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[Lecture 11: Fisher Information,](#)
[Asymptotic Normality of MLE;](#)

[Course](#) > [Unit 3 Methods of Estimation](#) > [Method of Moments](#)

> 3. Fisher Information

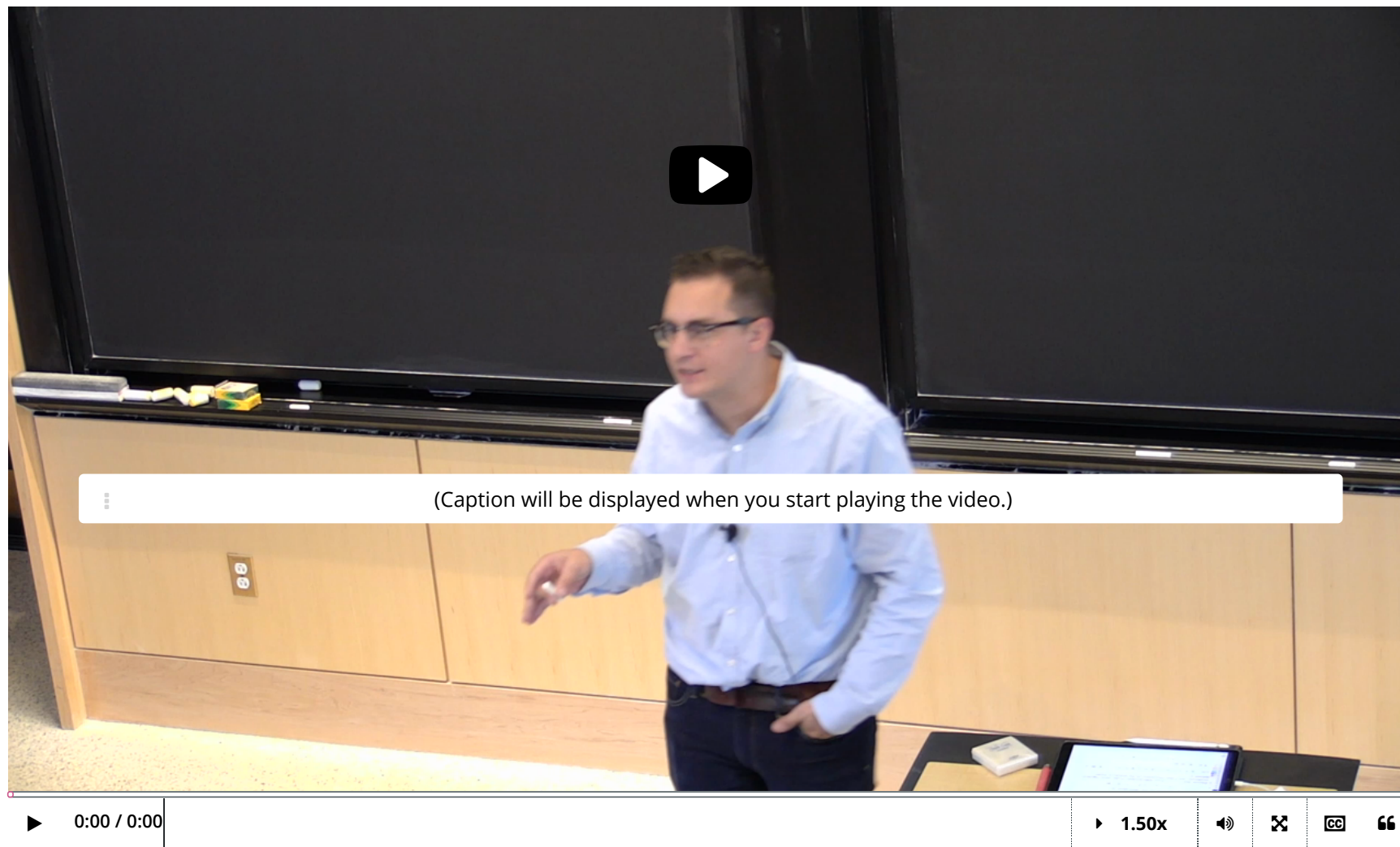
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3. Fisher Information

Fisher Information: Definitions



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Let $(\mathbb{R}, \{\mathbf{P}_\theta\}_{\theta \in \mathbb{R}})$ denote a continuous statistical model. Let $f_\theta(x)$ denote the pdf (probability density function) of the continuous distribution \mathbf{P}_θ . Assume that $f_\theta(x)$ is twice-differentiable as a function of the parameter θ .

In the next few problems, you will derive the formula

$$\mathcal{I}(\theta) = \int_{-\infty}^{\infty} \frac{\left(\frac{\partial f_{\theta}(x)}{\partial \theta}\right)^2}{f_{\theta}(x)} dx$$

using the definition $\mathcal{I}(\theta) = \text{Var}(\ell'(\theta))$ and the basic formula $\text{Var}(X) = \mathbb{E}[X^2] - \mathbb{E}[X]^2$ for any random variable X .

For computations, it is sometimes convenient to use the above formula for the Fisher information.

Note: The derivation in the next set of problems is presented as a proof in the video that follows, but we encourage you to attempt these problems before watching the video.

Deriving a Useful Formula for the Fisher Information I

2/2 points (graded)

Let $(\mathbb{R}, \{\mathbf{P}_{\theta}\}_{\theta \in \mathbb{R}})$ denote a statistical model for a continuous distribution \mathbf{P}_{θ} . Let f_{θ} denote the pdf (probability density function) of the continuous distribution \mathbf{P}_{θ} . Recall that

$$\int_{-\infty}^{\infty} f_{\theta}(x) dx = 1$$

for all $\theta \in \mathbb{R}$.

For the next two questions, assume that you are allowed to interchange derivatives and integrals.

What is

$$\int_{-\infty}^{\infty} \frac{\partial}{\partial \theta} f_{\theta}(x) dx \text{ ?}$$

✓ Answer: 0.0

What is

$$\int_{-\infty}^{\infty} \frac{\partial^2}{\partial \theta^2} f_{\theta}(x) dx ?$$

✓ Answer: 0.0

Solution:

Since we know $\int_{-\infty}^{\infty} f_{\theta}(x) dx = 1$, this implies that

$$\frac{\partial}{\partial \theta} \int_{-\infty}^{\infty} f_{\theta}(x) dx = \frac{\partial}{\partial \theta} 1 = 0.$$

Since we are allowed to interchange the integral and derivative, this implies that

$$\int_{-\infty}^{\infty} \frac{\partial}{\partial \theta} f_{\theta}(x) dx = 0.$$

Similarly for the second derivative,

$$\int_{-\infty}^{\infty} \frac{\partial^2}{\partial \theta^2} f_{\theta}(x) dx = \frac{\partial^2}{\partial \theta^2} \int_{-\infty}^{\infty} f_{\theta}(x) dx = \frac{\partial^2}{\partial \theta^2} 1 = 0.$$

Remark: If f is "nice enough," analytically speaking, then we can rigorously justify interchanging the integral and derivative.

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You have used 1 of 2 attempts

i Answers are displayed within the problem

Deriving a Useful Formula for the Fisher Information II

1/1 point (graded)

As before, let f_θ denote the pdf (probability density function) of the continuous distribution \mathbf{P}_θ . By definition,

$$\ell(\theta) = \ln L_1(X, \theta) = \ln f_\theta(X)$$

where $X \sim \mathbf{P}_\theta$. Differentiating, we see

$$\ell'(\theta) = \frac{\partial}{\partial \theta} \ln f_\theta(X) = \frac{\frac{\partial}{\partial \theta} f_\theta(X)}{f_\theta(X)}.$$

What is

$$\mathbb{E}[\ell'(\theta)] = \mathbb{E}\left[\frac{\frac{\partial}{\partial \theta} f_\theta(X)}{f_\theta(X)}\right]?$$

0

✓ Answer: 0.0

(Note that $X \sim \mathbf{P}_\theta$.)

STANDARD NOTATION

Solution:

Observe that

$$\mathbb{E} \left[\frac{\frac{\partial}{\partial \theta} f_{\theta}(X)}{f_{\theta}(X)} \right] = \int_{-\infty}^{\infty} \left(\frac{\frac{\partial}{\partial \theta} f_{\theta}(x)}{f_{\theta}(x)} \right) f_{\theta}(x) dx = \int_{-\infty}^{\infty} \frac{\partial}{\partial \theta} f_{\theta}(x) dx = 0,$$

by the computation in the previous question.

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Deriving a Useful Formula for the Fisher Information III

1/1 point (graded)

As before, let f_{θ} denote the pdf (probability density function) of the continuous distribution \mathbf{P}_{θ} . By definition,

$$\ell(\theta) = \ln L_1(X, \theta) = \ln f_{\theta}(X)$$

where $X \sim \mathbf{P}_{\theta}$.

Using the previous question, which of the following are equal to $\text{Var}(\ell'(\theta)) = \text{Var}\left(\frac{\partial}{\partial \theta} \ln f_{\theta}(X)\right)$? (Choose all that apply.)

☒ $\mathcal{I}(\theta)$

☐ $\mathbb{E} [\ell' (\theta)]$

☒ $\mathbb{E} [(\ell' (\theta))^2]$

☒ $\int_{-\infty}^{\infty} \frac{(\frac{\partial}{\partial \theta} f_{\theta}(x))^2}{f_{\theta}(x)} dx$



Solution:

We consider the choices in order.

- By definition, $\mathcal{I} (\theta) = \text{Var} (\ell' (\theta))$, so the first answer choice $\mathcal{I} (\theta)$ is correct.
- By the previous question, $\mathbb{E} [\ell' (\theta)] = 0$, so this answer choice is incorrect.
- By definition of variance,

$$\text{Var} (\ell' (\theta)) = \mathbb{E} [\ell' (\theta)^2] - \mathbb{E} [\ell' (\theta)]^2,$$

and $\mathbb{E} [\ell' (\theta)] = 0$, by the previous question. Hence, $\mathbb{E} [(\ell' (\theta))^2] = \text{Var} (\ell' (\theta))$, and so the answer choice $\mathbb{E} [(\ell' (\theta))^2]$ is correct.

- The last choice $\int_{-\infty}^{\infty} \frac{(\frac{\partial}{\partial \theta} f_{\theta}(x))^2}{f_{\theta}(x)} dx$ is correct because, using the previous bullet,

$$\begin{aligned} \text{Var} (\ell' (\theta)) &= \mathbb{E} [(\ell' (\theta))^2] \\ &= \mathbb{E} \left[\left(\frac{\frac{\partial}{\partial \theta} f_{\theta} (X)}{f_{\theta} (X)} \right)^2 \right] \end{aligned}$$

$$= \int_{-\infty}^{\infty} \frac{\left(\frac{\partial}{\partial \theta} f_{\theta}(x)\right)^2}{f_{\theta}(x)} dx.$$

Remark: A convenient way to compute the Fisher information is to use the fourth answer choice, which gives the useful formula

$$\mathcal{I}(\theta) = \int_{-\infty}^{\infty} \frac{\left(\frac{\partial}{\partial \theta} f_{\theta}(x)\right)^2}{f_{\theta}(x)} dx.$$

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You have used 1 of 3 attempts

i Answers are displayed within the problem

Proof of Fisher Information Equivalent Formulas for 1 Dimension

$$I(\theta) = \text{Cov}(\nabla \ell(\theta)) = E[\nabla \ell(\theta) \nabla \ell(\theta)^T] - E[\nabla \ell(\theta)] E[\nabla \ell(\theta)]^T$$

$$I(\theta) \text{ is a } d \times d \text{ matrix called Fisher Information}$$

Theorem:
$$I(\theta) = -E[H \ell(\theta)]$$

that the variance of L prime is minus the expectation of L

$$X \text{ is continuous with pdf } f_\theta(x) = L_1(x, \theta)$$

$$\int f_\theta(x) dx = 1 \quad \frac{\partial}{\partial \theta} \int f_\theta(x) dx = 0$$

also

$$\int \frac{\partial}{\partial \theta} L_1(x, \theta) dx = 0 \quad (1)$$

$$\int \frac{\partial^2}{\partial \theta^2} L_1(x, \theta) dx = 0 \quad (2)$$

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▶ 1.50x 🔊 🗨️ 📄 🗨️

Video

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Definition of Fisher Information

$$\mathcal{I}(\theta) = \text{Cov}(\nabla \ell(\theta)) = -\mathbb{E}[\mathbf{H}\ell(\theta)],$$

The definition when the distribution has a pmf $p_\theta(\mathbf{x})$ is also the same, with the expectation taken with respect to the pmf.

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