

# CDS: Machine Learning WK4-7

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## Week 6

### Graphical models

**a**

$$\begin{aligned} p(x_i|y) &= \frac{p(x_i, y)}{p(y)} \\ &= \frac{1}{p(y)} \sum_{x_1, \dots, x_{i-1}, x_{i+1}, \dots, x_n} p(x_1, x_2, \dots, x_n, y) \\ &= \frac{1}{p(y)} \sum_{x_1, \dots, x_{i-1}, x_{i+1}, \dots, x_n} p(x_1)p(x_2) \dots p(x_n)p(y|x_1, x_2, \dots, x_n) \end{aligned}$$

**b**

$$p(x_i|y, x_1 = 1, x_i = 1, i \neq 1) = \frac{1}{p(y)} \sum_{x_2, \dots, x_{i-1}, x_{i+1}, \dots, x_n} p(x_1)p(x_2) \dots p(x_n)p(y|x_1, x_2, \dots, x_n)$$

**a**

Mean is simple:

$$\mu = \begin{pmatrix} \mu_{x_1} \\ \mu_{x_2} \\ \mu_y \end{pmatrix} = \begin{pmatrix} 0 \\ 0 \\ \gamma(x_1 + x_2) \end{pmatrix}$$

Because  $x_1$  and  $x_2$  are independent, the upper left covariance matrix is just a diagonal matrix:

$$\Sigma_{x_1, x_2} = \begin{pmatrix} 1 & 0 \\ 0 & 1 \end{pmatrix}$$

$$\begin{aligned} \Sigma_{13} = \text{cov}(x_1, y) &= E(x_1 y) - E(x_1)E(y) = E(x_1)(\gamma(x_1 + x_2) + \chi) \\ &= E(x_1(\gamma(x_1 + x_2))) + E(x_1\chi) = E(\gamma x_1^2) + E(\gamma x_1 x_2) = \gamma \end{aligned}$$

As  $y$  depends in the same way on  $x_1$  and  $x_2$ , the covariances of  $y$  with either of them is the same. The variance of  $y$  is given by

$$\begin{aligned} E(y^2) - E(y)^2 &= E((\gamma x_1 + \gamma x_2 + \chi)^2) - E(\gamma(x_1 + x_2))^2 \\ &= E(\gamma^2 x_1^2 + 2\gamma^2 x_1 x_2 + 2\gamma x_1 \chi + \gamma^2 x_2^2 + 2\gamma x_2 \chi + \chi^2) - 0^2 \\ &= \gamma^2 E(x_1^2 + 2x_1 x_2 + x_2^2) + E(2\gamma x_1 \chi + 2\gamma x_2 \chi + \chi^2) \\ &= 2\gamma^2 + E(\chi^2) = 2\gamma^2 + \sigma^2 \end{aligned}$$

So we get for the covariance matrix  $\Sigma$ :

$$\Sigma = \begin{pmatrix} 1 & 0 & \gamma \\ 0 & 1 & \gamma \\ \gamma & \gamma & 2\gamma^2 + \sigma^2 \end{pmatrix}$$

Since the probability is a Gaussian:

$$p(x_1, x_2, y) = \mathcal{N}(x_1, x_2, y | \mu, \Sigma)$$

**b**

$$\mu_{\mathbf{x}_1, \mathbf{x}_2 | \mathbf{y}} = \begin{pmatrix} 0 \\ 0 \end{pmatrix} + \begin{pmatrix} \gamma \\ \gamma \end{pmatrix} \frac{\chi}{2\gamma^2 + \sigma^2} = \frac{\gamma \chi}{2\gamma^2 + \sigma^2} \begin{pmatrix} 1 \\ 1 \end{pmatrix}$$

$$\Sigma_{\mathbf{x}_1, \mathbf{x}_2 | \mathbf{y}} = \begin{pmatrix} 1 & 0 \\ 0 & 1 \end{pmatrix} - \begin{pmatrix} \gamma \\ \gamma \end{pmatrix} \frac{1}{2\gamma^2 + \sigma^2} (\gamma \quad \gamma) = \begin{pmatrix} 1 & 0 \\ 0 & 1 \end{pmatrix} - \frac{\gamma^2}{2\gamma^2 + \sigma^2} \begin{pmatrix} 1 & 1 \\ 1 & 1 \end{pmatrix} = \begin{pmatrix} 1 - \frac{\gamma^2}{2\gamma^2 + \sigma^2} & -\frac{\gamma^2}{2\gamma^2 + \sigma^2} \\ -\frac{\gamma^2}{2\gamma^2 + \sigma^2} & 1 - \frac{\gamma^2}{2\gamma^2 + \sigma^2} \end{pmatrix}$$

**c**

$$\rho = \frac{-\frac{\gamma^2}{2\gamma^2 + \sigma^2}}{1 - \frac{\gamma^2}{2\gamma^2 + \sigma^2}} = \frac{-\gamma^2}{2\gamma^2 + \sigma^2 - \gamma^2} = \frac{-\gamma^2}{\sigma^2 + \gamma^2}$$

## Week 7

### EM 2.3

1.

See python notebook

2.

Gaussian mixture model:

$$p(x, k) = \pi_k \mathcal{N}(x | a_k, \Sigma_k) \quad (1)$$

the logarithm of a Gaussian is

$$\log \mathcal{N}(x | a_k, \Sigma_k) = -\frac{(x - a_k)^2}{2\sigma_k^2} - \log \sigma_k - \frac{\log 2\pi}{2} \quad (3)$$

the responsibility is then

$$r_{\mu k} = \frac{\pi_k \mathcal{N}(x^\mu | \mu_k, \Sigma_k)}{\sum_{k'} \pi_{k'} \mathcal{N}(x^\mu | \mu_{k'}, \Sigma_{k'})} \quad (5)$$

the variational bound is then

$$Q(\theta, q^*) = L(\theta) = \sum_{\mu} \sum_k D(x) r_{\mu k} \log \frac{p(x, k | \theta)}{r_{\mu k}} \quad (7)$$

$$Q(\theta) = \sum_{\mu} D(x) \sum_k r_{\mu k} \log p(x, k | \theta) \quad (8)$$

$$Q(\theta) = \sum_{\mu} \sum_k r_{\mu k} (\log \pi_k + \log \mathcal{N}(x^\mu | \theta)) \quad (10)$$

next we approximate that the responsibility is a constant with regards to  $a_k$  and  $\sigma_k$ , use Lagrange multipliers and derive for a specific k

$$\sum_k P_k = 1 \quad (11)$$

$$Q(\theta_k) = \sum_{\mu} \sum_k r_{\mu k} (\log \pi_k + \log \mathcal{N}(x^\mu | \theta)) - \lambda (\sum_k \pi_k - 1) \quad (12)$$

$$\frac{\partial Q}{\partial \pi_k} = \sum_{\mu} \frac{r_{\mu k}}{\pi_k} - \lambda = 0 \quad (14)$$

$$\pi_k = \sum_{\mu} r_{\mu k} / \lambda \quad (15)$$

now we find  $\lambda$  by optimizing

(16)

$$\frac{\partial Q}{\partial \lambda} = - \sum_k \pi_k + 1 = 0 \quad (17)$$

$$\sum_k \pi_k = 1 = \sum_k \sum_\mu r_{\mu k} / \lambda \quad (18)$$

$$\lambda = \sum_k \sum_\mu r_{\mu k} \quad (19)$$

$$\pi_k = \frac{\sum_\mu r_{\mu k}}{\sum_k \sum_\mu r_{\mu k}} \quad (20)$$

$$r_k \equiv \sum_\mu r_{\mu k} \quad (21)$$

$$N \equiv \sum_k r_k \quad (22)$$

$$\pi_k = \frac{r_k}{N} \quad (23)$$

now for  $a_k$

$$\frac{\partial Q}{\partial a_k} = \sum_\mu r_{\mu k} \frac{x^\mu - a_k}{\sigma_k^2} = 0 \quad (24)$$

$$a_k \sum_\mu r_{\mu k} = \sum_\mu r_{\mu k} x^\mu \quad (25)$$

$$a_k = \frac{\sum_\mu r_{\mu k} x^\mu}{\sum_\mu r_{\mu k}} \quad (26)$$

$$a_k = \frac{\sum_\mu r_{\mu k} x^\mu}{r_k} \quad (27)$$

now for  $\sigma_k$

(28)

$$\frac{\partial Q}{\partial \sigma_k} = \sum_{\mu} r_{\mu k} \left( \frac{(x^{\mu} - a_k)^2}{\sigma_k^3} - \frac{1}{\sigma_k} \right) = 0 \quad (29)$$

$$\frac{\partial L}{\partial \sigma_k} = \sum_{\mu} r_{\mu k} (x^{\mu} - a_k)^2 - r_{\mu k} \sigma_k^2 = 0 \quad (30)$$

$$\sigma_k^2 \sum_{\mu} r_{\mu k} = \sum_{\mu} r_{\mu k} (x^{\mu} - a_k)^2 \quad (31)$$

$$\sigma_k^2 = \frac{\sum_{\mu} r_{\mu k} (x^{\mu} - a_k)^2}{\sum_{\mu} r_{\mu k}} \quad (32)$$

$$\sigma_k^2 = \frac{\sum_{\mu} r_{\mu k} (x^{\mu} - a_k)^2}{r_k} \quad (33)$$

now we fill in for  $a_k$

(34)

$$\sigma_k^2 = \frac{\sum_{\mu} r_{\mu k} (x^{\mu 2} - 2x^{\mu} a_k + a_k^2)}{r_k} \quad (35)$$

$$\sigma_k^2 = \frac{\sum_{\mu} r_{\mu k} x^{\mu 2}}{r_k} - \frac{\sum_{\mu} 2r_{\mu k} x^{\mu} a_k}{r_k} + \frac{\sum_{\mu} r_{\mu k} a_k^2}{r_k} \quad (36)$$

$$\sigma_k^2 = \frac{\sum_{\mu} r_{\mu k} x^{\mu 2}}{r_k} - 2a_k^2 + a_k^2 \quad (37)$$

$$\sigma_k^2 = \frac{\sum_{\mu} r_{\mu k} x^{\mu 2}}{r_k} - a_k^2 \quad (38)$$

This makes sense compared to the multidimensional case from the slides, when  $i == j$  we get the previous result.