

# Gamma distribution

In probability theory and statistics, the **gamma distribution** is a two-parameter family of continuous probability distributions. The exponential distribution, Erlang distribution, and chi-squared distribution are special cases of the gamma distribution. There are three different parametrizations in common use:

1. With a shape parameter  $k$  and a scale parameter  $\theta$ .
2. With a shape parameter  $\alpha = k$  and an inverse scale parameter  $\beta = 1/\theta$ , called a rate parameter.
3. With a shape parameter  $k$  and a mean parameter  $\mu = k\theta = \alpha/\beta$ .

In each of these three forms, both parameters are positive real numbers.

The gamma distribution is the maximum entropy probability distribution (both with respect to a uniform base measure and with respect to a  $1/x$  base measure) for a random variable  $X$  for which  $\mathbf{E}[X] = k\theta = \alpha/\beta$  is fixed and greater than zero, and  $\mathbf{E}[\ln(X)] = \psi(k) + \ln(\theta) = \psi(\alpha) - \ln(\beta)$  is fixed ( $\psi$  is the digamma function).<sup>[1]</sup>

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# Parameterizations

The parameterization with  $k$  and  $\theta$  appears to be more common in econometrics and certain other applied fields, where for example the gamma distribution is frequently used to model waiting times. For instance, in life testing, the waiting time until death is a random variable that is frequently modeled with a gamma distribution.<sup>[2]</sup>

The parameterization with  $\alpha$  and  $\beta$  is more common in Bayesian statistics, where the gamma distribution is used as a conjugate prior distribution for various types of inverse scale (aka rate) parameters, such as the  $\lambda$  of an exponential distribution or a Poisson distribution<sup>[3]</sup> – or for that matter, the  $\beta$  of the gamma distribution itself. (The closely related inverse gamma distribution is used as a conjugate prior for scale parameters, such as the variance of a normal distribution.)

If  $k$  is a positive integer, then the distribution represents an Erlang distribution; i.e., the sum of  $k$  independent exponentially distributed random variables, each of which has a mean of  $\theta$ .

## Characterization using shape $\alpha$ and rate $\beta$

The gamma distribution can be parameterized in terms of a shape parameter  $\alpha = k$  and an inverse scale parameter  $\beta = 1/\theta$ , called a rate parameter. A random variable  $X$  that is gamma-distributed with shape  $\alpha$  and rate  $\beta$  is denoted

$$X \sim \Gamma(\alpha, \beta) \equiv \text{Gamma}(\alpha, \beta)$$

The corresponding probability density function in the shape-rate parametrization is

$$f(x; \alpha, \beta) = \frac{\beta^\alpha x^{\alpha-1} e^{-\beta x}}{\Gamma(\alpha)} \quad \text{for } x > 0 \quad \alpha, \beta > 0,$$
  
$$f(x; \alpha, \theta) = \frac{x^{\alpha-1} e^{-x/\theta}}{\theta^\alpha \Gamma(\alpha)} \quad \text{for } x > 0, \quad \alpha, \theta > 0.$$

where  $\Gamma(\alpha)$  is the gamma function.

Both parametrizations are common because either can be more convenient depending on the situation.

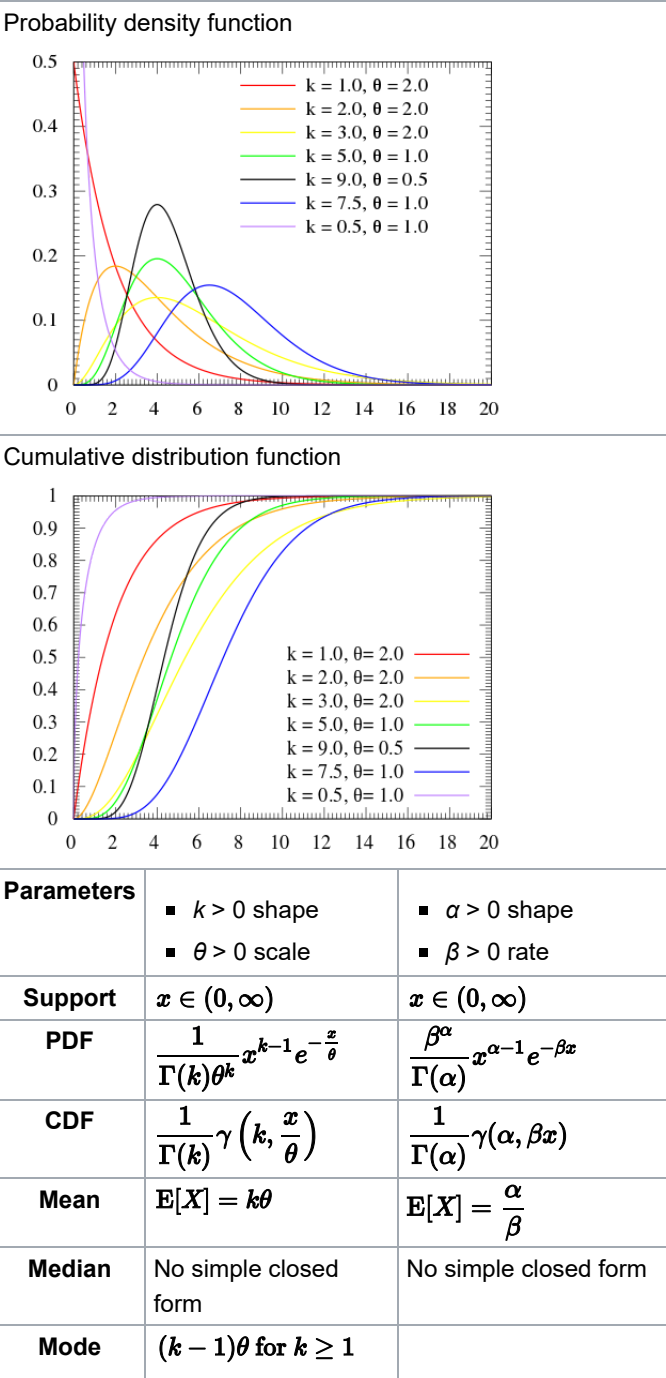
The cumulative distribution function is the regularized gamma function:

$$F(x; \alpha, \beta) = \int_0^x f(u; \alpha, \beta) du = \frac{\gamma(\alpha, \beta x)}{\Gamma(\alpha)},$$

where  $\gamma(\alpha, \beta x)$  is the lower incomplete gamma function.

If  $\alpha$  is a positive integer (i.e., the distribution is an Erlang distribution), the cumulative distribution function has the following series expansion:<sup>[4]</sup>

## Gamma



$$F(x; \alpha, \beta) = 1 - \sum_{i=0}^{\alpha-1} \frac{(\beta x)^i}{i!} e^{-\beta x} = e^{-\beta x} \sum_{i=\alpha}^{\infty} \frac{(\beta x)^i}{i!}.$$

Characterization using shape *k* and scale *θ*

A random variable *X* that is gamma-distributed with shape *k* and scale *θ* is denoted by

$$X \sim \Gamma(k, \theta) \equiv \text{Gamma}(k, \theta)$$

The [probability density function](#) using the shape-scale parametrization is

$$f(x; k, \theta) = \frac{x^{k-1} e^{-\frac{x}{\theta}}}{\theta^k \Gamma(k)} \quad \text{for } x > 0 \text{ and } k, \theta > 0.$$

Here  $\Gamma(k)$  is the [gamma function](#) evaluated at *k*.

The [cumulative distribution function](#) is the regularized gamma function:

$$F(x; k, \theta) = \int_0^x f(u; k, \theta) du = \frac{\gamma\left(k, \frac{x}{\theta}\right)}{\Gamma(k)},$$

where  $\gamma\left(k, \frac{x}{\theta}\right)$  is the lower [incomplete gamma function](#).

It can also be expressed as follows, if *k* is a positive [integer](#) (i.e., the distribution is an [Erlang distribution](#)):<sup>[4]</sup>

$$F(x; k, \theta) = 1 - \sum_{i=0}^{k-1} \frac{1}{i!} \left(\frac{x}{\theta}\right)^i e^{-x/\theta} = e^{-x/\theta} \sum_{i=k}^{\infty} \frac{1}{i!} \left(\frac{x}{\theta}\right)^i.$$

Properties

Skewness

The skewness of the gamma distribution only depends on its shape parameter, *k*, and it is equal to  $2/\sqrt{k}$ .

Median calculation

Unlike the mode and the mean which have readily calculable formulas based on the parameters, the median does not have an easy closed form equation. The median for this distribution is defined as the value *v* such that

$$\frac{1}{\Gamma(k)\theta^k} \int_0^v x^{k-1} e^{-x/\theta} dx = \frac{1}{2}.$$

		$\frac{\alpha - 1}{\beta}$ for $\alpha \geq 1$
Variance	$\text{Var}(X) = k\theta^2$	$\text{Var}(X) = \frac{\alpha}{\beta^2}$
Skewness	$\frac{2}{\sqrt{k}}$	$\frac{2}{\sqrt{\alpha}}$
Excess kurtosis	$\frac{6}{k}$	$\frac{6}{\alpha}$
Entropy	$k + \ln \theta + \ln \Gamma(k) + (1 - k)\psi(k)$	$\alpha - \ln \beta + \ln \Gamma(\alpha) + (1 - \alpha)\psi(\alpha)$
MGF	$(1 - \theta t)^{-k}$ for $t < \frac{1}{\theta}$	$\left(1 - \frac{t}{\beta}\right)^{-\alpha}$ for $t < \beta$
CF	$(1 - \theta it)^{-k}$	$\left(1 - \frac{it}{\beta}\right)^{-\alpha}$

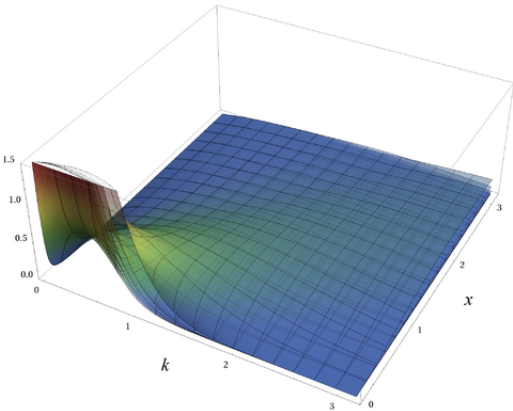


Illustration of the gamma PDF for parameter values over *k* and *x* with *θ* set to 1, 2, 3, 4, 5 and 6. One can see each *θ* layer by itself [here](#) [2] (<https://commons.wikimedia.org/wiki/File:Gamma-PDF-3D-by-k.png>) as well as by *k* [3] (<https://commons.wikimedia.org/wiki/File:Gamma-PDF-3D-by-Theta.png>) and *x*. [4] (<https://commons.wikimedia.org/wiki/File:Gamma-PDF-3D-by-x.png>).

A formula for approximating the median for any gamma distribution, when the mean is known, has been derived based on the fact that the ratio  $\mu/(\mu - \nu)$  is approximately a linear function of  $k$  when  $k \geq 1$ .<sup>[5]</sup> The approximation formula is

$$\nu \approx \mu \frac{3k - 0.8}{3k + 0.2},$$

where  $\mu = k\theta$  is the mean.

A rigorous treatment of the problem of determining an asymptotic expansion and bounds for the median of the Gamma Distribution was handled first by Chen and Rubin, who proved

$$m - \frac{1}{3} < \lambda(m) < m,$$

where  $\lambda(m)$  denotes the median of the **Gamma**( $m, 1$ ) distribution.<sup>[6]</sup>

K. P. Choi later showed that the first five terms in the asymptotic expansion of the median are

$$\lambda(m) = m - \frac{1}{3} + \frac{8}{405m} + \frac{184}{25515m^2} + \frac{2248}{3444525m^3} - \dots$$

by comparing the median to Ramanujan's  $\theta$  function.<sup>[7]</sup>

Later, it was shown that  $\lambda(m)$  is a convex function of  $m$ .<sup>[8]</sup>

## Summation

If  $X_i$  has a Gamma( $k_i, \theta$ ) distribution for  $i = 1, 2, \dots, N$  (i.e., all distributions have the same scale parameter  $\theta$ ), then

$$\sum_{i=1}^N X_i \sim \text{Gamma} \left( \sum_{i=1}^N k_i, \theta \right)$$

provided all  $X_i$  are independent.

For the cases where the  $X_i$  are independent but have different scale parameters see Mathai (1982) and Moschopoulos (1984).

The gamma distribution exhibits infinite divisibility.

## Scaling

If

$$X \sim \text{Gamma}(k, \theta),$$

then, for any  $c > 0$ ,

$$cX \sim \text{Gamma}(k, c\theta), \text{ by moment generating functions,}$$

or equivalently

$$cX \sim \text{Gamma}\left(\alpha, \frac{\beta}{c}\right),$$

Indeed, we know that if  $X$  is an exponential r.v. with rate  $\lambda$  then  $cX$  is an exponential r.v. with rate  $\lambda/c$ ; the same thing is valid with Gamma variates (and this can be checked using the moment-generating function, see, e.g., these notes ([http://www.stat.washington.edu/thompson/S341\\_10/Notes/week4.pdf](http://www.stat.washington.edu/thompson/S341_10/Notes/week4.pdf)), 10.4-(ii)): multiplication by a positive constant  $c$  divides the rate (or, equivalently, multiplies the scale).

## Exponential family

The gamma distribution is a two-parameter exponential family with natural parameters  $k - 1$  and  $-1/\theta$  (equivalently,  $\alpha - 1$  and  $-\beta$ ), and natural statistics  $X$  and  $\ln(X)$ .

If the shape parameter  $k$  is held fixed, the resulting one-parameter family of distributions is a natural exponential family.

## Logarithmic expectation and variance

One can show that

$$\mathbf{E}[\ln(X)] = \psi(\alpha) - \ln(\beta)$$

or equivalently,

$$\mathbf{E}[\ln(X)] = \psi(k) + \ln(\theta)$$

where  $\psi$  is the digamma function. Likewise,

$$\mathbf{var}[\ln(X)] = \psi^{(1)}(\alpha) = \psi^{(1)}(k)$$

where  $\psi^{(1)}$  is the polygamma function.

This can be derived using the exponential family formula for the moment generating function of the sufficient statistic, because one of the sufficient statistics of the gamma distribution is  $\ln(x)$ .

## Information entropy

The information entropy is

$$\begin{aligned} \mathbf{H}(X) &= \mathbf{E}[-\ln(p(X))] \\ &= \mathbf{E}[-\alpha \ln(\beta) + \ln(\Gamma(\alpha)) - (\alpha - 1) \ln(X) + \beta X] \\ &= \alpha - \ln(\beta) + \ln(\Gamma(\alpha)) + (1 - \alpha)\psi(\alpha). \end{aligned}$$

In the  $k, \theta$  parameterization, the information entropy is given by

$$\mathbf{H}(X) = k + \ln(\theta) + \ln(\Gamma(k)) + (1 - k)\psi(k).$$

## Kullback–Leibler divergence

The Kullback–Leibler divergence (KL-divergence), of  $\text{Gamma}(\alpha_p, \beta_p)$  ("true" distribution) from  $\text{Gamma}(\alpha_q, \beta_q)$  ("approximating" distribution) is given by<sup>[9]</sup>

$$D_{\text{KL}}(\alpha_p, \beta_p; \alpha_q, \beta_q) = (\alpha_p - \alpha_q)\psi(\alpha_p) - \log \Gamma(\alpha_p) + \log \Gamma(\alpha_q) \\ + \alpha_q(\log \beta_p - \log \beta_q) + \alpha_p \frac{\beta_q - \beta_p}{\beta_p}.$$

Written using the  $k, \theta$  parameterization, the KL-divergence of  $\text{Gamma}(k_p, \theta_p)$  from  $\text{Gamma}(k_q, \theta_q)$  is given by

$$D_{\text{KL}}(k_p, \theta_p; k_q, \theta_q) = (k_p - k_q)\psi(k_p) - \log \Gamma(k_p) + \log \Gamma(k_q) \\ + k_q(\log \theta_q - \log \theta_p) + k_p \frac{\theta_p - \theta_q}{\theta_q}.$$

## Laplace transform

The Laplace transform of the gamma PDF is

$$F(s) = (1 + \theta s)^{-k} = \frac{\beta^\alpha}{(s + \beta)^\alpha}.$$

## Parameter estimation

### Maximum likelihood estimation

The likelihood function for  $N$  iid observations  $(x_1, \dots, x_N)$  is

$$L(k, \theta) = \prod_{i=1}^N f(x_i; k, \theta)$$

from which we calculate the log-likelihood function

$$\ell(k, \theta) = (k - 1) \sum_{i=1}^N \ln(x_i) - \sum_{i=1}^N \frac{x_i}{\theta} - Nk \ln(\theta) - N \ln(\Gamma(k))$$

Finding the maximum with respect to  $\theta$  by taking the derivative and setting it equal to zero yields the maximum likelihood estimator of the  $\theta$  parameter:

$$\hat{\theta} = \frac{1}{kN} \sum_{i=1}^N x_i$$

Substituting this into the log-likelihood function gives

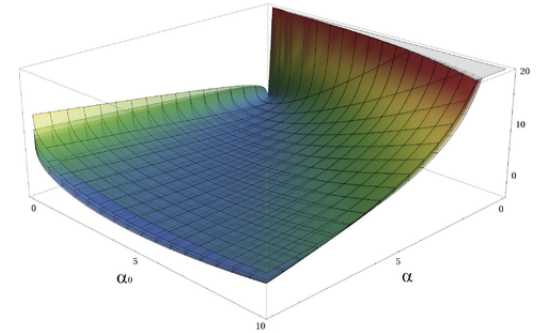


Illustration of the Kullback–Leibler (KL) divergence for two gamma PDFs. Here  $\beta = \beta_0 + 1$  which are set to 1, 2, 3, 4, 5 and 6. The typical asymmetry for the KL divergence is clearly visible.

$$\ell = (k-1) \sum_{i=1}^N \ln(x_i) - Nk - Nk \ln\left(\frac{\sum x_i}{kN}\right) - N \ln(\Gamma(k))$$

Finding the maximum with respect to  $k$  by taking the derivative and setting it equal to zero yields

$$\ln(k) - \psi(k) = \ln\left(\frac{1}{N} \sum_{i=1}^N x_i\right) - \frac{1}{N} \sum_{i=1}^N \ln(x_i)$$

There is no closed-form solution for  $k$ . The function is numerically very well behaved, so if a numerical solution is desired, it can be found using, for example, [Newton's method](#). An initial value of  $k$  can be found either using the [method of moments](#), or using the approximation

$$\ln(k) - \psi(k) \approx \frac{1}{2k} \left(1 + \frac{1}{6k+1}\right)$$

If we let

$$s = \ln\left(\frac{1}{N} \sum_{i=1}^N x_i\right) - \frac{1}{N} \sum_{i=1}^N \ln(x_i)$$

then  $k$  is approximately

$$k \approx \frac{3 - s + \sqrt{(s-3)^2 + 24s}}{12s}$$

which is within 1.5% of the correct value.<sup>[10]</sup> An explicit form for the Newton–Raphson update of this initial guess is:<sup>[11]</sup>

$$k \leftarrow k - \frac{\ln(k) - \psi(k) - s}{\frac{1}{k} - \psi'(k)}.$$

## Closed-form estimators

Consistent closed-form estimators of  $k$  and  $\theta$  exists that are derived from the likelihood of the [Generalized gamma distribution](#)<sup>[12]</sup>.

The estimate for the shape  $k$  is

$$\hat{k} = \frac{N \sum_{i=1}^N x_i}{N \sum_{i=1}^N x_i \ln(x_i) - \sum_{i=1}^N \ln(x_i) \sum_{i=1}^N x_i}$$

and the estimate for the scale  $\theta$  is

$$\hat{\theta} = \frac{1}{N^2} \left( N \sum_{i=1}^N x_i \ln(x_i) - \sum_{i=1}^N \ln(x_i) \sum_{i=1}^N x_i \right)$$

If the rate parameterization is used, the estimate of  $\hat{\beta} = \frac{1}{\hat{\theta}}$ .

These estimators are not strictly maximum likelihood estimators, but are instead referred to as mixed type log-moment estimators. They have however similar efficiency as the maximum likelihood estimators.

Although these estimators are consistent, they have a small bias. A bias-corrected variant of the estimator for the scale  $\theta$  is

$$\tilde{\theta} = \frac{N}{N-1} \hat{\theta}$$

A bias correction for the shape parameter  $k$  is given as <sup>[13]</sup>

$$\tilde{k} = \hat{k} - \frac{1}{N} \left( 3\hat{k} - \frac{2}{3} \left( \frac{\hat{k}}{1 + \hat{k}} \right) - \frac{4}{5} \frac{\hat{k}}{(1 + \hat{k})^2} \right)$$

## Bayesian minimum mean squared error

With known  $k$  and unknown  $\theta$ , the posterior density function for theta (using the standard scale-invariant prior for  $\theta$ ) is

$$P(\theta \mid k, x_1, \dots, x_N) \propto \frac{1}{\theta} \prod_{i=1}^N f(x_i; k, \theta)$$

Denoting

$$y \equiv \sum_{i=1}^N x_i, \quad P(\theta \mid k, x_1, \dots, x_N) = C(x_i) \theta^{-Nk-1} e^{-y/\theta}$$

Integration with respect to  $\theta$  can be carried out using a change of variables, revealing that  $1/\theta$  is gamma-distributed with parameters  $\alpha = Nk$ ,  $\beta = y$ .

$$\int_0^\infty \theta^{-Nk-1+m} e^{-y/\theta} d\theta = \int_0^\infty x^{Nk-1-m} e^{-xy} dx = y^{-(Nk-m)} \Gamma(Nk-m)$$

The moments can be computed by taking the ratio ( $m$  by  $m = 0$ )

$$\mathbb{E}[x^m] = \frac{\Gamma(Nk-m)}{\Gamma(Nk)} y^m$$

which shows that the mean  $\pm$  standard deviation estimate of the posterior distribution for  $\theta$  is

$$\frac{y}{Nk-1} \pm \sqrt{\frac{y^2}{(Nk-1)^2(Nk-2)}}.$$

## Generating gamma-distributed random variables



Given the scaling property above, it is enough to generate gamma variables with  $\theta = 1$  as we can later convert to any value of  $\beta$  with simple division.

Suppose we wish to generate random variables from  $\text{Gamma}(n + \delta, 1)$ , where  $n$  is a non-negative integer and  $0 < \delta < 1$ . Using the fact that a  $\text{Gamma}(1, 1)$  distribution is the same as an  $\text{Exp}(1)$  distribution, and noting the method of [generating exponential variables](#), we conclude that if  $U$  is [uniformly distributed](#) on  $(0, 1]$ , then  $-\ln(U)$  is distributed  $\text{Gamma}(1, 1)$ . Now, using the " $\alpha$ -addition" property of gamma distribution, we expand this result:

$$-\sum_{k=1}^n \ln U_k \sim \Gamma(n, 1)$$

where  $U_k$  are all uniformly distributed on  $(0, 1]$  and [independent](#). All that is left now is to generate a variable distributed as  $\text{Gamma}(\delta, 1)$  for  $0 < \delta < 1$  and apply the " $\alpha$ -addition" property once more. This is the most difficult part.

Random generation of gamma variates is discussed in detail by Devroye,<sup>[14]:401–428</sup> noting that none are uniformly fast for all shape parameters. For small values of the shape parameter, the algorithms are often not valid.<sup>[14]:406</sup> For arbitrary values of the shape parameter, one can apply the Ahrens and Dieter<sup>[15]</sup> modified acceptance–rejection method Algorithm GD (shape  $k \geq 1$ ), or transformation method<sup>[16]</sup> when  $0 < k < 1$ . Also see Cheng and Feast Algorithm GKM 3<sup>[17]</sup> or Marsaglia's squeeze method.<sup>[18]</sup>

The following is a version of the Ahrens-Dieter [acceptance–rejection method](#):<sup>[15]</sup>

1. Generate  $U$ ,  $V$  and  $W$  as [iid](#) uniform  $(0, 1]$  variates.
2. If  $U \leq \frac{e}{e + \delta}$  then  $\xi = V^{1/\delta}$  and  $\eta = W\xi^{\delta-1}$ . Otherwise,  $\xi = 1 - \ln V$  and  $\eta = We^{-\xi}$ .
3. If  $\eta > \xi^{\delta-1}e^{-\xi}$  then go to step 1.
4.  $\xi$  is distributed as  $\Gamma(\delta, 1)$ .

A summary of this is

$$\theta \left( \xi - \sum_{i=1}^{\lfloor k \rfloor} \ln(U_i) \right) \sim \Gamma(k, \theta)$$

where  $\lfloor k \rfloor$  is the integer part of  $k$ ,  $\xi$  is generated via the algorithm above with  $\delta = \{k\}$  (the fractional part of  $k$ ) and the  $U_k$  are all independent.

While the above approach is technically correct, Devroye notes that it is linear in the value of  $k$  and in general is not a good choice. Instead he recommends using either rejection-based or table-based methods, depending on context.<sup>[14]:401–428</sup>

For example, Marsaglia's simple transformation-rejection method relying on a one normal and one uniform random number:<sup>[19]</sup>

1. Setup:  $d = a - 1/3$ ,  $c = 1/\sqrt{9d}$ .
2. Generate:  $v = (1 + c \cdot x)^3$ , with  $x$  standard normal.
3. if  $v > 0$  and  $\log(\text{UNI}) < 0.5 \cdot x^2 + d - dv + d \log(v)$  return  $dv$ .
4. go back to step 2.

With  $1 \leq a = \alpha = k$  generates a gamma distributed random number in time that is approximately constant with  $k$ . The acceptance rate does depend on  $k$ , with an acceptance rate of 0.95, 0.98, and 0.99 for  $k=1, 2$ , and 4. For  $k < 1$ , one can use  $\gamma_\alpha = \gamma_{1+\alpha} U^{1/\alpha}$  to boost  $k$  to be usable with this method.

## Applications

The gamma distribution has been used to model the size of [insurance claims](#)<sup>[20]</sup> and [rainfalls](#).<sup>[21]</sup> This means that aggregate insurance claims and the amount of rainfall accumulated in a reservoir are modelled by a [gamma process](#) – much like the [exponential distribution](#) generates a [Poisson process](#).

The gamma distribution is also used to model errors in multi-level [Poisson regression](#) models, because the combination of the [Poisson distribution](#) and a gamma distribution is a [negative binomial distribution](#).

In wireless communication, the gamma distribution is used to model the [multi-path fading](#) of signal power; see also [Rayleigh distribution](#) and [Rician distribution](#).

In [oncology](#), the age distribution of [cancer incidence](#) often follows the gamma distribution, whereas the shape and scale parameters predict, respectively, the number of [driver events](#) and the time interval between them<sup>[22]</sup>.

In [neuroscience](#), the gamma distribution is often used to describe the distribution of [inter-spike intervals](#).<sup>[23][24]</sup>

In [bacterial gene expression](#), the [copy number](#) of a [constitutively expressed](#) protein often follows the gamma distribution, where the scale and shape parameter are, respectively, the mean number of bursts per cell cycle and the mean number of [protein molecules](#) produced by a single mRNA during its lifetime.<sup>[25]</sup>

In [genomics](#), the gamma distribution was applied in [peak calling](#) step (i.e. in recognition of signal) in [ChIP-chip](#)<sup>[26]</sup> and [ChIP-seq](#)<sup>[27]</sup> data analysis.

The gamma distribution is widely used as a [conjugate prior](#) in Bayesian statistics. It is the conjugate prior for the precision (i.e. inverse of the variance) of a [normal distribution](#). It is also the conjugate prior for the [exponential distribution](#).

## Related distributions

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### Conjugate prior

In Bayesian inference, the **gamma distribution** is the [conjugate prior](#) to many likelihood distributions: the [Poisson](#), [exponential](#), [normal](#) (with known mean), [Pareto](#), gamma with known shape  $\sigma$ , [inverse gamma](#) with known shape parameter, and [Gompertz](#) with known scale parameter.

The gamma distribution's [conjugate prior](#) is:<sup>[28]</sup>

$$p(k, \theta \mid p, q, r, s) = \frac{1}{Z} \frac{p^{k-1} e^{-\theta^{-1} q}}{\Gamma(k)^r \theta^{ks}},$$

where  $Z$  is the normalizing constant, which has no closed-form solution. The posterior distribution can be found by updating the parameters as follows:

$$\begin{aligned} p' &= p \prod_i x_i, \\ q' &= q + \sum_i x_i, \\ r' &= r + n, \\ s' &= s + n, \end{aligned}$$

where  $n$  is the number of observations, and  $x_i$  is the  $i$ th observation.

### Compound gamma

If the shape parameter of the gamma distribution is known, but the inverse-scale parameter is unknown, then a gamma distribution for the inverse-scale forms a conjugate prior. The compound distribution, which results from integrating out the inverse-scale, has a closed form solution, known as the compound gamma distribution.<sup>[29]</sup>

If instead the shape parameter is known but the mean is unknown, with the prior of the mean being given by another gamma distribution, then it results in K-distribution.

## Related distributions and properties

- If  $X \sim \text{Gamma}(1, 1/\lambda)$  (shape -scale parametrization), then  $X$  has an exponential distribution with rate parameter  $\lambda$ .
- If  $X \sim \text{Gamma}(v/2, 2)$  (shape -scale parametrization), then  $X$  is identical to  $\chi^2(v)$ , the chi-squared distribution with  $v$  degrees of freedom. Conversely, if  $Q \sim \chi^2(v)$  and  $c$  is a positive constant, then  $cQ \sim \text{Gamma}(v/2, 2c)$ .
- If  $k$  is an integer, the gamma distribution is an Erlang distribution and is the probability distribution of the waiting time until the  $k$ th "arrival" in a one-dimensional Poisson process with intensity  $1/\theta$ . If

$$X \sim \Gamma(k \in \mathbf{Z}, \theta), \quad Y \sim \text{Pois}\left(\frac{x}{\theta}\right),$$

then

$$P(X > x) = P(Y < k).$$

- If  $X$  has a Maxwell–Boltzmann distribution with parameter  $a$ , then

$$X^2 \sim \Gamma\left(\frac{3}{2}, 2a^2\right).$$

- If  $X \sim \text{Gamma}(k, \theta)$ , then  $\sqrt{X}$  follows a generalized gamma distribution with parameters  $p = 2$ ,  $d = 2k$ , and  $a = \sqrt{\theta}$ .
- More generally, if  $X \sim \text{Gamma}(k, \theta)$ , then  $X^q$  for  $q > 0$  follows a generalized gamma distribution with parameters  $p = 1/q$ ,  $d = k/q$ , and  $a = \theta^q$ .
- If  $X \sim \text{Gamma}(k, \theta)$ , then  $1/X \sim \text{Inv-Gamma}(k, \theta^{-1})$  (see Inverse-gamma distribution for derivation).
- Parametrization 1: If  $\mathbf{X}_k \sim \Gamma(\boldsymbol{\alpha}_k, \theta_k)$  are independent, then  $\frac{\alpha_2 \theta_2 X_1}{\alpha_1 \theta_1 X_2} \sim F(2\alpha_1, 2\alpha_2)$ , or equivalently,  $\frac{X_1}{X_2} \sim \beta'(\alpha_1, \alpha_2, 1, \frac{\theta_1}{\theta_2})$
- Parametrization 2: If  $\mathbf{X}_k \sim \Gamma(\boldsymbol{\alpha}_k, \beta_k)$  are independent, then  $\frac{\alpha_2 \beta_1 X_1}{\alpha_1 \beta_2 X_2} \sim F(2\alpha_1, 2\alpha_2)$ , or equivalently,  $\frac{X_1}{X_2} \sim \beta'(\alpha_1, \alpha_2, 1, \frac{\beta_2}{\beta_1})$
- If  $X \sim \text{Gamma}(\alpha, \theta)$  and  $Y \sim \text{Gamma}(\beta, \theta)$  are independently distributed, then  $X/(X + Y)$  has a beta distribution with parameters  $\alpha$  and  $\beta$ , and  $X/(X + Y)$  is independent of  $X + Y$ , which is  $\text{Gamma}(\alpha + \beta, \theta)$ -distributed.
- If  $X_i \sim \text{Gamma}(\alpha_i, 1)$  are independently distributed, then the vector  $(X_1/S, \dots, X_n/S)$ , where  $S = X_1 + \dots + X_n$ , follows a Dirichlet distribution with parameters  $\alpha_1, \dots, \alpha_n$ .
- For large  $k$  the gamma distribution converges to normal distribution with mean  $\mu = k\theta$  and variance  $\sigma^2 = k\theta^2$ .
- The gamma distribution is the conjugate prior for the precision of the normal distribution with known mean.
- The Wishart distribution is a multivariate generalization of the gamma distribution (samples are positive-definite matrices rather than positive real numbers).
- The gamma distribution is a special case of the generalized gamma distribution, the generalized integer gamma distribution, and the generalized inverse Gaussian distribution.
- Among the discrete distributions, the negative binomial distribution is sometimes considered the discrete analogue of the Gamma distribution.
- Tweedie distributions – the gamma distribution is a member of the family of Tweedie exponential dispersion models.

## Notes

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