

Quantitative Analytic Techniques (Class #12)

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1. Growth curve analysis

1. Growth curve analysis

An example of growth curves

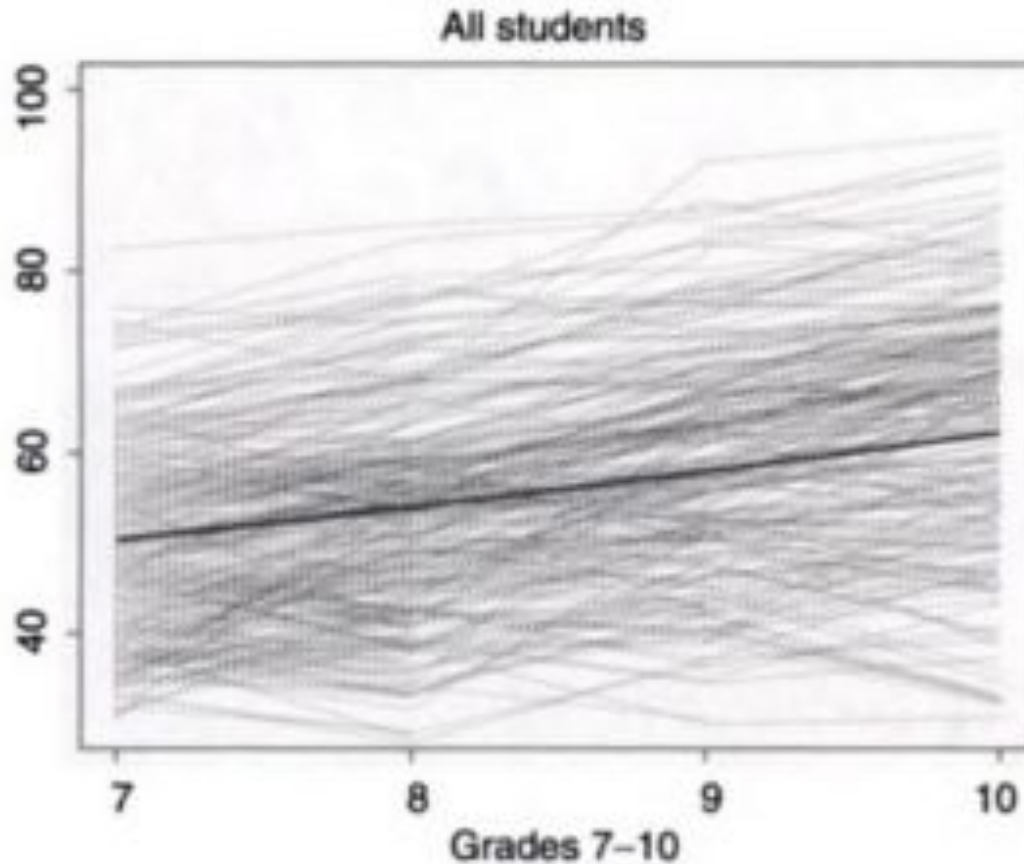
- Muthén, Bengt. "Latent variable analysis." *The Sage handbook of quantitative methodology for the social sciences*. Thousand Oaks, CA: Sage Publications (2004): 345-68.



Time often is necessary to see how people sort out

- Many processes take a long time to have effects
- Trajectories can converge or diverge
- We want to study groups of people as they change over time in similar and different ways
- Do certain events change people's trajectories up or down?

The substantive issue

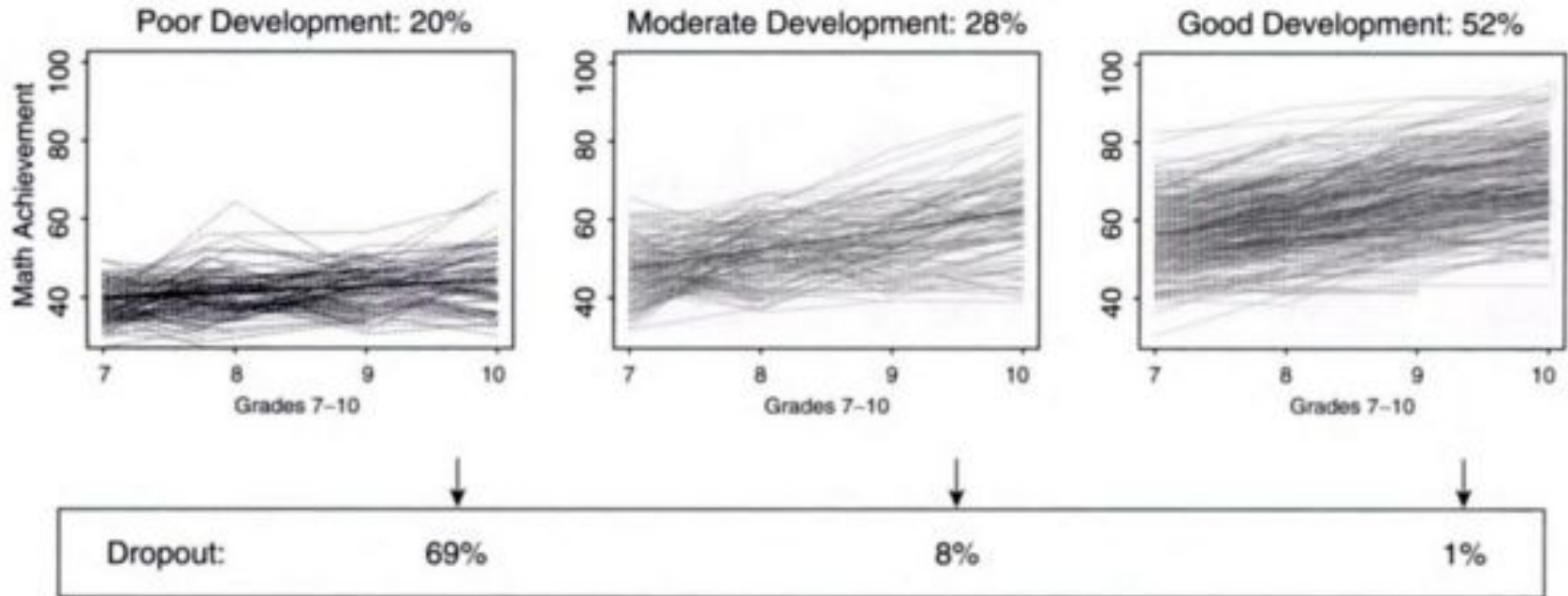


- Growth in math scores for Longitudinal Study of Youth
- Lots of variability! Lots

The question

Are there typical ways that math scores move? Can we see them?

Three basic patterns emerge



- Some covariates predict which path students are on and then that path corresponds with some outcomes, too, like drop-out

Why do growth curve analysis?

Why do growth curves I?

- We want to model the shape of growth over time
- One step more complicated than merely plotting the values over time
- How much does everyone follow the same path or do they deviate from it?
- Similar to trend analysis over pooled cross-sectional data, but now, we have the same people over time (i.e., panel data)

Why do growth curves II?

- We want to compare the starting points and trajectories (and final destinations) of groups over time
- We are usually looking for diverging trajectories
- Or converging ones
- Or ones with stable differences over time

Why do growth curves II?

- We want to see if there are inflection points, turning points, tipping points, etc. when something happens in the course of someone's life:
- How to criminal trajectories change when someone gets married?
- How do children's depression trajectories change after their parents get divorced?
- Etc.

What makes growth curves unique?

What makes growth curves unique I?

- We are modeling a full trajectory (a whole series of Y values over time) for each person as our dependent variable.
- Compare that to first differences, where we broke everything down to year-to-year changes only
- Growth curves take a longer-term perspective

What makes growth curves unique II?

- Tricky to think about how to incorporate time-varying characteristics, since we are modeling the whole growth curve
- Most unchanging variables can be entered “at baseline” – no problem
- Unchanging variables can change in importance over time; i.e., an interaction of that variable with time

What makes growth curves unique III?

- Must think hard about cumulative processes that take place over time
- How does one time period affect the next?

What makes growth curves unique IV?

- How much do we deal with individual heterogeneity and how much do we deal with overall patterns?

About growth curves

Growth curves can be modeled in a few ways

- Structural equation modeling (SEM)
- Hierarchical linear modeling (HLM)
- MANOVA

2. How do I do growth curve analysis?

How to do growth curves I

1. Look at the overall trend of Y over time
2. Plot the growth curves for each person
3. You will include time, but do you need a quadratic too?

How to do growth curves II

4. Do you need separate intercepts for each person? [*Almost always.*]
5. Do you need separate slopes (on time) for each person as well? [*Usually.*]
6. Do you need other Xs in the model? [*Sure.*]
7. Do you need interactions of some Xs with time? [*Hmmmm?*]

How to do growth curves I

1. Look at the overall trend of Y over time

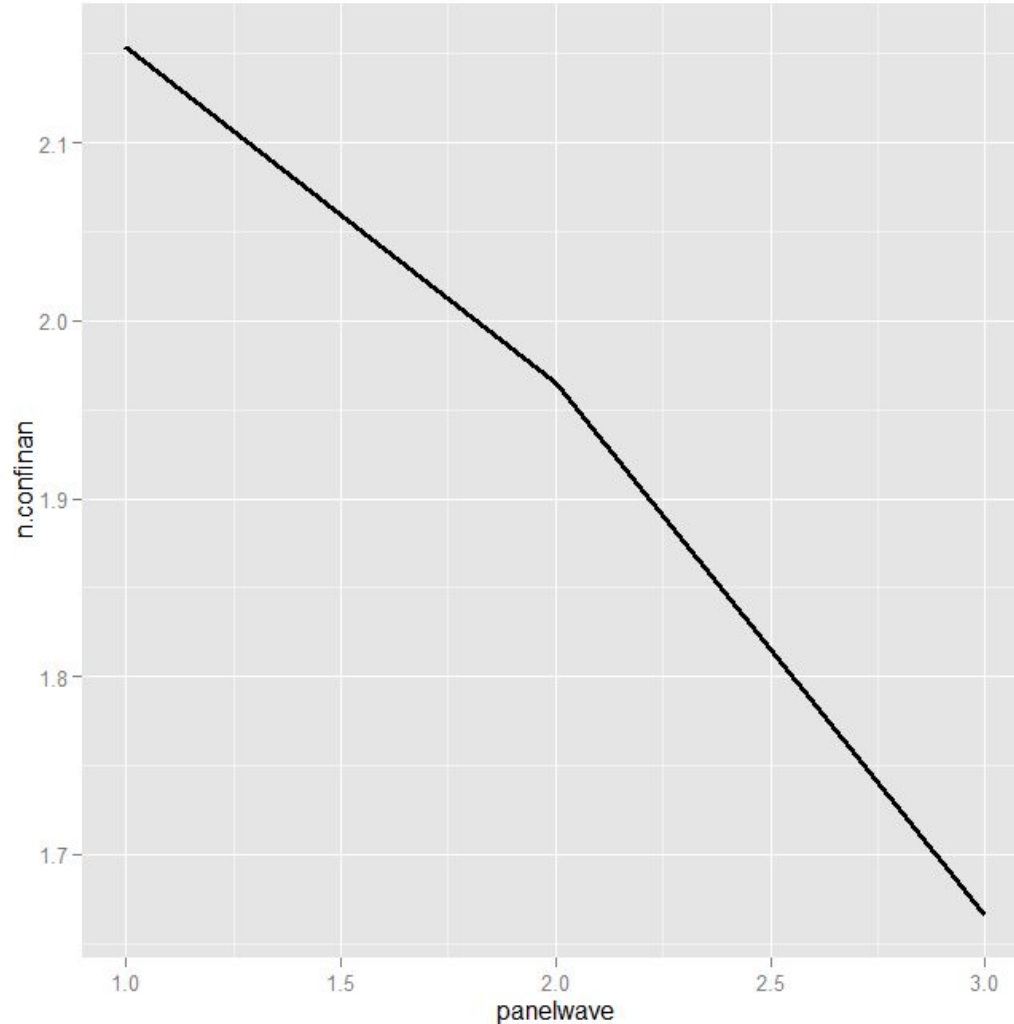
Setting up the data

```
library(QMSS)
library(ggplot2)

pd=read.csv(file.choose()) ### choose the GSS panel ###
```


Overall trend in confidence in banks

This is from Wave 1 (2006) to Wave 3 (2010)



How did I do this?

```
vars <- c("idnum", "panelwave", "sex", "age", "educ", "race", "polviews", "confinan")
sub <- pd[, vars]
sub$n.confinan <- ReverseThis(sub$confinan)

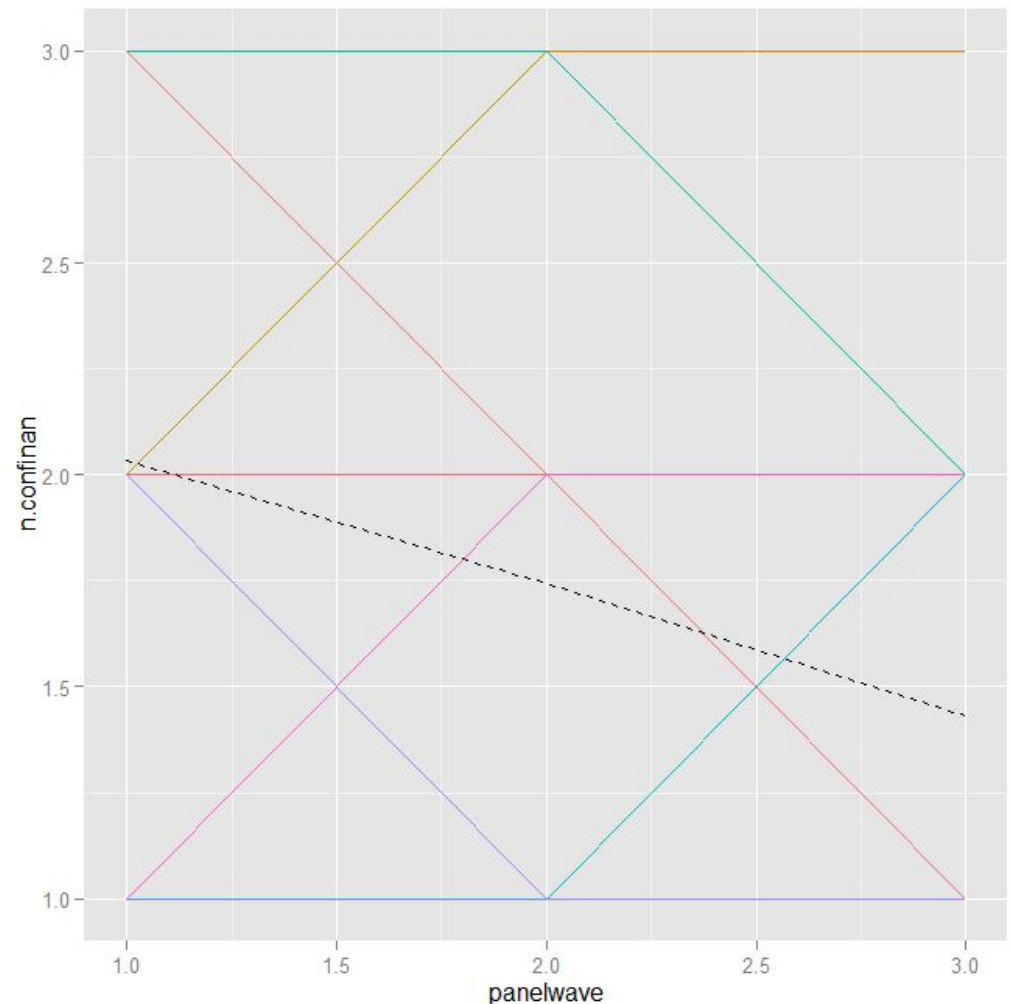
# Overall trend in confidence in banks
g_trend <- ggplot(sub, aes(x = panelwave, y = n.confinan))
(g_trend <- g_trend + stat_summary(fun.y=mean, geom="line", lwd = 1.25))
```

How to do growth curves I

2. Plot the growth curves for each person

Empirical growth curves

This is from Wave 1 (2006) to Wave 3 (2010); for the first hundred+ cases



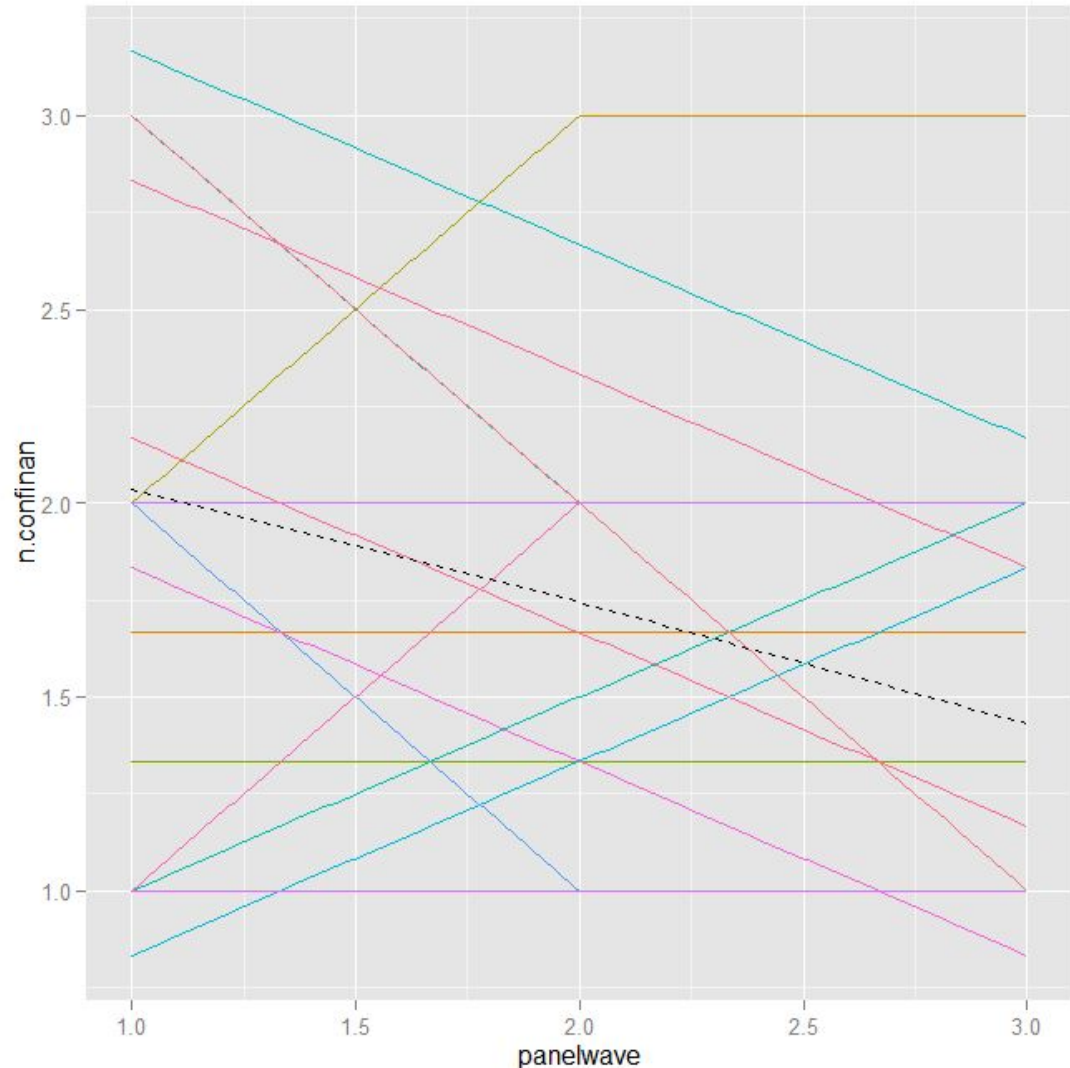
How did I do this?

```
# Empirical growth curves for idnum < 200 (& overall)
g_growth <- ggplot(subset(sub, idnum<200),
                  aes(x = panelwave, y = n.confinan, group = idnum, color =
factor(idnum)))
no_legend <- theme(legend.position="none")

g_id <- g_growth + geom_line() + no_legend
g_id + stat_summary(fun.y=mean, geom="line", aes(group=1), lty = 2, color="black")
```

Individual regressions for each person

This is from Wave 1 (2006) to Wave 3 (2010); for the first hundred+ cases



How did I do this?

```
# individual regression lines for idnum < 200 (& overall)
g_reg <- g_growth + stat_smooth(method = lm, se = F) + no_legend
g_reg + stat_summary(fun.y=mean, geom="smooth", aes(group=1), lty = 2, color = "black")
```

What do we see here?

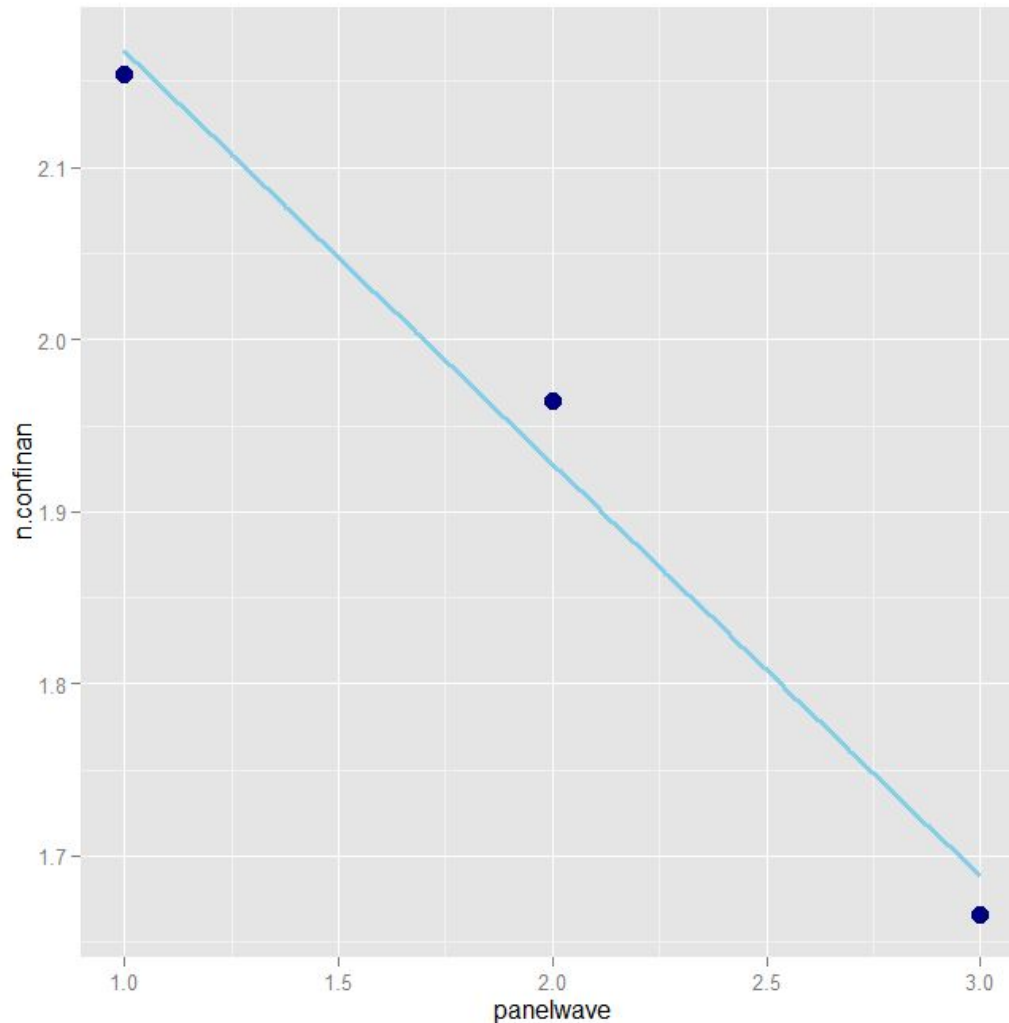
There also looks like a lot of individual variation over time regarding their opinions on banks

How to do growth curves I

3. You will include time, but do you need a quadratic too?

Overall linear prediction is this

Not a bad fit here

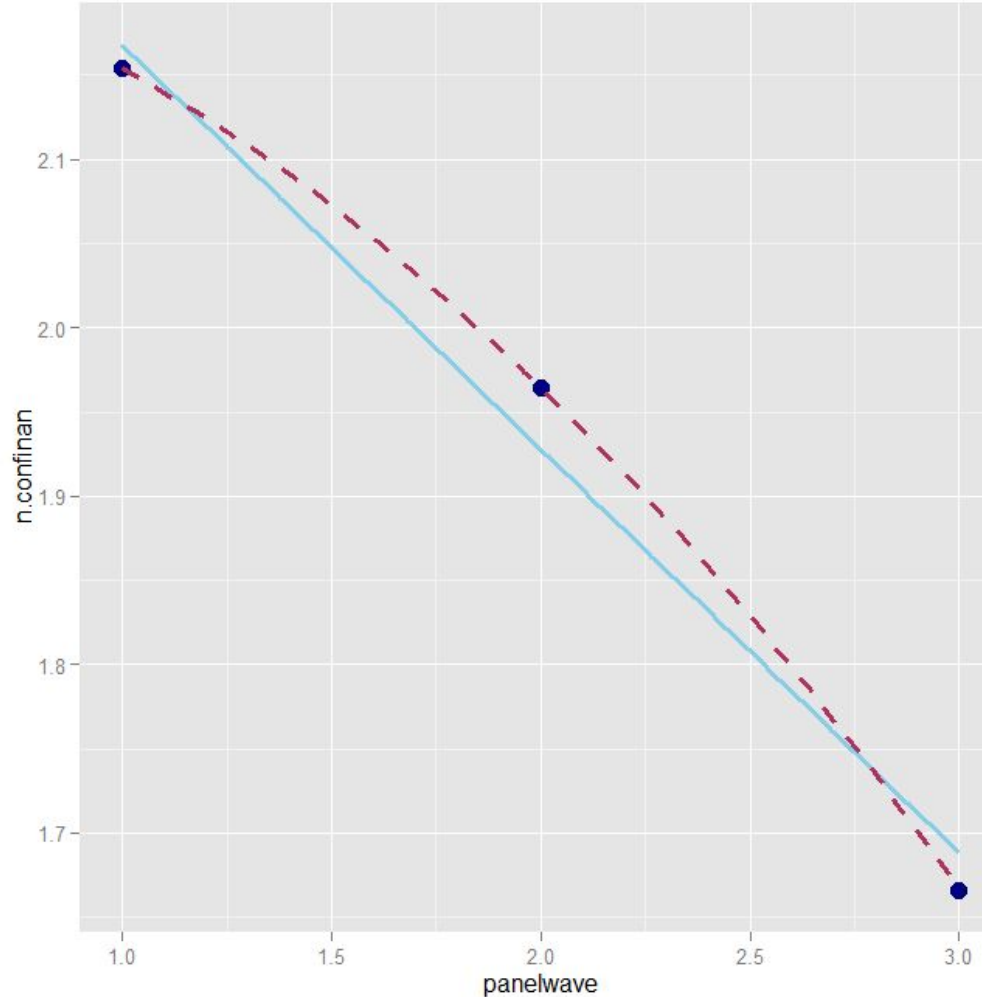


How did I do this?

```
# overall linear prediction
g_lm <- ggplot(sub, aes(x = panelwave, y = n.confinan))
g_lm <- g_lm + stat_summary(fun.y=mean, geom="point", aes(group=1), size=4, color =
"navyblue")
g_lm <- g_lm + stat_smooth(method = lm, se = F, color = "skyblue", lwd = 1.25)
g_lm
```

Overall quadratic prediction is this

This is an amazing fit, but still pretty close to linear ...



How did I do this?

```
# add quadratic prediction curve
g_lm <- g_lm + stat_smooth(formula = y ~ poly(x,2), method = lm, se = F,
                           color = "maroon", lty = 2, lwd = 1.25)
g_lm
```

How to do growth curves II

4. Do you need separate intercepts for each person? [*Almost always.*]

Model with no separate intercepts ...

Just an OLS with clustered and robust standard errors; at Wave 2, confidence in banks has dropped -.189, relative to Wave 1 – and at Wave 3, it has dropped -.487

```
library(plyr)
library(psych)
library(multcomp)
library(rms)
library(lme4)
```

```
# ols with clustered & robust SEs
robcov(ols(n.confinan ~ factor(panelwave), x = T, y = T, data = sub),
       cluster = sub$idnum)
```

		Model Likelihood Ratio Test	Discrimination Indexes
Obs	3163	LR chi2 280.89	R2 0.085
sigma	0.6433	d.f. 2	R2 adj 0.084
d.f.	3160	Pr(> chi2) 0.0000	g 0.208
Clusters	1324		

	Coef	S.E.	t	Pr(> t)
Intercept	2.1538	0.0180	119.83	<0.0001
panelwave=2	-0.1893	0.0229	-8.26	<0.0001
panelwave=3	-0.4876	0.0252	-19.34	<0.0001

Model with no separate intercepts ...

Just an OLS with clustered and robust standard errors;
with a continuous measure of time, this means that
confidence in banks drops -.239 points per Wave

```
robcov(ols(n.confinan ~ panelwave, x = T, y = T, data = sub), cluster = sub$idnum)
```

		Model Likelihood Ratio Test	Discrimination Indexes
Obs	3163	LR chi2 276.04	R2 0.084
sigma	0.6436	d.f. 1	R2 adj 0.083
d.f.	3161	Pr(> chi2) 0.0000	g 0.210
Clusters	1324		

	Coef	S.E.	t	Pr(> t)
Intercept	2.4077	0.0269	89.62	<0.0001
panelwave	-0.2398	0.0126	-19.03	<0.0001

Model with squared time ...

Just an OLS with clustered and robust standard errors;
confidence in banks starts to drop only a little and then
drops much further too

```
sub$panelwavesq = sub$panelwave^2
```

```
robcov(ols(n.confinan ~ panelwave + panelwavesq, x = T, y = T, data = sub), cluster =  
sub$idnum)
```

			Model Likelihood Ratio Test		Discrimination Indexes	
Obs	3163		LR chi2	280.89	R2	0.085
sigma	0.6433		d.f.	2	R2 adj	0.084
d.f.	3160		Pr(> chi2)	0.0000	g	0.208
Clusters	1324					

	Coef	S.E.	t	Pr(> t)
Intercept	2.2343	0.0690	32.40	<0.0001
panelwave	-0.0260	0.0801	-0.32	0.7453
panelwavesq	-0.0544	0.0200	-2.72	0.0065

Now, the random intercept model

The random intercept model is this:

$$y_i = \alpha_{i[j]} + \beta x_i + e_i$$

which provides a unique intercept for each person j

Our random intercept model

The random intercept model is this:

$$\text{nconf}_{ij} = \alpha_{i[j]} + \beta_1 \text{time}_{ij} + \beta_2 \text{time}_{ij}^2 + e_{ij}$$

which provides a unique intercept for each person j

A model with random intercepts ...

By typing “1 | idnum” this is like saying, for each unique idnum, give me a random parameter for each variable listed. If nothing is specified, it just gives for each idnum, a unique intercept

```
lmer.confinan <- lmer(n.confinan ~ panelwave + (1|idnum), data = sub, REML = F)
```

Linear mixed model fit by maximum likelihood ['lmerMod']

AIC	BIC	logLik	deviance	df.resid
5959.8	5984.1	-2975.9	5951.8	3159

Random effects:

Groups	Name	Variance	Std.Dev.
idnum	(Intercept)	0.1287	0.3588
Residual		0.2852	0.5341

Number of obs: 3163, groups: idnum, 1324

Fixed effects:

	Estimate	Std. Error	t value
(Intercept)	2.41046	0.02582	93.34
panelwave	-0.24151	0.01210	-19.95

Correlation of Fixed Effects:

	(Intr)
panelwave	-0.841

A model with random intercepts ...

Give me a random parameter for each variable listed. If nothing is specified, it just gives for each idnum, a unique intercept

```
> lmer.confinan <- lmer(n.confinan ~ panelwave + (1|idnum), data = sub, REML = F)
```

```
> summary(lmer.confinan)
```

```
Linear mixed model fit by maximum likelihood ['lmerMod']
```

```
Formula: n.confinan ~ panelwave + (1 | idnum)
```

```
Data: sub
```

AIC	BIC	logLik	deviance	df.resid
5959.8	5984.1	-2975.9	5951.8	3159

```
Scaled residuals:
```

Min	1Q	Median	3Q	Max
-2.26693	-0.64452	-0.03549	0.50991	2.74135

```
Random effects:
```

Groups	Name	Variance	Std.Dev.
idnum	(Intercept)	0.1287	0.3588
Residual		0.2852	0.5341

Number of obs: 3163, groups: idnum, 1324

```
Fixed effects:
```

	Estimate	Std. Error	t value
(Intercept)	2.41046	0.02582	93.34

A model with random intercepts ...

Confidence in banks drops -.242 points per Wave (slightly different from OLS earlier)

```
lmer.confinan <- lmer(n.confinan ~ panelwave + (1|idnum), data = sub, REML = F)
```

Linear mixed model fit by maximum likelihood ['lmerMod']

AIC	BIC	logLik	deviance	df.resid
5959.8	5984.1	-2975.9	5951.8	3159

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Correlation of Fixed Effects:

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panelwave	-0.841

What is all this “Random-effects” at bottom?

- *Std.Dev.(idnum)* is how much individual variation there is in the constant, once everyone gets their own intercept
- I.e., the constant has a mean of 2.41 and now a sd of 0.359

```
lmer.confinan <- lmer(n.confinan ~ panelwave + (1|idnum), data = sub, REML = F)
```

Linear mixed model fit by maximum likelihood ['lmerMod']

AIC	BIC	logLik	deviance	df.resid
5959.8	5984.1	-2975.9	5951.8	3159

Random effects:

Groups	Name	Variance	Std.Dev.
idnum	(Intercept)	0.1287	0.3588
Residual		0.2852	0.5341

Number of obs: 3163, groups: idnum, 1324

Fixed effects:

	Estimate	Std. Error	t value
(Intercept)	2.41046	0.02582	93.34
panelwave	-0.24151	0.01210	-19.95

Correlation of Fixed Effects:

	(Intr)
panelwave	-0.841

What is all this “Random-effects” at bottom?

- Is giving everyone their own intercept better than OLS?
Let's compare AICs between the OLS model (run as a GLS) and the random-intercept one.
- Random-intercept AIC (i.e., 5959.8) < OLS (i.e., 6192.9), so random-intercept is an improved fit.

```
> glm.confinan = glm(n.confinan ~ panelwave, data=sub)
> summary(glm.confinan)
```

Coefficients:

	Estimate	Std. Error	t value	Pr(> t)
(Intercept)	2.40768	0.02851	84.44	<2e-16 ***
panelwave	-0.23978	0.01412	-16.98	<2e-16 ***

Signif. codes: 0 '***' 0.001 '**' 0.01 '*' 0.05 '.' 0.1 ' ' 1

(Dispersion parameter for gaussian family taken to be 0.414281)

Null deviance: 1429.0 on 3162 degrees of freedom
Residual deviance: 1309.5 on 3161 degrees of freedom
(2837 observations deleted due to missingness)

AIC: 6192.9

What is all this “Random-effects” at bottom?

This is just the calculation of Rho again.

$$\begin{aligned}\text{Rho} &= \sigma_u^2 / (\sigma_u^2 + \sigma_e^2) \\ &= (.36^2) / (.36^2 + .53^2) \\ &= 0.31\end{aligned}$$

```
-- from earlier --
```

```
Random effects:
```

Groups	Name	Variance	Std.Dev.
idnum	(Intercept)	0.1287	0.3588
Residual		0.2852	0.5341

```
> rho(lmer.confinan)
[1] 0.3109777
```

We have already estimated this model

Using *plm*, with option="random": that gives same coefficients and same Rho as random intercepts model

Random intercepts but quadratic on time

This may have improved the fit marginally (lower AIC);
confidence in banks goes down and down even faster

```
> lmer.confinan2 <- update(lmer.confinan, ~ . + I(panelwave^2))  
> summary(lmer.confinan2)
```

AIC	BIC	logLik	deviance	df.resid
5954.3	5984.6	-2972.1	5944.3	3158

Random effects:

Groups	Name	Variance	Std.Dev.
idnum	(Intercept)	0.1293	0.3596
Residual		0.2840	0.5329

Number of obs: 3163, groups: idnum, 1324

Fixed effects:

	Estimate	Std. Error	t value
(Intercept)	2.22899	0.07077	31.496
panelwave	-0.01744	0.08226	-0.212
I(panelwave^2)	-0.05700	0.02070	-2.754

Correlation of Fixed Effects:

	(Intr)	panlww
panelwave	-0.966	
I(panlww^2)	0.931	-0.989

```
> rho(lmer.confinan2)  
[1] 0.3129095
```

How to do growth curves II

5. Do you need separate slopes (on time) for each person as well? [*Usually.*]

The equations -

The random intercept and random coefficient model is this:

$$y_i = \alpha_{i[j]} + \beta_{i[j]} x_i + u_{\alpha i} + u_{\beta i[j]} + e_i$$

Random intercepts + *random slopes*

A lot going on here. 1) By typing *panelwave* after the “1 +” this is like saying, allow *panelwave* to have a random coefficient (slope) for each idnum

```
> lmer.confinan3 <- lmer(n.confinan ~ panelwave + (1 + panelwave | idnum), data = sub, REML = F)
> summary(lmer.confinan3)
```

AIC	BIC	logLik	deviance	df.resid
5956.7	5993.1	-2972.3	5944.7	3157

Random effects:

Groups	Name	Variance	Std.Dev.	Corr
idnum	(Intercept)	0.24955	0.4995	
	panelwave	0.02236	0.1495	-0.69
Residual		0.26280	0.5126	

Number of obs: 3163, groups: idnum, 1324

Fixed effects:

	Estimate	Std. Error	t value
(Intercept)	2.41022	0.02680	89.94
panelwave	-0.24128	0.01245	-19.38

Correlation of Fixed Effects:

	(Intr)
panelwave	-0.854

Random intercepts + *random slopes*

The covariance between each idnum's intercept and slope can be unstructured; we don't want the default to be covariance of zero

```
> lmer.confinan3 <- lmer(n.confinan ~ panelwave + (1 + panelwave | idnum), data = sub, REML = F)
> summary(lmer.confinan3)
```

AIC	BIC	logLik	deviance	df.resid
5956.7	5993.1	-2972.3	5944.7	3157

Random effects:

Groups	Name	Variance	Std.Dev.	Corr
idnum	(Intercept)	0.24955	0.4995	
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Residual		0.26280	0.5126	

Number of obs: 3163, groups: idnum, 1324

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	Estimate	Std. Error	t value
(Intercept)	2.41022	0.02680	89.94
panelwave	-0.24128	0.01245	-19.38

Correlation of Fixed Effects:

	(Intr)
panelwave	-0.854

Random intercepts + *random slopes*

Interpretation of constant. At Wave 0, on average, people have a **2.41** score, but there is substantial variation around that mean, with a st dev of **.499**

```
> lmer.confinan3 <- lmer(n.confinan ~ panelwave + (1 + panelwave | idnum), data = sub, REML = F)
> summary(lmer.confinan3)
```

AIC	BIC	logLik	deviance	df.resid
5956.7	5993.1	-2972.3	5944.7	3157

Random effects:

Groups	Name	Variance	Std.Dev.	Corr
idnum	(Intercept)	0.24955	0.4995	
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Residual		0.26280	0.5126	

Number of obs: 3163, groups: idnum, 1324

Fixed effects:

	Estimate	Std. Error	t value
(Intercept)	2.41022	0.02680	89.94
panelwave	-0.24128	0.01245	-19.38

Correlation of Fixed Effects:

	(Intr)
panelwave	-0.854

Random intercepts + *random slopes*

Interpretation of slope. On average, people's scores drop by **-.24** per Wave, but there is substantial variation around that average slope, with a st dev of **.149**

```
> lmer.confinan3 <- lmer(n.confinan ~ panelwave + (1 + panelwave | idnum), data = sub, REML = F)
> summary(lmer.confinan3)
```

AIC	BIC	logLik	deviance	df.resid
5956.7	5993.1	-2972.3	5944.7	3157

Random effects:

Groups	Name	Variance	Std.Dev.	Corr
idnum	(Intercept)	0.24955	0.4995	
	panelwave	0.02236	0.1495	-0.69
Residual		0.26280	0.5126	

Number of obs: 3163, groups: idnum, 1324

Fixed effects:

	Estimate	Std. Error	t value
(Intercept)	2.41022	0.02680	89.94
panelwave	-0.24128	0.01245	-19.38

Correlation of Fixed Effects:

	(Intr)
panelwave	-0.854

Random intercepts + *random slopes*

Interpretation of `corr(panel~ve,_cons)` I. There is a high negative correlation between the constant and the slope for each idnum ($\rho = -.69$).

```
> lmer.confinan3 <- lmer(n.confinan ~ panelwave + (1 + panelwave | idnum), data = sub, REML = F)
> summary(lmer.confinan3)
```

AIC	BIC	logLik	deviance	df.resid
5956.7	5993.1	-2972.3	5944.7	3157

Random effects:

Groups	Name	Variance	Std.Dev.	Corr
idnum	(Intercept)	0.24955	0.4995	
	panelwave	0.02236	0.1495	-0.69
Residual		0.26280	0.5126	

Number of obs: 3163, groups: idnum, 1324

Fixed effects:

	Estimate	Std. Error	t value
(Intercept)	2.41022	0.02680	89.94
panelwave	-0.24128	0.01245	-19.38

Correlation of Fixed Effects:

	(Intr)
panelwave	-0.854

Random intercepts + *random slopes*

Interpretation of `corr(panel~ve,_cons)` II. When people started out with a high `nconfin` score, they fell more rapidly over time; conversely, those with low initial `nconfin` scores *increased* them rapidly

```
> lmer.confinan3 <- lmer(n.confinan ~ panelwave + (1 + panelwave | idnum), data = sub, REML = F)
```

AIC	BIC	logLik	deviance	df.resid
5956.7	5993.1	-2972.3	5944.7	3157

Random effects:

Groups	Name	Variance	Std.Dev.	Corr
idnum	(Intercept)	0.24955	0.4995	
	panelwave	0.02236	0.1495	-0.69
Residual		0.26280	0.5126	

Number of obs: 3163, groups: idnum, 1324

Fixed effects:

	Estimate	Std. Error	t value
(Intercept)	2.41022	0.02680	89.94
panelwave	-0.24128	0.01245	-19.38

Correlation of Fixed Effects:

	(Intr)
panelwave	-0.854

Are random slopes necessary?

We can determine if random slopes better fit our data by comparing the likelihood ratio from the random intercept model (LL=-2976) with the LR from the random intercept + random slopes model (LL=-2972)

```
> anova(lmer.confinan, lmer.confinan3)
```

Data: sub

Models:

lmer.confinan: n.confinan ~ panelwave + (1 | idnum)

lmer.confinan3: n.confinan ~ panelwave + (1 + panelwave | idnum)

	Df	AIC	BIC	logLik	deviance	Chisq	Chi	Df	Pr(>Chisq)
lmer.confinan	4	5959.8	5984.1	-2975.9	5951.8				
lmer.confinan3	6	5956.7	5993.1	-2972.3	5944.7	7.1503		2	0.02801 *

Signif. codes: 0 '***' 0.001 '**' 0.01 '*' 0.05 '.' 0.1 ' ' 1

Are random slopes necessary?

The random intercept + random slopes model provides a superior fit to the data than the random intercepts model ($p < .05$). But the AIC change is pretty minimal (=3 points)

```
> anova(lmer.confinan, lmer.confinan3)
```

```
Data: sub
```

```
Models:
```

```
lmer.confinan: n.confinan ~ panelwave + (1 | idnum)
```

```
lmer.confinan3: n.confinan ~ panelwave + (1 + panelwave | idnum)
```

	Df	AIC	BIC	logLik	deviance	Chisq	Chi	Df	Pr(>Chisq)
lmer.confinan	4	5959.8	5984.1	-2975.9	5951.8				
lmer.confinan3	6	5956.7	5993.1	-2972.3	5944.7	7.1503		2	0.02801 *

```
---
```

```
Signif. codes:  0 '***' 0.001 '**' 0.01 '*' 0.05 '.' 0.1 ' ' 1
```

How to do growth curves II

6. Do you need other Xs in the model? [Sure.]

Add in time-invariant characteristics

Net of time, men are $-.11$ lower on the nconfin scale. Easy enough.

```
> sub$male <- ifelse(sub$sex==1, "male", "female")
> lmer.confinan4 <- update(lmer.confinan3, ~ . + male)
> summary(lmer.confinan4)
```

AIC	BIC	logLik	deviance	df.resid
5943.3	5985.7	-2964.6	5929.3	3156

Random effects:

Groups	Name	Variance	Std.Dev.	Corr
idnum	(Intercept)	0.25402	0.5040	
	panelwave	0.02217	0.1489	-0.71
Residual		0.26298	0.5128	

Number of obs: 3163, groups: idnum, 1324

Fixed effects:

	Estimate	Std. Error	t value
(Intercept)	2.45701	0.02934	83.74
panelwave	-0.24174	0.01242	-19.46
malemale	-0.11033	0.02784	-3.96

Correlation of Fixed Effects:

	(Intr)	panlwv
panelwave	-0.788	
malemale	-0.403	0.010

Add in a time-varying characteristics too

Net of time and gender, each move up the political conservatism scale increases one's nconfin score by 0.016 ($p < .1$)

```
> lmer.confinan5 <- update(lmer.confinan4, ~ . + polviews)
```

AIC	BIC	logLik	deviance	df.resid
5759.6	5807.9	-2871.8	5743.6	3061

Random effects:

Groups	Name	Variance	Std.Dev.	Corr
idnum	(Intercept)	0.2612	0.5110	
	panelwave	0.0240	0.1549	-0.72
Residual		0.2610	0.5109	

Number of obs: 3069, groups: idnum, 1301

Fixed effects:

	Estimate	Std. Error	t value
(Intercept)	2.395228	0.046015	52.05
panelwave	-0.242785	0.012653	-19.19
malemale	-0.103955	0.028089	-3.70
polviews	0.015831	0.008601	1.84

Correlation of Fixed Effects:

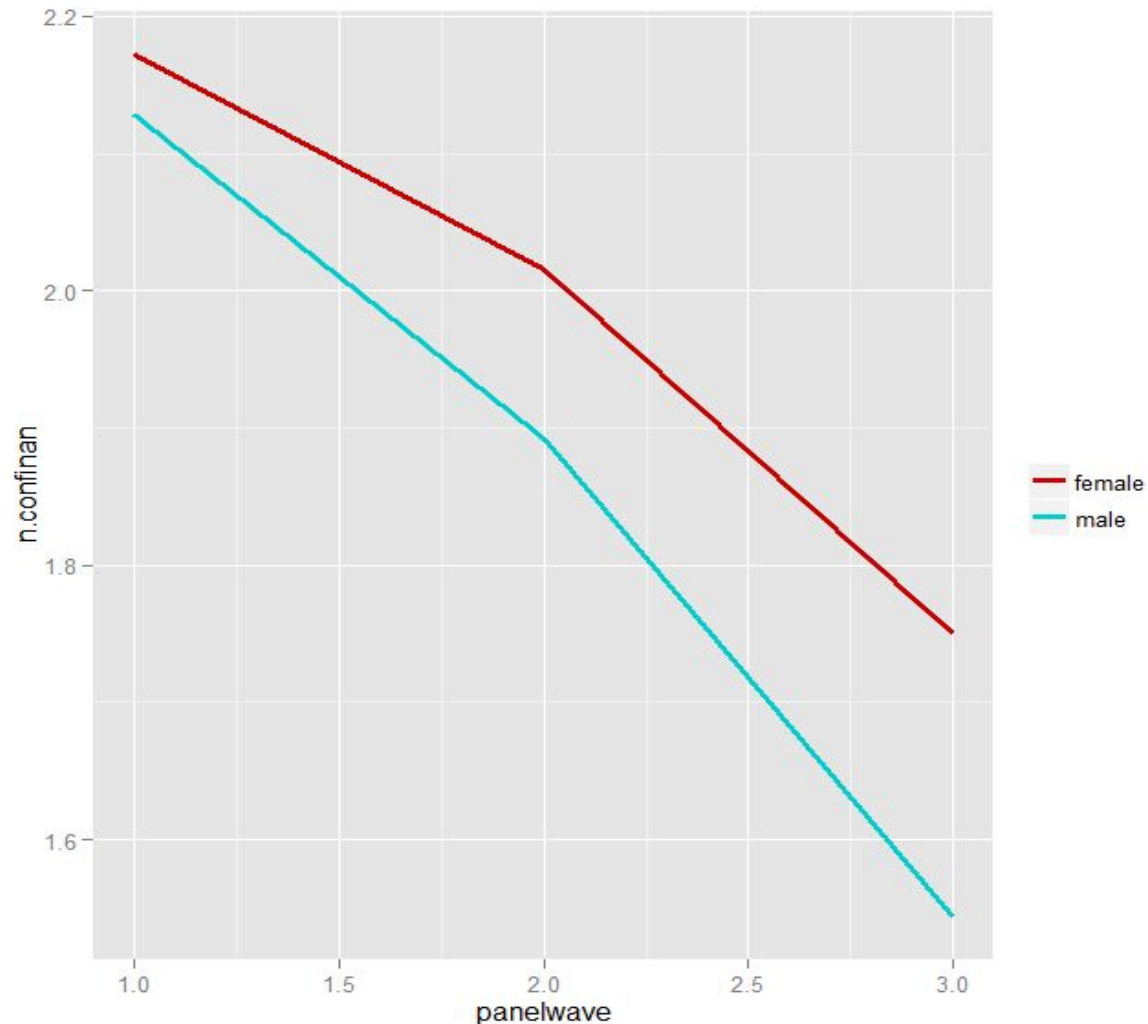
	(Intr)	panlwv	maleml
panelwave	-0.503		
malemale	-0.268	0.008	
polviews	-0.762	-0.012	0.013

How to do growth curves II

7. Do you need interactions of some Xs with time?

Changes in confidence in banks, by sex

This is from Wave 1 (2006) to Wave 3 (2010)



How did I do this?

```
colors_and_labels <- scale_color_manual(values = c("red3", "cyan3"), name = "")
g_sex <- ggplot(sub, aes(x = panelwave, y = n.confinan, color = male))
(g_sex <- (g_sex + stat_summary(fun.y=mean, geom="line", lwd = 1.25)))
g_sex + colors_and_labels
```

Add in a time-invariant characteristic

Males are losing confidence in banks faster than females
(B on the interaction = $-.081$, $p < .01$)

```
> lmer.confinan6 <- update(lmer.confinan5, ~ . + male:panelwave - polviews)
```

AIC	BIC	logLik	deviance	df.resid
5934.9	5983.3	-2959.4	5918.9	3155

Random effects:

Groups	Name	Variance	Std.Dev.	Corr
idnum	(Intercept)	0.2466	0.4966	
	panelwave	0.0200	0.1414	-0.70
Residual		0.2635	0.5133	

Number of obs: 3163, groups: idnum, 1324

Fixed effects:

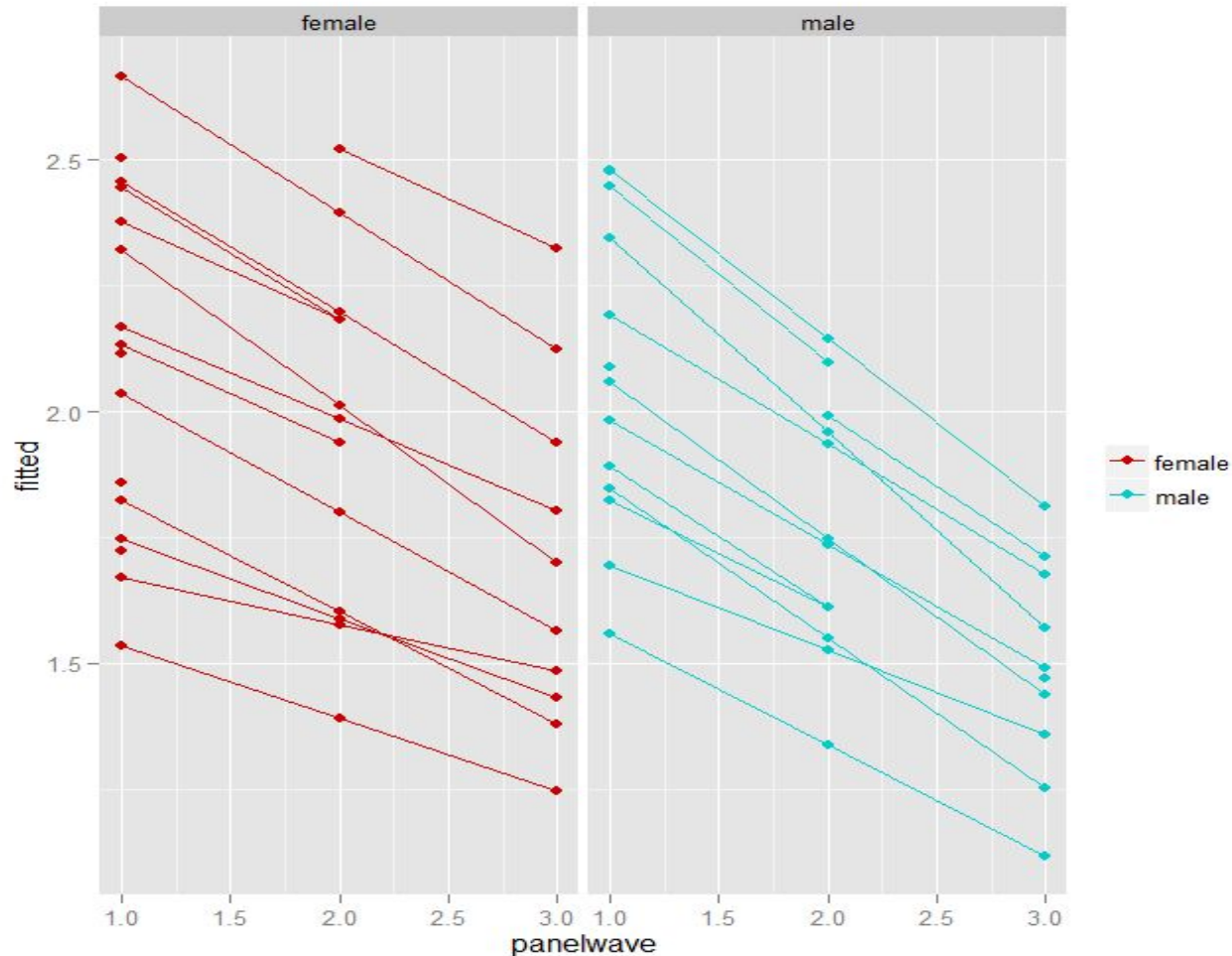
	Estimate	Std. Error	t value
(Intercept)	2.39458	0.03505	68.32
panelwave	-0.20828	0.01611	-12.93
malemale	0.04022	0.05424	0.74
panelwave:malemale	-0.08145	0.02516	-3.24

Correlation of Fixed Effects:

	(Intr)	panlwb	maleml
panelwave	-0.857		
malemale	-0.646	0.555	
panlwb:mlml	0.551	-0.642	-0.858

The predicted curves for males and females

What do you see?



How did I do this?

```
model.dat <- cbind(model.frame(lmer.confinan6), fitted = fitted(lmer.confinan6))
model.dat <- subset(model.dat, idnum < 200)
g_sex_fit <- ggplot(model.dat, aes(x = panelwave, y = fitted, group = idnum, color =
male))
(g_sex_fit <- g_sex_fit + geom_line() + geom_point() + facet_grid( . ~ male))
g_sex_fit + colors_and_labels
```

Maybe I should enter these interactions as random effects?

Why? Because perhaps the intercepts and/or slopes are more variable for one gender over the other.

Random intercepts/random slopes for men

Males are losing confidence in banks at $-.292$ per Wave, with a sd on that slope of $.105$; the constant has a mean of 2.43 and a sd of $.509$

```
> lmer.confinanM <- lmer(n.confinan ~ panelwave + (1 + panelwave | idnum), data = sub,  
subset = sex == 1)  
> summary(lmer.confinanM)
```

Random effects:

Groups	Name	Variance	Std.Dev.	Corr
idnum	(Intercept)	0.26126	0.5111	
	panelwave	0.01148	0.1072	-0.78
Residual		0.27684	0.5262	

Number of obs: 1309, groups: idnum, 568

Fixed effects:

	Estimate	Std. Error	t value
(Intercept)	2.43758	0.04240	57.49
panelwave	-0.29204	0.01914	-15.26

Correlation of Fixed Effects:

	(Intr)
panelwave	-0.856

Random intercepts/random slopes for women

Females are losing confidence in banks at $-.209$ per Wave, with a sd on that slope of $.171$; the constant has a mean of 2.397 and a sd of $.501$

```
> lmer.confinanW <- update(lmer.confinanM, subset = sex == 2)
> summary(lmer.confinanW)
```

Random effects:

Groups	Name	Variance	Std.Dev.	Corr
idnum	(Intercept)	0.25308	0.5031	
	panelwave	0.02961	0.1721	-0.70
Residual		0.24831	0.4983	

Number of obs: 1854, groups: idnum, 777

Fixed effects:

	Estimate	Std. Error	t value
(Intercept)	2.39741	0.03455	69.4
panelwave	-0.20966	0.01625	-12.9

Correlation of Fixed Effects:

	(Intr)
panelwave	-0.856

Are the intercepts and/or slopes more variable for one gender over the other?

Maybe a little bit. The standard deviation of the slope for males is only .10 but for females it is .17. This means that the growth curves for males are bunched a bit more together than for females. How much should we worry about this?

Not unlike sequence analysis too

Be creative about “types” of experiences, i.e.:

- “Early delinquents” vs. “late delinquents”
- “Young divorce” vs. “old divorce”
- Etc.