# **Expectation Maximisation** (EM)

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### Outline ... EM

- Importance
- Goal
- Idea
- Derivation
- Visualisation

### Importance ...

#### Maximum Likelihood from Incomplete Data via the EM Algorithm

By A. P. Dempster, N. M. Laird and D. B. Rubin

Harvard University and Educational Testing Service



Dempster, A. P., Laird, N. M. & Rubin, D. B. (1977). *Maximum likelihood from incomplete data via the EM algorithm*. **Journal of the Royal Statistical Society**: Series B, 39, 1-38.

### Importance ...

#### Maximum Likelihood from Incomplete Data via the EM Algorithm

By A. P. Dempster, N. M. Laird and D. B. Rubin

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https://www.jstor.org > stable



by AP Dempster · 1977 · Cited by 69463 — A broadly applicable algorithm for computing maximum likelihood estimates from incomplete data is presented at various levels of generality.

21, Feb, 2023



### Importance ...

Keywords ...

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# Setup

$$X:observable rv^*$$

$$\mathbf{X} = \{\mathbf{x}_1, \mathbf{x}_2, ..., \mathbf{x}_N\},$$

$$\mathbf{x}_i \in \mathbb{R}^{D_1}$$

$$X: observable ext{ rv}^*$$
  $Z: latent ext{ rv}$   $\mathbf{X} = \{\mathbf{x}_1, \mathbf{x}_2, ..., \mathbf{x}_N\}, \quad \mathbf{x}_i \in \mathbb{R}^{D_1}$   $\mathbf{Z} = \{\mathbf{z}_1, \mathbf{z}_2, ..., \mathbf{z}_N\}, \quad \mathbf{z}_i \in \mathbb{R}^{D_2}$ 

 $\mathbf{X}: incomplete \ \mathrm{data}$ 

 $\{\mathbf{X}, \mathbf{Z}\}$ : complete data

 $\theta$ : model parameters

### Setup

independent

$$p(\mathbf{X}|\theta) = p(\mathbf{x}_1, \mathbf{x}_2, ..., \mathbf{x}_N|\theta)$$
 identically distributed 
$$p(\mathbf{X}|\theta) = \sum_{i=1}^{N} \log p(\mathbf{x}_i|\theta)$$
  $p(\mathbf{x}_i|\theta)$ 

$$p(\mathbf{X}, \mathbf{Z}|\theta) = p(\mathbf{x}_1, \mathbf{z}_1, \mathbf{x}_2, \mathbf{z}_2, ..., \mathbf{x}_N, \mathbf{z}_N | \theta) \stackrel{i.i.d}{=} \prod_{i=1}^{N} p(\mathbf{x}_i, \mathbf{z}_i | \theta)$$
$$\log p(\mathbf{X}|\theta) = \log \sum_{\mathbf{Z}} p(\mathbf{X}, \mathbf{Z}|\theta)$$

### Setup

$$P(\mathbf{X}|\theta)$$
: incomplete data likelihood

$$P(\mathbf{X}, \mathbf{Z}|\theta)$$
: complete data likelihood

$$P(\mathbf{Z}|\mathbf{X}, \theta) : posterior \text{ probability}$$

#### Marginalisation

$$p(\mathbf{X}|\theta) = \sum_{\mathbf{Z}} p(\mathbf{X}, \mathbf{Z}|\theta)$$

#### Chain rule (probability)

$$p(\mathbf{X}|\theta) = \frac{p(\mathbf{X}, \mathbf{Z}|\theta)}{p(\mathbf{Z}|\mathbf{X}, \theta)}$$

### Goal ...

$$\theta_{ML}^* = \underset{\theta \in \Theta}{\operatorname{argmax}} \ p(\mathbf{X}|\theta) = \underset{\theta \in \Theta}{\operatorname{argmax}} \ \log p(\mathbf{X}|\theta)$$

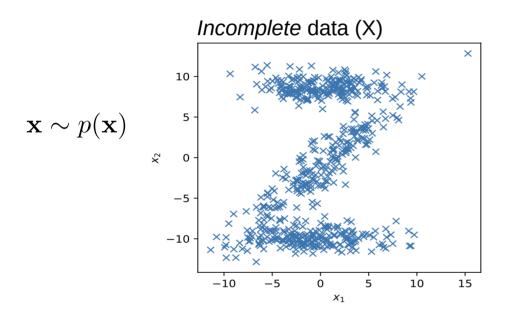
- log is *strictly increasing* → argmax remains identical
- Advantages:
  - ✓ Mathematical convenience → log[exp(.)]
  - Numerical Stability

### Goal ...

$$\theta_{ML}^*(\mathbf{X}|\theta) = \underset{\theta \in \Theta}{\operatorname{argmax}} \log p(\mathbf{X}|\theta)$$
$$= \underset{\theta \in \Theta}{\operatorname{argmax}} \log \sum_{\mathbf{Z}} p(\mathbf{X}, \mathbf{Z}|\theta)$$

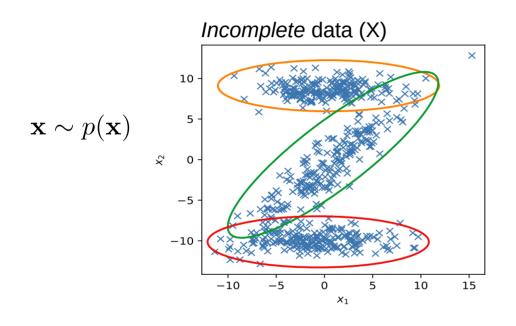
**EM assumes** model includes *latent variables* (**Z**).

# Latent Variable (Z)



**Interpretation**: *latent variable* is a part of a model ... *explains* **X**.

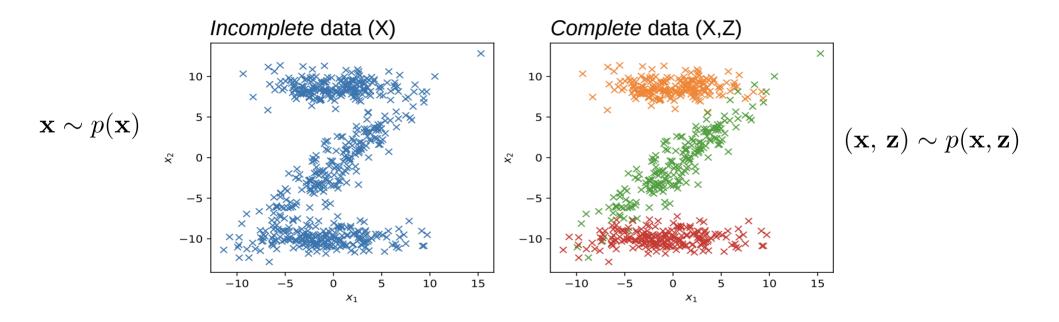
### Latent Variable (Z)



Consider clustering ...

**Interpretation**: *latent variable* is a part of a model ... *explains* **X**.

# Latent Variable (Z)



**Interpretation**: *latent variable* is a part of a model ... *explains X*.

### **Direct Solution**

$$\theta_{ML}^* = \underset{\theta \in \Theta}{\operatorname{argmax}} \log \sum_{\mathbf{Z}} p(\mathbf{X}, \mathbf{Z} | \theta)$$

$$\log \sum_{\mathbf{Z}} p(\mathbf{X}, \mathbf{Z} | \theta)$$

$$\frac{\partial \log p(\mathbf{X}|\theta)}{\partial \theta} = 0$$

Step 1

Step 2

$$\left. \frac{\partial^2 \log p(\mathbf{X}|\theta)}{\partial \theta^2} \right|_{\theta = \theta_0} < 0$$

Step 3

 $\theta_0$ : derivative roots

### **Direct Solution**

$$\theta_{ML}^* = \underset{\theta \in \Theta}{\operatorname{argmax}} \log \sum_{\mathbf{Z}} p(\mathbf{X}, \mathbf{Z} | \theta)$$

$$\log \sum_{\mathbf{Z}} p(\mathbf{X}, \mathbf{Z} | \theta)$$

Intractable (no closed-form solution)

$$\frac{\partial \log p(\mathbf{X}|\theta)}{\partial \theta} = 0$$

Step 1

Step 2

$$\left. \frac{\partial^2 \log p(\mathbf{X}|\theta)}{\partial \theta^2} \right|_{\theta = \theta_0} < 0$$

Step 3

 $\theta_0$ : derivative roots

### Direction solution does not work!

$$\log \sum_{z} p(x,z|\theta) = \log(\ldots + w_{z_i}e^{\frac{(x-\mu_i)^2}{2\sigma_i^2}} + \ldots + w_{z_j}e^{\frac{(x-\mu_j)^2}{2\sigma_j^2}} + \ldots)$$
 Consider simplest case ... Gaussian

$$\frac{\partial \log p(x|\theta)}{\partial \theta} = 0$$

Intractable

### Direction solution does not work!

$$\log \sum p(x, z | \theta) = \log(\dots + w_{z_i} e^{\frac{(x - \mu_i)^2}{2\sigma_i^2}} + \dots + w_{z_j} e^{\frac{(x - \mu_j)^2}{2\sigma_j^2}} + \dots)$$

 $\dots = 0 \rightarrow Intractable$ 

### Direction solution does not work!

$$\log \sum_{z} p(x,z|\theta) = \log(... + w_{z_i}e^{\frac{(x-\mu_i)^2}{2\sigma_i^2}} + ... + w_{z_j}e^{\frac{(x-\mu_j)^2}{2\sigma_j^2}} + ...)$$
 ... = 0  $\rightarrow$  Intractable if we could swap  $\log \& \Sigma$  ... ... = 0  $\rightarrow$  Tractable 
$$\sum_{z} \log p(x,z|\theta) = (... + w_i \log e^{\frac{(x-\mu_i)^2}{2\sigma_i^2}} + ... + w_j \log e^{\frac{(x-\mu_j)^2}{2\sigma_j^2}} + ...)$$
 
$$= (... + c_i(x-\mu_i)^2 + ... + c_j(x-\mu_j)^2 + ...)$$

#### Optimise ... $\theta^*$

$$\log \sum_{\mathbf{Z}} p(\mathbf{X}, \mathbf{Z} | \theta)$$



Optimise ...  $\theta^*$ 

$$\sum_{\mathbf{Z}} \log p(\mathbf{X}, \mathbf{Z} | \theta)$$

Intractable

**Tractable** 



$$\log \sum_{\mathbf{Z}} p(\mathbf{X}, \mathbf{Z} | \theta)$$



Optimise ...  $\theta^*$ 

$$\sum_{\mathbf{Z}} \log p(\mathbf{X}, \mathbf{Z} | \theta)$$

#### Intractable

#### **Tractable**

Your problem is to bridge the gap which exists between where you are now and the goal you intend to reach.

Earl Nightingale (1921-1989)

Optimise ...  $\theta^*$ 

$$\log \sum_{\mathbf{Z}} p(\mathbf{X}, \mathbf{Z} | \theta)$$



Optimise ... θ\*

$$\sum_{\mathbf{Z}} \log p(\mathbf{X}, \mathbf{Z} | \theta)$$

#### Intractable

**Tractable** 

Your problem is to bridge the gap which exists between where you are now and the goal you intend to reach.

Earl Nightingale (1921-1989)

$$p(\mathbf{X}|\theta) = \frac{p(\mathbf{X}, \mathbf{Z}|\theta)}{p(\mathbf{Z}|\mathbf{X}, \theta)}$$

**Step 0**: write  $p(X|\theta)$  using *chain rule* 

$$p(\mathbf{X}|\theta) = \frac{p(\mathbf{X}, \mathbf{Z}|\theta)}{p(\mathbf{Z}|\mathbf{X}, \theta)} \frac{q(\mathbf{Z})}{q(\mathbf{Z})}$$

$$= \frac{p(\mathbf{X}, \mathbf{Z}|\theta)}{q(\mathbf{Z})} \frac{q(\mathbf{Z})}{p(\mathbf{Z}|\mathbf{X}, \theta)}$$

**Step 1**: Multiply right-hand side in a *special 1* 

$$\log p(\mathbf{X}|\theta) = \log \left[ \frac{p(\mathbf{X}, \mathbf{Z}|\theta)}{q(\mathbf{Z})} \frac{q(\mathbf{Z})}{p(\mathbf{Z}|\mathbf{X}, \theta)} \right]$$
$$= \log \left[ \frac{p(\mathbf{X}, \mathbf{Z}|\theta)}{q(\mathbf{Z})} \right] + \log \left[ \frac{q(\mathbf{Z})}{p(\mathbf{Z}|\mathbf{X}, \theta)} \right]$$

**Step 2**: Take *log* from both sides

$$q(\mathbf{Z})\log p(\mathbf{X}|\theta) = q(\mathbf{Z})\log \left\lfloor \frac{p(\mathbf{X}, \mathbf{Z}|\theta)}{q(\mathbf{Z})} \right\rfloor + q(\mathbf{Z})\log \left\lfloor \frac{q(\mathbf{Z})}{p(\mathbf{Z}|\mathbf{X}, \theta)} \right\rfloor$$

**Step 3**: Multiply both sides by q(Z)

$$\sum_{\mathbf{Z}} q(\mathbf{Z}) \log p(\mathbf{X}|\theta) = \sum_{\mathbf{Z}} q(\mathbf{Z}) \log \left[ \frac{p(\mathbf{X}, \mathbf{Z}|\theta)}{q(\mathbf{Z})} \right] + \sum_{\mathbf{Z}} q(\mathbf{Z}) \log \left[ \frac{q(\mathbf{Z})}{p(\mathbf{Z}|\mathbf{X}, \theta)} \right]$$

Step 4: *Marginalise* over **Z** 

$$\sum_{\mathbf{Z}} q(\mathbf{Z}) \log p(\mathbf{X}|\theta) = \sum_{\mathbf{Z}} q(\mathbf{Z}) \log \left[ \frac{p(\mathbf{X}, \mathbf{Z}|\theta)}{q(\mathbf{Z})} \right] + \sum_{\mathbf{Z}} q(\mathbf{Z}) \log \left[ \frac{q(\mathbf{Z})}{p(\mathbf{Z}|\mathbf{X}, \theta)} \right]$$

$$\log p(\mathbf{X}|\theta) = \dots$$

Step 4: *Marginalise* over *Z* 

$$\log p(\mathbf{X}|\theta) = \sum_{\mathbf{Z}} q(\mathbf{Z}) \log \left[ \frac{p(\mathbf{X}, \mathbf{Z}|\theta)}{q(\mathbf{Z})} \right] + \sum_{\mathbf{Z}} q(\mathbf{Z}) \log \left[ \frac{q(\mathbf{Z})}{p(\mathbf{Z}|\mathbf{X}, \theta)} \right]$$

$$\log p(\mathbf{X}|\theta) = \sum_{\mathbf{Z}} q(\mathbf{Z}) \log \left[ \frac{p(\mathbf{X}, \mathbf{Z}|\theta)}{q(\mathbf{Z})} \right] + \sum_{\mathbf{Z}} q(\mathbf{Z}) \log \left[ \frac{q(\mathbf{Z})}{p(\mathbf{Z}|\mathbf{X}, \theta)} \right]$$

$$\log p(\mathbf{X}|\theta) = \sum_{\mathbf{Z}} q(\mathbf{Z}) \log \left[ \frac{p(\mathbf{X}, \mathbf{Z}|\theta)}{q(\mathbf{Z})} \right] + \sum_{\mathbf{Z}} q(\mathbf{Z}) \log \left[ \frac{q(\mathbf{Z})}{p(\mathbf{Z}|\mathbf{X}, \theta)} \right]$$

$$D_{KL}(q \mid\mid p) \triangleq \sum_{y} q(y) \log \frac{q(y)}{p(y)}$$

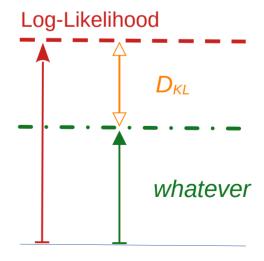
- *D<sub>KL</sub>* (*KL Divergence*) properties:
  - $\sim D_{KL}(q || p) \ge 0$
  - $VD_{KL}(q || p) = 0 \stackrel{\text{iff}}{\leftrightarrow} q = p$

$$\log p(\mathbf{X}|\theta) = \sum_{\mathbf{Z}} q(\mathbf{Z}) \log \left[ \frac{p(\mathbf{X}, \mathbf{Z}|\theta)}{q(\mathbf{Z})} \right] + \sum_{\mathbf{Z}} q(\mathbf{Z}) \log \left[ \frac{q(\mathbf{Z})}{p(\mathbf{Z}|\mathbf{X}, \theta)} \right]$$

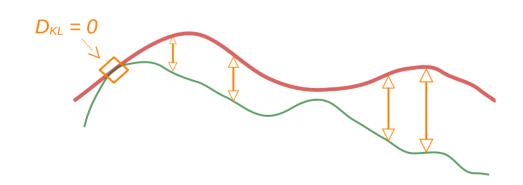
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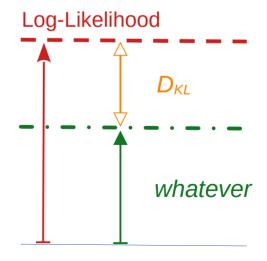
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$$\log p(\mathbf{X}|\theta) = \text{whatever} + \underbrace{D_{KL}(q(\mathbf{Z}) \mid\mid p(\mathbf{Z}|\mathbf{X}, \theta))}_{\geq 0}$$



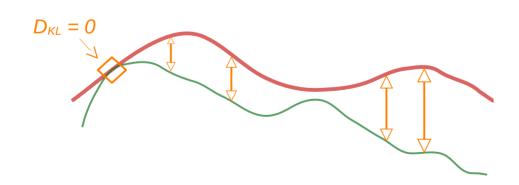
$$\log p(\mathbf{X}|\theta) = \text{whatever} + \underbrace{D_{KL}(q(\mathbf{Z}) \mid\mid p(\mathbf{Z}|\mathbf{X}, \theta))}_{\geq 0}$$

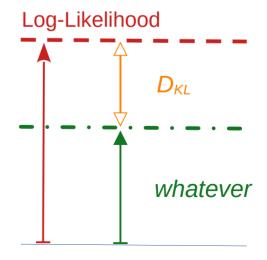




$$\log p(\mathbf{X}|\theta) = \text{whatever} + \underbrace{D_{KL}(q(\mathbf{Z}) \mid\mid p(\mathbf{Z}|\mathbf{X}, \theta))}_{\geq 0}$$

$$\log p(\mathbf{X}|\theta) \geq \text{whatever}$$



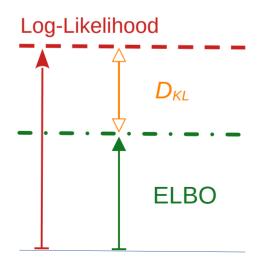


$$\log p(\mathbf{X}|\theta) = \text{ELBO} + \underbrace{D_{KL}(q(\mathbf{Z}) \mid\mid p(\mathbf{Z}|\mathbf{X}, \theta))}_{\geq 0}$$

$$\log p(\mathbf{X}|\theta) \geq \text{ELBO}$$

#### Evidence Lower Bound (ELBO)

Evidence ≡ log-likelihood



$$\log \sum_{\mathbf{Z}} p(\mathbf{X}, \mathbf{Z}) = \sum_{\mathbf{Z}} q(\mathbf{Z}) \log \left[ \frac{p(\mathbf{X}, \mathbf{Z} | \theta)}{q(\mathbf{Z})} \right] + \sum_{\mathbf{Z}} q(\mathbf{Z}) \log \left[ \frac{q(\mathbf{Z})}{p(\mathbf{Z} | \mathbf{X}, \theta)} \right]$$

$$ELBO(q, \theta)$$

$$ELBO(q, \theta)$$

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# EM Derivation – Step 6

$$\log \sum_{\mathbf{Z}} p(\mathbf{X}, \mathbf{Z}) = \sum_{\mathbf{Z}} q(\mathbf{Z}) \log \left[ \frac{p(\mathbf{X}, \mathbf{Z} | \theta)}{q(\mathbf{Z})} \right] + \sum_{\mathbf{Z}} q(\mathbf{Z}) \log \left[ \frac{q(\mathbf{Z})}{p(\mathbf{Z} | \mathbf{X}, \theta)} \right]$$

$$ELBO(q, \theta)$$

$$ELBO(q, \theta)$$

$$\theta_{ML}^* = \underset{\theta \in \Theta}{\operatorname{argmax}} \log \sum_{\mathbf{Z}} p(\mathbf{X}, \mathbf{Z} | \theta)$$

#### Direction solution does not work!

$$\log \sum_{z} p(x,z|\theta) = \log(... + w_{z_i}e^{\frac{(x-\mu_i)^2}{2\sigma_i^2}} + ... + w_{z_j}e^{\frac{(x-\mu_j)^2}{2\sigma_j^2}} + ...)$$

$$... = 0 \rightarrow \text{Intractable}$$

$$\sum_{z} \log p(x,z|\theta) = (... + w_i \log e^{\frac{(x-\mu_i)^2}{2\sigma_i^2}} + ... + w_j \log e^{\frac{(x-\mu_j)^2}{2\sigma_j^2}} + ...)$$

$$= (... + c_i(x-\mu_i)^2 + ... + c_j(x-\mu_j)^2 + ...)$$

# EM Derivation – Step 5

$$\underbrace{\log \sum_{\mathbf{Z}} p(\mathbf{X}, \mathbf{Z})}_{\log p(\mathbf{X}|\theta)} = \underbrace{\sum_{\mathbf{Z}} q(\mathbf{Z}) \log \left[ \frac{p(\mathbf{X}, \mathbf{Z}|\theta)}{q(\mathbf{Z})} \right]}_{ELBO(q,\theta)} + \underbrace{\sum_{\mathbf{Z}} q(\mathbf{Z}) \log \left[ \frac{q(\mathbf{Z})}{p(\mathbf{Z}|\mathbf{X}, \theta)} \right]}_{D_{KL}(q,\theta)}$$

#### Intractable!

 $\log \sum p(\mathbf{X}, \mathbf{Z}|\theta)$ 



$$\sum_{\mathbf{Z}} \log p(\mathbf{X}, \mathbf{Z} | \theta)$$

# EM Derivation – Step 6

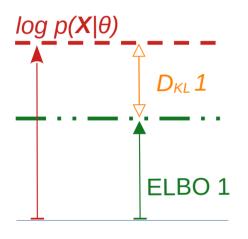
$$\frac{\log \sum_{\mathbf{Z}} p(\mathbf{X}, \mathbf{Z})}{\log p(\mathbf{X}|\theta)} = \underbrace{\sum_{\mathbf{Z}} q(\mathbf{Z}) \log \left[ \frac{p(\mathbf{X}, \mathbf{Z}|\theta)}{q(\mathbf{Z})} \right]}_{ELBO(q,\theta)} + \underbrace{\sum_{\mathbf{Z}} q(\mathbf{Z}) \log \left[ \frac{q(\mathbf{Z})}{p(\mathbf{Z}|\mathbf{X}, \theta)} \right]}_{D_{KL}(q,\theta)}$$

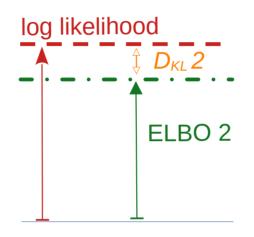
$$\theta^* = \underset{\theta \in \Theta}{\operatorname{argmax}} ELBO$$

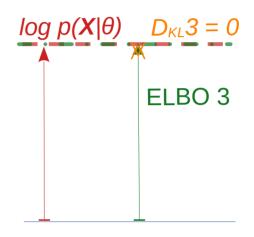


**EM** → Instead of log-likelihood ... maximise ELBO ...

### Best ELBO to optimise ...

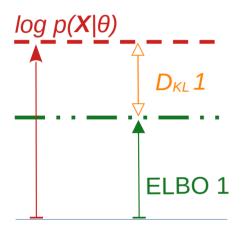


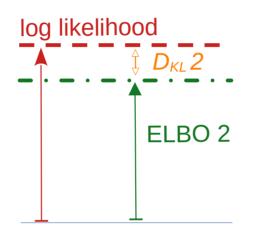


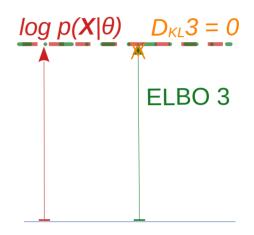


Which ELBO is better?

### Best ELBO to optimise ...

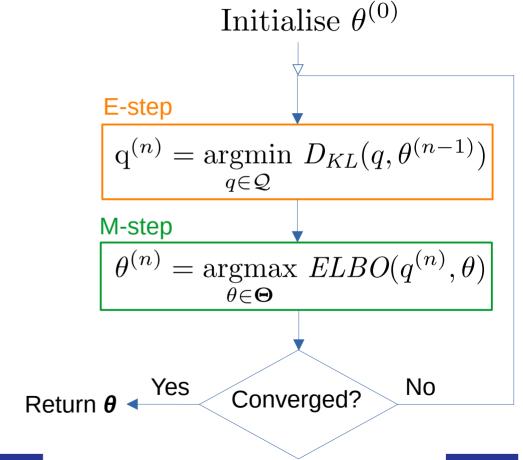






- Lower  $D_{KL} \rightarrow \text{Better ELBO (closer to } log p(X|\theta))$
- IDEAL:  $D_{KL} = 0 \rightarrow \text{Best ELBO} = \log p(X|\theta)$

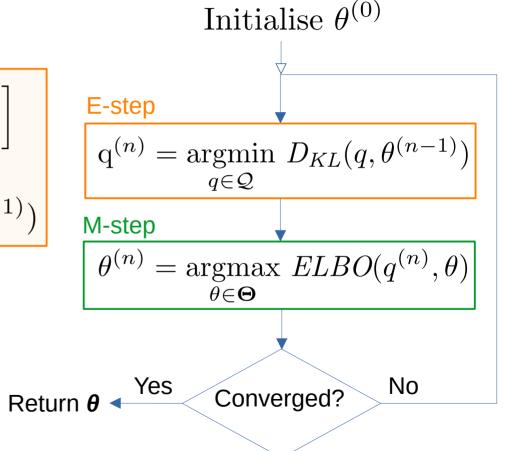
#### **EM Procedure**



### E-Step

$$D_{KL}(q, \theta^{(n-1)}) = \sum_{\mathbf{Z}} q(\mathbf{Z}) \log \left[ \frac{q(\mathbf{Z})}{p(\mathbf{Z}|\mathbf{X}, \theta^{(n-1)})} \right]$$

$$\operatorname{Min} D_{KL} = 0 \leftrightarrow q^{(n)}(\mathbf{Z}) = p(\mathbf{Z}|\mathbf{X}, \theta^{(n-1)})$$



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### M-Step

$$ELBO(q^{(n)}, \theta) = \sum_{\mathbf{Z}} q^{(n)}(\mathbf{Z}) \log \left[ \frac{p(\mathbf{X}, \mathbf{Z} | \theta)}{q^{(n)}(\mathbf{Z})} \right]$$
$$= \sum_{\mathbf{Z}} q^{(n)}(\mathbf{Z}) \log p(\mathbf{X}, \mathbf{Z} | \theta) - \sum_{\mathbf{Z}} q^{(n)}(\mathbf{Z}) \log q^{(n)}(\mathbf{Z})$$
constant

### M-Step

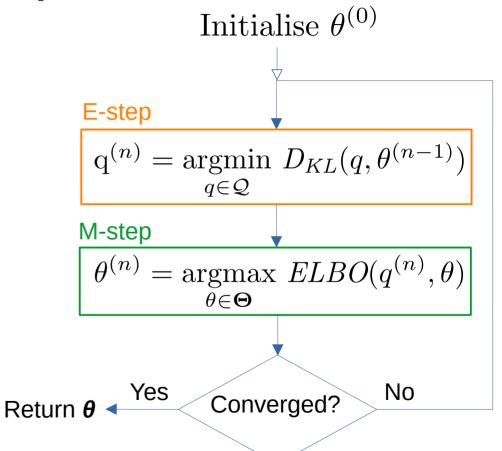
$$ELBO(q^{(n)}, \theta) = \sum_{\mathbf{Z}} q^{(n)}(\mathbf{Z}) \log \left[ \frac{p(\mathbf{X}, \mathbf{Z} | \theta)}{q^{(n)}(\mathbf{Z})} \right]$$
$$= \sum_{\mathbf{Z}} q^{(n)}(\mathbf{Z}) \log p(\mathbf{X}, \mathbf{Z} | \theta) - \sum_{\mathbf{Z}} q^{(n)}(\mathbf{Z}) \log q^{(n)}(\mathbf{Z})$$
constant

$$\theta^* = \underset{\theta \in \Theta}{\operatorname{argmax}} ELBO(q^{(n)}, \theta)$$
$$= \underset{\theta \in \Theta}{\operatorname{argmax}} \sum_{\mathbf{Z}} q^{(n)}(\mathbf{Z}) \log p(\mathbf{X}, \mathbf{Z} | \theta)$$

### M-Step

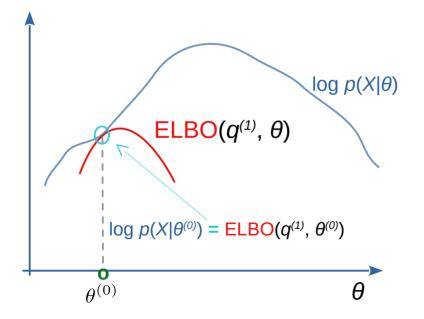
#### M-step

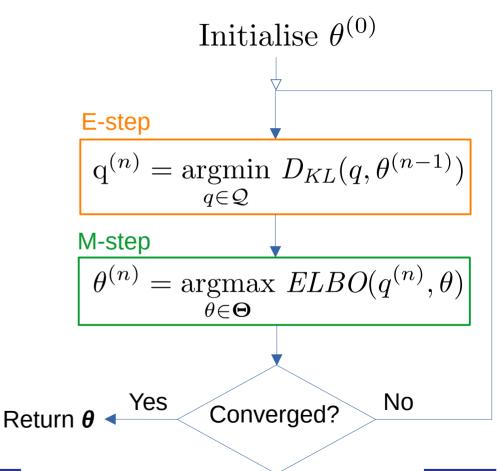
$$\theta^* = \underset{\theta \in \Theta}{\operatorname{argmax}} ELBO(q^{(n)}, \theta)$$
$$= \underset{\theta \in \Theta}{\operatorname{argmax}} \sum_{\mathbf{Z}} q^{(n)}(\mathbf{Z}) \log p(\mathbf{X}, \mathbf{Z} | \theta)$$



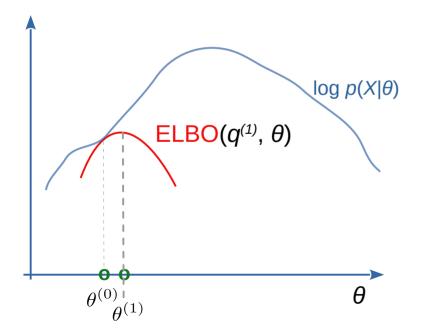
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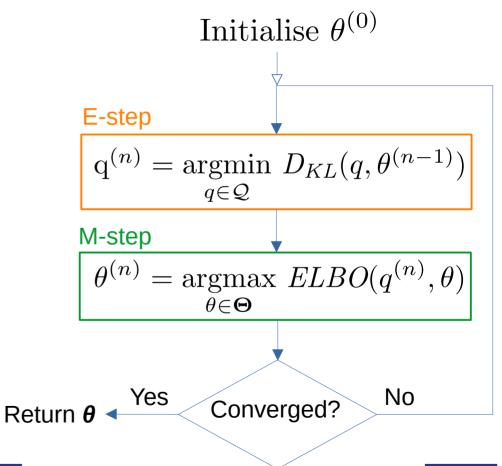
*E-step:*  $q^{(1)}$  = argmin  $D_{KL}(q, \theta^{(0)})$ 

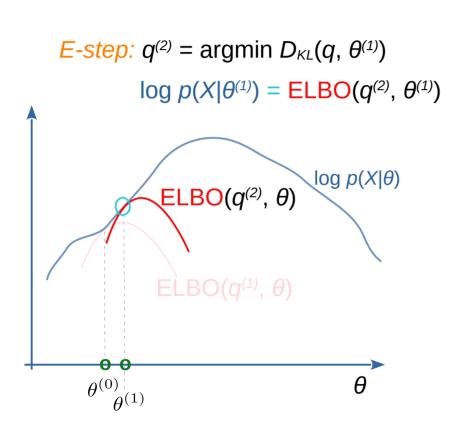




*M-step:*  $\theta^{(1)} = \operatorname{argmax} ELBO(q^{(1)}, \theta)$ 

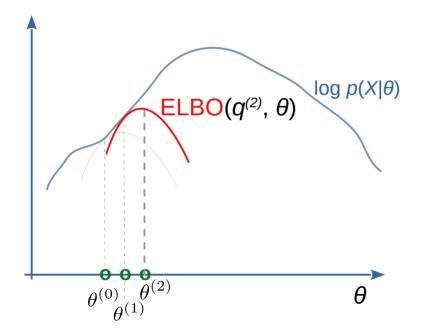






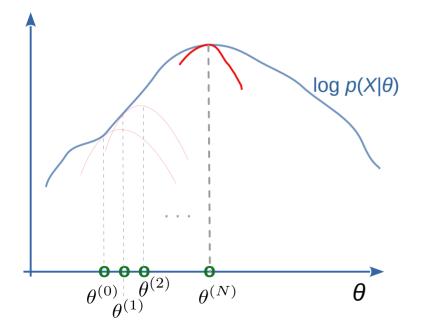


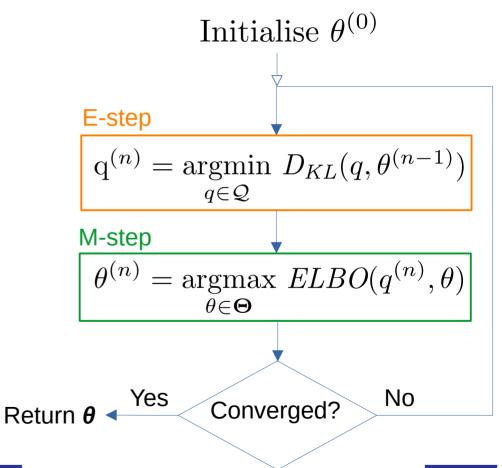
*M-step:*  $\theta^{(2)} = \operatorname{argmax} ELBO(q^{(2)}, \theta)$ 



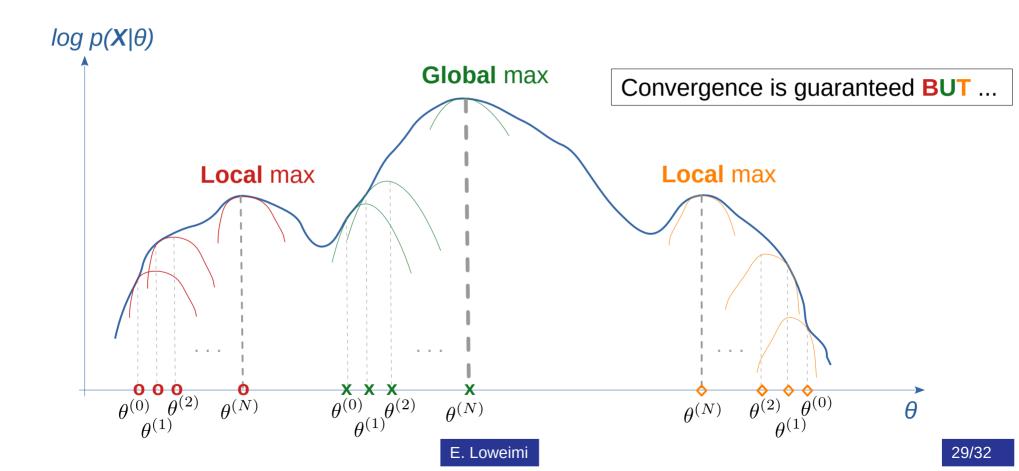


*M-step:*  $\theta^{(N)} = \operatorname{argmax} ELBO(q^{(N)}, \theta)$ 



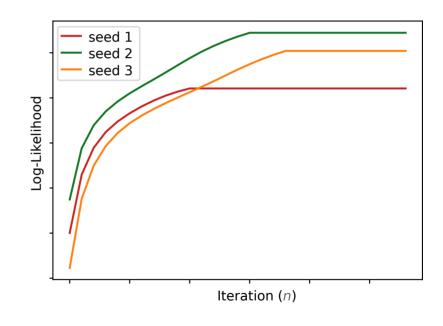


# Initialisation Matters (1)



# Initialisation Matters (2)

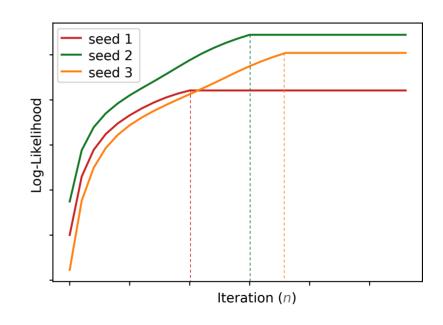
- $p(X|\theta^{(n)})$  is **ALWAYS** non-decreasing
  - $\checkmark ELBO*(q^{(n+1)}, \theta^{(n+1)}) \ge ELBO*(q^{(n)}, \theta^{(n)})$
  - Convergence is guaranteed ... but to local optimum ...



# Initialisation Matters (3)

•  $p(X|\theta^{(n)})$  is **ALWAYS** non-decreasing

 Initialisation affects ... convergence rate and final log-likelihood

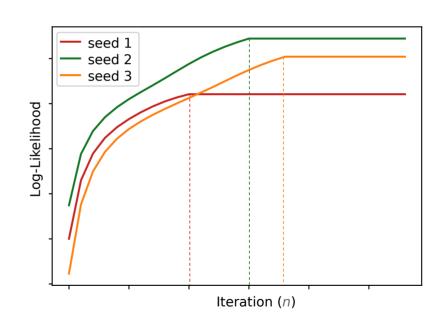


# Initialisation Matters (3)

•  $p(X|\theta^{(n)})$  is **ALWAYS** non-decreasing

 Initialisation affects ... convergence rate and final log-likelihood

Try multiple initialisations and pick up the best local optimum (highest log-likelihood).



#### Other Considerations

- EM is closely related to Coordinate Ascent [App B]
  - E-step: fix  $\theta$ , optimise q
  - M-step: fix  $\mathbf{q}$ , optimise  $\theta$
- EM shines when M-step can be solved analytically
- Alternatives when M-step is intractable ...
  - Generalised EM (GEM) → (conjugate) gradient ascent in M-Step
  - Expectation Conditional Maximisation (ECM) → coordinate ascent in M-step

### Wrap-up ... EM ...

- Goal: estimate  $\theta_{ML}$  for probabilistic models with latent var
- **How**: an iterative two-stage (E-step, M-step) procedure
- Applications: GMM, HMM, Computational biology, ...
- **Assignment**: estimate  $\theta_{MAP}$  using EM
- Appendices
  - (A) Further Reading
  - (B) Coordinate Ascent



# (A) Further Reading

- Murphy, Chapter 8, Section 7.2, Pages 306-310
- Bishop, Chapter 9, Section 4, Pages 450-455
- Andrew Ng's Lecture Notes, Chapter 11, Pages 142-147

Others: blog1 blog2

### (B) Coordinate Ascent

- Iterative optimisation method for multi-variate functions f(x)
- **Idea**: Maximise over <u>one</u> variable (or a block of variables) at a time, assuming others are constants

```
Procedure:

INITIALISE \mathbf{x}^{(0)} = [x_0, x_1, ..., x_{D-1}]

FOR i in range(#iterations):

FOR d in range(D):

\mathbf{x}^{(i)}[d] \leftarrow \operatorname{argmax} f(\mathbf{x}^{(i)}[:d], x_d, \mathbf{x}^{(i-1)}[d+1:])

X_d
```

