

# Homework 6

Qianbo Wang  
uni: qw2180

Consider the ChickWeight data in R. The body weights of the chicks were measured at birth (i.e., time=0) and every second day thereafter until day 20. They were also measured on day 21. There were four groups of chicks on different protein diets.

## Problem 1

Determine whether there is a significant difference in the mean weights of the four groups on Day 18.

- **Without adjusting for Birth Weight.**

Since without adjusting for Birth Weight, the model is just simple one-way anova. The anova table is as follows:

Anova without adjusting for Birth Weight

Call:

```
aov(formula = weight ~ Diet, data = sub_day18)
```

Terms:

	Diet	Residuals
Sum of Squares	36690.44	114840.84
Deg. of Freedom	3	43

Residual standard error: 51.67898

Estimated effects may be unbalanced

	Df	Sum Sq	Mean Sq	F value	Pr(>F)
Diet	3	36690	12230	4.579	0.0072 **
Residuals	43	114841	2671		

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Signif. codes: 0 \*\*\* 0.001 \*\* 0.01 \* 0.05 . 0.1 1

Since for F-test, the null and alternative hypothesis are:

$$H_0 : \mu_1 = \mu_2 = \mu_3 = \mu_4$$

$$H_1 : \mu_i \neq \mu_j \text{ for at least one pair } (i, j), i \neq j$$

And the  $p$ -value = 0.0072 < 0.05. Then we should reject the null hypothesis, i.e. there is a significant difference in the mean weights of the four groups on Day 18.

- **Adjusting for Birth Weight. Give the LS Means (i.e., adjusted for Birth Weight).**

Since we need do anova analysis based on adjusting for Birth Weight, then we simply add a new variable as Birth Weight, i.e. Give the LS means. Then this time the model is anovca. The anovca table is as follows:

Anova with adjusting for Birth Weight

Call:

```
aov(formula = weight ~ birthweight + Diet, data = sub_day18)
```

Terms:

	birthweight	Diet	Residuals
Sum of Squares	15728.23	27190.80	108612.25
Deg. of Freedom	1	3	42

Residual standard error: 50.85279

Estimated effects may be unbalanced

	Df	Sum Sq	Mean Sq	F value	Pr(>F)
birthweight	1	15728	15728	6.082	0.0178 *
Diet	3	27191	9064	3.505	0.0234 *
Residuals	42	108612	2586		

---

Signif. codes: 0 '\*\*\*' 0.001 '\*\*' 0.01 '\*' 0.05 '.' 0.1 ' ' 1

And the Birth Weight means of the four groups is as follows:

Table 1: Birth Weight LS Means

Diet	Birth Weight mean	LS mean
1	41.58824	164.2240
2	40.70000	183.2445
3	40.70000	229.7409
4	41.00000	201.7337

Since this time for F-test, the null and alternative hypothesis are:

$$H_0 : \mu_1 = \mu_2 = \mu_3 = \mu_4$$

$$H_1 : \mu_i \neq \mu_j \text{ for at least one pair } (i, j), i \neq j$$

And the  $p$ -value for Birth Weight is  $p$ -value = 0.0178 < 0.05 and  $p$ -value for Diet is  $p$ -value = 0.0234 < 0.05. Then we should reject the null hypothesis, i.e. there is a significant difference in the mean weights of the four groups on Day 18 regarding the Birth Weight differences.

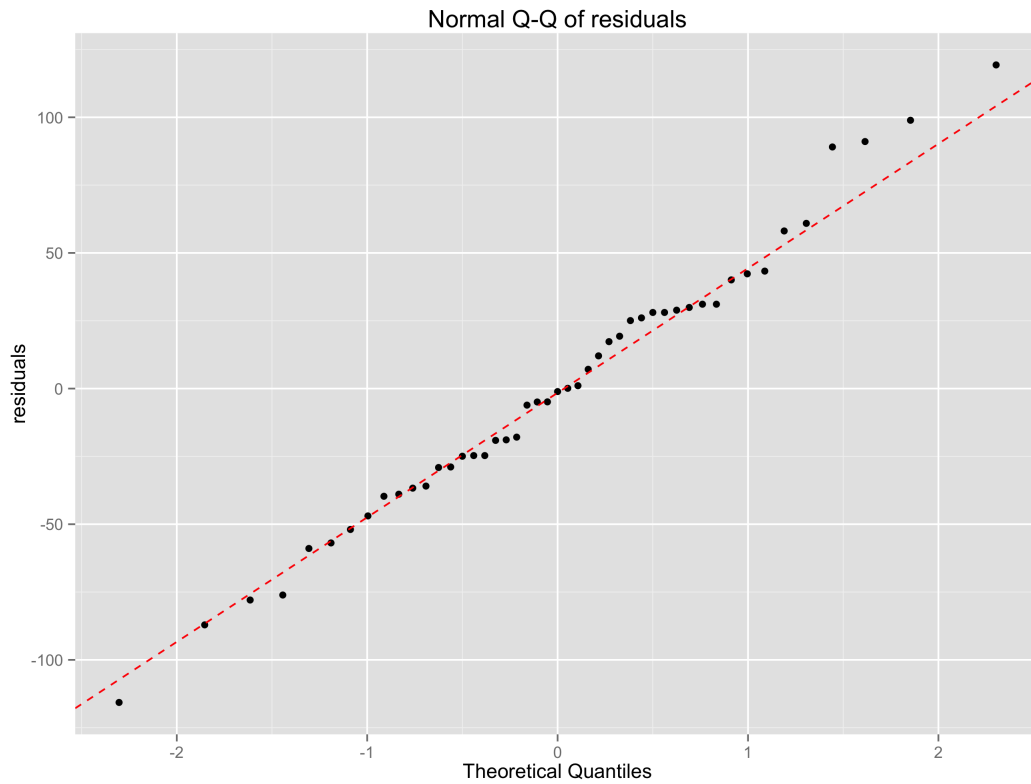
- **Check the validity of your assumptions, including parallelism. Suggest measures that you would take if the assumptions are not satisfied.**

Since the assumptions are:

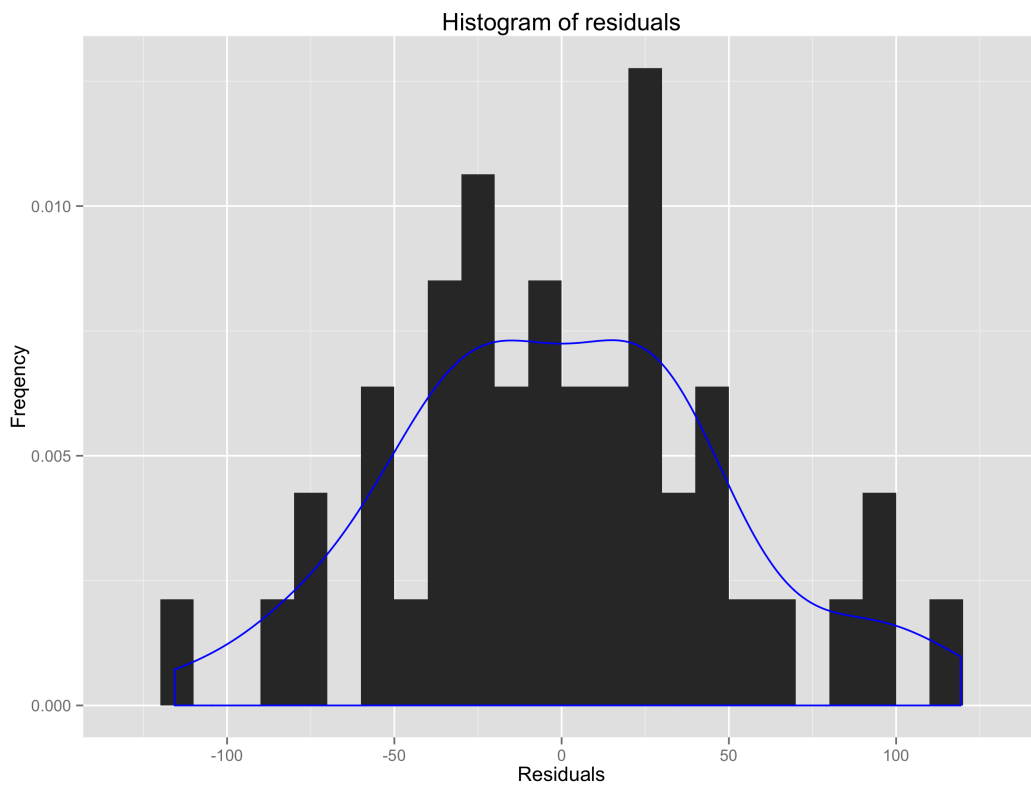
- i.i.d. normal
- constant variance

### 1. Check i.i.d. Normality

First, plot Q-Q plot on residuals, the plot is as follows:



Then, plot histogram with density on residuals, the plot is as follows:



Since from the plot, the residuals seem to satisfy normal distribution. Then do Shapiro-Wilk test on residuals, The null and alternative hypothesis of Shapiro-Wilk test are:

$$H_0 : \epsilon_i \sim \text{Normal}$$

$$H_1 : \epsilon_i \text{ not from Normal}$$

And the Shapiro-Wilk test result is as follows:

Shapiro-Wilk Test on residuals

Shapiro-Wilk normality test

```
data: residual$residuals
W = 0.9886, p-value = 0.9242
```

Since  $p - \text{value} = 0.9242$ , then we should not reject the null, i.e. the residuals satisfy normal distribution.

Then, check the i.i.d. assumption, use nonparametric Kruskal-Wallis test on raw data. The null and alternative hypothesis of Kruskal-Wallis test are:

$$H_0 : w_1, w_2, w_3, w_4 \text{ i.i.d.}$$

$$H_1 : w_1, w_2, w_3, w_4 \text{ not i.i.d.}$$

The result is as follows:

Kruskal-Wallis test on raw data

Kruskal-Wallis rank sum test

```
data: weight by Diet
Kruskal-Wallis chi-squared = 10.6234, df = 3, p-value = 0.01395
```

Since  $p - \text{value} = 0.01395$ , then we should reject the null, i.e. the weight data in the four Diet group are not from the same distribution.

## 2. Check Constant Variance

First, use Bartlett's on raw data to check the constant variance. The null and alternative hypothesis of Bartlett's test are:

$$H_0 : \sigma_1^2 = \dots = \sigma_4^2$$

$$H_1 : \sigma_i^2 \neq \sigma_j^2 \text{ for at least one pair } (i, j), i \neq j$$

And the Bartlett's test is highly dependent on the normal assumption, since we have check the validation of normality, then we can use Bartlett's test. The result is as follows:

Bartlett's test on raw data

Bartlett test of homogeneity of variances

```
data: sub_day18$weight and sub_day18$Diet
Bartlett's K-squared = 3.576, df = 3, p-value = 0.311
```

Since  $p - \text{value} = 0.3111 > 0.05$ , then we should not reject the null, i.e. the constant variance assumption satisfies.

Then use Levene's Test on raw data to check the constant variance. The null and alternative hypothesis of Levene's test are:

$$H_0 : \sigma_1^2 = \dots = \sigma_4^2$$

$$H_1 : \sigma_i^2 \neq \sigma_j^2 \text{ for at least one pair } (i, j), i \neq j$$

The result is as follows:

Levene's test on raw data

Levene's Test for Homogeneity of Variance (center = median)

	Df	F	value	Pr(>F)
group	3	1.059	0.3764	
	43			

Since  $p\text{-value} = 0.3764 > 0.05$ , then we should not reject the null, i.e. the constant variance assumption satisfies.

## 2. Check parallelism

Use two-way anova to check the parallelism, the null and alternative hypothesis are:

$H_0$  : there is joint effect with Diet and Birth Weight

$H_1$  : there is no joint effect with Diet and Birth Weight

The result is as follows:

Two-way Anova for parallelism check

	Df	Sum Sq	Mean Sq	F	value	Pr(>F)
birthweight	1	15728	15728	6.134	0.0177	*
Diet	3	27191	9064	3.535	0.0234	*
birthweight:Diet	3	8610	2870	1.119	0.3530	
Residuals	39	100002	2564			

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Signif. codes: 0 '\*\*\*' 0.001 '\*\*' 0.01 '\*' 0.05 '.' 0.1 ' ' 1

And the  $p\text{-value} = 0.3530 > 0.05$ . Then we should not reject the null hypothesis, i.e. there is no interaction effect between Diet group and the Birth Weight. So we do not need do marginal inference on BirthWeight and Diet.

## Problem 2

Perform an appropriate repeated measures ANOVA to determine whether there is a significant difference in the mean weights of the four groups using the measurements on Days 10, 18, and 21.

- Do the analyses assuming compound symmetry and unstructured covariance structures and compare the results.

### 1. Compound Symmetry

For this part of the problem, I simply use SAS to do the repeat measures ANOVA because SAS has a easily-used implementation of repeated measures ANOVA based on different structure assumptions on covariance. Since for the assumption, it assumes that the covariance between the data is compound symmetry or unstructured. Then the result is in the following page.

### 2. Unstructured Covariance

Since the assumption is unstructured covariance, then the result is in the following page.

### Repeated Measures on Chick Weight compound symmetry assumption

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#### The Mixed Procedure

Model Information	
Data Set	REPEAT.REPEATDATA
Dependent Variable	weight
Covariance Structure	Compound Symmetry
Subject Effect	Chick(Diet)
Estimation Method	REML
Residual Variance Method	Profile
Fixed Effects SE Method	Model-Based
Degrees of Freedom Method	Between-Within

Class Level Information		
Class	Levels	Values
Diet	4	1 2 3 4
Chick	49	1 10 11 12 13 14 15 16 17 19 2 20 21 22 23 24 25 26 27 28 29 3 30 31 32 33 34 35 36 37 38 39 4 40 41 42 43 44 45 46 47 48 49 5 50 6 7 8 9
Time	3	10 18 21

Dimensions	
Covariance Parameters	2
Columns in X	20
Columns in Z	0
Subjects	49
Max Obs Per Subject	3

Number of Observations	
Number of Observations Read	141
Number of Observations Used	141
Number of Observations Not Used	0

Iteration History			
Iteration	Evaluations	-2 Res Log Like	Criterion
0	1	1396.07743029	
1	2	1338.99956090	0.00000228
2	1	1338.99828947	0.00000000

Convergence criteria met.

### Repeated Measures on Chick Weight compound symmetry assumption

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#### The Mixed Procedure

Estimated R Correlation Matrix for Chick(Diet) 1 1			
Row	Col1	Col2	Col3
1	1.0000	0.6626	0.6626
2	0.6626	1.0000	0.6626
3	0.6626	0.6626	1.0000

Covariance Parameter Estimates		
Cov Parm	Subject	Estimate
CS	Chick(Diet)	1534.18
Residual		781.13

Fit Statistics	
-2 Res Log Likelihood	1339.0
AIC (smaller is better)	1343.0
AICC (smaller is better)	1343.1
BIC (smaller is better)	1346.8

Null Model Likelihood Ratio Test		
DF	Chi-Square	Pr > ChiSq
1	57.08	<.0001

Type 3 Tests of Fixed Effects				
Effect	Num DF	Den DF	F Value	Pr > F
Diet	3	45	5.95	0.0016
Time	2	84	192.58	<.0001
Diet*Time	6	84	3.76	0.0023

### Repeated Measures on Chick Weight unstructured correlation assumption

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#### The Mixed Procedure

Model Information	
Data Set	REPEAT.REPEATDATA
Dependent Variable	weight
Covariance Structure	Unstructured
Subject Effect	Chick(Diet)
Estimation Method	REML
Residual Variance Method	None
Fixed Effects SE Method	Model-Based
Degrees of Freedom Method	Between-Within

Class Level Information		
Class	Levels	Values
Diet	4	1 2 3 4
Chick	49	1 10 11 12 13 14 15 16 17 19 2 20 21 22 23 24 25 26 27 28 29 3 30 31 32 33 34 35 36 37 38 39 4 40 41 42 43 44 45 46 47 48 49 5 50 6 7 8 9
Time	3	10 18 21

Dimensions	
Covariance Parameters	6
Columns in X	20
Columns in Z	0
Subjects	49
Max Obs Per Subject	3

Number of Observations	
Number of Observations Read	141
Number of Observations Used	141
Number of Observations Not Used	0

Iteration History			
Iteration	Evaluations	-2 Res Log Like	Criterion
0	1	1396.07743029	
1	3	1177.78560780	0.00103860
2	1	1177.29205471	0.00004402
3	1	1177.27077572	0.00000009
4	1	1177.27073492	0.00000000



### Repeated Measures on Chick Weight unstructured correlation assumption

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#### The Mixed Procedure

Convergence criteria met.
---------------------------

Estimated R Correlation Matrix for Chick(Diet) 1 1			
Row	Col1	Col2	Col3
1	1.0000	0.7988	0.6615
2	0.7988	1.0000	0.9632
3	0.6615	0.9632	1.0000

Covariance Parameter Estimates		
Cov Parm	Subject	Estimate
UN(1,1)	Chick(Diet)	428.98
UN(2,1)	Chick(Diet)	882.94
UN(2,2)	Chick(Diet)	2848.34
UN(3,1)	Chick(Diet)	895.29
UN(3,2)	Chick(Diet)	3359.26
UN(3,3)	Chick(Diet)	4270.54

Fit Statistics	
-2 Res Log Likelihood	1177.3
AIC (smaller is better)	1189.3
AICC (smaller is better)	1190.0
BIC (smaller is better)	1200.6

Null Model Likelihood Ratio Test		
DF	Chi-Square	Pr > ChiSq
5	218.81	<.0001

Type 3 Tests of Fixed Effects				
Effect	Num DF	Den DF	F Value	Pr > F
Diet	3	45	6.03	0.0015
Time	2	45	100.88	<.0001
Diet*Time	6	45	2.47	0.0375

Compare the result:

Since both under compound symmetry covariance and unstructured covariance assumption, all of the  $p$ -values are significant, then we should conclude that the chicken have significant different of weight on at least two Diet group and between Day 10, Day18 and Day 21.

- **Check the validity of your assumptions.**

1. Normality

Since the assumptions for repeated measures are normality and constant variance. Then under the compound symmetry assumption and unstructured covariance assumption, the plots of the residuals given by SAS are as follows: Since from the plots, we see that the data is approximately normally distributed. Then do normal test on residuals, the result is also on the following pages.

Since under compound symmetry covariance assumption, for Shapiro-Wilk test on residuals gives the  $p - value = 0.9778 > 0.05$ , then we should not reject the null, i.e. the normality assumption satisfies.

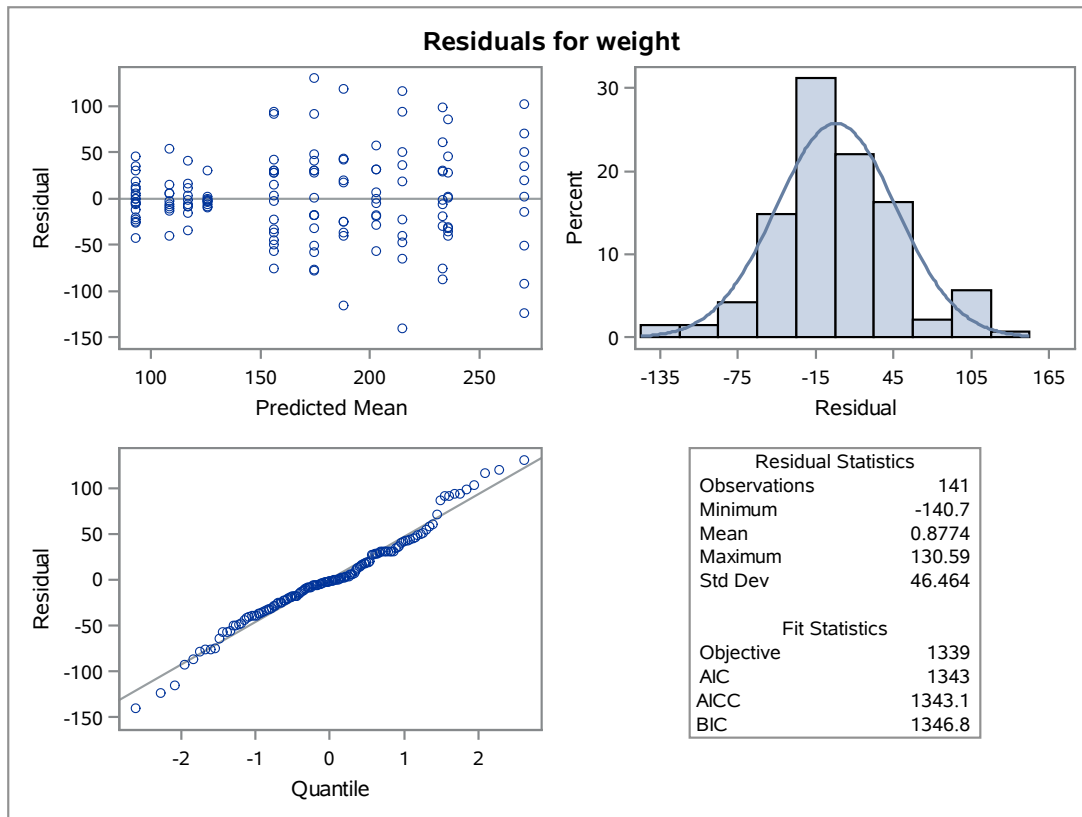
And under unstructured covariance assumption, for Shapiro-Wilk test on residuals gives the  $p - value = 0.9753 > 0.05$ , then we should not reject the null, i.e. the normality assumption satisfies.

2. Constant Variance

Since normality assumption satisfies then use Bartlett's test on raw data gives the  $p - value = 0.0561 > 0.05$ , then we should not reject the null, i.e. the Constant Variance assumption satisfies.

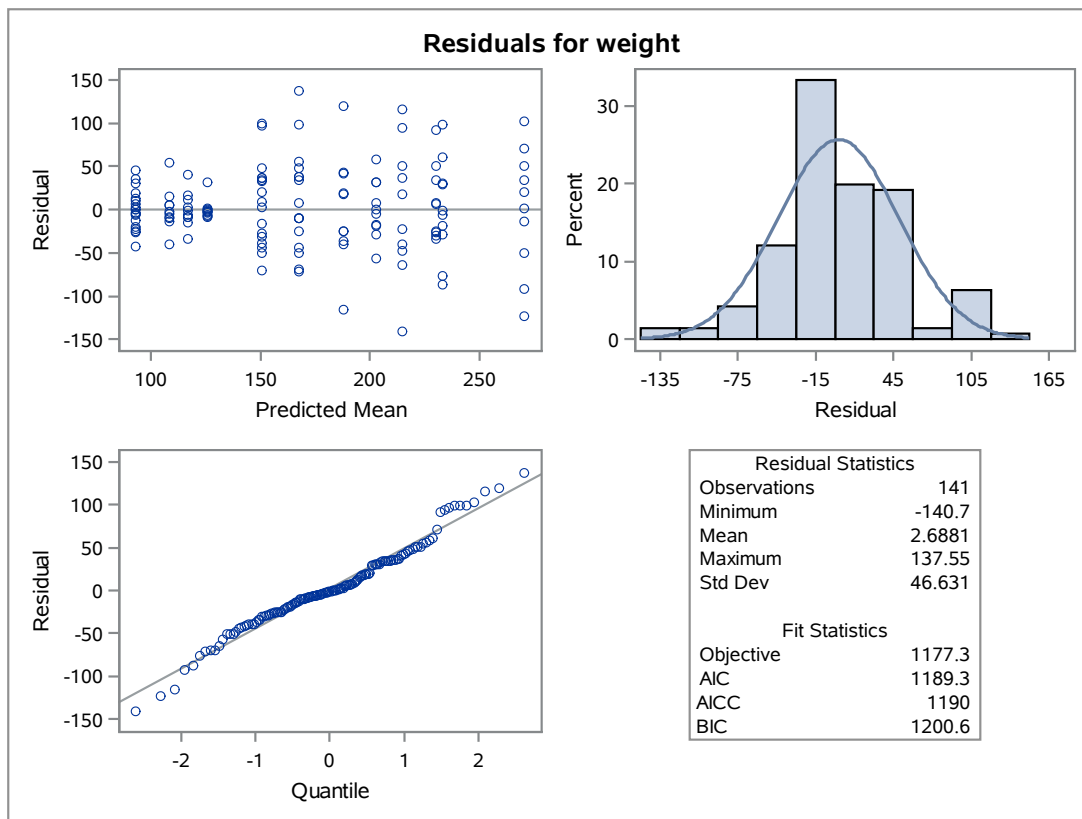
**Repeated Measures on Chick Weight  
compound symmetry assumption**

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**The Mixed Procedure**

**Repeated Measures on Chick Weight  
unstructured correlation assumption**

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**The Mixed Procedure**

### Noramlity test on residuals compound symmetry assumption

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#### The UNIVARIATE Procedure Variable: Resid (Residual)

Moments			
<b>N</b>	141	<b>Sum Weights</b>	141
<b>Mean</b>	0.87741402	<b>Sum Observations</b>	123.715377
<b>Std Deviation</b>	46.4638738	<b>Variance</b>	2158.89157
<b>Skewness</b>	0.09585474	<b>Kurtosis</b>	0.97181334
<b>Uncorrected SS</b>	302353.37	<b>Corrected SS</b>	302244.82
<b>Coeff Variation</b>	5295.5472	<b>Std Error Mean</b>	3.91296406

Basic Statistical Measures			
Location		Variability	
<b>Mean</b>	0.8774	<b>Std Deviation</b>	46.46387
<b>Median</b>	-1.1000	<b>Variance</b>	2159
<b>Mode</b>	-17.4135	<b>Range</b>	271.28652
		<b>Interquartile Range</b>	54.60000

Tests for Location: Mu0=0				
Test	Statistic		p Value	
<b>Student's t</b>	<b>t</b>	0.224233	<b>Pr &gt;  t </b>	0.8229
<b>Sign</b>	<b>M</b>	-4.5	<b>Pr &gt;=  M </b>	0.5006
<b>Signed Rank</b>	<b>S</b>	-51.5	<b>Pr &gt;=  S </b>	0.9160

Tests for Normality				
Test	Statistic		p Value	
<b>Shapiro-Wilk</b>	<b>W</b>	0.977752	<b>Pr &lt; W</b>	0.0212
<b>Kolmogorov-Smirnov</b>	<b>D</b>	0.080669	<b>Pr &gt; D</b>	0.0235
<b>Cramer-von Mises</b>	<b>W-Sq</b>	0.187511	<b>Pr &gt; W-Sq</b>	0.0078
<b>Anderson-Darling</b>	<b>A-Sq</b>	1.136853	<b>Pr &gt; A-Sq</b>	0.0056

Quantiles (Definition 5)	
Quantile	Estimate
<b>100% Max</b>	130.5865
<b>99%</b>	119.3000
<b>95%</b>	91.6745
<b>90%</b>	50.7000
<b>75% Q3</b>	29.9000
<b>50% Median</b>	-1.1000

**Normality test on residuals  
unstructured correlation assumption**

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**The UNIVARIATE Procedure  
Variable: Resid (Residual)**

Moments			
<b>N</b>	141	<b>Sum Weights</b>	141
<b>Mean</b>	2.68808255	<b>Sum Observations</b>	379.019639
<b>Std Deviation</b>	46.6314288	<b>Variance</b>	2174.49016
<b>Skewness</b>	0.11916932	<b>Kurtosis</b>	1.05114803
<b>Uncorrected SS</b>	305447.458	<b>Corrected SS</b>	304428.622
<b>Coeff Variation</b>	1734.74691	<b>Std Error Mean</b>	3.92707474

Basic Statistical Measures			
Location		Variability	
<b>Mean</b>	2.6881	<b>Std Deviation</b>	46.63143
<b>Median</b>	-1.1000	<b>Variance</b>	2174
<b>Mode</b>	-10.4541	<b>Range</b>	278.24589
		<b>Interquartile Range</b>	55.80000

Tests for Location: $\mu_0=0$				
Test	Statistic		p Value	
<b>Student's t</b>	<b>t</b>	0.6845	<b>Pr &gt;  t </b>	0.4948
<b>Sign</b>	<b>M</b>	-4.5	<b>Pr &gt;=  M </b>	0.5006
<b>Signed Rank</b>	<b>S</b>	183	<b>Pr &gt;=  S </b>	0.7079

Tests for Normality				
Test	Statistic		p Value	
<b>Shapiro-Wilk</b>	<b>W</b>	0.975347	<b>Pr &lt; W</b>	0.0119
<b>Kolmogorov-Smirnov</b>	<b>D</b>	0.079551	<b>Pr &gt; D</b>	0.0273
<b>Cramer-von Mises</b>	<b>W-Sq</b>	0.203494	<b>Pr &gt; W-Sq</b>	<0.0050
<b>Anderson-Darling</b>	<b>A-Sq</b>	1.25304	<b>Pr &gt; A-Sq</b>	<0.0050

Quantiles (Definition 5)	
Quantile	Estimate
<b>100% Max</b>	137.5459
<b>99%</b>	119.3000
<b>95%</b>	97.1833
<b>90%</b>	54.5000
<b>75% Q3</b>	31.1000
<b>50% Median</b>	-1.1000

**Constant Variance test on raw data**

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**The GLM Procedure**

Bartlett's Test for Homogeneity of weight Variance			
Source	DF	Chi-Square	Pr > ChiSq
Diet	3	7.5574	0.0561

**Constant Variance test on raw data**

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**The GLM Procedure**

Levene's Test for Homogeneity of weight Variance ANOVA of Squared Deviations from Group Means					
Source	DF	Sum of Squares	Mean Square	F Value	Pr > F
Diet	3	3.2505E8	1.0835E8	3.90	0.0103
Error	137	3.804E9	27766281		



**R Code:**

```

rm(list=ls())
data('ChickWeight')
#anova on day 18
sub_day18<-subset(ChickWeight,Time==18)
anova_day18<-aov(data = sub_day18, weight ~ Diet)
summary(anova_day18)

sink(' /Users/raymond/Drive/STAT W4201/HW6/anova.txt')
anova_day18
summary(anova_day18)
sink()

#anova adjust by LS mean on day 18
sub_day0<-subset(ChickWeight,Time==0)
sub_day18[, 'birthweight']<-sub_day0$weight[match(sub_day18$Chick,sub_day0$Chick)]
group_mean<-aggregate(sub_day18$birthweight,list(sub_day18$Diet),mean)
colnames(group_mean)<-c('Diet','birthweight_mean')
sub_day18<-merge(sub_day18,group_mean,by.y = 'Diet')
anova_adjust_day18<-aov(data=sub_day18, weight ~ birthweight + Diet)

sink(' /Users/raymond/Drive/STAT W4201/HW6/anova_adjust.txt')
anova_adjust_day18
summary(anova_adjust_day18)
sink()

#LS means
coef<-anova_adjust_day18$coefficients[2]
y_hat<-anova_adjust_day18$fitted.values
y_frame<-data.frame(yhat=y_hat,Diet=sub_day18$Diet)
yhat_mean<-aggregate(y_frame$yhat, list(y_frame$Diet),mean)
colnames(yhat_mean)<-c('Diet','yhat')
mu<-yhat_mean$yhat-coef*(group_mean$birthweight_mean-mean(sub_day18$birthweight))
lsmean_result<-cbind(group_mean,lsmean=mu)
sink(' /Users/raymond/Drive/STAT W4201/HW6/LSmeans.txt')
lsmean_result
sink()
#Or simply use lsmeans function in library lsmeans
library(lsmeans)
lsmean<-lsmeans(anova_adjust_day18,'Diet')
sink(' /Users/raymond/Drive/STAT W4201/HW6/LSmeans1.txt')
lsmean
sink()

library(ggplot2)
#check assumptions
#Normality
#Q-Q plot for residuals
residual<-data.frame(anova_day18[, 'residuals'],sub_day18[, 'Diet'])
normal_plot<-ggplot(data=residual,aes(sample=residuals))+
  stat_qq()+ggtitle('Normal Q-Q of residuals')+
  ylab('residuals')+xlab('Theoretical Quantiles')
resid_quantile<-quantile(residual$residuals,c(0.25, 0.75))
norm_quantile<-qnorm(c(0.25, 0.75))
slope<-diff(resid_quantile)/diff(norm_quantile)
inter<-resid_quantile[1]-slope*norm_quantile[1]
normal_plot+geom_abline(slope=slope,intercept=inter,linetype=2,color='red')
ggsave(filename = ' /Users/raymond/Drive/STAT W4201/HW6/qqplot.png')

#histogram for residuals
histogram_plot<-ggplot(data=residual,aes(residuals))+
  geom_bar(binwidth=10,aes(y=..density..))+ggtitle("Histogram of residuals")+
  ylab('Frequency')+xlab('Residuals')
histogram_plot+geom_density(color='blue')
ggsave(filename = ' /Users/raymond/Drive/STAT W4201/HW6/histogram.png')

#shapiro test on residuals

```

```
shapiro.test(residual$residuals)
sink('/Users/raymond/Drive/STAT W4201/HW6/normalcheck.txt')
shapiro.test(residual$residuals)
sink()

#check constant variance
#bartlett test on data
bartlett.test(x=sub.day18$weight,g=sub.day18$Diet)
sink('/Users/raymond/Drive/STAT W4201/HW6/constantcheck.txt')
bartlett.test(x=sub.day18$weight,g=sub.day18$Diet)
sink()

#levene test on data
library(car)
leveneTest(y=sub.day18$weight,group=as.factor(sub.day18$Diet))

#nonparametric check iid assumptions
#kruskal test on data
kruskal.test(weight~Diet,data=sub.day18)
sink('/Users/raymond/Drive/STAT W4201/HW6/identicalcheck.txt')
kruskal.test(weight~Diet,data=sub.day18)
sink()

#check for parallelism
#two-way anova
anova.inter<-aov(weight ~ birthweight*Diet,data = sub.day18)
summary(anova.inter)
sink('/Users/raymond/Drive/STAT W4201/HW6/paracheck.txt')
summary(anova.inter)
sink()

#repeated measures
#anova on day 10, 18, 21
sub_repeat<-subset(ChickWeight,Time %in% c(10,18,21))
repeated<-aov(weight ~ Diet*Time+Error(Chick),data=sub_repeat)

write.table(sub_repeat, file = '/Users/raymond/Drive/STAT W4201/HW6/ChickWeight.csv',col.names=TRUE,row.names=
FALSE,sep=",")
```

**SAS code:**

```

libname repeat 'C:\Users\QIANBO\Desktop\HW6';
*import data from csv;
proc import out=repeat.repeatdata
  Datafile = 'C:\Users\QIANBO\Desktop\HW6\ChickWeight.csv'
  Dbms =csv replace;
run;
proc print data = repeat.repeatdata;
run;

ods listing close;
ods graphics on;
options papersize=letter;
ods pdf file='C:\Users\QIANBO\Desktop\HW6\compoundsy.pdf';
title 'Repeated Measures on Chick Weight';
title2 'compound symmetry assumption';

*correlation compound symmetry assumption;
proc mixed data = repeat.repeatdata;
class Diet Chick Time;
model weight= Diet Time Diet*Time / outp=repeat.csparam residual;
repeated/type=cs sub=Chick(Diet) rcorr ;
run;
ods pdf close;

ods pdf file='C:\Users\QIANBO\Desktop\HW6\unstructured.pdf';
title 'Repeated Measures on Chick Weight';
title2 'unstructured correlation assumption';
*correlation unstructured assumption;
proc mixed data = repeat.repeatdata;
class Diet Chick Time;
model weight = Diet Time Diet*Time / outp=repeat.unparam residual;
repeated/type=un sub=Chick(Diet) rcorr;
run;
ods pdf close;
ods graphics off;

ods pdf file='C:\Users\QIANBO\Desktop\HW6\normalcs.pdf';
title 'Normality test on residuals';
title2 'compound symmetry assumption';
*test normality;
proc univariate data=repeat.csparam normal;
var resid;
run;
ods pdf close;

ods pdf file='C:\Users\QIANBO\Desktop\HW6\normalun.pdf';
title 'Normality test on residuals';
title2 'unstructured correlation assumption';
*test normality;
proc univariate data=repeat.unparam normal;
var resid;
run;
ods pdf close;

ods pdf file='C:\Users\QIANBO\Desktop\HW6\homo.pdf';
title 'Constant Variance test on raw data';
title2 ;
proc glm data=repeat.repeatdata;
class Diet Chick Time;
model weight = Diet;
means Diet / HOVTEST=BARTLETT;
means Diet / HOVTEST=LEVENE;
run;
ods pdf close;
ods listing ;
title ;

```