

AI1103 ASSIGNMENT 4

Name: MANNAM SARANDEEP, Rollno: CS20BTECH11030

Download the python code from

<https://github.com/sarandeepmannam/ASSIGNMENT4/blob/main/Assignment4.py>

and latex-tikz code from

<https://github.com/sarandeepmannam/ASSIGNMENT4/blob/main/Assignment4.tex>

1 QUESTION-CSIR UGC NET JUNE 2012,Q.50

Let X_1, X_2, \dots be i.i.d $N(1,1)$ random variables. Let $S_n = X_1^2 + X_2^2 + \dots + X_n^2$ for $n \geq 1$. Then

$$\lim_{n \rightarrow \infty} \frac{\text{Var}(S_n)}{n} =$$

- (A) 4
- (B) 6
- (C) 1
- (D) 0

2 SOLUTION-CSIR UGC NET JUNE 2012,Q.50

Definition 1 (NON-CENTRAL CHI SQUARE DISTRIBUTION). Let $X_1, X_2, X_3, \dots, X_i, \dots, X_n$ be n independent, normally distributed random variables with means μ_i and unit variances. Then the random variable

$$\sum_{i=1}^n X_i$$

is distributed according to the non-central chi square distribution. It has two parameters 'k' which specifies the number of degrees of freedom (i.e. the number of X_i), and ' λ ' which is called non-centrality parameter given by,

$$\lambda = \sum_{i=1}^n \mu_i^2 \quad (2.0.1)$$

From definition 1, Since X_1, X_2, \dots are i.i.d $N(1,1)$ normal random variables therefore the distribution of the random variable $S_n = X_1^2 + X_2^2 + \dots + X_n^2$ is a non-central chi square distribution with 'n' degrees

of freedom and non-centrality parameter ' λ ' given by

$$\lambda = \sum_{i=1}^n (\mu_i)^2 \quad (2.0.2)$$

$$\lambda = \sum_{i=1}^n 1 = n \quad (2.0.3)$$

Lemma 2.1. Moment generating function of a non-central chi square distributed random variable X is given by,

$$M_X(t) = \frac{e^{\frac{\lambda t}{1-2t}}}{(1-2t)^{\frac{n}{2}}} \quad (2.0.4)$$

From Lemma 2.1 Moment generating function of S_n is given by,

$$M_{S_n}(t) = \frac{e^{\frac{nt}{1-2t}}}{(1-2t)^{\frac{n}{2}}} \quad (2.0.5)$$

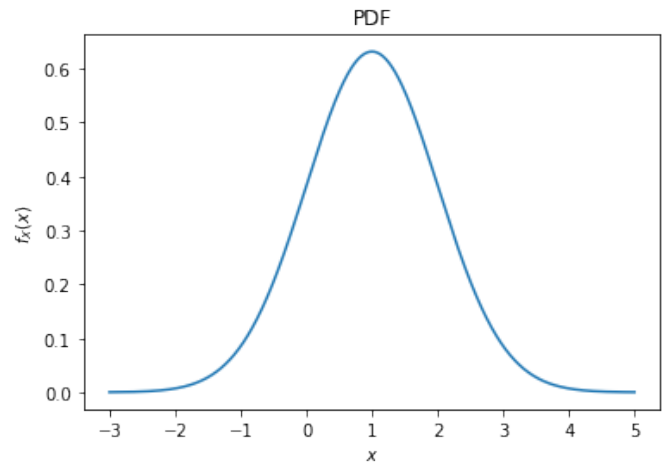


Fig. 4: PDF of X_1, X_2, \dots

Lemma 2.2. The n th moment of a random variable X whose MGF is $M_X(t) (\equiv E(e^{tX}) = \sum_{k=0}^{\infty} \frac{(tX)^k}{k!})$ is given by,

$$E(X^n) = \left(\frac{d^n}{dt^n} (M_X(t)) \right)_{t=0} \quad (2.0.6)$$

Proof.

$$M_X(t) \equiv E(e^{tX}) \quad (2.0.7)$$

Using Taylor series expansion,

$$M_X(t) = E\left(1 + \frac{tX}{1!} + \frac{(tX)^2}{2!} + \frac{(tX)^3}{3!} + \dots + \frac{(tX)^n}{n!} + \dots\right) \quad (2.0.8)$$

Taking n th derivative on both sides with respect to 't' at $t = 0$ we get,

$$\begin{aligned} \left(\frac{d^n}{dt^n} (M_X(t)) \right)_{t=0} &= (E(X^n))_{t=0} + \left(tE\left(\frac{X^{n+1}}{n+1}\right) \right)_{t=0} \\ &+ \left(t^2 E\left(\frac{X^{n+2}}{(n+1)(n+2)}\right) \right)_{t=0} + \dots \end{aligned} \quad (2.0.9)$$

$$\left(\frac{d^n}{dt^n} (M_X(t)) \right)_{t=0} = E(X^n) \quad (2.0.10)$$

□

From lemma 2.2 ,

$$E(S_n) = \left(\frac{d}{dt} (M_{S_n}(t)) \right)_{t=0} \quad (2.0.11)$$

$$E(S_n) = \left(\frac{2ne^{\frac{nt}{1-2t}}}{(1-2t)^{\frac{n}{2}+2}} (1-t) \right)_{t=0} \quad (2.0.12)$$

$$\implies E(S_n) = 2n \quad (2.0.13)$$

From lemma 2.2,

$$E(S_n^2) = \left(\frac{d^2}{dt^2} (M_{S_n}(t)) \right)_{t=0} \quad (2.0.14)$$

$$\begin{aligned} E(S_n^2) &= \frac{2ne^{\frac{nt}{1-2t}}}{(1-2t)^{\frac{n}{2}+4}} (n(1-t) - (1-2t)^2 \\ &+ (1-t)(1-2t)(n+4))|_{t=0} \end{aligned} \quad (2.0.15)$$

$$E(S_n^2) = \frac{2n}{1^{n/2+2}} (n - 1^2 + n + 4) \quad (2.0.16)$$

$$E(S_n^2) = 2n \times (2n + 3) \quad (2.0.17)$$

$$\implies E(S_n^2) = 4n^2 + 6n \quad (2.0.18)$$

Variance of S_n ,

$$\text{Var}(S_n) = E(S_n^2) - (E(S_n))^2 \quad (2.0.19)$$

Substituting equations (2.0.13) and (2.0.18) in (2.0.19), we get

$$\text{Var}(S_n) = 4n^2 + 6n - (2n)^2 \quad (2.0.20)$$

$$\implies \text{Var}(S_n) = 6n \quad (2.0.21)$$

and

$$\lim_{n \rightarrow \infty} \frac{\text{Var}(S_n)}{n} = \lim_{n \rightarrow \infty} \frac{6n}{n} = \lim_{n \rightarrow \infty} 6 = 6$$

Hence, option (B) is correct.