

EdX and its Members use cookies and other tracking technologies for performance, analytics, and marketing purposes. By using this website, you accept this use. Learn more about these technologies in the [Privacy Policy](#). ×



[Unit 0. Course Overview, Syllabus, Guidelines, and Homework on Course](#) > [Prerequisites](#) > [Homework 0: Probability and Linear algebra Review](#) > 4. Uniform random variables

**Audit Access Expires Dec 24, 2019**

You lose all access to this course, including your progress, on Dec 24, 2019.

Upgrade by Nov 4, 2019 to get unlimited access to the course as long as it exists on the site. [Upgrade now](#)

## 4. Uniform random variables

### Expectation, variance and probabilities

4/4 points (graded)

Let  $X$  be a uniform random variable in the interval  $[2, 8.5]$ . Find the following quantities (if needed, round to the nearest  $10^{-4}$ ):

$\mathbb{E}[X] =$

5.25



$$\text{Var}[X] = \boxed{3.520833} \quad \checkmark$$

$$\mathbf{P}(X > 4) = \boxed{0.6923076923076923} \quad \checkmark$$

$$\mathbf{P}(\log(X) \leq 1) = \boxed{0.11050489668600694} \quad \checkmark$$

(Note that the logarithm is the natural one to base  $e$ )

STANDARD NOTATION

Submit

You have used 3 of 4 attempts

---

✓ Correct (4/4 points)

---

## Two independent copies

3/3 points (graded)

Let  $U, V$  be i.i.d. random variables uniformly distributed in  $[0, 1]$ . Compute the following quantities:

$$\mathbb{E}[|U - V|] = \boxed{1/3} \quad \checkmark \text{ Answer: } 1/3$$

$$\mathbf{P}(U = V) = \boxed{0} \quad \checkmark \text{ Answer: } 0$$

$$\mathbf{P}(U \leq V) = \boxed{1/2}$$

✓ Answer: 1/2

STANDARD NOTATION

### Solution:

For the first quantity, we write the joint expectation as an iterated expectation and conditional expectation,

$$\mathbb{E}[|U - V|] = \mathbb{E}[\mathbb{E}[|U - V| | V]].$$

By independence, we can compute the inner expectation as

$$\begin{aligned} \mathbb{E}[|U - V| | V = v] &= \int_0^1 |u - v| \, du \\ &= \int_0^v (v - u) \, du + \int_v^1 (u - v) \, du \\ &= \left[ vu - \frac{1}{2}u^2 \right]_0^v + \left[ \frac{1}{2}u^2 - vu \right]_v^1 = v^2 - \frac{1}{2}v^2 + \frac{1}{2} - v + \frac{1}{2}v^2 + v^2 \\ &= v^2 - v + \frac{1}{2}, \end{aligned}$$

so

$$\mathbb{E}[|U - V|] = \mathbb{E}\left[V^2 - V + \frac{1}{2}\right] = \frac{1}{3} - \frac{1}{2} + \frac{1}{2} = \frac{1}{3}.$$

For the probability  $\mathbf{P}(U = V)$ , just write this as double expectation as well and notice that

$$\mathbf{P}(U = V) = \mathbb{E}[\mathbb{E}[\mathbf{1}(U = V) | V]] = \mathbb{E}[0] = 0,$$

because the probability of a uniform random variable being equal to any fixed number between 0 and 1 is zero.

For  $\mathbf{P}(U \leq V)$ , write it again as a double expectation,

$$\mathbf{P}(U \leq V) = \mathbb{E}[\mathbb{E}[\mathbf{1}(U \leq V) | V]] = \mathbb{E}[\mathbf{P}(U \leq V) | V] = \mathbb{E}[V] = \frac{1}{2}.$$

Alternatively, this can also be seen by symmetry of the two variables, i.e.,  $P(U \leq V) = P(V \leq U)$  and either one of the two must be true, counting double the zero-set of  $\mathbf{P}(U = V)$ .

: Uniform PDF in Lecture 8, *Probability density functions*.

Submit

You have used 2 of 3 attempts

**i** Answers are displayed within the problem

## Maximum and sum of independent copies

1/1 point (graded)

Let  $X, Y$  be independent random variables uniformly distributed in  $[0, 1]$ . In the graph below, sketch

1. the probability density  $f_{X+Y}(z)$  of  $X + Y$ ;
2. the probability density  $f_{\max(X,Y)}(z)$  of  $\max(X, Y)$ .

(Be sure to sketch on the **entire domain** shown on the graph.)

**Drawing tip:** The spline tool draws a smooth curve connecting the points you click. To draw sharp corners, click on the point where the corner would be, then click again very close to it, and then continue onto the next point of your function.



Answer: See solution.



STANDARD NOTATION

**Solution:**





The density of  $X + Y$  is given by the convolution of the density of a uniform random variable,

$$f(x) = \mathbf{1}_{[0,1]} = \begin{cases} 1, & \text{if } x \in [0,1] \\ 0, & \text{otherwise} \end{cases}$$

The density  $g$  of  $X + Y$  therefore is

$$\begin{aligned} g(z) &= \int_{\mathbb{R}} f(x) f(z-x) dx \\ &= \int_{\mathbb{R}} \mathbf{1}(x \in [0,1]) \mathbf{1}(z-x \in [0,1]) dx \\ &= \int_0^1 \mathbf{1}(z-1 \leq x \leq z) dx \\ &= \mathbf{1}(z \leq 2) \int_{\max\{0, z-1\}}^{\min\{1, z\}} dx \\ &= \begin{cases} 0, & z < 0 \\ z, & 0 < z < 1 \\ 2-z, & 1 < z < 2 \\ 0, & z > 2. \end{cases} \end{aligned}$$

For the density of  $\max\{X, Y\}$ , first note that it is supported in  $[0, 1]$ . Now, first compute the cdf on that interval:

$$\mathbf{P}(\max\{X, Y\} \leq y) = \mathbf{P}(X \leq y) \mathbf{P}(Y \leq y) \quad (\text{by independence})$$

$$= t^2.$$

Hence, the density  $h$  of  $\max\{X, Y\}$  is given by

$$h(z) = \begin{cases} 0, & z < 0 \\ 2z, & 0 \leq z \leq 1 \\ 0, & z > 1. \end{cases}$$

Submit

You have used 3 of 10 attempts

**i** Answers are displayed within the problem

## Maximum of uniform random variables

2/2 points (graded)

Let  $U_1, \dots, U_n$  be i.i.d. random variables uniformly distributed in  $[0, 1]$  and let  $M_n = \max_{1 \leq i \leq n} U_i$ .

Find the cdf of  $M_n$ , which we denote by  $G(t)$ , for  $t \in [0, 1]$ .

For  $t \in [0, 1]$ ,

$G(t) =$  $t^n$ ✓ Answer:  $t^n$  $t^n$ 

Now, let  $F_n(t)$  denote the cdf of  $n(1 - M_n)$ ; for  $t > 0$ , compute

 $\lim_{n \rightarrow \infty} F_n(t) =$  $1 - \exp(-t)$ ✓ Answer:  $1 - \exp(-t)$  $1 - \exp(-t)$ 

STANDARD NOTATION

**Solution:**

First, we compute the cdf. Let  $t \in [0, 1]$ . Then,

$$\mathbf{P}(M_n \leq t) = \mathbf{P}\left(\max_{i=1, \dots, n} U_i \leq t\right) = \mathbf{P}\left(\bigcap_{i=1}^n \{U_i \leq t\}\right) = \prod_{i=1}^n \mathbf{P}(U_i \leq t) = t^n,$$

where we used the independence of the  $U_i$  to write the intersection as a product.

Now,

$$\mathbf{P}(n(1 - M_n) \leq t) = \mathbf{P}\left(1 - M_n \leq \frac{t}{n}\right) = \mathbf{P}\left(M_n \geq 1 - \frac{t}{n}\right)$$

$$= 1 - \mathbf{P} \left( M_n < 1 - \frac{t}{n} \right) = 1 - \left( 1 - \frac{t}{n} \right)^n \xrightarrow{n \rightarrow \infty} 1 - \mathbf{e}^{-t}.$$

Hence,  $n(1 - M_n)$  converges in distribution to  $\text{Exp}(1)$ .

Submit

You have used 2 of 3 attempts

**i** Answers are displayed within the problem

## Discussion

Hide Discussion

**Topic:** Unit 0. Course Overview, Syllabus, Guidelines, and Homework on  
**Prerequisites:** Homework 0: Probability and Linear algebra Review / 4. Uniform random variables

Add a Post

◀ All Posts

### Not getting the pdf of Max(X,Y).

discussion posted 4 days ago by [sourabhXIII](#)

Can someone please give me some hint. Spent lot of time but couldn't get it right. Got the PDF as derivative of CDF; plotted it but grader disagrees. :(

Guide: [http://www.stankova.net/statistics\\_2012/lecture\\_12.pdf](http://www.stankova.net/statistics_2012/lecture_12.pdf) page 26.

This post is visible to everyone.



3 responses

[Add a Response](#)**mrBB**

4 days ago



Not sure what to add to the slide you refer to. Make sure you draw the pdf over it's entire (visible) domain, i.e. from  $-1$  to  $3$ .



I am clueless, how wrong this straightline can be.

posted 4 days ago by **sourabhXIII**

**GiacomoDemarie**

4 days ago



To find a PDF of  $\max\{X,Y\}$ , start evaluating the CDF.

$\text{CDF}(z) = P(z) = \text{Prob}(Z < z) = \text{Prob}(\max\{X,Y\} < z)$ .

Then think which condition must be satisfied so that  $\max\{X,Y\}$  is less than  $z$ .

Hope this helps.

...

I get the CDF as  $\text{Prob}(X < z)P(Y < z)$  Then differentiate it to get the PDF. How wrong my drawing of a straight line can be.

posted 4 days ago by [sourabhXIII](#)

Add a comment

**[sourabhXIII](#)**

4 days ago

+

...

Finally done with it. It was a classic stupidity from my end! Thank you guys.

...

Good to hear you figured it out!

posted 4 days ago by [mrBB](#)

Add a comment

Showing all responses

Add a response:

Preview

Submit

Learn About Verified Certificates

© All Rights Reserved