

# **Discussion of: The Subjective Belief Factor**

## **Cui, Delao, and Myers (2025)**

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## What This Paper Does

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**An SDF for beliefs:** just as the SDF prices many assets with one object  $M$ , the SBF explains many subjective forecasts with one probability distortion  $S$

If subjective expectations are coherent,  $\exists S_{t+1}$  with  $\mathbb{E}_t[S_{t+1}] = 1$  such that  $\mathbb{E}_t^*[X] = \mathbb{E}_t[S_{t+1}X]$

**The general formula.** For *any* target  $Y$ , given anchor variables  $\hat{X}$  with observed survey biases:

$$\hat{\mathbb{E}}_t^*[Y_{t+1}] = \mathbb{E}_t[Y_{t+1}] + \underbrace{\text{Cov}(\varepsilon_Y, \hat{\varepsilon})' \hat{\Sigma}^{-1}}_{\text{objective covariances only}} \cdot \underbrace{(\mathbb{E}_t^*[\hat{X}] - \mathbb{E}_t[\hat{X}])}_{\text{anchor survey biases } b_t}$$

- $\varepsilon_Y = Y_{t+1} - \mathbb{E}_t[Y_{t+1}]$ : target's **objective** forecast error
- $\hat{\varepsilon}$ : anchor objective forecast errors;  $\hat{\Sigma} = \text{Var}(\hat{\varepsilon})$
- $b_t = \mathbb{E}_t^*[\hat{X}] - \mathbb{E}_t[\hat{X}]$ : observed anchor biases from survey data

## Step by Step: One Anchor (TBILL $\rightarrow$ TBOND)

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anchor = TBILL survey, target = TBOND

$$\hat{\varepsilon} = \varepsilon^{\text{TBILL}}, \quad b_{t+h} = b_{t+h}^{\text{TBILL}} = \mathbb{E}_t^*[\text{TBILL}_{t+h}] - \mathbb{E}_t[\text{TBILL}_{t+h}]$$

Substituting:

$$\hat{\mathbb{E}}_t^*[\text{TBOND}_{t+h}] = \mathbb{E}_t[\text{TBOND}_{t+h}] + \frac{\text{Cov}(\varepsilon^{\text{TBOND}}, \varepsilon^{\text{TBILL}})}{\text{Var}(\varepsilon^{\text{TBILL}})} \cdot b_{t+h}^{\text{TBILL}}$$

$\beta \equiv$  OLS coefficient of TBOND forecast errors on TBILL forecast errors

So:  $\hat{\mathbb{E}}_t^*[\text{TBOND}_{t+h}] = \mathbb{E}_t[\text{TBOND}_{t+h}] + \beta \cdot b_{t+h}^{\text{TBILL}}$

$\beta$  answers: when the statistical model is surprised by 1 unit about short rates, how much is it *jointly* surprised about long yields?

## The Projection in the Data

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- In the data:  $\beta \approx 0.5$  — half of a short-rate belief distortion transmits to long yields
- $R^2 = 0.85$  of synthetic beliefs
- Also worked for: RGDP + TBILL  $\rightarrow$  RCONSUM, INDPROD (bias  $R^2 \approx 0.7$ )

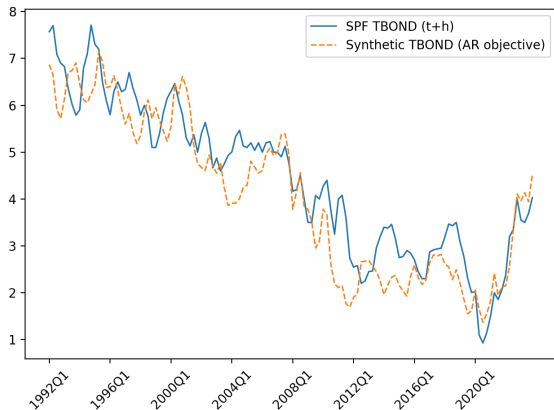
### Why this is powerful:

- One main assumption: coherent distortions
- $\beta$  is pinned by objective shock covariances
- Testable: synthetic forecast should track actual survey when it exists — *and it does*
- Agnostic about belief formation  $\Rightarrow$  portable across applications

Output: synthetic subjective beliefs for variables where no survey exists (earnings growth, equity risk premium, credit spreads)

## Replication: TBILL $\rightarrow$ Synthetic TBOND

**Exercise:** anchor = SPF TBILL, target = TBOND,  $h = 4$  quarters



Benchmark	$\beta$	Level $R^2$	Bias $R^2$
AR( $h$ )	0.479	0.853	0.428
VAR(1)	0.498	0.840	0.508

Bias  $R^2$ : regress  $\mathbb{E}_t^*[Y] - \mathbb{E}_t[Y]$  on  $\hat{\mathbb{E}}_t^*[Y] - \mathbb{E}_t[Y]$

## Replication: RGDP Growth + TBILL $\rightarrow$ Macro Targets

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**Multivariate rotation:** anchors = RGDP growth + TBILL wedges; targets = CPI, RCONSUM, INDPROD

Target	$\beta_{\text{RGDP}}$	$\beta_{\text{TBILL}}$	Level $R^2$	Bias $R^2$
CPI inflation	0.063	0.267	0.344	0.087
RCONSUM growth	0.929	-1.198	0.780	0.776
INDPROD growth	1.783	1.056	0.721	0.677

- Replicates paper's Table 2: RCONSUM and INDPROD bias  $R^2 \approx 0.7$
- CPI: spanning condition binds — RGDP + TBILL forecast errors do not capture inflation surprise variation

## Why This Paper Is Useful and Ambitious

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### Useful:

- Produces belief wedges for variables with no surveys (earnings growth, equity risk premium, credit spreads)
- Portable across applications (term structure, macro-to-finance, cross-country)
- Agnostic about belief formation — works for DE, sticky info, robust control, learning

### Ambitious:

- A general-purpose methodology for belief measurement

## Why This Paper Is Surprising

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- Surprising result 1: heterogeneous CG coefficients ( $-0.29$  to  $1.26$  across 24 variables) replicated by a *single* two-factor SBF
  - ⇒ Correlation between synthetic and actual CG coefficients:  $0.96$
  - ⇒ Observing heterogeneous behavioral patterns does *not* imply heterogeneous mechanisms
- Surprising result 2: variable-by-variable models rejected (Proposition 5)
  - ⇒ 17 of 22 variables exceed the upper bound on factor structure from independent distortions



## Why This Paper Is Surprising

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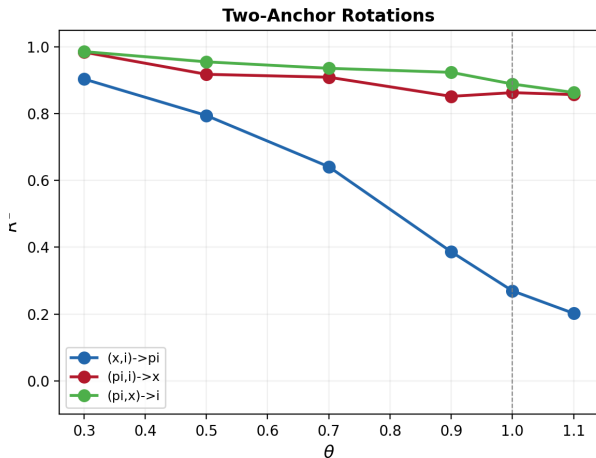
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Main implication: Biases have factor structure too strong to arise from independent distortions ⇒ you need a model with *joint* probability distortions

## Simulation Validation: NK with Diagnostic Expectations

**Exercise:** 3-equation NK with diagnostic expectations (BGS):

$$\mathbb{E}_t^\theta[y'] = \mathbb{E}_t[y'] + \theta(\mathbb{E}_t[y'] - \mathbb{E}_{t-1}[y'])$$

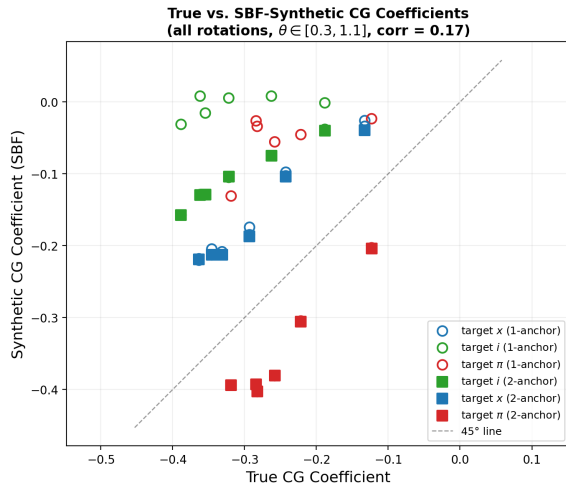


Two-anchor rotations, varying  $\theta$ :

- $(\pi, x) \rightarrow i$  and  $(\pi, i) \rightarrow x$ :  
 $R^2 > 0.85$  across all  $\theta$
- $(x, i) \rightarrow \pi$  **collapses** as  $\theta$  rises  
(0.90  $\rightarrow$  0.20)
- Cost-push shock partially unspanned  
by  $(x, i)$  under DE

## Simulation: How Well Does SBF Recover CG Coefficients?

**Exercise:** simulate NK-DE model across  $\theta \in [0.3, 1.1]$ ; compare true vs. SBF-synthetic CG coefficients (all rotations)



Open = single-anchor, filled = two-anchor; color = target variable

- Most rotations: synthetic CG **attenuated** toward zero — SBF projection strips overreaction signal
- Correlation: 0.17 in 3-variable model vs. paper's 0.96 across 24 variables with wide CG spread

## Suggestion: Perceived Policy Reaction Functions?

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**Motivation:** Bauer–Pflueger–Sunderam estimate perceived Taylor rules from survey expectations and find that perceived coefficients vary over time

**The SBF connection:** the SBF implies *specific differential biases* across Taylor rule inputs

- T-bill expectations underreact strongly ( $CG \approx 0.94$ )
- CPI expectations have a much smaller CG coefficient
- This differential bias across inputs could generate spurious time variation in perceived policy coefficients?

The horse race:

- Estimate perceived Taylor rule from *raw survey* expectations
- Re-estimate from *SBF-synthetic* expectations (which strip out the common distortion)
- How much time variation in perceived coefficients disappears?

## What the Structural Modeler Actually Learns

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### The SBF provides two distinct things:

1. **An empirical fact:** biases are low-rank (approximately rank 2)
  - Model-free — constrains the architecture of any belief specification
  - Models with many belief parameters must produce rank-2 biases
  - Single- $\theta$  models (standard DE): constraint automatically satisfied
2. **A spanning-dependent tool** for extending biases to unobserved variables
  - Reliability depends on whether anchor shocks span the target's shock space
  - Spanning quality is endogenous to  $\theta$  — degrades where distortion is largest
3. **An identification caveat:** data identifies  $\theta_k \rho_k$ , not the components
  - Single  $\theta$  + heterogeneous persistence  $\equiv$  shock-specific  $\theta_k$  + common persistence
  - Prop. 5 rejects variable-by-variable models but cannot separate these

CG validates beliefs;  $R^2$  validates shocks. Use SBF as *diagnostic*, not calibration target

## Summary

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- The SBF is an SDF for beliefs — one probability distortion  $S_{t+1}$  that “prices” many biases through objective shock covariances
- Key empirical finding: low-dimensional  $\hat{S}$  from two anchors explains biases across 24 variables; variable-by-variable models rejected; beliefs explain  $\sim 80\%$  of anomaly returns
- My simulations confirm: SBF works well structurally, but spanning quality is endogenous
- Interpreting the evidence:  $R^2$  validates shock structure; Prop. 5 and CG correlation validate beliefs — these are distinct tests
- Suggestions: horizon dependence, model discrimination via  $\beta_t$ , time-varying coefficients, conditional asset pricing
- For structural modelers: the rank constraint is portable; CG validates beliefs;  $R^2$  validates shocks; use SBF as diagnostic, not calibration target
- This is a very cool paper!

**Observation:** each SPF panelist  $i$  has their own  $S_{i,t+1}$  with loadings  $\beta_{i,t}$

**Key question:** is cross-forecaster disagreement about the *level* or the *composition* of the distortion?

- **Level disagreement:** everyone distorts the same shocks but by different amounts — single-factor model of disagreement may suffice
- **Composition disagreement:** some forecasters overweight demand shocks, others overweight supply shocks — requires multi-factor model of disagreement
- SPF has individual-level data  $\Rightarrow$  estimate  $\hat{\beta}_{i,t}$  for each panelist
- Study the *distribution* of  $\hat{\beta}_{i,t}$  across forecasters
- Connects to Patton–Timmermann (2010), Bordalo et al. (2020) disagreement literature