

[Unit 4: Continuous Random](#)[Course](#) > [Variables](#)> [4.1 Reading](#) > 4.4 Normal

4.4 Normal

Unit 4: Continuous Random Variables

Adapted from Blitzstein-Hwang Chapter 5.

The Normal distribution is a famous continuous distribution with a bell-shaped PDE. It is extremely widely used in statistics because of a theorem, the *central limit theorem*, which says that under very weak assumptions, the sum of a large number of i.i.d. random variables has an approximately Normal distribution, *regardless* of the distribution of the individual r.v.s. This means we can start with independent r.v.s from almost any distribution, discrete or continuous, but once we add up a bunch of them, the distribution of the resulting r.v. looks like a Normal distribution.

DEFINITION 4.4.1 (STANDARD NORMAL DISTRIBUTION).

A continuous r.v. Z is said to have the *standard Normal distribution* if its PDF φ is given by

$$\varphi(z) = \frac{1}{\sqrt{2\pi}} e^{-z^2/2}, \quad -\infty < z < \infty.$$

We write this as $Z \sim \mathcal{N}(0, 1)$.

The constant $\frac{1}{\sqrt{2\pi}}$ in front of the PDF may look surprising (why is something with π needed in front of something with e , when there are no circles in sight?), but it turns out to be what is needed to make the PDF integrate to 1. Such constants are called *normalizing constants* because they normalize the total area under the PDF to 1. The standard Normal CDF Φ is the accumulated area under the PDF:

$$\Phi(z) = \int_{-\infty}^z \varphi(t) dt = \int_{-\infty}^z \frac{1}{\sqrt{2\pi}} e^{-t^2/2} dt.$$

We need to leave this as an integral: it turns out to be mathematically impossible to find a closed-form expression for the antiderivative of φ , meaning that we cannot express Φ as a finite sum of more familiar functions like polynomials or exponentials. But closed-form or no, it's still a well-defined function: if we give Φ an input z , it returns the accumulated area under the PDF from $-\infty$ up to z .

Notation 4.4.2.

By convention, we use φ for the standard Normal PDF and Φ for the standard Normal CDF. We will often use Z to denote a standard Normal random variable. The standard Normal PDF and CDF are plotted in Figure 4.4.3. The PDF is bell-shaped and symmetric about 0, and the CDF is S-shaped. These have the same general shape as the Logistic PDF and CDF that we saw in a couple of previous examples, but the Normal PDF decays to 0 much more quickly: notice that nearly all of the area under φ is between -3 and 3 , whereas we had to go out to -5 and 5 for the Logistic PDF.

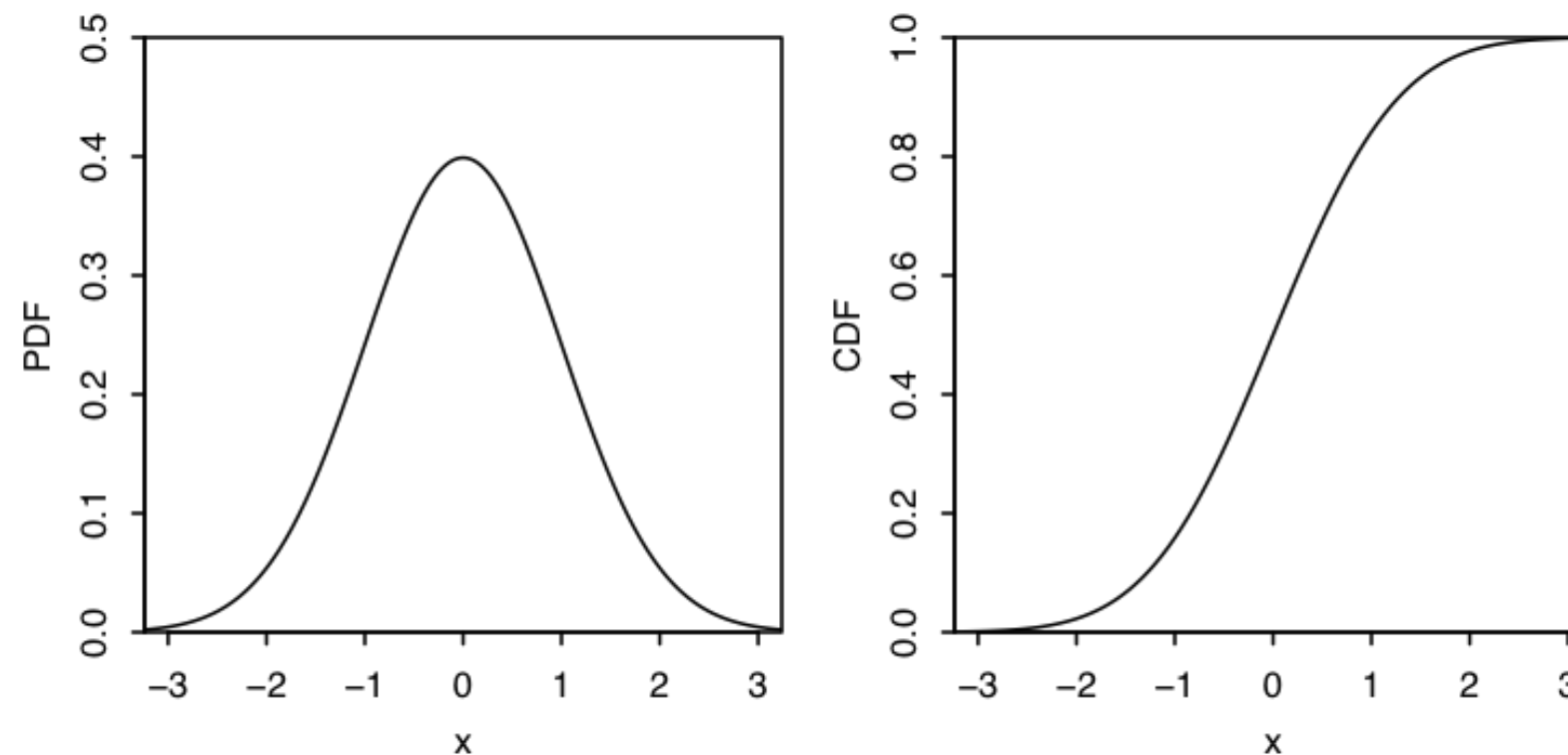


Figure 4.4.3: Standard Normal PDF φ (left) and CDF Φ (right).

[View Larger Image](#)

[Image Description](#)

There are several important symmetry properties that can be deduced from the standard Normal PDF and CDF.

1. *Symmetry of PDF:* φ satisfies $\varphi(z) = \varphi(-z)$, i.e., φ is an even function.
2. *Symmetry of tail areas:* The area under the PDF curve to the left of -2 , which is $P(Z \leq -2) = \Phi(-2)$ by definition, equals the area to the right of 2 , which is $P(Z \geq 2) = 1 - \Phi(2)$. In general, we have

$$\Phi(z) = 1 - \Phi(-z)$$

for all z . This can be seen visually by looking at the PDF curve, and mathematically by substituting $u = -t$ below and using the fact that PDFs integrate to 1:

$$\Phi(-z) = \int_{-\infty}^{-z} \varphi(t) dt = \int_z^{\infty} \varphi(u) du = 1 - \int_{-\infty}^z \varphi(u) du = 1 - \Phi(z).$$

3. *Symmetry of Z and $-Z$:* if $Z \sim \mathcal{N}(0, 1)$, then $-Z \sim \mathcal{N}(0, 1)$ as well. To see this, note that the CDF of $-Z$ is

$$P(-Z \leq z) = P(Z \geq -z) = 1 - \Phi(-z),$$

but that is $\Phi(z)$, according to what we just argued. So $-Z$ has CDF Φ .

The general Normal distribution has two parameters, denoted μ and σ^2 , which are the mean and variance (the mean and variance of a distribution are measures of the average and how spread out the distribution is, respectively; these are defined and explored in the next unit). Starting with a standard Normal r.v. $Z \sim \mathcal{N}(0, 1)$, we can convert to a Normal r.v. with any desired parameters μ and σ^2 by a location-scale transformation.

DEFINITION 4.4.4 (NORMAL DISTRIBUTION).

If $Z \sim \mathcal{N}(0, 1)$, then

$$X = \mu + \sigma Z$$

is said to have the *Normal distribution* with mean parameter μ and variance parameter σ^2 , for any real μ and σ^2 with $\sigma > 0$. We denote this by $X \sim \mathcal{N}(\mu, \sigma^2)$.

Of course, if we can get from Z to X , then we can get from X back to Z . The process of getting a standard Normal from a non-standard Normal is called, appropriately enough, *standardization*. For $X \sim \mathcal{N}(\mu, \sigma^2)$, the *standardized version* of X is

$$\frac{X - \mu}{\sigma} \sim \mathcal{N}(0, 1).$$

We can use standardization to find the CDF and PDF of X in terms of the standard Normal CDF and PDF.

THEOREM 4.4.5 (NORMAL CDF AND PDF).

Let $X \sim \mathcal{N}(\mu, \sigma^2)$. Then the CDF of X is

$$F(x) = \Phi\left(\frac{x - \mu}{\sigma}\right),$$

and the PDF of X is

$$f(x) = \varphi\left(\frac{x - \mu}{\sigma}\right) \frac{1}{\sigma}.$$

Proof



For the CDF, we start from the definition $F(x) = P(X \leq x)$, standardize, and use the CDF of the standard Normal:

$$F(x) = P(X \leq x) = P\left(\frac{X - \mu}{\sigma} \leq \frac{x - \mu}{\sigma}\right) = \Phi\left(\frac{x - \mu}{\sigma}\right).$$

Then we differentiate to get the PDF, remembering to apply the chain rule:

$$\begin{aligned} f(x) &= \frac{d}{dx} \Phi\left(\frac{x - \mu}{\sigma}\right) \\ &= \varphi\left(\frac{x - \mu}{\sigma}\right) \frac{1}{\sigma}. \end{aligned}$$

We can also write out the PDF as

$$f(x) = \frac{1}{\sqrt{2\pi}\sigma} \exp\left(-\frac{(x - \mu)^2}{2\sigma^2}\right).$$

Three important benchmarks for the Normal distribution are the probabilities of falling within one, two, and three standard deviations of the mean parameter μ . The 68-95-99.7 rule tells us that these probabilities are what the name suggests.

THEOREM 4.4.6 (68-95-99.7 RULE).

If $X \sim \mathcal{N}(\mu, \sigma^2)$, then

$$\begin{aligned} P(|X - \mu| < \sigma) &\approx 0.68 \\ P(|X - \mu| < 2\sigma) &\approx 0.95 \\ P(|X - \mu| < 3\sigma) &\approx 0.997. \end{aligned}$$

Example 4.4.7.

Let $X \sim \mathcal{N}(-1, 4)$. What is $P(|X| < 3)$, exactly (in terms of Φ) and approximately?

Solution

The event $|X| < 3$ is the same as the event $-3 < X < 3$. We use standardization to express this event in terms of the standard Normal r.v. $Z = (X - (-1))/2$, then apply the 68-95-99.7 rule to get an approximation. The exact answer is

$$P(-3 < X < 3) = P\left(\frac{-3 - (-1)}{2} < \frac{X - (-1)}{2} < \frac{3 - (-1)}{2}\right) = P(-1 < Z < 2),$$

which is $\Phi(2) - \Phi(-1)$. The 68-95-99.7 rule tells us that $P(-1 < Z < 1) \approx 0.68$ and $P(-2 < Z < 2) \approx 0.95$. In other words, going from ± 1 standard deviation to ± 2 standard deviations adds approximately $0.95 - 0.68 = 0.27$ to the area under the curve. By symmetry, this is evenly divided between the areas $P(-2 < Z < -1)$ and $P(1 < Z < 2)$. Therefore,

$$P(-1 < Z < 2) = P(-1 < Z < 1) + P(1 < Z < 2) \approx 0.68 + \frac{0.27}{2} = 0.815.$$

This is close to the correct value, $\Phi(2) - \Phi(-1) \approx 0.8186$.

One more useful property of the Normal distribution is that the sum of independent Normals is Normal.

THEOREM 4.4.8 (SUM OF INDEPENDENT NORMALS).

If $X_1 \sim \mathcal{N}(\mu_1, \sigma_1^2)$ and $X_2 \sim \mathcal{N}(\mu_2, \sigma_2^2)$ are independent, then

$$X_1 + X_2 \sim \mathcal{N}(\mu_1 + \mu_2, \sigma_1^2 + \sigma_2^2),$$

$$X_1 - X_2 \sim \mathcal{N}(\mu_1 - \mu_2, \sigma_1^2 + \sigma_2^2).$$

