# **Linear Regression**

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# So far ...

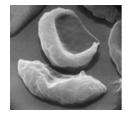
- Learning distributions
  - Maximum Likelihood Estimation (MLE)
  - Maximum A Posteriori (MAP)
- Learning classifiers
  - Naïve Bayes

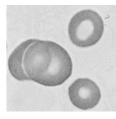
# **Discrete to Continuous Labels**

#### Classification



X = Document Y = Topic





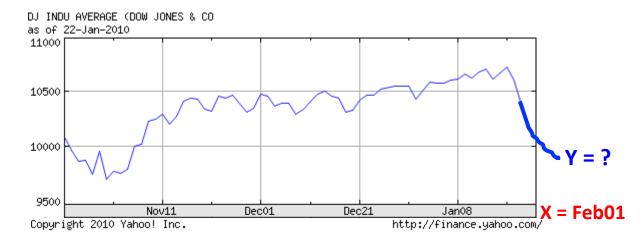
Anemic cell Healthy cell

X = Cell Image

Y = Diagnosis

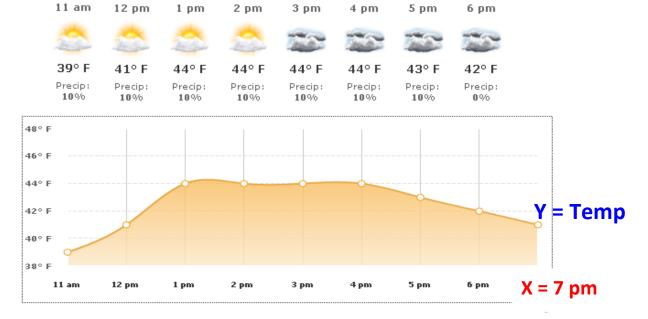
#### Regression

Stock Market Prediction

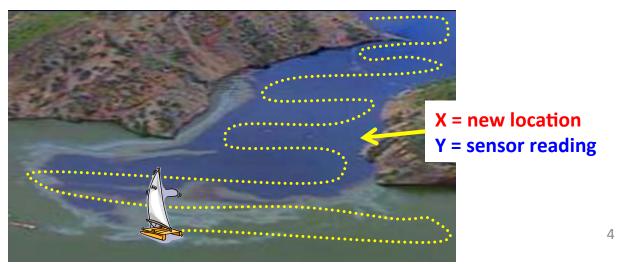


# **Regression Tasks**

Weather Prediction



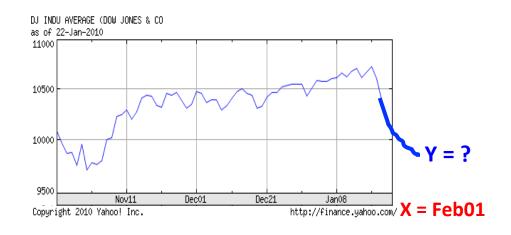
Estimating Contamination



# **Supervised Learning**

Goal: Construct a predictor  $f: X \to Y$  to minimize loss function (performance measure)





#### Classification:

$$P(f(X) \neq Y)$$

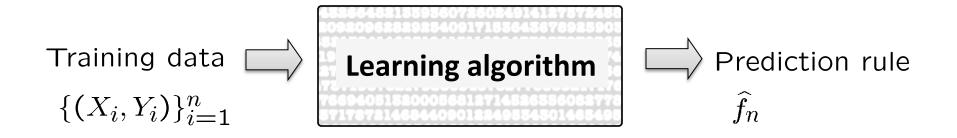
**Probability of Error** 

#### Regression:

$$\mathbb{E}[(f(X) - Y)^2]$$

**Mean Squared Error** 

# Regression algorithms



Linear Regression

Regularized Linear Regression – Ridge regression, Lasso

**Polynomial Regression** 

**Kernel Regression** 

Regression Trees, Splines, Wavelet estimators, ...

# Replace Expectation with Empirical Mean

Optimal predictor: 
$$f^* = \arg\min_{f} \mathbb{E}[(f(X) - Y)^2]$$

Empirical Minimizer: 
$$\widehat{f}_n = \arg\min_{f \in \mathcal{F}} \left( \frac{1}{n} \sum_{i=1}^n (f(X_i) - Y_i)^2 \right)$$

**Empirical mean** 

#### Law of Large Numbers:

$$\frac{1}{n} \sum_{i=1}^{n} \left[ loss(Y_i, f(X_i)) \right] \xrightarrow{\mathsf{n} \to \infty} \mathbb{E}_{XY} \left[ loss(Y, f(X)) \right]$$

# Restrict class of predictors

$$f^* = \arg\min_{f} \mathbb{E}[(f(X) - Y)^2]$$

#### Empirical Minimizer:

$$\widehat{f}_n = \arg\min_{f \in \mathcal{F}} \frac{1}{n} \sum_{i=1}^n (f(X_i) - Y_i)^2$$

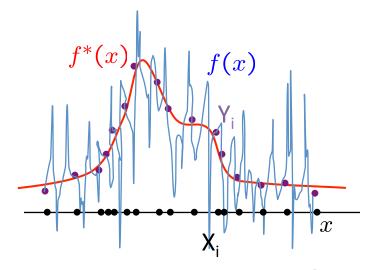
**Class of predictors** 

#### Why?

Overfitting!

Empiricial loss minimized by any function of the form

$$f(x) = \begin{cases} Y_i, & x = X_i \text{ for } i = 1, \dots, n \\ \text{any value,} & \text{otherwise} \end{cases}$$



# Restrict class of predictors

$$f^* = \arg\min_{f} \mathbb{E}[(f(X) - Y)^2]$$

$$\widehat{f}_n = \arg\min_{f \in \mathcal{F}} \frac{1}{n} \sum_{i=1}^n (f(X_i) - Y_i)^2$$

**Class of predictors** 

- ${\mathcal F}$  Class of Linear functions
  - Class of Polynomial functions
  - Class of nonlinear functions

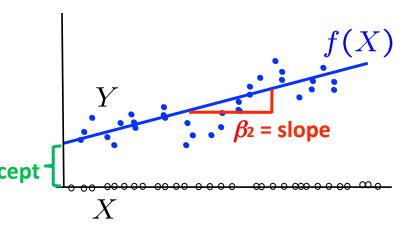
# **Linear Regression**

$$\widehat{f}_n^L = \arg\min_{f \in \mathcal{F}_L} \frac{1}{n} \sum_{i=1}^n (f(X_i) - Y_i)^2 \quad \text{Least Squares Estimator}$$

 $\mathcal{F}_L$  - Class of Linear functions

Uni-variate case:

$$f(X) = \beta_1 + \beta_2 X$$
  $\beta_1$  - intercept



Multi-variate case:

$$f(X) = f(X^{(1)}, \dots, X^{(p)}) = \beta_1 X^{(1)} + \beta_2 X^{(2)} + \dots + \beta_p X^{(p)}$$

$$= X\beta \qquad \text{where} \quad X = [X^{(1)} \dots X^{(p)}], \quad \beta = [\beta_1 \dots \beta_p]^T$$

# **Least Squares Estimator**

$$\widehat{f}_n^L = \arg\min_{f \in \mathcal{F}_L} \frac{1}{n} \sum_{i=1}^n (f(X_i) - Y_i)^2 \qquad f(X_i) = X_i \beta$$



$$\widehat{\beta} = \arg\min_{\beta} \frac{1}{n} \sum_{i=1}^{n} (X_i \beta - Y_i)^2$$
  $\widehat{f}_n^L(X) = X \widehat{\beta}$ 

$$= \arg\min_{\beta} \frac{1}{n} (\mathbf{A}\beta - \mathbf{Y})^T (\mathbf{A}\beta - \mathbf{Y})$$

$$\mathbf{A} = \begin{bmatrix} X_1 \\ \vdots \\ X_n \end{bmatrix} = \begin{bmatrix} X_1^{(1)} & \dots & X_1^{(p)} \\ \vdots & \ddots & \vdots \\ X_n^{(1)} & \dots & X_n^{(p)} \end{bmatrix} \quad \mathbf{Y} = \begin{bmatrix} \mathbf{Y}_1 \\ \vdots \\ \mathbf{Y}_n \end{bmatrix}$$

# **Least Squares Estimator**

$$\widehat{\beta} = \arg\min_{\beta} \frac{1}{n} (\mathbf{A}\beta - \mathbf{Y})^T (\mathbf{A}\beta - \mathbf{Y}) = \arg\min_{\beta} J(\beta)$$

$$J(\beta) = (\mathbf{A}\beta - \mathbf{Y})^T (\mathbf{A}\beta - \mathbf{Y})$$

$$\left. \frac{\partial J(\beta)}{\partial \beta} \right|_{\widehat{\beta}} = 0$$

# **Normal Equations**

$$(\mathbf{A}^T \mathbf{A})\widehat{\beta} = \mathbf{A}^T \mathbf{Y}$$

$$\mathbf{p} \times \mathbf{p} \times \mathbf{p} \times \mathbf{1} \qquad \mathbf{p} \times \mathbf{1}$$

If  $(\mathbf{A}^T\mathbf{A})$  is invertible,

$$\hat{\beta} = (\mathbf{A}^T \mathbf{A})^{-1} \mathbf{A}^T \mathbf{Y}$$
  $\hat{f}_n^L(X) = X \hat{\beta}$ 

When is  $(\mathbf{A}^T\mathbf{A})$  invertible? Recall: Full rank matrices are invertible. What is rank of  $(\mathbf{A}^T\mathbf{A})$ ?

What if  $(\mathbf{A}^T \mathbf{A})$  is not invertible ? Regularization (later)

### **Gradient Descent**

Even when  $(\mathbf{A}^T\mathbf{A})$  is invertible, might be computationally expensive if  $\mathbf{A}$  is huge.

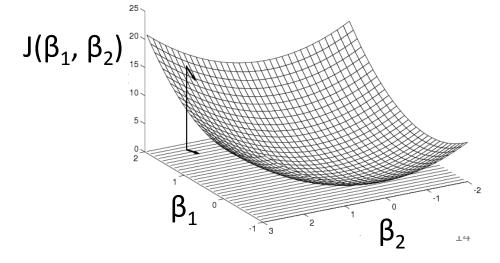
$$\widehat{\beta} = \arg\min_{\beta} \frac{1}{n} (\mathbf{A}\beta - \mathbf{Y})^T (\mathbf{A}\beta - \mathbf{Y}) = \arg\min_{\beta} J(\beta)$$

Treat as optimization problem

Observation:  $J(\beta)$  is convex in  $\beta$ .

# $\frac{\mathsf{J}(\beta_1)}{\beta_1}$

#### How to find the minimizer?



## **Gradient Descent**

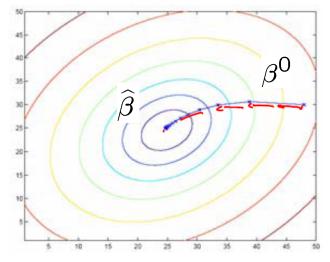
Even when  $(\mathbf{A}^T \mathbf{A})$  is invertible, might be computationally expensive if  $\mathbf{A}$  is huge.

$$\widehat{\beta} = \arg\min_{\beta} \frac{1}{n} (\mathbf{A}\beta - \mathbf{Y})^T (\mathbf{A}\beta - \mathbf{Y}) = \arg\min_{\beta} J(\beta)$$

#### Since $J(\beta)$ is convex, move along negative of gradient

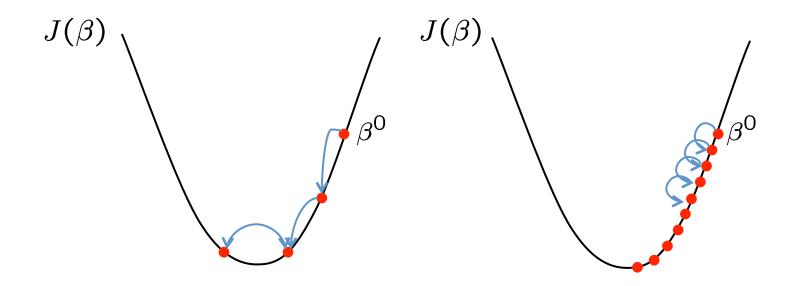
Initialize: 
$$\beta^0$$
 step size

Update:  $\beta^{t+1} = \beta^t - \frac{\alpha}{2} \frac{\partial J(\beta)}{\partial \beta}\Big|_t$ 
 $= \beta^t - \alpha \mathbf{A}^T (\mathbf{A}\beta^t - Y)$ 
 $0 \text{ if } \hat{\beta} = \beta^t$ 



Stop: when some criterion met e.g. fixed # iterations, or  $\frac{\partial J(\beta)}{\partial \beta}\Big|_{\beta^t} < \epsilon$ .

# Effect of step-size α



Large  $\alpha$  => Fast convergence but larger residual error Also possible oscillations

Small  $\alpha$  => Slow convergence but small residual error

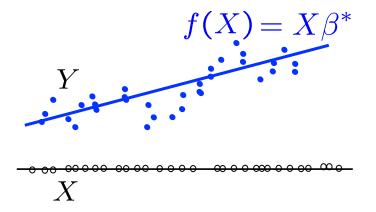
# **Least Squares and MLE**

Intuition: Signal plus (zero-mean) Noise model

$$Y = f^*(X) + \epsilon = X\beta^* + \epsilon$$

$$\epsilon \sim \mathcal{N}(0, \sigma^2 \mathbf{I}) \quad Y \sim \mathcal{N}(X\beta^*, \sigma^2 \mathbf{I})$$

$$\widehat{\beta}_{\text{MLE}} = \arg\max_{\beta} \log p(\{(X_i, Y_i)\}_{i=1}^n | \beta, \sigma^2)$$
 
$$\log \text{ likelihood}$$



$$= \arg\min_{\beta} \sum_{i=1}^{n} (X_i \beta - Y_i)^2 = \widehat{\beta}$$

Least Square Estimate is same as Maximum Likelihood Estimate under a Gaussian model!

# Regularized Least Squares and MAP

What if  $(\mathbf{A}^T \mathbf{A})$  is not invertible ?

$$\widehat{\beta}_{\mathsf{MAP}} = \arg\max_{\beta} \log p(\{(X_i, Y_i)\}_{i=1}^n | \beta, \sigma^2) + \log p(\beta)$$
 
$$\log \text{ likelihood} \qquad \log \text{ prior}$$

I) Gaussian Prior

$$\beta \sim \mathcal{N}(0, \tau^2 \mathbf{I})$$

$$p(\beta) \propto e^{-\beta^T \beta/2\tau^2}$$

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$$\widehat{\beta}_{\mathsf{MAP}} = \arg\min_{\beta} \sum_{i=1}^n (Y_i - X_i\beta)^2 + \lambda \|\beta\|_2^2 \qquad \mathsf{Ridge Regression}$$
 
$$\mathrm{constant}(\sigma^2,\tau^2)$$

$$\widehat{\beta}_{\text{MAP}} = (\boldsymbol{A}^{\top}\boldsymbol{A} + \lambda \boldsymbol{I})^{-1}\boldsymbol{A}^{\top}\boldsymbol{Y}$$

# Regularized Least Squares and MAP

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# Regularized Least Squares and MAP

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$$\log \text{ likelihood} \qquad \log \text{ prior}$$

II) Laplace Prior

$$eta_i \stackrel{iid}{\sim} \mathsf{Laplace}(\mathsf{0},t) \qquad \qquad p(eta_i) \propto e^{-|eta_i|/t}$$

$$p(\beta_i) \propto e^{-|\beta_i|/t}$$

$$\widehat{eta}_{\mathsf{MAP}} = \arg\min_{eta} \sum_{i=1}^n (Y_i - X_i eta)^2 + \lambda \|eta\|_1$$
 Lasso constant  $(\sigma^2, t)$ 

# Ridge Regression vs Lasso

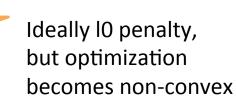
$$\min_{\beta} (\mathbf{A}\beta - \mathbf{Y})^T (\mathbf{A}\beta - \mathbf{Y}) + \lambda \mathrm{pen}(\beta) = \min_{\beta} J(\beta) + \lambda \mathrm{pen}(\beta)$$

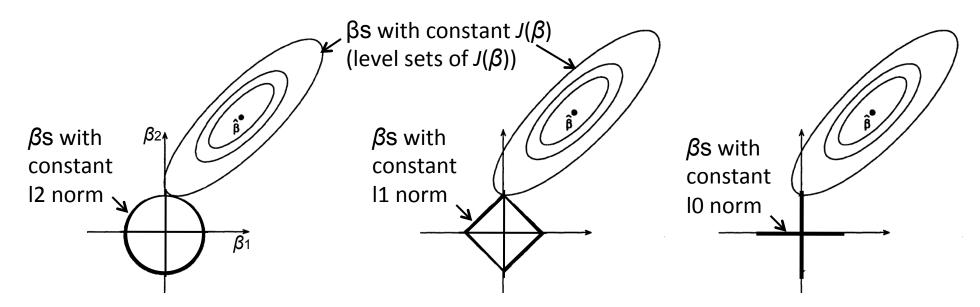
Ridge Regression:

$$pen(\beta) = \|\beta\|_2^2$$

Lasso:

$$pen(\beta) = \|\beta\|_1$$

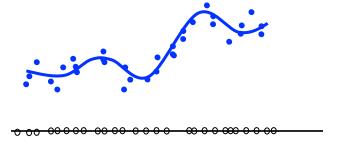




Lasso (I1 penalty) results in sparse solutions – vector with more zero coordinates Good for high-dimensional problems – don't have to store all coordinates!

# **Beyond Linear Regression**

Polynomial regression



Regression with nonlinear features

Later ...

Kernel regression - Local/Weighted regression

# Polynomial Regression

degree m

Univariate (1-dim)  $f(X) = \beta_0 + \beta_1 X + \beta_2 X^2 + \dots + \beta_m X^m = X\beta$ case:

where 
$$\mathbf{X} = [1 \ X \ X^2 \dots X^m], \ \beta = [\beta_1 \dots \beta_m]^T$$

$$\widehat{\beta} = (\mathbf{A}^T \mathbf{A})^{-1} \mathbf{A}^T \mathbf{Y}$$

$$\widehat{f}_n(X) = \mathbf{X} \widehat{\beta}$$

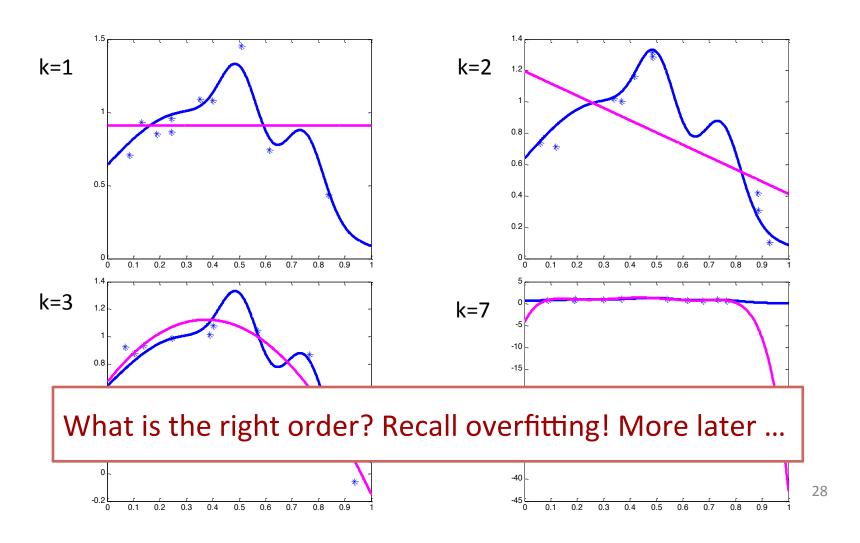
$$\widehat{f}_n(X) = \mathbf{X}\widehat{\beta}$$

$$\mathbf{A} = \begin{bmatrix} 1 & X_1 & X_1^2 & \dots & X_1^m \\ \vdots & & \ddots & \vdots \\ 1 & X_n & X_n^2 & \dots & X_n^m \end{bmatrix}$$

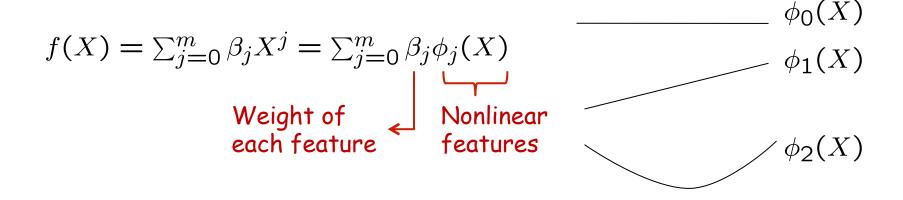
Multivariate (p-dim) 
$$f(X) = \beta_0 + \beta_1 X^{(1)} + \beta_2 X^{(2)} + \dots + \beta_p X^{(p)}$$
 case: 
$$+ \sum_{i=1}^p \sum_{j=1}^p \beta_{ij} X^{(i)} X^{(j)} + \sum_{i=1}^p \sum_{j=1}^p \sum_{k=1}^p X^{(i)} X^{(j)} X^{(k)} + \dots \text{ terms up to degree m}$$

# **Polynomial Regression**

Polynomial of order k, equivalently of degree up to k-1



# Regression with nonlinear features



In general, use any nonlinear features

# What you should know

#### **Linear Regression**

**Least Squares Estimator** 

**Normal Equations** 

**Gradient Descent** 

Probabilistic Interpretation (connection to MLE)

Regularized Linear Regression (connection to MAP)

Ridge Regression, Lasso

Polynomial Regression, Regression with Non-linear features