

# Profiling Female Exclusion from Labour Markets: A Latent Class Approach

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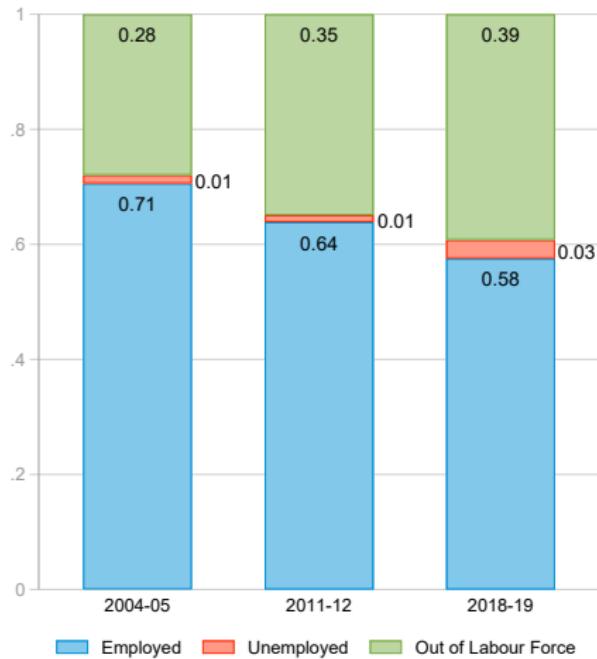
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# Context: Indian Workforce in the 21st Century

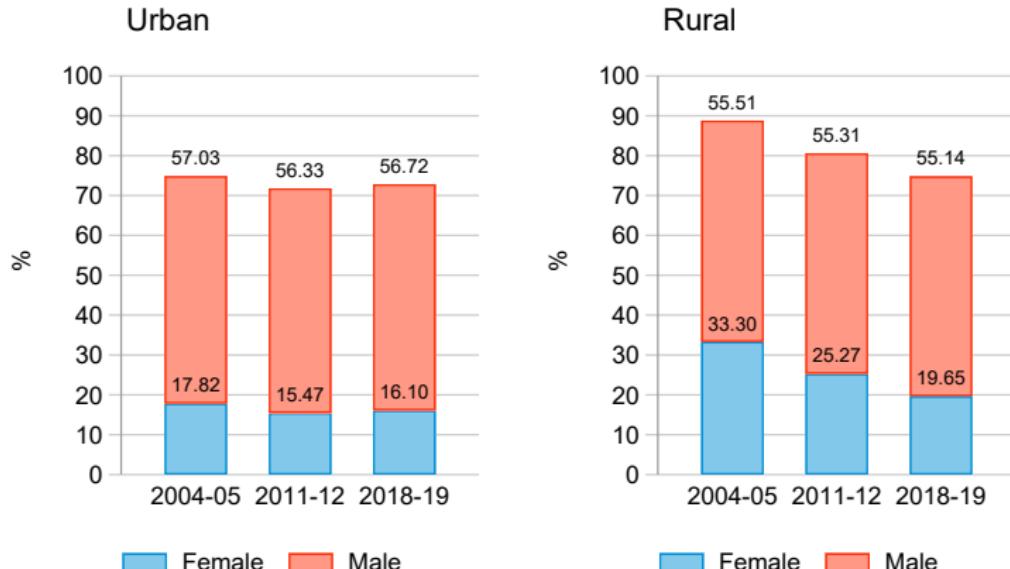
- Consistent decline in labour force participation rates despite having the second highest working age population in the world
  - 71% in 2004-05 → 58% in 2018-19
- Medium term economic prospects → trajectory of post-COVID job recoveries (Basole (2020), Deshpande (2020), Agarwal (2021))
- Critically important to identify vulnerable workers in the labour market

Demographic Composition of Working Age Population (20 - 59 years)



# Declining Female Labor Force Participation

- Consistent decline in female labor force participation rates
- Decline in FLFP more severe in rural areas (Pande and Fletcher (2018))



Source: NSS61, NSS68, PLFS 2018-19

# Motivation

- Vast dissimilarities in the gender composition of those not participating in labor markets (OLF)

Table: Duncan Dissimilarity Index for gender segregation among OLF

Year	2004	2011	2018
D-index	0.434	0.487	0.513

- Females face **excess burden** of being excluded from the labour market
- 57% women in India do not participate in the labour market (PLFS 2018-19)
- Vast heterogeneity in demographic characteristics of females excluded from the labour market, **making inclusionary policy targeting difficult and ineffective**

# Policy Problem

- Labour market policies uninformed by typologies of targeted females → expected outcomes remain unachieved
- Awareness of heterogeneities in OLF females is important for atleast three immediate reasons:
  - Greater ease in policy targeting
  - Tailoring policies and programs for specific typologies of OLF women
  - Focused prioritization of policies that can reintegrate females into the labor market at the lowest public expenditure

# Latent Class Approach

- ① Multidimensional Profiling of Female Exclusion
- ② Latent Class Model
- ③ Data and Covariates
- ④ Optimal Latent Classes

# Multidimensional Profiling of Female Exclusion

- Observed characteristics of females are used to enable identification of latent classes that capture different dimensions of exclusion
- **Person-oriented approach:** Group sub-types of females that exhibit similar patterns of characteristics using a probabilistic framework
- Statistically, we compute the score for each female according to the likelihood of belonging to each of the computed latent classes (Bergman and Magnusson (1997))
  - Subsequently assigned to the latent class to which they have the highest posterior probability of belonging
- Frontier research on labour market segmentation is increasingly using variations of latent class approach (Lukac (2019), Yoon and Chung (2015))

# Setting up the Latent Class Model

- Let  $j = 1, \dots, J$  be the observed indicator, and observed indicator has  $j$  has  $r_j = 1, \dots, R_j$  response categories.
- $L$  is categorical latent variable with  $c = 1, \dots, C$  latent classes.
- $P[L = c] = \gamma_c$
- Item-response probability of indicator  $j$  is referred to as  $\rho_j$

As latent classes are mutually exclusive and exhaustive, each female is a member of one and only one latent class

$$\sum_{c=1}^C \gamma_c = 1 \quad (1)$$

Item-response probability of a given indicator conditional on latent class is unity (since each individual provides a unique response to an indicator)

$$\sum_{r_j=1}^{R_j} \rho_{j, r_j | c} = 1 \quad (2)$$

# Estimation

- Let  $y_j$  represent element  $j$  of a response pattern  $\mathbf{y}$ .
- What is  $P(Y = \mathbf{y}|L = c)$ ?

Let  $I(Y_j = \gamma_j) = 1$  when  $j = \gamma_j$  and 0 otherwise

Probability of observing a particular vector of response is a function of the probabilities of membership in each latent class  $c$  and the probability of observing each response conditional on latent class membership

$$P(Y = \mathbf{y}|L = c) = \prod_{j=1}^J \prod_{\gamma_j=1}^{R_j} \rho_{j,\gamma_j|c}^{I(y_j=\gamma_j)} \quad (3)$$

Parameters  $\gamma$  and  $\rho$  are estimated via **maximum likelihood**

# Posterior Assignment

Using Bayes Theorem, we can calculate the posterior probability of belonging in latent class  $c$  conditional on the observed response pattern  $\mathbf{y}$

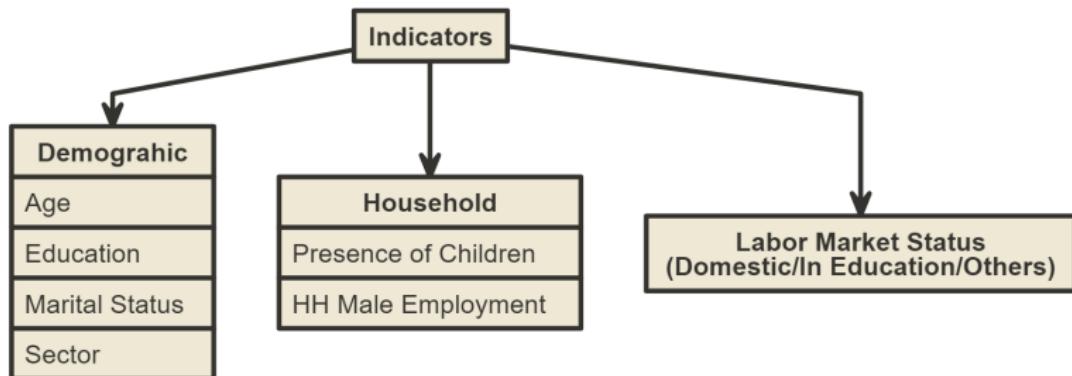
$$P(L = c \mid Y = \mathbf{y}) = \frac{P(Y = \mathbf{y} \mid L = c)P(L = c)}{P(Y = \mathbf{y})} \quad (4)$$

Since we know that  $P(L = c) = \gamma_c$ ,

$$P(L = c \mid Y = \mathbf{y}) = \frac{\left( \prod_{j=1}^J \prod_{\gamma_j=1}^{R_j} \rho_{j, \gamma_j|c}^{I(y_j=\gamma_j)} \right) \gamma_c}{\sum_{c=1}^C \gamma_c \prod_{j=1}^J \prod_{\gamma_j=1}^{R_j} \rho_{j, \gamma_j|c}^{I(y_j=\gamma_j)}} \quad (5)$$

# Data and Key Covariates

- **Data:** NSS Employment Survey 2004-05 and Periodic Labour Force Survey 2018-19 ( $N = 1,023,590$ )
- **Sub-sample:** Females (20-59 years) out of labour force ( $N = 164,214$ )
- **Indicator variables:**



These indicators collectively constitute the response pattern vector  $\mathbf{y}$  for each female in the sample

# Determination of Optimal Latent Classes

- Models with upto 9 latent classes are estimated
- To compare model fit and parsimony → **Akaike information criterion (AIC) and Bayesian Information criterion (BIC)**
- Smaller values of AIC and BIC represent optimal balance of model fit and parsimony
- Using elbow test, **4 classes** are optimal for 2004-05 and **6 classes** for 2018-19

# Classes	2004-05				2018-19			
	AIC	BIC	$\Delta(AIC)$	$\Delta(BIC)$	AIC	BIC	$\Delta(AIC)$	$\Delta(BIC)$
2	$1.31 \times 10^9$	$1.31 \times 10^9$			$2.38 \times 10^9$	$2.38 \times 10^9$		
3	$1.25 \times 10^9$	$1.25 \times 10^9$	-4.1%	-4.1%	$2.27 \times 10^9$	$2.27 \times 10^9$	-4.7%	-4.7%
<b>4</b>	$1.21 \times 10^9$	$1.21 \times 10^9$	-3.2%	-3.2%	$2.17 \times 10^9$	$2.17 \times 10^9$	-4.3%	-4.3%
5	$1.21 \times 10^9$	$1.21 \times 10^9$	-0.6%	-0.6%	$2.14 \times 10^9$	$2.14 \times 10^9$	-1.4%	-1.4%
<b>6</b>	$1.18 \times 10^9$	$1.18 \times 10^9$	-1.7%	-1.7%	$2.11 \times 10^9$	$2.11 \times 10^9$	-1.4%	-1.4%
7	$1.16 \times 10^9$	$1.16 \times 10^9$	-2.2%	-2.2%	$2.1 \times 10^9$	$2.1 \times 10^9$	-0.3%	-0.3%
8	$1.15 \times 10^9$	$1.15 \times 10^9$	-0.5%	-0.5%	$2.09 \times 10^9$	$2.09 \times 10^9$	-0.7%	-0.7%
9	$1.16 \times 10^9$	$1.16 \times 10^9$	0.6%	0.6%	$2.08 \times 10^9$	$2.08 \times 10^9$	-0.4%	-0.4%

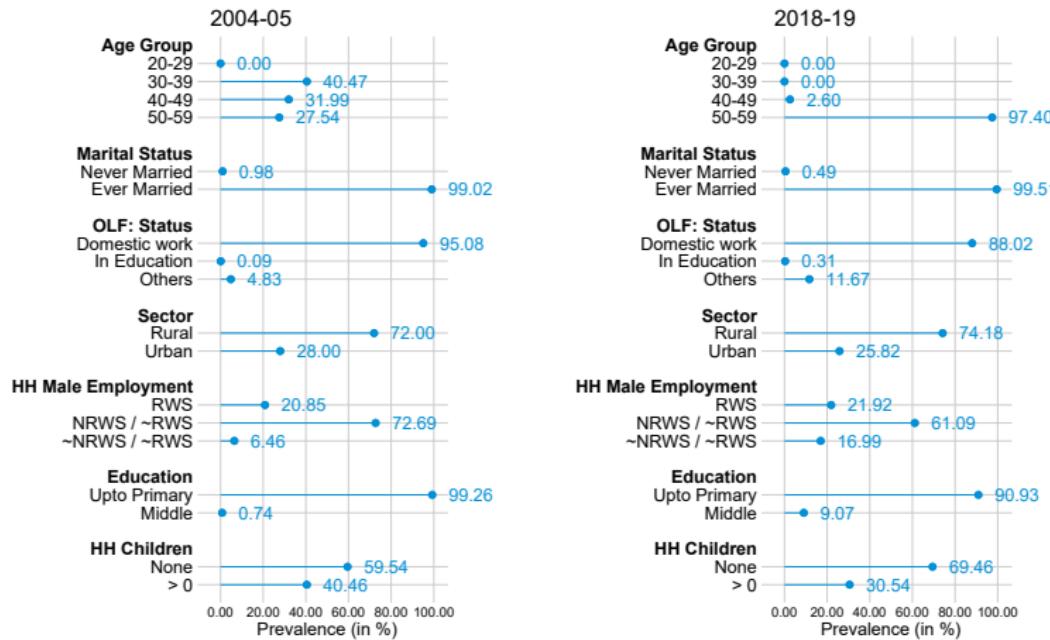
# Latent Classes

2004-05	2018-19
Elder, Married and Unlettered (EMU)	Elder, Married and Unlettered (EMU)
Young, Married and Unlettered (YMU)	Young, Married and Unlettered (YMU)
Middle-aged, married, urban and financially secured (MFS)	Middle-aged, married, urban and financially secured (MFS)
Urban, young and educated (UY)	Middle-aged, Married and Unlettered (MMU) Young and married with HH dependents (YEH) Young women in education (YE)

**Our focus:** Similar latent classes across time (EMU, YMU, MFS)

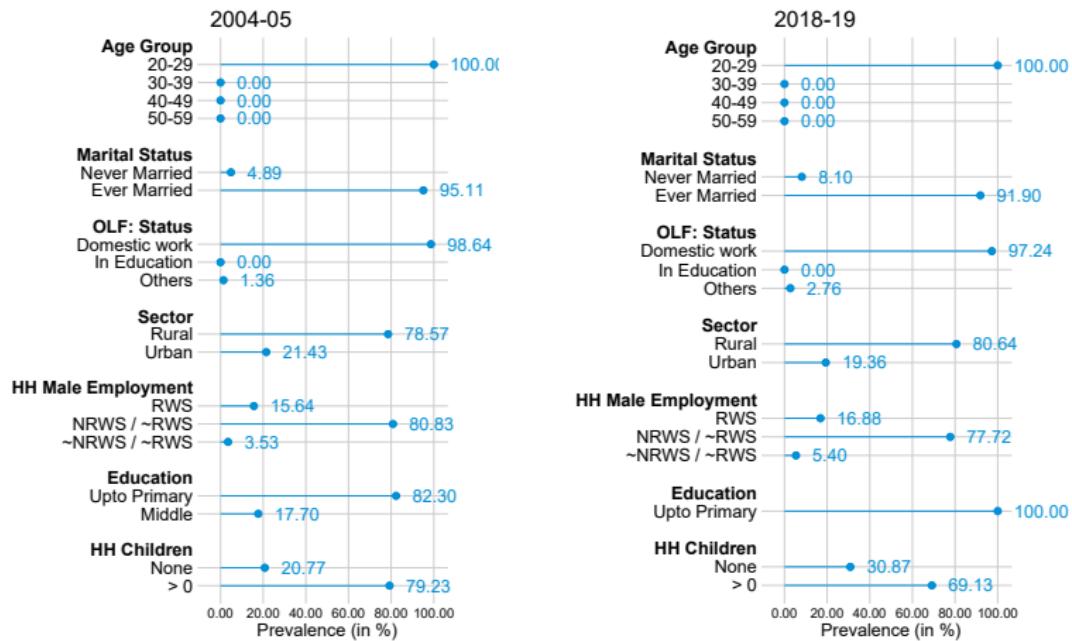
# Comparison: Elderly, Married and Unlettered (EMU)

Prevalence in 2004-05: 43.6%; 2018-19: 13.2%



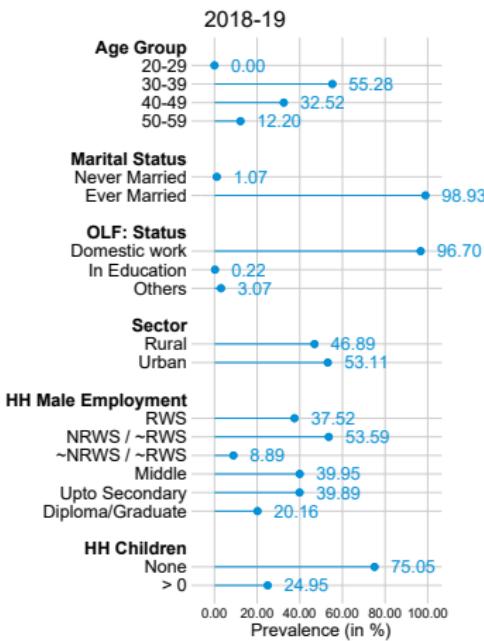
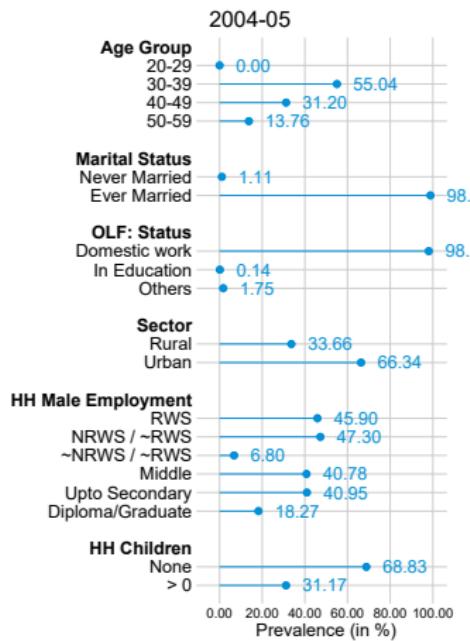
# Comparison: Young, Married and Unlettered (YMU)

Prevalence in 2004-05: 24.8%; 2018-19: 9.8%



# Comparison: (MFS)

Prevalence in 2004-05: 17.9%; 2018-19: 26.1%



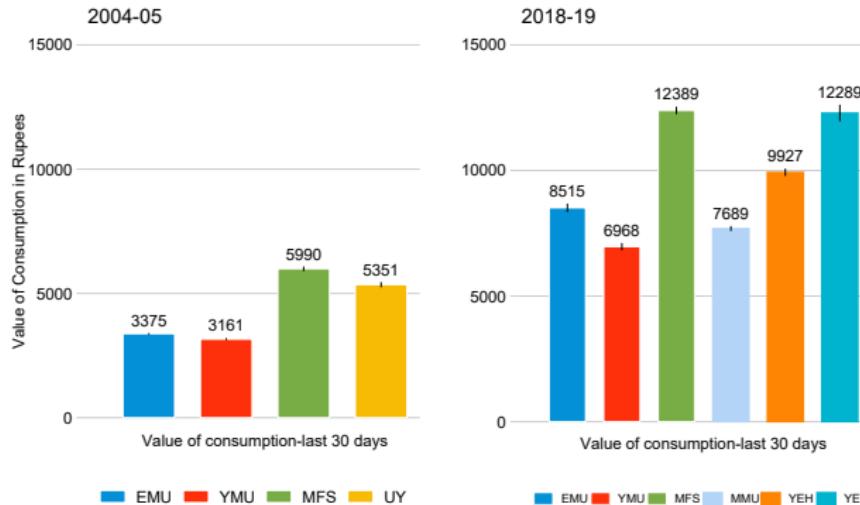
# Quantifying Vulnerability: Axes of Exclusion

- 3 axes of exclusion:
  - Economic deprivation: lack of access to goods
  - Social marginalization
  - Location barrier
- Identifying vulnerability:
  - Compare average HH consumption across latent classes
  - Compare prevalence in social groups (ST/SC/OBC/Others)
  - Between and within-state variation

# Household Consumption

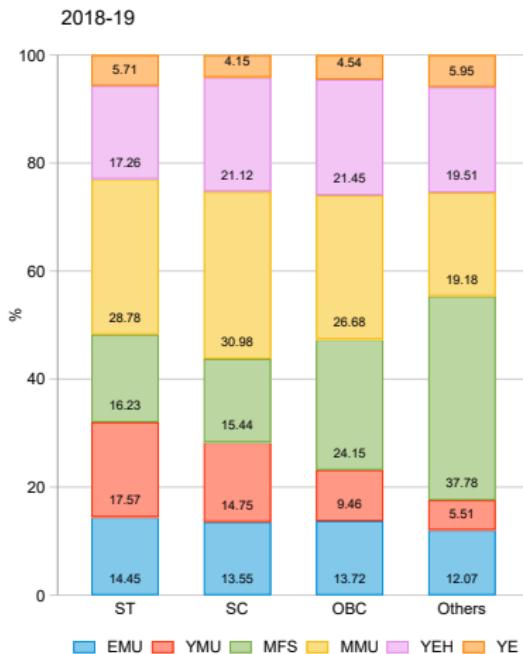
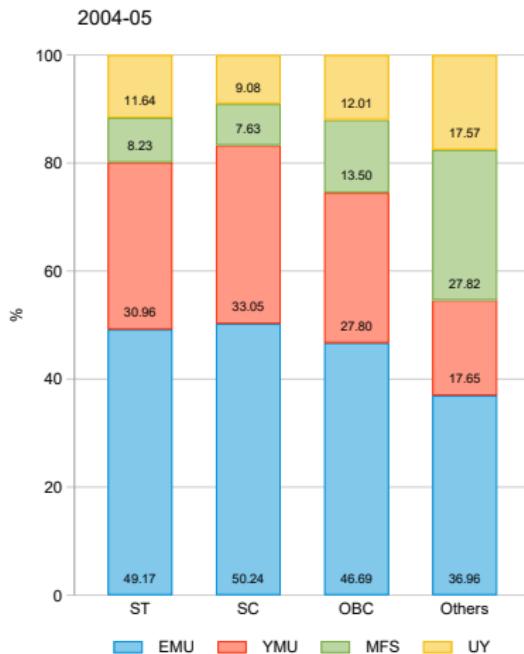
- Young, married and unlettered women (YMU) have lowest average HH consumption
- Average HH consumption increased for EMU but still low in 2018-19

Average HH Consumption by Latent Class



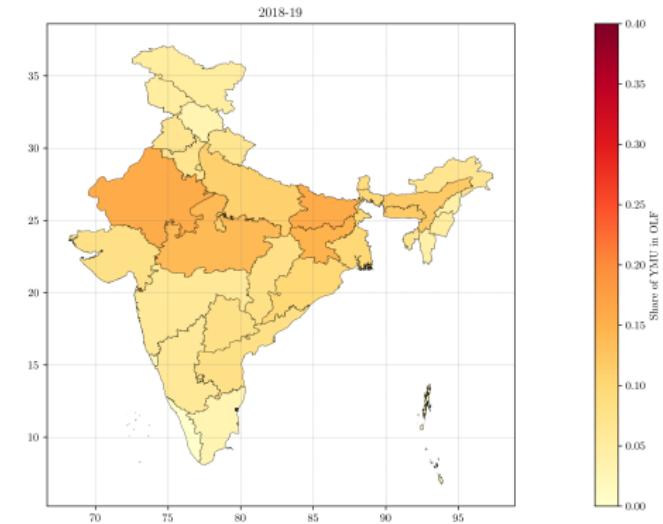
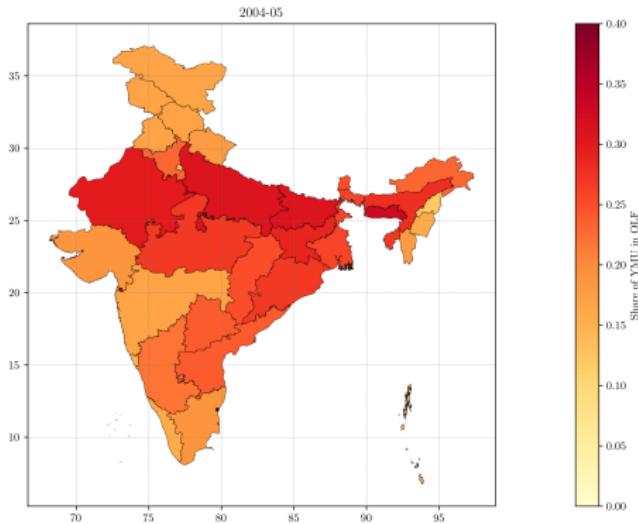
# Social Group

Composition of Social Groups by Latent Classes

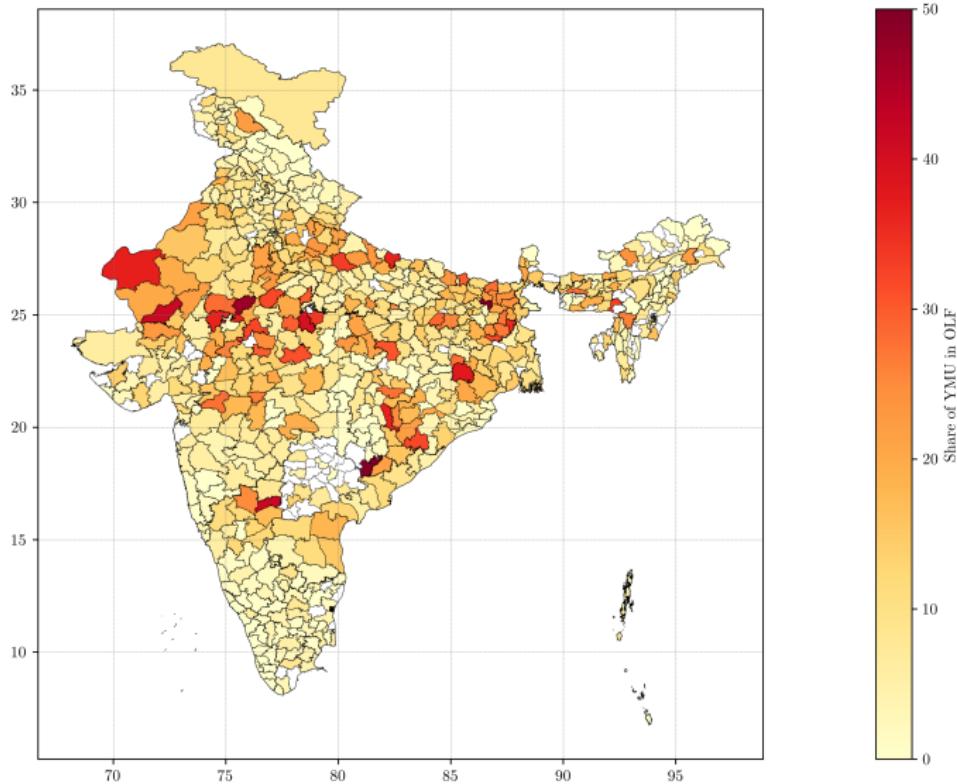


# Spatial Distribution: Share of Vulnerable Latent Classes

- Aggregate decline in share of socio-economically vulnerable latent classes over time

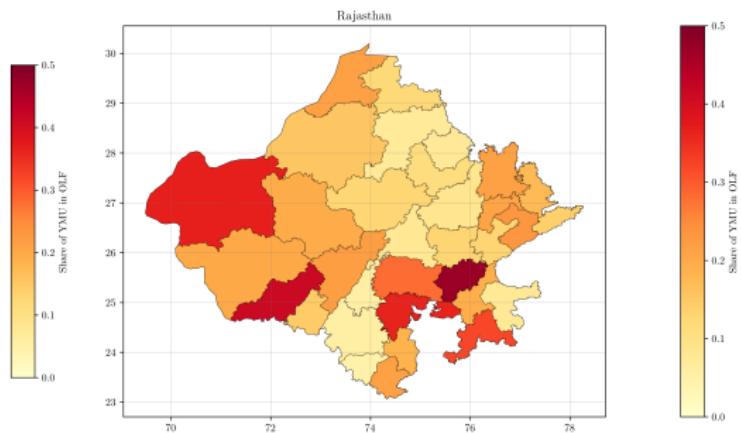
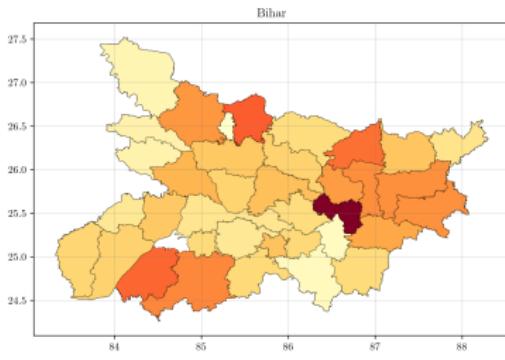


# Spatial Distribution: District-wise Variation



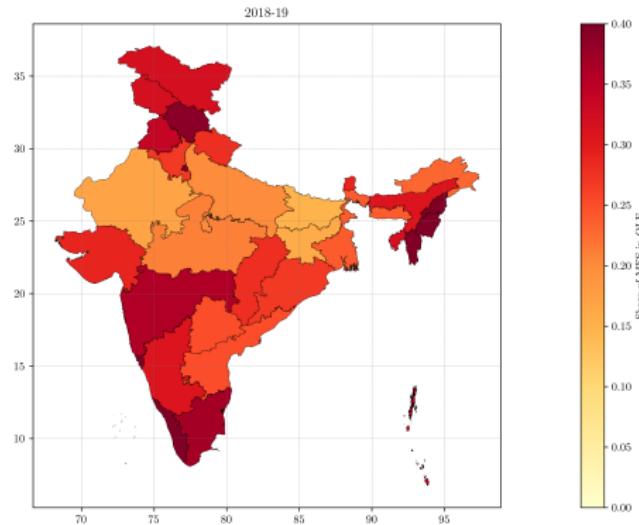
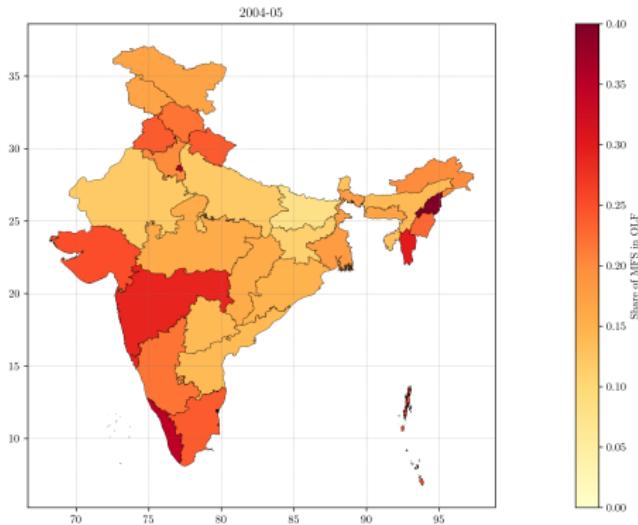
# Spatial Distribution: Within-State Variation

- Large within-state variation in prevalence of vulnerable latent classes



# Spatial Distribution: MFS

- Rise in MFS: more difficult to reintegrate in the labour market due to high substitution effects



# Key Takeaways

- Understanding composition of females not participating in the labor market is important for two reasons:
  - Demographically and socio-economically heterogeneous
  - Prevalence of each sub-type varies between and within state
- Large part of rise in declining FLFP can be attributed to the rise in MFS women
  - High substitution effects within household
  - More difficult to integrate due to lesser incentives to work
- Latent class approach can give us actionable insight to understand policy-relevant compositional differences

# Policy Recommendations

- EMU: Better social safety nets for women in regions with high prevalence of EMU women
- YMU: Increased public expenditure on public works programs in areas with high prevalence of YMU women
- Specific policy attention needed for reintegrating MFS women:
  - Targeted skill development
  - Improved design of market incentives
- Reduce barriers to female mobility

**Thank you!**