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Factor analysis is a generalization of PPCA, where we have a different noise variance per dimension.

$p(x_i|y_i) = N(x_i|W * y_i + \mu, \Psi)$, where Ψ is a diagonal matrix with diagonal elements $\sigma_1^2, \sigma_2^2, \dots, \sigma_F^2$.

$p(y_i) = N(y_i|0, I)$, as in standard PPCA.

Note that x_i has a dimension $F * 1$ and y_i has a dimension $d * 1$.

1. Finding the marginal, $p(x_i)$ and the posterior, $p(y_i|x_i)$

$$p(x_i|y_i) \sim (x_i - (W * y_i + \mu))' * \Psi^{-1} * (x_i - (W * y_i + \mu))$$

$$p(y_i) \sim y_i' * y_i$$

Since $p(x_i) = \int p(x_i|y_i) * p(y_i) * dy_i$, we need to use “completing the squares” technique, let's denote $\tilde{x} = x - \mu$

- a. Inside the integral, in the exponential we have:

$$(\tilde{x} - W * y_i)' * \Psi^{-1} * (\tilde{x} - W * y_i) \quad (1)$$

- b. Considering only terms containing the variable that we are marginalizing out (y_i)

$$y_i' * (I + W' * \Psi^{-1} * W) * y_i - 2 * \tilde{x}' * \Psi^{-1} * W * y_i$$

- c. For any arbitrary, but symmetric matrix B, we can write

$$= y_i' * (I + W' * \Psi^{-1} * W) * y_i - 2 * (B^{-1} * W' * \Psi^{-1} * \tilde{x})' * B * y_i \quad (2)$$

- d. Matching terms: a general quadratic expression in y_i : $(y_i - m)' * S * (y_i - m) =$

$$y_i' * S * y_i - 2 * m' * S * y_i + m' * S * m \quad (3), \text{ for symmetric } S$$

Comparison of (2) with (3) yields:

$$S = I + W' * \Psi^{-1} * W$$

$$m' * S = (B^{-1} * W' * \Psi^{-1} * \tilde{x})' * B$$

Solution:

$$B = S = I + W' * \Psi^{-1} * W$$

$$m = B^{-1} * W' * \Psi^{-1} * \tilde{x}$$

If we add the following term to (2):

$$m' * S * m = (B^{-1} * W' * \Psi^{-1} * \tilde{x})' * B * (B^{-1} * W' * \Psi^{-1} * \tilde{x})$$

$$= \tilde{x}' * \Psi^{-1} * W * B^{-1} * W' * \Psi^{-1} * \tilde{x}$$

This means that we identified that the quadratic term that cannot be taken out of the integral (since it is in y_i) is inside the exponential of a Gaussian:

$$N(y_i|B^{-1} * W' * \Psi^{-1} * \tilde{x}, B^{-1})$$

Since this is a Gaussian for y_i , but depends on x_i , this must be $p(y_i|x_i)$. Hence

$$p(y_i|x_i) = N(y_i|B^{-1} * W' * \Psi^{-1} * (x_i - \mu), B) \quad (4)$$

- e. We had a term in (1) that we did not consider, which is $\tilde{x}' * \Psi^{-1} * \tilde{x}$, plus we have to subtract what we added, which was $\tilde{x}' * \Psi^{-1} * W * B^{-1} * W' * \Psi^{-1} * \tilde{x}$. So the remaining terms (that obviously go outside the integral) are

$$\tilde{x}' * \Psi^{-1} * \tilde{x} - \tilde{x}' * \Psi^{-1} * W * B^{-1} * W' * \Psi^{-1} * \tilde{x}$$

$$= \tilde{x}' * (\Psi^{-1} - \Psi^{-1} * W * B^{-1} * W' * \Psi^{-1}) * \tilde{x}$$

Since whatever was inside the integral integrates to one (since the integral of a Gaussian is one), we have that

$$p(x_i) = N(\tilde{x}|0, (\Psi^{-1} - \Psi^{-1} * W * B^{-1} * W' * \Psi^{-1})^{-1})$$

$$= N(x_i|\mu, (\Psi^{-1} - \Psi^{-1} * W * B^{-1} * W' * \Psi^{-1})^{-1}) \quad (5)$$

Using the Woodbury identity,

$$(A + U * C * V)^{-1} = A^{-1} - A^{-1} * U * (C^{-1} + V * A^{-1} * U)^{-1} * V * A^{-1}$$

Choosing

$$A = \Psi$$

$$U = W$$

$$C = I$$

$$V = W'$$

We can easily see that $(\Psi + W * W')^{-1} = \Psi^{-1} - \Psi^{-1} * W * B^{-1} * W' * \Psi^{-1}$

Hence $(\Psi^{-1} - \Psi^{-1} * W * B^{-1} * W' * \Psi^{-1})^{-1} = \Psi + W * W'$

Substituting this into (5) yields:

$$p(x_i) = N(x_i | \mu, \Psi + W * W')$$

2. Devise an Expectation-Maximization algorithm for Factor Analysis

We aim to maximize the joint probability:

$$p(x_i, y_i) = p(x_i | y_i) * p(y_i) = N(x_i | W * y_i + \mu, \Psi) * N(y_i | 0, I) \quad (6)$$

Assuming independent samples:

$$p(X, Y) = \prod_{i=1}^N p(x_i, y_i) \quad (7)$$

Substituting (6) into (7) and taking the natural logarithm yields:

$$\begin{aligned} \ln(p(X, Y)) &= \frac{-1}{2} * \sum_{i=1}^N F * \ln(2\pi) + \ln(\Psi) + (x_i - (W * y_i + \mu))' * \Psi^{-1} \\ &\quad * (x_i - (W * y_i + \mu)) + d * \ln(2\pi) + y_i' * y_i \quad (8) \end{aligned}$$

Taking the expectation of (8) on the posterior

$$\begin{aligned} L(W, \mu, \Psi) &\stackrel{\text{def}}{=} E_{p(Y|X)}\{\ln(p(X, Y))\} \\ &= \frac{-1}{2} * \left[\sum_{i=1}^N +F * \ln(2\pi) + \ln(|\Psi|) + (x_i - \mu)' * \Psi^{-1} \right. \\ &\quad * (x_i - \mu) - 2 * E\{y_i\}' * W' * \Psi^{-1} * (x_i - \mu) + E\{(W * y_i)' \\ &\quad * \Psi^{-1} * (W * y_i)\} + d * \ln(2\pi) + E\{y_i' * y_i\} \left. \right] \quad (9) \end{aligned}$$

Note that $E\{(W * y_i)' * \Psi^{-1} * (W * y_i)\} = \text{tr}(E\{y_i * y_i'\} * W' * \Psi^{-1} * W)$, which – since Ψ^{-1} is a diagonal matrix – is $\text{tr}(\Psi^{-1} * E\{y_i * y_i'\} * W' * W)$

Substituting this to (9) yields:

$$\begin{aligned} L(W, \mu, \Psi) &\stackrel{\text{def}}{=} E_{p(Y|X)}\{\ln(p(X, Y))\} \\ &= \frac{-1}{2} * \left[\sum_{i=1}^N +F * \ln(2\pi) + \ln(\Psi) + (x_i - \mu)' * \Psi^{-1} * (x_i - \mu) \right. \\ &\quad - 2 * E\{y_i\}' * W' * \Psi^{-1} * (x_i - \mu) \\ &\quad + \text{tr}(\Psi^{-1} * E\{y_i * y_i'\} * W' * W) + d * \ln(2\pi) \\ &\quad \left. + E\{y_i' * y_i\} \right] \quad (10) \end{aligned}$$

We know $E\{y_i\}$ directly from (4):

$$E\{y_i\} = B^{-1} * W' * \Psi^{-1} * (x_i - \mu) \quad (11)$$

To calculate $E\{y_i' * y_i\}$, we will use that

$$E\{y_i' * y_i\} = \text{cov}(y_i) + E\{y_i\} * E\{y_i\}'$$

Hence

$$E\{y_i' * y_i\} = B^{-1} + E\{y_i\} * E\{y_i\}' \quad (12)$$

a) Finding the optimal W:

- a. Examining the term: $-2 * E\{y_i\}' * W' * \Psi^{-1} * (x_i - \mu)$

$$\begin{aligned} \frac{\delta}{\delta W} (-2 * E\{y_i\}' * W' * \Psi^{-1} * (x_i - \mu)) \\ = \frac{\delta}{\delta W} (-2 * tr(\Psi^{-1} * (x_i - \mu) * E\{y_i\}' * W')) \\ = -2 * \Psi^{-1} * (x_i - \mu) * E\{y_i\}' \quad (13) \end{aligned}$$

- b. Examining the term: $tr(\Psi^{-1} * E\{y_i * y_i'\} * W' * W)$

Using formula 118 from the Matrix Cookbook noting that matrix C in the formula is the identity matrix and that $\{y_i * y_i'\} = E\{y_i * y_i'\}$:

$$\frac{\delta}{\delta W} tr(\Psi^{-1} * E\{y_i * y_i'\} * W' * W) = 2 * \Psi^{-1} * W * E\{y_i * y_i'\} \quad (14)$$

Using (13) and (14) multiplied by $\frac{1}{2} * \Psi$ from the left and (10) we have

$$\begin{aligned} \frac{\delta L(W)}{\delta W} = 0 \rightarrow \\ W = \sum_i (x_i - \mu) * E\{y_i\}' * \left(\sum_i E\{y_i * y_i'\} \right)^{-1} \quad (15) \end{aligned}$$

- b) Finding the optimal Ψ :

- a. Examining the term $-2 * E\{y_i\}' * W' * \Psi^{-1} * (x_i - \mu)$

$$\begin{aligned} \frac{\delta}{\delta \Psi^{-1}} -2 * E\{y_i\}' * W' * \Psi^{-1} * (x_i - \mu) \\ = -2 * \frac{\delta}{\delta \Psi^{-1}} (x_i - \mu)' * \Psi^{-1} * W * E\{y_i\} \\ = -2 * \frac{\delta}{\delta \Psi^{-1}} tr(W * E\{y_i\} * (x_i - \mu)' * \Psi^{-1}) \end{aligned}$$

Which is –using formula (142) in the Matrix Cookbook –

$$= diag(-2 * W * E\{y_i\} * (x_i - \mu)') \quad (16)$$

- b. Examining the term $(x_i - \mu)' * \Psi^{-1} * (x_i - \mu)$

$$\begin{aligned} \frac{\delta}{\delta \Psi^{-1}} (x_i - \mu)' * \Psi^{-1} * (x_i - \mu) = \frac{\delta}{\delta \Psi^{-1}} tr((x_i - \mu) * (x_i - \mu)' * \Psi^{-1}) \\ = diag((x_i - \mu) * (x_i - \mu)') \quad (17) \end{aligned}$$

- c. Examining the term $tr(\Psi^{-1} * E\{y_i * y_i'\} * W' * W)$

$$\begin{aligned} \frac{\delta}{\delta \Psi^{-1}} tr(\Psi^{-1} * E\{y_i * y_i'\} * W' * W) = \frac{\delta}{\delta \Psi^{-1}} tr(W * E\{y_i * y_i'\} * W' * \Psi^{-1}) \\ = diag(W * E\{y_i * y_i'\} * W') \quad (18) \end{aligned}$$

- d. Examining the term $\ln(\Psi)$

$\frac{\delta}{\delta \Psi^{-1}} \ln(\Psi) = \frac{\delta}{\delta \Psi^{-1}} - \ln(|\Psi|^{-1})$, which is – due to property of the determinant –
 $\frac{\delta}{\delta \Psi^{-1}} - \ln(|\Psi^{-1}|)$, which is –using formula (57) from the matrix cookbook and the fact that Ψ is symmetric: $-\Psi$.

Alternatively, we can realize that since Ψ is diagonal, $\frac{\delta}{\delta \Psi^{-1}} \ln(\Psi)$ will be the same as if Ψ was a scalar. Hence $\frac{\delta}{\delta \Psi^{-1}} \ln(\Psi) = -\Psi \quad (19)$

Using (16)-(19) and (9) we have that

$$\frac{\delta L(\Psi)}{\delta \Psi^{-1}} = 0 \rightarrow$$

$$\Psi = \frac{1}{N} * \text{diag} \left(\sum_i [(x_i - \mu) * (x_i - \mu)' - 2 * W * E\{y_i\} * (x_i - \mu)' + W * E\{y_i * y_i'\} * W'] \right) \quad (20)$$

c) Finding the optimal μ :

a. Examining the term $(x_i - \mu)' * \Psi^{-1} * (x_i - \mu)$

$\frac{\delta}{\delta \mu} (x_i - \mu)' * \Psi^{-1} * (x_i - \mu)$ equals – using formula 86 from the Matrix Cookbook – to $-2 * \Psi^{-1} * (x_i - \mu)$ (21)

b. Examining the term $-2 * E\{y_i\}' * W' * \Psi^{-1} * (x_i - \mu)$

Denote $-2 * E\{y_i\}' * W' * \Psi^{-1}$ by d and note that it is a row vector. Then

$$\frac{\delta}{\delta \mu} d * (x_i - \mu) = -d' = 2 * \Psi^{-1} * W * E\{y_i\} \quad (22)$$

Using (21) and (22) multiplied by $\frac{1}{2} * \Psi$ from the left and (10) we have

$$\frac{\delta L(\mu)}{\delta \mu} = 0 \rightarrow$$

$$\mu = \frac{1}{N} \sum_i x_i - W * E\{y_i\} \quad (23)$$

Summary:

The EM updates are given by (15), (20) and (23) and the required expectations are given by (11) and (12).