Mixed effects models

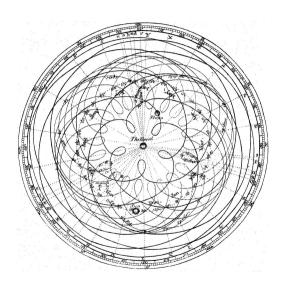
"Wheels within wheels"

Samuel Robinson, Ph.D.

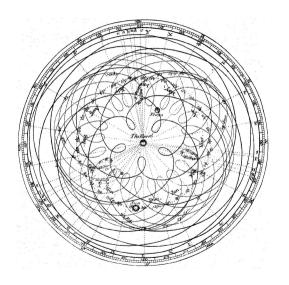
Oct 6, 2023

Part 1: Mixed effects models

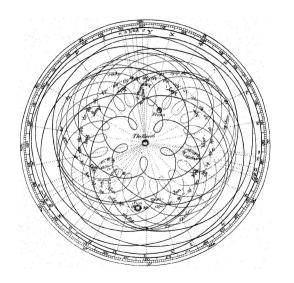
• Linear mixed effects models (LMMs)



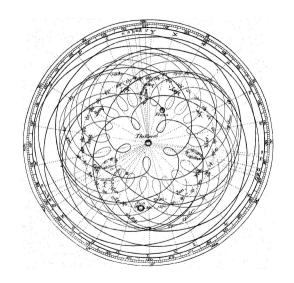
- Linear mixed effects models (LMMs)
 - A bit of math



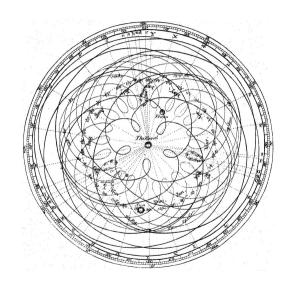
- Linear mixed effects models (LMMs)
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 - Fixed vs. random effects



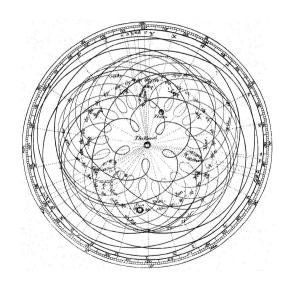
- Linear mixed effects models (LMMs)
 - A bit of math
 - Fixed vs. random effects
 - Random intercepts and slopes



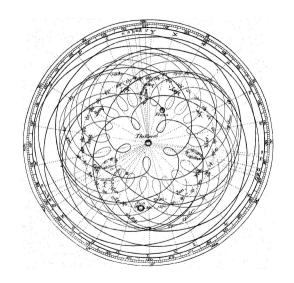
- Linear mixed effects models (LMMs)
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 - Fixed vs. random effects
 - Random intercepts and slopes
- Generalized linear mixed effects models (GLMMs)



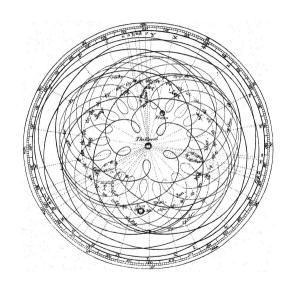
- Linear mixed effects models (LMMs)
 - A bit of math
 - Fixed vs. random effects
 - Random intercepts and slopes
- Generalized linear mixed effects models (GLMMs)
 - Residuals checks



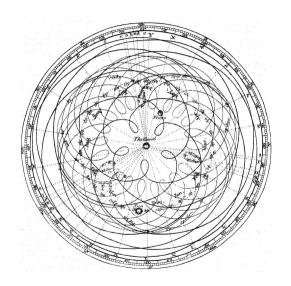
- Linear mixed effects models (LMMs)
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 - Some sage advice



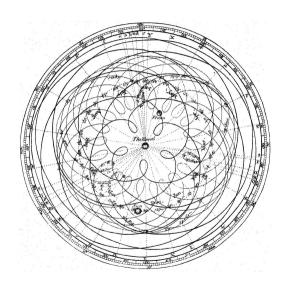
- Linear mixed effects models (LMMs)
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 - Fixed vs. random effects
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- Hypothesis testing and inference



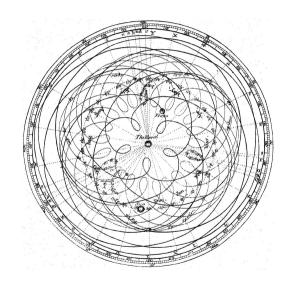
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 - AIC, R^2 , and Star-Gazing



Many different names:

Mixed effects models

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- Mixed effects models
- 2 Random effects models

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- 6 Split-plot models
- Variance partitioning

Scary math

Unfortunately, we need a review of matrix algebra in order to explain this:

• This is a matrix:

$$A = \begin{bmatrix} 1 & 4 & 7 \\ 2 & 5 & 8 \\ 3 & 6 & 9 \end{bmatrix}$$

$$A \times b = Ab = 1 \times \begin{bmatrix} 1 \\ 2 \\ 3 \end{bmatrix} + 2 \times \begin{bmatrix} 4 \\ 5 \\ 6 \end{bmatrix} + 3 \times \begin{bmatrix} 7 \\ 8 \\ 9 \end{bmatrix} = \begin{bmatrix} 30 \\ 36 \\ 42 \end{bmatrix}$$

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• Model matrix:

$$X = \begin{bmatrix} 1 & 1 & 10 \\ 1 & 1 & 12 \\ 1 & 0 & 9 \\ \vdots & \vdots & \vdots \end{bmatrix}$$

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Model matrix:

$$X = egin{bmatrix} 1 & 1 & 10 \ 1 & 1 & 12 \ 1 & 0 & 9 \ dots & dots & dots \end{bmatrix}$$

Multiplying them looks like:

$$\hat{y}=Xeta=egin{bmatrix} 1.60 \ 1.54 \ -0.17 \ dots \end{bmatrix}$$

This is exactly what R does to fit models:

head(dat)

```
## 1 1.5101095 -4.248450 g
## 2 3.7190900 5.766103 j
## 3 -4.3737644 -1.820462 f
## 4 30.1459331 7.660348 n
## 5 0.2777422 8.809346 o
## 6 -3.6978175 -9.088870 p

m1 <- lm(y~x,data=dat) #Uses x to predict y
summary(m1)
```

```
##
## Call:
## lm(formula = v \sim x. data = dat)
## Residuals:
       Min
                10 Median
                                         Max
## -18.2574 -4.1262 0.0296 3.1854 25.2780
## Coefficients:
              Estimate Std. Error t value Pr(>|t|)
## (Intercept) 2.7306 0.5772 4.731 4.92e-06 ***
## x
                0.7213
                       0.1020 7.074 4.60e-11 ***
## ---
## Signif. codes: 0 '***' 0.001 '**' 0.05 '.' 0.1 ' ' 1
## Residual standard error: 7.3 on 158 degrees of freedom
## Multiple R-squared: 0.2405, Adjusted R-squared: 0.2357
## F-statistic: 50.04 on 1 and 158 DF, p-value: 4.6e-11
```

This is exactly what R does to fit models (cont.):

```
head(model.matrix(m1))
    (Intercept)
             1 -4 248450
             1 5.766103
             1 -1.820462
             1 7.660348
             1 8.809346
## 5
             1 -9.088870
## 6
coef(m1)
## (Intercept)
    2 7305689 0 7212867
pred2 <- model.matrix(m1) %*% coef(m1) #predicted = matrix * coefs</pre>
head(data.frame(pred1=predict(m1),pred2)) #same thing!
        pred1
                  pred2
    -0.3337812 -0.3337812
     6.8895819 6.8895819
     1.4174942 1.4174942
     8.2558758 8.2558758
    9.0846325 9.0846325
## 6 -3.8251119 -3.8251119
```

Groups are coded by "dummy variables" (0s and 1s)

```
m2 <- lm(v~site.data=dat) #Use site to predict v
head(model.matrix(m2)) #0s and 1s used to identify groups
    (Intercept) siteb sitec sited sitee sitef siteg siteh sitei sitej sitek sitel
## 1
## 2
## 3
## 4
## 5
## 6
    sitem siten siteo sitep
## 1
## 2
## 3
## 4
             0
## 5
## 6
coef(m2) #This uses the 1st site as the "control" group
```

```
## (Intercept)
                     siteb
                                sitec
                                            sited
                                                        sitee
                                                                    sitef
     7.192416 -11.998464 -14.632803
                                         1.983649
                                                    -7.765354
                                                                -4.523079
                                sitei
                                                        sitek
                                                                    sitel
##
        siteg
                     siteh
                                            sitej
    -3.439621
                -8.280601
                            -4.306456
                                        -4.085855
                                                    -5.663021
                                                                -5.155112
        sitem
                     siten
                                siteo
                                            siten
##
    -6.226642
                 8.403599
                            -8.626661
                                       -10.934182
```

$$\hat{\mathbf{y}} = X\beta = b_0 1 + b_1 \mathbf{x}_1 \dots + b_i \mathbf{x}_i$$

 $\mathbf{y} \sim Normal(\hat{\mathbf{y}}, \boldsymbol{\sigma})$

• All linear models take the form:

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 $y \sim Normal(\hat{y}, \sigma)$

• y is a vector of data you want to predict

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- $X = \{1, x_1...\}$ is a matrix of *predictors* for y

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Structure of LMs... now with matrices!

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- $\beta = \{b_0, b_1, ...\}$ is a vector of *coefficients*
- $y \sim Normal(\hat{y}, \sigma)$ means:
 - "y follows a Normal distribution with mean \hat{y} and SD σ "

Say that X is a model matrix coding for 10 sites¹, and y is something we're interested in predicting

$$\hat{y} = b_0 + X\beta$$
 $\hat{y} = b_0 + X\zeta$ $y \sim Normal(\hat{y}, \sigma)$ $y \sim Normal(\hat{y}, \sigma)$ $\zeta \sim Normal(0, \sigma_{site})$

• Site coefficients (β) are unrelated to each other

¹Intercept is a separate variable

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- Site is a random effect

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A mixed effects model has both **fixed** and **random** effects

$$\hat{y} = X\beta + U\zeta$$
 $y \sim Normal(\hat{y}, \sigma)$
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- X =fixed effects matrix (e.g. intercept, temperature)
- β = fixed effects coefficients
- U = random effects matrix (e.g. sites)
- ζ = random effects coefficients
- σ , σ_{site} = variance terms

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• We're interested in predicting *y* using *x* (fixed effects)

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Let's go back to our earlier example:

- We're interested in predicting *y* using *x* (fixed effects)
- Data was collected at a number of sites, which may affect y "somehow"
- Effect of each site is normally distributed

head(dat)

```
summary(mm1)
## Linear mixed model fit by REML ['lmerMod']
## Formula: v ~ x + (1 | site)
     Data: dat
##
## REML criterion at convergence: 1040.4
##
## Scaled residuals:
        Min
                      Median
                                           Max
## -2 50816 -0 71380 -0 02682 0 69401 3 01951
##
## Random effects:
   Groups Name
                        Variance Std Dev
    site
            (Intercept) 26.78
                                 5.175
                        31.68
                                 5.628
   Residual
## Number of obs: 160, groups: site, 16
##
## Fixed effects:
              Estimate Std. Error t value
## (Intercept) 2.00331
                          1.38297
                                   1.449
               0.76192
                          0.08083
## x
                                   9.426
## Correlation of Fixed Effects:
    (Intr)
## x -0.002
```

Results from 1mer model:

Random effects:

²'Ime4' author doesn't think they can be calculated. I somewhat agree

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- Random effects:
 - residual and site variance $(\sigma, \sigma_{\text{site}})$
- Fixed effects:
 - Intercept and slope estimates (β)
 - No d.f. and p-value ²

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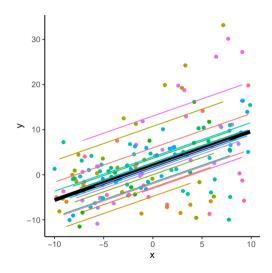
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- Random effects:
 - residual and site variance $(\sigma, \sigma_{\text{site}})$
- Fixed effects:
 - Intercept and slope estimates (β)
 - No d.f. and p-value ²
 - If you need p-values for parameters, you can use the *lmerTest* package (or just calculate them yourself using means/SEs)

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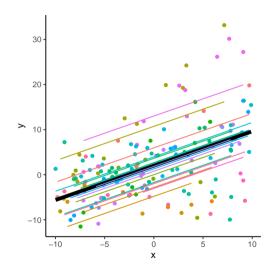
Mixed effect model results

 In a random intercepts model, the regression line of x on y is allowed to move up or down around the main regression line for each site



Mixed effect model results

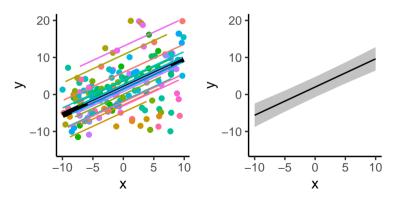
- In a random intercepts model, the regression line of x on y is allowed to move up or down around the main regression line for each site
- These changes in intercepts are normally distributed



Mixed effect model results (cont.)

For plotting, we want a partial effects plot that marginalizes across sites (i.e. "What does the trend look like at the average site?")

 ggpredict works well for this. If you want partial residuals, you'll have to add them in yourself using predict and residual



Suppose that y wasn't just higher or lower at each site, but that the effect of x on y was higher or lower at each site

$$\hat{y} = X\beta + U\zeta_{int} + U_{x}\zeta_{slope}$$
 $y \sim Normal(\hat{y}, \sigma)$
 $\zeta_{int} \sim Normal(0, \sigma_{int})$
 $\zeta_{slope} \sim Normal(0, \sigma_{slope})$

• X =fixed effects matrix (e.g. intercept, temperature)

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- X =fixed effects matrix (e.g. intercept, temperature)
- β = fixed effects coefficients

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- X =fixed effects matrix (e.g. intercept, temperature)
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- *U* = random intercept matrix (e.g. sites)

$$\hat{y} = X\beta + U\zeta_{int} + U_{x}\zeta_{slope}$$
 $y \sim Normal(\hat{y}, \sigma)$
 $\zeta_{int} \sim Normal(0, \sigma_{int})$
 $\zeta_{slope} \sim Normal(0, \sigma_{slope})$

- X =fixed effects matrix (e.g. intercept, temperature)
- β = fixed effects coefficients
- *U* = random intercept matrix (e.g. sites)
- $U_x = \text{random slopes matrix (e.g. temperature)}$

$$\hat{y} = X\beta + U\zeta_{int} + U_{x}\zeta_{slope}$$
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- U_X = random slopes matrix (e.g. temperature)
- $\zeta_{int}, \zeta_{slope} = \text{random intercept and slope coefficients}$

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- X =fixed effects matrix (e.g. intercept, temperature)
- β = fixed effects coefficients
- *U* = random intercept matrix (e.g. sites)
- $U_x = \text{random slopes matrix (e.g. temperature)}$
- ζ_{int} , ζ_{slope} = random intercept and slope coefficients
- σ , σ_{int} , σ_{slope} = variance terms

```
#Intercept varies with site, and slope of x can
# also vary with site (both hierarchical)
mm2 \leftarrow lmer(y \sim x + (x|site), data=dat)
summary(mm2)
## Linear mixed model fit by REML ['lmerMod']
## Formula: v ~ x + (x | site)
     Data: dat
##
## REML criterion at convergence: 900.6
##
## Scaled residuals:
       Min
                     Median
## -2.10500 -0.64857 0.02414 0.61137 2.22996
## Random effects:
   Groups Name
                   Variance Std.Dev. Corr
   site (Intercept) 35,2210 5,9347
                      0.7889 0.8882
                                       0.82
                       9.3162 3.0522
  Residual
## Number of obs: 160, groups: site, 16
##
## Fixed effects:
             Estimate Std. Error t value
## (Intercept) 2.0383 1.5091 1.351
             0.6438
                         0.2275 2.830
## Y
## Correlation of Fixed Effects:
   (Intr)
## x 0.790
```

Results from 1mer model:

Random effects:

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 - residual, slope, and site variance (σ , σ_{int} , σ_{slope})
 - Correlation b/w intercept and slope = 0.82
 - Sites with higher intercept also have a higher slope
- Fixed effects:
 - Intercept and slope estimates

• X

X

U

X

U

Ux

sitea siteb sitec sited sitee sitef siteg siteh sitei sitej sitek

X

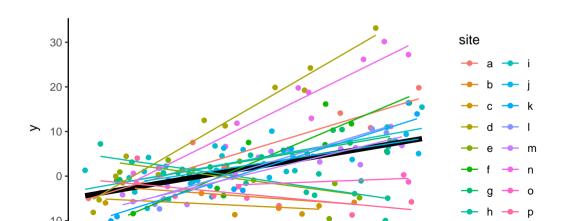
U

Ux

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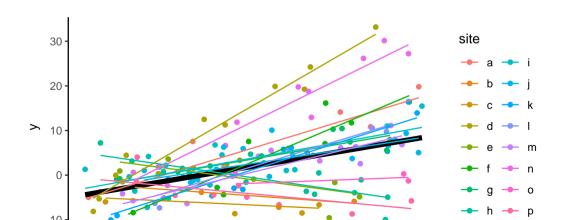
Mixed effect model results

• Regression line of x on y is allowed to move up or down around the main regression line for each site (random intercepts)



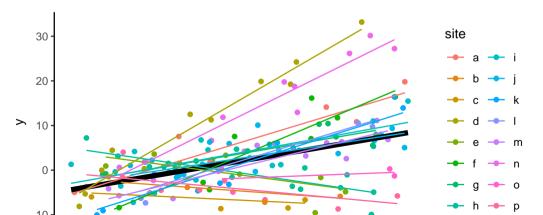
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Mixed effect model results

- Regression line of x on y is allowed to move up or down around the main regression line for each site (random intercepts)
- Slope of regression line can be more or less steep for each site (random slopes)
- Changes in intercepts and slopes are *normally distributed*, and in this example, are correlated with each other



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 - Easier to interpret (p-values, ANOVA, etc.)

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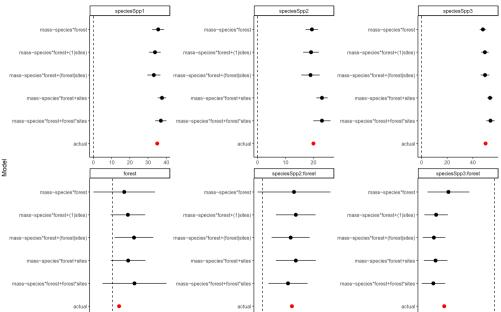
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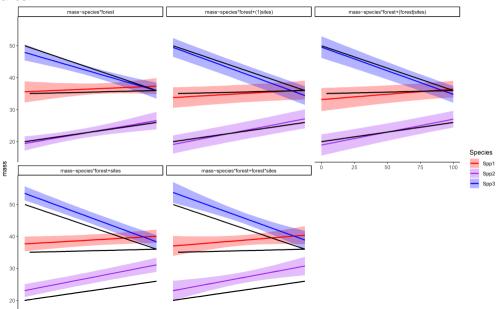
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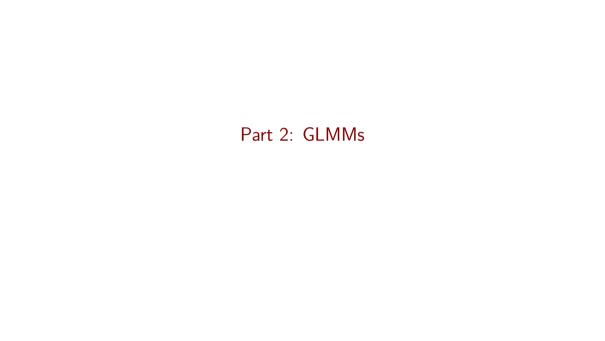
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- Fit a mixed effects model with the fixed effects you're interested in (forest cover, species), and include site as a random effect (intercept or slope)
- Your supervisor doesn't like hierarchical models, and tells you to just use site as another fixed term in an 1m model. Do you get different results if you use their approach?

Results:



Results:





Linear model (LM)

 $\hat{\mathbf{v}} = X\beta$

$$y \sim Normal(\hat{y}, \sigma)$$
 $\hat{y} = X\beta + U\zeta$ $y \sim Normal(\hat{y}, \sigma)$ $\zeta \sim Normal(0, \sigma_{site})$

$$logit(\hat{\phi}) = X\beta$$
 $y \sim Binomial(\hat{\phi})$

$$egin{aligned} logit(\hat{m{\phi}}) &= m{X}m{eta} + m{U}m{\zeta} \ & m{y} \sim Binomial(\hat{m{\phi}}) \ & m{\zeta} \sim Normal(0, m{\sigma_{site}}) \end{aligned}$$

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• Generalized linear model (GLM)

$$logit(\hat{\phi}) = oldsymbol{X}eta \ y \sim Binomial(\hat{\phi})$$

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Linear mixed effects model (LMM)

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 Generalized linear mixed effects model (GLMM)

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 $y \sim Binomial(\hat{\phi})$
 $\zeta \sim Normal(0, \sigma_{site})$

How do I fit GLMMs?

 glmer and glmer.nb from library(lme4) work for Binomial, Poisson, and Negative Binomial data

```
library(lme4)
glmm1 <- glmer.nb(y~x+(x|site),data=dat2) #Negative binomial GLMM
summary(glmm1) #glmer.nb takes a long time to run
```

```
library(glmmTMB)
glmm2 <- glmmTMB(y~x+(x|site),data=dat2,family=nbinom2())
summary(glmm2) #Similar results, but quicker
## Family: nbinom2 (log)
              v \sim x + (x \mid site)
## Formula:
## Data: dat2
       AIC
          BIC logLik deviance df.resid
     627.8
            646.2 -307.9 615.8
                                      154
## Random effects:
## Conditional model:
   Groups Name
                    Variance Std.Dev. Corr
   site (Intercept) 1.43543 1.1981
                    0.02892 0.1701 0.92
```

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 e.g. Zero-inflation, Beta-binomial, Spatial Models

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## Data: dat2
##
## AIC BIC logLik deviance df.resid
## 627.8 646.2 -307.9 615.8 154
```

Conditional model: ## Groups Name Variance Std.Dev. Corr ## site (Intercept) 1.43543 1.1981 ## x 0.02892 0.1701 0.92

Random effects:

Fitting GLMMs - glmer.nb

```
## Generalized linear mixed model fit by maximum likelihood (Laplace
    Approximation) [glmerMod]
## Family: Negative Binomial(5.1294) (log)
## Formula: y ~ x + (x | site)
     Data: dat2
       AIC
               BIC logLik deviance df.resid
##
     627.8
             646.3 -307.9 615.8
                                      154
##
## Scaled residuals:
      Min
          1Q Median
                             30
                                    Max
## -1.3745 -0.7098 -0.3946 0.5108 2.5367
##
## Random effects:
## Groups Name
                    Variance Std.Dev. Corr
## site (Intercept) 1.43500 1.1979
                     0.02878 0.1697 0.92
          v
## Number of obs: 160, groups: site, 16
##
## Fixed effects:
              Estimate Std. Error z value Pr(>|z|)
##
## (Intercept) 0.32746
                         0.32166
                                  1.018 0.3087
## v
              0.10830
                       0.04681
                                  2.314
                                         0.0207 *
## ---
## Signif. codes: 0 '***' 0.001 '**' 0.05 '.' 0.1 ' ' 1
##
## Correlation of Fixed Effects:
## (Intr)
## x 0.799
```

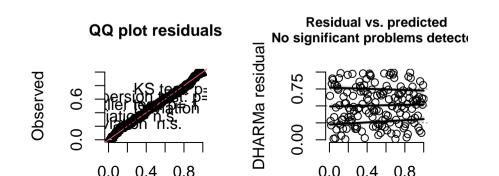
Fitting GLMMs - glmmTMB

```
## Family: nbinom2 (log)
## Formula:
                   v ~ x + (x | site)
## Data: dat2
       AIC
               BIC
                    logLik deviance df.resid
     627.8
            646.2 -307.9
                               615.8
                                         154
##
##
## Random effects:
##
## Conditional model:
## Groups Name
                     Variance Std.Dev. Corr
   site (Intercept) 1.43543 1.1981
                     0.02892 0.1701 0.92
## Number of obs: 160, groups: site, 16
##
## Dispersion parameter for nbinom2 family (): 5.12
##
## Conditional model:
             Estimate Std. Error z value Pr(>|z|)
## (Intercept) 0.34132 0.32172 1.061 0.2887
## x
              0.11026
                       0.04697
                                   2.348 0.0189 *
## ---
## Signif. codes: 0 '***' 0.001 '**' 0.01 '*' 0.05 '.' 0.1 ' ' 1
```

Residual checks on glmmTMB models

 Extract residuals and make your own plots, or use simulateResiduals from library(DHARMa) (see here)

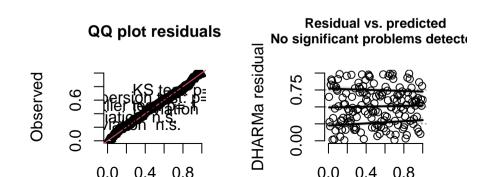
DHARMa residual



Residual checks on glmmTMB models

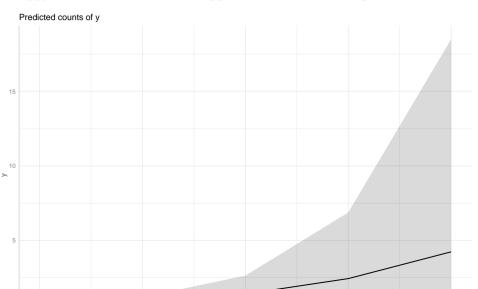
- Extract residuals and make your own plots, or use simulateResiduals from library(DHARMa) (see here)
- DHARMa also has useful functions for checking overdispersion and zero-inflation (found here)

DHARMa residual



Partial residual plots for glmmTMB models

• ggpredict() from library(ggeffects) works with glmmTMB models



Part 3: Hypothesis testing and inference

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- How well did your model fit your data (overall)?
- Some other bits of advice

Step 1: was the term "important"?

In linear models, this is done using an ANOVA F-test (also shown at the bottom of a summary() statement):

```
## Analysis of Variance Table
##
## Response: mpg
## Df Sum Sq Mean Sq F value Pr(>F)
## disp 1 808.89 808.89 73.9959 1.788e-09 ***
## gear 1 0.14 0.14 0.0132 0.9093
## Residuals 29 317.01 10.93
## ---
## Signif. codes: 0 '***' 0.001 '**' 0.01 '*' 0.05 '.' 0.1 ' ' 1
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But wait, there's more!

Unfortunately, if we have more than 1 term in the model, the order of terms can change your answer:

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## Response: mpg
           Df Sum Sq Mean Sq F value
                                      Pr(>F)
## gear 1 259.75 259.75 23.762 3.595e-05 ***
## disp
           1 549.28 549.28 50.248 8.465e-08 ***
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## Single term deletions
##
## Model:
## mpg ~ disp + gear
## Df Sum of Sq RSS AIC F value Pr(>F)
## <none> 317.01 79.383
## disp 1 549.28 866.30 109.552 50.2476 8.465e-08 ***
## gear 1 0.14 317.16 77.397 0.0132 0.9093
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## Signif. codes: 0 '***' 0.001 '**' 0.01 '*' 0.05 '.' 0.1 ' ' 1
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## Single term deletions
##
## Model:
## mpg ~ gear * disp
## | Df Sum of Sq RSS AIC F value Pr(>F)
## <none> 278.60 77.250
## gear:disp 1 38.412 317.01 79.383 3.8604 0.05943 .
## ---
## Signif. codes: 0 '***' 0.001 '**' 0.01 '*' 0.05 '.' 0.1 ' ' 1
```

 Anova from the car package tests other terms after interactions, but the meaning isn't the same. I prefer to keep things simple and just use drop1()

```
## Anova Table (Type II tests)
##
## Response: mpg
## Sum Sq Df F value Pr(>F)
## gear 0.14 1 0.0145 0.90502
## disp 549.28 1 55.2038 4.312e-08 ***
## gear:disp 38.41 1 3.8604 0.05943 .
## Residuals 278.60 28
## ---
## Signif. codes: 0 '***' 0.001 '**' 0.01 '*' 0.05 '.' 0.1 ' ' 1
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- Unfortunately, different numbers of data points change your likelihood, so this can wreck LRTs. This happens a lot if you have NAs in your predictor columns, so clean up your data before using it in models.

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Solution:

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- ① Use ML if comparing between models with different fixed effects, then...
- 2 Re-fit with REML once you've decided on a model

How many terms should be in my model?

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 - AIC tests usually say the same thing as LR tests

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