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Are the random variables $X + Y$ and $X - Y$ independent if X, Y are distributed normal?

Asked 3 years, 10 months ago Active 10 months ago Viewed 4k times



Let X, Y be independent random variables such that $X, Y \sim N(\mu, \sigma^2)$, show that $X + Y$ and $X - Y$ are independent using the moment generating function.

2

I know that the moment generating function of a sum of independent random variables is the product of the MGF.



So, I'm trying to solve that but i don't know if my process is correct



1

$$M_{X+Y}(t_1, t_2) = M_X(t_1)M_Y(t_2) = M_{N(\mu, \sigma^2)}^2(t) ?$$

[probability](#)

[statistics](#)

[moment-generating-functions](#)

edited Jan 25 '16 at 0:46



Zhanxiong

9,512

1

13

34

asked Jan 25 '16 at 0:36



A P

572

4

15

2 What you need to show is the MGF for the random vector $(X + Y, X - Y)$ can be factorized to the product of MGFs of $X + Y$ and $X - Y$. – Zhanxiong Jan 25 '16 at 0:44

3 Answers



Recall that for an $\mathcal{N}(\mu, \sigma^2)$ random variable, the moment generating function of it is

6

$$M(t) = \exp\left(\mu t + \frac{1}{2}\sigma^2 t^2\right). \quad (1)$$



By condition, $X + Y \sim \mathcal{N}(2\mu, 2\sigma^2)$ and $X - Y \sim \mathcal{N}(0, 2\sigma^2)$. Therefore by (1), we have:

$$M_{X+Y}(t) = \exp(2\mu t + \sigma^2 t^2), \quad M_{X-Y}(t) = \exp(\sigma^2 t^2).$$

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$$\begin{aligned}
 &= E \{ \exp[(t_1 + t_2)X] \times \exp[(t_1 - t_2)Y] \} \\
 &= E \{ \exp[(t_1 + t_2)X] \} \times E \{ \exp[(t_1 - t_2)Y] \} \quad \text{by independence of } X \text{ and } Y. \\
 &= M_X(t_1 + t_2) M_Y(t_1 - t_2) \\
 &= \exp \left(\mu(t_1 + t_2) + \frac{1}{2} \sigma^2 (t_1 + t_2)^2 \right) \exp \left(\mu(t_1 - t_2) + \frac{1}{2} \sigma^2 (t_1 - t_2)^2 \right) \\
 &= \exp(2\mu t_1 + \sigma^2 t_1^2) \exp(\sigma^2 t_2^2) \\
 &= M_{X+Y}(t_1) M_{X-Y}(t_2).
 \end{aligned}$$

Hence $X + Y$ and $X - Y$ are independent.

edited Sep 28 '16 at 13:13

answered Jan 25 '16 at 1:03



Zhanxiong

9,512 1 13 34

▲ I have one question, when you say: $E \{ \exp[(t_1 + t_2)X] \} \times E \{ \exp[(t_1 - t_2)Y] \}$ by independence of X and Y . it is a property?, I only know that $E(XY)=E(X)E(Y)$ but I didn't know that property. – A P Jan 25 '16 at 1:24

1 ▲ Yes, it is exactly this property, where you treat $(t_1 + t_2)X$ and $(t_1 - t_2)Y$ are independent random variables. Notice that t_1 and t_2 are all constants. – Zhanxiong Jan 25 '16 at 1:36

▲ so, when I have 2 independent random variables X, Y $E[g(aX)h(bY)] = E[g(aX)] \cdot E[h(bY)]$ is always true? – A P Jan 25 '16 at 1:44

1 ▲ Yes, you made a good guess. Of course, rigorous, g and h needs to be measurable, which is usually guaranteed. – Zhanxiong Jan 25 '16 at 1:45

▲ it's an interesant result that i didn't know, can you suggest me a site where find the proof to this property? – A P Jan 25 '16 at 1:52

|



No. If X and $2Y = X$, if X is normally distributed so is Y . And $X + Y = 3Y$ and $X - Y = Y$ aren't independent at all.

-1

answered Jan 25 '16 at 0:39



vonbrand

21.2k 6 33 61

▲ OP did mentioned independence of X and Y – Francis Jan 25 '16 at 0:44

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$E[(X + Y)(X - Y)] = E[X^2 - Y^2] = 0 - 0 = 0 = (2\mu) \cdot 0 = E[X + Y]E[X - Y]$. Thus $X + Y$ and $X - Y$ are independent.



answered Jan 25 at 21:12



nullUser

21k 4 45 110