



## $N(\theta, \theta)$ : MLE for a Normal where mean=variance

Asked 6 years, 6 months ago   Active 1 year, 5 months ago   Viewed 6k times

▲ For an  $n$ -sample following a Normal( $\mu = \theta, \sigma^2 = \theta$ ), how do we find the mle?

7 I can find the root of the score function

$$\theta = \frac{1 \pm \sqrt{1 - 4\frac{s}{n}}}{2}, s = \sum x_i^2,$$



1 but I don't see which one is the maximum.  
I tried to substitute in the second derivative of the log-likelihood, without success.

For the likelihood, with  $x = (x_1, x_2, \dots, x_n)$ ,

$$f(x) = (2\pi)^{-n/2} \theta^{-n/2} \exp\left(-\frac{1}{2\theta} \sum (x_i - \theta)^2\right),$$

then, with  $s = \sum x_i^2$  and  $t = \sum x_i$ ,

$$\ln f(x) = -\frac{n}{2} \ln(2\pi) - \frac{n}{2} \ln \theta - \frac{s}{2\theta} - t + \frac{n}{2} \theta,$$

so that

$$\partial_\theta \ln f(x) = -\frac{n}{2} \frac{1}{\theta} + \frac{s}{2\theta^2} + \frac{n}{2},$$

and the roots are given by

$$\theta^2 - \theta + \frac{s}{n} = 0.$$

Also,



$$\partial_{\theta\theta} \ln f(x) = -\frac{n}{2} \frac{1}{\theta^2} - \frac{s}{\theta^3}.$$

$$-\frac{1}{2} \log \theta - \frac{(x-\theta)^2}{2\theta}$$

maximum-likelihood

asked Apr 16 '13 at 19:39

user21186

2  It looks to me like there might be an error in your calculation of  $\log f(x)$ . I think it should be  $\text{const} - \frac{n}{2} \log(\theta) - \frac{s}{2\theta} + t - \frac{n\theta}{2}$ . As is, there is a positive probability chance that  $1 - 4\frac{s}{n} < 0$  which is a problem. – [guy](#) Apr 16 '13 at 19:57 

## 2 Answers



There are some typos (or algebraical mistakes) in the signs of the log-likelihood, followed by the corresponding unpleasant consequences.

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Since this is a well-known problem, I will only point out a reference with the solution:



[Asymptotic Theory of Statistics and Probability](#) pp. 53, by Anirban DasGupta.



answered Apr 16 '13 at 19:59

[Gustav](#)

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Recall that the normal distribution  $N(\mu, \sigma^2)$  has pdf  $f(x | \mu, \sigma^2) = \frac{1}{\sqrt{2\pi\sigma^2}} \exp\left(-\frac{(x-\mu)^2}{2\sigma^2}\right)$ , Note here that  $\mu = \theta$  and  $\sigma^2 = \theta$  and therefore

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$\sigma = \sqrt{\theta}$



$$\begin{aligned}
L(x_1, x_2, \dots, x_n | \theta) &= \prod_{i=1}^n f(x_i | \theta) \\
&= \prod_{i=1}^n \frac{1}{\sqrt{2\pi\theta}} \exp \left\{ -\frac{1}{2\theta} (x_i - \theta)^2 \right\} \\
&= (2\pi)^{-n/2} (\theta)^{-n/2} \prod_{i=1}^n \exp \left\{ -\frac{1}{2\theta} (x_i - \theta)^2 \right\} \\
&= (2\pi)^{-n/2} (\theta)^{-n/2} \exp \left\{ -\frac{1}{2\theta} \sum_{i=1}^n (x_i - \theta)^2 \right\} \\
\log L &= -\frac{n}{2} \log(2\pi) - \frac{n}{2} \log(\theta) - \frac{1}{2\theta} \sum_{i=1}^n (x_i - \theta)^2
\end{aligned}$$

Consider the term  $\frac{1}{2\theta} \sum_{i=1}^n (x_i - \theta)^2$  which can be expanded and simplified

$$\begin{aligned}
\frac{1}{2\theta} \sum_{i=1}^n (x_i - \theta)^2 &= \frac{1}{2\theta} \sum_{i=1}^n (x_i - \theta)(x_i - \theta) \\
&= \frac{1}{2\theta} \sum_{i=1}^n (x_i^2 - 2\theta x_i + \theta^2) \\
&= \frac{1}{2\theta} \left( \sum_{i=1}^n (x_i^2) - 2\theta \sum_{i=1}^n (x_i) + n\theta^2 \right) \\
&= \frac{1}{2\theta} \sum_{i=1}^n (x_i^2) - \sum_{i=1}^n (x_i) + \frac{n\theta}{2}
\end{aligned}$$

We can now compute the derivative with respect to  $\theta$ , equate to zero and solve for  $\theta$

$$\begin{aligned}\log L &= -\frac{n}{2}\log(2\pi) - \frac{n}{2}\log(\theta) - \left( \frac{1}{2\theta} \sum_{i=1}^n (x_i^2) - \sum_{i=1}^n (x_i) + \frac{n\theta}{2} \right) \\ \frac{d}{d\theta} \log L &= \frac{-n}{2\theta} - \left( \frac{-1}{2\theta^2} \sum_{i=1}^n (x_i^2) + \frac{n}{2} \right) = 0 \\ &= \frac{-n}{2\theta} + \frac{1}{2\theta^2} \sum_{i=1}^n (x_i^2) - \frac{n}{2} \\ &= -\theta^2 - \theta + \frac{1}{n} \sum_{i=1}^n (x_i^2) \\ \text{let } s &= \frac{1}{n} \sum_{i=1}^n (x_i^2) \\ 0 &= -\theta^2 - \theta + s \\ \hat{\theta} &= \frac{\sqrt{1+4s} - 1}{2}\end{aligned}$$

edited Apr 23 '18 at 14:30

answered Apr 21 '18 at 17:21



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