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Matrix Notation for Simple Linear Regression (Ch.5)

Simple Linear Regression Model

Recall the simple Linear Regression model:

$$Y_i = \beta_0 + \beta_1 X_i + \varepsilon_i$$

where ε_i 's are independent Normally distributed random variables with mean 0 and variance σ^2 .

Consider writing the observations:

$$Y_1 = \beta_0 + \beta_1 X_1 + \varepsilon_1$$

$$Y_2 = \beta_0 + \beta_1 X_2 + \varepsilon_2$$

$$\vdots$$

$$Y_n = \beta_0 + \beta_1 X_n + \varepsilon_n$$

The above equation group can be written in matrix form:

$$\begin{pmatrix} Y_1 \\ Y_2 \\ \vdots \\ Y_n \end{pmatrix} = \begin{pmatrix} \beta_0 + \beta_1 X_1 \\ \beta_0 + \beta_1 X_2 \\ \vdots \\ \beta_0 + \beta_1 X_n \end{pmatrix} + \begin{pmatrix} \varepsilon_1 \\ \varepsilon_2 \\ \vdots \\ \varepsilon_n \end{pmatrix}$$

$$= \begin{pmatrix} 1 & X_1 \\ 1 & X_2 \\ \vdots & \vdots \\ 1 & X_n \end{pmatrix} \begin{pmatrix} \beta_0 \\ \beta_1 \end{pmatrix} + \begin{pmatrix} \varepsilon_1 \\ \varepsilon_2 \\ \vdots \\ \varepsilon_n \end{pmatrix}.$$

➤ $Y_{n \times 1} = \begin{pmatrix} Y_1 \\ Y_2 \\ \vdots \\ Y_n \end{pmatrix}_{n \times 1}$ be the vector of responses.

➤ $X_{n \times 2} = \begin{bmatrix} 1 & X_1 \\ 1 & X_2 \\ \vdots & \vdots \\ 1 & X_n \end{bmatrix}_{n \times 2}$ be the design matrix.

➤ $\beta_{2 \times 1} = \begin{pmatrix} \beta_0 \\ \beta_1 \end{pmatrix}_{2 \times 1}$ be the vector of parameters.

➤ $\varepsilon_{n \times 1} = \begin{pmatrix} \varepsilon_1 \\ \varepsilon_2 \\ \vdots \\ \varepsilon_n \end{pmatrix}_{n \times 1}$ be the vector of error terms is:

The simple linear regression model can then be written as:

$$Y_{n \times 1} = X_{n \times 2} \beta_{2 \times 1} + \varepsilon_{n \times 1}$$

$$Y = X\beta + \varepsilon$$

Estimations

$$\hat{\beta}_{2 \times 1} = b_{2 \times 1} = \begin{pmatrix} b_0 \\ b_1 \end{pmatrix}_{2 \times 1} = (X_{n \times 2}^T X_{n \times 2})^{-1} (X_{n \times 2}^T Y_{n \times 1}) = (X^T X)^{-1} (X^T Y),$$

and

$$\text{cov}(b_{2 \times 1}) = \sigma^2 (X^T X)^{-1}.$$

The variance σ^2 can be estimated by MSE. $\widehat{\sigma^2} = MSE$.

Prediction

$$\hat{Y} = \begin{pmatrix} \hat{Y}_1 \\ \hat{Y}_2 \\ \vdots \\ \hat{Y}_n \end{pmatrix} = \begin{pmatrix} b_0 + b_1 X_1 \\ b_0 + b_1 X_2 \\ \vdots \\ b_0 + b_1 X_n \end{pmatrix} = \begin{pmatrix} 1 & X_1 \\ 1 & X_2 \\ \vdots & \vdots \\ 1 & X_n \end{pmatrix} \begin{pmatrix} b_0 \\ b_1 \end{pmatrix} = Xb$$

To predict Y at X_h , $\hat{Y}_h = \begin{pmatrix} 1 & X_h \end{pmatrix} \begin{pmatrix} b_0 \\ b_1 \end{pmatrix}$.

The Hat Matrix

Note that $\hat{Y} = Xb$, and $b = (X^T X)^{-1} (X^T Y)$. Hence,

$$\hat{Y} = Xb = X (X^T X)^{-1} (X^T Y) = \left(X (X^T X)^{-1} X^T \right) Y = HY$$

where $H = X(X^T X)^{-1} X^T$ is called the “**hat matrix**.”

Partitioning Sum of Squares: ANOVA

Let $I_{n \times n}$ be a diagonal matrix with 1's on the diagonal and 0 everywhere else (i.e., Identity matrix). Let $J_{n \times n}$ be a matrix of 1's. Then:

- The sum of squares explained by the regression: $SSR = \sum (\hat{Y}_i - \bar{Y})^2 = Y'(H - \frac{1}{n}J)Y$
- The sum of squares of error: $SSE = \sum (Y_i - \hat{Y}_i)^2 = Y'(I - H)Y$
- Total sum of squares: $SSTO = \sum (Y_i - \bar{Y})^2 = Y'(I - \frac{1}{n}J)Y$
- $df_R = 1 = \text{rank} \left(H - \frac{1}{n}J \right)$
- $df_E = n - 2 = \text{rank} (I - H)$
- $df_{TO} = n - 1 = \text{rank} \left(I - \frac{1}{n}J \right)$

Recall that: $SSR + SSE = SSTO$, $df_R + df_E = df_{TO}$, and $MSR = SSR/df_R$, $MSE = SSE/df_E$.

Multiple Linear Regression (Ch. 6)

The work is very similar to simple linear regression.

Multiple Linear Regression Model

$$\begin{pmatrix} Y_1 \\ Y_2 \\ \vdots \\ Y_n \end{pmatrix} = \begin{pmatrix} \beta_0 + \beta_1 X_{1,1} + \beta_2 X_{1,2} + \cdots + \beta_{p-1} X_{1,p-1} \\ \beta_0 + \beta_1 X_{2,1} + \beta_2 X_{2,2} + \cdots + \beta_{p-1} X_{2,p-1} \\ \vdots \\ \beta_0 + \beta_1 X_{n,1} + \beta_2 X_{n,2} + \cdots + \beta_{p-1} X_{n,p-1} \end{pmatrix} + \begin{pmatrix} \varepsilon_1 \\ \varepsilon_2 \\ \vdots \\ \varepsilon_n \end{pmatrix}$$

$$= \begin{pmatrix} 1 & X_{1,1} & X_{1,2} & \cdots & X_{1,p-1} \\ 1 & X_{2,1} & X_{2,2} & \cdots & X_{2,p-1} \\ \vdots & \vdots & \vdots & \cdots & \vdots \\ 1 & X_{n,1} & X_{n,2} & \cdots & X_{n,p-1} \end{pmatrix} \begin{pmatrix} \beta_0 \\ \beta_1 \\ \beta_2 \\ \vdots \\ \beta_{p-1} \end{pmatrix} + \begin{pmatrix} \varepsilon_1 \\ \varepsilon_2 \\ \vdots \\ \varepsilon_n \end{pmatrix}.$$

➤ $X_{n \times p} = \begin{pmatrix} 1 & X_{1,1} & X_{1,2} & \cdots & X_{1,p-1} \\ 1 & X_{2,1} & X_{2,2} & \cdots & X_{2,p-1} \\ \vdots & \vdots & \vdots & \cdots & \vdots \\ 1 & X_{n,1} & X_{n,2} & \cdots & X_{n,p-1} \end{pmatrix}$ be the design matrix.

➤ $\beta_{p \times 1} = \begin{pmatrix} \beta_0 \\ \beta_1 \\ \beta_2 \\ \vdots \\ \beta_{p-1} \end{pmatrix}$ be the vector of parameters.

The multiple linear regression model can then be written as:

$$Y_{n \times 1} = X_{n \times p} \beta_{p \times 1} + \varepsilon_{n \times 1}$$

$$Y = X \beta + \varepsilon$$

Estimations

$$\hat{\beta}_{p \times 1} = b_{p \times 1} = \begin{pmatrix} b_0 \\ b_1 \\ b_2 \\ \vdots \\ b_{p-1} \end{pmatrix}_{p \times 1} = (X_{n \times p}^T X_{n \times p})^{-1} (X_{n \times p}^T Y_{n \times 1}) = (X^T X)^{-1} (X^T Y),$$

and

$$\text{cov}(b) = \sigma^2 (X^T X)^{-1}.$$

The variance σ^2 can be estimated by MSE. $\widehat{\sigma^2} = MSE$

Prediction

- The fitted values (for observations in the data set):

$$\begin{aligned}\hat{Y} = \begin{pmatrix} \hat{Y}_1 \\ \hat{Y}_2 \\ \vdots \\ \hat{Y}_n \end{pmatrix} &= \begin{pmatrix} b_0 + b_1X_{1,1} + b_2X_{1,2} + \cdots + b_{p-1}X_{1,p-1} \\ b_0 + b_1X_{2,1} + b_2X_{2,2} + \cdots + b_{p-1}X_{2,p-1} \\ \vdots \\ b_0 + b_1X_{n,1} + b_2X_{n,2} + \cdots + b_{p-1}X_{n,p-1} \end{pmatrix} \\ &= \begin{pmatrix} 1 & X_{1,1} & X_{1,2} & \cdots & X_{1,p-1} \\ 1 & X_{2,1} & X_{2,2} & \cdots & X_{2,p-1} \\ \vdots & \vdots & \vdots & \cdots & \vdots \\ 1 & X_{n,1} & X_{n,2} & \cdots & X_{n,p-1} \end{pmatrix} \begin{pmatrix} b_0 \\ b_1 \\ b_2 \\ \vdots \\ b_{p-1} \end{pmatrix} = Xb.\end{aligned}$$

- The predicted value (aka. the estimate of Y, or the fitted value) for a new set of x-values: $(X_{h,1} \ X_{h,2} \ \cdots \ X_{h,p-1})$.

$$\hat{Y}_h = (1 \ X_{h,1} \ X_{h,2} \ \cdots \ X_{h,p-1})b.$$

Partitioning Sum of Squares: ANOVA

Let $H = X(X^T X)^{-1} X^T$ denote the hat matrix. Let $I_{n \times n}$ be a diagonal matrix with 1's on the diagonal and 0 everywhere else (i.e., Identity matrix). Let $J_{n \times n}$ be a matrix of 1's. Then:

- The sum of squares explained by the regression: $SSR = \sum (\hat{Y}_i - \bar{Y})^2 = Y'(H - \frac{1}{n}J)Y$
- The sum of squares of error: $SSE = \sum (Y_i - \hat{Y}_i)^2 = Y'(I - H)Y$
- Total sum of squares: $SSTO = \sum (Y_i - \bar{Y})^2 = Y'(I - \frac{1}{n}J)Y$
- $df_R = p - 1 = \text{rank}(H - \frac{1}{n}J)$
- $df_E = n - p = \text{rank}(I - H)$
- $df_{TO} = n - 1 = \text{rank}(I - \frac{1}{n}J)$

Recall that: $SSR + SSE = SSTO$, $df_R + df_E = df_{TO}$, and $MSR = SSR/df_R$, $MSE = SSE/df_E$.