

Part 3 Assignment 1

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Part 1. Correlation between GDP and access to personal computers in the year of 2005

Question 1

Introduction

We are analyzing the relationship between GDP per capita and access to the personal computers per 100 measured in 2005 (data # Ref 1). The data set from two data (from GDP_per_capita_2005.xlsx, Personalcomputer_2005.xlsx) were combined in the form of `gdp_pc` (data import). Data for the countries missing either of the data, i.e. GDP per capita or access to personal computer were removed from the dataset for this analysis.

Method

Initially, we see the general summary statistics using `summary(gdp_pc)` as shown in result section. Then we calculate the mean, median and standard deviation (sd) for both the columns in the dataset (i.e. `gdp_pc`) and we report Report the Mean, SD and Median.

Result

The mean GDPperCPITA is 7963.5138784 and mean PC_per_100 is 15.7809032. The median GDPperCPITA and PC_per_100 is 2196.247196 and 5.9 respectively. The standard deviation for GDPperCPITA and PC_per_100 is 1.2328808×10^4 and 22.1197352, respectively.

```
summary(gdp_pc)
```

##	Data	GDPperCAPITA	PC_per_100
##	Albania	: 1 Min. : 128.3	Min. : 0.070
##	Algeria	: 1 1st Qu.: 618.6	1st Qu.: 1.755
##	Angola	: 1 Median : 2196.2	Median : 5.900
##	Antigua and Barbuda:	1 Mean : 7963.5	Mean :15.781
##	Argentina	: 1 3rd Qu.: 9271.1	3rd Qu.:16.480
##	Armenia	: 1 Max. :81828.0	Max. :88.660
##	(Other)	:149	

```
sd(gdp_pc$GDPperCAPITA)
```

```
## [1] 12328.81
```

```
sd(gdp_pc$PC_per_100)
```

```
## [1] 22.11974
```

Histogram

From the histogram below, both the GDPperCAPITA and PC_per_100 does not appear to be normally distributed. Since the data was collected over same countries, and to see the relationship between the continuous variables GDPperCAPITA and PC_per_100 across the data set, we need to perform correlation analysis. To check the relationship between GDPperCAPITA and PC_per_100, scatter plot was made which shows linear correlation between the chosen data. In this case, Pearson product-moment correlation was used for the analysis, as Spearman Rank-Order correlation is for the non-parametric analysis, requiring at least one ordinal variable.

```
with(gdp_pc, Hist(GDPperCAPITA, scale="frequency", breaks="Sturges",  
                  col="darkgray", xlab = "GDP Per Capita", ylab = "Frequency"))
```

```
with(gdp_pc, Hist(PC_per_100, scale="frequency", breaks="Sturges",  
                  col="darkgray", xlab = "PC Per 100", ylab = "Frequency"))
```

Analysis

Pearson product-moment correlation:

- Correlation coefficient: 0.7882171
- df: 153
- p-value: < 2.2e-16
- name of test: Pearson's product-moment correlation

```
with(gdp_pc, cor.test(GDPperCAPITA, PC_per_100, alternative="two.sided",  
                      method="pearson"))
```

```
##  
## Pearson's product-moment correlation  
##  
## data: GDPperCAPITA and PC_per_100  
## t = 15.843, df = 153, p-value < 2.2e-16  
## alternative hypothesis: true correlation is not equal to 0  
## 95 percent confidence interval:  
## 0.7200420 0.8413218  
## sample estimates:  
## cor  
## 0.7882171
```

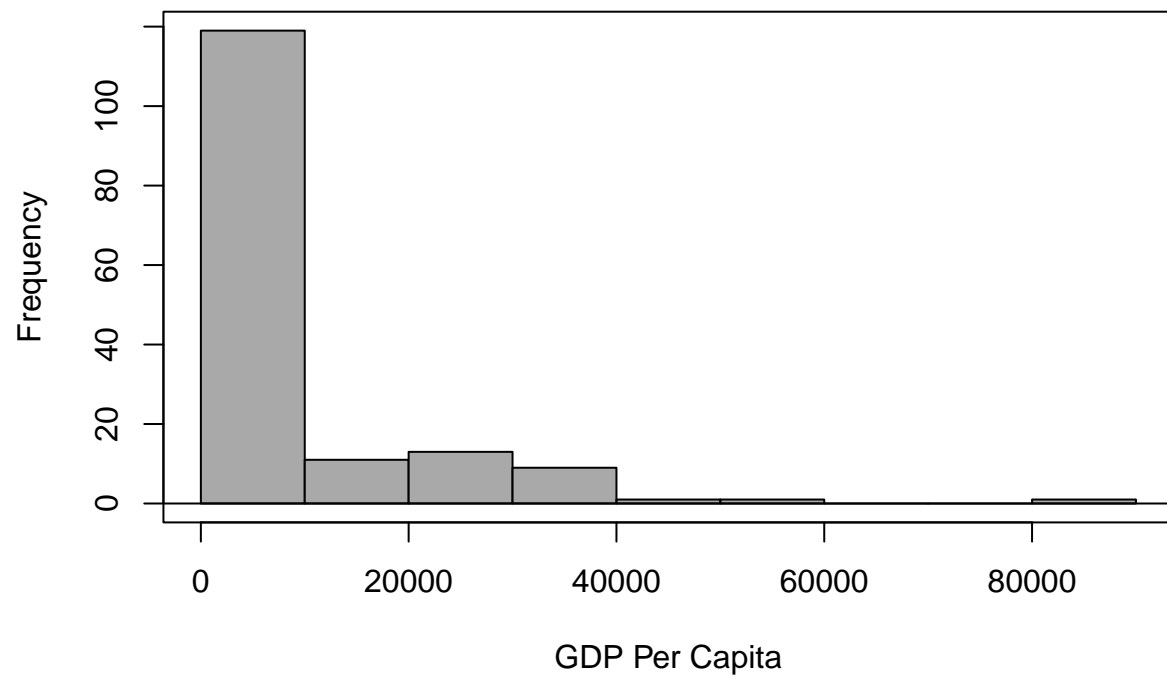


Figure 1: Histogram of GDPperCAPITA

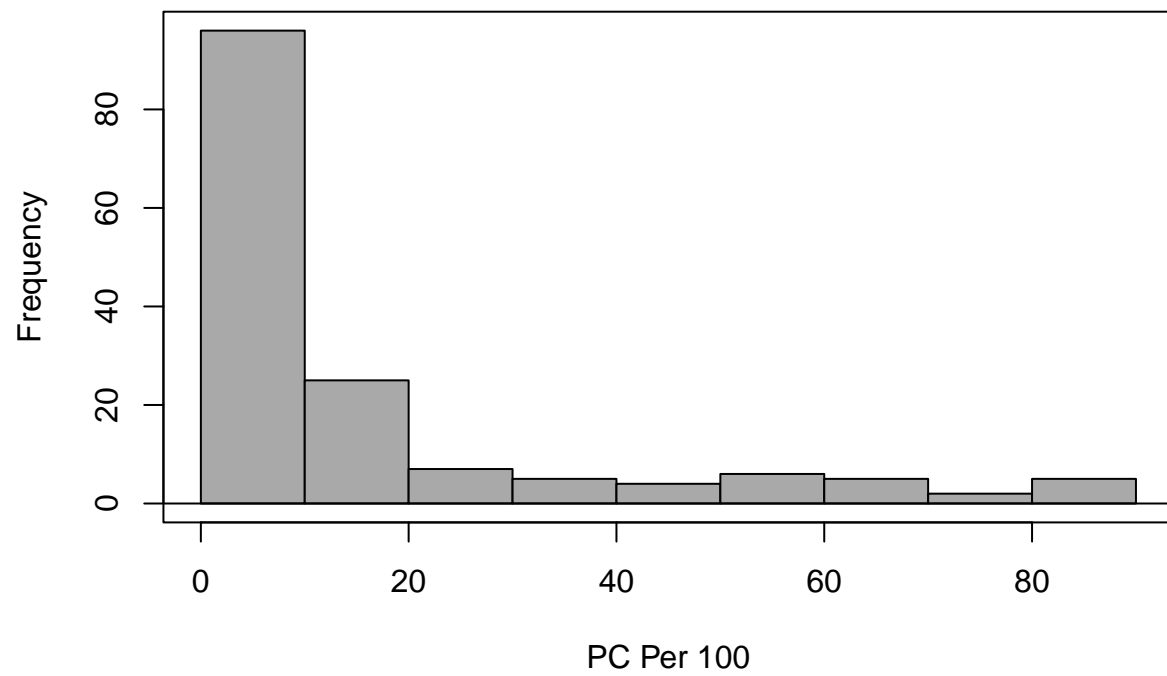


Figure 2: Histogram of PC per 100

Scatterplot

```
scatterplot(PC_per_100~GDPperCAPITA, regLine=FALSE, smooth=FALSE,  
           boxplots=FALSE, data=gdp_pc, xlab = "GDP Per Capita",  
           ylab = "PC Per 1000")
```

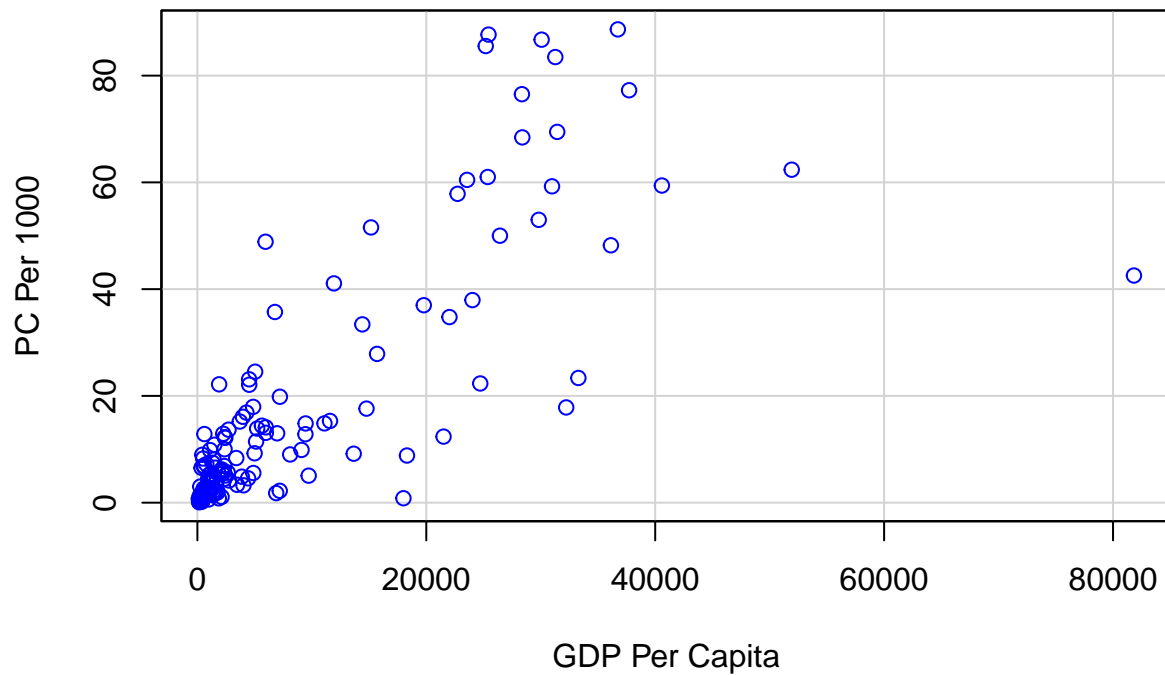


Figure 3: Scatterplot of GDPperCAPITA vs PC_per_100

Discussion

The data shows significant high linear correlation between the GDP Per Capita and PC per 1000. However, from the dataset, the data points are not uniformly distributed but are concentrated near x, y intercept (as apparent from scatterplot). Also shapiro.test, shows $p < 0.005$ (and therefore, hypothesis of normally distributed is rejected) for both GDPperCAPITA and PC_per_100. Therefor, for this data we may also choose Spearman Rank-Order correlation, which also shows similar trend and conclusion. In this case:

- Correlation rho: 0.8547637
- p-value: $< 2.2e-16$
- name of test: Spearman's rank correlation rho

Both the correlation methods give positive correlation between GDPperCAPITA and PC_per_100.

```
shapiro.test(gdp_pc$GDPperCAPITA)
```

```
##  
## Shapiro-Wilk normality test  
##  
## data:  gdp_pc$GDPperCAPITA  
## W = 0.66031, p-value < 2.2e-16
```

```
shapiro.test(gdp_pc$PC_per_100)
```

```
##  
## Shapiro-Wilk normality test  
##  
## data:  gdp_pc$PC_per_100  
## W = 0.69647, p-value < 2.2e-16
```

```
with(gdp_pc, cor.test(GDPperCAPITA, PC_per_100, alternative="two.sided",  
  method="spearman", exact=FALSE))
```

```
##  
## Spearman's rank correlation rho  
##  
## data:  GDPperCAPITA and PC_per_100  
## S = 90137, p-value < 2.2e-16  
## alternative hypothesis: true rho is not equal to 0  
## sample estimates:  
##      rho  
## 0.8547637
```

References

1. GDP_per_capita_2005.xlsx, Personalcomputer_2005.xlsx from Gapminder (<https://www.gapminder.org/data/>).

Part 2. Regression analysis

Question 2 Has the electricity generation per capita in China increased from 1990 to 2005?

Introduction

Here we analyse a linear regression model to study the electricity generation per capita from 1990 to 2005 (data # Ref 1).

Method

Here we make three different regression models i.e. normal data, log10 transformed and square root transformed data to make these models.

Result

Three different regression models were made as follows.

```
Reg.Model.1 <- lm(China~Year, data=chinaElectricity)
summary(Reg.Model.1)

##
## Call:
## lm(formula = China ~ Year, data = chinaElectricity)
##
## Residuals:
##      Min       1Q   Median       3Q      Max
## -187.50  -98.50   14.09   62.55  300.70
##
## Coefficients:
##              Estimate Std. Error t value Pr(>|t|)
## (Intercept) -1.569e+05  1.446e+04  -10.85 3.36e-08 ***
## Year         7.908e+01  7.239e+00   10.92 3.09e-08 ***
## ---
## Signif. codes:  0 '***' 0.001 '**' 0.01 '*' 0.05 '.' 0.1 ' ' 1
##
## Residual standard error: 133.5 on 14 degrees of freedom
## Multiple R-squared:  0.895, Adjusted R-squared:  0.8875
## F-statistic: 119.3 on 1 and 14 DF,  p-value: 3.092e-08
```

```
#plot(Reg.Model.1)
```

- r-squared: 0.887505
- b (regression coefficient): -1.5694468×10^5
- SEb: 7.2392692
- t: 10.9242484
- df: 2, 14, 2
- p: 3.0922238×10^{-8}

Histogram

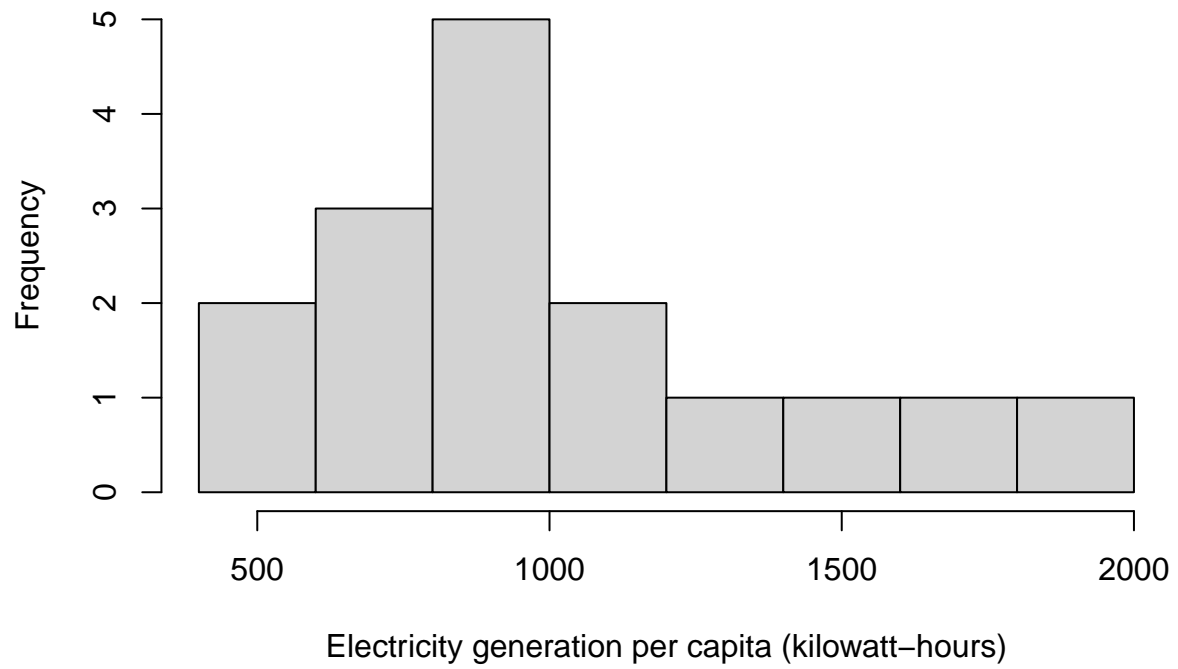
Distribution of electricity production looks like normal distribution, skewed towards left.

```
summary(chinaElectricity$China)

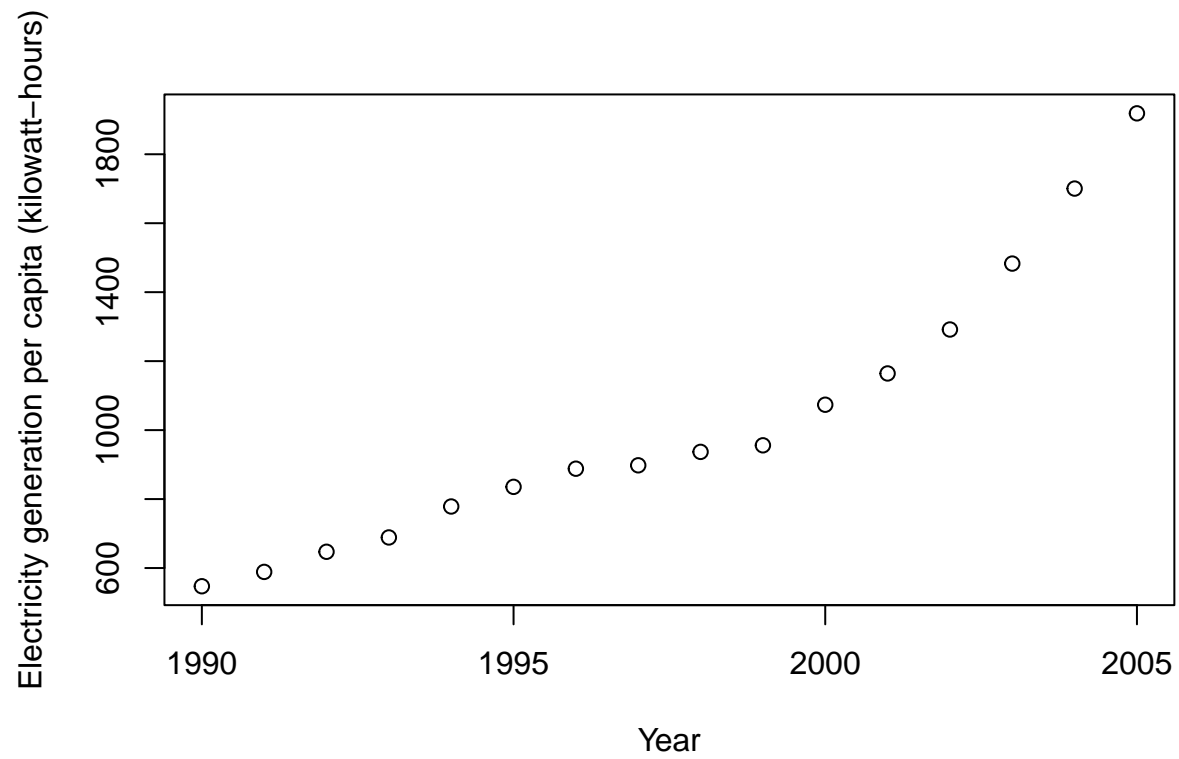
##      Min. 1st Qu.  Median      Mean 3rd Qu.      Max.
##   547.2   756.0   917.4  1024.8  1196.2  1918.6

hist(chinaElectricity$China,
      xlab = "Electricity generation per capita (kilowatt-hours)",
      ylab = "Frequency")
```

Histogram of chinaElectricity\$China

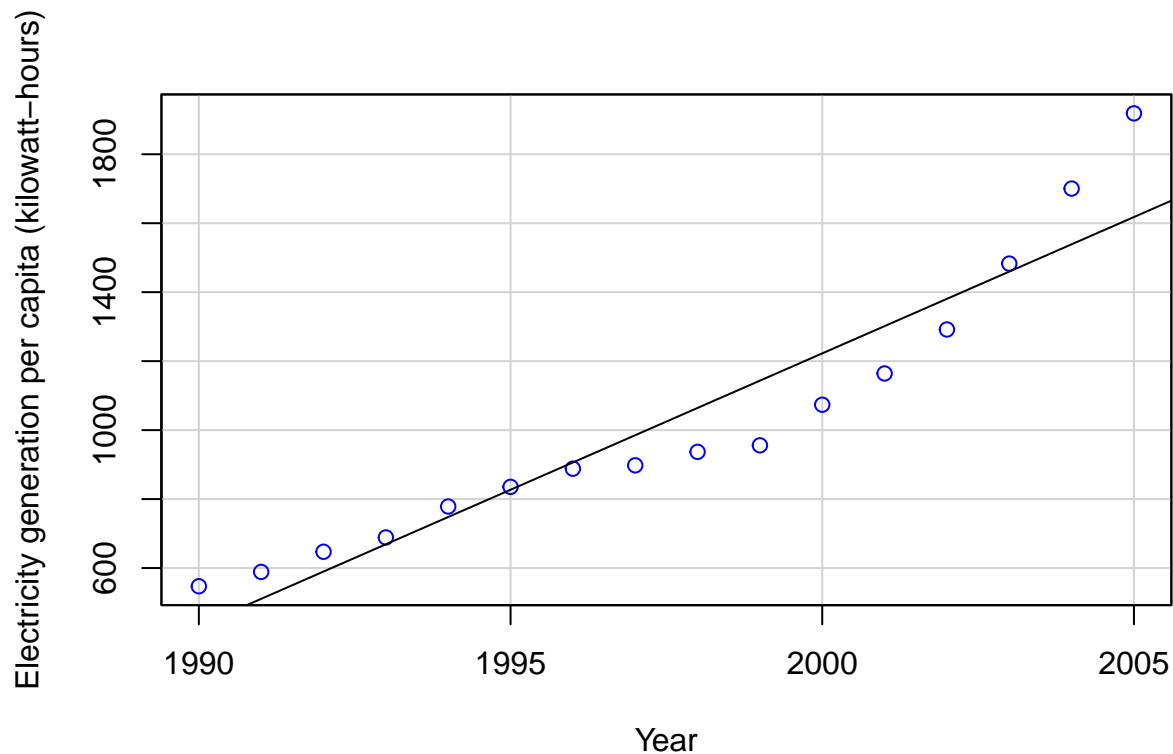


```
plot(chinaElectricity$Year, chinaElectricity$China, xlab = "Year",  
      ylab = "Electricity generation per capita (kilowatt-hours)")
```

Scatterplot

```
scatterplot(chinaElectricity$China~chinaElectricity$Year, regLine=FALSE,
            smooth=FALSE, boxplots=FALSE, xlab = "Year",
            ylab = "Electricity generation per capita (kilowatt-hours)")
abline(Reg.Model.1)
```



Does the relationship look linear? Yes, there is linear increase in the electricity generation per capita from 1990 to 2005 in China.

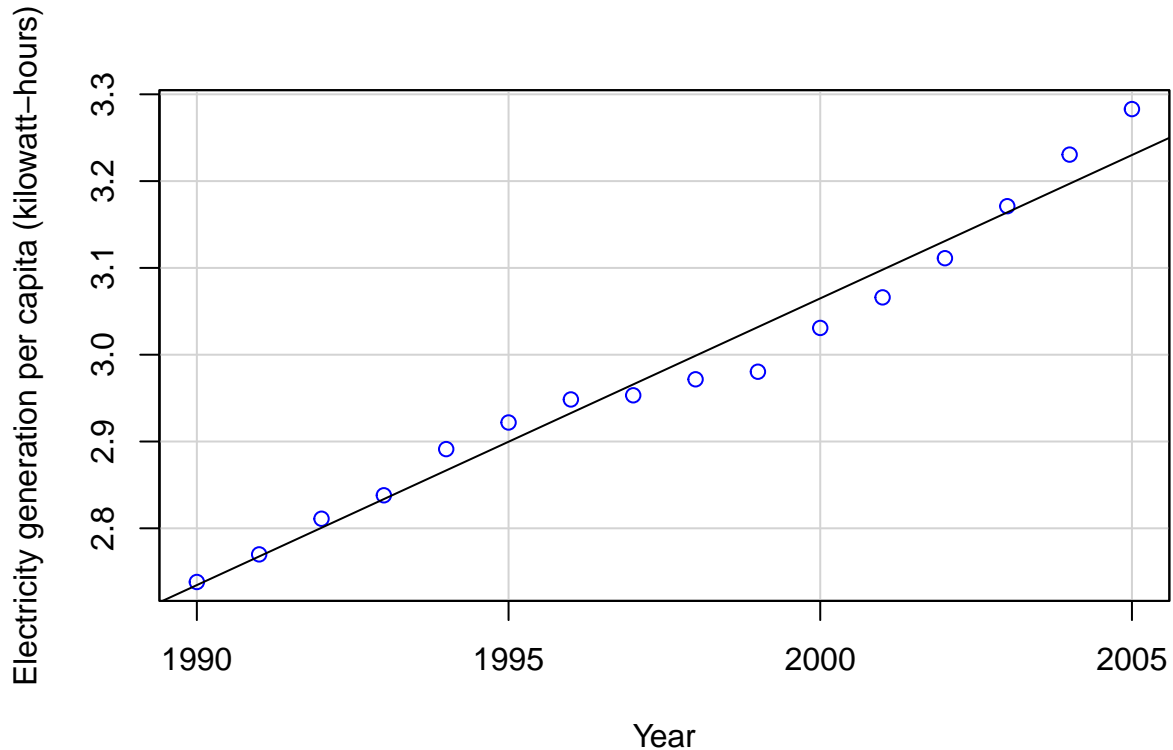
log10 transformation

```
Reg.Model.2 <-lm(log10(China)~Year, data=chinaElectricity)
summary(Reg.Model.2)
```

```
##
## Call:
## lm(formula = log10(China) ~ Year, data = chinaElectricity)
##
## Residuals:
##      Min       1Q   Median       3Q      Max
## -0.051445 -0.021607  0.003993  0.017326  0.052931
##
## Coefficients:
##              Estimate Std. Error t value Pr(>|t|)
## (Intercept) -63.000976   3.114238  -20.23 9.21e-12 ***
## Year          0.033033   0.001559   21.19 4.91e-12 ***
## ---
## Signif. codes:  0 '***' 0.001 '**' 0.01 '*' 0.05 '.' 0.1 ' ' 1
##
## Residual standard error: 0.02875 on 14 degrees of freedom
```

```
## Multiple R-squared:  0.9698, Adjusted R-squared:  0.9676
## F-statistic: 448.9 on 1 and 14 DF,  p-value: 4.913e-12
```

```
scatterplot(chinaElectricity$Year, log10(chinaElectricity$China), regLine=FALSE,
            smooth=FALSE, boxplots=FALSE, xlab = "Year",
            ylab = "Electricity generation per capita (kilowatt-hours)")
abline(Reg.Model.2)
```



```
#plot(Reg.Model.2)
```

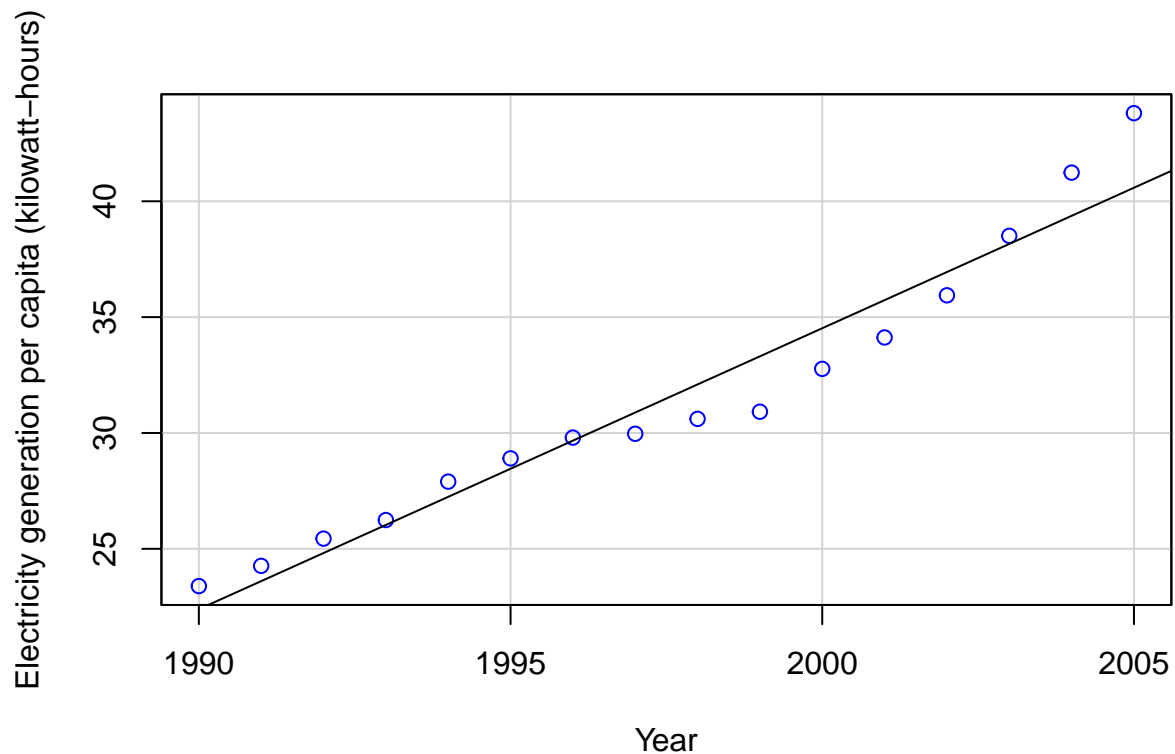
- r-squared: 0.9675968
- b (regression coefficient): -63.0009764
- SEb: 0.0015591
- t: 21.1876694
- df: 2, 14, 2
- p: $4.9127545 \times 10^{-12}$

Square root transformation

```
Reg.Model.3 <-lm(sqrt(China)~Year, data=chinaElectricity)
summary(Reg.Model.3)
```

```
##
## Call:
## lm(formula = sqrt(China) ~ Year, data = chinaElectricity)
##
## Residuals:
##      Min       1Q   Median       3Q      Max
## -2.3899 -1.1247  0.2812  0.6577  3.2179
##
## Coefficients:
##              Estimate Std. Error t value Pr(>|t|)
## (Intercept) -2.391e+03  1.636e+02  -14.62 7.16e-10 ***
## Year          1.213e+00  8.188e-02   14.81 6.02e-10 ***
## ---
## Signif. codes:  0 '***' 0.001 '**' 0.01 '*' 0.05 '.' 0.1 ' ' 1
##
## Residual standard error: 1.51 on 14 degrees of freedom
## Multiple R-squared:  0.94, Adjusted R-squared:  0.9357
## F-statistic: 219.4 on 1 and 14 DF, p-value: 6.02e-10
```

```
scatterplot(chinaElectricity$Year, sqrt(chinaElectricity$China), regLine=FALSE,
            smooth=FALSE, boxplots=FALSE, xlab = "Year",
            ylab = "Electricity generation per capita (kilowatt-hours)")
abline(Reg.Model.3)
```



```
#plot(Reg.Model.3)
```

- r-squared: 0.9357204
- b (regression coefficient): -2390.9385343
- SEb: 0.0818822
- t: 14.8106541
- df: 2, 14, 2
- p: $6.0202529 \times 10^{-10}$

Comparison of different models

- r-squared(linear): 0.887505
- r-squared(log10): 0.9675968
- r-squared (sqrt): 0.9357204

From above value, the r-squared(log10) is the highest (i.e. 0.9675968) and therefore this regression model explains the data optimally. Also see below (mtable):

```
mtable(Reg.Model.1, Reg.Model.2, Reg.Model.3) # Ref 2
```

```
##
## Calls:
## Reg.Model.1: lm(formula = China ~ Year, data = chinaElectricity)
## Reg.Model.2: lm(formula = log10(China) ~ Year, data = chinaElectricity)
## Reg.Model.3: lm(formula = sqrt(China) ~ Year, data = chinaElectricity)
##
## =====
##              Reg.Model.1   Reg.Model.2   Reg.Model.3
##              -----
##              China      log10(China)  sqrt(China)
## -----
## (Intercept) -156944.684*** -63.001*** -2390.939***
##              (14460.479)   (3.114)    (163.560)
## Year         79.084***     0.033***   1.213***
##              (7.239)      (0.002)    (0.082)
## -----
## R-squared    0.895         0.970         0.940
## N            16           16           16
## =====
## Significance: *** = p < 0.001; ** = p < 0.01;
##              * = p < 0.05
```

References

1. Electricity Generation per capita.xlsx from Gapminder (<https://www.gapminder.org/data/>).
2. mtable from <https://bookdown.org/josiesmith/labbook/bivariate-linear-regression.html>

Part 3. Testing differences between groups

Question 3 Is there a difference in income between the New York districts, Manhattan and Brooklyn?

Introduction

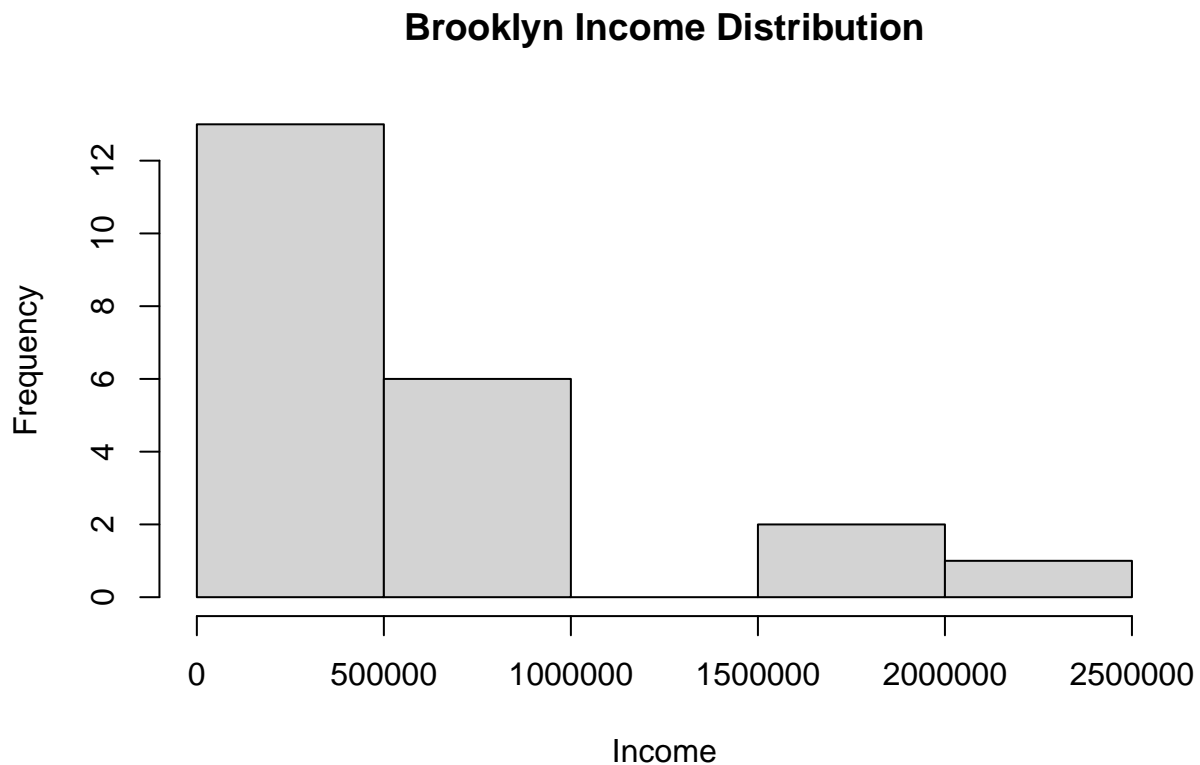
Here we analyse the data from Lander (2019) (Ref 1) to study the difference in the income between two districts in New York i.e. Manhattan and Brooklyn.

Histogram

Histogram by cities

Brooklyn

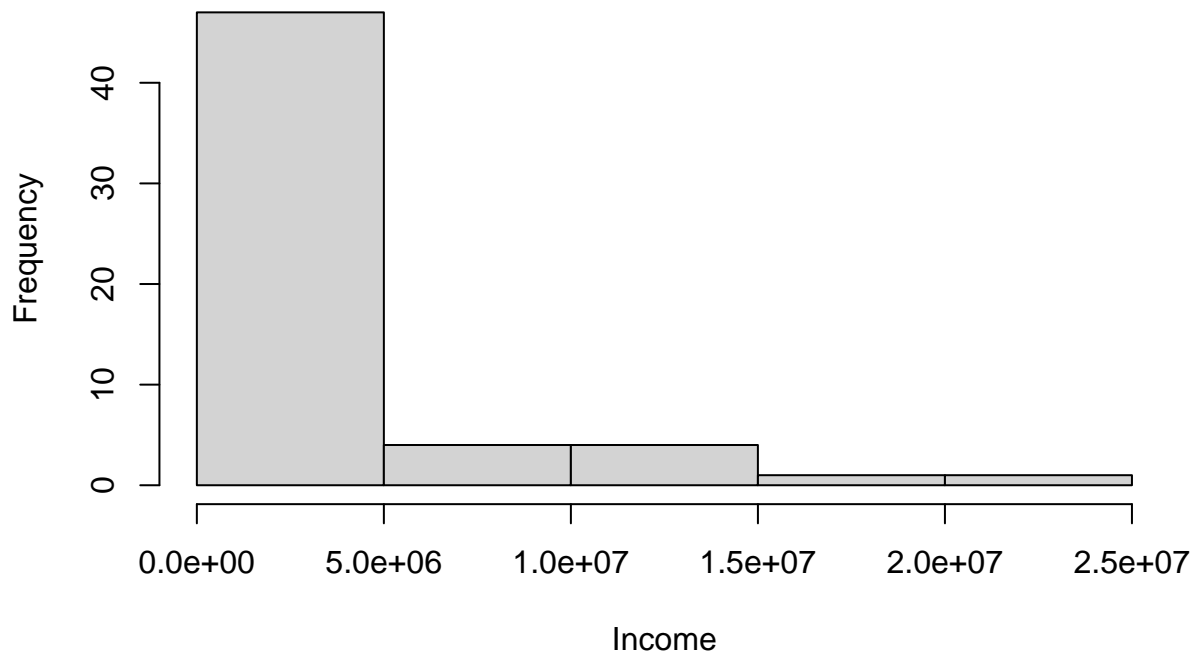
```
hist(subset(LanderHousingNew, Boro == "Brooklyn")$Income,  
     main = "Brooklyn Income Distribution",  
     xlab = "Income")
```



Manhattan

```
hist(subset(LanderHousingNew, Boro == "Manhattan")$Income,  
     main = "Manhattan Income Distribution",  
     xlab = "Income")
```

Manhattan Income Distribution



Which test is most appropriate to use? The LanderHousingNew Income data is not normally distributed data, it is left skewed and therefore, non-parametric test Wilcoxon's rank-sum test would be used in this case.

This is also supported by shapiro.test.

For Brooklyn

```
shapiro.test(subset(LanderHousingNew, Boro == "Brooklyn")$Income)
```

```
##
##  Shapiro-Wilk normality test
##
## data:  subset(LanderHousingNew, Boro == "Brooklyn")$Income
## W = 0.7567, p-value = 0.0001131
```

For Manhattan

```
shapiro.test(subset(LanderHousingNew, Boro == "Manhattan")$Income)
```

```
##
##  Shapiro-Wilk normality test
##
## data:  subset(LanderHousingNew, Boro == "Manhattan")$Income
## W = 0.67163, p-value = 5.125e-10
```

Summary of the data

```
summary(LanderHousingNew)
```

```
##           Neighborhood           Class           Units
## LOWER EAST SIDE           : 6   R2-CONDOMINIUM: 9   Min.    : 5.00
## WILLIAMSBURG-CENTRAL      : 6   R4-CONDOMINIUM:57   1st Qu.: 13.00
## MIDTOWN EAST              : 5   R9-CONDOMINIUM:11   Median : 28.00
## HARLEM-CENTRAL            : 4   RR-CONDOMINIUM: 2   Mean   : 62.61
## TRIBECA                   : 4                               3rd Qu.: 81.00
## UPPER EAST SIDE (79-96): 4                               Max.    :372.00
## (Other)                   :50
##   YearBuilt           SqFt           Income           IncomePerSqFt
## Min.    :1900   Min.    : 5700   Min.    : 147206   Min.    :12.51
## 1st Qu.:1917   1st Qu.: 21587   1st Qu.: 445324   1st Qu.:21.84
## Median :1986   Median : 43065   Median : 1393233   Median :32.86
## Mean    :1965   Mean    : 73678   Mean    : 2591555   Mean    :30.76
## 3rd Qu.:2004   3rd Qu.: 93001   3rd Qu.: 2496518   3rd Qu.:38.30
## Max.    :2009   Max.    :512280   Max.    :22673513   Max.    :61.11
## NA's      :1
##   Expense           ExpensePerSqFt           NetIncome           Value
## Min.    : 51987   Min.    : 6.03   Min.    : 76001   Min.    : 536802
## 1st Qu.: 167086   1st Qu.: 8.16   1st Qu.: 274948   1st Qu.: 2002000
## Median : 443851   Median :10.45   Median : 907440   Median : 6328000
## Mean    : 801372   Mean    :10.20   Mean    : 1790183   Mean    : 13122118
## 3rd Qu.:1023117   3rd Qu.:11.39   3rd Qu.: 1652922   3rd Qu.: 12163499
## Max.    :5609466   Max.    :17.45   Max.    :17064047   Max.    :124320032
##
##   ValuePerSqFt           Boro
## Min.    : 37.40   Brooklyn :22
## 1st Qu.: 92.01   Manhattan:57
## Median :160.47
## Mean    :151.13
## 3rd Qu.:207.76
## Max.    :366.64
##
```

Wilcoxon rank-sum test

```
wilcox.test(LanderHousingNew$Income~LanderHousingNew$Boro)
```

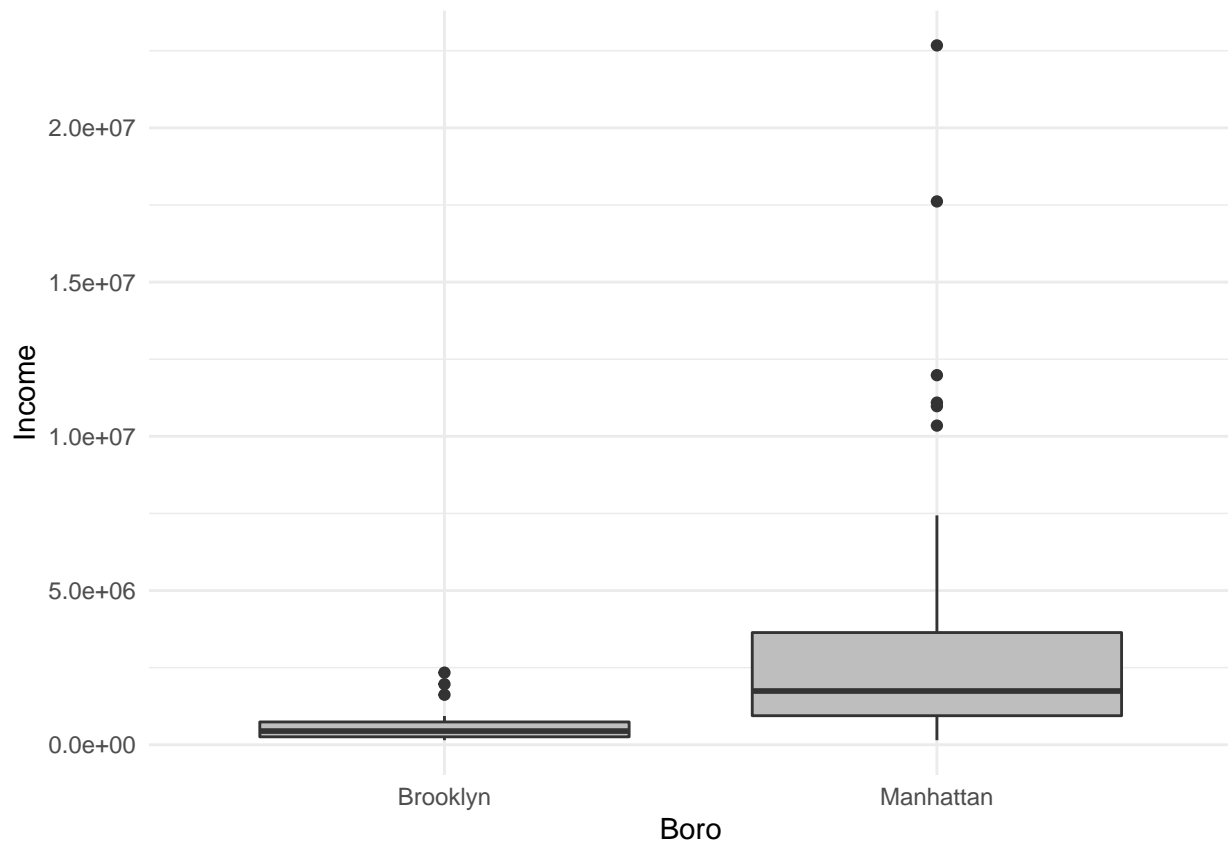
```
##
## Wilcoxon rank sum test with continuity correction
##
## data:  LanderHousingNew$Income by LanderHousingNew$Boro
## W = 223, p-value = 1.019e-05
## alternative hypothesis: true location shift is not equal to 0
```

Since, the p-value 1.0191702×10^{-5} is less than the significance level 0.05, it is therefore concluded that the difference in the Income between the two group i.e. Brooklyn and Manhattan is significant. $W = 223$.

Distribution by cities

Graph representing differences among the districts.

```
ggplot(LanderHousingNew) +  
  aes(x = Boro, y = Income) +  
  geom_boxplot(fill = "grey") +  
  theme_minimal()
```



Above plot shows that income at Manhattan is higher than the income at Brooklyn.

What are the measured central tendencies for income in the two districts?

```
LanderHousingNew %>%  
  group_by(Boro) %>%  
  summarise("n" = length(Income), "Mean" = mean(Income), "Median" = median(Income),  
            "Mode" = Mode(Income), "SD" = sd(Income)) %>%  
  kbl(caption = "Measured central tendencies for income in the two districts") %>%  
  kable_minimal()
```

References

1. Lander (2019) (<https://www.jaredlander.com/datasets/>)

Table 1: Measured central tendencies for income in the two districts

Boro	n	Mean	Median	Mode	SD
Brooklyn	22	639548.1	443850.5	147206	598728.4
Manhattan	57	3344961.0	1742515.0	147229	4345456.9

Question 4 Are there differences in house pricing (SEK/m2) in Sweden between 2016 and 2017?

Introduction

Here we analyse the data from “Svensk mäklarstatistik” (Ref 1) to study the difference in housing pricing in Sweden between 2016 and 2017.

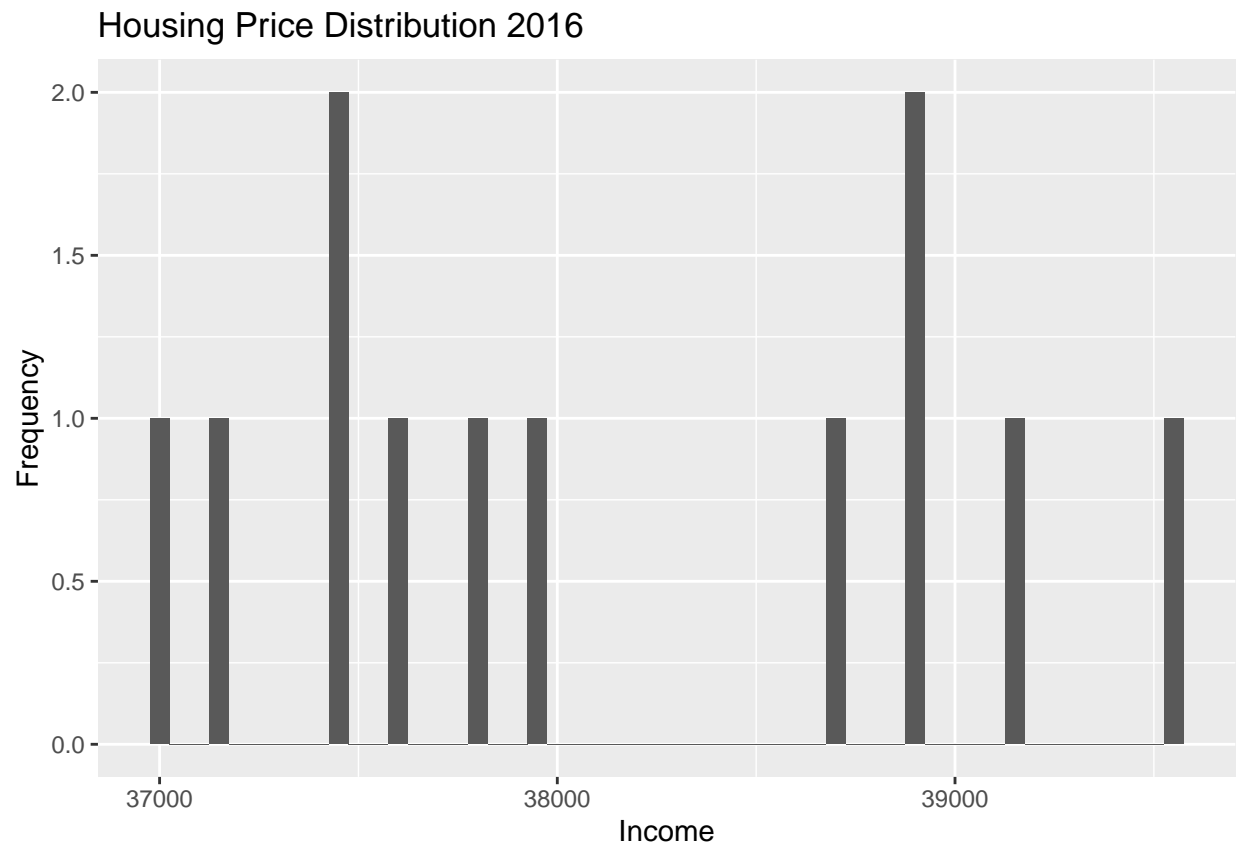
Methods and Results

Histogram

Histogram by year

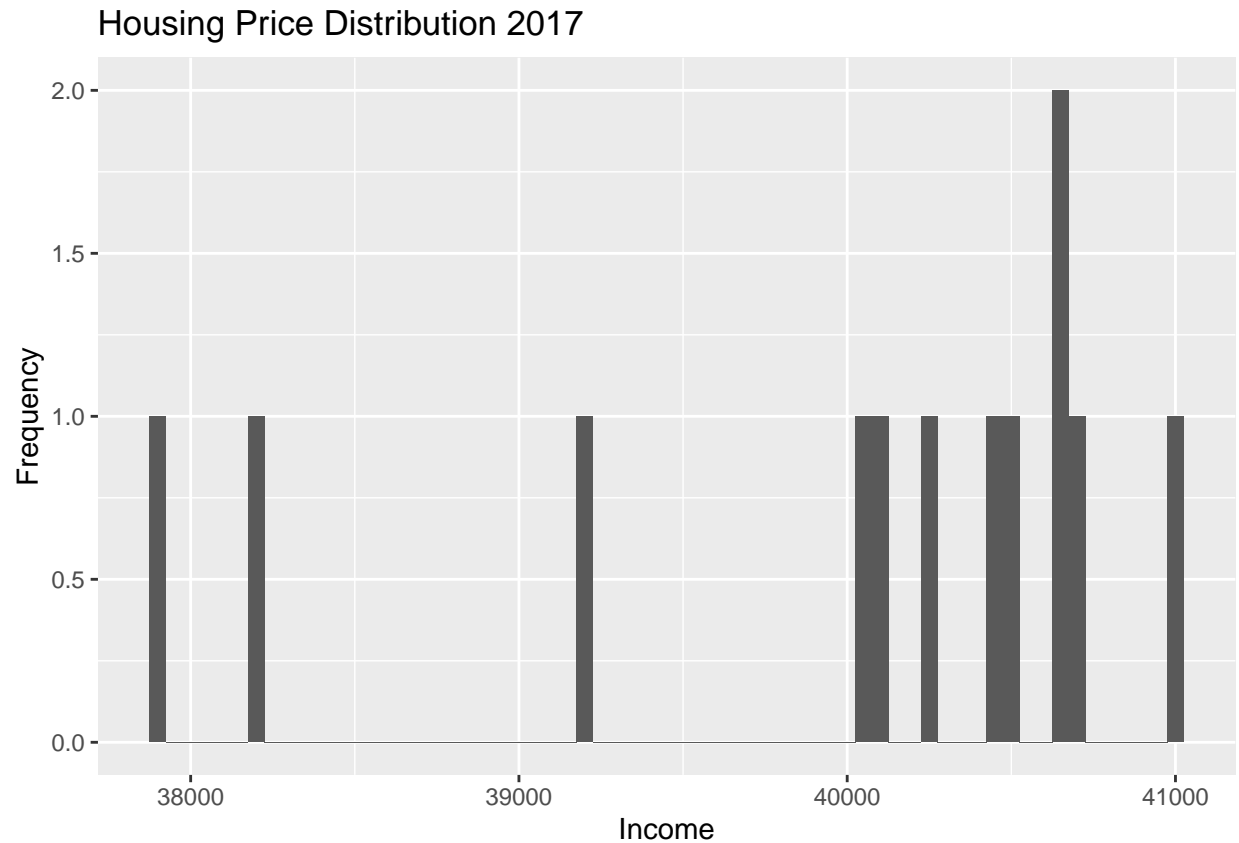
Year 2016: The data is normally distributed as apparent from the histogram below.

```
ggplot(Housepricing_sweden, aes(x=X2016_sek_sqrm)) +
  geom_histogram(binwidth = 50) +
  labs(title="Housing Price Distribution 2016", x="Income", y="Frequency")
```



Year 2017: The data is not-normally distributed (skewed).

```
ggplot(Housepricing_sweden, aes(x=X2017_sek_sqrm)) +  
  geom_histogram(binwidth = 50) +  
  labs(title="Housing Price Distribution 2017", x="Income", y="Frequency")
```



On an average, there is an increase in minimum, mean and maximum house pricing in 2017 as compared to 2016.

```
summary(Housepricing_sweden)
```

```
##      Month  X2016_sek_sqm X2017_sek_sqm
## apr      :1   Min.    :36982   Min.    :37924
## aug      :1   1st Qu.:37458   1st Qu.:39850
## dec      :1   Median :37873   Median :40346
## feb      :1   Mean    :38129   Mean    :39977
## jan      :1   3rd Qu.:38909   3rd Qu.:40643
## july     :1   Max.    :39551   Max.    :41006
## (Other) :6
```

```
sd(Housepricing_sweden$X2016_sek_sqm)
```

```
## [1] 862.8587
```

```
sd(Housepricing_sweden$X2017_sek_sqm)
```

```
## [1] 1001.695
```

```
shapiro.test(Housepricing_sweden$X2016_sek_sqrm)
```

```
##  
## Shapiro-Wilk normality test  
##  
## data: Housepricing_sweden$X2016_sek_sqrm  
## W = 0.9216, p-value = 0.2994
```

```
shapiro.test(Housepricing_sweden$X2017_sek_sqrm)
```

```
##  
## Shapiro-Wilk normality test  
##  
## data: Housepricing_sweden$X2017_sek_sqrm  
## W = 0.81025, p-value = 0.01228
```

The two data set from 2016 and 2017 are not identical with ($V = 6$, p-value 0.006836)

```
wilcox.test(Housepricing_sweden$X2016_sek_sqrm, Housepricing_sweden$X2017_sek_sqrm,  
            paired=TRUE)
```

```
##  
## Wilcoxon signed rank exact test  
##  
## data: Housepricing_sweden$X2016_sek_sqrm and Housepricing_sweden$X2017_sek_sqrm  
## V = 6, p-value = 0.006836  
## alternative hypothesis: true location shift is not equal to 0
```

Paired samples test

```
t.test(Housepricing_sweden$X2017_sek_sqrm, Housepricing_sweden$X2016_sek_sqrm,  
       paired = TRUE, alternative = "two.sided")
```

```
##  
## Paired t-test  
##  
## data: Housepricing_sweden$X2017_sek_sqrm and Housepricing_sweden$X2016_sek_sqrm  
## t = 3.8116, df = 11, p-value = 0.002885  
## alternative hypothesis: true difference in means is not equal to 0  
## 95 percent confidence interval:  
## 780.8864 2915.1136  
## sample estimates:  
## mean of the differences  
## 1848
```

Discussion

The p-value (0.002885) of the test is less than the p-value (0.05). Therefore, we can reject the null hypothesis. There is a significant increase in the housing prices between 2016 and 2017 in Sweden ($t = 3.8116$, $df = 11$, p-value = 0.002885).

References

1. “Svensk mäklarstatistik” <https://www.maklarstatistik.se/>