

Uniformly most powerful test

In statistical hypothesis testing, a **uniformly most powerful (UMP) test** is a hypothesis test which has the **greatest power** $1 - \beta$ among all possible tests of a given size α . For example, according to the Neyman–Pearson lemma, the likelihood-ratio test is UMP for testing simple (point) hypotheses.

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Setting

Let \mathbf{X} denote a random vector (corresponding to the measurements), taken from a parametrized family of probability density functions or probability mass functions $f_{\boldsymbol{\theta}}(\mathbf{x})$, which depends on the unknown deterministic parameter $\boldsymbol{\theta} \in \boldsymbol{\Theta}$. The parameter space $\boldsymbol{\Theta}$ is partitioned into two disjoint sets $\boldsymbol{\Theta}_0$ and $\boldsymbol{\Theta}_1$. Let H_0 denote the hypothesis that $\boldsymbol{\theta} \in \boldsymbol{\Theta}_0$, and let H_1 denote the hypothesis that $\boldsymbol{\theta} \in \boldsymbol{\Theta}_1$. The binary test of hypotheses is performed using a test function $\varphi(\mathbf{x})$ with a reject region R (a subset of measurement space).

$$\varphi(\mathbf{x}) = \begin{cases} 1 & \text{if } \mathbf{x} \in R \\ 0 & \text{if } \mathbf{x} \in R^c \end{cases}$$

meaning that H_1 is in force if the measurement $\mathbf{X} \in R$ and that H_0 is in force if the measurement $\mathbf{X} \in R^c$. Note that $R \cup R^c$ is a disjoint covering of the measurement space.

Formal definition

A test function $\varphi(\mathbf{x})$ is UMP of size α if for any other test function $\varphi'(\mathbf{x})$ satisfying

$$\sup_{\boldsymbol{\theta} \in \boldsymbol{\Theta}_0} \mathbb{E}[\varphi'(\mathbf{X})|\boldsymbol{\theta}] = \alpha' \leq \alpha = \sup_{\boldsymbol{\theta} \in \boldsymbol{\Theta}_0} \mathbb{E}[\varphi(\mathbf{X})|\boldsymbol{\theta}]$$

we have

$$\forall \boldsymbol{\theta} \in \boldsymbol{\Theta}_1, \quad \mathbb{E}[\varphi'(\mathbf{X})|\boldsymbol{\theta}] = 1 - \beta'(\boldsymbol{\theta}) \leq 1 - \beta(\boldsymbol{\theta}) = \mathbb{E}[\varphi(\mathbf{X})|\boldsymbol{\theta}].$$

The Karlin–Rubin theorem

The Karlin–Rubin theorem can be regarded as an extension of the Neyman–Pearson lemma for composite hypotheses.^[1] Consider a scalar measurement having a probability density function parameterized by a scalar parameter θ , and define the likelihood ratio $l(x) = f_{\theta_1}(x)/f_{\theta_0}(x)$. If $l(x)$ is monotone non-decreasing, in x , for any pair $\theta_1 \geq \theta_0$ (meaning that the greater x is, the more likely H_1 is), then the threshold test:

$$\varphi(x) = \begin{cases} 1 & \text{if } x > x_0 \\ 0 & \text{if } x < x_0 \end{cases}$$

where x_0 is chosen such that $\mathbf{E}_{\theta_0} \varphi(X) = \alpha$

is the UMP test of size α for testing $H_0 : \theta \leq \theta_0$ vs. $H_1 : \theta > \theta_0$.

Note that exactly the same test is also UMP for testing $H_0 : \theta = \theta_0$ vs. $H_1 : \theta > \theta_0$.

Important case: exponential family

Although the Karlin-Rubin theorem may seem weak because of its restriction to scalar parameter and scalar measurement, it turns out that there exist a host of problems for which the theorem holds. In particular, the one-dimensional exponential family of probability density functions or probability mass functions with

$$f_{\theta}(x) = g(\theta)h(x) \exp(\eta(\theta)T(x))$$

has a monotone non-decreasing likelihood ratio in the sufficient statistic $T(x)$, provided that $\eta(\theta)$ is non-decreasing.

Example

Let $\mathbf{X} = (\mathbf{X}_0, \dots, \mathbf{X}_{M-1})$ denote i.i.d. normally distributed N -dimensional random vectors with mean $\theta \mathbf{m}$ and covariance matrix \mathbf{R} . We then have

$$\begin{aligned} f_{\theta}(\mathbf{X}) &= (2\pi)^{-MN/2} |\mathbf{R}|^{-M/2} \exp \left\{ -\frac{1}{2} \sum_{n=0}^{M-1} (\mathbf{X}_n - \theta \mathbf{m})^T \mathbf{R}^{-1} (\mathbf{X}_n - \theta \mathbf{m}) \right\} \\ &= (2\pi)^{-MN/2} |\mathbf{R}|^{-M/2} \exp \left\{ -\frac{1}{2} \sum_{n=0}^{M-1} (\theta^2 \mathbf{m}^T \mathbf{R}^{-1} \mathbf{m}) \right\} \\ &\quad \exp \left\{ -\frac{1}{2} \sum_{n=0}^{M-1} \mathbf{X}_n^T \mathbf{R}^{-1} \mathbf{X}_n \right\} \exp \left\{ \theta \mathbf{m}^T \mathbf{R}^{-1} \sum_{n=0}^{M-1} \mathbf{X}_n \right\} \end{aligned}$$

which is exactly in the form of the exponential family shown in the previous section, with the sufficient statistic being

$$T(\mathbf{X}) = \mathbf{m}^T \mathbf{R}^{-1} \sum_{n=0}^{M-1} \mathbf{X}_n.$$

Thus, we conclude that the test

$$\varphi(T) = \begin{cases} 1 & T > t_0 \\ 0 & T < t_0 \end{cases} \quad \mathbb{E}_{\theta_0} \varphi(T) = \alpha$$

is the UMP test of size α for testing $H_0 : \theta \leq \theta_0$ vs. $H_1 : \theta > \theta_0$

Further discussion

Finally, we note that in general, UMP tests do not exist for vector parameters or for two-sided tests (a test in which one hypothesis lies on both sides of the alternative). The reason is that in these situations, the most powerful test of a given size for one possible value of the parameter (e.g. for θ_1 where $\theta_1 > \theta_0$) is different from the most powerful test of the same size for a different value of the parameter (e.g. for θ_2 where $\theta_2 < \theta_0$). As a result, no test is **uniformly** most powerful in these situations.

References

1. Casella, G.; Berger, R.L. (2008), *Statistical Inference*, Brooks/Cole. ISBN 0-495-39187-5 (Theorem 8.3.17)

Further reading

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