## Weak Law of Large Numbers

**Theorem (WLLN)**. If  $\{X_1,...,X_n\}$  are iid with  $E|X_i| < \infty$  and then  $\overline{X}_n = \frac{1}{n} \sum_{i=1}^n X_i \to_p E(X_i)$ .

**Proof.** Without loss of generality, we can set  $E(X_i) = 0$  (by recentering  $X_i$  on its expectation). We need to show that for all  $\delta > 0$  and  $\eta > 0$  there is some  $\overline{n} < \infty$  so that for all  $n \geq \overline{n}$ ,  $P(|\overline{X}_n| > \delta) \leq \eta$ . Fix  $\delta$  and  $\eta$ . Set  $\varepsilon = \delta \eta/3$ . Pick  $C < \infty$  large enough so that

$$E(|X|1(|X|>C)) \le \varepsilon \tag{1}$$

(where  $1(\cdot)$  is the indicator function) which is possible since  $E|X| < \infty$ . Then set

$$\overline{n} \ge 4C^2/\varepsilon^2. \tag{2}$$

Define the random vectors

$$W_{i} = X_{i}1 (|X_{i}| \leq C) - E(X_{i}1 (|X_{i}| \leq C))$$
  

$$Z_{i} = X_{i}1 (|X_{i}| > C) - E(X_{i}1 (|X_{i}| > C)).$$

Since  $X_i$  is iid,  $W_i$  and  $Z_i$  are also.

By Jensen's inequality and (1),

$$|E(X_i 1(|X_i| > C))| \le E(|X_i| 1(|X_i| > C)) \le \varepsilon.$$

By the triangle inequality and (1),

$$E|\overline{Z}_n| \le E|Z_i| \le E|X_i| 1(|X_i| > C) + |E(X_i 1(|X_i| > C))| \le 2\varepsilon.$$

Note that  $|W_i| \leq 2C$ . Thus (crudely)  $EW_i^2 \leq 4C^2$ . Since the  $W_i$  are iid and mean zero,

$$E\overline{W}_{n}^{2} = \frac{EW_{i}^{2}}{n} \le \frac{4C^{2}}{n} \le \varepsilon^{2}$$

the final inequality holding for  $n \geq \overline{n}$  by (2). Thus by Jensen's inequality

$$(E|\overline{W}_n|)^2 \le E\overline{W}_n^2 \le \varepsilon^2.$$

Finally, by Markov's inequality, the fact that  $\overline{X}_n = \overline{W}_n + \overline{Z}_n$ , the triangle inequality, and these two bounds,

$$P\left(\left|\overline{X}_{n}\right| > \delta\right) \leq \frac{E\left|\overline{X}_{n}\right|}{\delta} \leq \frac{E\left|\overline{W}_{n}\right| + E\left|\overline{Z}_{n}\right|}{\delta} \leq \frac{3\varepsilon}{\delta} = \eta,$$

the equality by the definition of  $\varepsilon$ . We have shown that for any  $\delta > 0$  and  $\eta > 0$  there is some  $\overline{n} < \infty$  so that for all  $n \ge \overline{n}$ ,  $P(|\overline{X}_n| > \delta) \le \eta$ , as needed.

## Strong Law of Large Numbers

**Theorem (SLLN).** If  $\{X_1,...,X_n\}$  are iid with  $E|X_i| < \infty$  and  $EX_i = \mu$  then  $\overline{X}_n \to_{a.s.} \mu$ as  $n \to \infty$ .

Classical proofs of strong laws are based on convergence results from analysis. Two powerful results are known as the Toeplitz Lemma and the Kronecker Lemma.

A **Toeplitz array**  $\{a_{ni}\}$  satisfies the following three characteristics:

- (i) For all  $n \ge 1$ ,  $\sum_{i=1}^{\infty} |a_{ni}| \le c < \infty$ (ii) As  $n \to \infty$ ,  $\sum_{i=1}^{\infty} a_{ni} \to 1$
- (iii) For all  $i \geq \overline{1}$ , as  $n \to \infty$ ,  $a_{ni} \to 0$

An example of a Toeplitz array is  $a_{ni} = 1/n$  if  $i \le n$ , else  $a_{ni} = 0$ .

**Toeplitz Lemma.** If  $\{a_{ni}\}$  is a Toeplitz array and  $x_n$  is a real sequence such that  $x_n \to x$  as  $n \to \infty$  then as  $n \to \infty$ 

$$y_n = \sum_{i=1}^{\infty} a_{ni} x_i \to x.$$

**Proof**: Using property (ii) WLOG assume x=0. Fix  $\varepsilon>0$  and pick N so that  $|x_i|\leq \varepsilon/2c$ for all  $i \geq N$ . Then by property (i)

$$|y_n| \leq \sum_{i=1}^{N} |a_{ni}| |x_i| + \sum_{i=N+1}^{\infty} |a_{ni}| |x_i|$$

$$\leq \sum_{i=1}^{N} |a_{ni}| |x_i| + \varepsilon/2$$

$$\leq \varepsilon$$

the final inequality holding for n sufficiently large by property (iii).

**Kronecker Lemma**. If  $b_n$  is an increasing real sequence with  $b_n \to \infty$ , and  $x_n$  is a real sequence such that  $\sum_{i=1}^{\infty} x_i$  exists (that is,  $\sum_{i=1}^{n} x_i$  converges to a finite limit as  $n \to \infty$ ), then

$$\frac{1}{b_n} \sum_{i=1}^n b_i x_i \to 0.$$

**Proof.** Let  $s_n = \sum_{i=1}^n x_i$  and define  $s_0 = 0$  and  $b_0 = 0$ . Now

$$\sum_{i=1}^{n} b_{i} x_{i} = \sum_{i=1}^{n} b_{i} s_{i} - \sum_{i=1}^{n} b_{i} s_{i-1}$$

$$= \sum_{i=1}^{n} b_{i} s_{i} - \sum_{i=1}^{n} b_{i-1} s_{i-1} - \sum_{i=1}^{n} (b_{i} - b_{i-1}) s_{i-1}$$

$$= b_{n} s_{n} - \sum_{i=1}^{n} (b_{i} - b_{i-1}) s_{i-1}.$$

Thus

$$\frac{1}{b_n} \sum_{i=1}^n b_i x_i = s_n - \sum_{i=1}^n a_{ni} s_{i-1}$$
(3)

where

$$a_{ni} = \frac{b_i - b_{i-1}}{b_n}$$

and we define  $a_{ni}=0$  for i>n. Note that  $|a_{ni}|\leq 1$ ,  $\sum_{i=1}^{\infty}a_{ni}=1$ , and  $a_{ni}\to 0$  as  $n\to\infty$ , so  $a_{ni}$  is a Toeplitz array. Since  $s_n\to x$  then  $\sum_{i=1}^na_{ni}s_{i-1}\to x$  by the Toeplitz Lemma and (3) converges to x-x=0.

We also need a strengthening of Markov's inequality.

**Kolmogorov's Inequality.** Assume  $U_1, ..., U_n$  are independent (but not necessarily iid) with  $EU_i = 0$ . Set  $S_j = \sum_{i=1}^j U_i$ . Then for any  $\lambda > 0$ 

$$P\left(\max_{1\leq i\leq n}|S_i|>\lambda\right)\leq \frac{ES_n^2}{\lambda^2}=\frac{1}{\lambda^2}\sum_{i=1}^n EU_i^2. \tag{4}$$

**Proof:** Define

$$I_{i-1} = \left\{ |S_i| > \lambda; \max_{j < i} |S_j| \le \lambda \right\},\,$$

the event that the sequence  $|S_j|$  first exceeds  $\lambda$  at j=i. Since these events are disjoint,

$$P\left(\max_{1 \le i \le n} |S_i| > \lambda\right) = P\left(\bigcup_{i=1}^n I_{i-1}\right) = \sum_{i=1}^n P(I_{i-1}) \le \sum_{i=1}^n P(I_{i-1}|S_i| > \lambda) \le \lambda^{-2} \sum_{i=1}^n E(I_{i-1}S_i^2).$$
(5)

The first inequality holds since  $I_{i-1} = 1$  implies  $I_{i-1} |S_i| > \lambda$ , and the last inequality is Markov's. Let  $\tilde{U}_i = (U_1, ..., U_i)$  and note that

$$E\left(S_n^2\mid \tilde{U}_i\right) = E\left(S_i^2\mid \tilde{U}_i\right) + 2E\left(S_i(S_n - S_i)\mid \tilde{U}_i\right) + E\left((S_n - S_i)^2\mid \tilde{U}_i\right) = S_i^2 + E(S_n - S_i)^2 \geq S_i^2$$

so using iterated expectations,

$$E\left(I_{i-1}S_i^2\right) \le E\left(I_{i-1}E\left(S_n^2 \mid \tilde{U}_i\right)\right) = E\left(E\left(I_{i-1}S_n^2 \mid \tilde{U}_i\right)\right) = E\left(I_{i-1}S_n^2\right). \tag{6}$$

Together, (5) and (6) show that

$$\lambda^2 P\left(\max_{1\leq i\leq n}|S_i|>\lambda\right)\leq \sum_{i=1}^n E\left(I_{i-1}S_n^2\right)=E\left(\left(\sum_{i=1}^n I_{i-1}\right)S_n^2\right)\leq ES_n^2.$$

Given the Kronecker Lemma and Kolmogorov's inequality, it is straightforward to establish the SLLN if  $Var(X) < \infty$ .

## Almost Sure Convergence Theorem. If

$$\sum_{i=1}^{\infty} \frac{Var(X_i)}{i^2} < \infty \tag{7}$$

then  $\overline{X}_n \to 0$  almost surely.

Before we prove this theorem, we state the following implication.

**Kolmogorov SLLN**. If  $X_i$  is iid and  $Var(X_i) < \infty$  then  $\overline{X}_n \to 0$  almost surely.

## Proof of Kolmogorov SLLN.

$$\sum_{i=1}^{\infty} \frac{Var(X_i)}{i^2} = Var(X_i) \sum_{i=1}^{\infty} \frac{1}{i^2} < \infty$$

since  $\sum_{i=1}^{\infty} \frac{1}{i^2} = \pi^2/6 < \infty$ . Then by the almost sure convergence theorem,  $\overline{X}_n \to 0$  almost surely.

**Proof of Almost Sure Convergence Theorem**. WLOG assume  $EX_i = 0$ . Let  $U_i = i^{-1}X_i$ . The Kronecker Lemma implies that if  $S_n = \sum_{i=1}^n U_i$  converges to a finite random limit as  $n \to \infty$ , then  $\overline{X}_n \to 0$ . (To see this, set  $x_i = U_i$  and  $b_i = i$ .) We now show that  $S_n$  converges almost surely as  $n \to \infty$ , so  $\overline{X}_n \to 0$  almost surely.

One characterization of convergence is that  $S_n$  converges iff  $S_{m+k} - S_m \to 0$  as  $m, k \to \infty$ . In other words,  $S_n$  converges if for all  $\varepsilon > 0$ , there is a sufficiently large  $\overline{m} < \infty$  such that for all  $m \geq \overline{m}$ ,  $|S_{m+k} - S_m| < \varepsilon$  for all  $k \geq 1$ . But for all  $\varepsilon > 0$  and  $m < \infty$ 

Under (7), (9) tends to 0 as  $m \to \infty$ , as required. Note that inequality (8) is Kolmogorov's inequality (4).

For a proof of the SLLN without assuming the variance is finite, we need another intermediate result.

**Lemma**.  $E|X| < \infty$  iff

$$\sum_{i=1}^{\infty} P(|X| > i) < \infty. \tag{10}$$

**Proof.** Let Y = |X|. By expansion

$$EY = \sum_{i=1}^{\infty} E(|X| 1 (i - 1 < Y \le i))$$

$$\leq \sum_{i=1}^{\infty} iE(1 (i - 1 < Y \le i))$$

$$= \sum_{i=1}^{\infty} iP(i - 1 < Y \le i)$$

$$= \sum_{i=1}^{\infty} iP(Y > i - 1) - \sum_{i=1}^{\infty} iP(Y > i)$$

$$= \sum_{i=1}^{\infty} (i + 1) P(Y > i) - \sum_{i=0}^{\infty} iP(Y > i)$$

$$= \sum_{i=1}^{\infty} P(Y > i)$$

Thus  $\sum_{i=1}^{\infty} P(Y > i) < \infty$  implies  $EY < \infty$ . The converse can be shown similarly.

General Proof of SLLN. WLOG assume  $EX_i = 0$ . By the previous Lemma,  $E|X_i| < \infty$  and  $X_i$  identically distributed implies

$$\sum_{i=1}^{\infty} P(|X_i| > i) < \infty. \tag{11}$$

The Borel-Cantelli Lemma states that (11) implies that  $P(\{|X_i| > i\} \text{ infinitely often}) = 0$ . This means that  $\overline{X}_n \to 0$  almost surely iff

$$\frac{1}{n} \sum_{i=1}^{n} X_i 1 (|X_i| \le i) \to 0$$

almost surely, which occurs iff

$$\frac{1}{n} \sum_{i=1}^{n} \left[ X_i 1 \left( |X_i| \le i \right) - E \left( X_i 1 \left( |X_i| \le i \right) \right) \right] \to 0$$

almost surely, since  $EX_i = 0$  and identically distributed implies  $E(X_i 1 (|X_i| \le i)) \to 0$  as  $i \to \infty$ , and an application of the Toeplitz Lemma yields

$$\frac{1}{n}\sum_{i=1}^{n}E\left(X_{i}1\left(\left|X_{i}\right|\leq i\right)\right)\rightarrow0.$$

By the almost sure convergence theorem, it is therefore sufficient to show that

$$\sum_{i=1}^{\infty} \frac{Var(X_{i}1(|X_{i}| \le i))}{i^{2}} \le \sum_{i=1}^{\infty} \frac{E(X_{i}^{2}1(|X_{i}| \le i))}{i^{2}} < \infty.$$

Let

$$A_j = \sum_{i=j}^{\infty} \frac{1}{i^2} \le \frac{2}{j}.$$

The inequality holds since for  $j=1, A_1=\pi^2/6<2$ , and for  $j\geq 2$ , by comparing  $A_j$  to the sum of rectangles beneath the curve  $x^{-2}$ ,

$$\sum_{i=j}^{\infty} \frac{1}{i^2} \le \int_{j-1}^{\infty} x^{-2} dx = \frac{1}{j-1} \le \frac{2}{j}.$$

Then by expanding and changing the order of summation

$$\sum_{i=1}^{\infty} \frac{E(X_i^2 1 (|X_i| \le i))}{i^2} = \sum_{i=1}^{\infty} \sum_{j=1}^{i} \frac{E(X_i^2 1 (j-1 < |X_i| \le j))}{i^2}$$

$$= \sum_{j=1}^{\infty} \sum_{i=j}^{\infty} \frac{E(X_i^2 1 (j-1 < |X_i| \le j))}{i^2}$$

$$= \sum_{j=1}^{\infty} E(X_i^2 1 (j-1 < |X_i| \le j)) A_j$$

$$\leq 2 \sum_{j=1}^{\infty} \frac{E(X_i^2 1 (j-1 < |X_i| \le j))}{j}$$

$$\leq 2 \sum_{j=1}^{\infty} E(|X_i| 1 (j-1 < |X_i| \le j))$$

$$= 2E|X| < \infty$$

which is what we wanted to show.