



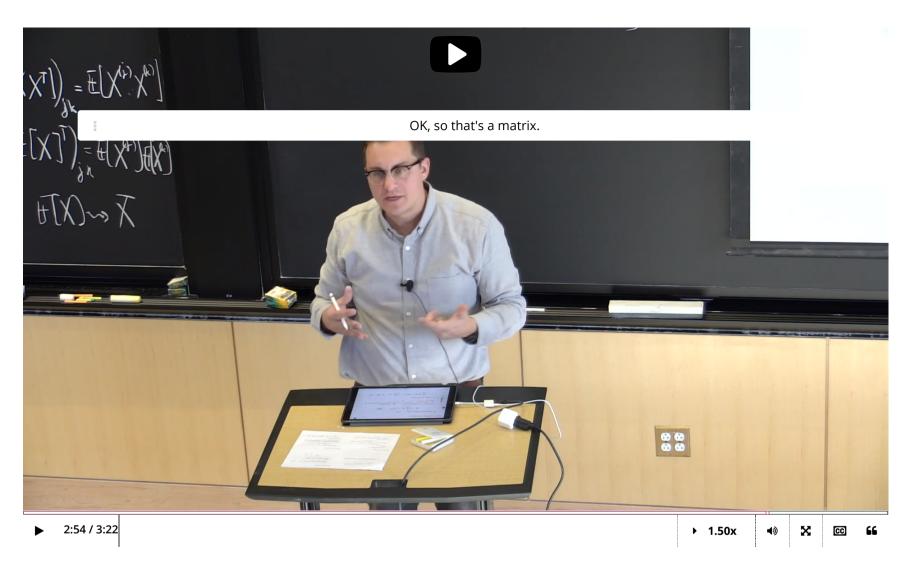
(Optional) Unit 8 Principal
Course > component analysis

(Optional) Lecture 23: Principal

> Component Analysis

- 3. Multivariate Statistics and Geometry Behind the Empirical
- > Covariance

# 3. Multivariate Statistics and Geometry Behind the Empirical Covariance Empirical Covariance Matrix



Video

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Transcripts

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### Bias

1/1 point (ungraded)

Let  $\mathbf{X}_i, i=1,\ldots,n$  be iid data points in  $\mathbb{R}^d$ . As presented in the lecture and given in the slides, let S be the empirical covariance matrix

$$S riangleq rac{1}{n} \sum_{i=1}^n \left( \mathbf{X}_i \mathbf{X}_i^T 
ight) - \overline{\mathbf{X}} \ \overline{\mathbf{X}}^T,$$

where  $\overline{\mathbf{X}}$  is the empirical or sample mean  $\frac{1}{n} \sum_{i=1}^{n} \mathbf{X}_{i}$ .

Is the following **true or false**. "S is an unbiased estimator of the covariance matrix  $\Sigma$  of  $\mathbf{X}_i$ 's."

True



False



#### Solution:

The answer is **False**. We have seen before in <u>Lecture 10</u> that the correct fraction for obtaining an unbiased estimator for the sample variance or the sample covariance is  $\frac{1}{n-1}$  and not  $\frac{1}{n}$ . The same carries over to covariance matrices, as well (Self-exercise: Verify this statement).

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You have used 1 of 1 attempt

**1** Answers are displayed within the problem

## Projection Onto a Subspace: Example

0/1 point (ungraded)

**Note:** This problem is discussed briefly at the beginning of the following video.

Recall the matrix solution to the linear regression problem  $\min_{\beta} \|\mathbf{Y} - \mathbb{X}\boldsymbol{\beta}\|^2$ :

$$\widehat{oldsymbol{eta}} = \left(\mathbb{X}^T\mathbb{X}
ight)^{-1}\mathbb{X}^T\mathbf{Y},$$

where  $\mathbb{X}\in\mathbb{R}^{n imes d}$ ,  $\mathbf{Y}\in\mathbb{R}^n$ , and  $m{eta}\in\mathbb{R}^d$  and where we assume that  $\mathbb{X}$  is full column rank, i.e.  $\mathrm{rank}\left(\mathbb{X}\right)=d$ .

Choose from the following the correct geometric interpretation of  $\mathbb{X}\widehat{\beta}$ . That is,

 $igcup \mathbb{X}\widehat{m{eta}}$  is a vector in the column space of  $\mathbb{X}$  and is the orthogonal projection of  ${f Y}$  onto the column space of  $\mathbb{X}$ .  $m{\checkmark}$ 

 $lackbox{} \mathbb{X}\widehat{oldsymbol{eta}}$  is a vector in the row space of  $\mathbb{X}$  and is the orthogonal projection of  $\mathbf{Y}$  onto the row space of  $\mathbb{X}$ .



#### Solution:

By definition of column space,  $\mathbb{X}\widehat{\beta}$  is a vector in the column space of  $\mathbb{X}$ .

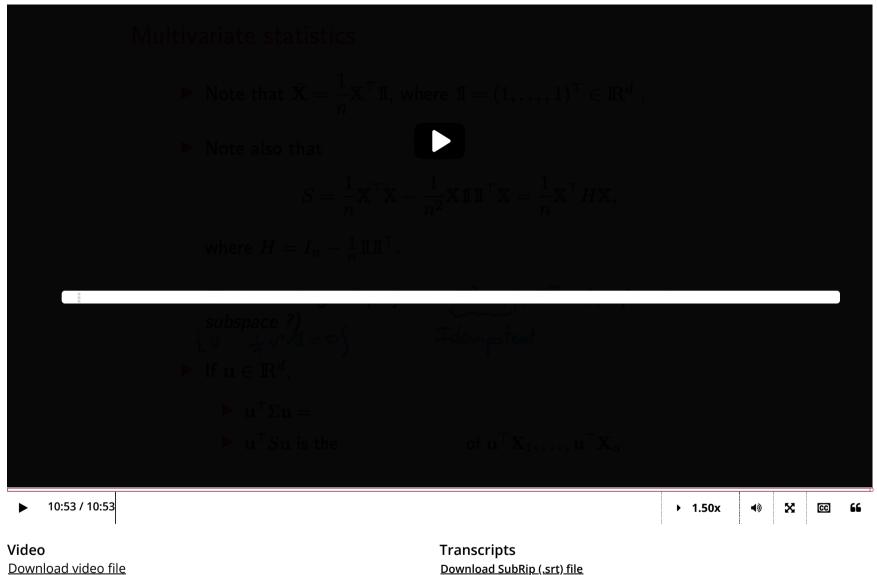
The orthogonal projection of a vector  $\mathbf{Y}$  onto the column space of a matrix such as  $\mathbb{X}$ , which is full column rank, is (in linear algebra theory) equal to  $\mathbb{X}(\mathbb{X}^T\mathbb{X})^{-1}\mathbb{X}^T\mathbf{Y}$ . The important observation in the linear least squares problem is that the solution  $\widehat{\boldsymbol{\beta}}$  to the linear least squares problem turns out to have the property that  $\mathbb{X}\widehat{\boldsymbol{\beta}}$  is the orthogonal projection of  $\mathbf{Y}$  onto the column space of  $\mathbb{X}$ .

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**1** Answers are displayed within the problem

# **Geometric View of Empirical Covariance**



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A matrix  $P \in \mathbb{R}^{n imes n}$  is called an **orthogonal projection matrix** if and only if  $P^2 = P = P^T$ . A matrix with the property that  $P^2 = P$  is also called **idempotent** .

# Projection: Zero Mean Vectors in 2 Dimensions

1/1 point (ungraded)

Consider the projection matrix introduced in the above video:

$$H=\mathbf{I}_n-rac{1}{n}egin{bmatrix}1\1\ dots\1\end{bmatrix}_{n imes 1} egin{bmatrix}1&1&\cdots&1\end{bmatrix}_{1 imes n}$$

Let n=2. What is the subspace that H , with n=2 , projects any vector  $\mathbf{x}\in\mathbb{R}^2$  onto?

- $igcap \{ {f y} : y^{(1)} y^{(2)} = 0 \}$
- $ullet \{ {f y}: rac{y^{(1)}+y^{(2)}}{2}=0 \}$
- $igcup \{ \mathbf{y} : y^{(1)} = rac{1}{2} y^{(2)} \}$



#### **Solution:**

As shown in the video, the projection matrix H projects any given vector  $\mathbf{x}$  onto the space of mean-removed vectors, which is choice 2.

The same solution can also be arrived at by looking at the columns of H:

$$H=egin{bmatrix} rac{1}{2} & -rac{1}{2} \ -rac{1}{2} & rac{1}{2} \end{bmatrix}$$

Notice that the two columns are linearly dependent (through a sign change). The column space of H is all vectors that are scalar multiples of

$$\left[\begin{array}{c}\frac{1}{2}\\-\frac{1}{2}\end{array}\right],$$

which is the line  $y^{(1)}=-y^{(2)}.$ 

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You have used 1 of 2 attempts

• Answers are displayed within the problem

## Discussion

**Topic:** (Optional) Unit 8 Principal component analysis:(Optional) Lecture 23: Principal Component Analysis / 3. Multivariate Statistics and Geometry Behind the Empirical Covariance

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