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Based on the exercise provided by Dr. Stefanos Zafeiriou as part of CO-495, Imperial College London

Factor analysis is a generalization of PPCA, where we have a different noise variance per dimension.

 $p(x_i|y_i) = N(x_i|W*y_i + \mu, \Psi)$ , where  $\Psi$  is a diagonal matrix with diagonal elements  $\sigma_1^2, \sigma_2^2, \dots, \sigma_F^2$ .  $p(y_i) = N(y_i|0,I)$ , as in standard PPCA.

Note that  $x_i$  has dimension F\*1 and  $y_i$  has dimension d\*1, d < F. Accordingly, W has dimension F\*d and Y has dimension F\*F.

1. Finding the marginal,  $p(x_i)$  and the posterior,  $p(y_i|x_i)$ 

$$p(x_i|y_i) \sim (x_i - (W * y_i + \mu))' * \Psi^{-1} * (x_i - (W * y_i + \mu))$$
$$p(y_i) \sim y_i' * y_i$$

Since  $p(x_i) = \int p(x_i|y_i) * p(y_i) * dy_i$ , we need to use "completing the squares" technique, let's denote  $\tilde{x} = x_i - \mu$ 

a. Inside the integral, in the exponential we have:

$$(\tilde{x} - W * y_i)' * \Psi^{-1} * (\tilde{x} - W * y_i) + y_i' * y_i$$
 (1)

b. Considering only terms containing the variable that we are marginalizing out  $(y_i)$ 

$$y_i' * (I + W' * \Psi^{-1} * W) * y_i - 2 * \tilde{x}' * \Psi^{-1} * W * y_i$$

c. For any arbitrary, but symmetric matrix B, we can write

$$= y_i' * (I + W' * \Psi^{-1} * W) * y_i - 2 * (B^{-1} * W' * \Psi^{-1} * \tilde{x})' * B * y_i$$
 (2)

d. Matching terms: a general quadratic expression in  $y_i$ :  $(y_i - m)' * S * (y - m) = y_i' * S * y_i - 2 * m' * S * y_i + m' * S * m$  (3), for symmetric S Comparison of (2) with (3) yields:

$$S = I + W' * \Psi^{-1} * W$$
  
$$m' * S = (B^{-1} * W' * \Psi^{-1} * \tilde{x})' * B$$

Solution:

$$B = S = I + W' * \Psi^{-1} * W$$
  
 $m = B^{-1} * W' * \Psi^{-1} * \tilde{x}$ 

If we add the following term to (2):

$$m' * S * m = (B^{-1} * W' * \Psi^{-1} * \tilde{x})' * B * (B^{-1} * W' * \Psi^{-1} * \tilde{x})$$
$$= \tilde{x}' * \Psi^{-1} * W * B^{-1} * W' * \Psi^{-1} * \tilde{x}$$

Then we can complete the square and it yields the quadratic term in (3) that is in the exponential of the following Gaussian (after substituting m and S):

$$N(y_i|B^{-1}*W'*\Psi^{-1}*\tilde{x},B^{-1})$$

Since this is a Gaussian for  $y_i$ , but depends on  $x_i$ , this must be  $p(y_i|x_i)$ . Hence

$$p(y_i|x_i) = N(y_i|B^{-1} * W' * \Psi^{-1} * (x_i - \mu), B^{-1})$$
 (4)

e. We had a term in (1) that we did not consider, which is  $\tilde{x}' * \Psi^{-1} * \tilde{x}$ , plus we have to subtract what we added, which was  $\tilde{x}' * \Psi^{-1} * W * B^{-1} * W' * \Psi^{-1} * \tilde{x}$ . So the remaining terms (that obviously go outside the integral) are

Since whatever was inside the integral integrates to one (since the integral of a Gaussian is one), we have that

$$p(x_i) = N(\tilde{x}|0, (\Psi^{-1} - \Psi^{-1} * W * B^{-1} * W' * \Psi^{-1})^{-1})$$
  
=  $N(x_i|\mu, (\Psi^{-1} - \Psi^{-1} * W * B^{-1} * W' * \Psi^{-1})^{-1})$  (5)

Using the Woodbury identity,

$$(A + U * C * V)^{-1} = A^{-1} - A^{-1} * U(C^{-1} + V * A^{-1} * U)^{-1} * V * A^{-1}$$
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Choosing

$$A = \Psi$$

$$U = W$$

$$C = I$$

$$V = W'$$

We can easily see that  $(\Psi + W * W')^{-1} = \Psi^{-1} - \Psi^{-1} * W * B^{-1} * W' * \Psi^{-1}$  Hence  $(\Psi^{-1} - \Psi^{-1} * W * B^{-1} * W' * \Psi^{-1})^{-1} = \Psi + W * W'$ 

Substituting this into (5) yields:

$$p(x_i) = N(x_i | \mu, \Psi + W * W')$$

2. Devise an Expectation-Maximization algorithm for Factor Analysis We aim to maximize the joint probability:

 $p(x_i,y_i) = p(x_i|y_i) * p(y_i) = N(x_i|W*y_i + \mu, \Psi) * N(y_i|0,I)$  (6) Assuming independent samples:

$$p(X,Y) = \prod_{i=1}^{N} p(x_i, y_i)$$
 (7)

Substituting (6) into (7) and taking the natural logarithm yields:

$$\ln(p(X,Y)) = \frac{-1}{2} * \sum_{i=1}^{N} F * \ln(2\pi) + \ln(\Psi) + (x_i - (W * y_i + \mu))' * \Psi^{-1}$$
$$* (x_i - (W * y_i + \mu)) + d * \ln(2\pi) + * y_i' * y_i$$
(8)

Taking the expectation of (8) on the posterior

$$\begin{split} L(W,\mu,\Psi) &\stackrel{\text{def}}{=} E_{p(Y|X)} \{ \ln(p(X,Y)) \} \\ &= \frac{-1}{2} * [\sum_{i=1}^{N} + F * \ln(2\pi) + \ln(|\Psi|) + (x_i - \mu)' * \Psi^{-1} \\ &\quad * (x_i - \mu) - 2 * E\{y_i\}' * W' * \Psi^{-1} * (x_i - \mu) + E\{(W * y_i)' \\ &\quad * \Psi^{-1} * (W * y_i) \} + d * \ln(2\pi) + E\{y_i' * y_i\} ] \end{split}$$

Note that  $E\{(W*y_i)'*\Psi^{-1}*(W*y_i)\} = tr(E\{y_i*y_i'\}*W'*\Psi^{-1}*W)$ , which—since  $\Psi^{-1}$  is a diagonal matrix—is  $tr(\Psi^{-1}*E\{y_i*y_i'\}*W'*W)$  Substituting this to (9) yields:

$$\begin{split} L(W,\mu,\Psi) &\stackrel{\text{def}}{=} E_{p(Y|X)} \{ \ln(p(X,Y)) \} \\ &= \frac{-1}{2} * [\sum_{i=1}^{N} +F * \ln(2\pi) + \ln(\Psi) + (x_{i} - \mu)' * \Psi^{-1} * (x_{i} - \mu) \\ &- 2 * E\{y_{i}\}' * W' * \Psi^{-1} * (x_{i} - \mu) \\ &+ tr(\Psi^{-1} * E\{y_{i} * y_{i}'\} * W' * W) + d * \ln(2\pi) \\ &+ E\{y_{i}' * y_{i}\} ] \end{split}$$

We know  $E\{y_i\}$  directly from (4):

$$E\{y_i\} = B^{-1} * W' * \Psi^{-1} * (x_i - \mu)$$
 (11)

To calculate  $E\{y_i' * y_i\}$ , we will use that

$$E\{y_i' * y_i\} = cov(y_i) + E\{y_i\} * E\{y_i\}'$$

Hence

$$E\{y_i' * y_i\} = B^{-1} + E\{y_i\} * E\{y_i\}' \quad (12)$$

- a) Finding the optimal W:
  - a. Examining the term:  $-2 * E\{y_i\}' * W' * \Psi^{-1} * (x_i \mu)$

$$\begin{split} \frac{\delta}{\delta W} \left( -2 * E \{ y_i \}' * W' * \Psi^{-1} * (x_i - \mu) \right) \\ &= \frac{\delta}{\delta W} \left( -2 * tr(\Psi^{-1} * (x_i - \mu) * E \{ y_i \}' * W') \right) \\ &= -2 * \Psi^{-1} * (x_i - \mu) * E \{ y_i \}' \end{split}$$
 (13)

b. Examining the term:  $tr(\Psi^{-1}*E\{y_i*y_i'\}*W'*W)$ Using formula 118 from the Matrix Cookbook noting that matrix C in the formula is the identity matrix and that  $\{y_i*y_i'\}=E\{y_i*y_i'\}'$ :

$$\frac{\delta}{\delta W} tr(\Psi^{-1} * E\{y_i * y_i'\} * W' * W) = 2 * \Psi^{-1} * W * E\{y_i * y_i'\}$$
 (14)

Using (13) and (14) multiplied by  $\frac{1}{2} * \Psi$  from the left and (10) we have

$$\frac{\delta L(W)}{\delta W} = 0 \rightarrow$$

$$W = \sum_{i} (x_i - \mu) * E\{y_i\}' * \left(\sum_{i} E\{y_i * y_i'\}\right)^{-1}$$
 (15)

- b) Finding the optimal  $\Psi$ :
  - a. Examining the term  $-2 * E\{y_i\}' * W' * \Psi^{-1} * (x_i \mu)$

$$\begin{split} \frac{\delta}{\delta \Psi^{-1}} - 2 * E \{ y_i \}' * W' * \Psi^{-1} * (x_i - \mu) \\ &= -2 * \frac{\delta}{\delta \Psi^{-1}} (x_i - \mu)' * \Psi^{-1} * W * E \{ y_i \} \\ &= -2 * \frac{\delta}{\delta \Psi^{-1}} tr(W * E \{ y_i \} * (x_i - \mu)' * \Psi^{-1}) \end{split}$$

Which is -using formula (142) in the Matrix Cookbook -

$$= diag(-2 * W * E\{y_i\} * (x_i - \mu)') \quad (16)$$

b. Examining the term  $(x_i - \mu)' * \Psi^{-1} * (x_i - \mu)$ 

$$\frac{\delta}{\delta \Psi^{-1}} (x_i - \mu)' * \Psi^{-1} * (x_i - \mu) = \frac{\delta}{\delta \Psi^{-1}} * tr((x_i - \mu) * (x_i - \mu)' * \Psi^{-1})$$

$$= diag((x_i - \mu) * (x_i - \mu)') \quad (17)$$

c. Examining the term  $tr(\Psi^{-1} * E\{y_i * y_i'\} * W' * W)$ 

$$\frac{\delta}{\delta \Psi^{-1}} tr(\Psi^{-1} * E\{y_i * y_i'\} * W' * W) = \frac{\delta}{\delta \Psi^{-1}} tr(W * E\{y_i * y_i'\} * W' * \Psi^{-1})$$

$$= diag(W * E\{y_i * y_i'\} * W') \quad (18)$$

d. Examining the term  $ln(\Psi)$ 

$$\frac{\delta}{\delta\Psi^{-1}}\ln(\Psi) = \frac{\delta}{\delta\Psi^{-1}} - \ln(|\Psi|^{-1}), \text{ which is } - \text{ due to property of the determinant } - \frac{\delta}{\delta\Psi^{-1}} - \ln(|\Psi^{-1}|), \text{ which is } - \text{using formula (57) from the matrix cookbook and the fact that } \Psi \text{ is symmetric: } -\Psi.$$

Alternatively, we can realize that since  $\Psi$  is diagonal,  $\frac{\delta}{\delta\Psi^{-1}}\ln(\Psi)$  will be the same as if  $\Psi$  was a scalar. Hence  $\frac{\delta}{\delta\Psi^{-1}}\ln(\Psi)=-\Psi$  (19)

Using (16)-(19) and (9) we have that

$$\frac{\delta L(\Psi)}{\delta \Psi^{-1}} = 0 \rightarrow$$

$$\Psi = \frac{1}{N} * diag(\sum_{i} [(x_i - \mu) * (x_i - \mu)' - 2 * W * E\{y_i\} * (x_i - \mu)' + W * E\{y_i * y_i'\} * W'])$$
(20)

- c) Finding the optimal  $\mu$ :
  - a. Examining the term  $(x_i-\mu)'*\Psi^{-1}*(x_i-\mu)$   $\frac{\delta}{\delta\mu}(x_i-\mu)'*\Psi^{-1}*(x_i-\mu) \text{ equals-using formula 86 from the Matrix Cookbook-to } -2*\Psi^{-1}*(x_i-\mu) \tag{21}$
  - b. Examining the term  $-2*E\{y_i\}'*W'*\Psi^{-1}*(x_i-\mu)$  Denote  $-2*E\{y_i\}'*W'*\Psi^{-1}$  by d and note that it is a row vector. Then  $\frac{\delta}{\delta\mu}d*(x_i-\mu)=-d'=2*\Psi^{-1}*W*E\{y_i\} \qquad (22)$

Using (21) and (22) multiplied by  $\frac{1}{2} * \Psi$  from the left and (10) we have

$$\frac{\delta L(\mu)}{\delta \mu} = 0 \to$$

$$\mu = \frac{1}{N} \sum_{i} x_{i} - W * E\{y_{i}\} = \frac{1}{N} \sum_{i} x_{i}$$
 (23)

Summary:

The EM updates are given by (15), (20) and (23) and the required expectations are given by (11) and (12).