

Review: Logistic regression, Gaussian naïve Bayes, linear regression, and their connections

New: Bias-variance decomposition, bias-variance tradeoff, overfitting, regularization, and feature selection

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Parts of the slides are from previous 10-701 lectures

Outline

- Logistic regression
- Decision surface (boundary) of classifiers
- Generative vs. discriminative classifiers
- Linear regression
- Bias-variance decomposition and tradeoff
- Overfitting and regularization
- Feature selection

Outline

- Logistic regression
 - Model assumptions: $P(Y|X)$
 - Decision making
 - Estimating the model parameters
 - Multiclass logistic regression
- Decision surface (boundary) of classifiers
- Generative vs. discriminative classifiers
- Linear regression
- Bias-variance decomposition and tradeoff
- Overfitting and regularization
- Feature selection

Logistic regression: assumptions

- Binary classification
 - $f: X = (X_1, X_2, \dots, X_n) \rightarrow Y \in \{0, 1\}$
- Logistic regression: assumptions on $P(Y|X)$:

$$P(Y = 0|X, w) = \frac{1}{1 + \exp(w_0 + \sum_i w_i X_i)}$$

- And thus:

$$\begin{aligned} P(Y = 1|X, w) &= \frac{\exp(w_0 + \sum_i w_i X_i)}{1 + \exp(w_0 + \sum_i w_i X_i)} \\ &\quad (= \frac{1}{1 + \exp(-w_0 - \sum_i w_i X_i)}) \end{aligned}$$

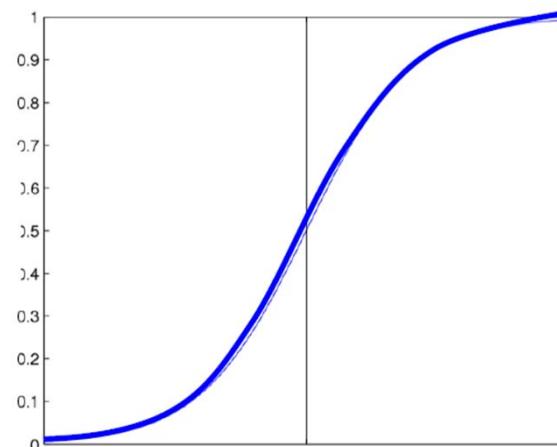
Logistic regression: assumptions

- Model assumptions: the form of $P(Y|X)$

$$P(Y = 0|X, w) = \frac{1}{1 + \exp(w_0 + \sum_i w_i X_i)}$$

- “Logistic” regression
 - $P(Y|X)$ is the **logistic function** applied to a linear function of X

$$\frac{1}{1 + \exp(-z)}$$



Decision making

- Given a logistic regression \mathbf{w} and an \mathbf{X} :

$$P(Y = 0|\mathbf{X}, \mathbf{w}) = \frac{1}{1 + \exp(w_0 + \sum_i w_i X_i)}$$

$$P(Y = 1|\mathbf{X}, \mathbf{w}) = \frac{\exp(w_0 + \sum_i w_i X_i)}{1 + \exp(w_0 + \sum_i w_i X_i)}$$

- Decision making on \mathbf{Y} :

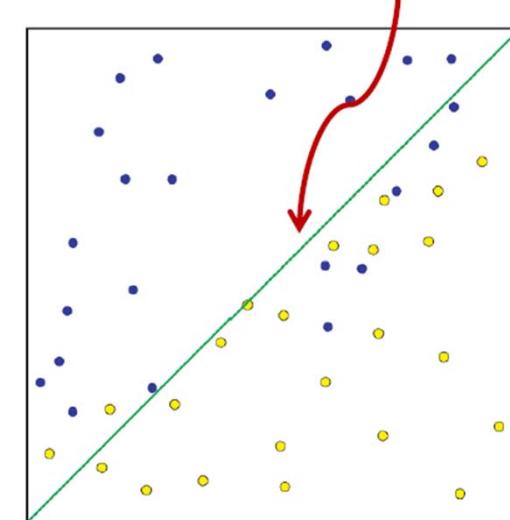
$$P(Y = 0|X) \stackrel{0}{\gtrless} P(Y = 1|X)$$

1

Linear decision boundary !

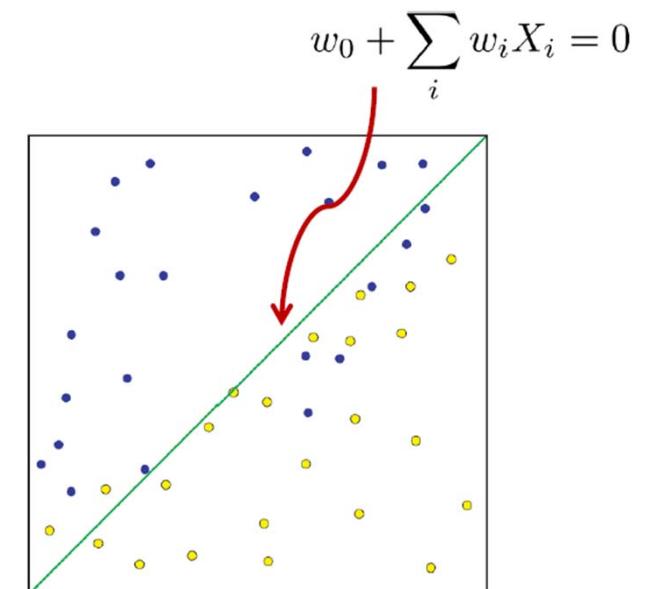
$$0 \stackrel{0}{\gtrless} w_0 + \sum_i w_i X_i$$

$$w_0 + \sum_i w_i X_i = 0$$



Estimating the parameters \mathbf{w}

- Given $\{(X^{(1)}, Y^{(1)}), \dots, (X^{(L)}, Y^{(L)})\}$
 - where $X^{(j)} = (X_1^{(j)}, X_2^{(j)}, \dots, X_n^{(j)})$
- How to **estimate** $\mathbf{w} = (w_0, w_1, \dots, w_n)$?



[Aarti, 10-701]

Estimating the parameters \mathbf{w}

- Given $\{(X^{(j)}, Y^{(j)})\}_{j=1}^L$, $X^{(j)} = (X_1^{(j)}, X_2^{(j)}, \dots, X_n^{(j)})$
- Assumptions: $P(Y|X, \mathbf{w})$
- Maximum ***conditional*** likelihood on data!
$$\hat{\mathbf{w}}_{MCLE} = \arg \max_{\mathbf{w}} \prod_{j=1}^L P(Y^{(j)} | X^{(j)}, \mathbf{w})$$
 - Logistic regression only models $P(Y|X)$
 - So we only maximize $P(Y|X)$, ***ignoring*** $P(X)$

Estimating the parameters \mathbf{w}

- Given $\{(X^{(j)}, Y^{(j)})\}_{j=1}^L$, $X^{(j)} = (X_1^{(j)}, X_2^{(j)}, \dots, X_n^{(j)})$
- Assumptions:
$$P(Y = 0 | \mathbf{X}, \mathbf{w}) = \frac{1}{1 + \exp(w_0 + \sum_i w_i X_i)}$$
$$P(Y = 1 | \mathbf{X}, \mathbf{w}) = \frac{\exp(w_0 + \sum_i w_i X_i)}{1 + \exp(w_0 + \sum_i w_i X_i)}$$

- Maximum ***conditional*** likelihood on data!

$$\hat{\mathbf{w}}_{MCLE} = \arg \max_{\mathbf{w}} \prod_{j=1}^L P(Y^{(j)} | X^{(j)}, \mathbf{w})$$

- Let's maximize conditional ***log***-likelihood

$$\max_{\mathbf{w}} l(\mathbf{w}) \equiv \ln \prod_j^L P(y^j | \mathbf{x}^j, \mathbf{w})$$

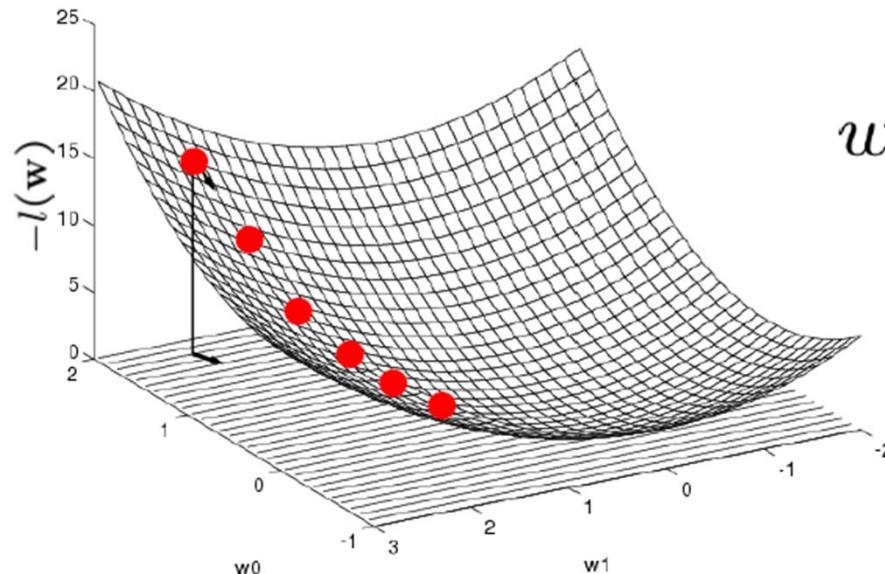
$$= \sum_j^L y^j (w_0 + \sum_i^n w_i x_i^j) - \ln(1 + \exp(w_0 + \sum_i^n w_i x_i^j))$$

Estimating the parameters \mathbf{w}

- Max conditional log-likelihood on data

$$\begin{aligned}\max_{\mathbf{w}} l(\mathbf{w}) &\equiv \ln \prod_{j=1}^L P(y^j | \mathbf{x}^j, \mathbf{w}) \\ &= \sum_{j=1}^L y^j (w_0 + \sum_{i=1}^{n_j} w_i x_i^j) - \ln(1 + \exp(w_0 + \sum_{i=1}^{n_j} w_i x_i^j))\end{aligned}$$

- A concave function (beyond the scope of class)
- No local optimum: **gradient ascent** (descent) ☺



$$w_i^{(t+1)} \leftarrow w_i^{(t)} + \eta \frac{\partial l(\mathbf{w})}{\partial w_i^{(t)}}$$

Estimating the parameters \mathbf{w}

- Max conditional log-likelihood on data

$$\begin{aligned} \max_{\mathbf{w}} l(\mathbf{w}) &\equiv \ln \prod_{j=1}^L P(y^j | \mathbf{x}^j, \mathbf{w}) \\ &= \sum_{j=1}^L y^j (w_0 + \sum_{i=1}^{n_j} w_i x_i^j) - \ln(1 + \exp(w_0 + \sum_{i=1}^{n_j} w_i x_i^j)) \end{aligned}$$

- A concave function (beyond the scope of class)
- No local optimum: **gradient ascent** (descent) ☺

$$w_i^{(t+1)} \leftarrow w_i^{(t)} + \eta \frac{\partial l(\mathbf{w})}{\partial w_i^{(t)}}$$

$$w_i^{(t+1)} \leftarrow w_i^{(t)} + \eta \sum_{j=1}^L x_i^j [y^j - \hat{P}(Y^j = 1 | \mathbf{x}^j, \mathbf{w}^{(t)})]$$

Multiclass logistic regression

- **Binary classification**

$$P(Y = 0 | \mathbf{X}, \mathbf{w}) = \frac{1}{1 + \exp(w_0 + \sum_i w_i X_i)}$$

$$P(Y = 1 | \mathbf{X}, \mathbf{w}) = \frac{\exp(w_0 + \sum_i w_i X_i)}{1 + \exp(w_0 + \sum_i w_i X_i)}$$

- **K-class classification**

- For each class $k < K$

$$P(Y = y_k | \mathbf{X}) = \frac{\exp(w_{k0} + \sum_{i=1}^d w_{ki} X_i)}{1 + \sum_{j=1}^{K-1} \exp(w_{j0} + \sum_{i=1}^d w_{ji} X_i)}$$

- For class K

$$P(Y = y_K | \mathbf{X}) = \frac{1}{1 + \sum_{j=1}^{K-1} \exp(w_{j0} + \sum_{i=1}^d w_{ji} X_i)}$$

Outline

- Logistic regression
- Decision surface (boundary) of classifiers
 - Logistic regression
 - Gaussian naïve Bayes
 - Decision trees
- Generative vs. discriminative classifiers
- Linear regression
- Bias-variance decomposition and tradeoff
- Overfitting and regularization
- Feature selection

Logistic regression

- Model assumptions on $P(Y|X)$:

$$P(Y = 0|X, w) = \frac{1}{1 + \exp(w_0 + \sum_i w_i X_i)}$$

$$P(Y = 1|X, w) = \frac{\exp(w_0 + \sum_i w_i X_i)}{1 + \exp(w_0 + \sum_i w_i X_i)}$$

- Deciding Y given X :

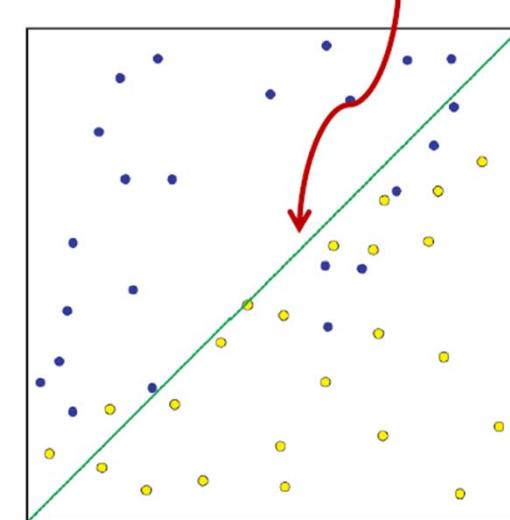
$$P(Y = 0|X) \stackrel{0}{\gtrless} P(Y = 1|X)$$

1

Linear decision
boundary!

$$0 \stackrel{0}{\gtrless} w_0 + \sum_i w_i X_i$$

$$w_0 + \sum_i w_i X_i = 0$$



Gaussian naïve Bayes

- Model assumptions $P(\mathbf{X}, \mathbf{Y}) = P(\mathbf{Y})P(\mathbf{X}|\mathbf{Y})$
 - Bernoulli on \mathbf{Y} : $P(Y = 1) = \pi$
 - Conditional independence of \mathbf{X}
$$P(X = (X_1, X_2, \dots, X_n)|Y = k) = \prod_{i=1}^n P(X_i|Y = k)$$
 - Gaussian for X_i given \mathbf{Y} : $P(X_i|Y = k) \sim N(\mu_{ik}, \sigma_{ik}^2)$

- Deciding \mathbf{Y} given \mathbf{X}

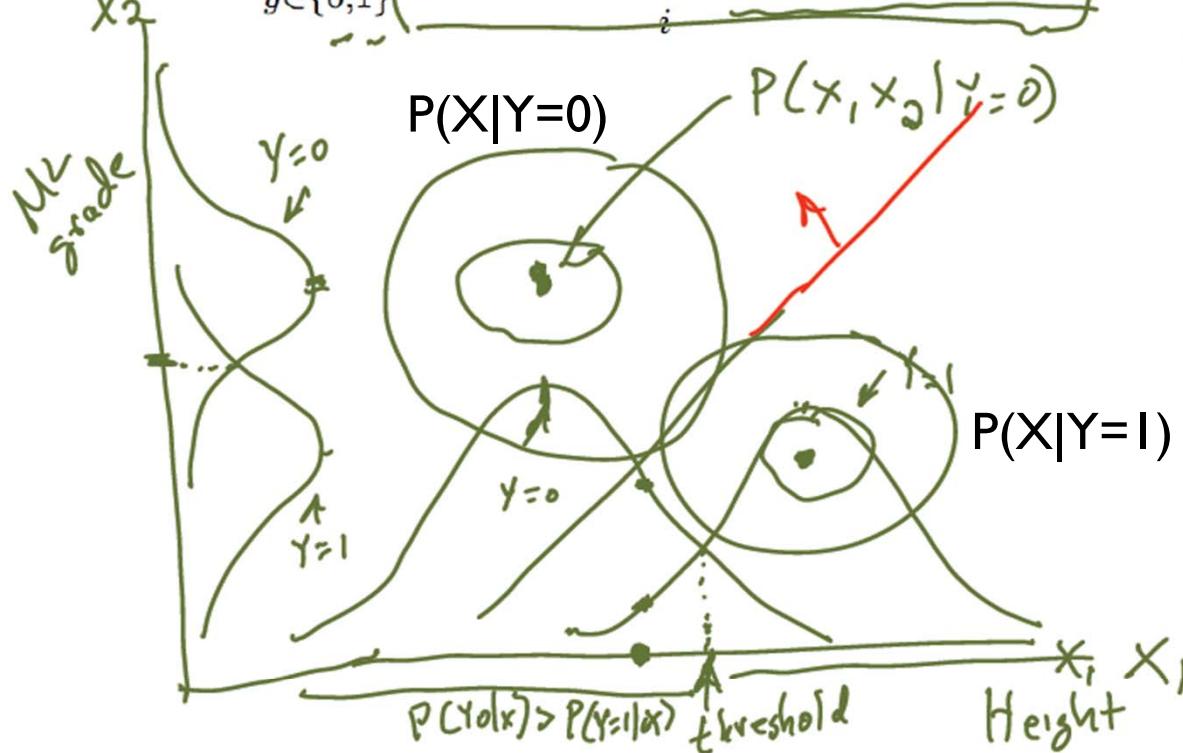
$$P(Y = 0|X) \stackrel{0}{\underset{1}{\gtrless}} P(Y = 1|X)$$

$$P(Y = 0)P(X|Y = 0) \stackrel{0}{\underset{1}{\gtrless}} P(Y = 1)P(X|Y = 1)$$

Gaussian Naïve Bayes – Big Picture

Consider boolean Y , continuous X_i . Assume $P(Y=1)=0.5 = P(Y=0)$

$$Y^{new} \leftarrow \arg \max_{y \in \{0,1\}} P(Y=y) \prod_i P(X_i^{new}|Y=y)$$

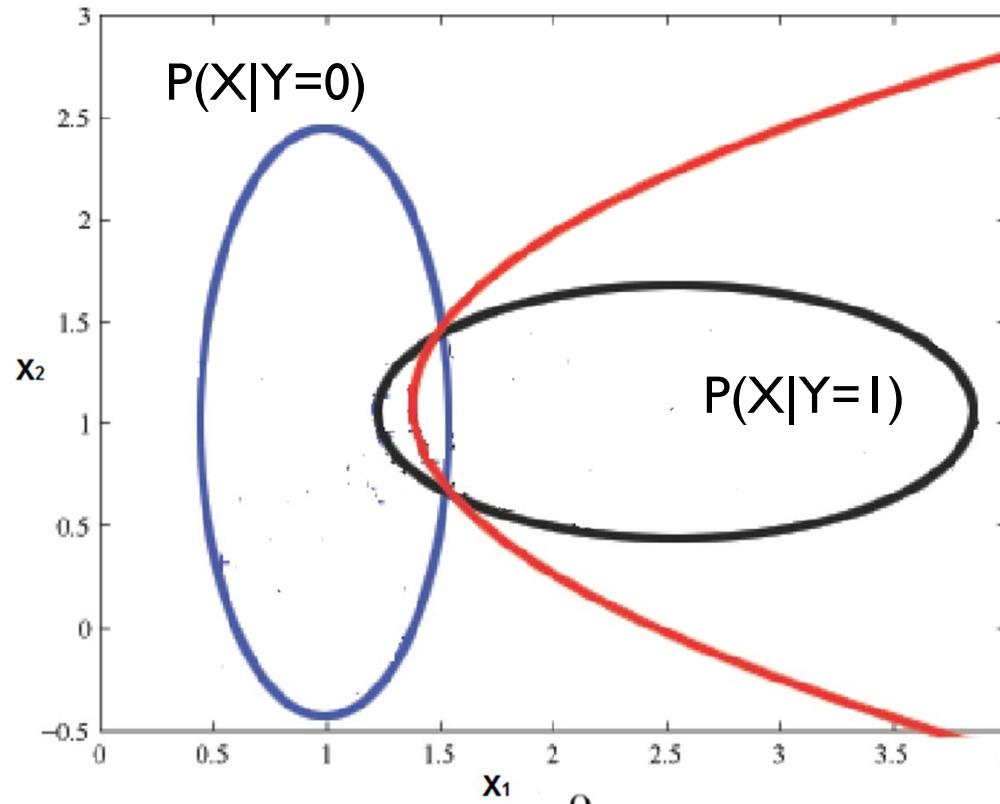


assume
 $P(X_i | Y=k) \sim N(\mu_{ik}, \sigma_{ik}^2)$

$$P(Y=0)P(X|Y=0) \stackrel{0}{\underset{1}{\gtrless}} P(Y=1)P(X|Y=1)$$

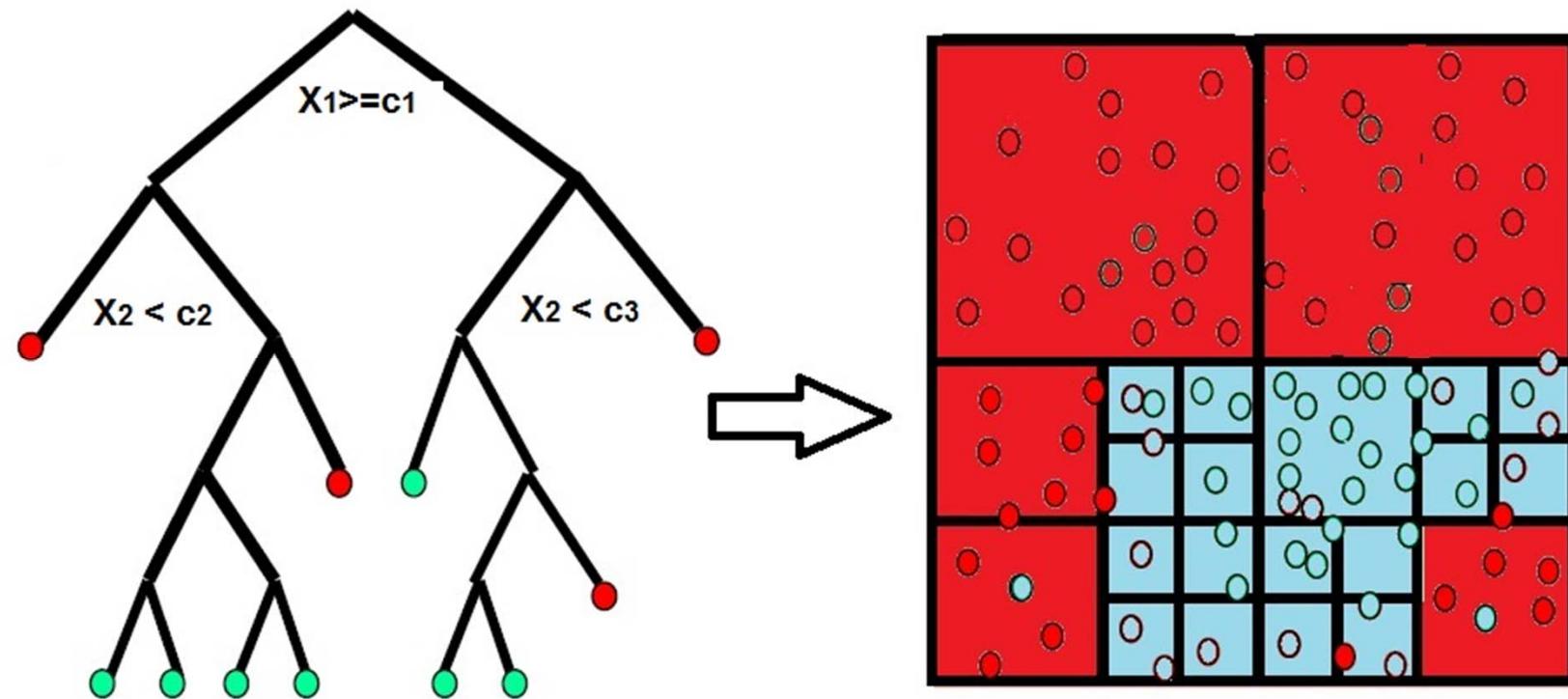
Gaussian naïve Bayes: nonlinear case

- Again, assume $P(Y=1) = P(Y=0) = 0.5$



$$P(Y = 0)P(X|Y = 0) \underset{1}{\overset{0}{\gtrless}} P(Y = 1)P(X|Y = 1)$$

Decision trees



- Decision making on Y: follow the tree structure to a leaf

Outline

- Logistic regression
- Decision surface (boundary) of classifiers
- **Generative vs. discriminative classifiers**
 - Definitions
 - How to compare them
 - GNB-1 vs. logistic regression
 - GNB-2 vs. logistic regression
- Linear regression
- Bias-variance decomposition and tradeoff
- Overfitting and regularization
- Feature selection

Generative and discriminative classifiers

- Generative classifiers
 - Modeling the joint distribution $P(X, Y)$
 - Usually via $P(X, Y) = P(Y) P(X|Y)$
 - Examples: Gaussian naïve Bayes ☺
- Discriminative classifiers
 - Modeling $P(Y|X)$ or simply $f: X \rightarrow Y$
 - Do not care about $P(X)$
 - Examples: logistic regression, support vector machines (later in this course)

Generative vs. discriminative

- How can we compare, say, Gaussian naïve Bayes and a logistic regression?
 - $P(X,Y) = P(Y) P(X|Y)$ vs. $P(Y|X)$?
- Hint: decision making is based on $P(Y|X)$
 - Compare the $P(Y|X)$ they can represent !

Two versions: GNB-1 and GNB-2

- Model assumptions on $P(X, Y) = P(Y)P(X|Y)$

- Bernoulli on Y : $P(Y = 1) = \pi$

- Conditional independence of X

$$P(X = (X_1, X_2, \dots, X_n)|Y = k) = \prod_{i=1}^n P(X_i|Y = k)$$

GNB-1 ◦ Gaussian on $X_i|Y$: $P(X_i|Y = k) \sim N(\mu_{ik}, \underline{\sigma}_{ik}^2)$

GNB-2 ◦ (Additionally,) class-independent variance

$$P(X_i|Y = k) \sim N(\mu_{ik}, \underline{\sigma}_i^2)$$

Two versions: GNB-1 and GNB-2

- Model assumptions on $P(X, Y) = P(Y)P(X|Y)$

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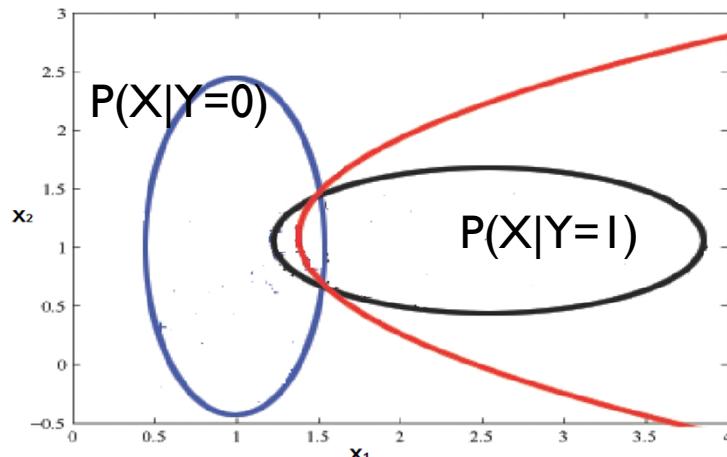
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GNB-1 ◦ Gaussian on $X_i|Y$: $P(X_i|Y = k) \sim N(\mu_{ik}, \underline{\sigma_{ik}^2})$

GNB-2 ◦ (Additionally,) class-independent variance

$$P(X_i|Y = k) \sim N(\mu_{ik}, \underline{\sigma_i^2})$$

Impossible for GNB-2



GNB-2 vs. logistic regression

- GNB-2: $P(X, Y) = P(Y)P(X|Y)$
 - Bernoulli on Y : $P(Y = 1) = \pi$
 - Conditional independence of X , and Gaussian on X_i
 - Additionally, class-independent variance

$$P(X_i|Y = k) \sim N(\mu_{ik}, \underline{\sigma_i^2})$$

- It turns out, $P(Y|X)$ of GNB-2 has the form:

$$P(Y = 1|X) = \frac{1}{1 + \exp\left(\ln \frac{1-\pi}{\pi} + \sum_i \left(\frac{\mu_{i0}-\mu_{i1}}{\sigma_i^2} X_i + \frac{\mu_{i1}^2 - \mu_{i0}^2}{2\sigma_i^2}\right)\right)}$$

GNB-2 vs. logistic regression

- It turns out, $P(Y|X)$ of GNB-2 has the form:

$$P(Y = 1|X) = \frac{1}{1 + \exp\left(\ln \frac{1-\pi}{\pi} + \sum_i \left(\frac{\mu_{i0}-\mu_{i1}}{\sigma_i^2} X_i + \frac{\mu_{i1}^2 - \mu_{i0}^2}{2\sigma_i^2} \right)\right)}$$

- See [Mitchell: Naïve Bayes and Logistic Regression], section 3.1 (page 8 – 10)

- Recall: $P(Y|X)$ of logistic regression:

$$P(Y = 1|X) = \frac{1}{1 + \exp(w_0 + \sum_{i=1}^n w_i X_i)}$$

GNB-2 vs. logistic regression

- $P(Y|X)$ of GNB-2 is **subset** of $P(Y|X)$ of LR
- Given *infinite* training data
 - We claim: $LR \geq GNB-2$

GNB-I vs. logistic regression

- GNB-I: $P(X, Y) = P(Y)P(X|Y)$
 - Bernoulli on Y : $P(Y = 1) = \pi$
 - Conditional independence of X_i and Gaussian on X_i

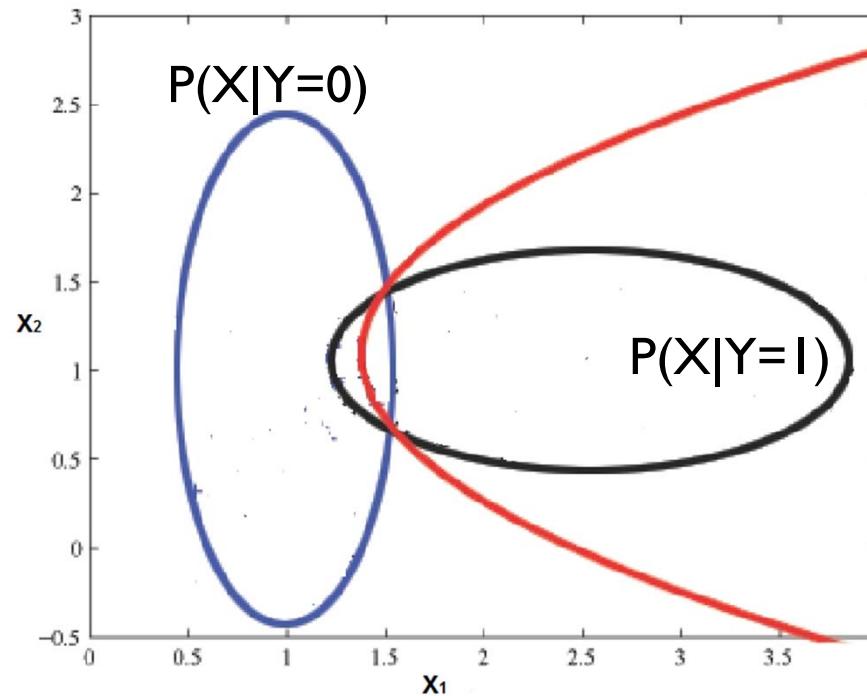
$$P(X_i|Y = k) \sim N(\mu_{ik}, \underline{\sigma_{ik}^2})$$

- Logistic regression: $P(Y|X)$

$$P(Y = 1|X) = \frac{1}{1 + \exp(w_0 + \sum_{i=1}^n w_i X_i)}$$

GNB-I vs. logistic regression

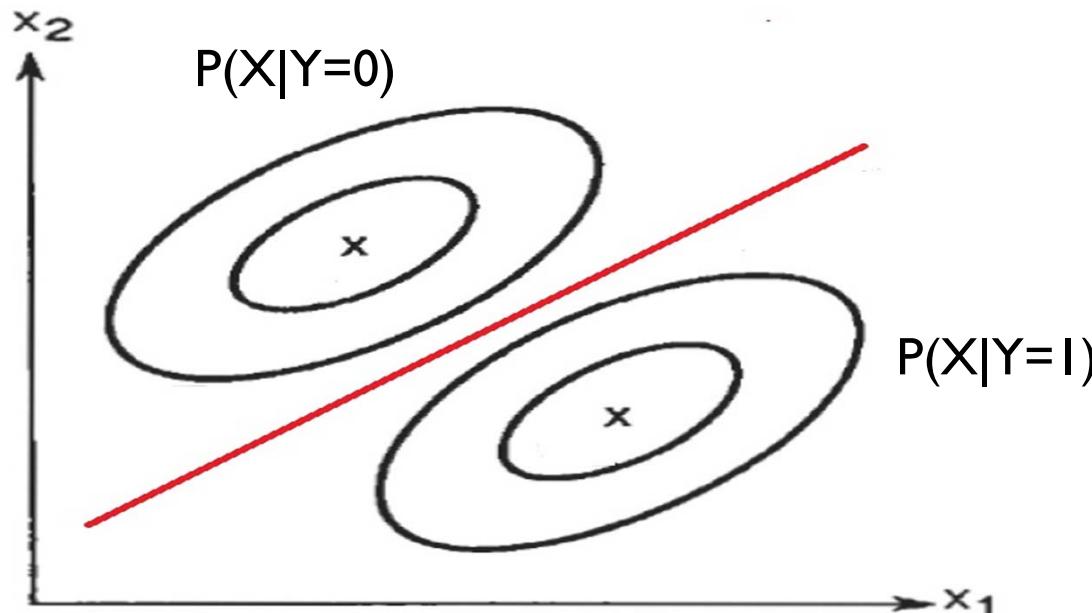
- **None of them** encompasses the other
- First, find a $P(Y|X)$ from GNB-I that **cannot** be represented by LR



- LR only represents linear decision surfaces

GNB-I vs. logistic regression

- **None of them** encompasses the other
- Second, find a $P(Y|X)$ represented by LR that **cannot** be derived from GNB-I assumptions



- GNB-I cannot represent any correlated Gaussian
- But can still possibly be represented by LR (HW2)

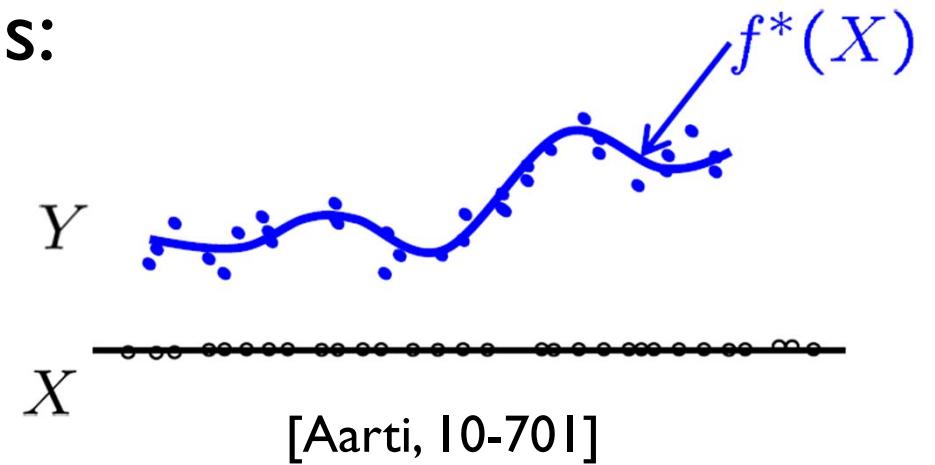
Outline

- Logistic regression
- Decision surface (boundary) of classifiers
- Generative vs. discriminative classifiers
- Linear regression
 - Regression problems
 - Model assumptions: $P(Y|X)$
 - Estimate the model parameters
- Bias-variance decomposition and tradeoff
- Overfitting and regularization
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Regression problems

- Regression problems:
 - Predict Y given X
 - Y is continuous
 - General assumption:

$$Y = f^*(X) + \epsilon, \quad \epsilon \sim Distribution()$$



Linear regression: assumptions

- Linear regression assumptions
 - Y is generated from $f(X)$ plus Gaussian noise

$$Y = f(X) + \epsilon, \quad \epsilon \sim N(0, \sigma^2)$$

- $f(X)$ is a linear function

$$f(X) = w_0 + \sum_{i=1}^n w_i X_i$$

Linear regression: assumptions

- Linear regression assumptions
 - Y is generated from $f(X)$ plus Gaussian noise

$$Y = f(X) + \epsilon, \quad \epsilon \sim N(0, \sigma^2)$$

- $f(X)$ is a linear function

$$f(X) = w_0 + \sum_{i=1}^n w_i X_i$$

- Therefore, assumptions on $P(Y|X, w)$:

$$\begin{aligned} P(Y|X = (X_1, X_2, \dots, X_n)) &\sim N(f(X), \sigma^2) \\ &\sim N(w_0 + \sum_{i=1}^n w_i X_i, \sigma^2) \end{aligned}$$

Linear regression: assumptions

- Linear regression assumptions
 - \mathbf{Y} is generated from $f(\mathbf{X})$ plus Gaussian noise

$$Y = f(X) + \epsilon, \quad \epsilon \sim N(0, \sigma^2)$$

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- Therefore, assumptions on $P(\mathbf{Y}|\mathbf{X}, \mathbf{w})$:

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$$P(Y|X, \mathbf{w}) = \frac{1}{\sqrt{2\pi\sigma^2}} \exp -\frac{(Y - w_0 - \sum_{i=1}^n w_i X_i)^2}{2\sigma^2}$$

Estimating the parameters \mathbf{w}

- Given $\{(X^{(j)}, Y^{(j)})\}_{j=1}^L$, $X^{(j)} = (X_1^{(j)}, X_2^{(j)}, \dots, X_n^{(j)})$
- Assumptions:

$$P(Y|X, \mathbf{w}) = \frac{1}{\sqrt{2\pi\sigma^2}} \exp -\frac{(Y - w_0 - \sum_{i=1}^n w_i X_i)^2}{2\sigma^2}$$

- Maximum ***conditional*** likelihood on data!

$$\hat{\mathbf{w}}_{MCLL} = \arg \max_{\mathbf{w}} \prod_{j=1}^L P(Y^{(j)} | X^{(j)}, \mathbf{w})$$

Estimating the parameters \mathbf{w}

- Given $\{(X^{(j)}, Y^{(j)})\}_{j=1}^L$, $X^{(j)} = (X_1^{(j)}, X_2^{(j)}, \dots, X_n^{(j)})$
- Assumptions:

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- Maximum ***conditional*** likelihood on data!

$$\hat{\mathbf{w}}_{MCLE} = \arg \max_{\mathbf{w}} \prod_{j=1}^L P(Y^{(j)} | X^{(j)}, \mathbf{w})$$

- Let's maximize conditional ***log***-likelihood

Estimating the parameters \mathbf{w}

- Given $\{(X^{(j)}, Y^{(j)})\}_{j=1}^L$, $X^{(j)} = (X_1^{(j)}, X_2^{(j)}, \dots, X_n^{(j)})$
- Assumptions:

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- Maximum ***conditional*** likelihood on data!

$$\hat{\mathbf{w}}_{MCLE} = \arg \max_{\mathbf{w}} \prod_{j=1}^L P(Y^{(j)} | X^{(j)}, \mathbf{w})$$

- Let's maximize conditional ***log***-likelihood

$$\begin{aligned} \max_{\mathbf{w}} l(\mathbf{w}) &\equiv \ln \prod_j^L P(y^j | \mathbf{x}^j, \mathbf{w}) \\ &= \sum_{j=1}^L -\frac{(Y^{(j)} - w_0 - \sum_{i=1}^n w_i X_i^{(j)})^2}{2\sigma^2} + const \end{aligned}$$

Estimating the parameters \mathbf{w}

- Max the conditional log-likelihood over data

$$\begin{aligned} & \operatorname{argmax}_{\mathbf{w}} \sum_{j=1}^L -\frac{(Y^{(j)} - w_0 - \sum_{i=1}^n w_i X_i^{(j)})^2}{2\sigma^2} \\ &= \operatorname{argmax}_{\mathbf{w}} \sum_{j=1}^L -(Y^{(j)} - w_0 - \sum_{i=1}^n w_i X_i^{(j)})^2 \end{aligned}$$

- OR minimize the sum of “squared errors”

$$= \operatorname{argmin}_{\mathbf{w}} \sum_{j=1}^L (Y^{(j)} - w_0 - \sum_{i=1}^n w_i X_i^{(j)})^2$$

- Gradient ascent (descent) is easy
- Actually, a closed form solution exists ☺

Estimating the parameters \mathbf{w}

- Max the conditional log-likelihood over data

$$\operatorname{argmax}_{\mathbf{w}} \sum_{j=1}^L -(Y^{(j)} - w_0 - \sum_{i=1}^n w_i X_i^{(j)})^2$$

- OR

$$\operatorname{argmin}_{\mathbf{w}} \sum_{j=1}^L (Y^{(j)} - w_0 - \sum_{i=1}^n w_i X_i^{(j)})^2$$

- Actually, a closed form solution exists ☺

- $\mathbf{w} = (\mathbf{A}^T \mathbf{A})^{-1} \mathbf{A}^T \mathbf{Y}$

- \mathbf{A} is an L by n matrix: m examples, n variables

- \mathbf{Y} is an L by 1 vector: m examples

Outline

- Logistic regression
- Decision surface (boundary) of classifiers
- Generative vs. discriminative classifiers
- Linear regression
- **Bias-variance decomposition and tradeoff**
 - True risk for a (regression) model
 - Risk of the perfect model
 - Risk of a learning method: bias and variance
 - **Bias-variance tradeoff**
- Overfitting and regularization
- Feature selection

(True) risk

- Consider a regression problem
- (True) risk of a prediction model $f(\cdot)$

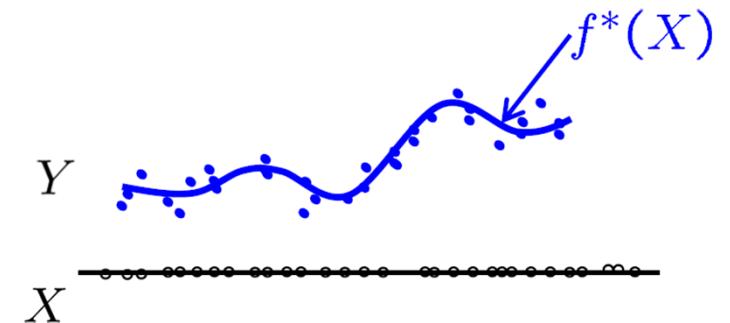
$$R(f) = E_{X,Y}[(f(X) - Y)^2]$$

- **Expected** squared error when we use $f(\cdot)$ to make prediction on future examples

Risk of the perfect model

- The “true” model $f^*()$

$$Y = f^*(X) + \epsilon, \quad \epsilon \sim N(0, \sigma^2)$$



[Aarti, 10-70]

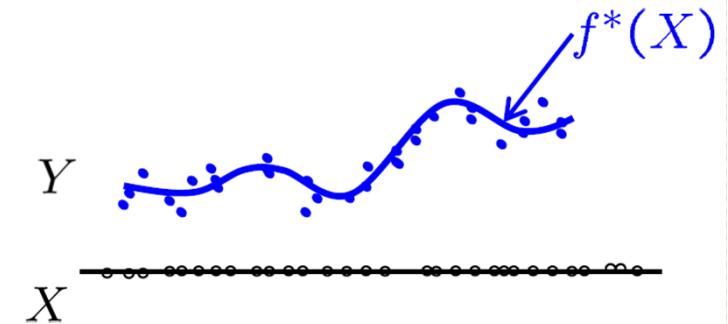
- The risk of $f^*()$

$$R^* = \mathbb{E}_{XY}[(f^*(X) - Y)^2] = \mathbb{E}[\epsilon^2] = \sigma^2$$

- The best we can do !
- σ^2 is the “unavoidable risk”
- Is this achievable ? ... well ...
 - Model makes perfect assumptions
 - $f^*()$ belongs to the class of functions we consider
 - Infinite training data

A learning method

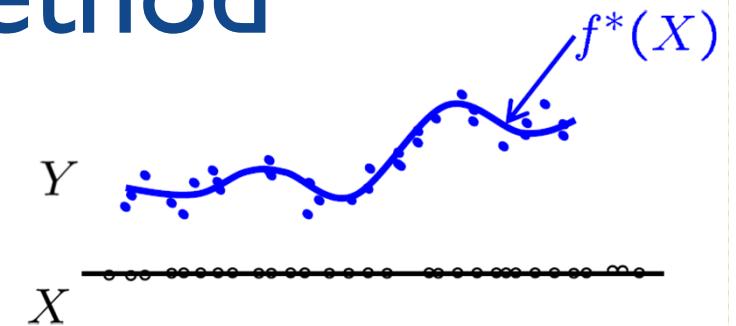
Regression: $Y = f^*(X) + \epsilon$ $\epsilon \sim \mathcal{N}(0, \sigma^2)$



- A learning method:
 - Model assumptions, i.e., the space of models
 - e.g., the form of $P(Y|X)$ in linear regression
 - An optimization/search algorithm
 - e.g., maximum conditional likelihood on data
- Given a set of L training samples
 - $D = \{(X^{(j)}, Y^{(j)})\}_{j=1}^L$
 - A learning method outputs: $\hat{f}_D(\)$

Risk of a learning method

Regression: $Y = f^*(X) + \epsilon$ $\epsilon \sim \mathcal{N}(0, \sigma^2)$

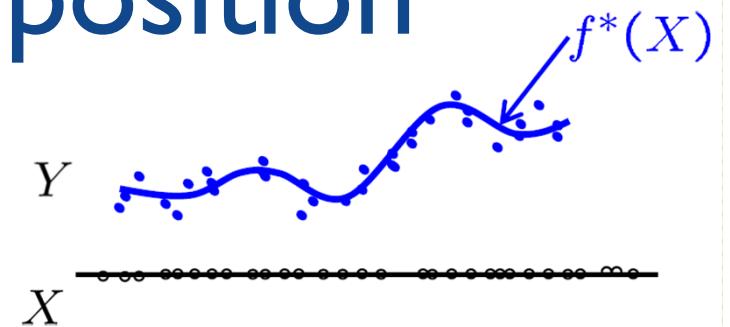


- Risk of a learning method

$$E_D[R(\hat{f}_D)] = E_{X,Y,D}[(\hat{f}_D(X) - Y)^2]$$

Bias-variance decomposition

Regression: $Y = f^*(X) + \epsilon$ $\epsilon \sim \mathcal{N}(0, \sigma^2)$



- Risk of a learning method

$$E_D[R(\hat{f}_D)] = E_{X,Y,D}[(\hat{f}_D(X) - Y)^2]$$

$$= E_{X,Y}[(E_D[\hat{f}_D(X)] - f^*(X))^2]$$

+

Bias²

bias : $E[Z] - Z^*$

variance :

$E[(Z - E[Z])^2]$

$$E_{X,Y}[\underbrace{E_D[(\hat{f}_D(X) - E_D[\hat{f}_D(X)])^2]}_{\text{Variance}}]$$

+

Variance

σ^2

Unavoidable error

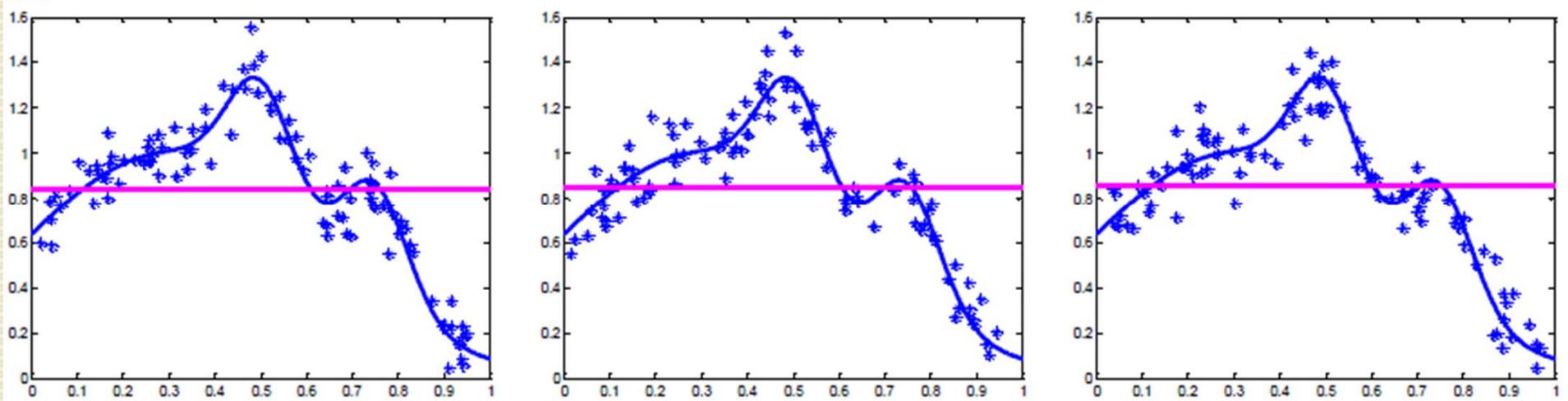
Bias-variance decomposition

$$\begin{aligned} E_D[R(\hat{f}_D)] &= E_{X,Y,D}[(\hat{f}_D(X) - Y)^2] \\ &= E_{X,Y}[(E_D[\hat{f}_D(X)] - f^*(X))^2] \quad \text{Bias}^2 \\ &\quad + E_{X,Y}[\underbrace{E_D[(\hat{f}_D(X) - E_D[\hat{f}_D(X)])^2]}_{+ \sigma^2} \quad \text{Variance}] \end{aligned}$$

- **Bias:** how much is the “mean” estimation different from the true function f^*
 - Does f^* belong to our model space? If not, how far?
- **Variance:** how much is a single estimation different from the “mean” estimation
 - How sensitive is our method to a “bad” training set D ?

Bias-variance tradeoff

- Minimize both bias and variance ? No free lunch
- Simple models: low variance but high bias

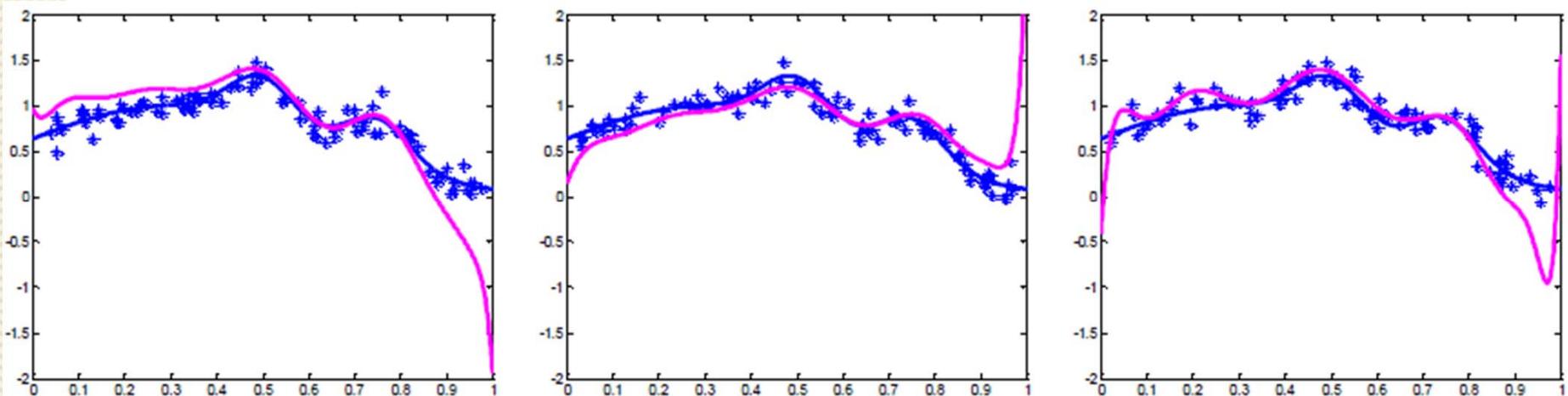


[Aarti]

- Results from 3 random training sets D
- Estimation is very stable over 3 runs (low variance)
- But estimated models are **too simple** (high bias)

Bias-variance tradeoff

- Minimize both bias and variance ? No free lunch
- Complex models: low bias but high variance



- Results from 3 random training sets D
- Estimated models complex enough (low bias)
- But estimation is ***unstable*** over 3 runs (high variance)

Bias-variance tradeoff

- We need a good tradeoff between bias and variance
 - Class of models are not too simple (so that we can **approximate** the true function well)
 - But not too complex to overfit the training samples (so that the **estimation** is **stable**)

Outline

- Logistic regression
- Decision surface (boundary) of classifiers
- Generative vs. discriminative classifiers
- Linear regression
- Bias-variance decomposition and tradeoff
- **Overfitting and regularization**
 - Empirical risk minimization
 - Overfitting
 - Regularization
- Feature selection

Empirical risk minimization

- Many learning methods essentially minimize a loss (risk) function over training data

$$\operatorname{argmin}_{\mathbf{w}} \sum_{j=1}^L R(Y^{(j)}, X^{(j)}, \mathbf{w})$$

- Both linear regression and logistic regression

$$\operatorname{argmax}_{\mathbf{w}} \ln \prod_{j=1}^L P(Y^{(j)} | X^{(j)}, \mathbf{w})$$

$$= \operatorname{argmin}_{\mathbf{w}} -\ln \prod_{j=1}^L P(Y^{(j)} | X^{(j)}, \mathbf{w})$$

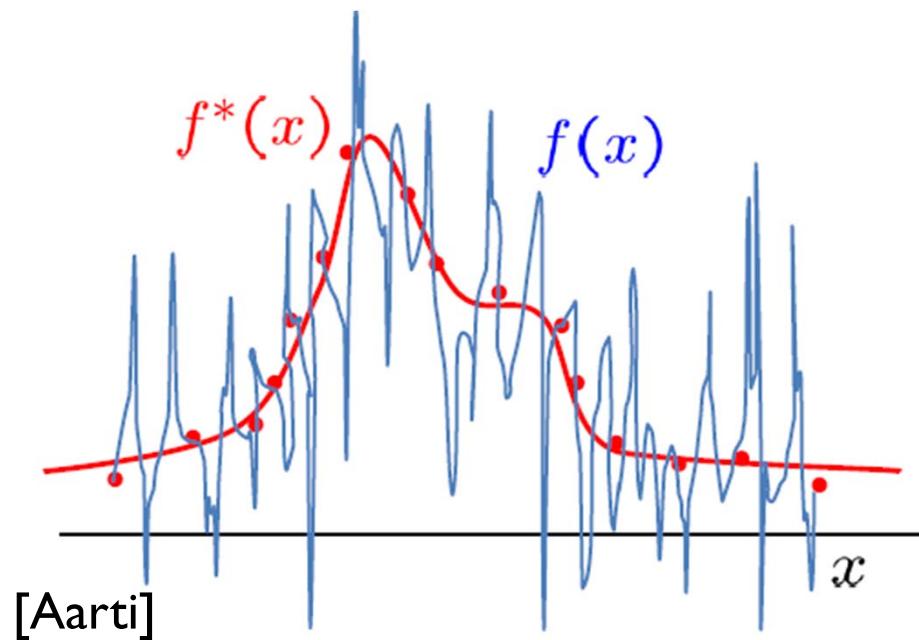
$$= \operatorname{argmin}_{\mathbf{w}} \sum_{j=1}^L -\ln P(Y^{(j)} | X^{(j)}, \mathbf{w})$$

- (linear regression: squared errors)

$$= \operatorname{argmin}_{\mathbf{w}} \sum_{j=1}^L (Y^{(j)} - w_0 - \sum_{i=1}^n w_i X_i^{(j)})^2$$

Overfitting

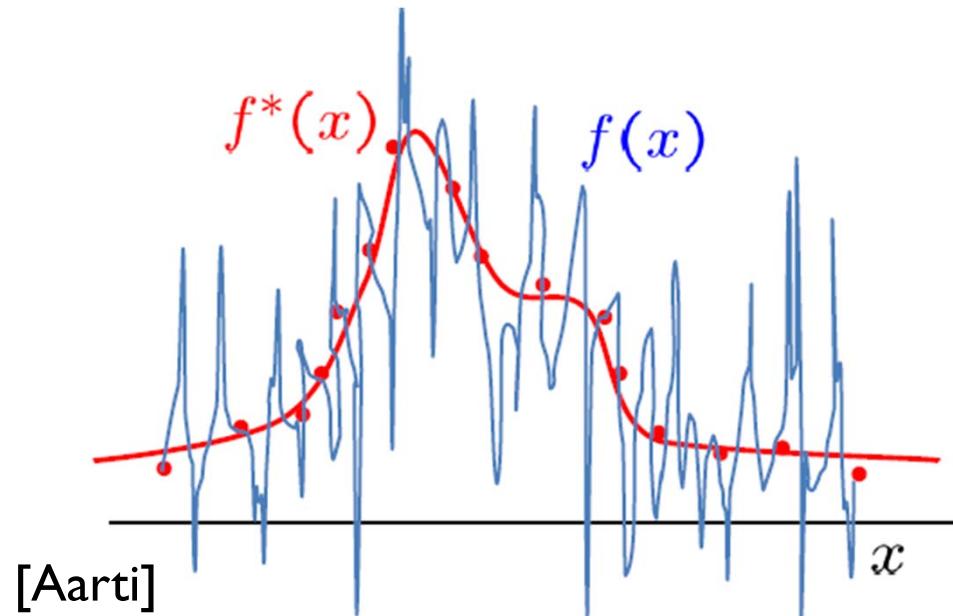
- Minimize the loss over **finite** samples is not always a good thing ... overfitting!



- The blue function has **zero** error on training samples, but is **too complicated** (and “crazy”)

Overfitting

- How to avoid overfitting?



- Hint: complicated and “crazy” models like the blue one usually have **large** model coefficients w

Regularization: L-2 regularization

- Minimize the loss + a regularization penalty

$$\operatorname{argmin}_{\mathbf{w}} - \ln \prod_{j=1}^L P(Y^{(j)}|X^{(j)}) + \lambda \sum_{i=1}^n w_i^2$$

- Intuition: prefer **small** coefficients
- Prob. interpretation: a Gaussian prior on \mathbf{w}
 - Minimize the penalty $\lambda \sum_{i=1}^n w_i^2$ is:

$$\operatorname{min}_{\mathbf{w}} - \ln \prod_{i=1}^n p(w_i), \quad p(w_i) \sim N(0, 1/\lambda)$$

- So minimize loss+penalty is max (log-)posterior

$$\operatorname{argmin}_{\mathbf{w}} - \ln \prod_{j=1}^L P(Y^{(j)}|X^{(j)}) - \ln \prod_{i=1}^n p(w_i)$$

$$\operatorname{argmax}_{\mathbf{w}} \underline{\ln \prod_{j=1}^L P(Y^{(j)}|X^{(j)})} + \underline{\ln \prod_{i=1}^n p(w_i)}$$

Regularization: L-1 regularization

- Minimize a loss + a regularization penalty

$$\operatorname{argmin}_{\mathbf{w}} - \ln \prod_{j=1}^L P(Y^{(j)} | X^{(j)}) + \underline{\lambda \sum_{i=1}^n |w_i|}$$

- Intuition: prefer ***small*** and ***sparse*** coefficients
- Prob. interpretation: a Laplacian prior on \mathbf{w}

Outline

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- **Feature selection**

Feature selection

- Feature selection: select a subset of “useful” features
 - Less features → less complex models
 - Lower “variance” in the risk of the learning method ☺
 - But might lead to higher “bias” ☹

Feature selection

- **Score each feature** and select a subset
 - The mutual information between X_i and Y

$$\hat{I}(X_i, Y) = \sum_k \sum_y \hat{P}(X_i = k, Y = y) \log \frac{\hat{P}(X_i = k, Y = y)}{\hat{P}(X_i = k)\hat{P}(Y = y)}$$

- Accuracy of single-feature classifier $f: X_i \rightarrow Y$
- etc.

Feature selection

- Score each feature and **select a subset**
 - One-step method: select k highest score features
 - Iterative method:
 - Select a highest score feature from the pool
 - **Re-score** the rest, e.g., mutual information **conditioned** on already-selected features
 - Iterate

Outline

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Homework 2 due tomorrow

- Feb. 4th (Friday), 4pm
- Sharon Cavlovich's office (GHC 8215).
- 3 separate sets (each question for a TA)

The last slide

- **Go Steelers !**
- Question ?