TIME SERIES AUTO REGRESSION AND ERROR CORRECTION MODELS

# Multi-Variate Econometric Model Configuration

Vector auto regression, Volatility, Granger causality & Error Correction



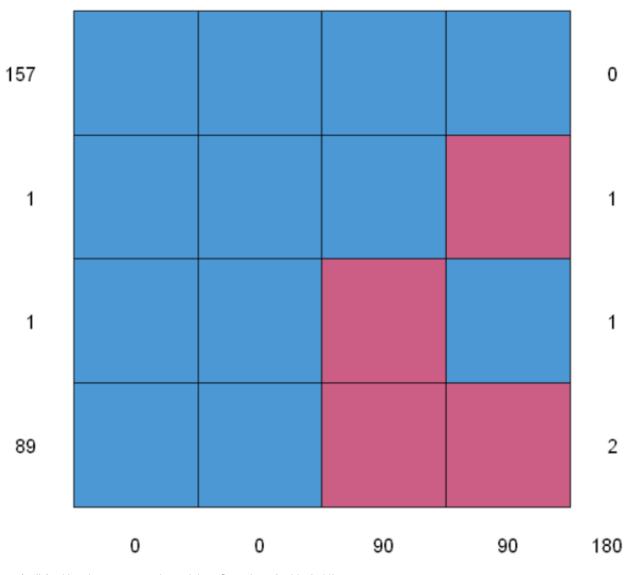


Image by author

ector auto regression (VAR) to first difference generally creates integrated timeseries (TS) models. But we may eliminate valuable information about the relationship among variables by differencing, where Vector Error Correction model (VECM) is applicable.

### Granger causality:

VAR involves multiple exog variables which are important to predict future state of endog variable. Using Granger causality (GC) we can determine the importance of multiple variables and GC is only relevant with TS variables. We will use VAR to investigate GC here.

Here our use case is that, we have data of Western Texas Intermediate, Brent Crude oil and HenryHub Spot price and we shall forecast future 15 time steps of each. We shall use R program to solve this. Let us load the data:

```
wti = as_tibble(get_fred_series("DCOILWTICO", "WTI",
  observation_start = "2000-01-02"))
brent = as_tibble(get_fred_series("DCOILBRENTEU", "Brent",
  observation_start = "2000-01-02"))
hh = as_tibble(get_fred_series("MHHNGSP", 'HenryHub',
  observation_start = "2000-01-02"))
oil_prices = wti %>% left_join(brent, by="date")
prices = hh %>% left_join(oil_prices, by="date")
head(prices)
```

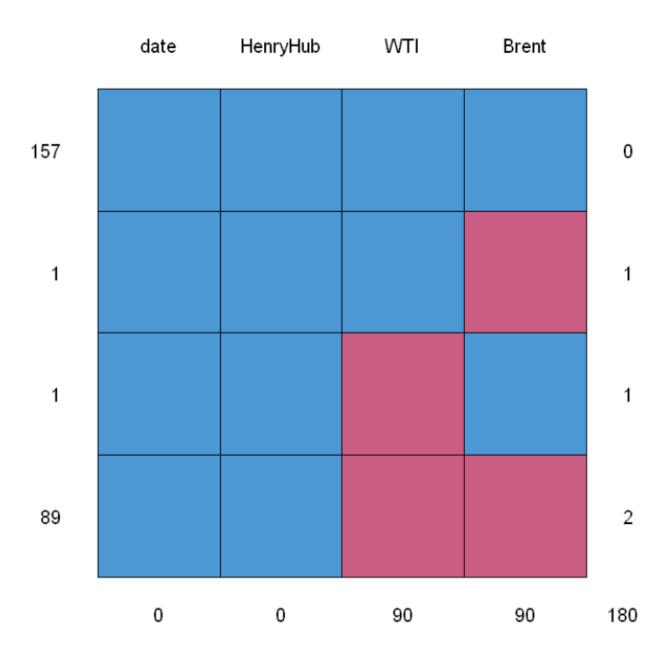
date	HenryHub	WTI	Brent
2000-02-01	2.66	28.28	27.35
2000-03-01	2.79	31.71	29.78
2000-04-01	3.04	NA	NA
2000-05-01	3.59	25.84	NA
2000-06-01	4.29	30.19	29.69
2000-07-01	3.99	NA	NA

# Pre-processing data:

\*

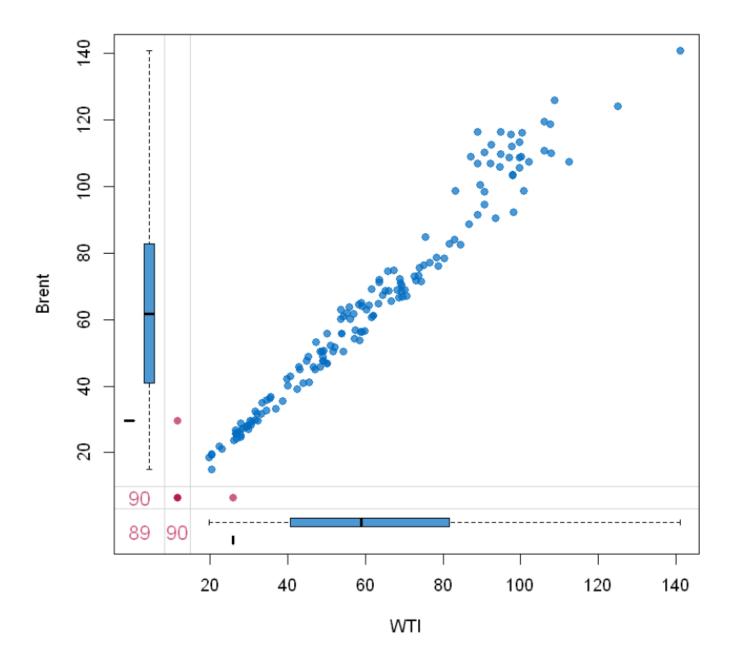
# #understand the missing value pattern md.pattern(prices)

	date	HenryHub	WTI	Brent	
157	1	1	1	1	0
1	1	1	1	0	1
1	1	1	0	1	1
89	1	1	0	0	2
	0	0	90	90	180



# Margin plot to visualize missing values:

```
marginplot(prices[, c("WTI", "Brent")], col = mdc(1:2), cex.numbers =
1.2, pch = 19)
```



# Imputing missing values:

```
imputes = mice(prices, m=5, maxit = 40)
# methods used for imputing
imputes
```

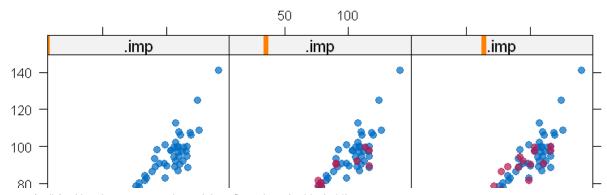
```
Class: mids
Number of multiple imputations:
Imputation methods:
    date HenryHub
                                Brent
                         WTI
                       "pmm"
                                "pmm"
PredictorMatrix:
         date HenryHub WTI Brent
date
             0
                       1
                           1
                                  1
HenryHub
             1
                       0
                           1
                                  1
WTI
                                  1
             1
                       1
                           0
Brent
             1
                       1
                           1
                                  0
```

```
#Imputed dataset
data = complete(imputes,5)
head(data)
```

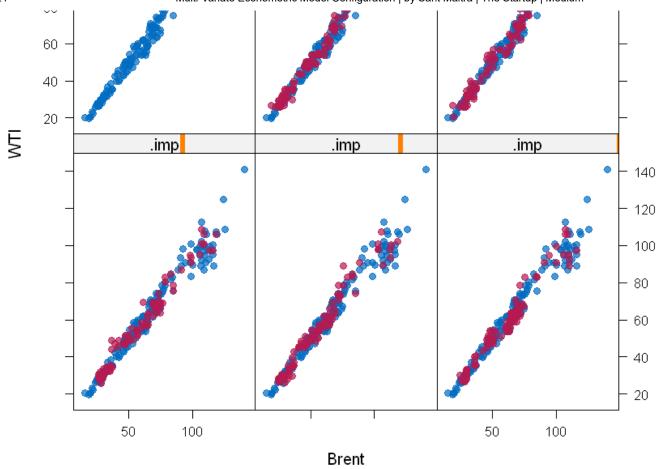
date	HenryHub	WTI	Brent
2000-02-01	2.66	28.28	27.35
2000-03-01	2.79	31.71	29.78
2000-04-01	3.04	29.89	29.42
2000-05-01	3.59	25.84	24.76
2000-06-01	4.29	30.19	29.69
2000-07-01	3.99	58.91	66.79

# Goodness of fit:

```
xyplot(imputes, WTI ~ Brent | .imp, pch = 20, cex = 1.4)
```

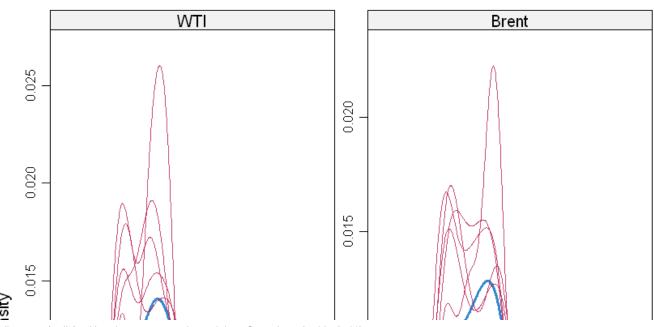




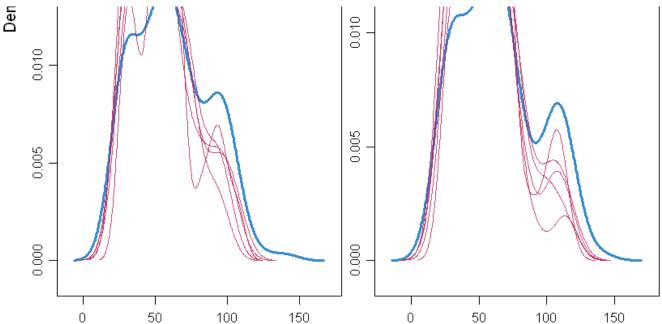


# **Density plot:**

densityplot(imputes)

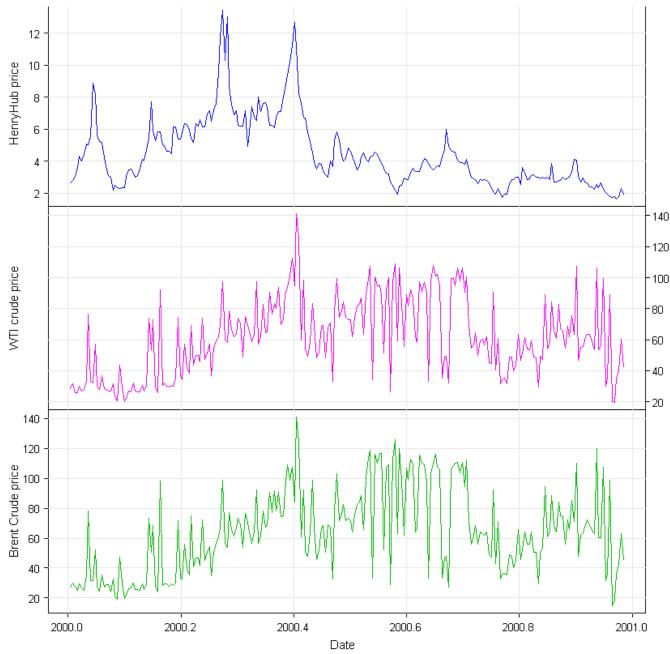






# Time-series & Line plots:

```
# converting to time -series
henryhub = ts(data$HenryHub, start = c(2000,02,01), frequency = 252)
wti = ts(data\$WII, start = c(2000,02,01), frequency = 252)
brent = ts(data\$Brent, start = c(2000,02,01), frequency = 252)
# time series line plot
par(mfrow=c(3,1), mar=c(0,3.5,0,3), oma=c(3.5,0,2,0), mgp=c(2,.6,0),
cex.lab=1.1, tcl=-.3, las=1)
plot(henryhub, ylab=expression('HenryHub price'), xaxt="no",
type='n')
  grid(lty=1, col=gray(.9))
  lines (henryhub, col=rgb(0,0,.9))
plot(wti, ylab=expression('WTI crude price'), xaxt="no", yaxt='no',
type='n')
  grid(lty=1, col=gray(.9))
  lines(wti, col=rgb(.9,0,.9))
  axis(4)
plot(brent, ylab=expression('Brent Crude price'))
  grid(lty=1, col=gray(.9))
  lines (brent, col=rgb (0, .7, 0))
title(xlab="Date", outer=TRUE)
df = cbind(henryhub, wti, brent)
```



# **ADF Unit root test**

Thumb rule: If calculated statistics > tabulated values, null hyposes can be rejected.

- ADF test (H0: series has unit root)
- PP test (H0: series has unit root)
- KPSS test (H0: series has no unit root)
- Zivot & Andrew test (H0: series has unit root)

```
adf.test(log(data[, "HenryHub"]))
adf.test(log(data[, "WTI"]))
adf.test(log(data[, "Brent"]))
```

### Augmented Dickey-Fuller Test

```
data: log(data[, "HenryHub"])
Dickey-Fuller = -3.4604, Lag order = 6, p-value = 0.04712
alternative hypothesis: stationary
```

### Augmented Dickey-Fuller Test

```
data: log(data[, "WTI"])
Dickey-Fuller = -2.5962, Lag order = 6, p-value = 0.3249
alternative hypothesis: stationary
```

### Augmented Dickey-Fuller Test

```
data: log(data[, "Brent"])
Dickey-Fuller = -2.4523, Lag order = 6, p-value = 0.3854
alternative hypothesis: stationary
```

```
pp.test(log(data[, "HenryHub"]), type = "Z(t_alpha)")
pp.test(log(data[, "WTI"]), type = "Z(t_alpha)")
pp.test(log(data[, "Brent"]), type = "Z(t_alpha)")
```

### Phillips-Perron Unit Root Test

```
data: log(data[, "HenryHub"])
Dickey-Fuller Z(t_alpha) = -3.4877, Truncation lag parameter = 5,
p-value = 0.04449
alternative hypothesis: stationary
Warning message in pp.test(log(data[, "WTI"]), type = "Z(t_alpha)"):
"p-value smaller than printed p-value"
```

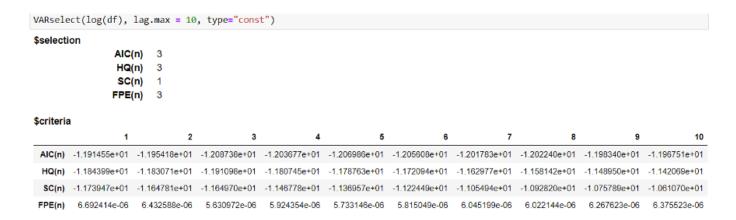
### Phillips-Perron Unit Root Test

```
data: log(data[, "WTI"])
Dickey-Fuller Z(t_alpha) = -9.248, Truncation lag parameter = 5,
p-value = 0.01
alternative hypothesis: stationary
Warning message in pp.test(log(data[, "Brent"]), type = "Z(t alpha)"):
"p-value smaller than printed p-value"
        Phillips-Perron Unit Root Test
data: log(data[, "Brent"])
Dickey-Fuller Z(t alpha) = -9.3893, Truncation lag parameter = 5,
p-value = 0.01
alternative hypothesis: stationary
  kpss.test(log(data[, "HenryHub"]))
  kpss.test(log(data[, "WTI"]))
  kpss.test(log(data[, "Brent"]))
 Warning message in kpss.test(log(data[, "HenryHub"])):
 "p-value smaller than printed p-value"
         KPSS Test for Level Stationarity
 data: log(data[, "HenryHub"])
 KPSS Level = 1.899, Truncation lag parameter = 5, p-value = 0.01
 Warning message in kpss.test(log(data[, "WTI"])):
 "p-value smaller than printed p-value"
         KPSS Test for Level Stationarity
 data: log(data[, "WTI"])
 KPSS Level = 1.3095, Truncation lag parameter = 5, p-value = 0.01
 Warning message in kpss.test(log(data[, "Brent"])):
 "p-value smaller than printed p-value"
```

KPSS Test for Level Stationarity

```
data: log(data[, "Brent"])
KPSS Level = 1.5159, Truncation lag parameter = 5, p-value = 0.01
```

### Number of lags



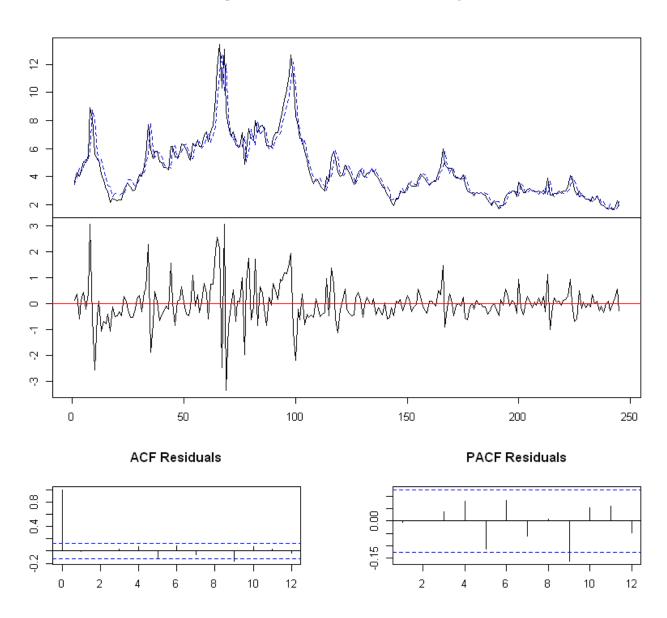
According to the AIC, HQ and FPE the optimal lag number is p=3, whereas the SC criterion indicates an optimal lag length of p=1. They estimated for all three lag orders a VAR including a constant and a trend as deterministic regressors and conducted diagnostic tests with respect to the residuals.

# **Estimating VAR Model:**

```
var_mod <- VAR(df, p = 3, type = "both")
summary(var_mod)
plot(var_mod, names = "HenryHub")</pre>
```

```
VAR Estimation Results:
Endogenous variables: HenryHub, WTI, Brent
Deterministic variables: both
Sample size: 245
Log Likelihood: -2012.016
Roots of the characteristic polynomial:
0.902 0.8221 0.7166 0.6837 0.6837 0.4529 0.4529 0.3309 0.2936
VAR(y = data, p = 3, type = "both")
Estimation results for equation HenryHub:
HenryHub = HenryHub.l1 + WTI.l1 + Brent.l1 + HenryHub.l2 + WTI.l2 + Brent.l2 + HenryHub.l3 + WTI.l3 + Brent.l3 + const + trend
             Estimate Std. Error t value Pr(>|t|)
HenryHub.l1 0.9222864 0.0655510 14.070 < 2e-16 ***
WTI.l1 -0.0011835 0.0143208 -0.083 0.93421
Brent.l1
          -0.0012326 0.0125694 -0.098 0.92197
HenryHub.l2 0.0490920 0.0897412
                                   0.547
                                           0.58487
WTI.12 0.0085102 0.0148313 0.574 0.56665
Brent.12 -0.0031738 0.0131398 -0.242 0.80935
HenryHub.l3 -0.0830773 0.0688170
                                  -1.207
           -A AA16739 A A1/219A -A 118 A 9A639
```

### Diagram of fit and residuals for HenryHub



# **Model Diagnostics:**

### Residuals:

Portmanteau goodness of-fit test to test the adequacy of the fitted model by checking whether the residuals are approximately white noise.

```
residuals = serial.test(var_mod, lags.pt=3, type="PT.asymptotic") # residuals
residuals$serial
```

```
Portmanteau Test (asymptotic)

data: Residuals of VAR object var_mod

Chi-squared = 5.4563, df = 0, p-value < 2.2e-16
```

The null hypothesis of no autocorrelation is rejected since the p-value < than the significance level of 0.05.

### Normality test:

```
norm <- normality.test(var mod)</pre>
norm$jb.mul
$JB
        JB-Test (multivariate)
       Residuals of VAR object var mod
Chi-squared = 246.88, df = 6, p-value < 2.2e-16
$Skewness
        Skewness only (multivariate)
       Residuals of VAR object var mod
Chi-squared = 9.9389, df = 3, p-value = 0.01909
$Kurtosis
        Kurtosis only (multivariate)
       Residuals of VAR object var mod
Chi-squared = 236.94, df = 3, p-value < 2.2e-16
```

- The null hypothesis of the Jarque-Bera test is a joint hypothesis of the skewness being zero and the excess kurtosis being zero
- The result of the p-value shows the null hypothesis is rejected.

• Thus, we can conclude residuals does not follow a normal distribution.

# Conditional volatility model:

```
arch <- arch.test(var_mod, lags.multi = 5, multivariate.only = TRUE)
arch$arch.mul</pre>
```

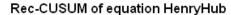
```
ARCH (multivariate)
```

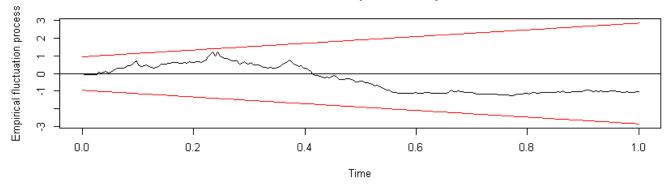
```
data: Residuals of VAR object var_mod
Chi-squared = 375.04, df = 180, p-value = 7.772e-16
```

Here p < 0.05, therefore, so, we can assume that model suffers from heteroscedasticity.

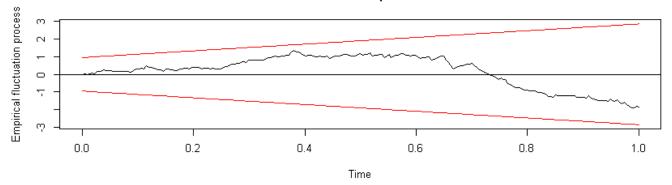
# Testing structural breaks:

```
## Stability
plot(stability(var_mod, type = "Rec-CUSUM"))
```

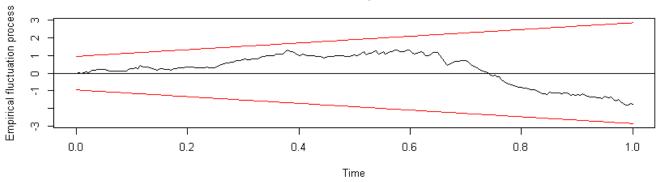




#### Rec-CUSUM of equation WTI



#### Rec-CUSUM of equation Brent



Here we find the system is stable because line plots do not exceed the red lines.

### **Granger Causality:**

```
causality(var mod, cause = 'wti')
causality(var mod, cause = 'brent')
causality(var mod, cause = 'henryhub')
$Granger
Granger causality HO: WTI do not Granger-cause HenryHub Brent
data: VAR object var mod
F-Test = 1.4459, df1 = 6, df2 = 702, p-value = 0.1944
$Instant
HO: No instantaneous causality between: WTI and HenryHub Brent
data: VAR object var mod
Chi-squared = 120.3, df = 2, p-value < 2.2e-16
$Granger
Granger causality HO: Brent do not Granger-cause HenryHub WTI
data: VAR object var mod
F-Test = 1.1576, df1 = 6, df2 = 702, p-value = 0.3273
$Instant
HO: No instantaneous causality between: Brent and HenryHub WTI
data: VAR object var mod
Chi-squared = 120.29, df = 2, p-value < 2.2e-16
```

```
$Granger

Granger causality H0: HenryHub do not Granger-cause WTI Brent

data: VAR object var_mod
F-Test = 6.4644, df1 = 6, df2 = 702, p-value = 1.197e-06

$Instant

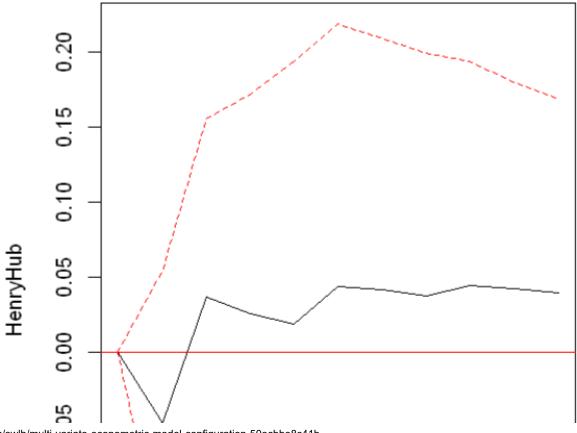
H0: No instantaneous causality between: HenryHub and WTI Brent

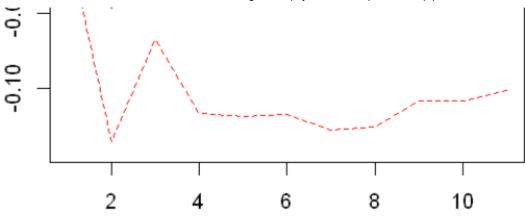
data: VAR object var_mod
Chi-squared = 3.9472, df = 2, p-value = 0.139
```

# Impulse response:

```
## IRF
plot(irf(var_mod, impulse = "wti", response = c("henryhub"),
n_ahead=15, boot = TRUE))
```

# Orthogonal Impulse Response from WTI



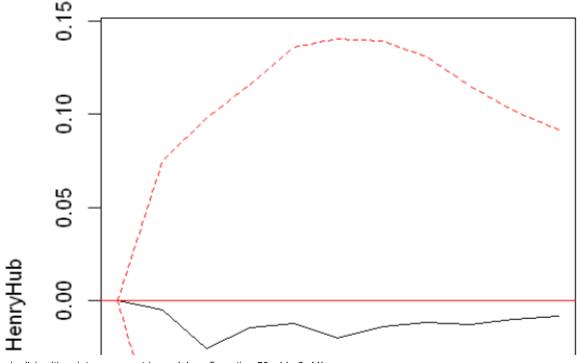


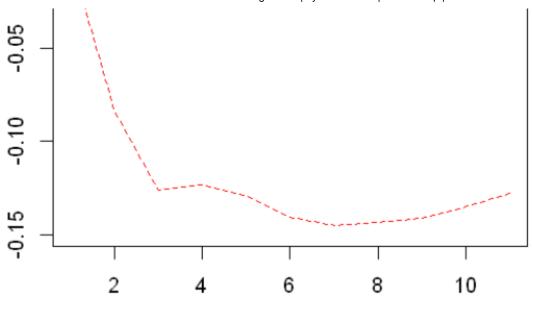
95 % Bootstrap CI, 100 runs

Confidence interval span is quite big indicating we have relatively bigger room for error; so, to effect there could be no effect between WTI & HenryHub; however, on the basis of point estimate, WTI will have positive effect on HenryHub.

```
## IRF
plot(irf(var_mod, impulse = "Brent", response = c("HenryHub"),
n_ahead=15, boot = TRUE))
```

# Orthogonal Impulse Response from Brent





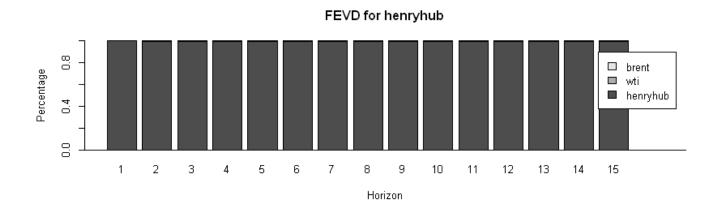
95 % Bootstrap CI, 100 runs

here too, we find there is no impact of Brent on HenryHub; we have a relatively bigger room for error; however, Brent having negative effect on HenryHub.

# Forecast error variance decomposition:

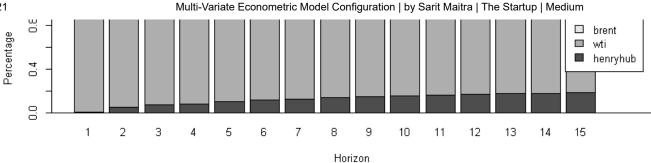
Forecast Error Variance Decomposition examines the impact of variables on one another.

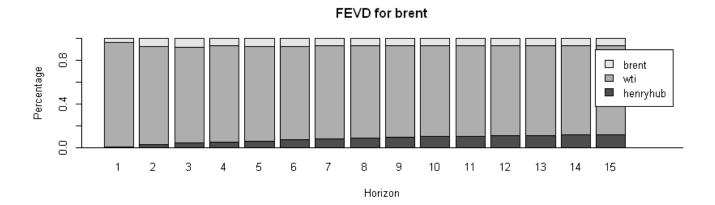
```
plot(fevd(var_mod, n.ahead = 15))
```



FEVD for wti







FEVD estimates the contribution of a shock in each variable to the response in both variables. Almost 100 % of the variance in HenryHub is caused by HenryHub itself, while only about 80 % in the variance of WTI and Brent caused by themselves and others.

### **Forecast VAR:**

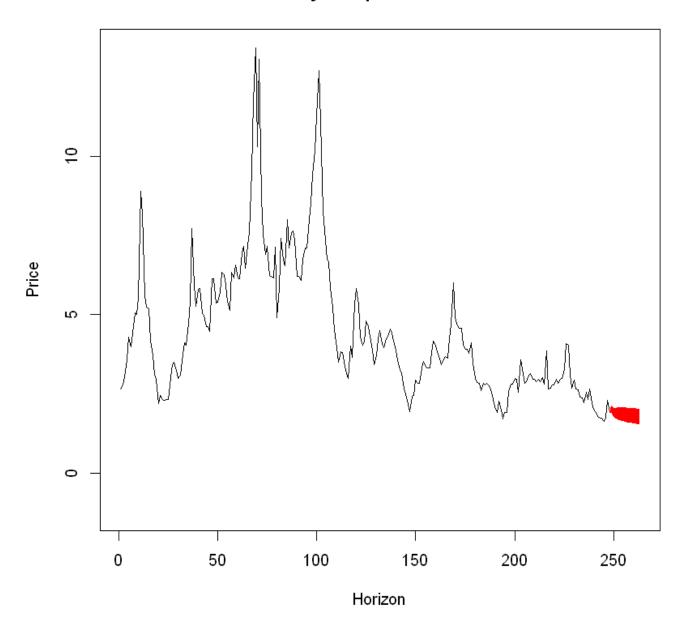
```
forecast <- predict(var_mod, n.ahead = 15, ci = 0.95)
fanchart(forecast, names = "henryhub", main = "HenryHub price
forecast", xlab = "Horizon", ylab = "Price")
forecast</pre>
```

### \$HenryHub

	fcst	lower	upper	CI
[1,	2.023945	0.5085243	3.539366	1.515421
[2,	1.918098	-0.1408333	3.977030	2.058932
[3,	1.898657	-0.5686141	4.365928	2.467271
[4,	1.906046	-0.8471416	4.659233	2.753187
[5,	1.877952	-1.0853022	4.841206	2.963254
[6,	] 1.865119	-1.2529830	4.983221	3.118102
[7,	1.864559	-1.3717686	5.100886	3.236328
[8,	1.851493	-1.4754498	5.178436	3.326943
ľΩ	1 1 0/10//	1 55/200/	5 220070	2 207120

```
[9,] 1.642646 -1.5542964 5.259979 3.397159
[10,] 1.837991 -1.6141955 5.290177 3.452186
[11,] 1.828779 -1.6666943 5.324253 3.495474
[12,] 1.820372 -1.7093403 5.350085 3.529713
[13,] 1.812809 -1.7441342 5.369753 3.556944
[14,] 1.803387 -1.7752512 5.382025 3.578638
[15,] 1.793695 -1.8022861 5.389676 3.595981
```

## HenryHub price forecast



# **Cointegration test:**

If variables are found to be cointegrated, then we should work with an error correction model (ECM) involving these variables.

### **Eigen test:**

```
coin = ca.jo(df, type = "eigen", ecdet = "none", K = 3, spec =
 "transitory")
summary(coin)
# Johansen-Procedure #
Test type: maximal eigenvalue statistic (lambda max), with linear trend
Eigenvalues (lambda):
[1] 0.12310724 0.05612077 0.03222678
Values of teststatistic and critical values of test:
         test 10pct 5pct 1pct
r <= 2 | 8.03 6.50 8.18 11.65
r <= 1 | 14.15 12.91 14.90 19.19
r = 0 \mid 32.19 \mid 18.90 \mid 21.07 \mid 25.75
Eigenvectors, normalised to first column:
(These are the cointegration relations)
           henryhub.l1
                                      brent.l1
                          wti.l1
henryhub.l1
             1.0000000 1.00000000 1.000000000
wti.l1
            -0.7789157 -0.94858843 0.008402311
brent.l1
             0.6724960 0.05803292 -0.023625046
Weights W:
(This is the loading matrix)
           henryhub.l1
                           wti.l1
                                     brent.l1
henryhub.d -0.002444656 0.004290066 -0.05996139
wti.d
           0.385850853 0.233467332 0.54049420
```

• r=2 & r=1: the critical value @ 95% are > test statistics, so we cannot reject H0; means our series are cointegrated and they have long run relations. Hence we can

-0.252220780 0.266163009 0.65840636

brent.d

run ECM.

• r = 0: critical value @95% < test statistic, so we reject H0

#### **Trace test:**

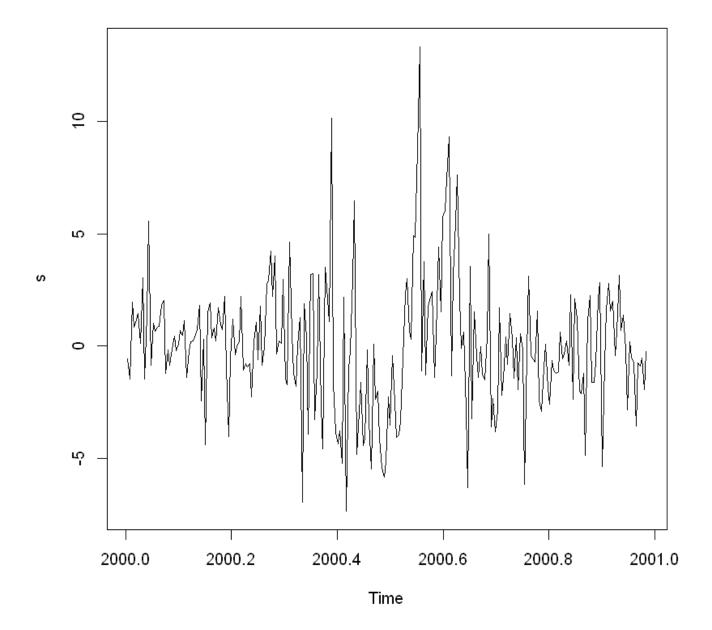
```
coin = ca.jo(df, type = "trace", ecdet = "none", K = 3, spec =
"transitory")
# endet = 'none' means there is a linear trend in data
summary(coin)
```

```
# Johansen-Procedure #
Test type: trace statistic, with linear trend
Eigenvalues (lambda):
[1] 0.12310724 0.05612077 0.03222678
Values of teststatistic and critical values of test:
         test 10pct 5pct 1pct
r <= 2 | 8.03 6.50 8.18 11.65
r <= 1 | 22.18 15.66 17.95 23.52
r = 0 | 54.36 28.71 31.52 37.22
Eigenvectors, normalised to first column:
(These are the cointegration relations)
           henryhub.l1 wti.l1
                                    brent.l1
henryhub.l1 1.0000000 1.00000000 1.0000000000
wti.l1
           -0.7789157 -0.94858843 0.008402311
brent.l1
            0.6724960 0.05803292 -0.023625046
Weights W:
(This is the loading matrix)
           henryhub.l1 wti.l1 brent.l1
henryhub.d -0.002444656 0.004290066 -0.05996139
wti.d
           0.385850853 0.233467332 0.54049420
brent.d
          -0.252220780 0.266163009 0.65840636
```

s = 1\*henryhub - 0.7788110 \* wti + 0.6723858\*brent + 0.4075001 plot(s, type = 'l')

# Augmented Dickey-Fuller Test

data: s
Dickey-Fuller = -3.9344, Lag order = 6, p-value = 0.01283
alternative hypothesis: stationary



### **Error correction model:**

```
model_vecm = VECM(df, lag=3, r=1, estim = 'ML')
summary(model_vecm)
```

Equation hen Equation wti Equation bre	•	71) 0.9615(1.4087	7) 4.0662(1.6759)*
Equation hen	nryhub 0.0031(0.02		
Equation wti	•	,	,
Equation bre	•		,
•	wti -2 `	brent -2	henryhub -3
Equation hen	ryhub 0.0090(0.01	.86) -0.0045(0.01	161) -0.0352(0.0691)
Equation wti	-0.7466(0.4	482). 0.2381(0.386	0.8158(1.6632)
Equation bre	ent -0.6502(0.5	136) 0.0942(0.442	28) 0.8632(1.9057)
	wti -3	brent -3	
Equation hen	ryhub 0.0060(0.01	49) -0.0022(0.013	31)
Equation wti	0.1549(0.35	94) -0.1605(0.314	12)
Equation bre	ent 0.1666(0.41	18) -0.1871(0.366	<del>3</del> 0)

# **VAR representation of VECM**

```
VARrep(model_vecm)
```

	constant	henryhub.l1	wti.l1	brent.l1	henryhub.l2	wti.l2	brent.l2	henryhub.l3	wti.l3	brent.l3	henryhub.
henryhub	-0.0004806127	0.9715251	-0.001415281	-7.319949e- 05	0.06092865	0.005881203	-0.0006319013	-0.06205003	-0.002984156	0.002376428	0.035188
wti	0.9614912589	4.7073087	-0.541764120	7.453706e-01	-1.13314517	0.278697424	-0.0673222273	-2.11727341	0.901514891	-0.398614006	-0.815774
brent	0.3144215925	4.2411088	-0.915974879	1.120904e+00	-1.24383218	0.215857487	0.0158534913	-2.07206761	0.816784389	-0.281293110	-0.863201

### **Forecast VECM:**

```
fcast = predict(model_vecm, n.ahead = 15)
fcast
```

	henryhub	wti	brent		
249	2.051539	45.32164	48.27112		
250	2.006353	52.23220	55.64675		
251	1.950159	45.78249	49.17743		
252	1.982784	46.95041	50.39440		
253	1.969997	49.37090	53.00246		
254	1.942418	47.18312	50.82095		
255	1.950695	47.63156	51.29926		
256	1.941458	48.50981	52.26203		
257	1.927774	47.80119	51.57116		
258	1.926214	48.00356	51.79891		
259	1.918714	48.34895	52.18861		
260	1.909685	48.15135	52.01215		
261	1.904833	48.26792	52.15220		
262	1.897930	48.43174	52.34589		
263	1.890522	48.40989	52.34586		

# **Summary:**

Above we have shown a simplified approach. However, in real-life scenario building a VAR/ECM model involves several steps as stated below:

1. Information criterion (IC) to identify the optimal order; I have used AIC here; however HQ, FPE and SC can also be tried if the lag values are different in such cases.

- 2. Once optimal lag is identified, we need to estimate the model using OLS method and check the test statistic of the residuals to determine the goodness of fit o the model.
- 3. Forecast using fitted model.
- 4. Error correction model can be used if variables are found to be cointegrated.

### Connect me here.

Note: The programs described here are experimental and should be used with caution for any commercial purpose. All such use at your own risk.

