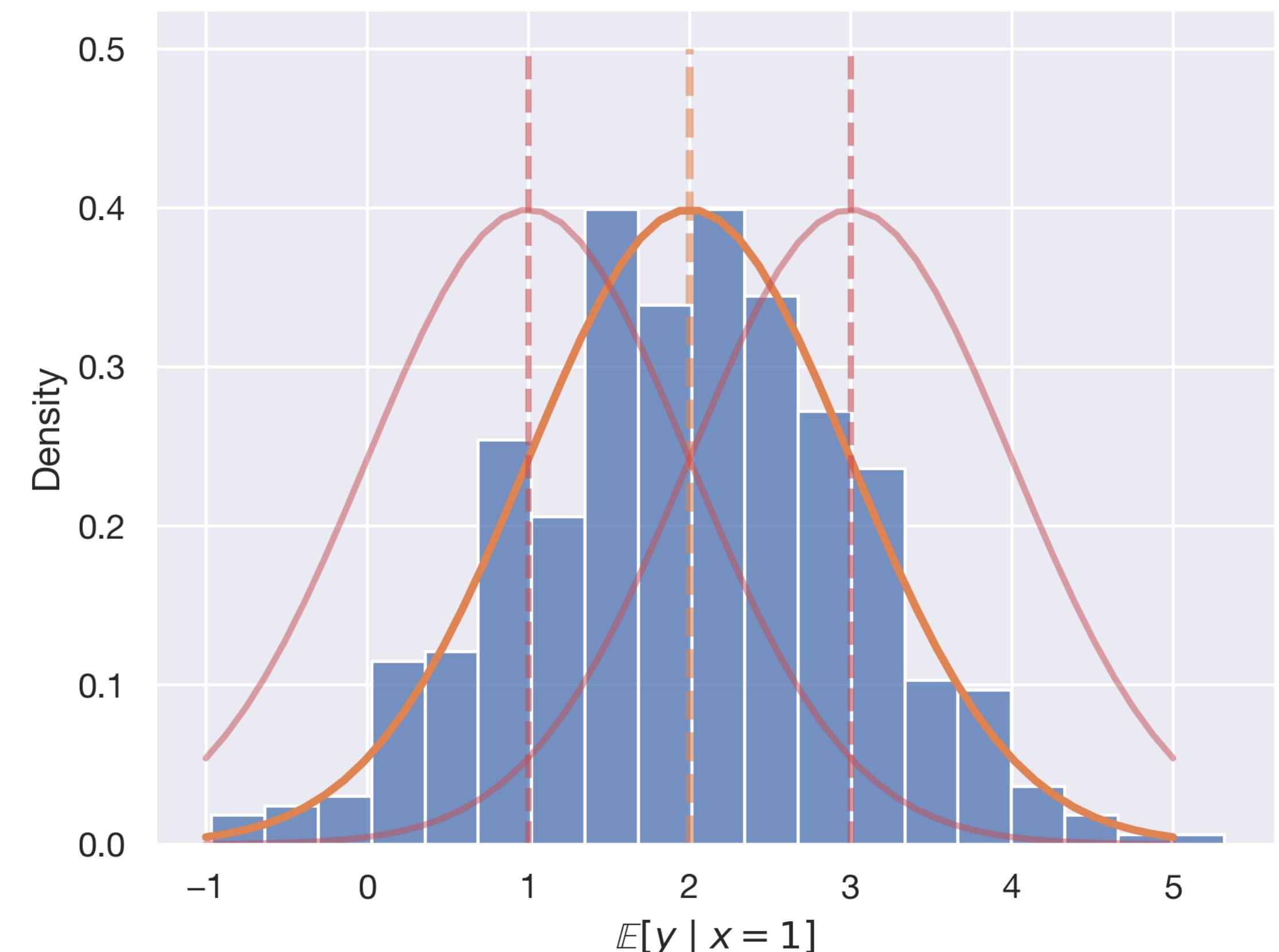
## Big Picture: Least Squares

We introduced the [Gaussian distribution](#), and we motivated its importance by the [Central Limit Theorem](#). The Gaussian distribution is just one of many “named distributions” that model common phenomena.

When we have a guess at a [parametrized model](#) or generating our i.i.d. data  $(\mathbf{x}_1, y_1), \dots, (\mathbf{x}_n, y_n)$ , an alternative perspective on our problem of finding a good model is [maximum likelihood estimation \(MLE\)](#).

This let us prove that, under the [Gaussian error model](#), maximizing the likelihood for the conditional distribution  $y | \mathbf{x}$  again gives us back the OLS estimator:

$$\hat{\mathbf{w}}_{MLE} = \arg \max L_n(\mathbf{w}) = (\mathbf{X}^\top \mathbf{X})^{-1} \mathbf{X}^\top \mathbf{y}$$



# Week 6.2

## Multivariate Gaussian Distribution

# Multivariate Gaussian Distribution

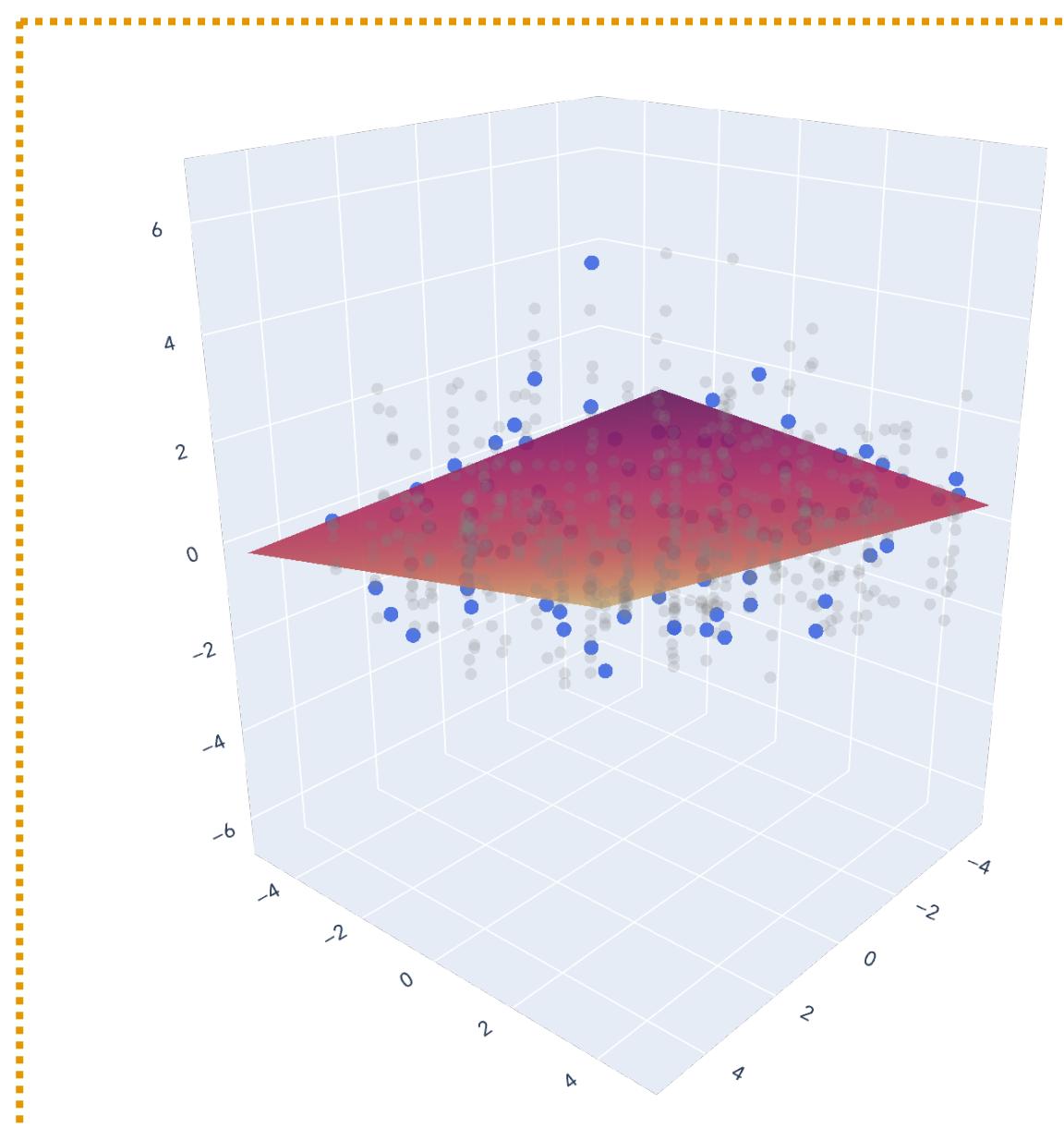
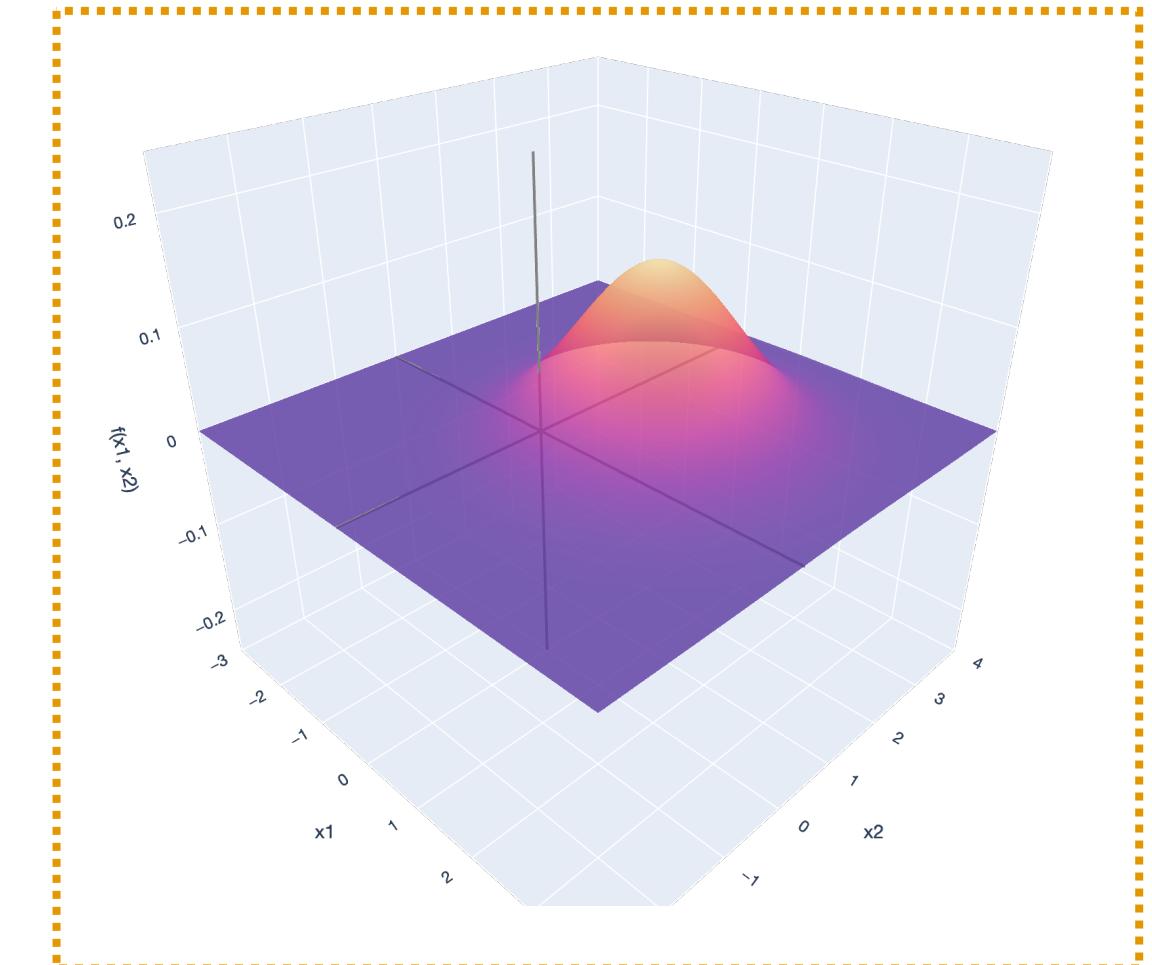
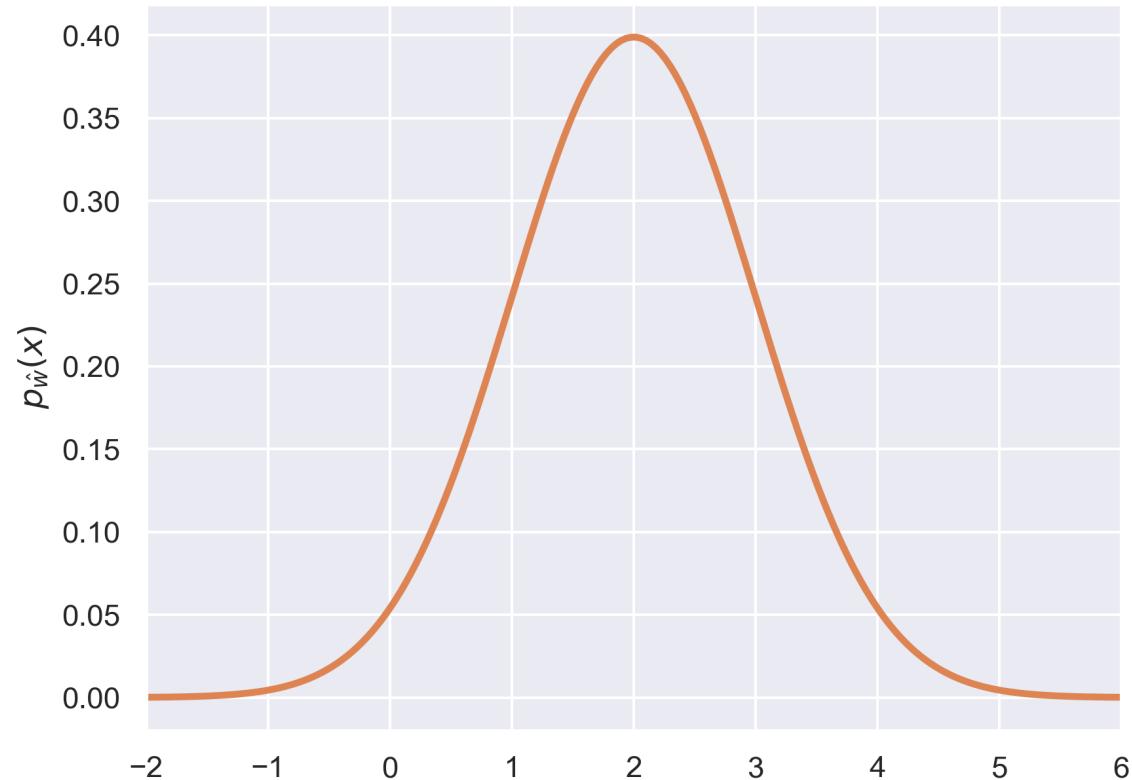
## Big Picture: Least Squares

We found that, under the [Gaussian error model](#), the distribution of the OLS estimator *itself* is [multivariate Normal/Gaussian](#).

$$\hat{\mathbf{w}} \sim N(\mathbf{w}^*, \sigma^2(\mathbf{X}^\top \mathbf{X})^{-1})$$

This motivated our study for the MVN distribution, which has properties:

1. Factorization under diagonal covariance.
2. Ellipsoidal geometry from eigendecomposition.
3. Affine transformations bridge standard MVN and general MVN.



What about the rest of ML?  
OLS and GD as a “Home Base”

# What about the rest of ML?

OLS and GD as a “Home Base”

$$\hat{\mathbf{w}} = (\mathbf{X}^\top \mathbf{X})^{-1} \mathbf{X}^\top \mathbf{y}$$

$$\mathbf{w}_t \leftarrow \mathbf{w}_{t-1} - \eta \nabla f(\mathbf{w}_{t-1})$$



# Extension 1: Nonlinear Models

## Feature transformations

# Nonlinear Models

## Feature Transformations

Now, consider the following nonlinear function,  $\phi : \mathbb{R}^2 \rightarrow \mathbb{R}^3$

$$\phi(x_1, x_2) = (x_1^2, x_1 x_2, x_2^2).$$

Because  $\phi(\cdot, \cdot)$  takes inputs in  $\mathbb{R}^2$ , we can feed it each row (sample) in our data matrix. This allows us to “transform” our data matrix to a new data matrix,  $\mathbf{X}' \in \mathbb{R}^{5 \times 3}$  by applying  $\phi(\cdot, \cdot)$  row by row. By doing so, we are constructing 3 new features from the  $d = 2$  old features.

**Problem 4(e) [4 points]** Find the transformed data matrix  $\mathbf{X}' \in \mathbb{R}^{5 \times 3}$  obtained by applying  $\phi(\cdot, \cdot)$  to each of the 5 rows. Find  $\mathbf{w} \in \mathbb{R}^d$  by least squares regression on  $\mathbf{X}'$  and the original  $\mathbf{y}$ . Also compute the sum of squared residuals error of your solution,  $\text{err}(\mathbf{w})$  (you should find that, now,  $\text{err}(\mathbf{w}) = 0$ ). You may use numpy or any other

It turns out that the true relationship between  $y_i$  and  $\mathbf{x}_i = (x_{i1}, x_{i2})$  for the data in (14) is actually:

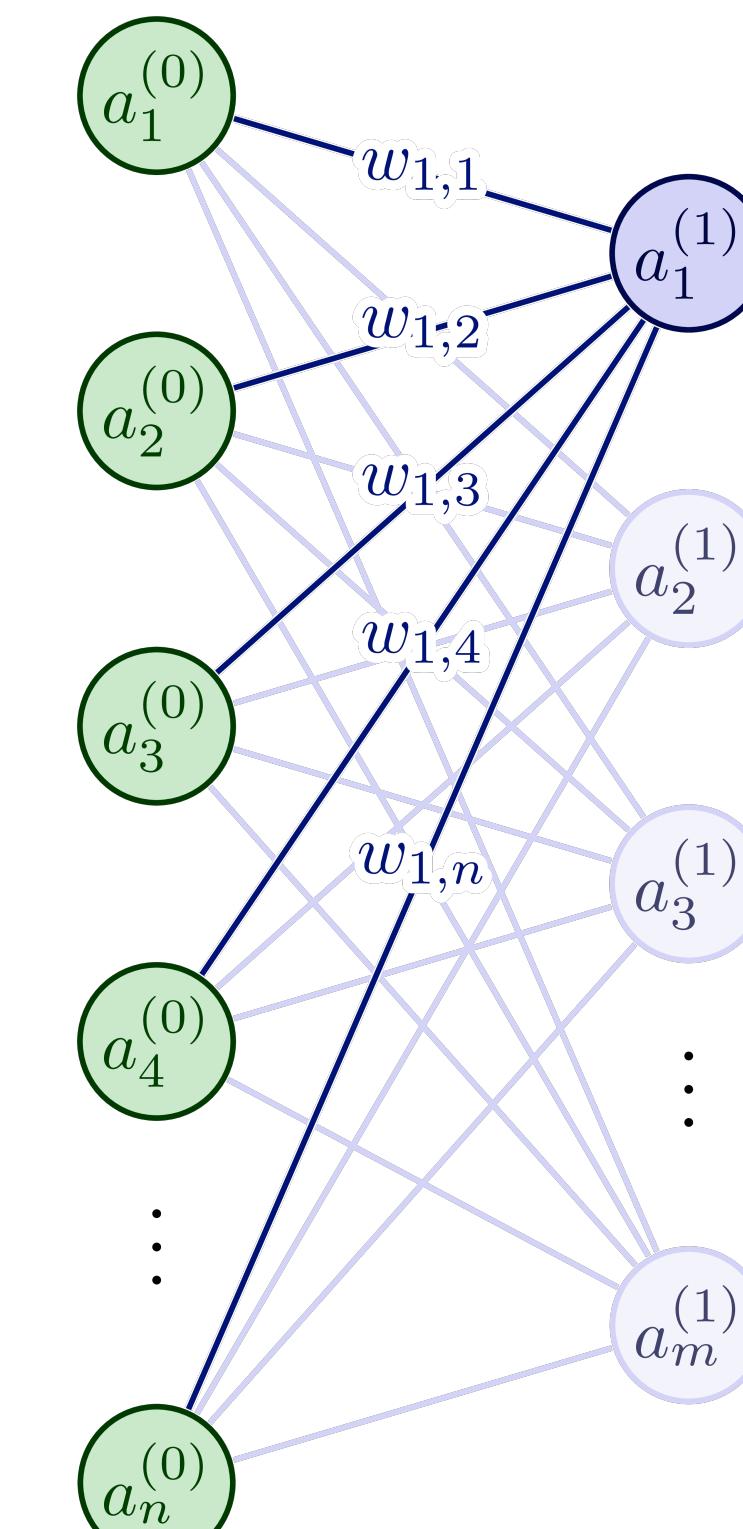
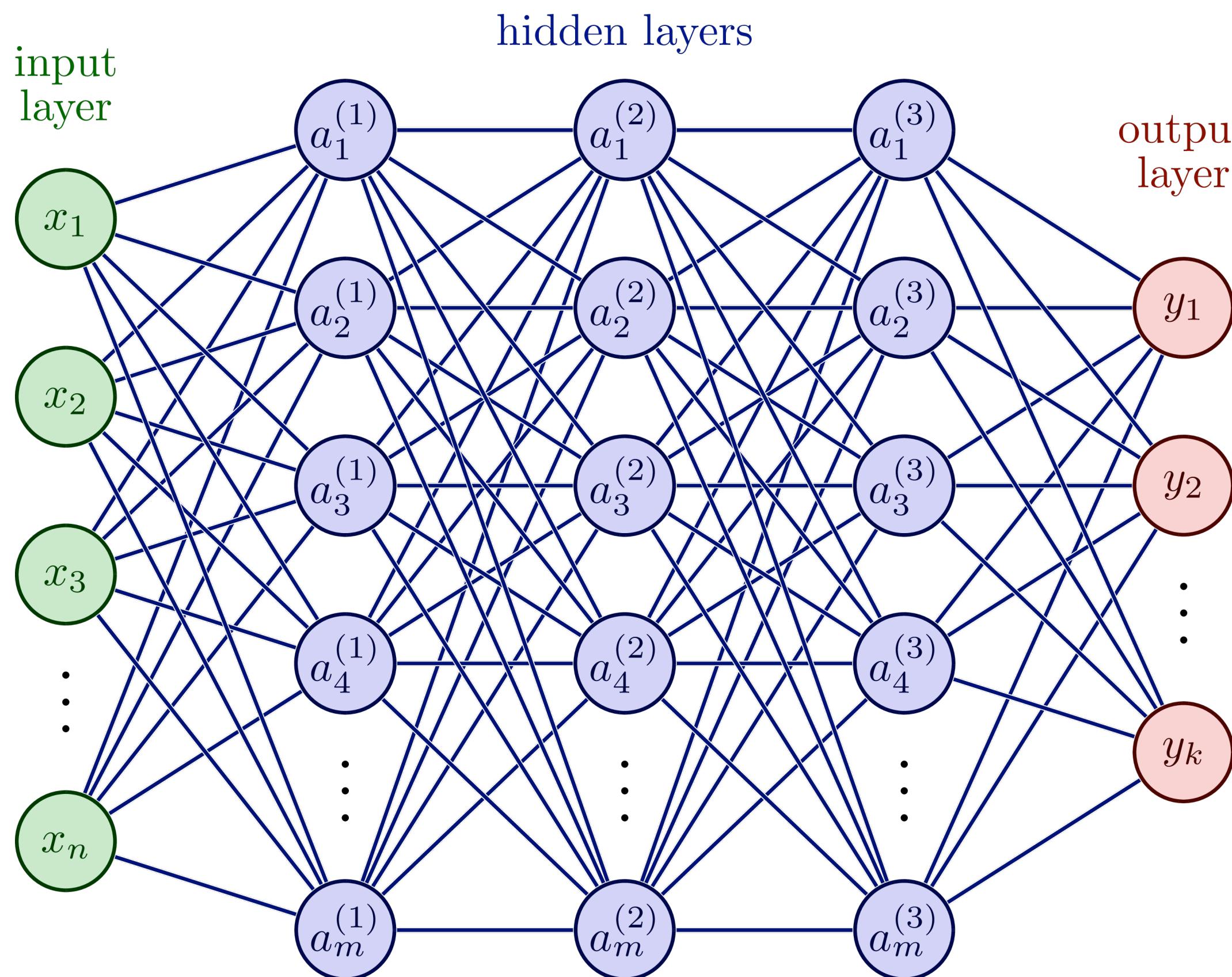
$$y_i = x_{i1}^2 + 2x_{i1}x_{i2} - x_{i2}^2 \quad \text{for all } i \in [n]. \quad (16)$$

By finding the feature transformation  $\phi(\cdot, \cdot)$  above, we turned a problem with a nonlinear relationship into a problem where a linear model is again useful (and, in fact, perfectly fits  $\mathbf{X}'$ ). We are back in our ideal scenario in Equation (12), but there now exists some  $\mathbf{w}^* \in \mathbb{R}^d$  such that

$$y_i = (\mathbf{w}^*)^\top \phi(\mathbf{x}_i).$$

# Nonlinear Models

## Neural Networks



$$a_1^{(1)} = \sigma(w_{1,1}a_1^{(0)} + w_{1,2}a_2^{(0)} + \dots + w_{1,n}a_n^{(0)} + b_1^{(0)})$$

$$= \sigma\left(\sum_{i=1}^n w_{1,i}a_i^{(0)} + b_1^{(0)}\right)$$

$$\begin{pmatrix} a_1^{(1)} \\ a_2^{(1)} \\ \vdots \\ a_m^{(1)} \end{pmatrix} = \sigma \left[ \begin{pmatrix} w_{1,1} & w_{1,2} & \dots & w_{1,n} \\ w_{2,1} & w_{2,2} & \dots & w_{2,n} \\ \vdots & \vdots & \ddots & \vdots \\ w_{m,1} & w_{m,2} & \dots & w_{m,n} \end{pmatrix} \begin{pmatrix} a_1^{(0)} \\ a_2^{(0)} \\ \vdots \\ a_n^{(0)} \end{pmatrix} + \begin{pmatrix} b_1^{(0)} \\ b_2^{(0)} \\ \vdots \\ b_m^{(0)} \end{pmatrix} \right]$$

$$\mathbf{a}^{(1)} = \sigma(\mathbf{W}^{(0)}\mathbf{a}^{(0)} + \mathbf{b}^{(0)})$$

# Extension 2: Loss Functions

## Beyond squared loss

# Loss Functions

## Beyond Squared Loss

# Extension 3: Algorithms

## Beyond gradient descent

# Algorithms

## Beyond Gradient Descent

# Extension 4: Learning Theory

## Other issues in generalization

# Learning Theory

Other issues in generalization

Thank you for listening!

Hope you enjoyed the class :)

