18.2.a This is a regression of 4 on D with state fixed effects. So 0 = Et=0 Ziro (Dit - Di) (Yit - T) 5, T=0 5, i=0 (Die-D)2

* where Di and Yi are sample averages over time of Dib and lit, respectively.

18.2.6 For the untreated sub-sample, Dot=0, So Do=0 and for the treated sul-sample Dio=0, Di=1, Di=0.5 = (D00 - D0) (Y00 - \(\bar{Y}_0\) + (D01 - \(\bar{D}_0\)) (Y01 - \(\bar{Y}_0\)) + (D10 - \(\bar{D}_1\)) (Y10 - \(\bar{Y}_1\)) + (D11 - \(\bar{D}_1\)) (Yu-\(\bar{Y}_1\))

(Do - Do)2 + (Do - Do)2 + (Do - Di)2 + (Du - Di)2 $= (D_{10} - \overline{D_{1}})(Y_{10} - \overline{Y_{1}}) + (D_{11} - \overline{D_{1}})(Y_{11} - \overline{Y_{1}})$ $(D_{10} - \overline{D_1})^2 + (D_{11} - \overline{D_1})^2$

 $= \frac{(-0.5)(\lambda^{(0}-\lambda^{(1)})+(0.2)(\lambda^{(1)}-\lambda^{(1)})}{(-0.2)(\lambda^{(0)}-\lambda^{(1)})}$ (0.25) + (0.25)= Y11 - Y10

19.2.C. No, since & doesn't depend on the untreated sub-samples this is only a difference estimator.

18.2.d. If there is no time trend, e.g. E[Yoo]=E[Yoi].

18.5.a. WI MN

Pre 15.23
$$16.42$$
 1.19

Post 16.72 18.16 1.38

Diff $+1.49 - +1.68 = -0.19$

17.1.a.
$$E[X^{4}] = \int x \int_{hh} \sum_{i=1}^{h} F(x_{i}-x_{i}) dx$$

$$= \frac{1}{nh} \int X \sum_{i=1}^{n} E\left(\frac{x_i - x}{h}\right) dx$$

Let
$$y = \left(\frac{x_i - x}{h}\right)$$
, hay = dx, $x = x_i - hy$

$$\Rightarrow = \frac{1}{n} \sum_{i=1}^{n} \int (x_i - y_h) K(y) dy$$

$$=\overline{\chi}_n$$

17.1.6. Var
$$(X^{\pm}) = E[X^{\pm 2}] - E[X^{\pm}]^{2}$$

$$= \int X^{2} \frac{1}{nh} \sum_{i=1}^{n} F(\frac{X_{i} - X}{h}) dX - \overline{X}_{n}^{2}$$

$$= \frac{1}{nh} \int X^{2} \sum_{i=1}^{n} F(\frac{X_{i} - X}{h}) dX - \overline{X}_{n}^{2}$$
Let $y^{2} \left(\frac{X_{i} - X}{h}\right) + h dy = dX$, $x = X_{i} - yh$

$$= \frac{1}{n} \sum_{i=1}^{n} \int (X_{i}^{2} - 2X_{i}yh + y^{2}h^{2}) F(y) dy - \overline{X}_{n}^{2}$$

$$= \frac{1}{n} \sum_{i=1}^{n} \int X_{i}^{2} F(y) dy - \frac{2h}{n} \sum_{i=1}^{n} \int X_{i}^{2} F(y) dy - \overline{X}_{n}^{2}$$

$$= \frac{1}{n} \sum_{i=1}^{n} \int X_{i}^{2} F(y) dy - \frac{2h}{n} \sum_{i=1}^{n} \int X_{i}^{2} F(y) dy - \overline{X}_{n}^{2}$$

$$= \frac{1}{n} \sum_{i=1}^{n} X_{i}^{2} + h^{2} - \overline{X}_{n}^{2}$$

$$= \frac{1}{n} \sum_{i=1}^{n} X_{i}^{2} + h^{2} - \overline{X}_{n}^{2}$$

$$= \frac{1}{n} \sum_{i=1}^{n} X_{i}^{2} + h^{2} - \overline{X}_{n}^{2}$$

17.3 The optimal bandwidth is
$$h = \left(\frac{R_E}{R(f'')}\right)^{1/5} n^{1/5}$$
.

Since the distribution is uniform, $f''=0$, so the optimal bandwidth is whatever is the

largest feasible.

17.4 Using the Same boundwidth will result in a much wider, flatter plot than is appropriate since the Scaling changed. The bandwidth should be divided by 1,000,100 to get the same shape density plot.

19.3. When m(x) is increasing and concave > neg. bias. increasing and convex - pos. bias. decreasing and concave - neg.bias ? decreasing and convex - pos. bias & This is because you're averaging around the point of interest. 19.4. Bias is: $B(X) = \frac{1}{2} m''(X) + m'(X) \frac{f'(X)}{f(X)}$ = B f '(x) If $\beta>0$, then B(x)>0 if f'(x)>0 and B(x)<0it ticx) <0. If B < 0, then B(x)>0 if f'(x)<0 and B(X) < îc f \(\x) >0. f(x) is the marginal density of X. for $\beta>0$, if f'(x)>0,

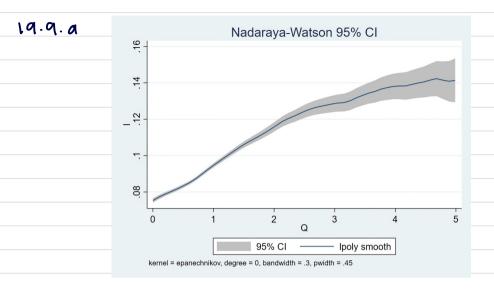
right of x, so the bias is positive, if f'(x) < 0, then the mass of x is concentrated to the left of x, so the bias is negative.

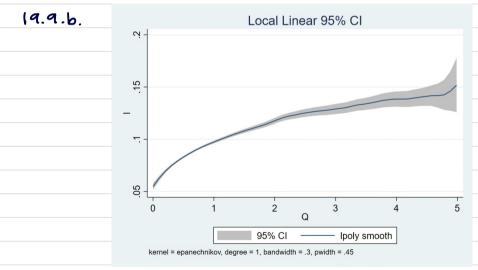
For g < 0, if f'(x) > 0, then the mass of x is concentrated to the left of $x \rightarrow$ negative bias. If f'(x) < 0, the mass is concentrated to the

then the mass of x is concentrated to the

right -> positive bias.

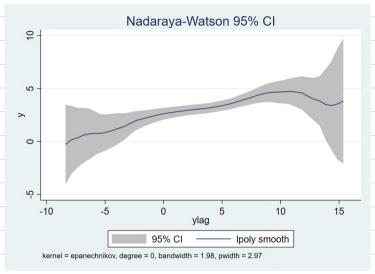
In short, if the mass is to the left of x, there's negative bias. If the mass is concentrated to the right, there is positive bias.





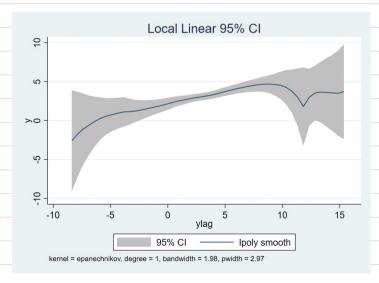
19.9.c. It appears that there is non-linearity in the relationship.

19.11.a. Done in stata.





19.11.6.



in the relationship.