



Untitled

ARCHIVE

Random Forest

Data Description (from R):

This famous (Fisher's or Anderson's) iris data set gives the measurements in centimeters of the variables sepal length and width and petal length and width, respectively, for 50 flowers from each of 3 species of iris. The species are Iris setosa, versicolor, and virginica. Out of these species, setosa was ignored to convert it into binary classification problem.

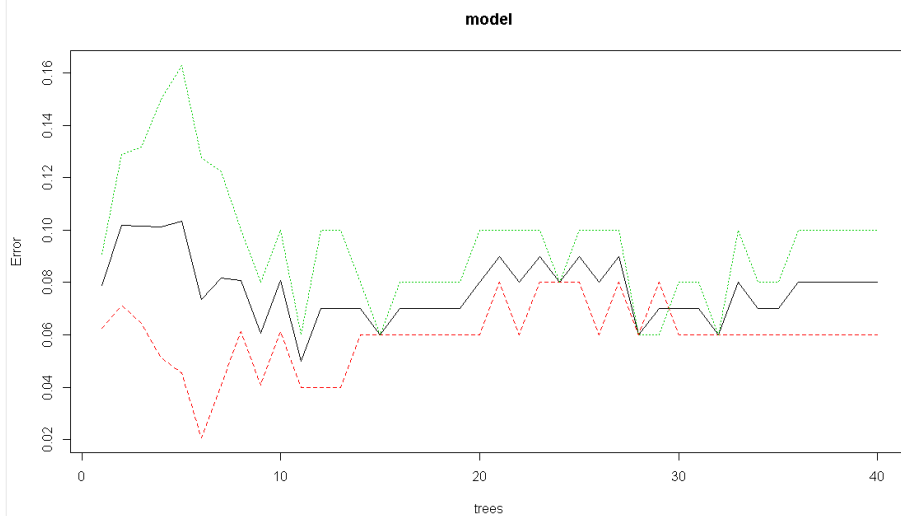
Variable Description:

Iris data set comprises of the following variables"

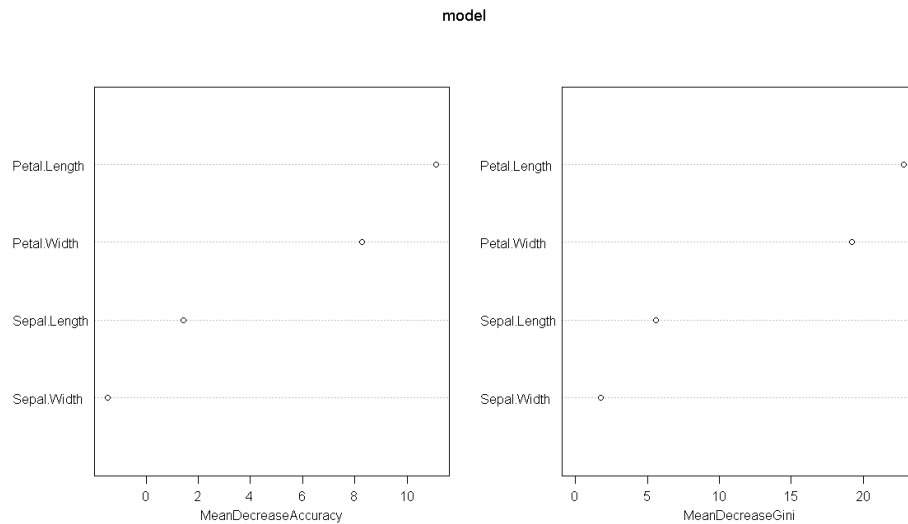
Independent variables: Sepal.Length, Sepal.Width, Petal.Length, Petal.Width

Dependent variable: Species

Error Rate vs Number of Trees:



Variable Importance:



Code:

```
data(iris)
iris <- iris[iris$Species!="setosa", ]
iris$Species <- as.factor(as.character(iris$Species))
library(randomForest)
library(SDMTools)
set.seed(123)
model <- randomForest(formula = Species ~ ., data=iris, ntree=40, importance=T)
plot(model)
prediction <- predict(model, iris)
actual <- rep(1, nrow(iris))
actual[iris$Species=="virginica"] <- 0
pred <- rep(1, nrow(iris))
pred[prediction=="virginica"] <- 0
confusion.matrix(obs = actual, pred = pred)
varImpPlot(model)
```

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Classification Results:

Key: 1="versicolor", 0="virginica"

```
obs
pred 0 1
0 50 0
1 0 50
```

Discussion:

Random forest model builds collection of decision trees, each decision tree built using randomly selected set of variables as predictors. The results (votes or probabilities) of all trees are averaged to reduce the variance of the prediction. Hence it can be used for variable selection.

In case of iris data, the classification accuracy in training data is 100%, which is an improvement to decision tree model (96%). As in decision tree model, order of predictor importance is Petal.Length > Petal.Width > (Sepal.Length > Sepal.Width) - variables in bracket were not used by decision tree algorithm. Least error rate is observed when number of trees is more than 1. Hence growing a forest is more advantageous.

Feb 8th, 2016

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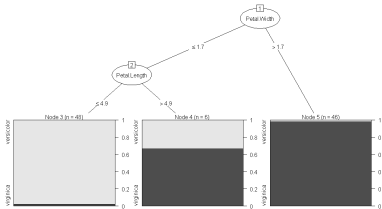
Data Description (from R):

This famous (Fisher's or Anderson's) iris data set gives the measurements in centimeters of the variables sepal length and width and petal length and width, respectively, for 50 flowers from each of 3 species of iris. The species are Iris setosa, versicolor, and virginica. Out of these species, setosa was ignored to convert it into binary classification problem.

Variable Description:

Iris data set comprises of the following variables”
Independent variables: Sepal.Length, Sepal.Width, Petal.Length, Petal.Width
Dependent variable: Species

Classification Tree:



Code (written in R v 3.2.3):

```
data(iris)
iris <- iris[iris$Species!="setosa", ]
library(C50)
set.seed(1)
model <- C5.0(formula = Species ~ .,
data=iris)
plot(model)
print(summary(model))
```

Discussion:

Classification tree was built in order to check for non-linearity of relationship between outcome variable (species) and predictors. (Maximum) Information gain criteria is computed in each step for generating the binary splits in the tree.

The training set has highest entropy $(-0.5 \cdot \log_2(0.5) - 0.5 \cdot \log_2(0.5) = 1)$. Hence the first iteration will always be an improvement, irrespective of how close the probabilities are to 0.5. However, the first iteration yielded good classification. Out of 46 cases in the right leaf, 1 is versicolor and 45 are virginica for $\text{Petal.Width} > 1.7$. Out of 54 cases on the left leaf, 5 are virginica and 49 are versicolor for $\text{Petal.Width} \leq 1.7$.

Admission Dataset

Data Description:

The data was provided in UCLA's website. It contains data on 400 applications to the university with the objective to study the likelihood of getting admission in graduate program at UCLA. **Hypothesis:** Likelihood of admission depends on GRE score, GPA and rank in undergraduate program.

Variable Descriptions:

admit Binary - whether or not the candidate got admission
gre Candidate's GRE score (average of section scores)
gpa Candidate's undergraduation GPA taken with common denominator of 4
rank Candidate's undergraduation rank

Program and Output (Analysis performed in R version 3.2.3):

Road Tests Dataset

Data Description (from R documentation):

The data was extracted from the 1974 Motor Trend US magazine, and comprises fuel consumption and 10 aspects of automobile design and performance for 32 automobiles (1973–74 models).

Variable Descriptions (from R documentation):

The dataset is in data frame form with 11 variables described below:
mpg Miles/(US) gallon
cyl Number of cylinders
displacement Displacement (cu.in.)
hp Gross horsepower
drat Rear axle ratio
wt Weight (1000 lbs)
qsec 1/4 mile time
vs V/S
am Transmission (0 = automatic, 1 = manual)
gear Number of forward gears
carb Number of carburetors
Out of these, only mpg (outcome), hp, wt and gear (independent) were chosen for analysis based on correlation.

Program and Output (Analysis performed in R version 3.2.3):

Dataset

Data Description

This dataset is or to be used for line demonstrates sin It was used by st study the height c height of parent (i mother and father complications). T observations = 9; concept known a mediocrity'.

Variable Descrip

There are two vai variable - height c with mean = 68.0 standard deviation and independent parent (continuou 68.30819 units ar = 1.787333 units. categorical variat

Program and Ou performed in R :

```
library(Usingf
data(galton)
mean(galton$ch
sd(galton$chil
mean(galton$pa
sd(galton$pare
galton$parent=
nt)
galton$child=s
)
model=lm(chilc
data=galton)
summary(model)
```

```
Call:
lm(formula = child ~ parent, data = gals)

Residuals:
    Min       3Q   Max
-3.09976 -0.54889  0.64889  2.35111

Coefficients:
(Intercept)      parent 
  15.72222    -0.45875 

Error t-value
parent    0.45875
15.72    <2e-16

---
Signif. codes:
  '***' 0.01 '**' 0.05 '*' 0.1 '.' 0.1 '1'

Residual standard error: 1.097 on 927 degrees of freedom
Multiple R-squared:  0.0008675
Adjusted R-squared:  -0.0008675
F-statistic: 2.167 on 1 and 927 DF, p-value: 0.1425
```

```

#           Estimate Std.
Error z value Pr(>|z|)
#(Intercept) -0.8098
0.1120 -7.233 4.74e-13 ***
#gre          0.3108
0.1222  2.544  0.0109 *
#gpa          0.2872
0.1216  2.361  0.0182 *
#---
#Signif. codes:  0 '***' 0.001
'***' 0.01 '**' 0.05 '.' 0.1 ' '
1
#
#(Dispersion parameter for
binomial family taken to be 1)
#
# Null deviance: 499.98 on
399 degrees of freedom
#Residual deviance: 480.34 on
397 degrees of freedom
#AIC: 486.34
#
#Number of Fisher Scoring
iterations: 4
#
> model2 <- glm(admit ~ gre +
gpa + rank, data=data,
family=binomial)
> summary(model2)
#
#Call:
#glm(formula = admit ~ gre +
gpa + rank, family = binomial,
data = data)
#
#Deviance Residuals:
#   Min       1Q   Median
       3Q      Max
#-1.4290 -0.8902 -0.6552
 1.1937  2.1122
#
#Coefficients:
#           Estimate Std.
Error z value Pr(>|z|)
#(Intercept) -0.4109
0.1447 -2.839  0.00453 **
#gre          0.2633
0.1254  2.101  0.03568 *
#gpa          0.3288
0.1247  2.638  0.00834 **
#rank1        -0.9383
0.2329 -4.028 5.62e-05 ***
#---
#Signif. codes:  0 '***' 0.001
'***' 0.01 '**' 0.05 '.' 0.1 ' '
1
#
#(Dispersion parameter for
binomial family taken to be 1)
#
# Null deviance: 499.98 on
399 degrees of freedom
#Residual deviance: 463.37 on
396 degrees of freedom
#AIC: 471.37
#
#Number of Fisher Scoring

```

```

Residual standard error: 0.6305
on 31 degrees of freedom
Multiple R-squared:  0.6024,
Adjusted R-squared:  0.5896
F-statistic: 46.98 on 1 and 31
DF, p-value: 1.11e-07

```

```

>
model2=lm(mtcars$mpg~indep$hp+i
ndep$wt-1)
> summary(model2)

```

```

Call:
lm(formula = mtcars$mpg ~
indep$hp + indep$wt - 1)

```

```

Residuals:
      Min       1Q   Median       3Q      Max
-0.6539 -0.2655 -0.0302  0.1742
 0.9713

```

```

Coefficients:
           Estimate Std. Error t
value Pr(>|t|)
indep$hp  -0.3615      0.1010
-3.579   0.0012 **
indep$wt  -0.6296      0.1010
-6.233 7.27e-07 ***
---
Signif. codes:  0 '***' 0.001
'***' 0.01 '**' 0.05 '.' 0.1 ' '
1

```

```

Residual standard error: 0.4231
on 30 degrees of freedom
Multiple R-squared:  0.8268,
Adjusted R-squared:  0.8152
F-statistic: 71.6 on 2 and 30
DF, p-value: 3.791e-12

```

```

>
model3=lm(mtcars$mpg~indep$hp+i
ndep$wt+indep$gear-1)
> summary(model3)

```

```

Call:
lm(formula = mtcars$mpg ~
indep$hp + indep$wt +
indep$gear -
1)

```

```

Residuals:
      Min       1Q   Median       3Q      Max
-0.55936 -0.31554 -0.05714
 0.16399  1.00640

```

```

Coefficients:
           Estimate Std. Error
t value Pr(>|t|)
indep$hp  -0.4185      0.1106
-3.785 0.000715 ***
indep$wt  -0.5192      0.1350
-3.844 0.000609 ***
indep$gear  0.1249      0.1024

```

be tall, but not as
Similarly child of
expected to be st
as the parent. Thi
as regression to i
behavior).

```

iterations: 3
#
> # Confidence intervals using
standard errors of log
likelihoods
> confint.default(model2)
#                2.5 %
97.5 %
#(Intercept) -0.69459942
-0.1272177
#gre          0.01761952
0.5090337
#gpa          0.08450421
0.5731589
#rank1        -1.39488233
-0.4817514
> # Wald test for significance
of effect of rank
> wald.test(b=coef(model2),
Sigma=vcov(model2), Terms=4)
#Wald test:
#-----
#
#Chi-squared test:
#X2 = 16.2, df = 1, P(> X2) =
5.6e-05

```

Interpretation of Output:

All quantitative variable have been normalized $((X - \text{mean})/\text{sd})$. Rank group '0' is treated as base group and the log likelihood coefficients are reported with respect to group '0'.

Let us consider p-value > 0.05 for not rejecting the null hypothesis in model 1. Intercept is significant and its estimate is -0.8098. gre is significant and its coefficient is estimated to be 0.3108. gpa is significant and its coefficient is estimated to be 0.2872. However, there is evidence of 'rank' being a confounder to model 1.

Analysis of model 2 suggests that intercept is significant with estimate = -0.4109 (CI= -0.695 to -0.127, p-value=0.00453). gre is significant with coefficient estimate = 0.2633 (unit increase in scaled gre increases log likelihood by 0.2633, CI=0.01762 to 0.509, p-value=0.0357). gpa is significant with coefficient estimate = 0.3288 (unit increase in scaled gre increases log likelihood by 0.3288, CI=0.0845 to 0.5732, p-value=0.008). rank1 is significant with coefficient estimate = -0.9383 (moving from rank group 0 to rank group 1 decrease log likelihood by 0.9383, CI= -1.395 to -0.482, p-value=5.62e-05)

Confounding:

```

1.219 0.232593
---
Signif. codes:  0 '***' 0.001
***' 0.01 '**' 0.05 '.' 0.1 ' '
1

Residual standard error: 0.4197
on 29 degrees of freedom
Multiple R-squared:  0.8352,
Adjusted R-squared:  0.8182
F-statistic: 49 on 3 and 29
DF, p-value: 1.788e-11

> plot(model3)

```

Interpretation of Output:

All variables - dependent and independent - have been scaled. Intercept has been explicitly removed from linear model call as the estimate will pass through (X-avg, Y-avg) for quantitative X irrespective of the number of independent variables in X. Independent variables are correlated, but the magnitude of correlations is not very high. However, this introduces the possibility of confounding.

The first concept to be checked is the significance of the model. p-value(F-stat)< 1.788e-11, which is less than 0.05 (chosen as cutoff for most significance tests). Therefore, the regression model is significant. It explains 83.52% of the variance in the outcome variable. There are 32 observations out of which 3 have been used for model building. Hence degrees of freedom for error in ANOVA = 32-3 = 29.

The independent variable 'hp' has p-value(t-stat) [two sided test] = 0.000715, which is less than 0.05. Therefore, the variable 'hp' is significant. The coefficient is estimated to be -0.4185. The independent variable 'wt' has p-value(t-stat) [two sided test] = 0.000609, which is less than 0.05. Therefore, the variable 'wt' is significant. The coefficient is estimated to be -0.5192. The independent variable 'gear' has p-value(t-stat) [two sided test] = 0.232593, which is greater than 0.05. Therefore, the variable '**gear**' is **insignificant**. The coefficient is estimated to be 0.1249.

Confounding:

Adding rank to model 1 proves that rank is a significant predictor. Wald test of significance shows that rank is a significant independent variable to explain the dependent variable. Hence confounding is present in model 1.

Discussion:

Let us consider a candidate with rank group = 0, scaled gre score = 1 and scaled gpa = 1. Let p be the probability of getting admission as predicted by the model

$\log_e(p/(1-p)) = 0.1812496 \Rightarrow p = 1.198714341/2.198714341$
 $= 0.545188758$ = probability of getting admission as predicted by the model

There is clear association between the predictor and outcome variables. Hence the hypothesis has been validated. Higher gpa and higher gre score are desirable. Worse rank group is not desirable.

Show more

```
>
model14=lm(mtcars$mpg~indep$gear
-1)
> summary(model14)
```

Call:

```
lm(formula = mtcars$mpg ~
indep$gear - 1)
```

Residuals:

Min	1Q	Median	3Q	Max
-1.69904	-0.46347	-0.03401	0.35272	2.08784

Coefficients:

	Estimate	Std. Error	t value	Pr(> t)
indep\$gear	0.4803	0.1575	3.049	0.00467 **

Signif. codes: 0 '***' 0.001
'**' 0.01 '*' 0.05 '.' 0.1 ' ' 1

Residual standard error: 0.8771
on 31 degrees of freedom
Multiple R-squared: 0.2307,
Adjusted R-squared: 0.2059
F-statistic: 9.295 on 1 and 31
DF, p-value: 0.004672

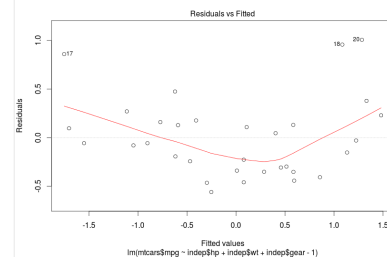
Discussion:

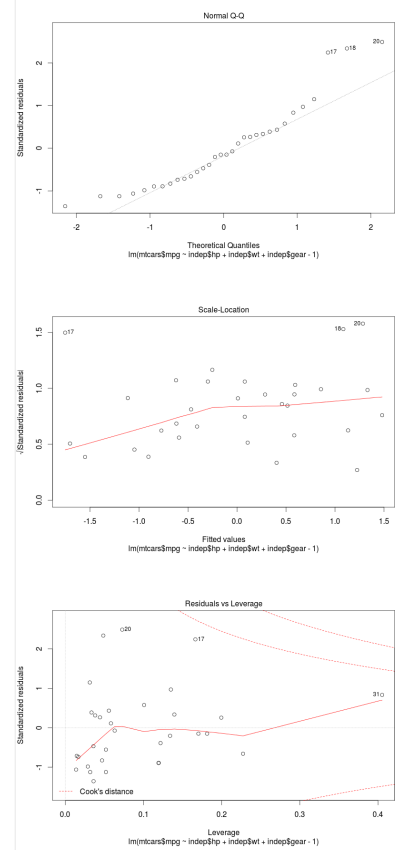
$\text{mpg}(\text{est}) = -0.4185 \cdot \text{hp} - 0.5192 \cdot \text{wt} + 0.1249 \cdot \text{gear}$
p-values = (0.000715, 0.000609, 0.232593)

There is clear (significant) association between dependent variable and independent variables.

There is evidence of confounding as 'gear' is correlated with both independent (hp and wt) and dependent (mpg) variables and is insignificant (p-value > 0.05).

Diagnostic Plots:





Points 17, 18 and 20 deviate significantly from model behavior ("extreme point w.r.t y"). Additionally, point 31 has unusually high leverage with respect to predictors ("extreme point w.r.t x").