

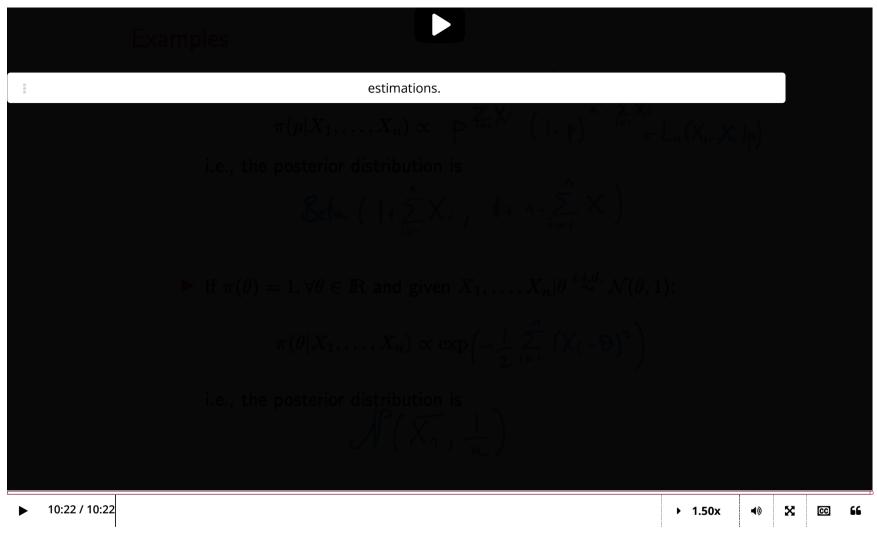


Lecture 18: Jeffreys Prior and

<u>Course</u> > <u>Unit 5 Bayesian statistics</u> > <u>Bayesian Confidence Interval</u>

> 5. Improper Prior: Example

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Gaussian Prior on Gaussian Observations

2/2 points (graded)

Let X_1, X_2, \ldots, X_n be random variables; which are i.i.d., conditional on heta, and such that,

$$p\left(X_{i}| heta
ight)=N\left(heta,1
ight),$$

where $p\left(X_i| heta
ight)$ is the conditonal density of X_i given heta. Furthermore, we assume the prior $\pi\left(heta
ight)\sim N\left(\mu,1
ight)$. Let

$$\pi\left(heta|X_{1},\ldots,X_{n}
ight)\sim N\left(lpha,eta^{2}
ight).$$

Find, α and β^2 .

- α =
 - $igodesign rac{1}{n+1}((\sum_{i=1}^n X_i) + \mu)$
 - igcirc $rac{1}{n}(\sum_{i=1}^n X_i + \mu)$
 - $\bigcap rac{1}{n+1}(\sum_{i=1}^n X_i) + \mu$
 - $igcup_{n} rac{1}{n} (\sum_{i=1}^{n} X_i)$
 - ~
- $\beta^2 =$
 - $\bigcirc \frac{1}{n+1}$
 - $\bigcirc \frac{1}{n}$
 - $\bigcirc \frac{\mu}{n}$

$$\bigcirc \frac{\mu}{n+1}$$



Solution:

We begin by recalling that,

$$\pi\left(heta|X_1,\ldots,X_n
ight) \, \propto p_n\left(X_1,\ldots,X_n| heta
ight)\pi\left(heta
ight) \ = \left(-rac{1}{2}\sum_{i=1}^n\left(X_i- heta
ight)^2
ight)\exp\left(-rac{1}{2}(heta-\mu)^2
ight).$$

We now study the last quantity, keeping in mind that, Gaussian distribution is a conjugate prior of itself; hence, we expect the resulting distribution to be a Gaussian. For this, we need to arrive at a formula of the form,

$$\exp\left(-rac{1}{2}\sum_{i=1}^n{(X_i- heta)^2}
ight)\exp\left(-rac{1}{2}(heta-\mu)^2
ight)\propto \exp\left(-rac{1}{2eta^2}(heta-lpha)^2
ight).$$

Now, we begin doing the algebra, after removing \exp 's from both sides.

$$egin{aligned} &-rac{1}{2}\sum_{i=1}^{n}\left(X_{i}- heta
ight)^{2}-rac{1}{2}(heta-\mu)^{2}\ &=-rac{1}{2}igg((n+1)\, heta^{2}-2 heta\left(\left[\sum_{i=1}^{n}X_{i}
ight]+\mu
ight)igg)+C\ &=-rac{n+1}{2}igg(heta^{2}-2 hetarac{\left[\sum_{i=1}^{n}X_{i}
ight]+\mu}{n+1}igg)+C\ &=-rac{1}{2(n+1)^{-1}}(heta-lpha)^{2}+C', \end{aligned}$$

where C and C''s are constants; and.

$$lpha = rac{1}{n+1} \Biggl(\Biggl[\sum_{i=1}^n X_i \Biggr] + \mu \Biggr) \quad ext{and} \quad eta^2 = rac{1}{n+1}.$$

This is, essentially, the same formula as derived in the lecture, except that n is replaced with n+1 (since we have a prior now), and the summation also takes μ into account. In a sense, you may view this as, n+1 observations, where the first n are X_1,\ldots,X_n and the last one is from N ($\mu,1$).

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• Answers are displayed within the problem

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lf you are stuck on this one, consider referencing Lecture 15 from 6.431x

It discusses recognizing normal PDFs and their mean and variance when they are represented in perhaps a less familiar fashion.

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