Exercise of Supervised Learning: Regularization Part 1

Yawei Li

yawei.li@stat.uni-muenchen.de

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Exercise 1: L0 Regularization

Consider the regression learning setting, i.e., $\mathcal{Y} = \mathbb{R}$, and the feature space $\mathcal{X} = \mathbb{R}^p$. Let the hypothesis space be the linear models:

$$\mathcal{H} = \{ f(\mathbf{x}) = \boldsymbol{\theta}^\mathsf{T} \mathbf{x} \mid \boldsymbol{\theta} \in \mathbb{R}^p \}.$$

Suppose your loss function of interest is the L2 loss $L(y, f(\mathbf{x})) = \frac{1}{2}(y - f(\mathbf{x}))^2$. Consider the L_0 -regularized empirical risk of a model $f(\mathbf{x} \mid \theta)$:

$$\mathcal{R}_{\mathsf{reg}}(oldsymbol{ heta}) = \mathcal{R}_{\mathsf{emp}}(oldsymbol{ heta}) + \lambda ||oldsymbol{ heta}||_0 = rac{1}{2} \sum_{i=1}^n (y^{(i)} - oldsymbol{ heta}^\mathsf{T} \mathbf{x}^{(i)})^2 + \lambda \sum_{i=1}^p I_{|oldsymbol{ heta}_i|
eq 0} \qquad riangleleft$$

Assume that $\mathbf{X}^T\mathbf{X} = \mathbf{I}$, which holds if \mathbf{X} has orthonormal columns. Show that the minimizer $\hat{\theta}_{L0} = (\hat{\theta}_{L0,1}, \dots, \hat{\theta}_{L0,p})^T$ is given by

$$\hat{\theta}_{L0,i} = \hat{\theta}_i \mathbf{I}_{|\hat{\theta}_i| > \sqrt{2\lambda}}, \qquad i = 1, \dots, p,$$

where $\hat{\boldsymbol{\theta}} = (\hat{\theta}_1, \dots, \hat{\theta}_p)^T = (\mathbf{X}^T \mathbf{X})^{-1} \mathbf{X}^T \mathbf{y}$ is the minimizer of the unregularized empirical risk. For this purpose, using the following steps:

Exercise 1 (i)

(i) Derive that

$$\mathop{\mathrm{arg\,min}}_{oldsymbol{ heta}} \mathcal{R}_{\mathsf{emp}}(oldsymbol{ heta}) = \mathop{\mathrm{arg\,min}}_{oldsymbol{ heta}} \sum_{i=1}^{p} -\hat{ heta}_{i} heta_{i} + rac{ heta_{i}^{2}}{2} + \lambda oldsymbol{I}_{| heta_{i}|
eq 0}.$$

Note that $heta_i$ is from the minimizer $\hat{m{ heta}}$ of the unregularized empirical risk :

$$\hat{\boldsymbol{\theta}} = (\mathbf{X}^T \mathbf{X})^{-1} \mathbf{X}^T \mathbf{y} = \mathbf{X}^T \mathbf{y}$$

because we assume that $\mathbf{X}^T\mathbf{X} = \mathbf{I}$.

$$\mathop{\mathsf{arg\,min}}_{m{ heta}} \mathcal{R}_{\mathsf{reg}}(m{ heta}) = \mathop{\mathsf{arg\,min}}_{m{ heta}} rac{1}{2} ||\mathbf{X}m{ heta} - \mathbf{y}||_2^2 + \lambda \sum_{i=1}^{
ho} \mathbf{I}_{|m{ heta}_i|
eq 0}$$

$$\begin{split} \arg\min_{\boldsymbol{\theta}} \mathcal{R}_{\text{reg}}(\boldsymbol{\theta}) &= \arg\min_{\boldsymbol{\theta}} \frac{1}{2} ||\mathbf{X}\boldsymbol{\theta} - \mathbf{y}||_2^2 + \lambda \sum_{i=1}^p \mathbf{I}_{|\boldsymbol{\theta}_i| \neq 0} \\ &= \arg\min_{\boldsymbol{\theta}} \frac{1}{2} \mathbf{y}^T \mathbf{y} - \mathbf{y}^T \mathbf{X} \boldsymbol{\theta} + \frac{1}{2} \boldsymbol{\theta}^T \mathbf{X}^T \mathbf{X} \boldsymbol{\theta} + \lambda \sum_{i=1}^p \mathbf{I}_{|\boldsymbol{\theta}_i| \neq 0} \end{split}$$

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$$\begin{split} \arg\min_{\boldsymbol{\theta}} \mathcal{R}_{\text{reg}}(\boldsymbol{\theta}) &= \arg\min_{\boldsymbol{\theta}} \frac{1}{2} ||\mathbf{X}\boldsymbol{\theta} - \mathbf{y}||_2^2 + \lambda \sum_{i=1}^p \mathbf{I}_{|\boldsymbol{\theta}_i| \neq 0} \\ &= \arg\min_{\boldsymbol{\theta}} \frac{1}{2} \mathbf{y}^T \mathbf{y} - \mathbf{y}^T \mathbf{X} \boldsymbol{\theta} + \frac{1}{2} \boldsymbol{\theta}^T \mathbf{X}^T \mathbf{X} \boldsymbol{\theta} + \lambda \sum_{i=1}^p \mathbf{I}_{|\boldsymbol{\theta}_i| \neq 0} \\ &= \arg\min_{\boldsymbol{\theta}} - \underbrace{\mathbf{y}^T \mathbf{X}}_{\hat{\boldsymbol{\theta}}^T} \boldsymbol{\theta} + \frac{1}{2} \boldsymbol{\theta}^T \underbrace{\mathbf{X}^T \mathbf{X}}_{=\mathbf{I}} \boldsymbol{\theta} + \lambda \sum_{i=1}^p \mathbf{I}_{|\boldsymbol{\theta}_i| \neq 0} \\ &= \arg\min_{\boldsymbol{\theta}} - \hat{\boldsymbol{\theta}}^T \boldsymbol{\theta} + \frac{1}{2} \boldsymbol{\theta}^T \boldsymbol{\theta} + \lambda \sum_{i=1}^p \mathbf{I}_{|\boldsymbol{\theta}_i| \neq 0} \end{split}$$

$$\begin{split} \arg\min_{\boldsymbol{\theta}} \mathcal{R}_{\text{reg}}(\boldsymbol{\theta}) &= \arg\min_{\boldsymbol{\theta}} \frac{1}{2} ||\mathbf{X}\boldsymbol{\theta} - \mathbf{y}||_2^2 + \lambda \sum_{i=1}^p \mathbf{I}_{|\boldsymbol{\theta}_i| \neq 0} \\ &= \arg\min_{\boldsymbol{\theta}} \frac{1}{2} \mathbf{y}^T \mathbf{y} - \mathbf{y}^T \mathbf{X} \boldsymbol{\theta} + \frac{1}{2} \boldsymbol{\theta}^T \mathbf{X}^T \mathbf{X} \boldsymbol{\theta} + \lambda \sum_{i=1}^p \mathbf{I}_{|\boldsymbol{\theta}_i| \neq 0} \\ &= \arg\min_{\boldsymbol{\theta}} - \underbrace{\mathbf{y}^T \mathbf{X}}_{\hat{\boldsymbol{\theta}}^T} \boldsymbol{\theta} + \frac{1}{2} \boldsymbol{\theta}^T \underbrace{\mathbf{X}^T \mathbf{X}}_{=1} \boldsymbol{\theta} + \lambda \sum_{i=1}^p \mathbf{I}_{|\boldsymbol{\theta}_i| \neq 0} \\ &= \arg\min_{\boldsymbol{\theta}} - \hat{\boldsymbol{\theta}}^T \boldsymbol{\theta} + \frac{1}{2} \boldsymbol{\theta}^T \boldsymbol{\theta} + \lambda \sum_{i=1}^p \mathbf{I}_{|\boldsymbol{\theta}_i| \neq 0} \\ &= \arg\min_{\boldsymbol{\theta}} \sum_{i=1}^p - \hat{\boldsymbol{\theta}}_i \boldsymbol{\theta}_i + \frac{\boldsymbol{\theta}_i^2}{2} + \lambda \sum_{i=1}^p \mathbf{I}_{|\boldsymbol{\theta}_i| \neq 0} \end{split}$$

Exercise 1 (ii)

(ii) Note that the minimization problem on the right-hand side of (i) can be written as $\sum_{i=1}^{p} g_i(\theta)$, where

$$g_i(heta) = -\hat{ heta}_i heta + rac{ heta^2}{2} + \lambda I_{| heta|
eq 0}.$$

What is the advantage of this representation if we seek to find the θ with entries $\theta_1, \dots, \theta_p$ minimizing $\mathcal{R}_{reg}(\theta)$?

Advantage: we can minimize each g_i separately to obtain the optimal entries $\theta_1, \ldots, \theta_p$.

Exercise 1 (iii)

(iii) Consider the first case that $|\hat{\theta}_i| > \sqrt{2\lambda}$ and infer that for the minimizer θ_i^* of g_i it must hold that $\theta_i^* = \hat{\theta}_i$.

Hint: Show that $g_i(\theta_i) < 0 = g_i(0)$ and argue that the minimizer must have the same sign as $\hat{\theta}_i$. (**Personally I find this hint is not so useful.**)

In other words, if $|\hat{\theta}_i|$ is larger than the threshold, $\sqrt{2\lambda}$, then the optimal θ_i^* is the consistent between the regularized and un-regularized empirical risk.

We start with computing the arg min $g_i(\theta_i) = \frac{\theta_i^2}{2} - \hat{\theta}_i \theta_i + \lambda I_{|\theta_i| \neq 0}$.

- ightharpoonup Case 1: $\theta_i = 0$. Then, $g_i(\theta_i) = 0$
- ightharpoonup Case 2: $\theta_i \neq 0$. Then

$$g_i(\theta_i) = \frac{\theta_i^2}{2} - \hat{\theta}_i \theta_i + \lambda,$$

which is a quadratic function, and its minimizer is

$$heta_i^* = \mathsf{arg}\,\mathsf{min}_{ heta_i}\, g_i(heta_i) = \hat{ heta}_i$$

and the minimal value of g_i in this case is

$$g_i(\theta_i^*) = -\frac{\hat{\theta}_i^2}{2} + \lambda$$

Which optimal g_i is smaller? Case 1 or Case 2? It depends on λ . Note that we are given with $|\hat{\theta}_i| > \sqrt{2\lambda}$. So $-\frac{\hat{\theta}_i^2}{2} + \lambda < -\frac{2\lambda}{2} + \lambda = 0$. So the optimal g_i in Case 2 is smaller. So for the minimizer of g_i it holds that $\theta_i^* = \hat{\theta}_i$.

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So for the minimizer of g_i it holds that $\theta_i^* = \hat{\theta}_i$.

Exercise 1 (iv)

(iv) Derive that $\theta_i^* = \hat{\theta}_i \mathbf{I}_{|\hat{\theta}_i| > \sqrt{2\lambda}}$, by using (iii) (and also still considering the case $|\hat{\theta}_i| > \sqrt{2\lambda}$).

Solution: In the solution of (iii) we have proven that

$$heta_i^* = \hat{ heta}_i \quad ext{if} \quad |\hat{ heta}_i| > \sqrt{2\lambda},$$

The optimal θ_i can be written as

$$heta_i^* = \hat{ heta}_i \cdot 1 = \hat{ heta}_i I_{|\hat{ heta}_i| > \sqrt{2\lambda}}$$

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Exercise 1 (v)

(v) Consider the complementary case of (iii) and (iv), i.e., $|\hat{\theta}_i| \leq \sqrt{2\lambda}$, and infer that for the minimizer θ_i^* of q_i it must hold that $\theta_i^* = 0$.

Hint: What is $g_i(0)$? Consider $\tilde{g}_i(0) = \hat{\theta}_i \theta + \frac{\theta^2}{2} + \lambda$ which is the smooth extensiion of g_i .

What is the relationship between the minimizer of g_i and the minimizer of \tilde{g}_i ?

(We do not need this hint in the solution presented in the subsequent slides)

Similarly, we start with computing arg min $g_i(\theta_i) = \frac{\theta_i^2}{2} - \hat{\theta}_i \theta_i + \lambda I_{|\theta_i| \neq 0}$.

- ightharpoonup Case 1: $\theta_i = 0$. Then, $g_i(\theta_i) = 0$
- ightharpoonup Case 2: $\theta_i \neq 0$. Then

$$g_i(\theta_i) = \frac{\theta_i^2}{2} - \hat{\theta}_i \theta_i + \lambda$$

We have shown that the minimizer in **this case** is $\theta_i^* = \hat{\theta}_i$ and min $g_i(\theta_i^*) = g_i(\hat{\theta}_i) = -\frac{\hat{\theta}_i^2}{2} + \lambda$.

Since we consider the constraint $|\hat{\theta}_i| \leq \sqrt{2\lambda}$. Then in Case 2

$$g_i(heta_i^*) \geq -rac{2\lambda}{2} + \lambda = 0.$$

$$\theta_i^* = 0$$

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$$g_i(heta_i^*) \geq -rac{2\lambda}{2} + \lambda = 0.$$

So the minimal g_i in Case 2 is **not smaller** than the minimal g_i in Case 1. (**Plot** $g_i(\theta_i)$ **vs.** θ_i).

Therefore, combining two cases, for the minimizer θ_i^* of g_i it holds that

$$\theta_i^* = 0.$$

Similarly, we start with computing $g_i(\theta_i) = \frac{\theta_i^2}{2} - \hat{\theta}_i \theta_i + \lambda I_{|\theta_i| \neq 0}$.

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We have shown that the minimizer in **this case** is $\theta_i^* = \hat{\theta}_i$ and $\min g_i(\theta_i^*) = g_i(\hat{\theta}_i) = -\frac{\hat{\theta}_i^2}{2} + \lambda$.

Since we consider the constraint $|\hat{\theta}_i| \leq \sqrt{2\lambda}$. Then in Case 2

$$g_i(heta_i^*) \geq -rac{2\lambda}{2} + \lambda = 0.$$

$$\theta_i^* = 0.$$

Exercise 2: Regularization

Directly show the standard solution.