

# Cardano.R

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```
library(tseries)

## Registered S3 method overwritten by 'quantmod':
##   method      from
##   as.zoo.data.frame zoo

library(tidyverse)

## -- Attaching packages ----- tidyverse 1.3.2 --
## v ggplot2 3.4.0     v purrr    0.3.5
## v tibble   3.1.8     v dplyr    1.0.10
## v tidyverse 1.2.1    v stringr  1.4.1
## v readr    2.1.3     vforcats 0.5.2
## -- Conflicts ----- tidyverse_conflicts() --
## x dplyr::filter() masks stats::filter()
## x dplyr::lag()   masks stats::lag()

library(ggplot2)
library(stats)
library(forecast)

## Warning: package 'forecast' was built under R version 4.2.3
library(TSA)

## Registered S3 methods overwritten by 'TSA':
##   method      from
##   fitted.Arima forecast
##   plot.Arima  forecast
##
## Attaching package: 'TSA'
##
## The following object is masked from 'package:readr':
## 
##   spec
## 
## The following objects are masked from 'package:stats':
## 
##   acf, arima
## 
## The following object is masked from 'package:utils':
## 
##   tar
```

```

library(urca)

## Warning: package 'urca' was built under R version 4.2.3
library(FinTS)

## Warning: package 'FinTS' was built under R version 4.2.3

## Loading required package: zoo
##
## Attaching package: 'zoo'
##
## The following objects are masked from 'package:base':
##
##     as.Date, as.Date.numeric
##
##
## Attaching package: 'FinTS'
##
## The following object is masked from 'package:forecast':
##
##     Acf

df = read_csv("E:\\Stevens institute\\SecondSem\\MA641 Time Series\\Project\\archive\\coin_Cardano.csv")
df = select(df, Date, Close)

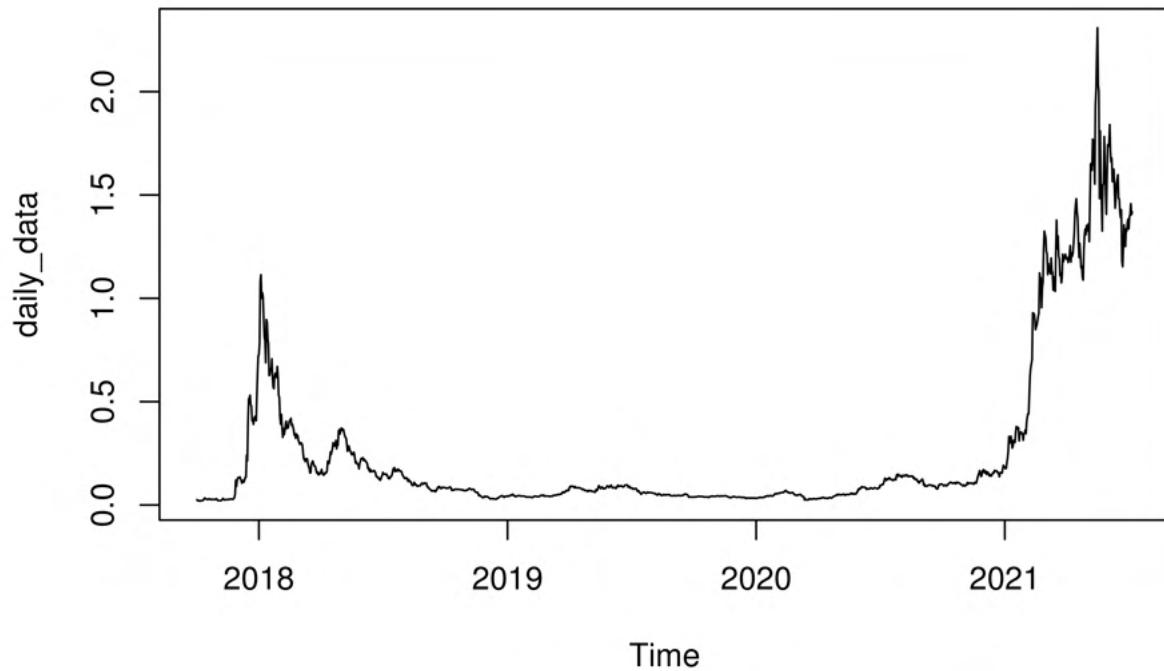
df$date <- as.Date(df$date, "%Y-%m-%d")

## Warning in as.POSIXlt.POSIXct(x, tz = tz): unknown timezone '%Y-%m-%d'

# Create a time series with daily data from 2017-10-02 to 2021-07-06
start_date <- as.Date("2017-10-02")
end_date <- as.Date(max(df$date))
daily_data <- ts(df$Close, start = c(2017, 275), frequency = 365)
plot(daily_data, main = 'original_dataset')

```

## original\_dataset



```
#####
#####
# Select data from 2018 to 2020
start_date <- as.Date("2018-10-01")
end_date <- as.Date("2020-8-31")

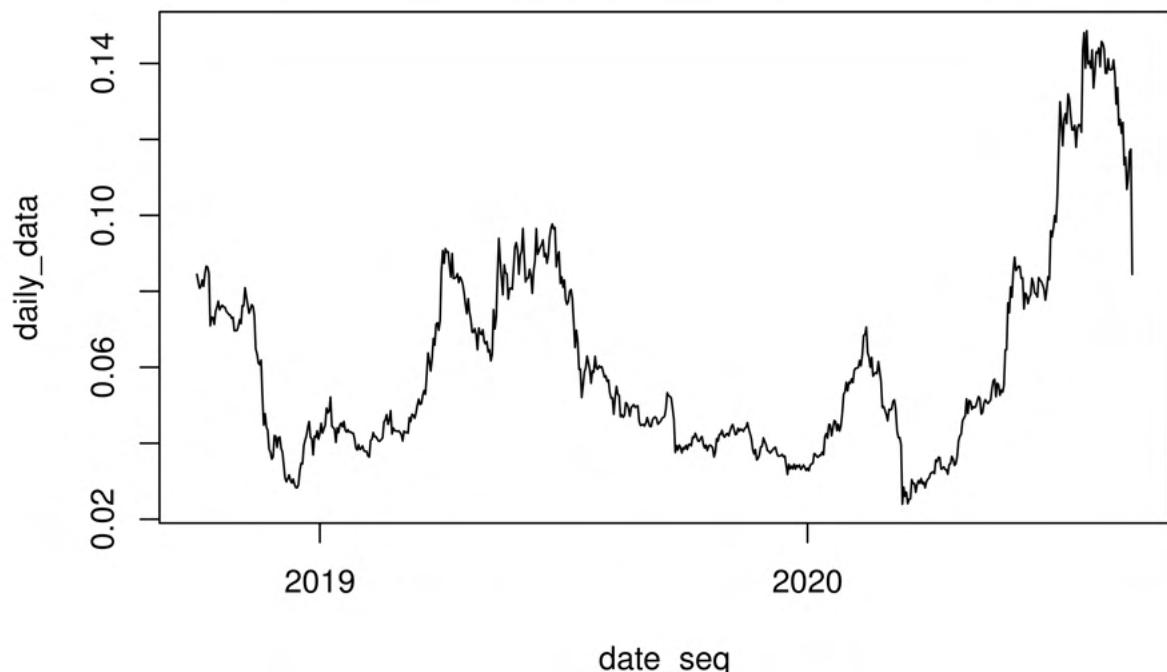
#####
#for prediction purpose: ## just next 60 values
df2 = df[df$Date >= end_date & df$Date <= as.Date("2020-10-31"),]
#####

## for model fitting purpose
df <- df[df$Date >= start_date & df$Date < end_date,]

#####
# Create a time series with daily data from 2018 to 2020
as.integer(format(as.Date("2018-10-01"), "%j"))

## [1] 274
daily_data <- ts(df$Close, start = c(2018, 274), end = c(2020, 244), frequency = 365)
date_seq <- seq(from = start_date, to = end_date, by = "day")
plot(date_seq,daily_data, type='l',main='Data between October-2018 and August-2020')
```

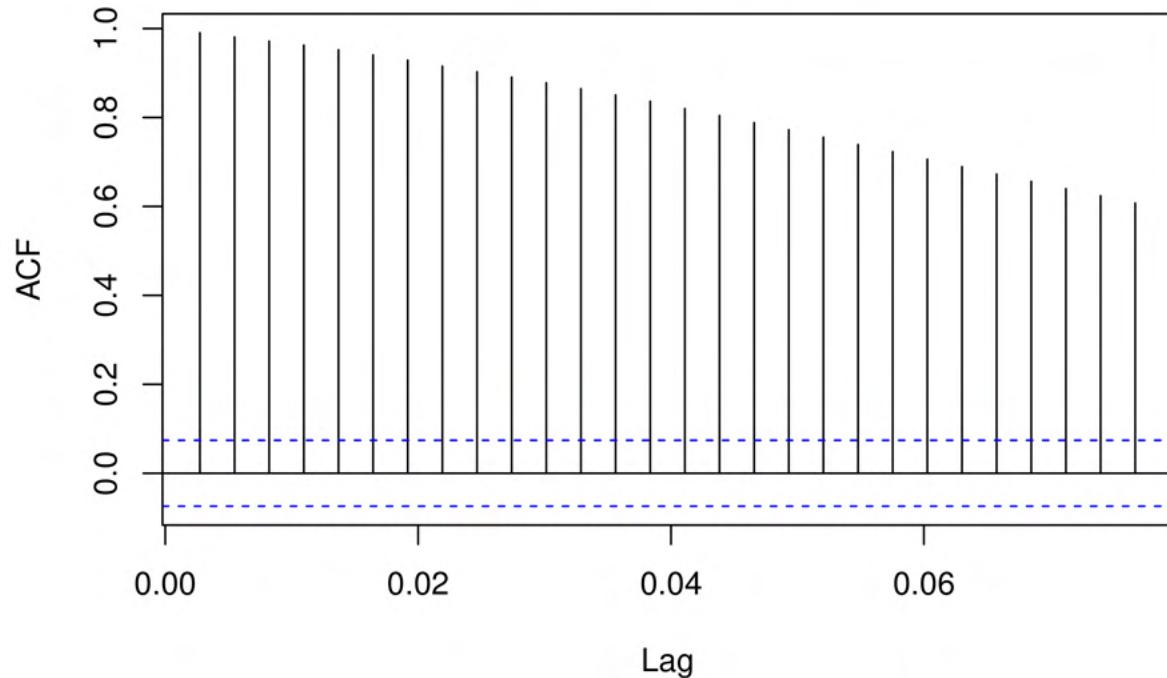
## Data between October–2018 and August–2020



```
#####
df3 <- data.frame(Date = seq(as.Date("2018-10-01"), as.Date("2020-08-31"), by = "day"),
                    Close = daily_data)
#####

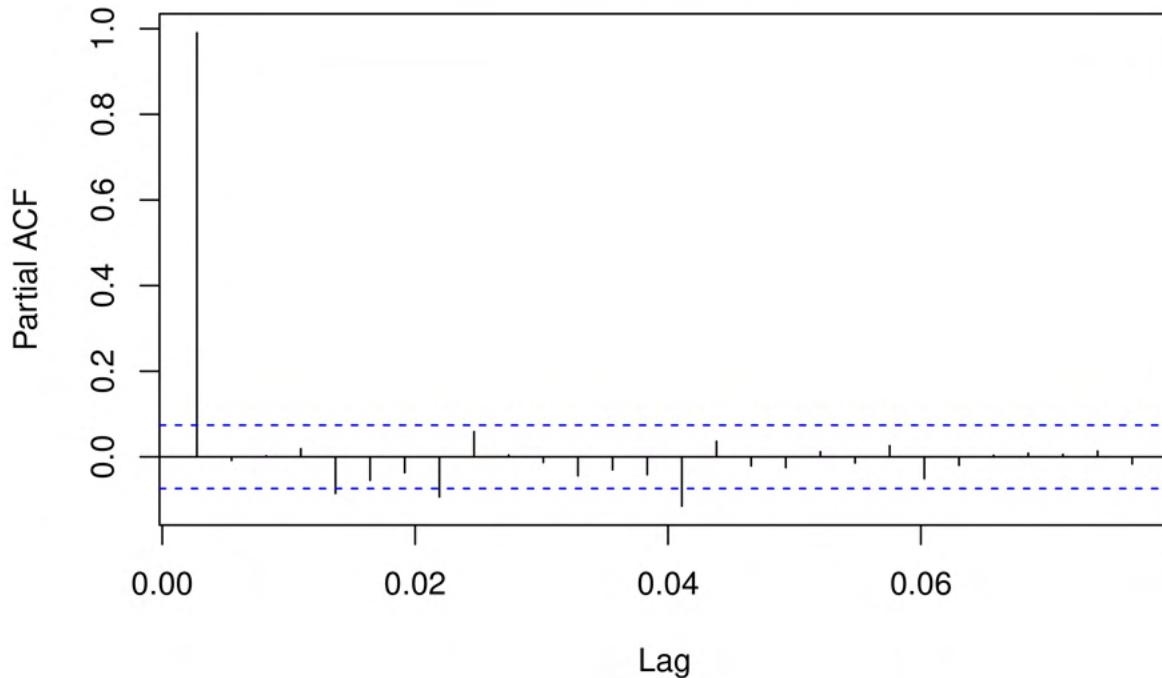
data1 = df3$Close
acf(data1, main = 'ACF of original data')
```

### ACF of original data



```
## exponentially dying
pacf(data1,main = 'PACF of original data' )
```

## PACF of original data



```
## Seems like AR(1) model

## Performing dickey Fuller Test

adf_test = adf.test(data1)
print(adf_test)

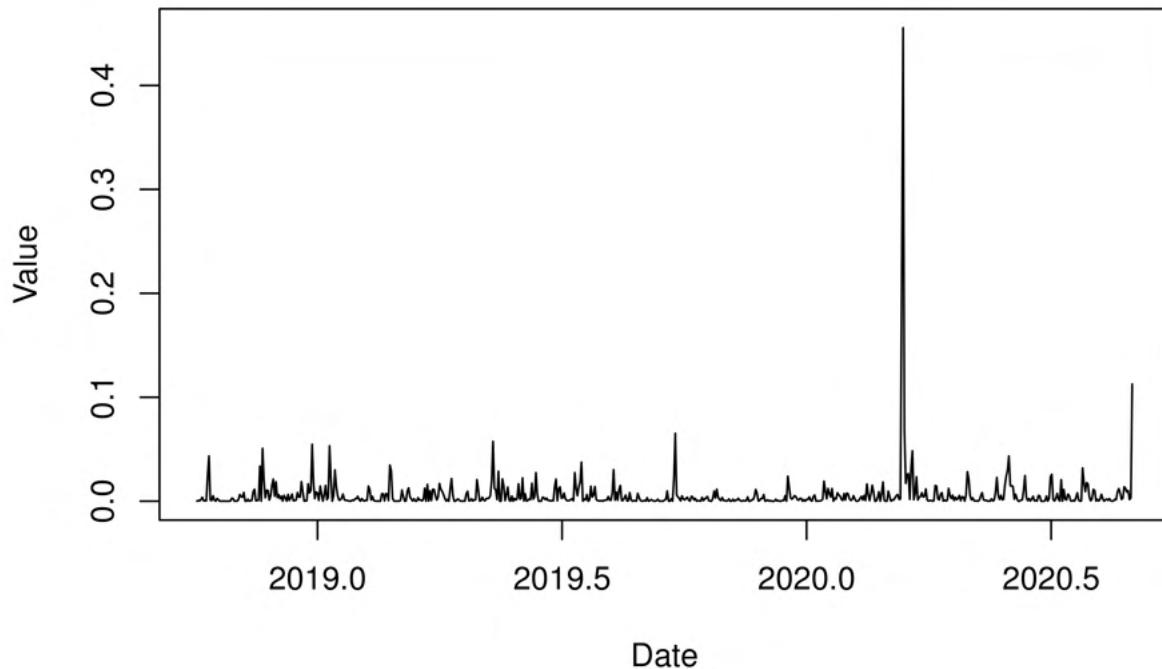
##
## Augmented Dickey-Fuller Test
##
## data: data1
## Dickey-Fuller = -2.2735, Lag order = 8, p-value = 0.4626
## alternative hypothesis: stationary
## Since p value > 0.05: fail to reject H0: it means non-stationary

#####
## Log differencing

close_diff = (diff(diff(log(data1))))^2

plot(close_diff, type = "l", xlab = "Date", ylab = "Value", main = "Line Graph for square of log-double
```

## Line Graph for square of log–double differenced data



```
adf_test = adf.test(close_diff)

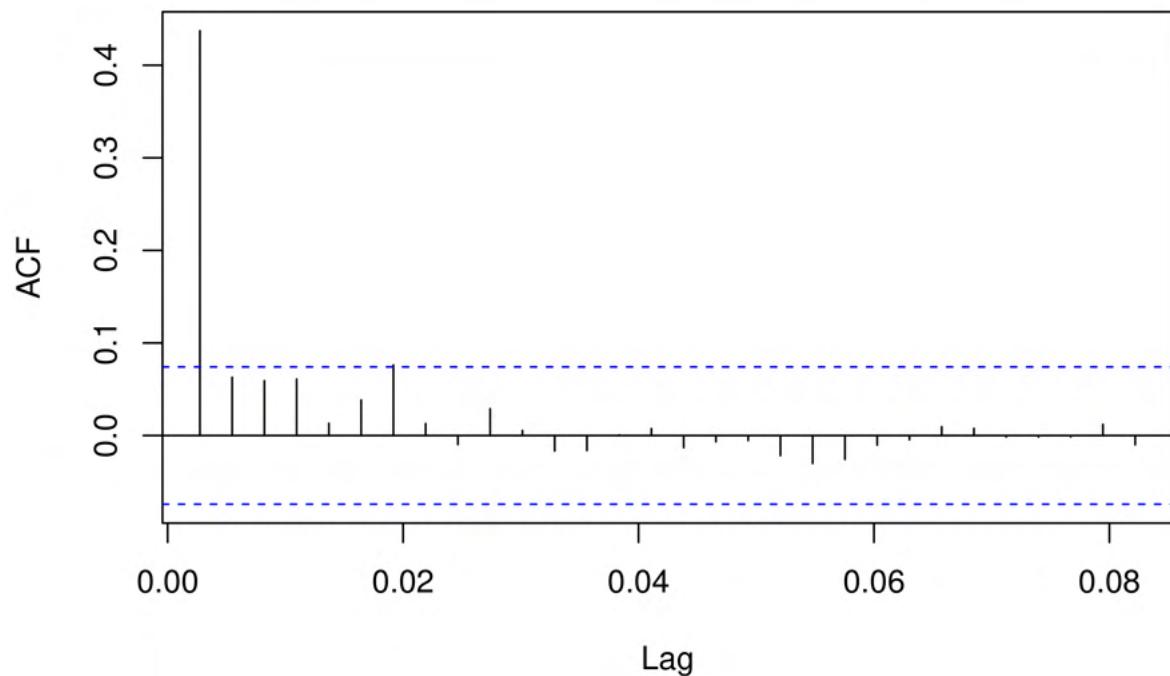
## Warning in adf.test(close_diff): p-value smaller than printed p-value
print(adf_test)

##
##  Augmented Dickey-Fuller Test
##
## data: close_diff
## Dickey-Fuller = -7.5784, Lag order = 8, p-value = 0.01
## alternative hypothesis: stationary
## Means the data has become stationary

## acf and pacf plots

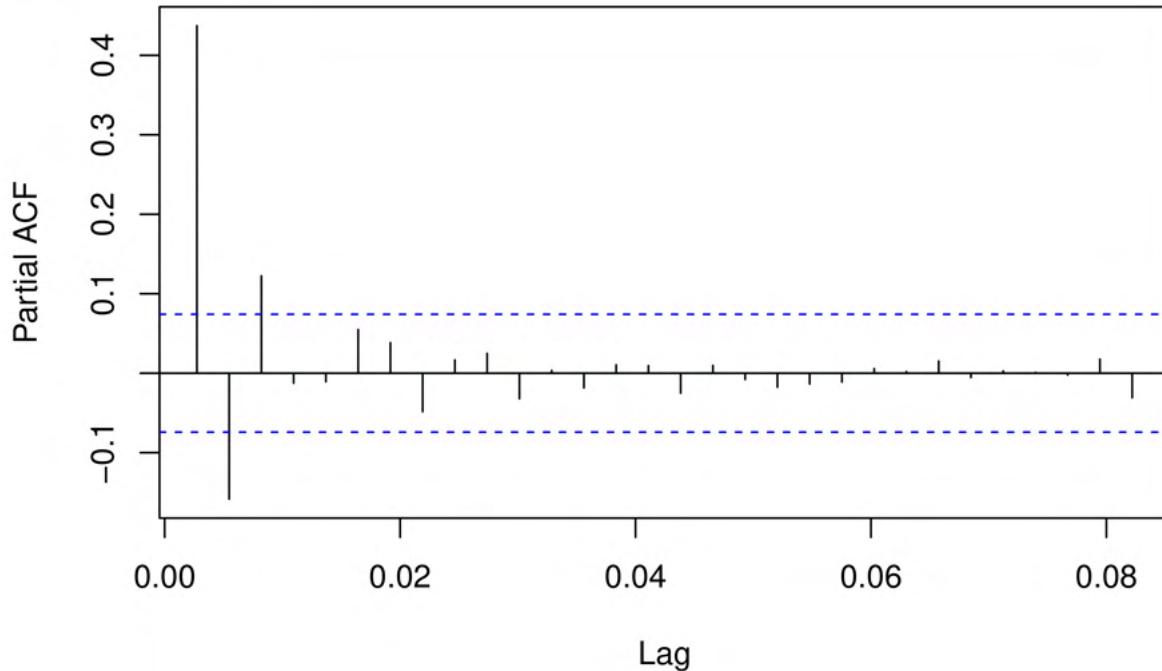
acf(close_diff, lag.max = 30, main = 'ACF of square of log-double differenced data')
```

### ACF of square of log–double differenced data



```
pacf(close_diff, lag.max = 30, main = 'PACF of square of log-double differenced data')
```

## PACF of square of log-double differenced data



```
eacf(close_diff)

## AR/MA
## 0 1 2 3 4 5 6 7 8 9 10 11 12 13
## 0 x o o o o o x o o o o o o o
## 1 x x o o o o x o o o o o o o
## 2 x x o o o o x o o o o o o o
## 3 x x x x o o o o o o o o o
## 4 x o x x x o o o o o o o o o
## 5 x x x x x o o o o o o o o o
## 6 x x x x x o x o o o o o o o
## 7 x x x o o o x o o o o o o o

#####
## suppose Models:
## ARIMA(3,2,2) ## actually from the model
## ARIMA(0,2,2) ## from the eacfs
## ARIMA(1,2,2)
## ARIMA(3,2,3)
## ARIMA(2,2,3)

#####
```

```
## Defining Function:
```

```

plot_arima_resid <- function(model) {
  # Extract the residuals from the model
  resid <- resid(model)

  # Create a time series of the residuals
  resid_ts <- ts(resid, start = c(2018, 274), end = c(2020, 365), frequency = 365)

  # Get the values of p, d, and q from the ARIMA model
  p <- model$arma[1]
  q <- model$arma[2]

  # Plot the residuals with custom x-axis labels and title with ARIMA values
  plot(resid_ts, xlab = "Year", ylab = "Residuals", main = paste("Residuals Plot (ARIMA(", p, ", ", 2, "))

  acf(resid_ts[1:length(resid_ts)], lag.max = 150, main = paste("Residuals ACF Plot (ARIMA(", p, ", ", 2,
  pacf(resid_ts[1:length(resid_ts)], lag.max = 150, main = paste("Residuals PACF Plot (ARIMA(", p, ", ", 2, "))

}

#####
# create an empty data frame with column names
# create an empty data frame with column names
my_df <- data.frame(model = character(),
                     AIC = numeric(),
                     BIC = numeric(),
                     Shapiro = round(numeric(), 3),
                     Ljung = round(numeric(), 3))

new_row <- data.frame(model = "", AIC = 0, BIC = 0, Shapiro = 0, Ljung = 0)

plot_arima_residuals <- function(arima_model, my_df) {
  # extract residuals and create time series
  resid <- residuals(arima_model)
  resid_ts <- ts(resid, start = c(2018, 274), end = c(2020, 365), frequency = 365)
  plot_arima_resid(arima_model)

  # normal Q-Q plot of residuals
  qqnorm(resid_ts, main = "Residuals plot ")
  qqline(resid_ts)

  # Shapiro-Wilk test of normality
  shap = shapiro.test(resid_ts)

  # Ljung-Box test of autocorrelation
  ljung = LB.test(arima_model)

  # AIC and BIC
  cat("AIC:", AIC(arima_model), "\n")
  cat("BIC:", BIC(arima_model), "\n")

  # create a new row for the data frame and append it
}

```

```

new_row$model <- as.character(arima_model)
new_row$AIC <- AIC(arima_model)
new_row$BIC <- BIC(arima_model)
new_row$Shapiro <- shap$p.value
new_row$Ljung <- round(ljung$p.value,3)
my_df <- rbind(my_df, new_row)

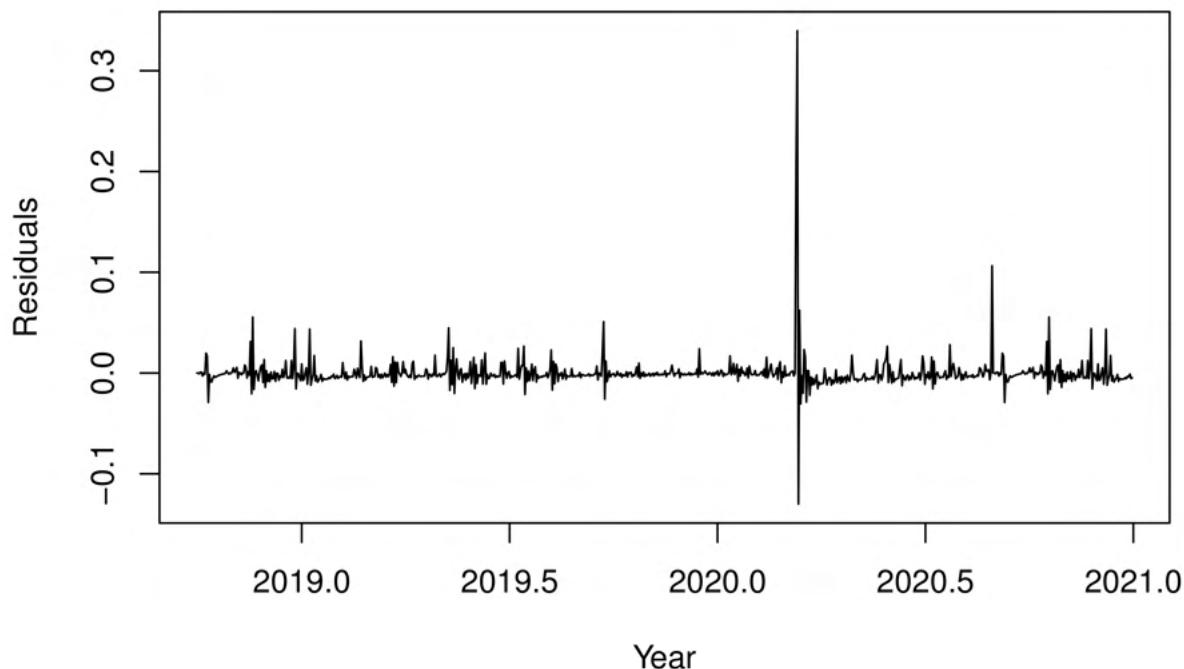
# return the updated data frame
return(my_df)
}

# call the function and update the data frame
arima322 <- arima(x = close_diff, order = c(3, 2, 2))
arima022 <- arima(x = close_diff, order = c(0, 2, 2))
arima122 <- arima(x = close_diff, order = c(1, 2, 2))
arima323 <- arima(x = close_diff, order = c(3, 2, 3))
arima223 <- arima(x = close_diff, order = c(2, 2, 3))

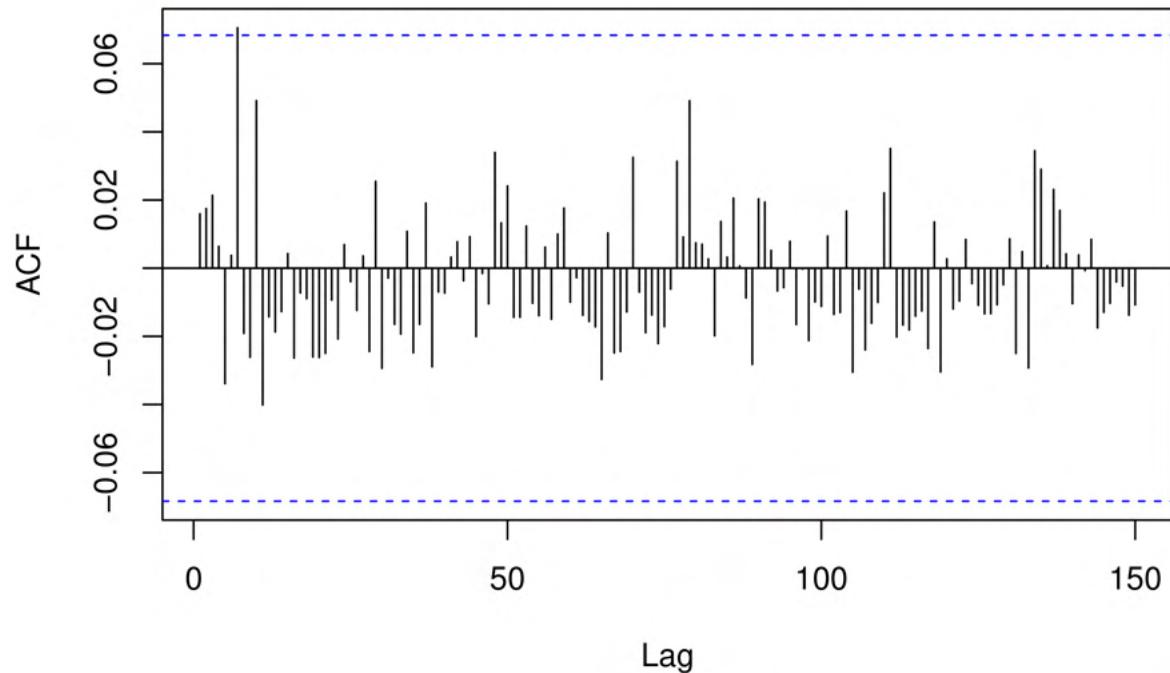
my_df <- plot_arima_residuals(arima322, my_df)

```

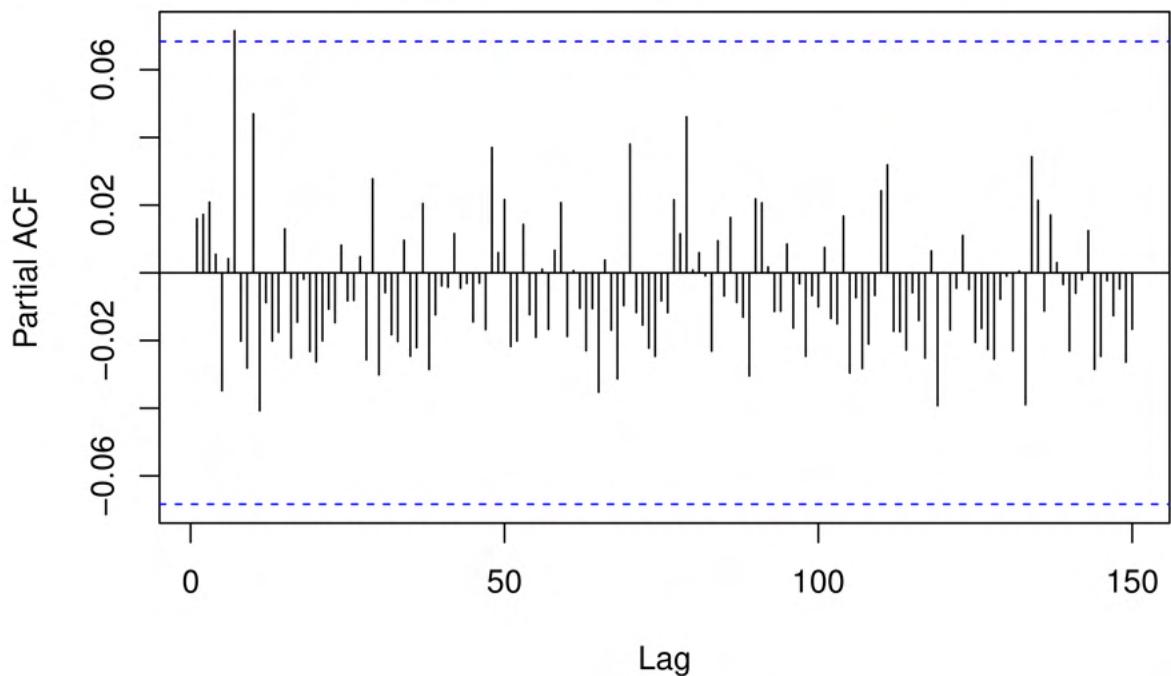
### Residuals Plot (ARIMA( 3 , 2 , 2 ))



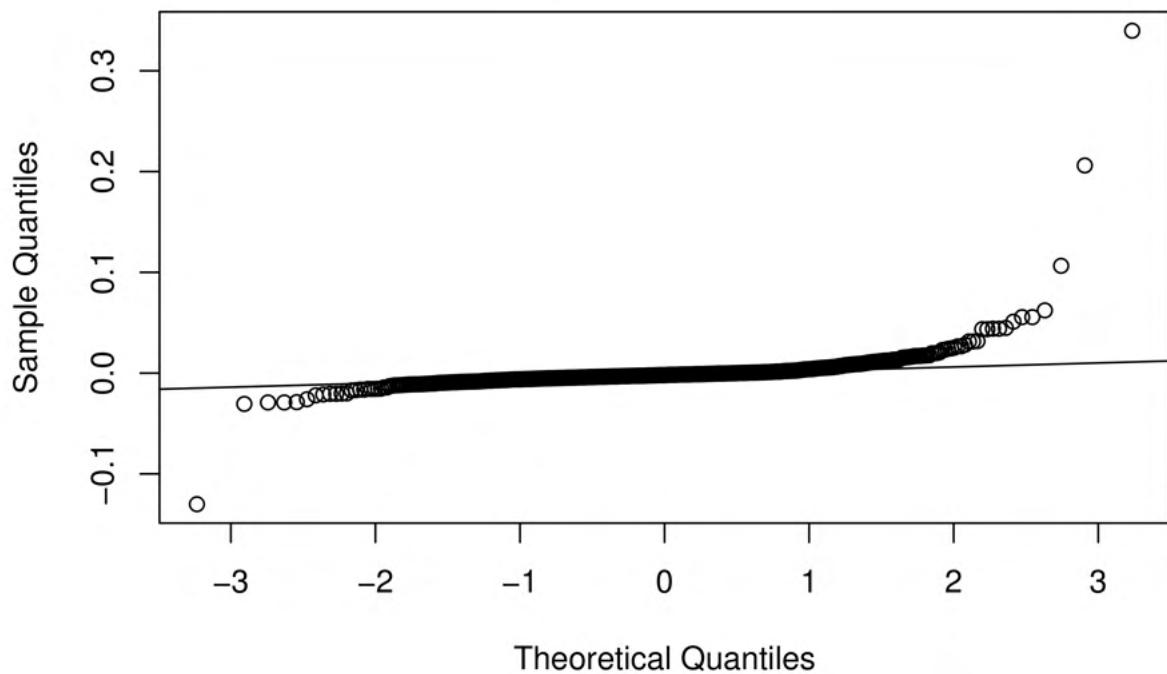
### Residuals ACF Plot (ARIMA( 3 , 2 , 2 ))



### Residuals PACF Plot (ARIMA( 3 , 2 , 2 ))

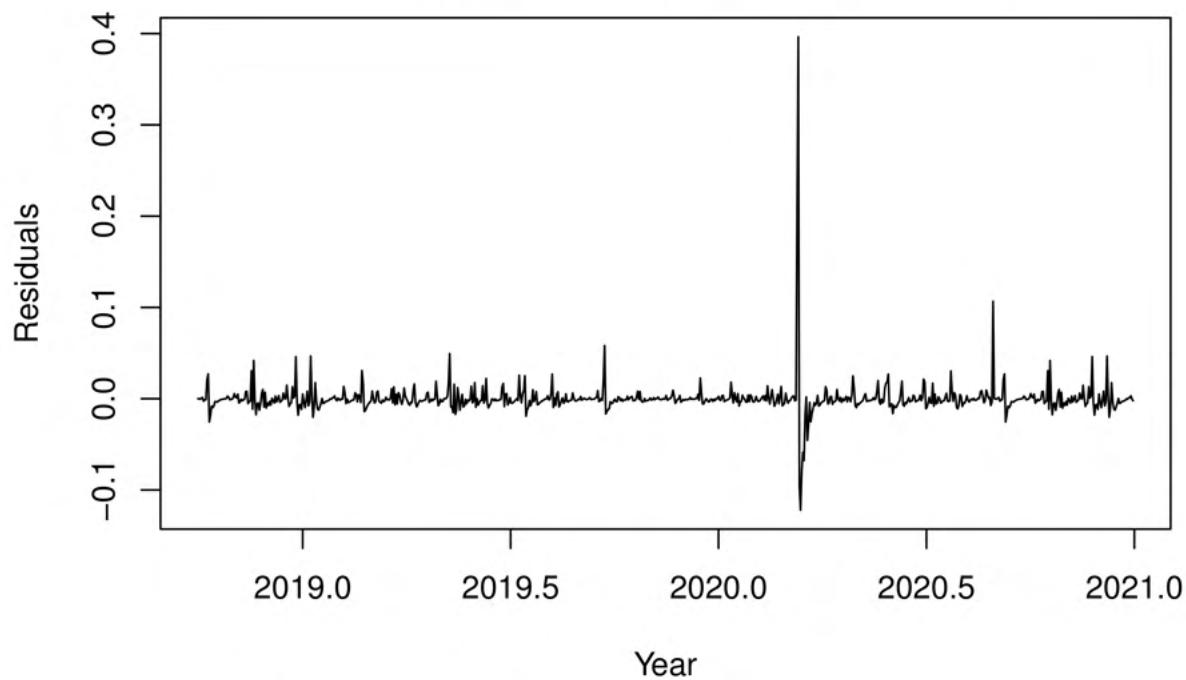


## Residuals plot

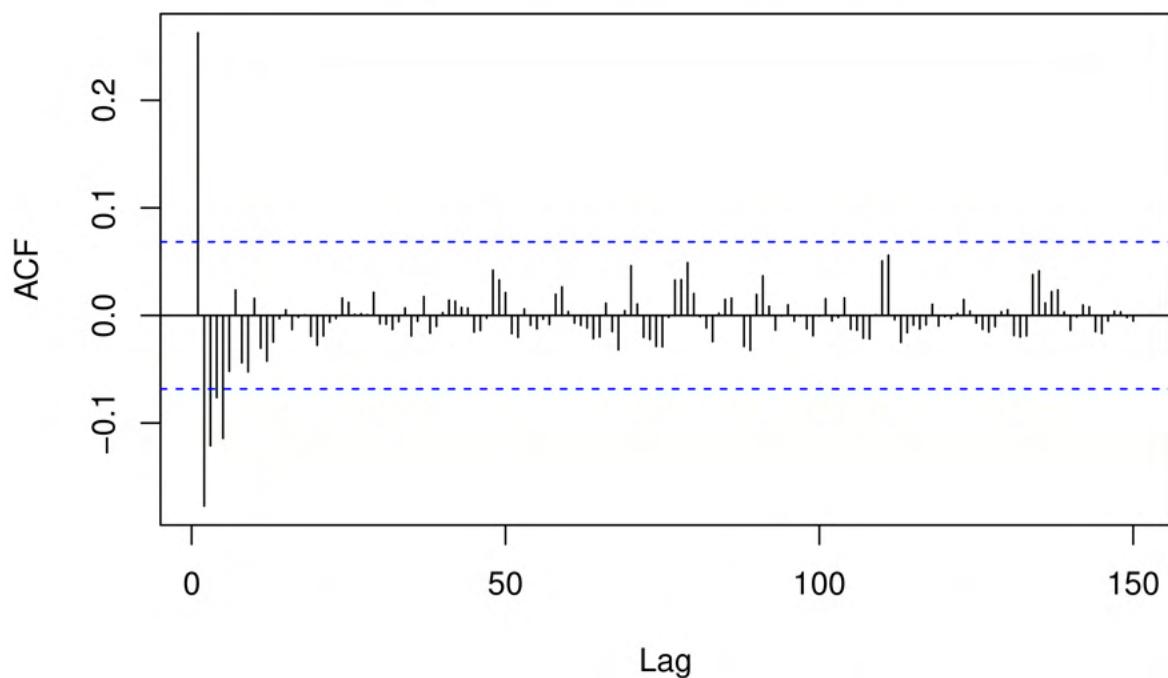


```
## AIC: -3559.729
## BIC: -3532.449
my_df <- plot_arima_residuals(arima022, my_df)
```

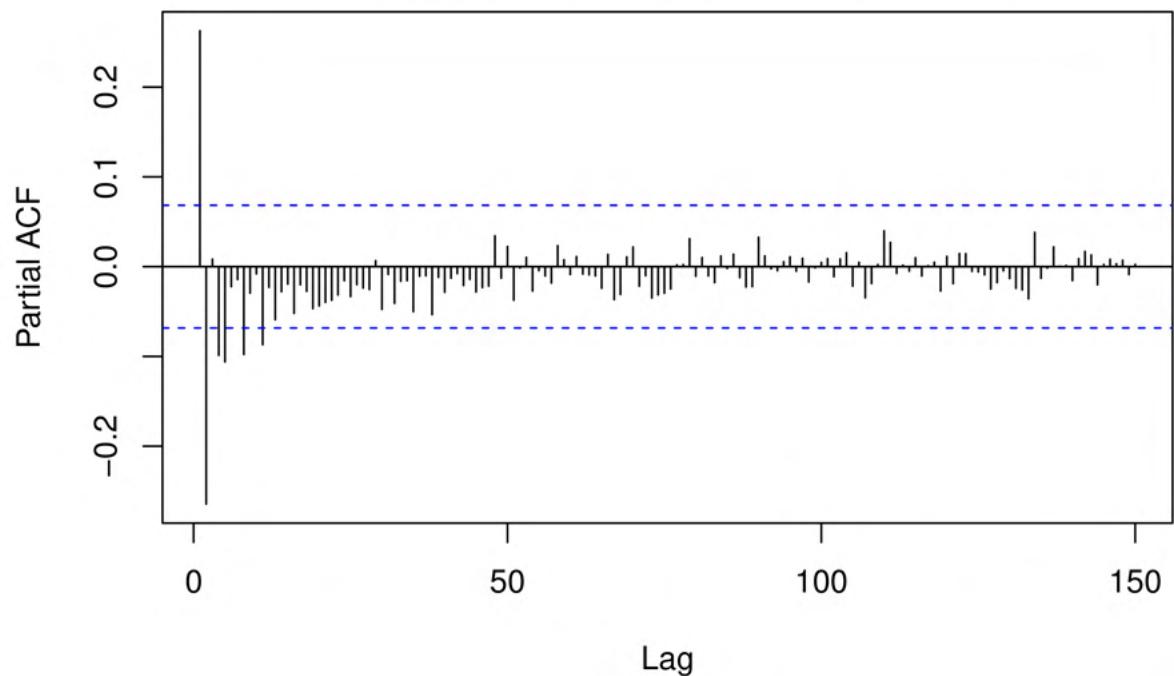
### Residuals Plot (ARIMA( 0 , 2 , 2 ))



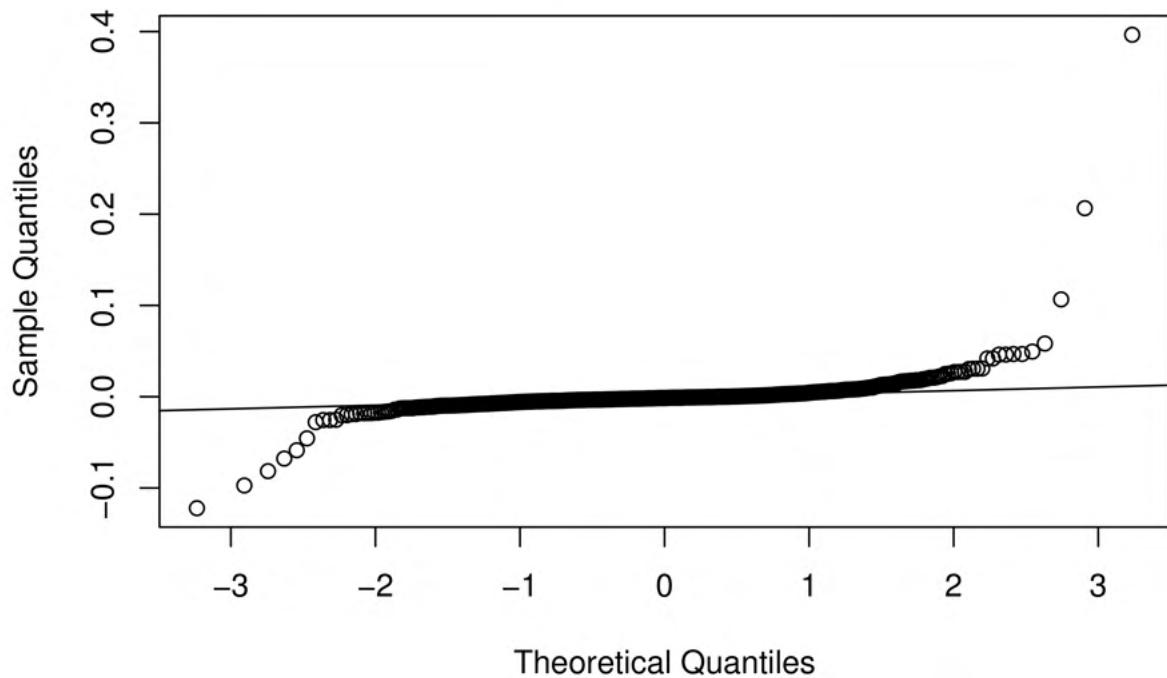
### Residuals ACF Plot (ARIMA( 0 , 2 , 2 ))



**Residuals PACF Plot (ARIMA( 0 , 2 , 2 ))**

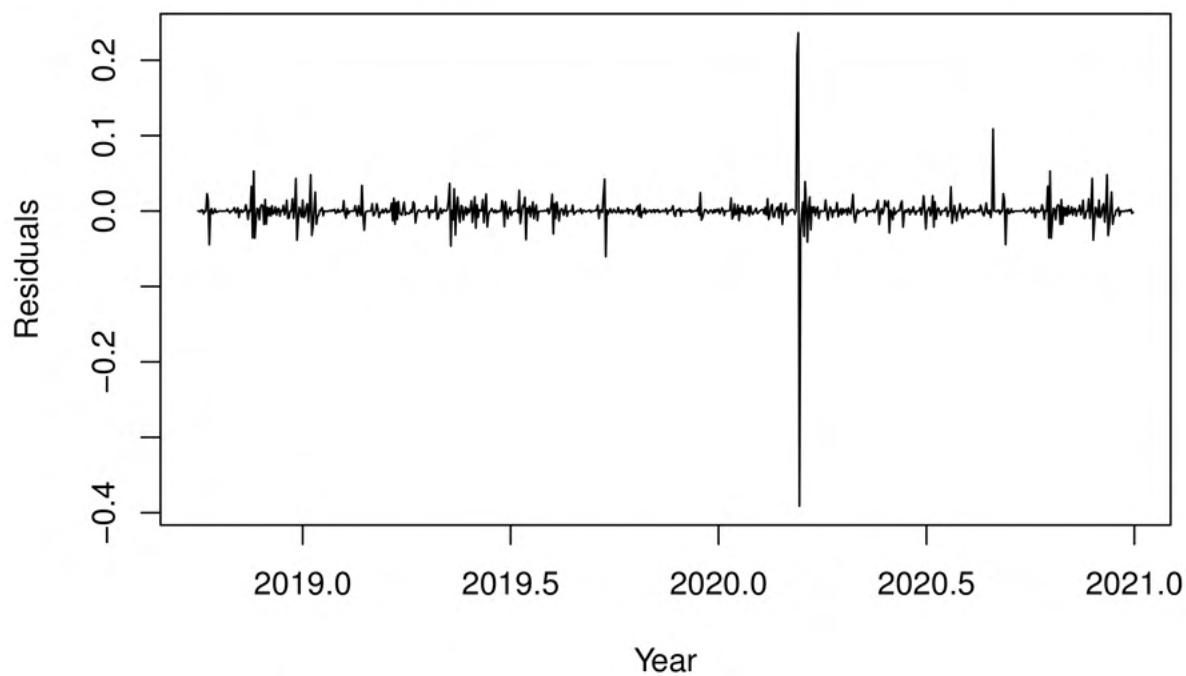


## Residuals plot

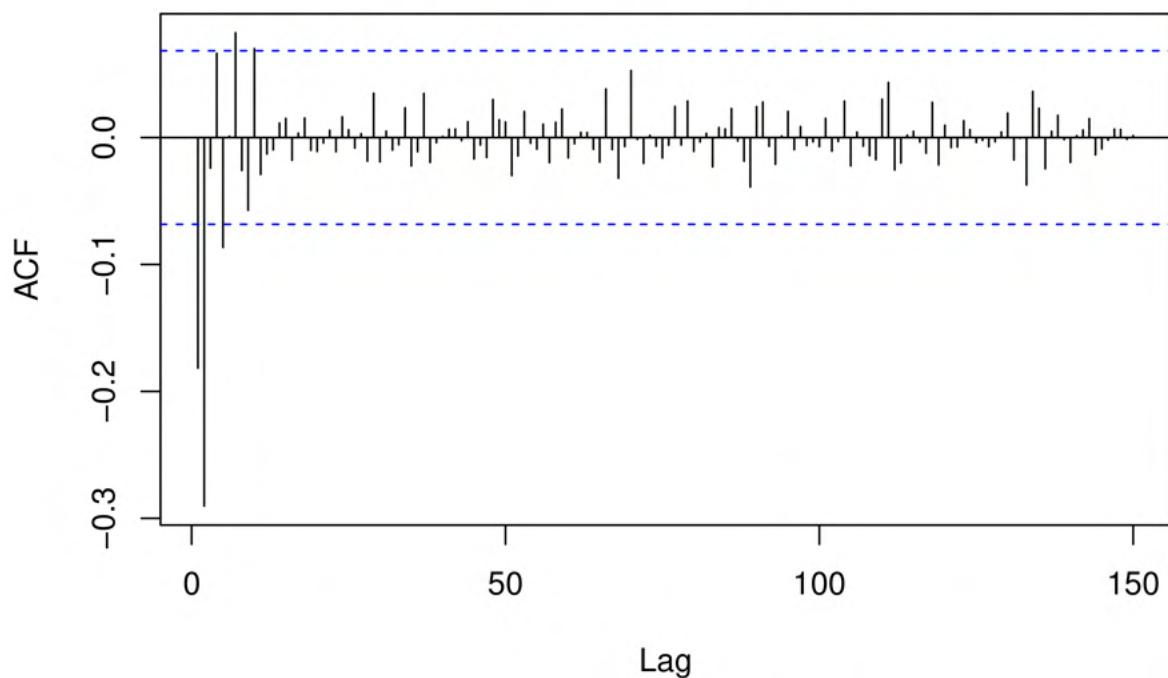


```
## AIC: -3409.742
## BIC: -3396.101
my_df <- plot_arima_residuals(arima122, my_df)
```

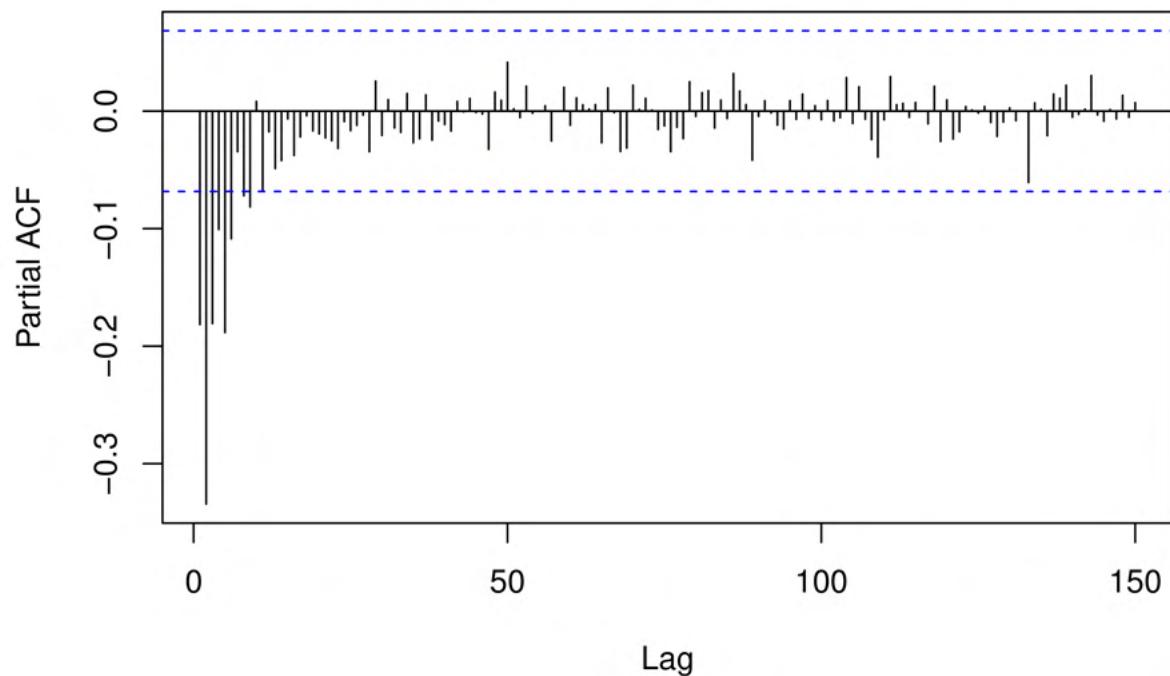
### Residuals Plot (ARIMA( 1 , 2 , 2 ))



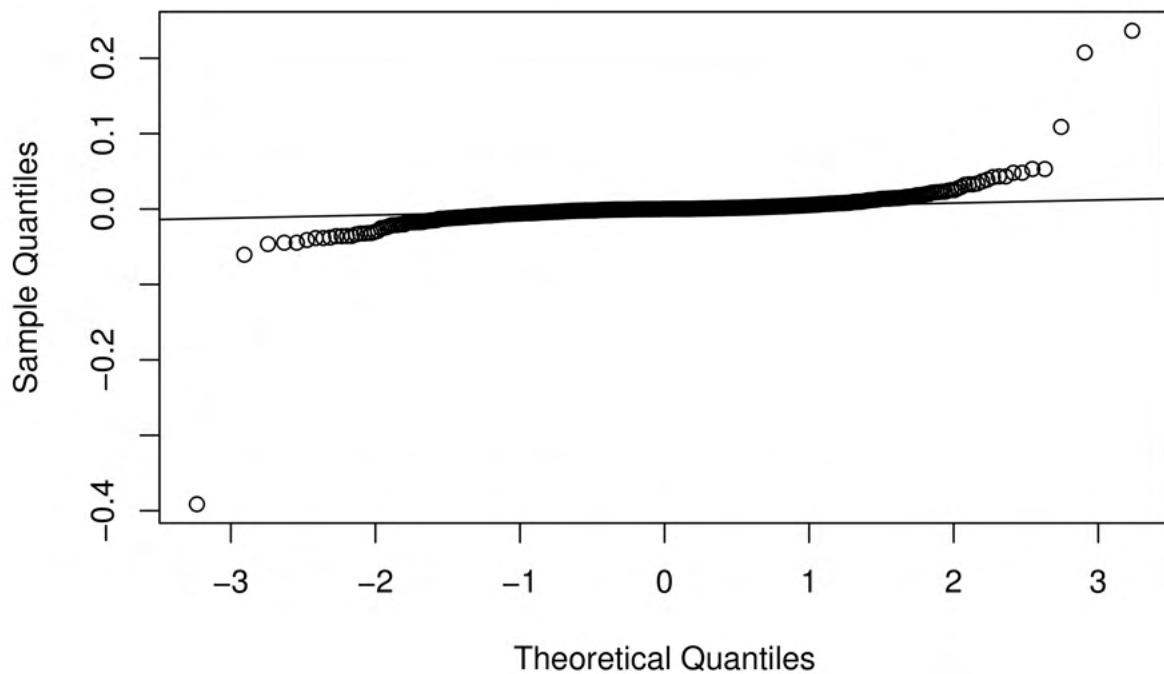
### Residuals ACF Plot (ARIMA( 1 , 2 , 2 ))



### Residuals PACF Plot (ARIMA( 1 , 2 , 2 ))

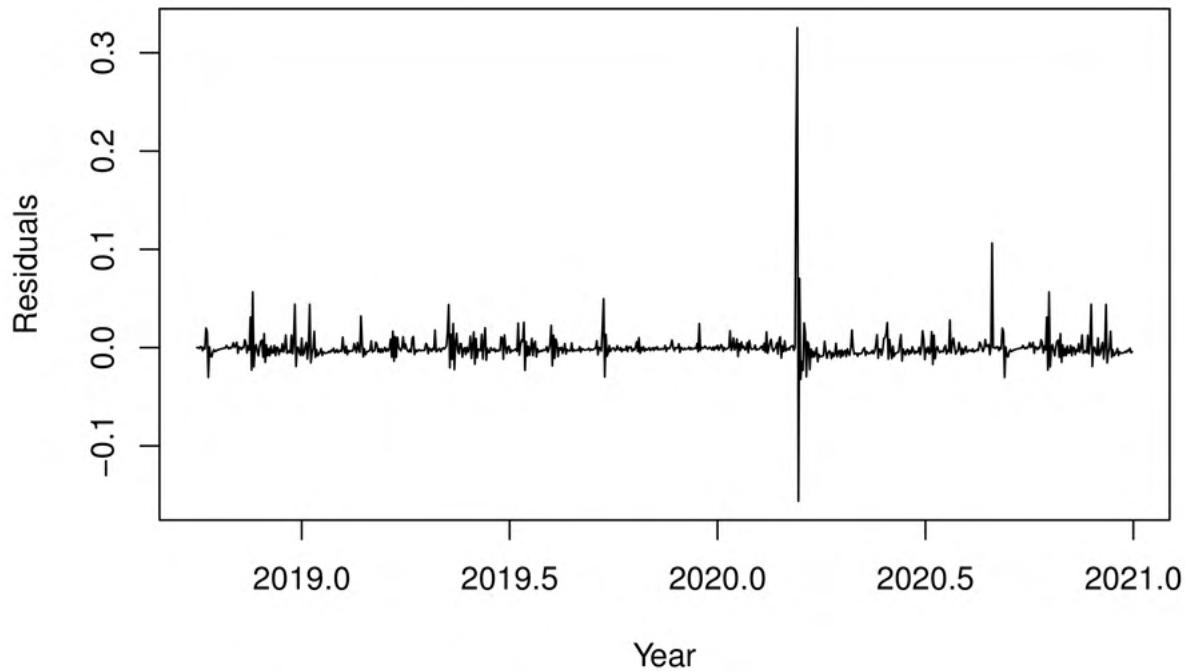


## Residuals plot

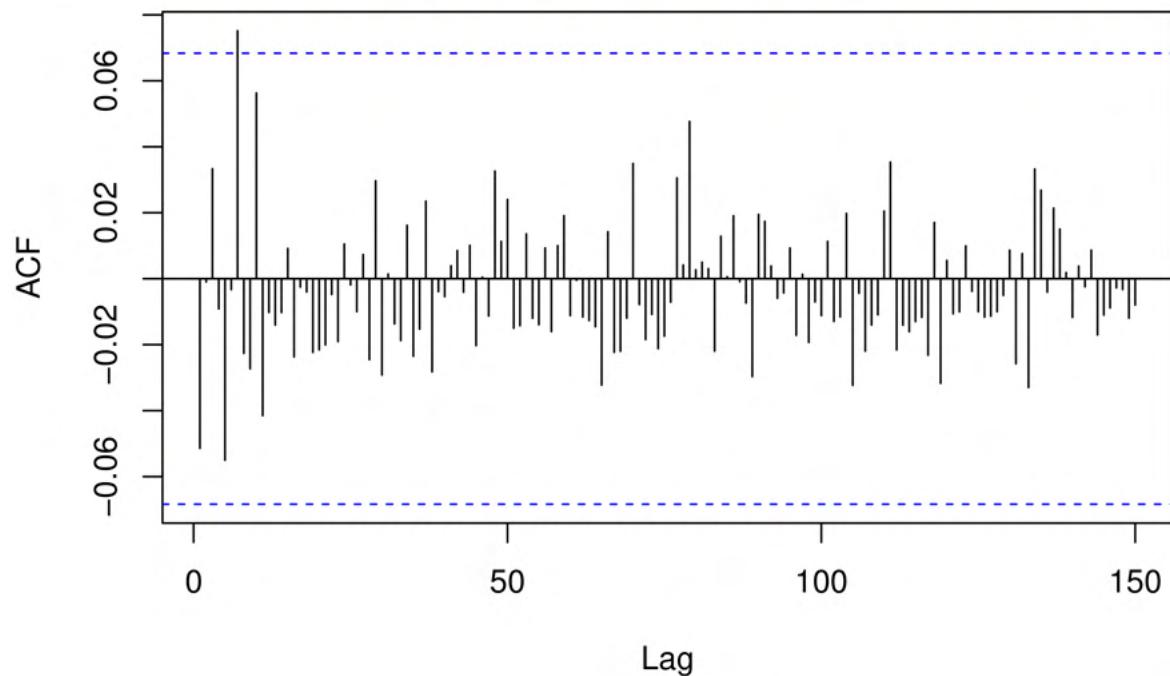


```
## AIC: -3336.977
## BIC: -3318.79
my_df <- plot_arima_residuals(arima323, my_df)
```

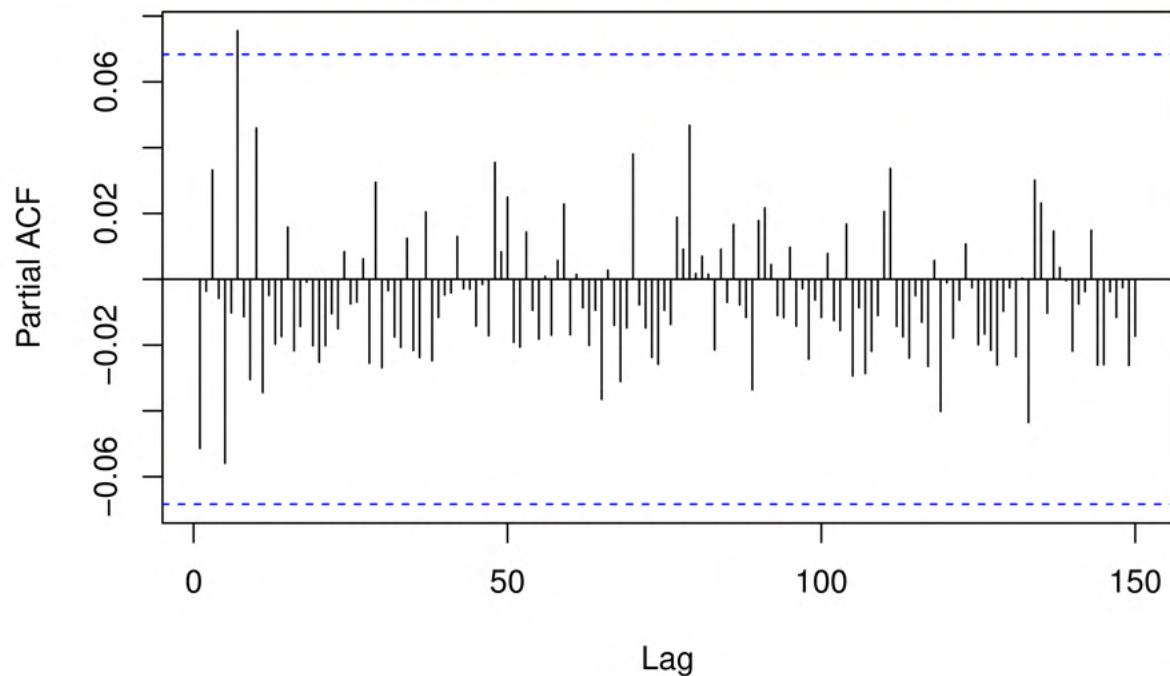
### Residuals Plot (ARIMA( 3 , 2 , 3 ))



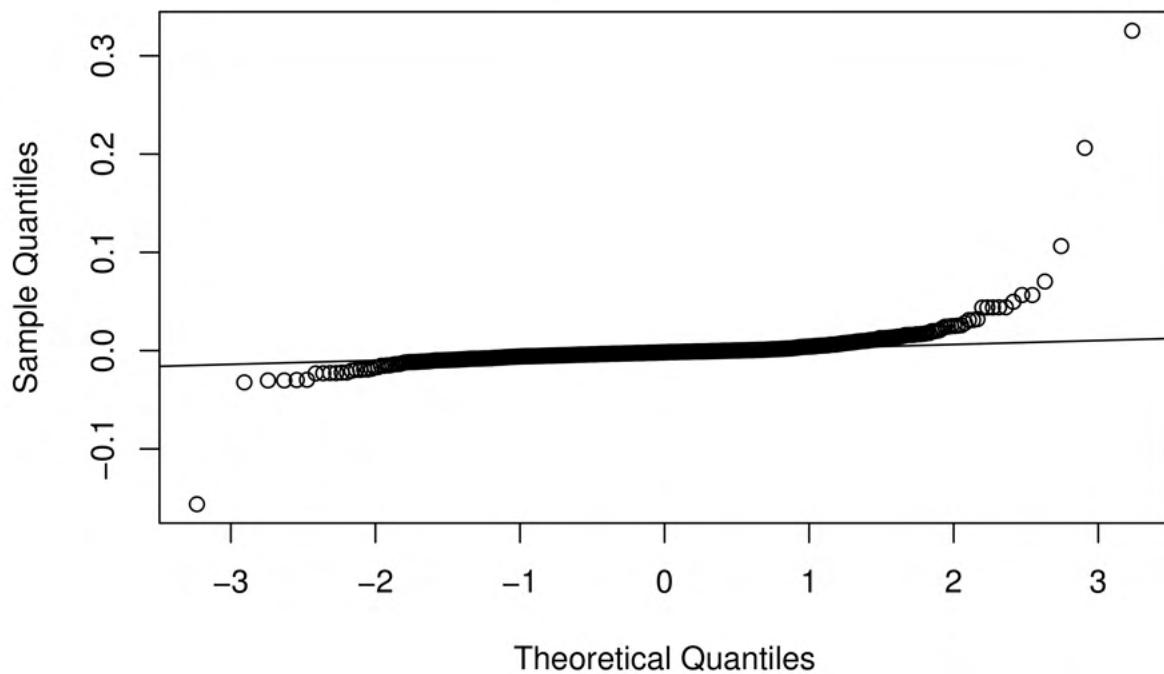
### Residuals ACF Plot (ARIMA( 3 , 2 , 3 ))



### Residuals PACF Plot (ARIMA( 3 , 2 , 3 ))

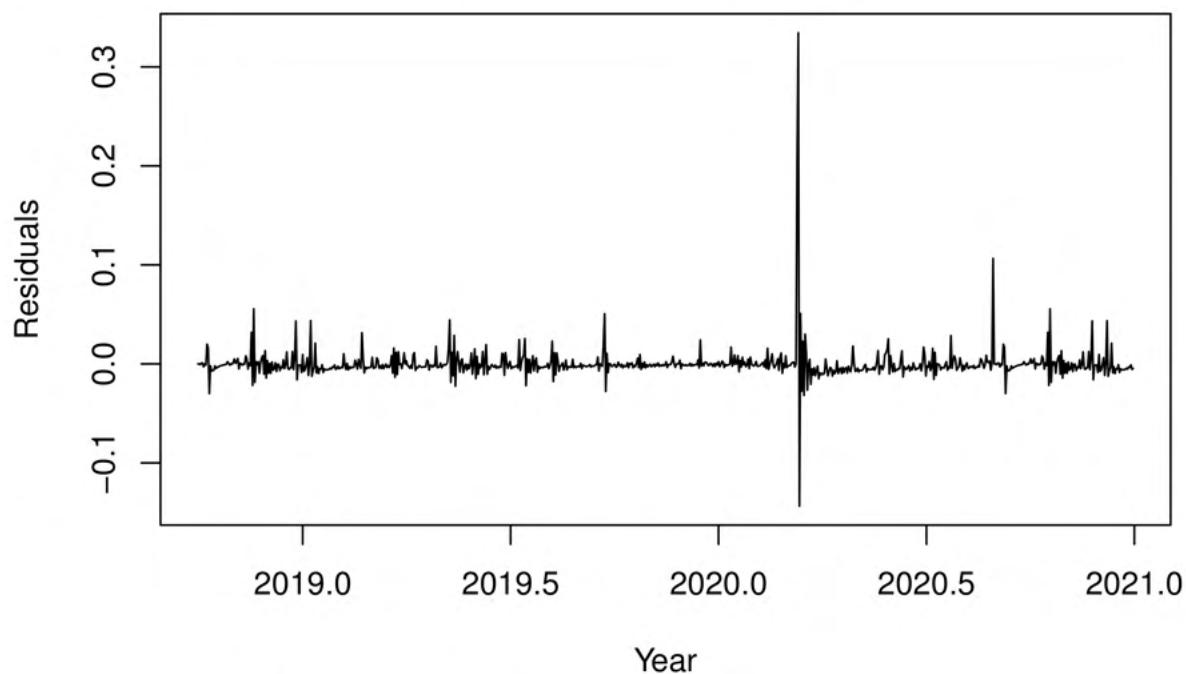


## Residuals plot

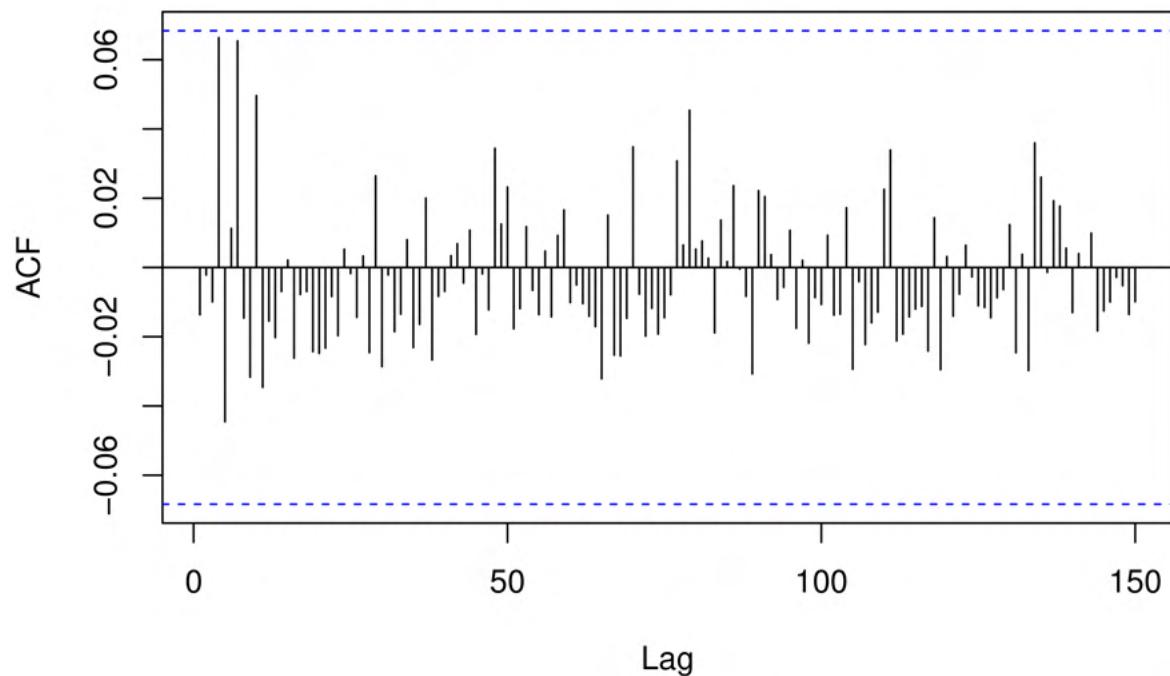


```
## AIC: -3556.466
## BIC: -3524.639
my_df <- plot_arima_residuals(arima223, my_df)
```

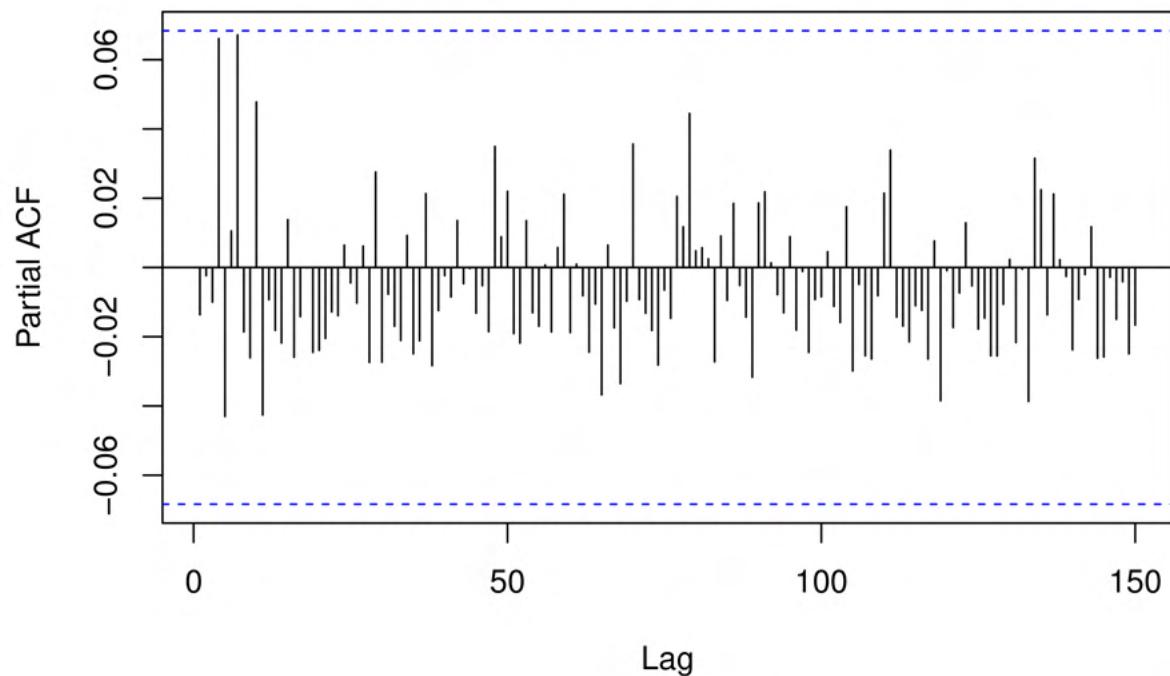
### Residuals Plot (ARIMA( 2 , 2 , 3 ))



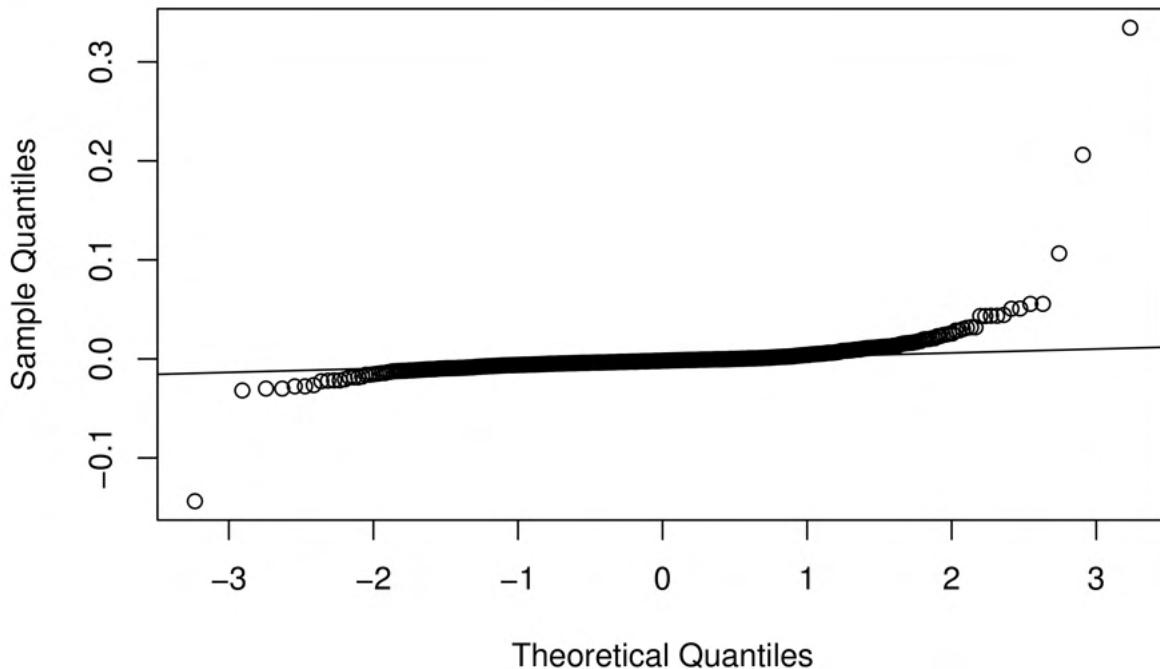
### Residuals ACF Plot (ARIMA( 2 , 2 , 3 ))



### Residuals PACF Plot (ARIMA( 2 , 2 , 3 ))



## Residuals plot



```
## AIC: -3558.017
## BIC: -3530.736
# view updated data frame
print(my_df)

##          model      AIC      BIC   Shapiro Ljung
## 1 ARIMA(3,2,2) -3559.729 -3532.449 1.556828e-46 0.341
## 2 ARIMA(0,2,2) -3409.742 -3396.101 8.646146e-47 0.000
## 3 ARIMA(1,2,2) -3336.977 -3318.790 4.176565e-46 0.000
## 4 ARIMA(3,2,3) -3556.466 -3524.639 3.768687e-46 0.099
## 5 ARIMA(2,2,3) -3558.017 -3530.736 2.529240e-46 0.213
#####
#####
cat("## Garch Model ")

## ## Garch Model
return = diff(log(df3$Close))

ArchTest(return)

##
## ARCH LM-test; Null hypothesis: no ARCH effects
##
## data: return
## Chi-squared = 17.207, df = 12, p-value = 0.142
```

```

## No ARCH EFFECT is there

#####
#####

## finding the GARCH order

garch(x=return,grad = 'numerical', trace = FALSE)

##
## Call:
## garch(x = return, grad = "numerical", trace = FALSE)
##
## Coefficient(s):
##      a0      a1      b1
## 0.002689  0.050000  0.050000
## a1 and b1: MEANS : alpha = 1 and beta = 1
## GARCH(1,1)

#####
#####

## Garch Model - Normal
## Normal: Pearson value < 0.05: not good
## sstd: pearson > 0.05: good

library(rugarch)

## Warning: package 'rugarch' was built under R version 4.2.3
## Loading required package: parallel
##
## Attaching package: 'rugarch'
##
## The following object is masked from 'package:purrr':
## 
##     reduce
##
## The following object is masked from 'package:stats':
## 
##     sigma

## A) using Normal distribution model

return = diff(log(df3$Close))

s_norm <- ugarchspec(mean.model = list(armaOrder = c(3,2)),variance.model = list(garchOrder = c(1,1)),
                      distribution.model = 'sstd')

m_norm <- ugarchfit(data = return, spec = s_norm, solver ='hybrid')
m_norm

##
## *-----*

```

```

## *          GARCH Model Fit      *
## *-----*
##
## Conditional Variance Dynamics
## -----
## GARCH Model : sGARCH(1,1)
## Mean Model  : ARFIMA(3,0,2)
## Distribution : sstd
##
## Optimal Parameters
## -----
##           Estimate Std. Error t value Pr(>|t|)
## mu      -0.000156   0.002024 -0.076878 0.938720
## ar1      0.383922   0.397342  0.966227 0.333930
## ar2      0.438508   0.304523  1.439983 0.149872
## ar3      0.033814   0.063401  0.533335 0.593802
## ma1     -0.506525   0.396307 -1.278112 0.201210
## ma2     -0.342009   0.352932 -0.969051 0.332520
## omega    0.000429   0.000187  2.292001 0.021906
## alpha1   0.130492   0.058347  2.236495 0.025319
## beta1    0.777131   0.069215 11.227800 0.000000
## skew     0.976442   0.049275 19.816091 0.000000
## shape    3.048401   0.418177  7.289738 0.000000
##
## Robust Standard Errors:
##           Estimate Std. Error t value Pr(>|t|)
## mu      -0.000156   0.002195 -0.070908 0.943471
## ar1      0.383922   0.096069  3.996334 0.000064
## ar2      0.438508   0.097503  4.497395 0.000007
## ar3      0.033814   0.029899  1.130940 0.258080
## ma1     -0.506525   0.095816 -5.286426 0.000000
## ma2     -0.342009   0.101531 -3.368520 0.000756
## omega    0.000429   0.000124  3.451614 0.000557
## alpha1   0.130492   0.048416  2.695198 0.007035
## beta1    0.777131   0.043756 17.760681 0.000000
## skew     0.976442   0.051900 18.813896 0.000000
## shape    3.048401   0.403132  7.561794 0.000000
##
## LogLikelihood : 1145.442
##
## Information Criteria
## -----
##           Akaike      -3.2413
##           Bayes      -3.1697
##           Shibata   -3.2417
##           Hannan-Quinn -3.2136
##
## Weighted Ljung-Box Test on Standardized Residuals
## -----
##           statistic p-value
## Lag[1]            2.714 0.09948
## Lag[2*(p+q)+(p+q)-1][14] 6.913 0.83702
## Lag[4*(p+q)+(p+q)-1][24] 10.629 0.74966

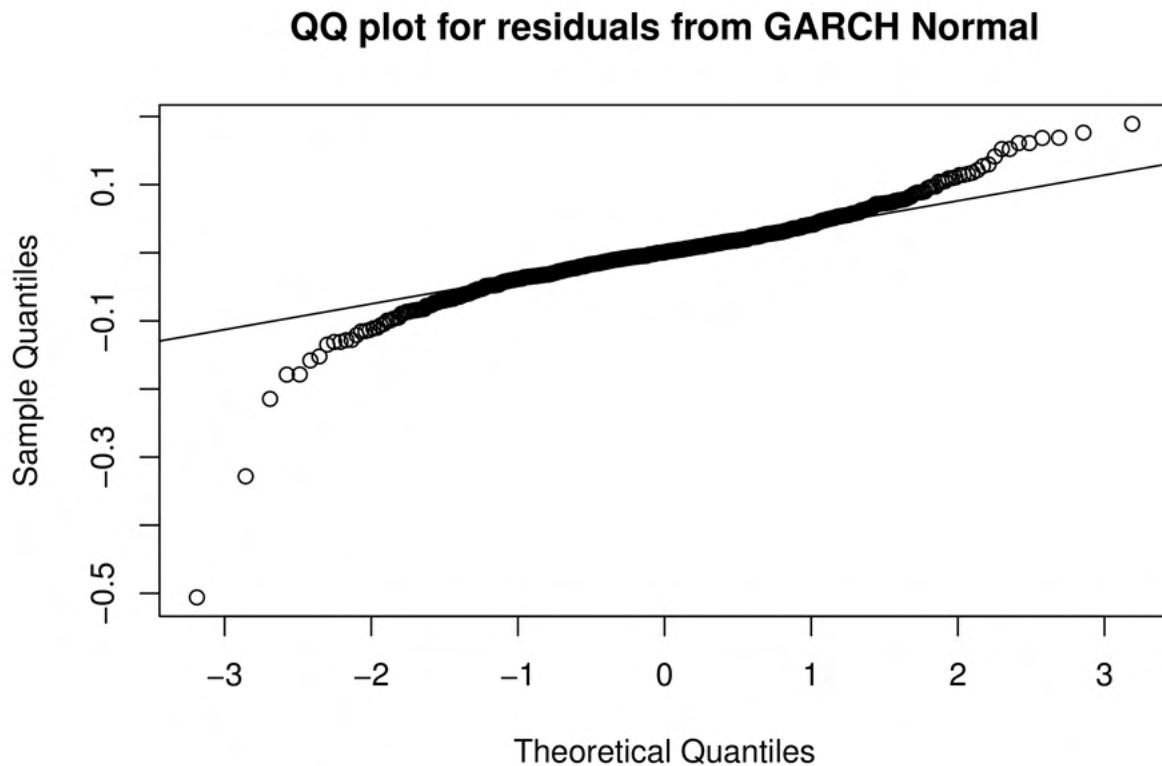
```

```

## d.o.f=5
## H0 : No serial correlation
##
## Weighted Ljung-Box Test on Standardized Squared Residuals
## -----
##                      statistic p-value
## Lag[1]                 0.1491  0.6994
## Lag[2*(p+q)+(p+q)-1][5] 1.0300  0.8529
## Lag[4*(p+q)+(p+q)-1][9] 1.7129  0.9364
## d.o.f=2
##
## Weighted ARCH LM Tests
## -----
##          Statistic Shape Scale P-Value
## ARCH Lag[3]    0.1013 0.500 2.000  0.7503
## ARCH Lag[5]    1.4514 1.440 1.667  0.6052
## ARCH Lag[7]    1.6245 2.315 1.543  0.7961
##
## Nyblom stability test
## -----
## Joint Statistic: 1.6203
## Individual Statistics:
## mu      0.24498
## ar1     0.12582
## ar2     0.07218
## ar3     0.09464
## ma1     0.12610
## ma2     0.07947
## omega   0.09535
## alpha1   0.16483
## beta1   0.10484
## skew    0.22460
## shape   0.10170
##
## Asymptotic Critical Values (10% 5% 1%)
## Joint Statistic:        2.49 2.75 3.27
## Individual Statistic:   0.35 0.47 0.75
##
## Sign Bias Test
## -----
##          t-value prob sig
## Sign Bias       0.008354 0.9933
## Negative Sign Bias 0.090984 0.9275
## Positive Sign Bias 0.665173 0.5062
## Joint Effect     0.748124 0.8618
##
## 
## 
## Adjusted Pearson Goodness-of-Fit Test:
## -----
## group statistic p-value(g-1)
## 1      20      10.57      0.9374
## 2      30      32.00      0.3199
## 3      40      27.77      0.9101
## 4      50      41.71      0.7604

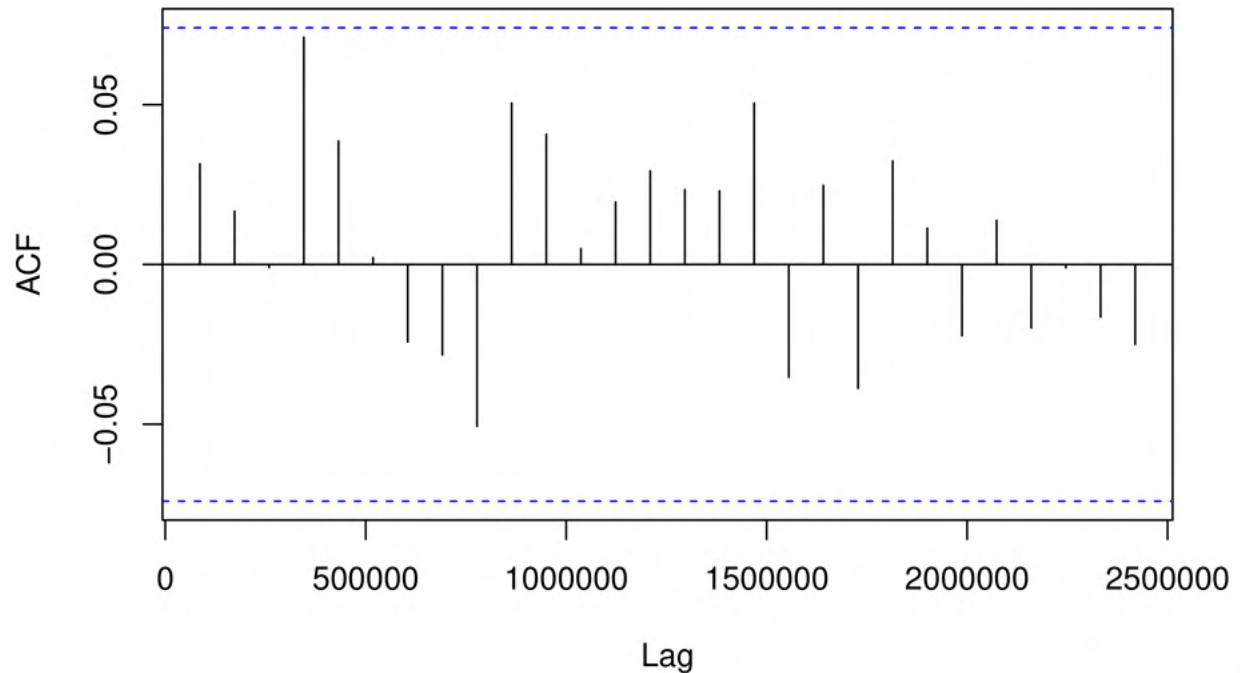
```

```
##  
##  
## Elapsed time : 0.4915371  
resid_norm = residuals(m_norm)  
qqnorm(resid_norm, main = 'QQ plot for residuals from GARCH Normal')  
qqline(resid_norm)
```



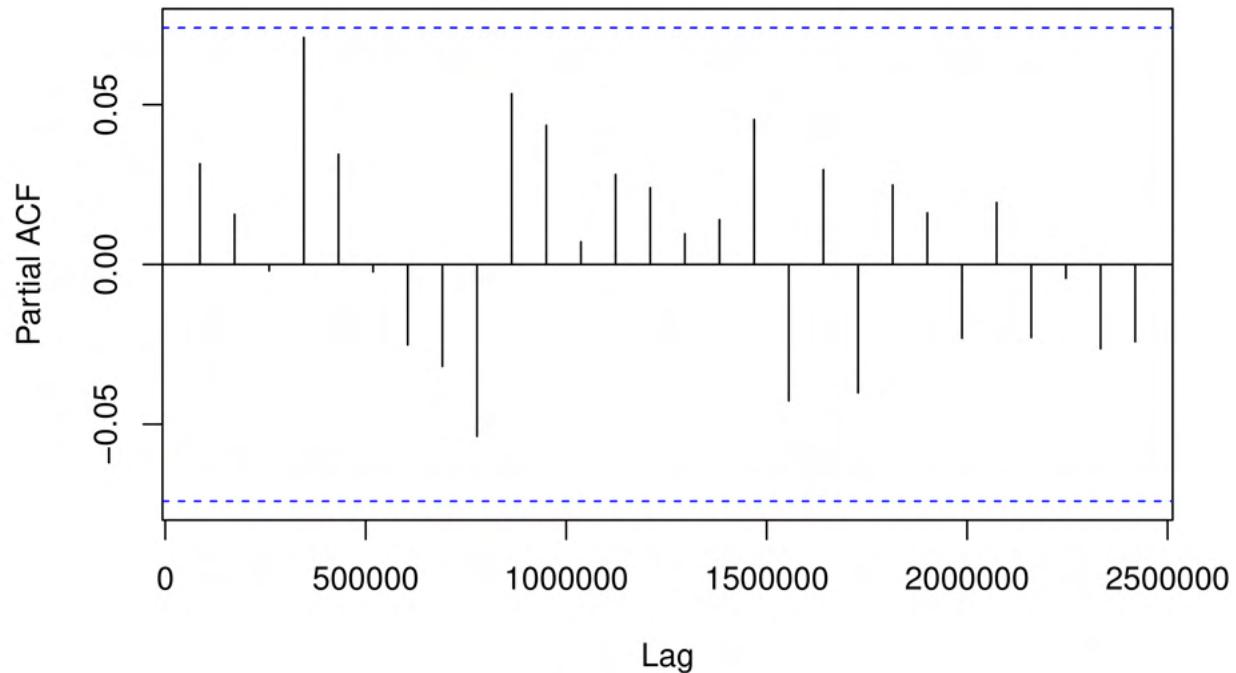
```
acf(resid_norm, main="ACf of garch residuals")
```

### ACf of garch residuals



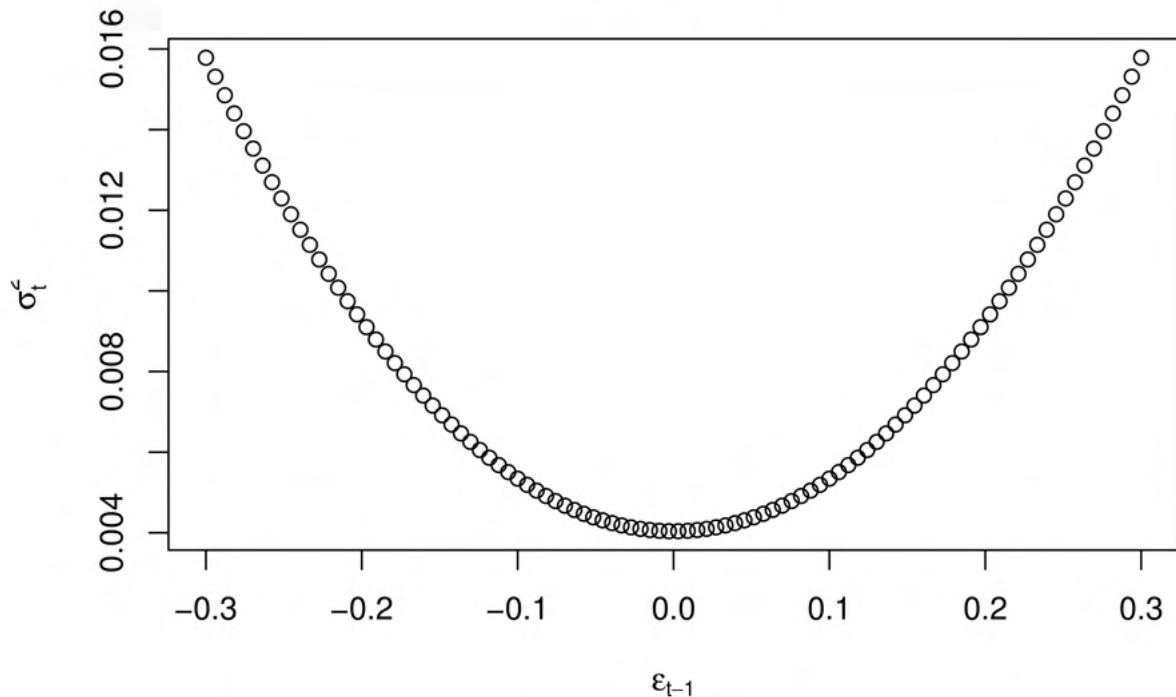
```
pacf(resid_norm, main = "PACF of garch residuals")
```

## PACF of garch residuals



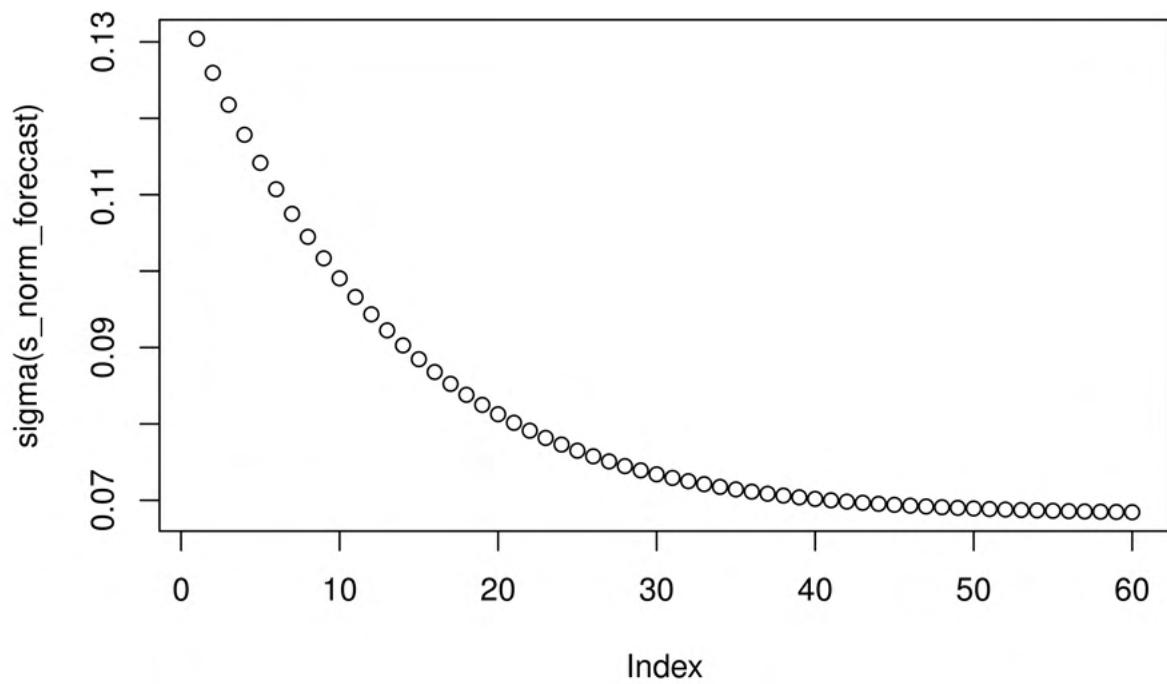
```
## Pearson goodness < 0.05: Normal is not a good model then  
## impact of news:  
## lets see whether the model is symmetric or not  
  
news_norm = newsimpact(m_norm)  
plot(news_norm$zx, news_norm$zy, ylab = news_norm$yexpr, xlab = news_norm$xexpr, main='News Impact')
```

## News Impact



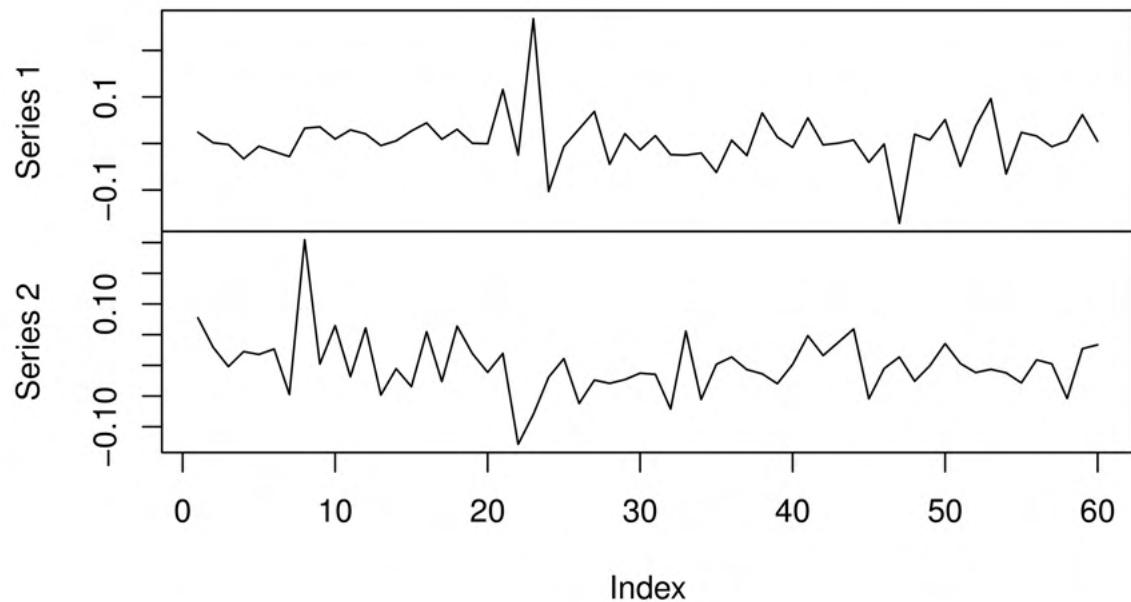
```
## symmetric Curve: positive and negative news have the same impact
```

```
## Volatility Forecast for next 60 days
sfinal = s_norm
setfixed(sfinal) = as.list(coef(m_norm))
s_norm_forecast = ugarchforecast(data = return, n.ahead = 60, fit0Rspec = sfinal)
plot(sigma(s_norm_forecast))
```



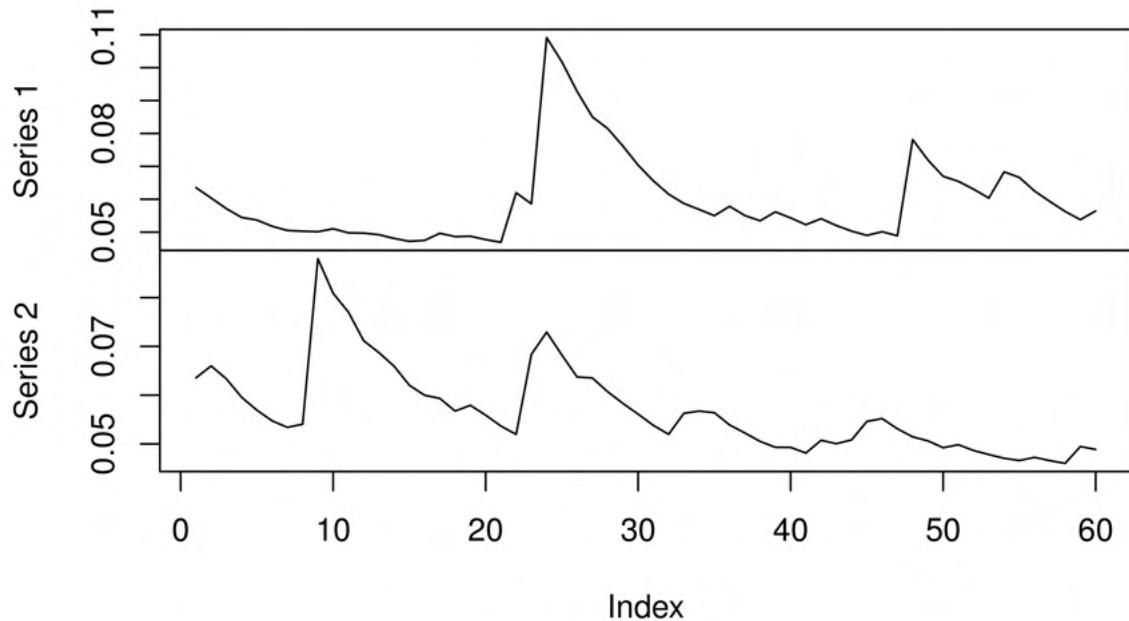
```
sim <- ugarchpath(spec = sfinal, m.sim = 2, n.sim = 1*60, rseed = 123)
plot.zoo(fitted(sim))
```

**fitted(sim)**



```
plot.zoo(sigma(sim))
```

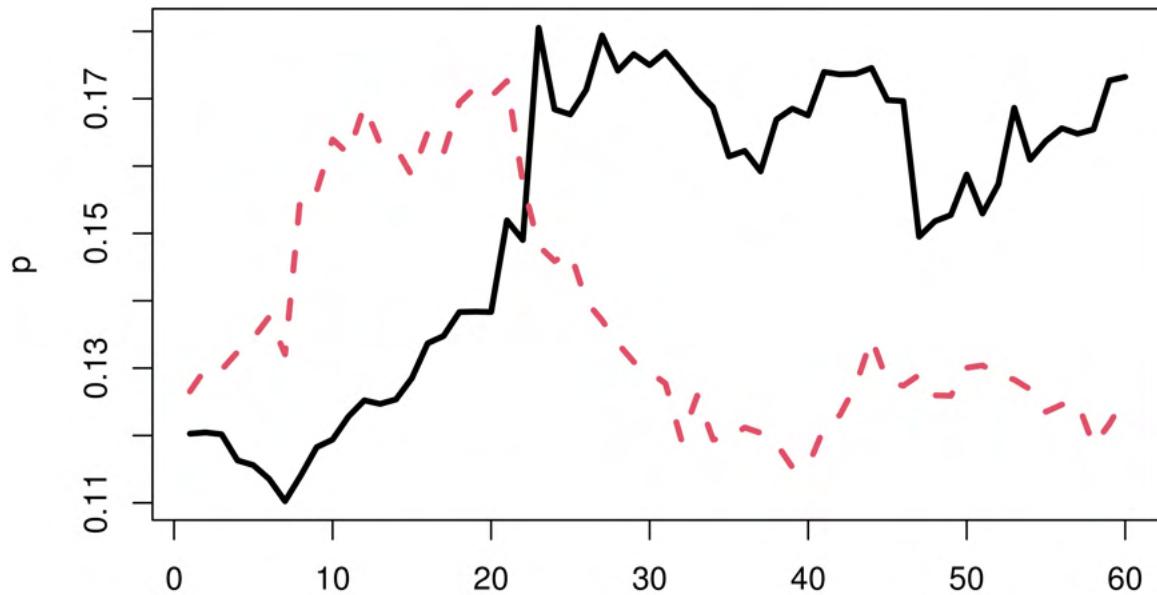
### **sigma(sim)**



```
print(tail(df$Close))

## [1] 0.1132921 0.1152852 0.1068243 0.1094526 0.1167117 0.1174374
last_value = tail(df$Close,1)

p <- last_value*apply(fitted(sim), 2, 'cumsum') + last_value
matplot(p, type = "l", lwd = 3)
```



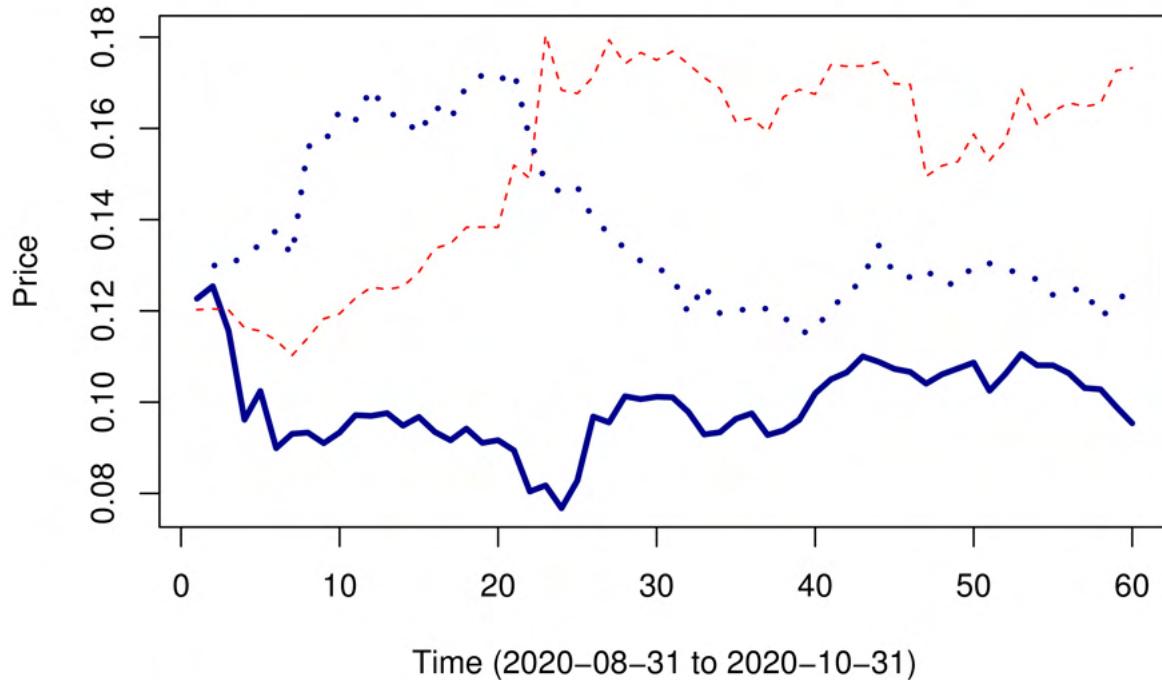
```

matplot(cbind(df2$Close, p), type = "l", lwd = c(3, 1), col = c("darkblue", "red"),
       xlab = paste("Time (", end_date, " to ", as.Date("2020-10-31"), ")",
       ylab = "Price", main = "Actual Prices vs Simulated Price Paths with Simple GARCH Model")

## Warning in cbind(df2$Close, p): number of rows of result is not a multiple of
## vector length (arg 1)

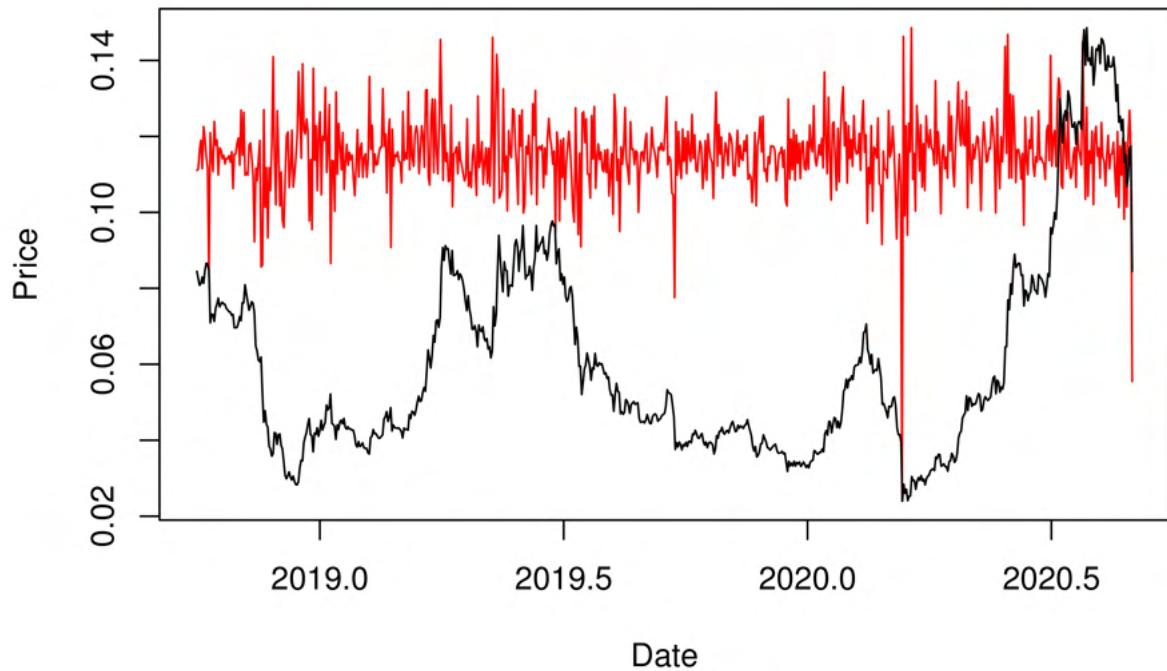
```

## Actual Prices vs Simulated Price Paths with Simple GARCH Model



```
#####
## Leverage effect is there: so apply the E-Garch Model
## gamma is the leverage variable
## if gamma positive: good news decreases the volatility

plot(return, type="l", col="red", xlab="", ylab="", axes=FALSE)
par(new=TRUE)
plot(df3$Close, type="l", xlab="Date", ylab="Price")
```



```
#####
s_norm <- ugarchspec(mean.model = list(armaOrder = c(3,2)),variance.model = list(model = 'eGARCH',garch
                           distribution.model = 'sstd')

m_norm <- ugarchfit(data = return, spec = s_norm, solver ='hybrid')
m_norm

##
## *-----*
## *      GARCH Model Fit      *
## *-----*
##
## Conditional Variance Dynamics
## -----
## GARCH Model   : eGARCH(1,1)
## Mean Model    : ARFIMA(3,0,2)
## Distribution  : sstd
##
## Optimal Parameters
## -----
##           Estimate Std. Error   t value Pr(>|t|)
## mu     -0.000228   0.001046  -0.21826 0.827226
## ar1     0.384556   0.048580   7.91594 0.000000
## ar2     0.442573   0.068318   6.47816 0.000000
## ar3     0.030571   0.026709   1.14457 0.252386
## ma1    -0.510891   0.051054 -10.00691 0.000000
```

```

## ma2      -0.338453   0.062611  -5.40569 0.000000
## omega    -0.623208   0.315055  -1.97810 0.047918
## alpha1    0.016514   0.041401   0.39888 0.689982
## beta1     0.890580   0.054442  16.35817 0.000000
## gamma1    0.240069   0.077666   3.09104 0.001995
## skew      0.974663   0.041541  23.46282 0.000000
## shape     3.061490   0.422240   7.25059 0.000000
##
## Robust Standard Errors:
##           Estimate Std. Error t value Pr(>|t|)
## mu      -0.000228   0.000619 -0.36862 0.712409
## ar1      0.384556   0.014502 26.51675 0.000000
## ar2      0.442573   0.040427 10.94755 0.000000
## ar3      0.030571   0.019843  1.54061 0.123413
## ma1     -0.510891   0.019157 -26.66898 0.000000
## ma2     -0.338453   0.020096 -16.84152 0.000000
## omega    -0.623208   0.240774 -2.58836 0.009643
## alpha1    0.016514   0.035542  0.46462 0.642202
## beta1     0.890580   0.041139 21.64822 0.000000
## gamma1    0.240069   0.061985  3.87302 0.000107
## skew      0.974663   0.036449  26.74072 0.000000
## shape     3.061490   0.409005  7.48520 0.000000
##
## LogLikelihood : 1146.073
##
## Information Criteria
## -----
## 
## Akaike      -3.2402
## Bayes       -3.1622
## Shibata     -3.2408
## Hannan-Quinn -3.2101
## 
## Weighted Ljung-Box Test on Standardized Residuals
## -----
##                      statistic p-value
## Lag[1]                  2.624  0.1053
## Lag[2*(p+q)+(p+q)-1][14] 6.793  0.8840
## Lag[4*(p+q)+(p+q)-1][24] 10.382 0.7818
## d.o.f=5
## H0 : No serial correlation
## 
## Weighted Ljung-Box Test on Standardized Squared Residuals
## -----
##                      statistic p-value
## Lag[1]                  0.07946 0.7780
## Lag[2*(p+q)+(p+q)-1][5] 0.94329 0.8724
## Lag[4*(p+q)+(p+q)-1][9] 1.82823 0.9245
## d.o.f=2
## 
## Weighted ARCH LM Tests
## -----
##          Statistic Shape Scale P-Value
## ARCH Lag[3] 0.09212 0.500 2.000 0.7615

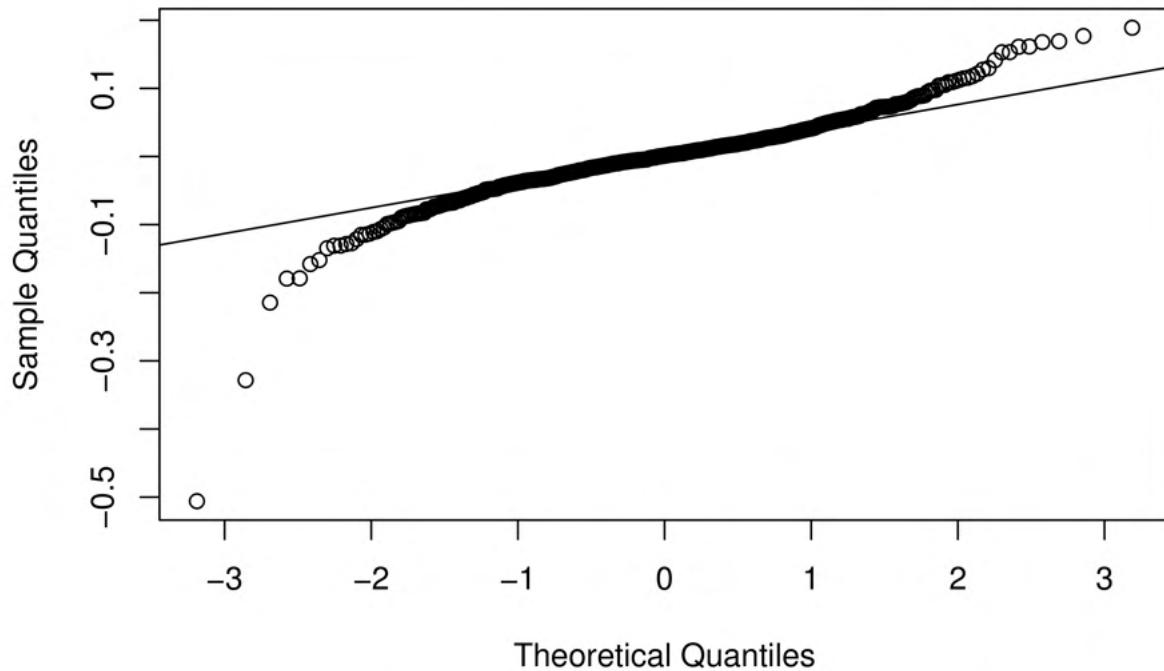
```

```

## ARCH Lag[5]    1.45430 1.440 1.667  0.6044
## ARCH Lag[7]    1.88086 2.315 1.543  0.7425
##
## Nyblom stability test
## -----
## Joint Statistic:  1.6802
## Individual Statistics:
## mu      0.22876
## ar1     0.11357
## ar2     0.07487
## ar3     0.09899
## ma1     0.11306
## ma2     0.08123
## omega   0.08143
## alpha1   0.05822
## beta1   0.07974
## gamma1  0.17912
## skew    0.22231
## shape   0.09010
##
## Asymptotic Critical Values (10% 5% 1%)
## Joint Statistic:          2.69 2.96 3.51
## Individual Statistic:     0.35 0.47 0.75
##
## Sign Bias Test
## -----
##                  t-value  prob sig
## Sign Bias        0.03231 0.9742
## Negative Sign Bias 0.17505 0.8611
## Positive Sign Bias 0.70845 0.4789
## Joint Effect      0.96946 0.8086
##
## 
## Adjusted Pearson Goodness-of-Fit Test:
## -----
## group statistic p-value(g-1)
## 1    20       16.23      0.6420
## 2    30       35.60      0.1855
## 3    40       45.14      0.2307
## 4    50       37.43      0.8863
##
## 
## Elapsed time : 0.620018
resid_norm = residuals(m_norm)
qqnorm(resid_norm, main = 'QQ plot for residuals from GARCH Normal')
qqline(resid_norm)

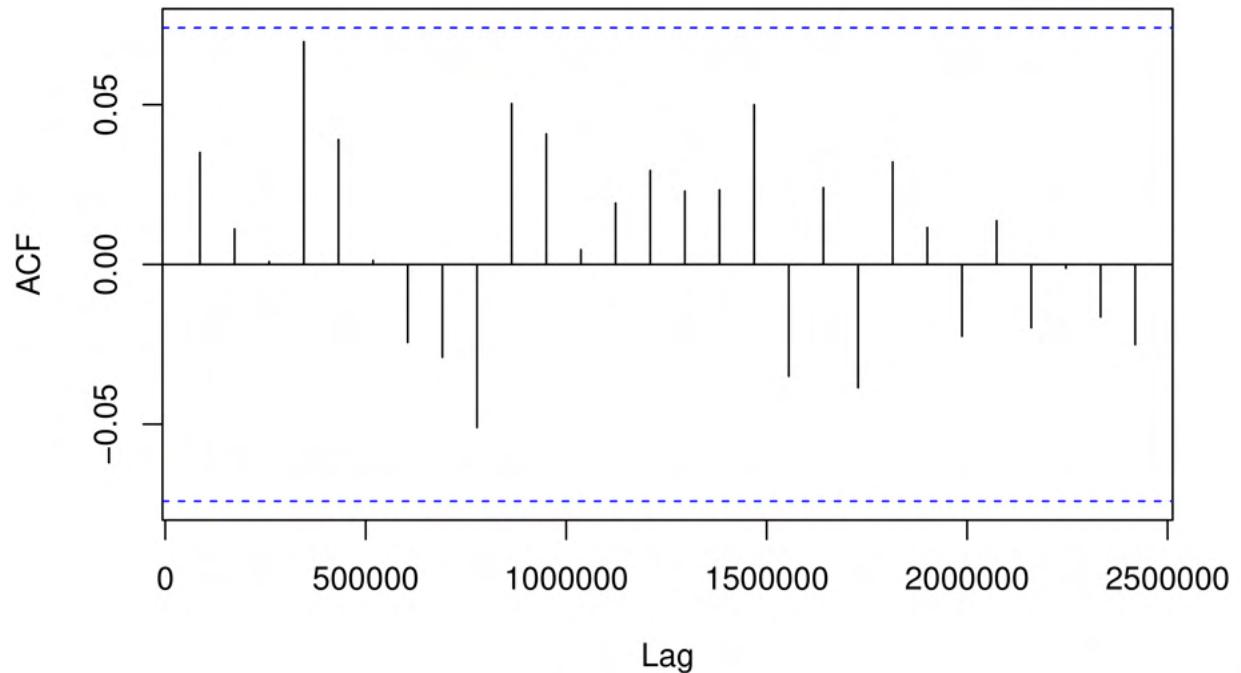
```

**QQ plot for residuals from GARCH Normal**



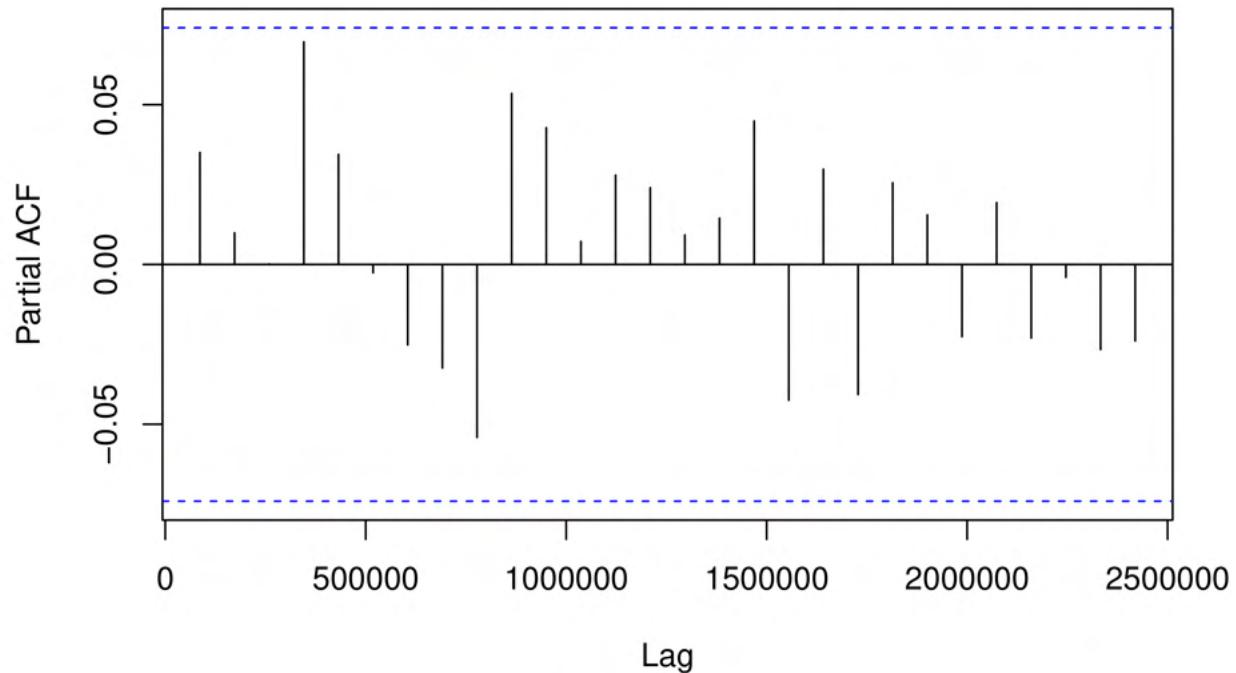
```
acf(resid_norm, main="ACf of garch residuals")
```

### ACf of garch residuals



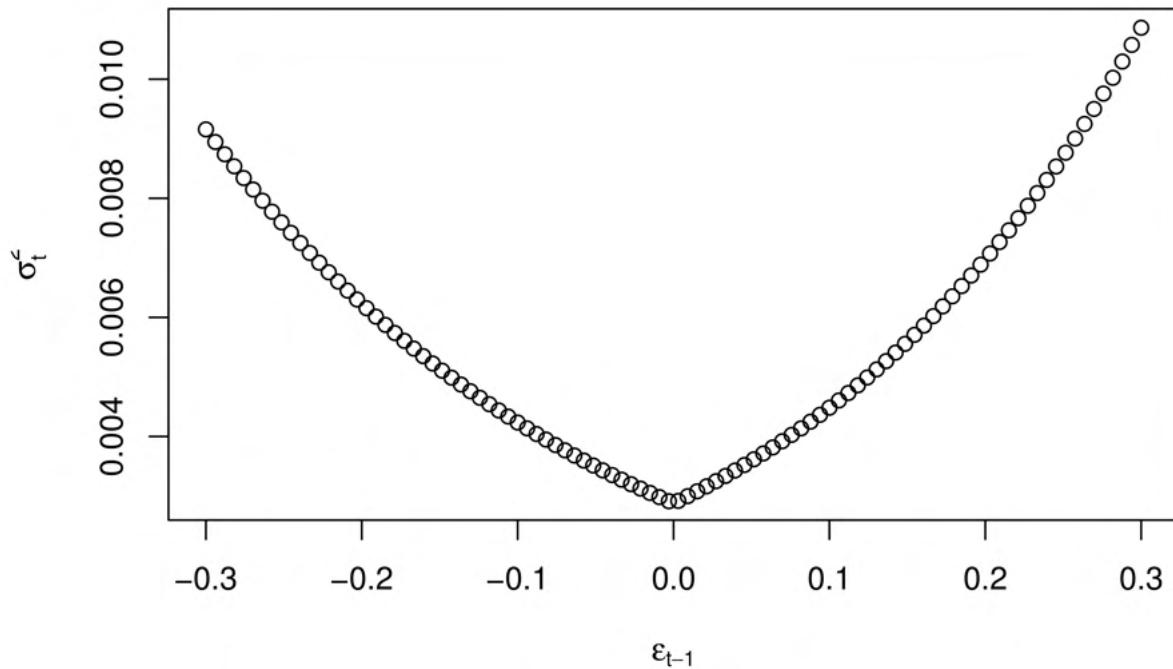
```
pacf(resid_norm, main = "PACF of garch residuals")
```

## PACF of garch residuals



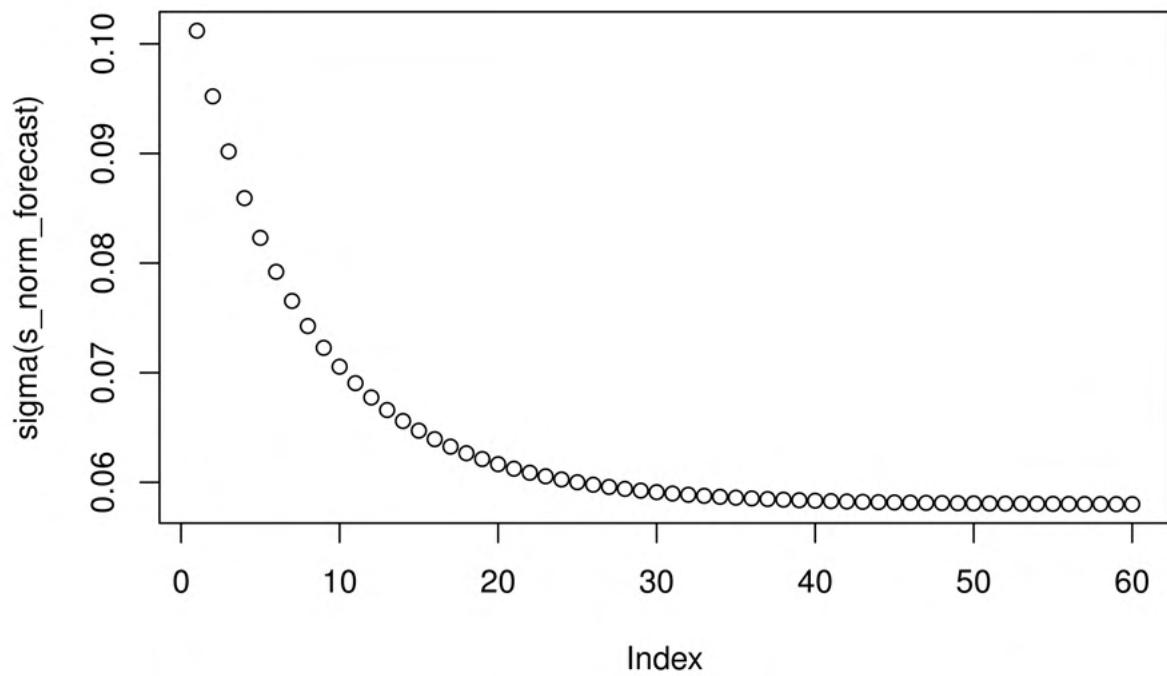
```
## Pearson goodness < 0.05: Normal is not a good model then  
## impact of news:  
## lets see whether the model is symmetric or not  
  
news_norm = newsimpact(m_norm)  
plot(news_norm$zx, news_norm$zy, ylab = news_norm$yexpr, xlab = news_norm$xexpr, main='News Impact')
```

## News Impact



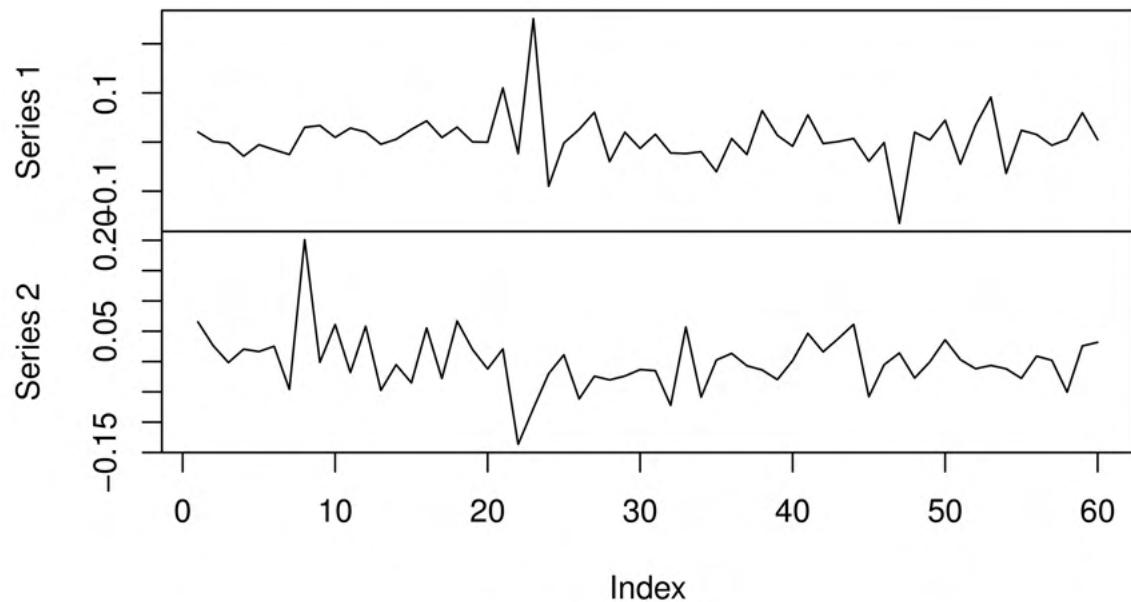
```
## symmetric Curve: positive and negative news have the same impact

## Volatility Forecast for next 20 days
sfinal = s_norm
setfixed(sfinal) = as.list(coef(m_norm))
s_norm_forecast = ugarchforecast(data = return, n.ahead = 60, fitOrSpec = sfinal)
plot(sigma(s_norm_forecast))
```



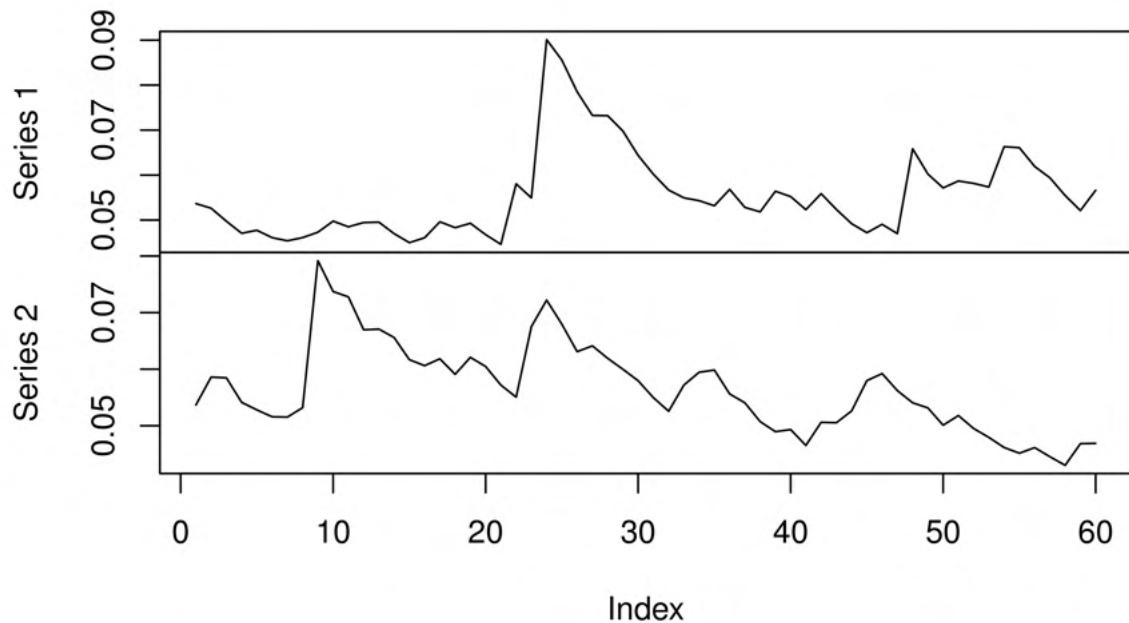
```
sim <- ugarchpath(spec = sfinal, m.sim = 2, n.sim = 1*60, rseed = 123)
plot.zoo(fitted(sim))
```

**fitted(sim)**



```
plot.zoo(sigma(sim))
```

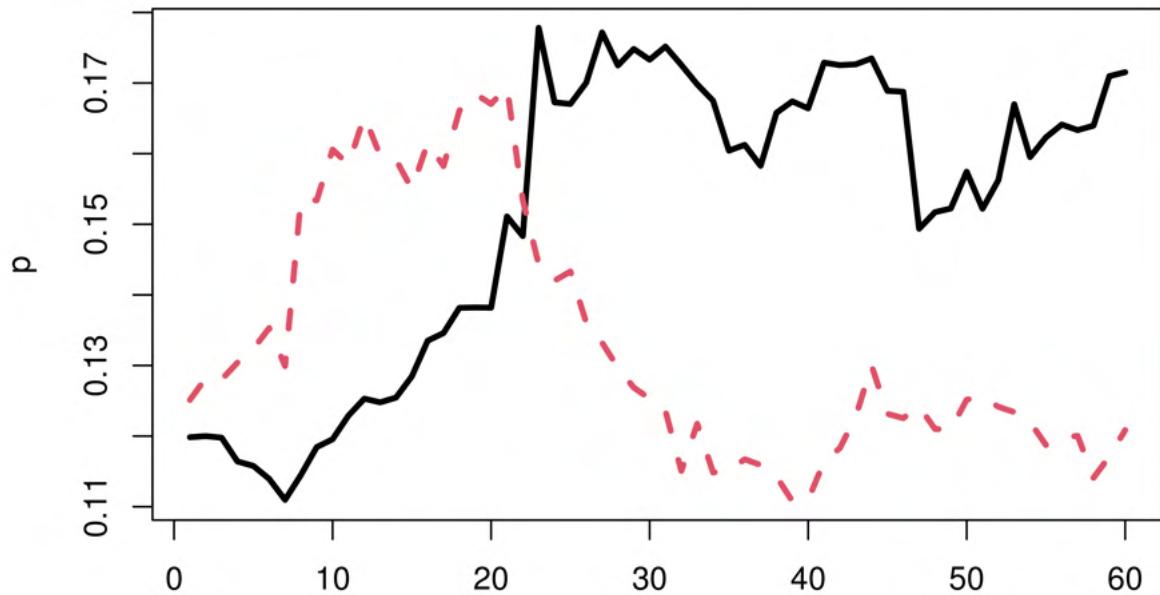
### **sigma(sim)**



```
print(tail(df$Close))

## [1] 0.1132921 0.1152852 0.1068243 0.1094526 0.1167117 0.1174374
last_value = tail(df$Close,1)

p <- last_value*apply(fitted(sim), 2, 'cumsum') + last_value
matplot(p, type = "l", lwd = 3)
```



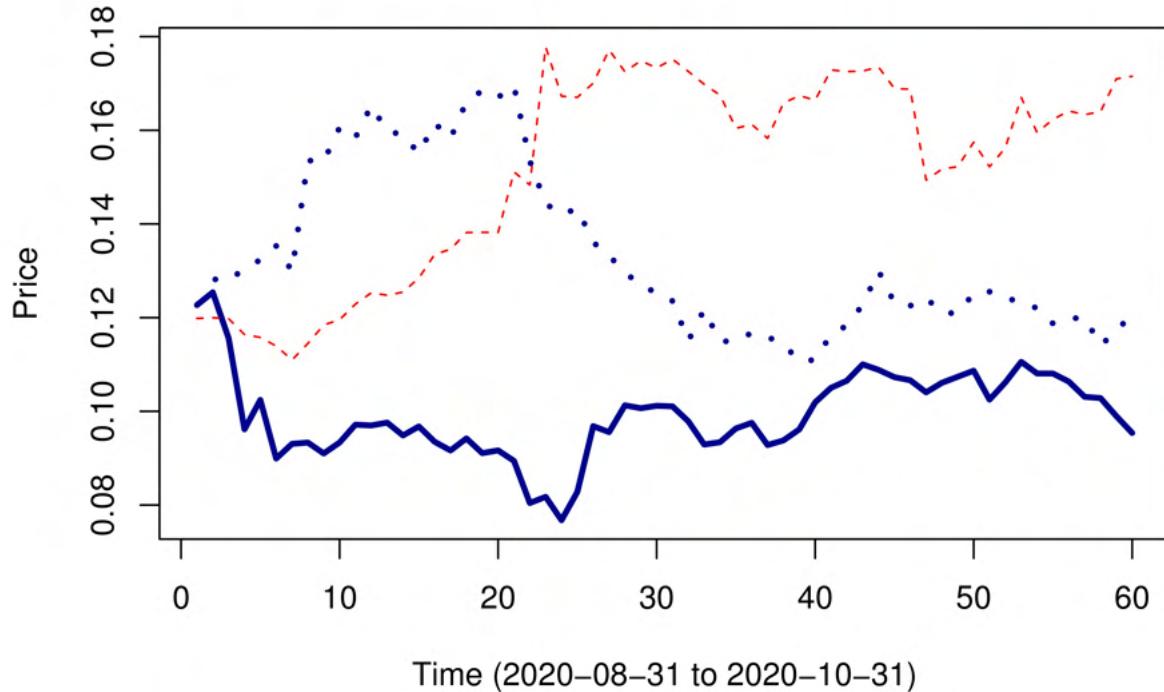
```

matplot(cbind(df2$Close, p), type = "l", lwd = c(3, 1), col = c("darkblue", "red"),
       xlab = paste("Time (", end_date, " to ", as.Date("2020-10-31"), ")"),
       ylab = "Price", main = "Actual Prices vs Simulated Price Paths considering Leverage Effect")

## Warning in cbind(df2$Close, p): number of rows of result is not a multiple of
## vector length (arg 1)

```

## Actual Prices vs Simulated Price Paths considering Leverage Effect



```
#####
#####
return = (diff(log(df3$Close)))

s_norm <- ugarchspec(mean.model = list(armaOrder = c(3,2)),variance.model = list(model = 'gjrGARCH',gar
distribution.model = 'sstd')

m_norm <- ugarchfit(data = return, spec = s_norm, solver ='hybrid')
m_norm

##
## *-----*
## *      GARCH Model Fit      *
## *-----*
##
## Conditional Variance Dynamics
## -----
## GARCH Model  : gjrGARCH(1,1)
## Mean Model   : ARFIMA(3,0,2)
## Distribution : sstd
##
## Optimal Parameters
## -----
##           Estimate Std. Error t value Pr(>|t|)
## mu     -0.000045   0.002032 -0.022337 0.982179
```

```

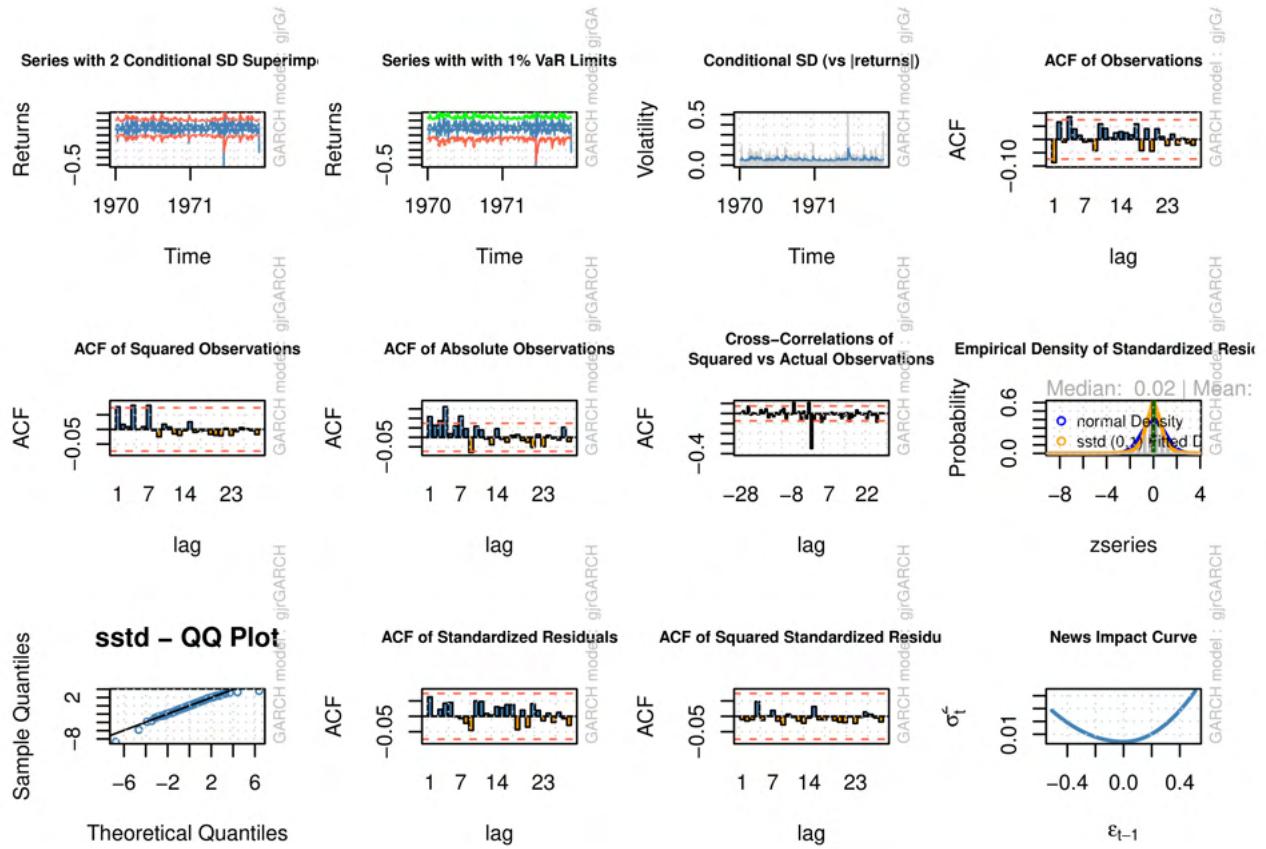
## ar1      0.375339   0.394653   0.951062  0.341573
## ar2      0.447581   0.298878   1.497538  0.134253
## ar3      0.034598   0.064820   0.533758  0.593509
## ma1     -0.502071   0.393382  -1.276294  0.201852
## ma2     -0.347480   0.348737  -0.996396  0.319058
## omega    0.000405   0.000186   2.177804  0.029421
## alpha1   0.152348   0.070091   2.173580  0.029737
## beta1    0.790889   0.069959  11.305024  0.000000
## gamma1  -0.058789   0.067396  -0.872291  0.383050
## skew     0.975428   0.049405  19.743454  0.000000
## shape    3.030699   0.413115   7.336205  0.000000
##
## Robust Standard Errors:
##           Estimate Std. Error t value Pr(>|t|)
## mu      -0.000045   0.002240 -0.020257 0.983838
## ar1      0.375339   0.107005  3.507693 0.000452
## ar2      0.447581   0.102094  4.384007 0.000012
## ar3      0.034598   0.030061  1.150918 0.249766
## ma1     -0.502071   0.105558 -4.756359 0.000002
## ma2     -0.347480   0.106885 -3.250986 0.001150
## omega    0.000405   0.000129  3.133640 0.001727
## alpha1   0.152348   0.052214  2.917748 0.003526
## beta1    0.790889   0.049349 16.026514 0.000000
## gamma1  -0.058789   0.063596 -0.924414 0.355271
## skew     0.975428   0.052148 18.704971 0.000000
## shape    3.030699   0.401407  7.550184 0.000000
##
## LogLikelihood : 1145.817
##
## Information Criteria
## -----
## 
## Akaike       -3.2395
## Bayes        -3.1615
## Shibata     -3.2401
## Hannan-Quinn -3.2093
## 
## Weighted Ljung-Box Test on Standardized Residuals
## -----
##                      statistic p-value
## Lag[1]                  2.778 0.09555
## Lag[2*(p+q)+(p+q)-1][14] 7.079 0.75550
## Lag[4*(p+q)+(p+q)-1][24] 10.918 0.70988
## d.o.f=5
## H0 : No serial correlation
## 
## Weighted Ljung-Box Test on Standardized Squared Residuals
## -----
##                      statistic p-value
## Lag[1]                  0.07931 0.7782
## Lag[2*(p+q)+(p+q)-1][5] 0.93261 0.8747
## Lag[4*(p+q)+(p+q)-1][9] 1.61975 0.9452
## d.o.f=2
## 
```

```

## Weighted ARCH LM Tests
## -----
##           Statistic Shape Scale P-Value
## ARCH Lag[3]    0.08496 0.500 2.000  0.7707
## ARCH Lag[5]    1.44017 1.440 1.667  0.6082
## ARCH Lag[7]    1.65548 2.315 1.543  0.7897
##
## Nyblom stability test
## -----
## Joint Statistic: 1.8188
## Individual Statistics:
## mu      0.25568
## ar1     0.14087
## ar2     0.07705
## ar3     0.11093
## ma1     0.13824
## ma2     0.08519
## omega   0.08382
## alpha1   0.18502
## beta1   0.10255
## gamma1  0.10182
## skew    0.22207
## shape   0.10426
##
## Asymptotic Critical Values (10% 5% 1%)
## Joint Statistic:        2.69 2.96 3.51
## Individual Statistic:   0.35 0.47 0.75
##
## Sign Bias Test
## -----
##           t-value  prob sig
## Sign Bias       0.03581 0.9714
## Negative Sign Bias 0.27053 0.7868
## Positive Sign Bias 0.71192 0.4768
## Joint Effect     0.97928 0.8063
##
## 
## Adjusted Pearson Goodness-of-Fit Test:
## -----
## group statistic p-value(g-1)
## 1      20      13.43      0.8159
## 2      30      34.83      0.2102
## 3      40      33.26      0.7286
## 4      50      40.14      0.8124
##
## 
## Elapsed time : 1.428219
plot(m_norm, which='all')

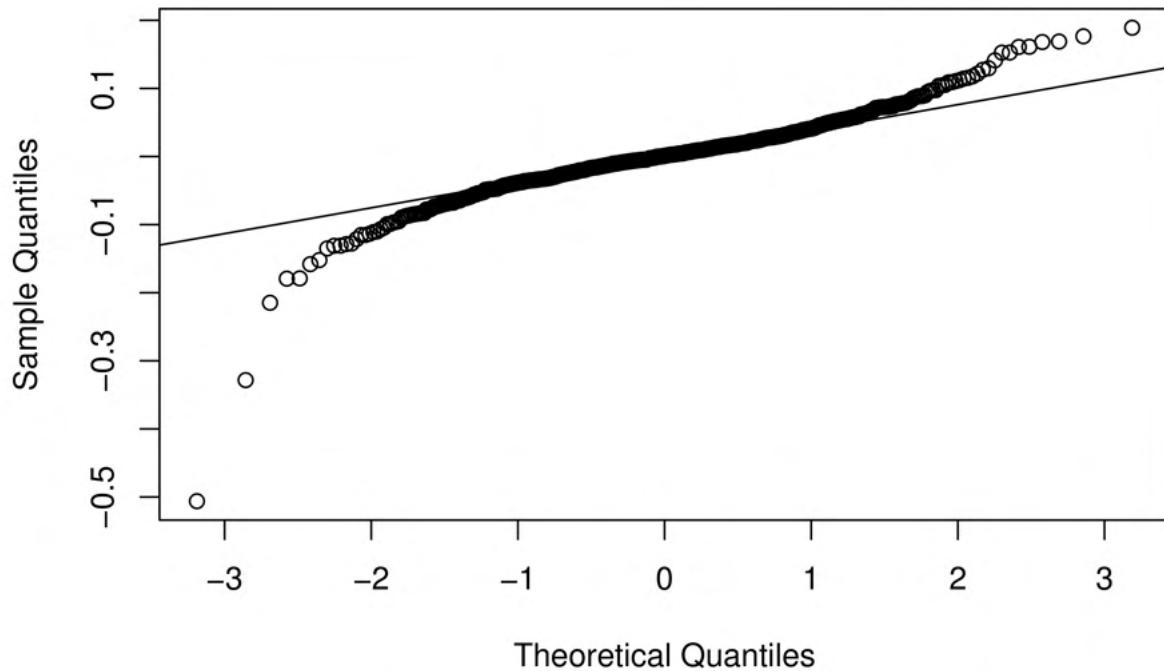
##
## please wait...calculating quantiles...

```



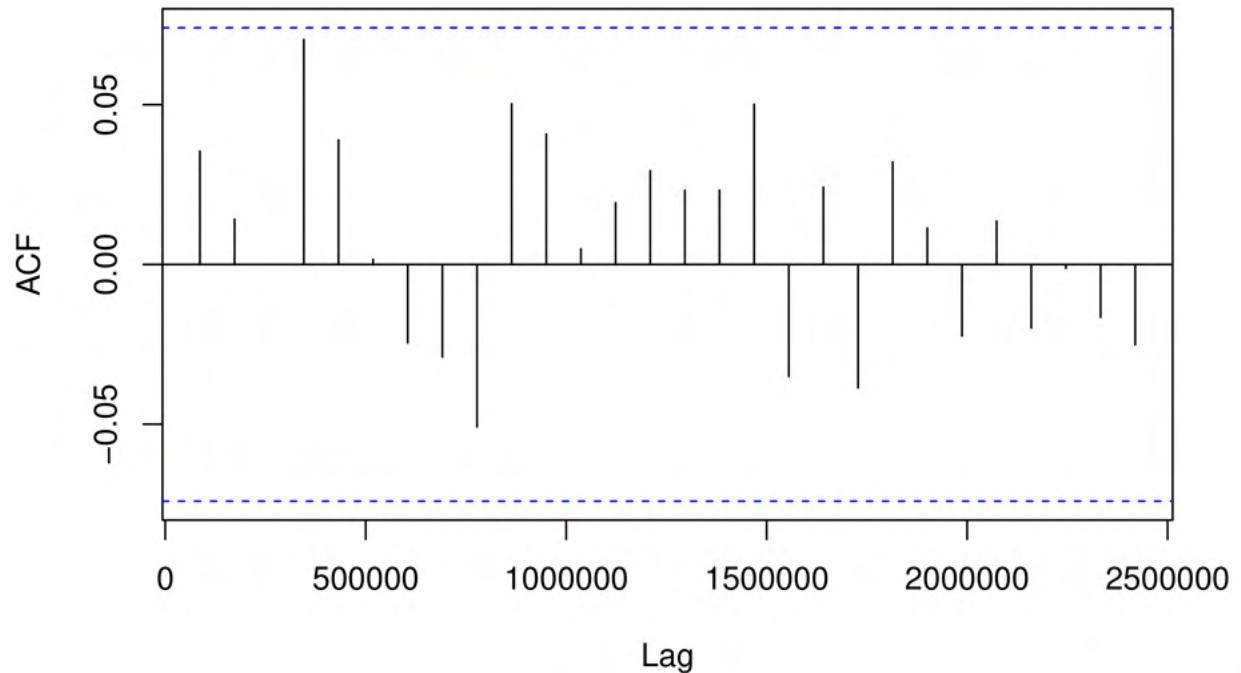
```
resid_norm = residuals(m_norm)
qqnorm(resid_norm, main = 'QQ plot for residuals from ggjGARCH Normal')
qqline(resid_norm)
```

### QQ plot for residuals from ggjGARCH Normal



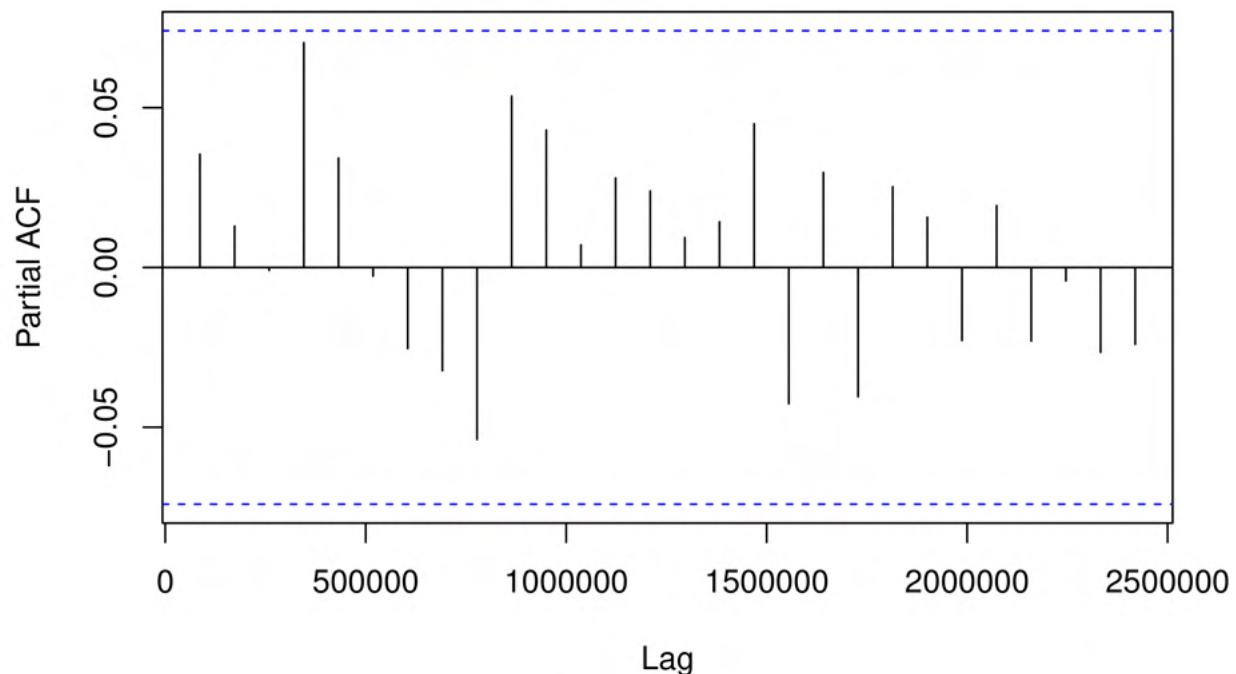
```
acf(resid_norm, main="ACF of ggjgarch residuals")
```

### ACF of ggjgarch residuals



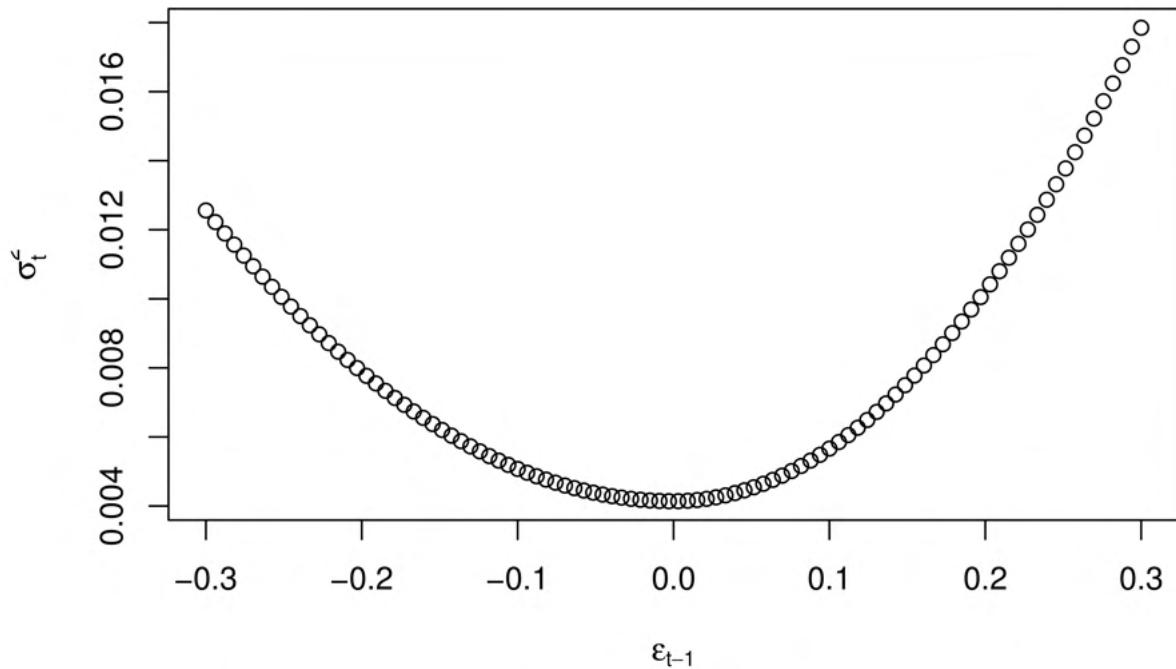
```
pacf(resid_norm, main = "PACF of ggjgarch residuals")
```

## PACF of ggjgarch residuals



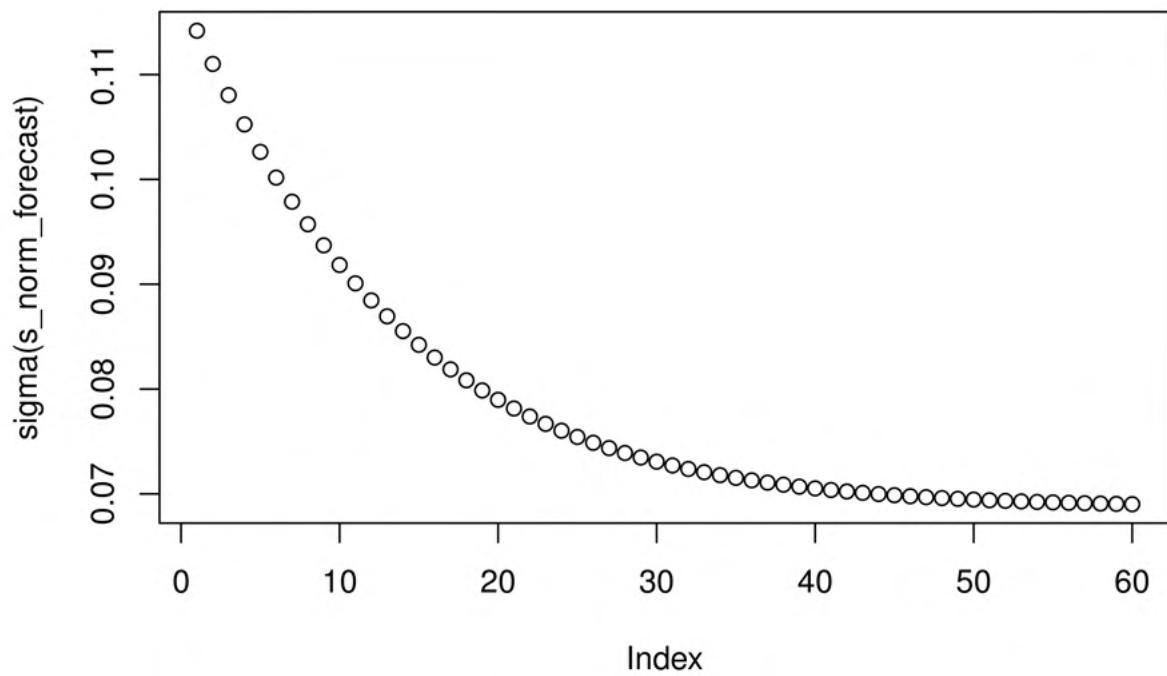
```
## Pearson goodness < 0.05: Normal is not a good model then  
##  
## impact of news:  
## lets see whether the model is symmetric or not  
  
news_norm = newsimpact(m_norm)  
plot(news_norm$zx, news_norm$zy, ylab = news_norm$yexpr, xlab = news_norm$xexpr, main='News Impact')
```

## News Impact



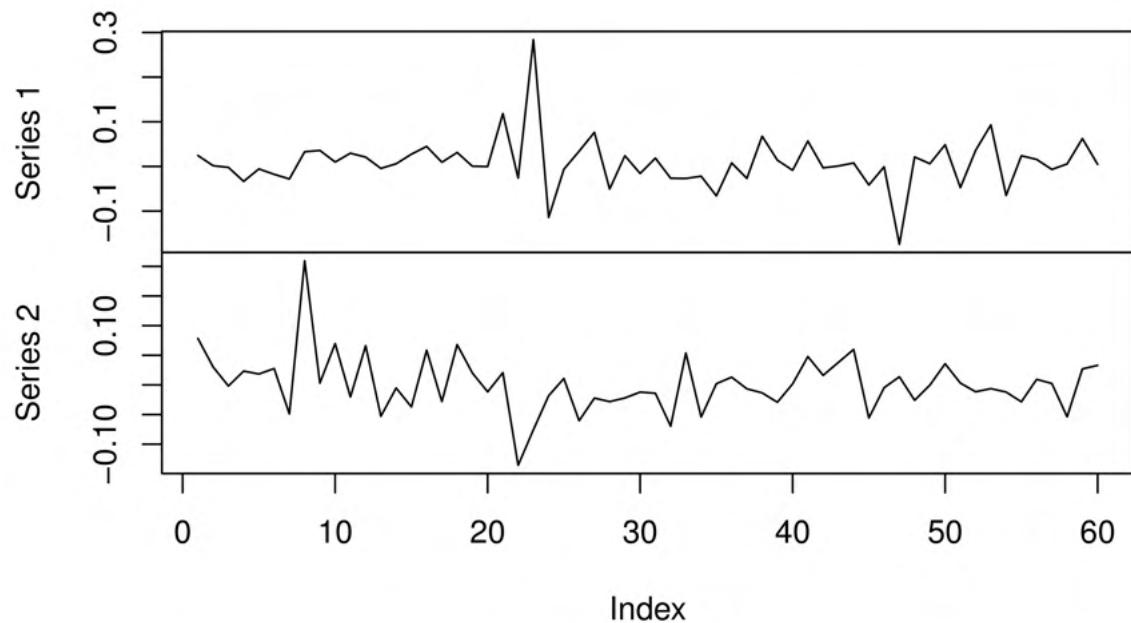
```
## symmetric Curve: positive and negative news have the same impact

## Volatility Forecast for next 20 days
sfinal = s_norm
setfixed(sfinal) = as.list(coef(m_norm))
s_norm_forecast = ugarchforecast(data = return, n.ahead = 60, fitOrSpec = sfinal)
plot(sigma(s_norm_forecast))
```



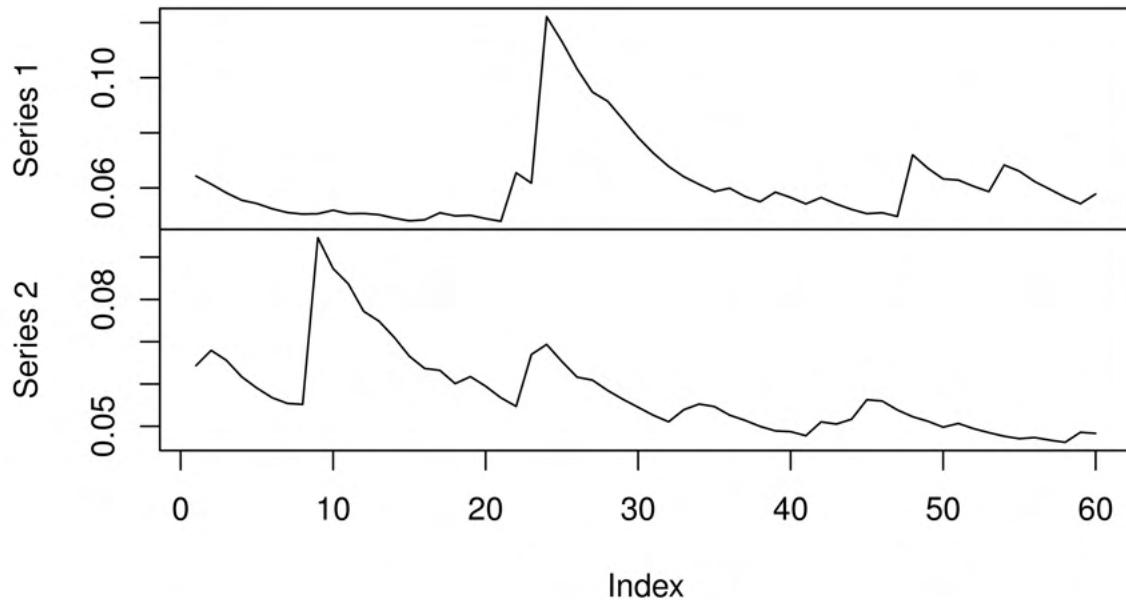
```
sim <- ugarchpath(spec = sfinal, m.sim = 2, n.sim = 1*60, rseed = 123)
plot.zoo(fitted(sim))
```

**fitted(sim)**



```
plot.zoo(sigma(sim))
```

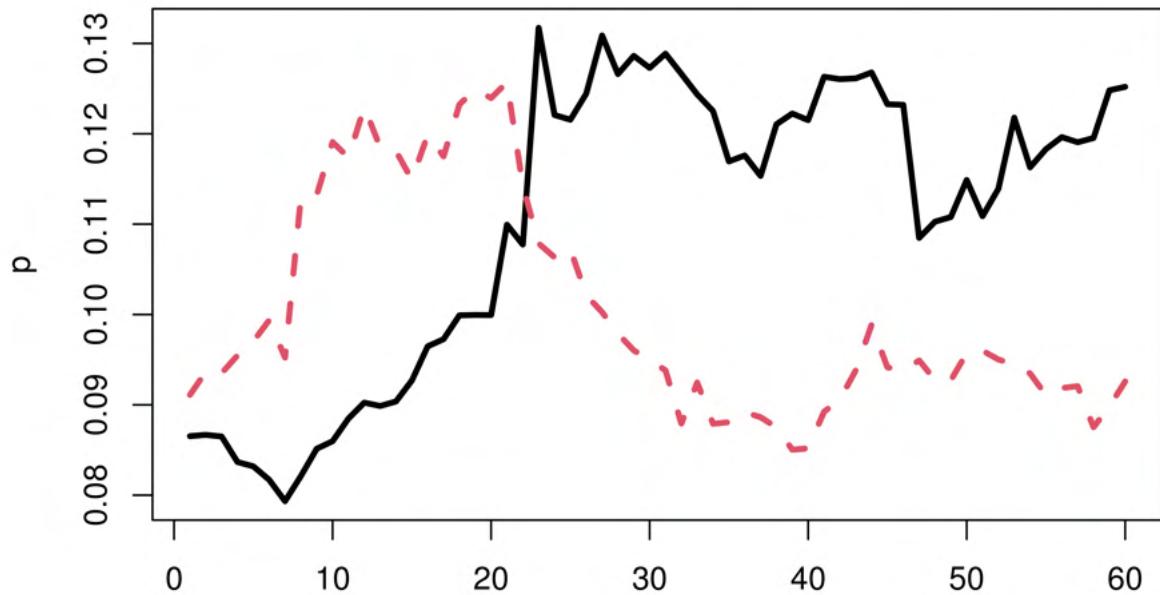
## sigma(sim)



```
print(tail(df3$Close))

## Time Series:
## Start = c(2020, 239)
## End = c(2020, 244)
## Frequency = 365
## [1] 0.1152852 0.1068243 0.1094526 0.1167117 0.1174374 0.0844693
last_value = tail(df3$Close,1)[1]
last_value

## [1] 0.0844693
p <- last_value*apply(fitted(sim), 2, 'cumsum') + last_value
matplot(p, type = "l", lwd = 3)
```



```

matplot(cbind(df2$Close, p), type = "l", lwd = c(3, 1), col = c("blue", "red"),
       xlab = paste("Time (", end_date, " to ", as.Date("2020-10-31"), ")",
       ylab = "Price", main = "Actual Prices vs Simulated Price Paths using ggrGarch")

## Warning in cbind(df2$Close, p): number of rows of result is not a multiple of
## vector length (arg 1)

```

## Actual Prices vs Simulated Price Paths using ggrGarch

