

# A Data-Driven Approach to Political and Health Indicators on Maternal Mortality

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Since the *Dobbs v. Jackson Women's Health Organization* decision was handed down by the US Supreme Court in 2022, the effect of access to abortion care on maternal mortality has become an increasingly salient topic. This paper asks whether there is a relationship between political indicators and maternal mortality rates when controlling for other predictive factors such as health indicator data. Moreover, we make the case for the necessity of continuing to collect accurate data on the issue to improve and expand this research. Through several specifications using OLS regressions, data sources ranging from estimate type data sets for maternal mortality, political indexes on abortion restrictions and state governments, as well as maternal health indicators, we found there was a positive correlation of at least 0.42 more deaths per 100K, even when controlling for other possible indicators of maternal mortality risk. The same result was obtained for the increasing level of abortion restriction- the stricter a state's abortion policies, the higher the rate of maternal mortality at the county-level when controlling for other maternal mortality indicators (an increase of 1.02 deaths in the "Most Restrictive" category).

Maternal Mortality - trends | Health and Political Indicators | Abortion Policy | Politics and Maternal Mortality

Abortion laws have changed with remarkable speed in the past decade compared to when they first reached the high court with the decision of *Roe v. Wade*. Leading up to the landmark decision being struck down in the summer of 2022, states across the country were already passing more stringent abortion laws, affecting access to abortion-related healthcare. Challenges exist in this area of study in terms of availability of data, due to both privacy concerns and variations in state data practices, placing limitations on conclusions that can be drawn about patterns in states with tighter abortion restrictions or anti-abortion state leadership.

## Adding to Existing Research

State abortion policies have been known to affect maternal mortality. A study published in the *American Public Health Association* outlines this relationship, concluding that restricting access to abortion care at the state level may increase the risk of maternal mortality.\*

A more grim study explored the relationship between state-level variations in mortality among young children and state abortion restriction policies, concluding that stricter policies were associated with an average increase of 5.7 deaths per state.†

\*Dovile Vilda, Maeve E. Wallace, Clare Daniel, Melissa Goldin Evans, Charles Stoecker, Katherine P. Theall, "State Abortion Policies and Maternal Death in the United States, 2015–2018", *American Journal of Public Health* 111, no. 9 (September 1, 2021): pp. 1696-1704

†Bisakha Sen, Martha Slay Wingate, and Russell Kirby, "The Relationship Between State Abortion-Restrictions and Homicide Deaths Among Children Under 5 Years of Age: A Longitudinal Study," *Social Science and Medicine* 75, no. 1 (July 2012): 156–64, <https://doi.org/10.1016/j.socscimed.2012.01.037>.

## Significance Statement

Collecting maternal mortality data is crucial for policy analysts and researchers, but its is doubly important that this data be organized, accessible, and accurate. In a post-Dobbs decision environment, data now exists from a pre- and post-Roe time period, by which the actual health and societal effects of abortion access can be measured in a way that historically has not been possible.

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95 The implications of these results not only illustrate how impor- 142  
96 tant these findings are to maternal health, but also the health 143  
97 and well-being of children. 144

## 98 145 99 **The Effect of Government, Abortion Protections, and** 146

100 **Restrictions.** Ten states had abortion measures on the ballot in 147  
101 2024. Seven of those states adopted protections for abortion rights, 148  
102 six of which were implemented immediately (Nevada is required to 149  
103 vote twice to pass ballot measures, and therefore will vote again in 150  
104 2026), and the remaining three did not pass abortion protections 151  
105 or passed complete bans.<sup>‡</sup> 152

106 The effects of abortion bans are seemingly well-known. Our 153  
107 goal was to generate a dataset that could be easily replicated and 154  
108 expanded for future research in order to analyze the effects of 155  
109 restrictive state abortion policies. We in turn used this data to 156  
110 assess the relationship between political indicators of maternal 157  
111 mortality, controlling for other predictive factors, specifically 158  
112 health indicator data. 159

113 **Our Findings.** When examining the relationship between maternal 160  
114 mortality rate and residence in a GOP controlled state, we found 161  
115 that there was a positive correlation of at least 0.42 more deaths 162  
116 per 100K at the county level, even when correcting for other 163  
117 possible indicators of maternal mortality risk. The same result 164  
118 was obtained for the increasing level of abortion restriction- the 165  
119 stricter a state's abortion policies, the higher the rate of maternal 166  
120 mortality at the county level when controlling for other maternal 167  
121 mortality indicators (an increase of 1.02 deaths in the "Most 168  
122 Restrictive" category). Certain limitations must be carefully 169  
123 considered when interpreting this data, as many factors such 170  
124 as economic disparities, maternal age, characterization of the 171  
125 pregnancy (wanted, unwanted), were not included in the analysis 172  
126 but can be reasonably assumed to affect maternal mortality rates 173  
127 and the impacts of restrictive abortion policies. 174  
128 175

## 129 176 130 **Data and Methods** 177

131 **Data Sources and Documentation. Maternal Mortality.** We 178  
132 acquired the data set for estimated maternal mortality and demo- 179  
133 graphics from the Global Health Data Exchange (GHDx) data 180  
134 catalog, coordinated by the University of Washington's Institute 181  
135 for Health Metrics and Evaluation (IHME). This dataset provides 182  
136 estimated maternal mortality rates at the county level across 183  
137 multiple years.<sup>§</sup> 184  
138 185

139 <sup>‡</sup>Erin Geiger Smith and Kathrina Szymborski Wolfkot, "Voters in Seven States Pass Measures to Protect Abortion," State Court Report, November 6, 2024, <https://statecourtreport.org/our-work/analysis-opinion/voters-seven-states-pass-measures-protect-abortion>. 186

140 <sup>§</sup>Institute for Health Metrics and Evaluation, "Causes of Death (COD) Visualization," updated April 3, 2024, <https://vizhub.healthdata.org/cod/>. 187

189	<b>Maternal Health Indicators.</b> We acquired our data set for	236
190	maternal health indicators from the Health Resources and Services	237
191	Administration’s (HRSA) Maternal and Infant Health Mapping	238
192	(MIHM) Tool. <sup>¶</sup> HRSA’s MIHM Tool explores a diverse range of	239
193	maternal and infant population information and health resources	240
194	indexed by county FIPS codes. We included the following health	241
195	indicators, all of which were cataloged as incidence per 100K	242
196	births: prenatal care in the first trimester, diabetes pre-pregnancy,	243
197	hypertension pre-pregnancy, and obesity pre-pregnancy.	244
198	<b>Abortion Restrictions/Protections Index.</b> We acquired	245
199	our data set for state abortion laws from the Guttmacher Insti-	246
200	tute’s Interactive Map, which groups states according to seven	247
201	categories. <sup>  </sup> The data set is current as of November 26, 2024.	248
202	<b>State Government.</b> Finally, our data set for state legislature	249
203	party control was scraped from Ballotpedia and filtered to the	250
204	year 2017 to allow for a lagged comparison of party control and	251
205	resulting health outcomes, indexed by state. <sup>**</sup> The data for the	252
206	state governor party came from Wikipedia for the year 2017. Since	253
207	these data sets reflected election results, we indicated party control	254
208	began the year following the election year, not the election year	255
209	itself.	256
210		257
211	<b>Limitations Among Chosen Datasets.</b> The primary limitation	258
212	of our data is the necessity to rely on estimated data given the	259
213	inconsistency in measurement and availability of health data. We	260
214	also note that our data does not span across a consistent time	261
215	period with maternal mortality rates estimated for 2019 and our	262
216	health indicator covariates ranging across three years (2017-2020).	263
217	We pull from enough sources and include enough overlapping and	264
218	adjacent indicators that this is not a large concern for our project.	265
219	Generally, medical conditions, such as diabetes and hyperten-	266
220	sion, are likely under-reported, and availability and accuracy	267
221	depends on geographic location. HRSA notes the degree of	268
222	smoothing is inversely proportional to the number of events,	269
223	borrowing from data in neighboring counties, to ensure usable	270
224	data, and estimates were suppressed if there were fewer than 10	271
225	events in the county and adjacent counties <sup>††</sup> . HRSA worked in	272
226	agreement with some counties directly, but cannot disclose which	273
227	counties specifically due to confidentiality. This is the nature of	274
228	health data that is not collected from a primary resource.	275
229	Additionally, there may be measurement error in terms of how	276
230	pregnancy related deaths are recorded. A study by the American	277
231		278
232	<sup>¶</sup> U.S. Department of Health and Human Services, Health Resources and Services Administration (HRSA). "Maternal and Child Health Bureau Interactive Map Tool." <a href="https://data.hrsa.gov/maps/mchb">https://data.hrsa.gov/maps/mchb</a> .	279
233	<sup>  </sup> Guttmacher Institute, "Interactive Map: US Abortion Policies and Access After Roe," <a href="https://states.guttmacher.org/policies/">https://states.guttmacher.org/policies/</a> .	280
234	<sup>**</sup> Ballotpedia, "Main Page," <a href="https://ballotpedia.org/Main_Page">https://ballotpedia.org/Main_Page</a> .	281
235	<sup>††</sup> Data downloaded from the Maternal and Infant Health Mapping Tool, a website developed by the Health Resources and Services Administration, Maternal and Child Health Bureau. <a href="https://data.hrsa.gov/maps/mchb/">https://data.hrsa.gov/maps/mchb/</a> . Accessed Dec 2024.	282

## DATA AND METHODS

<b>Maternal Health Indicators</b>	<ul style="list-style-type: none"> <li>- U.S. Department of Health and Human Services, Health Resources and Services Administration (<b>HRSA</b>). "Maternal and Child Health Bureau Interactive Map Tool."</li> <li>- Indexed by county <b>FIPS</b> codes (and thus HRSA region, state, and county name)</li> </ul>
<b>Abortion Restrictions</b>	<ul style="list-style-type: none"> <li>- <b>Guttmacher Institute</b>, "Interactive Map: US Abortion Policies and Access After Roe," current as of Nov 26, 2024</li> <li>- Indexed by <b>state</b></li> </ul>
<b>Maternal Mortality</b>	<ul style="list-style-type: none"> <li>- University of Washington's Institute for Health Metrics and Evaluation (<b>IHME</b>), "Causes of Death (COD) Visualization," updated April 3, 2024</li> <li>- Indexed by <b>year</b></li> </ul>
<b>State Government</b>	<ul style="list-style-type: none"> <li>- Scraped from Ballotpedia</li> <li>- Indexed by <b>state</b></li> </ul>

**Fig. 1.** Data Sources and Indexes

Journal of Obstetrics and Gynecology suggests that rising rates between 2018 and 2021 can partially be attributed to changes in maternal mortality surveillance. In 2002, a "pregnancy-related death" checkbox was added on the national death certificate. The study found that this checkbox was unreliable, including deaths of male patients or female patients who were not even pregnant. This led to an over-inflated figure for data collected from 2002-2018.<sup>††</sup> Assuming these errors are evenly distributed, we assess the risk of the error creating significant bias is low.

**Variables and Covariates. Maternal Mortality (dependent variable):** The maternal mortality rate in this dataset is reported as a fraction of live births. To interpret the rates in the standard unit of deaths per 100,000 live births, values must be multiplied by 100,000

**Guttmacher Index (independent variable):** Dummy variables based on state's level of restriction/protection are as follows:  
 - Most Restrictive - Very Restrictive - Restrictive - Some Restrictions/Protections - Protective - Very Protective - Most Protective

**GOP Trifecta (independent variable):** dummy variable indicating a Republican controlled state legislature and executive branch.

**Health Indicators (covariates):** certain indicators of overall health in the population to control for other factors that may

<sup>††</sup>Jane Smith, "Maternal Mortality Overestimate: Deaths, Births, Health Disparities," NPR, March 13, 2024, <https://www.npr.org/sections/health-shots/2024/03/13/1238269753/maternal-mortality-overestimate-deaths-births-health-disparities>.

377	increase risk of mortality. - OBGYN Provider Rate - providers per	424
378	100K - Prenatal Care First Trimester - Percentage - Obesity Pre-	425
379	Pregnancy - Percentage - Diabetes Pre-Pregnancy - Percentage -	426
380	Hypertension Pre-Pregnancy - Percentage	427
381		428
382	<b>Data Wrangling.</b> All data required some level of cleaning, whether	429
383	it was standardizing column names or ensuring the county-level	430
384	data was able to be matched with other sources.	431
385	<b>Maternal Mortality</b> (maternal_mortality_county_data_wrangling.ipynb)	432
386	Maternal Mortality was provided at national, state and county	433
387	levels and had observations for various racial and ethnic groups	434
388	over multiple years. The entire data set was split into multiple	435
389	files to manage the file size. The first steps were to combine	436
390	all the files and filter the resulting data frame to only county-	437
391	level rows. Because Louisiana and Alaska both use different	438
392	geographical regions, Parishes and Boroughs, respectively, initially,	439
393	we accidentally filtered both states out of our data set and had to	440
394	readjust our code. Since the original data set included data across	441
395	multiple years and broken down by race, we filtered the data set	442
396	to 2019, and to observations that reflected the total values across	443
397	all races. We also extracted the state name from the location	444
398	column to help us join this dataset to state-level datasets.	445
399	<b>Guttmacher Index</b> (project_political_data - Center_for_repro_rights.csv)	446
400	The Guttmacher index was provided at the state level. This	447
401	index examines a range of policies relating to abortion, selecting	448
402	approximately 20 types of abortion restrictions and 10 protective	449
403	policies. These include, but are not limited to, gestational duration	450
404	bans, waiting periods, insurance coverage bans, and medication	451
405	abortion restrictions, state constitutional protections, abortion	452
406	funding, insurance coverage for abortion, and protections for	453
407	patients and clinic staff. The states are then assigned to one of	454
408	the seven categories previously mentioned based on the policies	455
409	currently in effect (as of November 2024). <sup>§§</sup>	456
410	<b>GOP Trifecta</b> (ballotpedia_state_leg_scraper.ipynb) GOP tri-	457
411	fecta was provided on the state level. This dataset used election	458
412	results data scraped from ballotpedia (for state legislative races),	459
413	and gathered from Wikipedia (for gubernatorial races) to deter-	460
414	mine whether each state had a Republican government trifecta in	461
415	a given year.	462
416	Because these data sources reflected election results and not	463
417	governing control, we took the election results data and used a	464
418	lagged variable based on election year to code the years following	465
419	the election year as having had that party controlling the respective	466
420	office. Next, using a fact table of years*state as our base table,	467
421	we used the <i>pandasql</i> package to employ a SQL query so that we	468
422		469
423	<sup>§§</sup> Guttmacher Institute. "Methodology: State Policies in Brief." <a href="https://states.guttmacher.org/policies/methodology.html">https://states.guttmacher.org/policies/methodology.html</a> .	470

could do a range join over years and state. This, together with the election year lag variable, allowed us to fill in the years between elections with data on party control.

**Health Indicators** Each health indicator was downloadable as a CSV file from HRSA, and each of these was combined using a python script that cleaned and organized the data. Health indicators were provided at the county level.

When merging the data, some data sets had variables for full state name and some used state abbreviations. To address this we also joined a small states data set from a csv to make the connection between state name and abbreviation where needed.

## Analysis

We based our specifications on two theories:

1. Having a GOP trifecta in a state will increase maternal mortality rates, other things equal
2. More restrictive abortion policies will increase maternal mortality rates, other things equal

To test these theories, we ran regressions using Ordinary Least Squares (OLS).

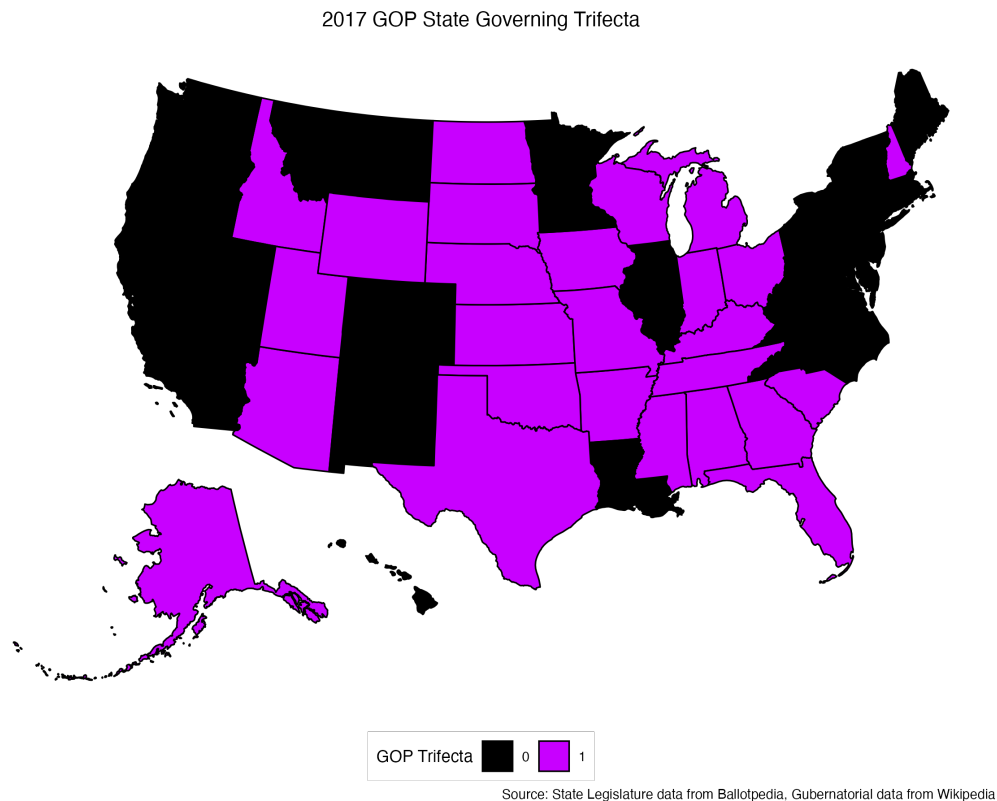


Fig. 2. GOP Trifecta

**GOP Trifecta.** We tested our first theory on the impact of

having a GOP trifecta at the state level using three specifications. First, we ran a regression with the effect of a GOP trifecta on maternal mortality rate. Next, we show standardized Maternal Mortality Rate (MMR) for easier interpretability, which was used for all subsequent measures of maternal mortality. Finally, we add health indicator covariates.

**For our first theory and specification**, we estimate the following Ordinary Least Squares (OLS) regression model to assess the relationship between maternal mortality rate ( $y$ ) and the GOP trifecta variable ( $X$ ):

$$y_i = \beta_0 + \beta_1 X_i + \epsilon_i$$

Where:

- $y_i$  is the maternal mortality rate (MMR rate) for observation  $i$ ,
- $X_i$  is the GOP trifecta variable for observation  $i$ ,
- $\beta_0$  is the intercept (constant),
- $\beta_1$  is the coefficient for the GOP trifecta variable,
- $\epsilon_i$  is the error term for observation  $i$ .

We add a constant term to the GOP trifecta variable matrix  $X$  before fitting the model:  $X = \text{sm.add.constant}(X)$

The model is then fitted using OLS, and the estimated coefficients are computed:

$$\hat{y} = \hat{\beta}_0 + \hat{\beta}_1 X$$

Where  $\hat{y}$  is the predicted maternal mortality rate, and  $\hat{\beta}_0$  and  $\hat{\beta}_1$  are the estimated coefficients from the regression.

**For our first theory, second specification**, we first standardize the Maternal Mortality Rate (MMR) by subtracting the mean of MMR and dividing by the standard deviation.

Next, we estimate the following Ordinary Least Squares (OLS) regression model to assess the effect of a GOP trifecta on the standardized MMR and add a constant to the GOP trifecta variable matrix  $X$  before fitting the model:  $X = \text{sm.add.constant}(X)$  The model is then fitted using OLS, and the estimated coefficients are computed as:

$$\hat{\text{MMR}}_{\text{standardized}} = \hat{\beta}_0 + \hat{\beta}_1 \text{GOP\_Trifecta}$$



Where  $\hat{\text{MMR}}_{\text{standardized}}$  represents the predicted standardized maternal mortality rate, and  $\hat{\beta}_0$  and  $\hat{\beta}_1$  are the estimated coefficients from the regression. Additionally, one standard deviation of MMR on the original scale corresponds to a value of:

$$\text{Deaths per 100,000} = \sigma_{\text{MMR}} \times 100,000$$

**In the third specification**, we control for various health indicators in addition to the GOP trifecta.

The covariates  $X$  include the GOP trifecta variable and several health indicators, which, to reiterate, are:

- *GOP\_Trifecta*
- *Obgyn\_Provider\_Rate\_100k*
- *Prenatal\_Care\_First\_Trimester\_Pct*
- *Obesity\_Pre\_Pregnancy\_Pct*
- *Diabetes\_Pre\_Pregnancy\_Pct*
- *Hypertension\_Pre\_Pregnancy\_Pct*

The full model, including an intercept  $\beta_0$ , is specified with the constant as

$$\hat{\text{MMR}}_{\text{standardized}} = \hat{\beta}_0 + \hat{\beta}_1 \text{GOP\_Trifecta} + \hat{\beta}_2 \text{Obgyn\_Provider\_Rate\_100k} + \hat{\beta}_3 \text{Prenatal\_Care\_First\_Trimester\_Pct} + \hat{\beta}_4 \text{Obesity\_Pre\_Pregnancy\_Pct} + \hat{\beta}_5 \text{Diabetes\_Pre\_Pregnancy\_Pct} + \hat{\beta}_6 \text{Hypertension\_Pre\_Pregnancy\_Pct}$$

Where  $\hat{\text{MMR}}_{\text{standardized}}$  is the predicted standardized maternal mortality rate, and  $\hat{\beta}_0, \hat{\beta}_1, \dots, \hat{\beta}_6$  are the estimated coefficients from the regression. The results are shown in Fig. 4.

**Guttmacher Index.** As we did with GOP trifecta theory, we tested our theory on the impact of abortion policies using three specifications. The first one looked at the comparative effect of states with the highest level of abortion restriction compared to the states most protective on maternal mortality rates. Next, we include a specification for all levels of abortion restriction that the index captures, providing a coefficient for each index. Finally, we looked again at the effect of all levels of abortion restriction on maternal mortality, this time controlling for health indicators.



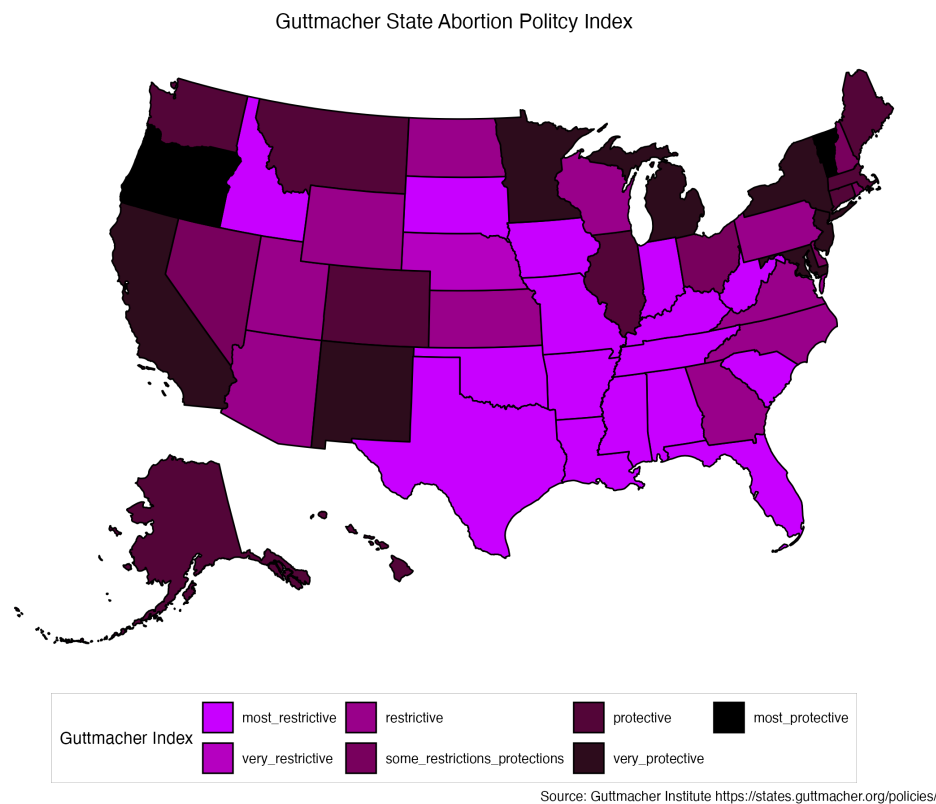


Fig. 3. Guttmacher Index

To handle missing values in the dataset, NA values were handled through column-wise operations to include as many observations as possible. All 50 states were represented, with states like Alaska and Louisiana retrofitted to conform with “county” columns (Alaska has “boroughs” and Louisiana has “parishes”). The resulting dataset produced 2,975 observations, encompassing 95 percent of the 3,144 counties in the Continental United States. Checks were performed to ensure the same number of observations were included in each regression.

**For our second theory** involving the Guttmacher Index, we examine the effect of abortion policy restrictions, specifically the Guttmacher “most restrictive” category, on standardized maternal mortality. The regression model we estimate is:

$$y_i = \beta_0 + \beta_1 \cdot \text{Guttmacher\_Most\_Restrictive}_i + \epsilon_i$$

Where:

- $y_i$  is the standardized maternal mortality rate for state  $i$ .
- $\beta_0$  is the intercept.
- $\beta_1$  is the coefficient for the Guttmacher “most restrictive” indicator.

- $\epsilon_i$  is the error term.

We estimate this model using Ordinary Least Squares (OLS).

**For the second specification**, we included all levels of abortion restrictions as explanatory variables. The model can be represented as:

$$\text{MMR}_i = \beta_0 + \beta_1 \cdot \text{Guttmacher}_{\text{most restrictive},i} + \beta_2 \cdot \text{Guttmacher}_{\text{very restrictive},i} + \beta_3 \cdot \text{Guttmacher}_{\text{restrictive},i} + \beta_4 \cdot \text{Guttmacher}_{\text{some restrictions protections},i} + \beta_5 \cdot \text{Guttmacher}_{\text{very protective},i} + \beta_6 \cdot \text{Guttmacher}_{\text{protective},i} + \epsilon_i$$

Where:

- $\text{MMR}_i$  is the standardized maternal mortality rate for state  $i$ ,
- $\text{Guttmacher}_{\text{most restrictive},i}$ ,  $\text{Guttmacher}_{\text{very restrictive},i}$ ,  $\text{Guttmacher}_{\text{restrictive},i}$ ,  $\text{Guttmacher}_{\text{some restrictions protections},i}$ ,  $\text{Guttmacher}_{\text{very protective},i}$ , and  $\text{Guttmacher}_{\text{protective},i}$  are dummy variables representing various levels of abortion restrictions in state  $i$ ,
- $\beta_0$  is the intercept, and
- $\epsilon_i$  is the error term.

**For our final and most comprehensive specification**, we include all levels of abortion restrictions and health indicators as explanatory variables. The model can be represented as:

$$\text{MMR}_i = \beta_0 + \beta_1 \cdot \text{Guttmacher}_{\text{most restrictive},i} + \beta_2 \cdot \text{Guttmacher}_{\text{very restrictive},i} + \beta_3 \cdot \text{Guttmacher}_{\text{restrictive},i} + \beta_4 \cdot \text{Guttmacher}_{\text{some restrictions protections},i} + \beta_5 \cdot \text{Guttmacher}_{\text{protective},i} + \beta_6 \cdot \text{Guttmacher}_{\text{very protective},i} + \beta_7 \cdot \text{ObGyn Provider Rate}_i + \beta_8 \cdot \text{Prenatal Care (First Trimester)}_i + \beta_9 \cdot \text{Obesity Pre-Pregnancy}_i + \beta_{10} \cdot \text{Diabetes Pre-Pregnancy}_i + \beta_{11} \cdot \text{Hypertension Pre-Pregnancy}_i + \epsilon_i$$

Where:

- $\text{MMR}_i$  is the standardized maternal mortality rate for state  $i$ ,
- $\text{Guttmacher}_{\text{most restrictive},i}$ ,  $\text{Guttmacher}_{\text{very restrictive},i}$ ,  $\text{Guttmacher}_{\text{restrictive},i}$ ,  $\text{Guttmacher}_{\text{some restrictions protections},i}$ ,  $\text{Guttmacher}_{\text{protective},i}$ , and  $\text{Guttmacher}_{\text{very protective},i}$  are dummy variables representing various levels of abortion restrictions in state  $i$ ,

- ObGyn Provider Rate<sub>*i*</sub> is the number of ob/gyn providers per 100,000 population in state *i*,
- Prenatal Care (First Trimester)<sub>*i*</sub> is the percentage of prenatal care initiated in the first trimester in state *i*,
- Obesity Pre-Pregnancy<sub>*i*</sub>, Diabetes Pre-Pregnancy<sub>*i*</sub>, and Hypertension Pre-Pregnancy<sub>*i*</sub> are the respective health indicators for state *i*,
- $\beta_0$  is the intercept, and
- $\epsilon_i$  is the error term.

**Standardization of MMR.** In order to improve readability of our coefficients, the regression was run on standard deviation of the maternal mortality rate column in our dataset. This was calculated using the z-score formula in pandas, with the raw MMR value subtracted from the mean of the MMR column and then dividing that result by the standard deviation of the column. In the interpretation of our data, a standard deviation equals roughly 1.36 deaths per 100K maternal deaths, therefore coefficients can be multiplied by this factor when interpreting results.

**Results.** Our regression results for GOP Trifectas are shown in the following figures:

Theory One: A GOP trifecta has a positive relationship with maternal mortality.

	(1)	(2)	(3)
const	0.000*** (0.000)	-0.350*** (0.030)	0.758*** (0.205)
diabetes_pre_pregnancy_pct			-0.192*** (0.037)
gop_trifecta	0.000*** (0.000)	0.533*** (0.037)	0.316*** (0.029)
hypertension_pre_pregnancy_pct			0.279*** (0.015)
obesity_pre_pregnancy_pct			0.068*** (0.003)
obgyn_provider_rate_100k			0.013*** (0.001)
prenatal_care_first_trimester_pct			-0.049*** (0.002)
Observations	2975	2975	2975
R <sup>2</sup>	0.064	0.064	0.468
Adjusted R <sup>2</sup>	0.064	0.064	0.467
Residual Std. Error	0.000 (df=2973)	0.968 (df=2973)	0.730 (df=2968)
F Statistic	203.159*** (df=1; 2973)	203.159*** (df=1; 2973)	434.932*** (df=6; 2968)
Note:	*p<0.1; **p<0.05; ***p<0.01		

Fig. 4. Theory One

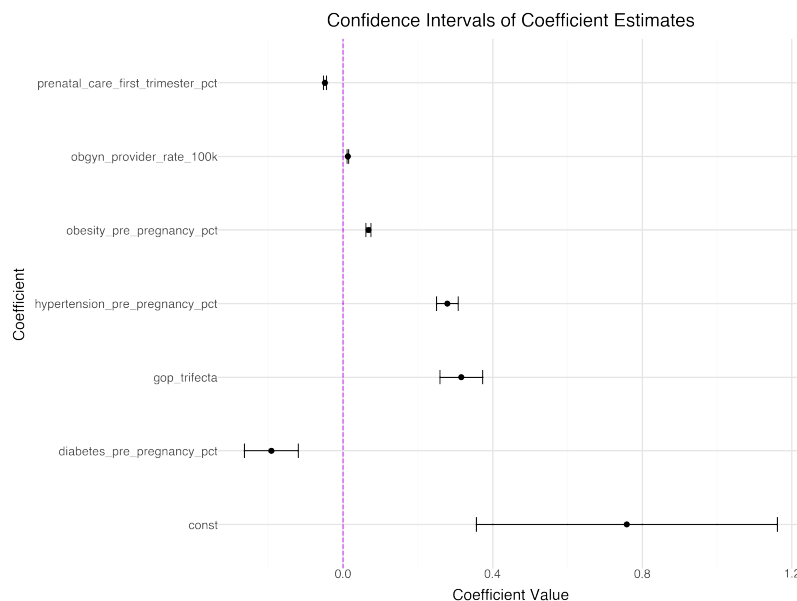


Fig. 5. Theory One Confidence Intervals

**The Effect of a GOP Trifecta on MMR.** Holding everything else constant, the presence of a GOP Trifecta seems to have a positive relationship with maternal mortality rates in a given state. All coefficients were statistically significant. The magnitude of the actual coefficient, as previously mentioned, is extremely small, making the case that a standardized variable will improve interpretability.

**The Effect of a GOP Trifecta on MMR (standardized).** After standardizing the maternal mortality figure, our model indicates that having a GOP trifecta is associated with an increase of 0.533 standard deviations in MMR (about a 0.71 increase in deaths per 100K). All coefficients were statistically significant.

**The Effect of a GOP Trifecta on MMR with added covariates (health indicators).** Continuing to use the standardized maternal mortality rates, we found in this specification that GOP Trifecta continue to be positively correlated with maternal mortality, to a lesser degree. The presence of a GOP Trifecta continues to have more influence than other covariates. OBGYN Provider Rate and Prenatal Care Percentage had a small negative effect on Maternal Mortality, suggesting that access to healthcare reduces the overall rate of maternal death, in keeping with conventional wisdom. A result worth noting was the coefficient associated with Pre-Pregnancy Diabetes, which appeared to suggest that higher diabetes rates are associated with a lower MMR. Pre-Pregnancy Hypertension provided a strong positive association with MMR, influencing maternal death rates almost as much as a GOP Trifecta.

Our results for the Guttmacher Index regressions are as follows:

Theory Two: The highest level of abortion restriction corresponds with the highest level of maternal mortality in a state

Dependent variable: standardized_mmr			
	(1)	(2)	(3)
const	-0.249*** (0.024)	-1.062*** (0.131)	0.327 (0.235)
diabetes_pre_pregnancy_pct			-0.208*** (0.037)
guttmacher_most_restrictive	0.534*** (0.035)	1.348*** (0.133)	0.786*** (0.106)
guttmacher_protective		0.533*** (0.141)	0.496*** (0.111)
guttmacher_restrictive		1.214*** (0.136)	0.876*** (0.107)
guttmacher_some_restrictions_protections		0.782*** (0.156)	0.421*** (0.122)
guttmacher_very_protective		0.446*** (0.140)	0.384*** (0.109)
guttmacher_very_restrictive		0.663*** (0.173)	0.821*** (0.136)
hypertension_pre_pregnancy_pct			0.262*** (0.015)
obesity_pre_pregnancy_pct			0.066*** (0.003)
obgyn_provider_rate_100k			0.013*** (0.001)
prenatal_care_first_trimester_pct			-0.048*** (0.002)
Observations	2975	2975	2975
R <sup>2</sup>	0.071	0.144	0.483
Adjusted R <sup>2</sup>	0.071	0.142	0.481
Residual Std. Error	0.964 (df=2973)	0.926 (df=2968)	0.721 (df=2963)
F Statistic	227.033*** (df=1; 2973)	83.198*** (df=6; 2968)	251.195*** (df=11; 2963)

Fig. 6. Theory Two

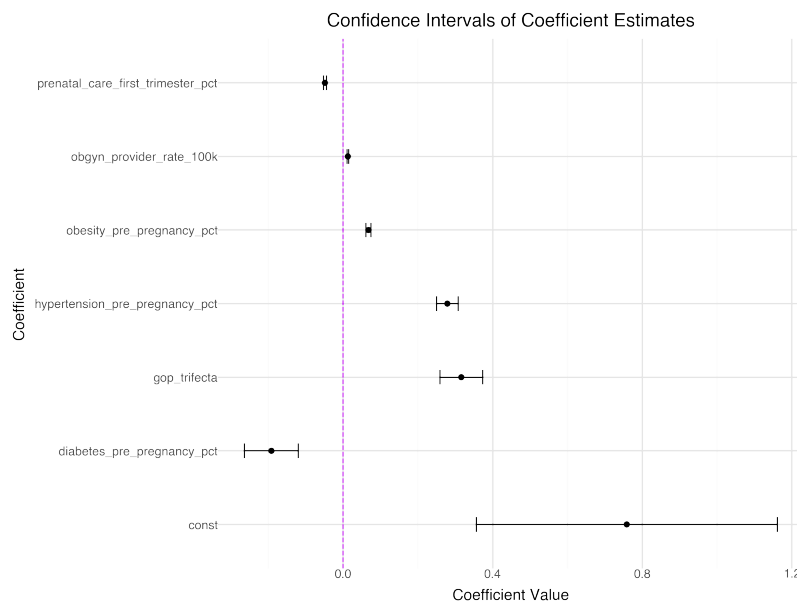


Fig. 7. Theory Two Confidence Intervals

### The Effect of Abortion Restriction Level on MMR.

When compared to states that had less stringent abortion restrictions, pregnancy in a state with the Most Restrictive category led to a 0.533 standard deviation increase in maternal mortality rates- the exact same value as a GOP trifecta (associated with an increase of 0.71 deaths per 100K). This coefficient is almost identical to the first specification of GOP Trifecta vs. MMR. The intercept was -0.249 (or 0.33 fewer deaths per 100K), which indicates a negative correlation between being in a state that did not fall under the “Most Restrictive” category. This result suggests that policies that restrict abortion are correlated with an increase in maternal mortality. All coefficients were statistically significant.

**The Effect of Abortion Restriction Level on MMR (all levels).** In this specification, the largest effect came from “Most Restrictive” states, in keeping with results from the first Guttmacher Index regression. Although higher in magnitude in terms of restriction levels, the coefficient for “Very Restrictive” had a much lower effect (0.663) than the coefficient for “Restrictive” (1.213). This result raises questions about the differentiation between categories and possible uncaptured confounding variables. The remainder of the restriction levels are consistent with the hypothesis that higher levels of abortion restriction lead to higher maternal mortality rates.

**The Effect of Abortion Restriction Level on MMR (all levels), with added covariates (health indicators).** When controlling for health indicators, the general narrative appears to follow the pattern that the more restrictive a state’s policies are,

1317 the higher the maternal mortality rate is likely to be. Diabetes 1364  
1318 Pre-Pregnancy continues to suggest, against conventional wisdom, 1365  
1319 a decrease in MMR. All other covariates behave similarly to the 1366  
1320 third GOP Trifecta specification. 1367

1321 1368

1322 **Discussion** 1369

1323 Overall, the presence of a GOP Trifecta and increasing levels of 1370  
1324 abortion restriction correlated with higher maternal mortality 1371  
1325 rates, including results calculated after adding health indicator 1372  
1326 covariates for expectant mothers. 1373

1327 1374

1328 **Diabetes Pre-Pregnancy as an Indicator.** The unexpected result 1375  
1329 of diabetes decreasing the incidence of maternal mortality could be 1376  
1330 associated with the several confounding factors. Among these are 1377  
1331 access to quality healthcare, under-reporting among populations 1378  
1332 with lower rates of healthcare access, and interactions with other 1379  
1333 health indicators. Women with diabetes might receive more 1380  
1334 frequent and specialized care during pregnancy, such as high- 1381  
1335 risk monitoring or early screening for complications, potentially 1382  
1336 mitigating risks associated with diabetes. More likely, the result 1383  
1337 is due to under-reporting among populations with less healthcare 1384  
1338 access. Correcting for socioeconomic indicators may improve 1385  
1339 this result. Interactions with other health indicators are also a 1386  
1340 possibility. Women with diabetes may be more likely to have 1387  
1341 blood pressure and obesity issues, which can cause an interaction 1388  
1342 of the variables and produce unexpected results. 1389

1343 1390

1344 **Statistical Insignificance of “Most Protections” Category.** This 1391  
1345 result could potentially be explained by the small number of 1392  
1346 observations in this category. Only three states were classified as 1393  
1347 “Most Protections” states. Those included Maryland, Vermont, 1394  
1348 and Oregon. There were potentially not enough data points for 1395  
1349 the result to be statistically significant, due to the low number of 1396  
1350 counties in these states (69 total across all three, meaning only 69 1397  
1351 observations). 1398

1352 1399

1353 **Implications and Importance of This Dataset and Research.** Our 1400  
1354 most significant goal for this project was to create a holistic dataset 1401  
1355 based on available data. Any effort to gather health data generates 1402  
1356 privacy concerns, and while these are valid considerations, a 1403  
1357 higher volume of more detailed data could bolster efforts to reduce 1404  
1358 alarming maternal mortality rates. Research continues to show a 1405  
1359 relationship between maternal mortality and abortion restricting 1406  
1360 policies. Maternal deaths after Roe v. Wade was passed in 1973 1407  
1361 decreased 30-40 percent for people of color.<sup>14</sup> Without important 1408

1362 <sup>14</sup>Kheyfets, Anna, Shubhecchha Dhaurali, Paige Feyock, Farinaz Khan, April Lockley, Brenna Miller, Lauren Cohen, Eimaan Anwar, and Ndidiama Amutah-Onukagha. "In 2021, Over 90 Restrictive Abortion 1409

1363 Laws Were Passed." *Frontiers in Public Health*. National Library of Medicine. <https://pmc.ncbi.nlm.nih.gov/articles/PMC10728320/>. 1410



1411 figures and information such as those provided here, lawmakers  
1412 will not be fully equipped to make informed policy decisions, nor  
1413 will voters be as informed.

1414 Several states want to discontinue examining this type of data.  
1415 Each state has a Maternal Mortality Committee comprised of  
1416 medical professionals who publish an annual report attempting  
1417 to determine if any deaths were preventable. Some committees  
1418 provide recommendations based on demographics. The Texas  
1419 Maternal Mortality and Morbidity Review Committee announced  
1420 in November of 2024 that they would not review cases from  
1421 2022 and 2023, which are years immediately following the imple-  
1422 mentation of laws that placed Texas in the “Most Restrictive”  
1423 Guttmacher Index category.<sup>\*\*\*</sup> Following suit, Georgia’s state  
1424 public health commissioner dismissed all members of the state’s  
1425 maternal mortality committee, a state that banned abortions after  
1426 six weeks. In 2023, Idaho’s maternal mortality committee was  
1427 dissolved completely and then re-convened this year, missing an  
1428 entire year of analysis for deaths post-roe<sup>††</sup>.

1429 Our results show the importance of accurate and organized data.  
1430 Without the ability to attribute deaths to abortion restrictions,  
1431 establishing a causal relationship between abortion restrictions  
1432 and a higher rate of maternal mortality becomes a challenge.  
1433 With our dataset, we were unable to make the claim that abortion  
1434 restrictions impact maternal mortality- we were only able to  
1435 establish that states with higher levels of restrictions and a  
1436 GOP trifecta correlated with higher maternal mortality rates,  
1437 taking into account other health indicators. In this context,  
1438 maternal mortality committees become exceedingly important  
1439 in determining the causes of death in these cases that were not  
1440 due to specific health issues. The limitations of our results may  
1441 be more informative than the results themselves, as it is apparent  
1442 that there is a gap in data collection.

1443  
1444 **Future Considerations.** Understanding of the relationship be-  
1445 tween maternal mortality rates and restrictive abortion policies,  
1446 typically attributed to GOP control, will allow lawmakers to  
1447 make more informed and nuanced decisions about health-care  
1448 policy in their states. These decisions are not limited to policies  
1449 related to elective abortions, but can affect healthcare of mothers  
1450 who want to have children and miscarry at some point in their  
1451 pregnancy. When a miscarriage occurs, doctors often perform a  
1452 DnC, or dilation and curettage, which is categorized as abortion  
1453 care. An article published in December of 2024 outlines a situation

1454\* Kitchener, Caroline. “Texas Committee Won’t Examine Maternal Deaths in First Years after Abortion Ban.” The Washington Post, November 26, 2024.  
1455 <https://www.washingtonpost.com/investigations/2024/11/26/texas-committee-wont-examine-maternal-deaths-first-years-after-abortion-ban/>.

1456††† Corbin, Clark. “Reconstituted Idaho Maternal Mortality Committee Will Release a New Report by Jan. 31.” Idaho Capitol Sun, 13 Dec. 2024, [www.idahocapitalsun.com/2024/12/13/reconstituted-idaho-maternal-mortality-committee-will-release-a-new-report-by-jan-31/](http://www.idahocapitalsun.com/2024/12/13/reconstituted-idaho-maternal-mortality-committee-will-release-a-new-report-by-jan-31/).

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in which a female officer in the military nearly lost her life when

TRICARE (the military’s health insurance) refused to pay for

this procedure.<sup>†††</sup> While this particular case is more likely due

to Hyde Amendment restrictions, it illustrates how much of an

impact similar policies have on the lives of prospective mothers

outside of the military.

**Further Analysis.** Beyond obtaining more accurate and complete

datasets, many more regressions and analyses can be conducted

from the data we do have. Accounting for lower income and

rural populations could help explain more of the data and refine

relationships between independent and dependent variables in our

model. Subsetting the data by race could also help clarify the

disparities between populations, which are known to exist. A study

conducted by researchers in 2023 at Southern Connecticut State

University found that black women have a maternal mortality rate

of 2.9 times that of White women.<sup>§§§</sup> Addressing these disparities

in future analyses is essential in gaining a complete understanding

of the issue.

**Further Data-Collection.** Shifting to gaps in existing data, the

impact of COVID-19, incidence of home-births, maternal age,

and documentation of wanted vs. unwanted pregnancies are data

points of interest that don’t exist yet. As with all other health

data phenomena, COVID-19 is an anomaly that may be difficult

to address without careful consideration. The incidence of home-

births and deaths due to COVID itself are factors that could

make causal relationships more difficult to establish. Maternal

age could be easier to establish a relationship with, however it

also could be considered a privacy issue in the quest to obtain it.

Another consideration in examining the impact certain policies

have on maternal mortality is the number of pregnancies that were

wanted vs those that were unwanted. Women who were forced

to carry unwanted pregnancies to term due to state restrictions

and then die as a result of their pregnancy have a causal link to

these policies compared to women who suffer complications due to

miscarriages or difficult births. Without sufficient data collected

on the sentiment of women’s pregnancies, the connection remains

anecdotal and ambiguous.

**Conclusion**

In closing, while there appears to be a correlation between mater-

nal mortality, GOP leadership and restrictive healthcare policies,

<sup>†††</sup> Edwards, Erin, and Robin Fields. "A Coast Guard Commander Miscarried. She Nearly Died After Being Denied Care." ProPublica, December 13, 2023.<https://www.propublica.org/article/elizabeth-nakagawa-miscarriage-military-tricare-abortion-policy>

<sup>§§§</sup> Njoku, Anuli, Marian Evans, Lillian Nimo-Sefah, and Jonell Bailey. "Listen to the Whispers before They Become Screams: Addressing Black Maternal Morbidity and Mortality in the United States." Healthcare 11, no. 3 (2023): 438. <https://doi.org/10.3390/healthcare11030438>.

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more research needs to be done to make a robust causal link to maternal mortality rates especially deaths relating to denial of abortion care. The combination of disparate data sources has allowed us to draw this conclusion among others. This analysis requires a more focused approach in data-gathering, especially in states that are characterized as having the “Most-Restrictive” policies. Researchers ought to be provided with the most up-to-date datasets to conduct the research, and lawmakers with the most up-to-date analyses to allow for more informed policy decisions.

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