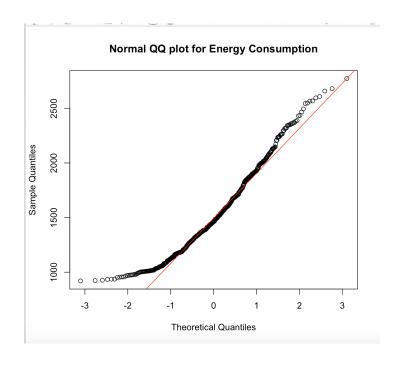
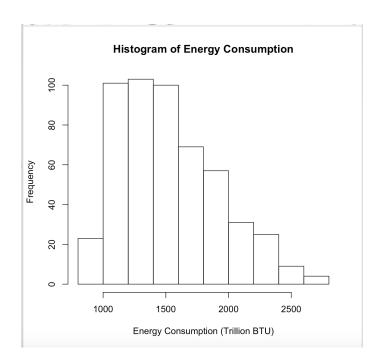
## **CSC 425** HW 4 **Sarah Cummings**

## 1) Energy Consumption Problem a. Analysis of distribution of energy consumption





For this data, we have the following summary statistics:

Minimum 919.805000 Maximum 2772.805000 1. Quartile 1215.065000 3. Quartile 1771.142250

Mean 1529.458182

Median 1463.134000

Skewness 0.661541 Kurtosis -0.114761

Title: Jarque - Bera Normalality Test

Test Results:

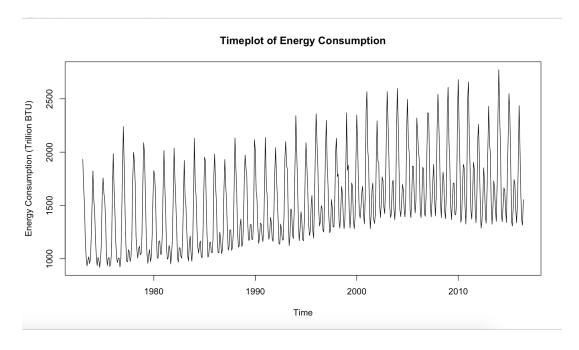
STATISTIC: X-squared: 38.5278

P VALUE:

Asymptotic p Value: 4.303e-09

Looking at the histogram, there appears to be a right skew in the data. The Jarque-Bera Normality test has a significant statistic, so we reject the null hypothesis of a normal distribution.

## b. Analysis of time plot of energy consumption

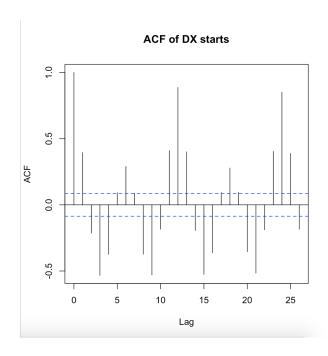


Looking at the time plot, the energy consumption appears to be increasing over time. There also is likely some seasonality, since there are repetitive flocculations that seem may coincide with them of year. The energy consumption appears to have regular cycles of high values (around 2000) and low values (around 1000) that likely correspond to the seasons where we use more or less energy due to weather. The variance seems seems to be increasing over time.

### c. Should we transform the data?

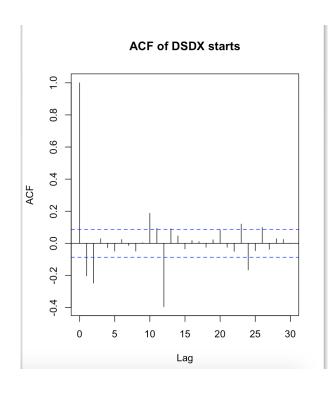
I would not transform the data.

### d. Analyze the ACF of the first difference of the time series data



Looking at the ACF of the first difference, we see that lags that are multiples of three are significantly different than zero. Because of the patterns in the ACF, we can conclude there is seasonality in this data.

### e. After de-trending and de-seasonalizing, do you obtain a stationary time series?



Box-Ljung test

data: sdx

X-squared = 53.276, df = 3, p-value = 1.601e-11

Box-Ljung test

data: sdx

X-squared = 55.225, df = 6, p-value = 4.175e-10

Yes. Given the small p values of both of our Box Ljung tests, we reject the null hypothesis of independence and conclude energy consumptions are serially correlated.

### f. Find initial SARIMA model with autoartima:

Series: ts1

ARIMA(2,0,3)(0,0,2)[12] with non-zero mean

### Coefficients:

ar1 ar2 ma1 ma2 ma3 sma1 sma2 intercept 0.6671 -0.5847 0.271 0.5025 0.3666 0.6980 0.3979 1527.4942 s.e. 0.1236 0.0824 0.123 0.0727 0.0543 0.0537 0.0433 28.9166

sigma^2 estimated as 19198: log likelihood=-3315.3 AIC=6648.6 AICc=6648.95 BIC=6686.91 z test of coefficients:

```
intercept 1527.494150 28.916581 52.8242 < 2.2e-16 ***
Signif. codes: 0 '***' 0.001 '**' 0.01 '*' 0.05 '.' 0.1 ''1
g. Find adequate model:
Alt model 1:
Series: ts1
ARIMA(0,1,1)(0,1,1)[12]
Coefficients:
     ma1
            sma1
   -0.5208 -0.7819
s.e. 0.0892 0.0261
sigma<sup>2</sup> estimated as 8470: log likelihood=-3028.84
AIC=6063.69 AICc=6063.74 BIC=6076.39
z test of coefficients:
   Estimate Std. Error z value Pr(>|z|)
ma1 -0.520765 0.089178 -5.8396 5.233e-09 ***
sma1 -0.781927  0.026119 -29.9375 < 2.2e-16 ***
Alt model 2:
Series: ts1
ARIMA(0,1,2)(0,1,1)[12]
Coefficients:
     ma1
             ma2 sma1
   -0.4898 -0.4307 -0.7983
s.e. 0.0370 0.0370 0.0278
sigma<sup>2</sup> estimated as 7099: log likelihood=-2985.05
AIC=5978.1 AICc=5978.18 BIC=5995.03
z test of coefficients:
   Estimate Std. Error z value Pr(>|z|)
ma1 -0.489841 0.037049 -13.221 < 2.2e-16 ***
ma2 -0.430743 0.037013 -11.637 < 2.2e-16 ***
sma1 -0.798329  0.027772 -28.746 < 2.2e-16 ***
Signif. codes: 0 '***' 0.001 '**' 0.01 '*' 0.05 '.' 0.1 ' '1
```

### Alt model 3:

Series: ts1

ARIMA(1,1,2)(0,1,1)[12]

### Coefficients:

ar1 ma1 ma2 sma1 0.2928 -0.7329 -0.2195 -0.7959 s.e. 0.0941 0.0964 0.0887 0.0284

sigma^2 estimated as 6990: log likelihood=-2980.74 AIC=5971.47 AICc=5971.59 BIC=5992.64

#### z test of coefficients:

Estimate Std. Error z value Pr(>|z|)
ar1 0.292757 0.094107 3.1109 0.001865 \*\*
ma1 -0.732883 0.096445 -7.5990 2.984e-14 \*\*\*
ma2 -0.219543 0.088687 -2.4755 0.013306 \*
sma1 -0.795913 0.028372 -28.0526 < 2.2e-16 \*\*\*
--Signif. codes: 0 '\*\*\*' 0.001 '\*\*' 0.01 '\*' 0.05 '.' 0.1 ' '1

### a. Model selected:

The autoarima and the alt models 2 and 3 do not pass the box-jung test on model residuals, thus I will proceed with alt model 1:

 $(1-B^{12})(1-B)$  xt =  $(1-0.5208 B)(1-0.7819^2)$ at, with var(at) = 8470.

# **b.** Does model provides an adequate explanation of the time process ? Analysis of model. ARIMA(0,1,1)(0,1,1)[12]

Box-Ljung test data: m2\$residuals

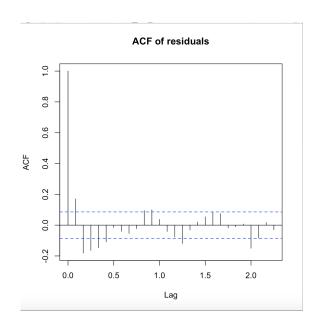
X-squared = 64.8, df = 4, p-value = 2.836e-13

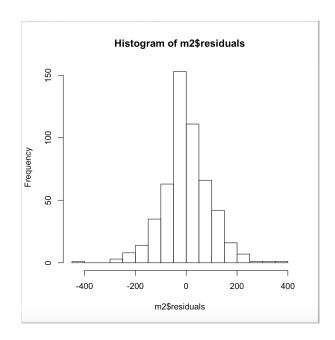
Box-Ljung test

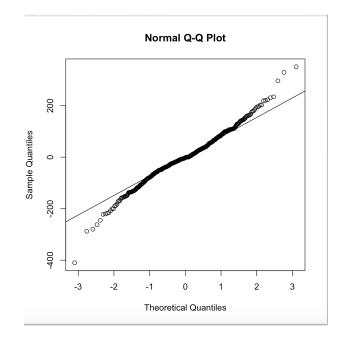
data: m2\$residuals X-squared = 72.359, df = 8, p-value = 1.664e-12

Box-Ljung test data: m2\$residuals

X-squared = 78.368, df = 10, p-value = 1.048e-12







This model has normally distributed residuals as seen above in the histogram and the QQ plot. It also passes the Box- Ljung test of residuals.

## h. Compute forecasts for energy consumption for the next five months using the selected model.

f1
Point Forecast Lo 80 Hi 80 Lo 95 Hi 95

Jul 2016
1732.559 1614.613 1850.506 1552.176 1912.943

Aug 2016
1689.049 1558.257 1819.840 1489.021 1889.076

Sep 2016
1421.324 1278.841 1563.806 1203.415 1639.232

Oct 2016
1318.842 1165.557 1472.127 1084.413 1553.271

Nov 2016
1581.857 1418.482 1745.231 1331.997 1831.717

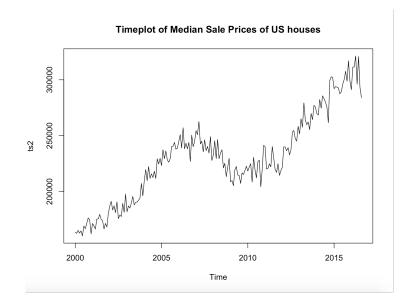
# i. Apply backtesting to compute the MAPE for the fitted model. Discuss the accuracy of your model forecasts.

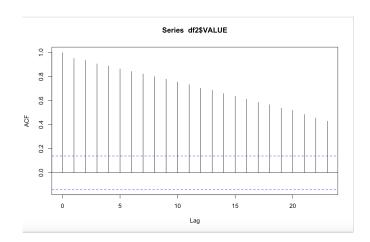
- [1] "Mean Absolute Percentage error"
- [1] 0.05096556

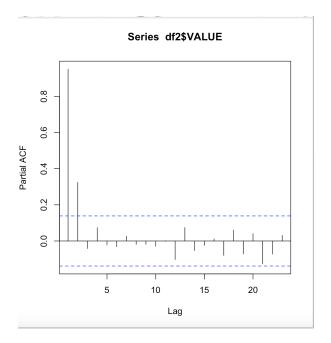
This is a fairly low MAPE, thus our model is fairly accurate.

## 2) Home sales problem

## 1. Plot observed time series and its ACFs, 20 lags







## 2. Analyze if it is stationary using the ACF and Dickey Fuller Test

Looking at the time plot, it looks like this is a unit root non stationary TS. The ACF plot decays, but it decays very slowly. It looks like the ACF plot does not reach zero.

The dickey fuller tests are as follows:

Title: Title: Augmented Dickey-Fuller Test Augmented Dickey-Fuller Test Augmented Dickey-Fuller Test Test Results: Test Results: Test Results: PARAMETER: PARAMETER: PARAMETER: Lag Order: 5 Lag Order: 3 Lag Order: 7 STATISTIC: STATISTIC: STATISTIC: Dickey-Fuller: -0.7586 Dickey-Fuller: -0.6088 Dickey-Fuller: -0.9588 P VALUE: P VALUE: P VALUE: 0.7742 0.8298 0.6998 Description: Description: Description: Sat Nov 5 17:50:50 2016 by user: Sat Nov 5 17:51:03 2016 by user:

Since we have large p values, we cannot reject the null hypothesis. Thus, the process can be considered non stationary and its dynamic behavior can be explained by an ARIMA(p,1,q) model.

## 3. Specify an ARIMA model that describes the behavior over time of median home prices.

Series: ts2

ARIMA(0,1,1) with drift

Sat Nov 5 17:50:45 2016 by user:

Coefficients:

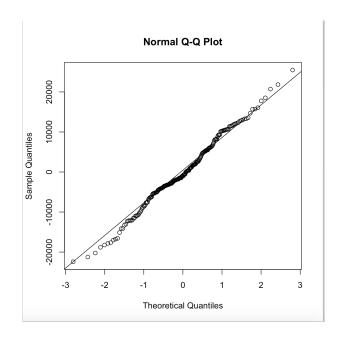
ma1 drift -0.6245 680.3049 s.e. 0.0541 239.0084

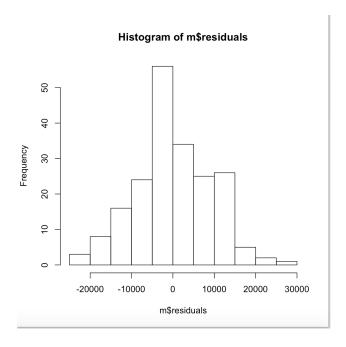
sigma^2 estimated as 79945151: log likelihood=-2092.2 AIC=4190.39 AICc=4190.52 BIC=4200.27

#### a. Write Down model

(1-B)  $Xt = 680.3049 + a_t - 0.6245 a_{t-1}$ 

### b. Residual analysis:





Looking at the histogram and QQ plot of residuals, the residuals appear to be fairly normal. However, they do not pass the Box Ljung test.

```
Box-Ljung test
```

```
data: m$residuals
```

X-squared = 3.0165, df = 4, p-value = 0.5551

Box-Ljung test

data: m\$residuals

X-squared = 6.9049, df = 8, p-value = 0.5469

Box-Ljung test

data: m\$residuals

X-squared = 7.9727, df = 10, p-value = 0.6315

I tried did the same model with the natural log of the values hoping it would improve the residuals test, but that did not work.

## 4. Modeling discussion

This data is unit root non stationary with unit-root of order 1. This means there is a zero intercept to the data, and the time series has a linear trend. The linear trend is described in the drift. The median home value in the US has a linear relationship with time. Home values are increasing over time.

## 5. 5-step ahead forecasts using the fitted model. Write down the forecasts and their standard errors. Do the forecasts show an increasing trend?

```
Point Forecast Lo 80 Hi 80 Lo 95 Hi 95

Sep 2016 298679.9 287221.3 310138.6 281155.5 316204.4

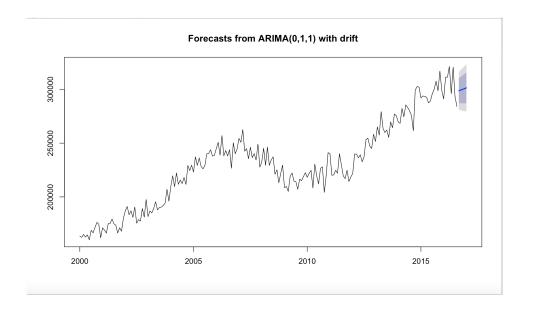
Oct 2016 299360.2 287120.3 311600.2 280640.8 318079.7

Nov 2016 300040.5 287066.2 313014.9 280197.9 319883.2

Dec 2016 300720.9 287051.5 314390.2 279815.3 321626.4

Jan 2017 301401.2 287070.4 315731.9 279484.2 323318.1
```

The forecasts do continue an increasing trend. I'm not sure where to find the errors for the forecasts, but the error for the model is: 0.05407



## 6. Use backtesting procedures to compute the RMSE and the MAPE $\,$

- [1] "RMSE of out-of-sample forecasts"
- [1] 98.01205
- [1] "Mean Absolute Percentage error"
- [1] 0.04451203

Our model is expected to be off by about 4%. On average, the expected values = actual values  $\pm$  0.04(actual values)