# Lab 3: ACF and CCF

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### Stationarity

Stationarity is a property of a **process**, not of a time series.

There are two types of stationarity:

▶ Strong / Strict / Strict-sense stationarity : A stochastic process  $\{X_t\}_{t\in\mathbb{R}}$  is strongly stationary iif

$$\forall K \in \mathbb{N}, h, t_1, \dots t_K \in \mathbb{R}, \qquad F_{X_{(t_1)}, \dots X_{(t_K)}} = F_{X_{(t_1+h)}, \dots X_{(t_K+h)}}.$$

This is a condition on the **distributions**.

Weak / Wide-sense / covariance stationarity: A stochastic process is weakly stationary iif:

$$\mu_X(t) = \mathrm{cst}$$
 $R_X(\tau) = \mathrm{cov}(X_t, X_{t+\tau})$ 
 $\mathbb{E}\left[|X(t)|^2\right] < +\infty.$ 

This is a condition on the **moments**.

#### Examples

# A strongly stationary white noise

$$w_t \sim \mathcal{N}(0, \sigma^2)$$
, i.i.d.

The i.i.d. assumption is more restrictive than the strong stationarity.

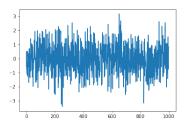


Figure 1: Gaussian noise  $\mathcal{N}(0,1)$ 

#### Examples

## A weakly stationary white noise

$$w_t = \sin(2\pi t U),$$

where  $U \sim \mathcal{U}(0,1)$ .

We can show that  $\mathbb{E}[w_t] = 0$ ,  $cov(w_t, w_{t+h}) = 0$  if  $h \neq 0$  and  $\mathbb{V}[w_t] = \frac{1}{2}$ . However it is not strongly stationary.

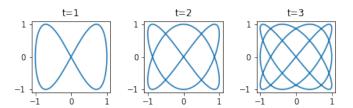


Figure 1: Graph of  $w_{t+1} = f(w_t)$  for  $u \in [0, 1]$ 

# Stationarity: Properties

#### Remarks

- Weak stationarity does not imply strong stationarity: a condition on the moments is less restrictive than a condition on the distribution.
  - Example: Keep the same first two moments, but change the generating distribution on parity considering independence.
- Strong stationarity is equivalent to weak stationarity iif the first two moments exist.
  - Example : Let  $X_t$  be iid Cauchy. X is then strongly stationary but not weakly stationary.
- For a Gaussian Process (any subset of variables is gaussian) with a stationary kernel, weak stationarity is equivalent to strong stationarity.

# Testing stationarity: (Augmented) Dickey-Fuller (ADF) test

The Dickey-Fuller test tests the null hypothesis that a unit root is present in an AR model.

With an AR(1) model we would have :

$$x_t = \lambda x_{t-1} + w_t$$

$$\Delta x_t = (\lambda - 1)x_{t-1} + w_t$$

$$= \delta x_{t-1} + w_t. \qquad (\delta := \lambda - 1)$$

The t-statistic  $\frac{\hat{\lambda}-\lambda}{\sqrt{1-\lambda^2}}$  does not follow a standard distribution when the null hypothesis holds; instead we compare the value to a Dickey-Fuller table.

#### **Definitions**

If we have X, Y stationary processes, we can define the following quantities :

$$\mu_X(t) = \mathbb{E}[X_t] \qquad \text{(mean)}$$

$$R_X(t,u) = \operatorname{cov}(X_t,X_u) \qquad \text{(covariance)}$$

$$\gamma_X(\tau) = R_X(t,t+\tau) \qquad \text{(autocovariance)}$$

$$\rho_X(t,u) = \frac{R_X(t,u)}{\sqrt{R_X(t,t)R_X(u,u)}} \text{(autocorrelation)}$$

$$R_{X,Y}(t,u) = \operatorname{cov}(X_t,Y_u) \qquad \text{(cross-covariance)}$$

$$\rho_{X,Y}(t,u) = \frac{R_X(t,u)}{\sqrt{R_X(t,t)R_X(u,u)}} \text{(cross-correlation)}$$

#### **Definitions**

If we have X, Y stationary processes, we can define the following quantities :

$$\begin{split} \mu_X(t) &= \mathbb{E}[X_t] & \text{(mean)} \\ R_X(t,u) &= \text{cov}(X_t,X_u) & \text{(covariance)} \\ \gamma_X(\tau) &= R_X(t,t+\tau) & \text{(autocovariance)} \\ \rho_X(\tau) &= \frac{\gamma_X(\tau)}{\sigma_X^2} & \text{(autocorrelation)} \\ R_{X,Y}(t,u) &= \text{cov}(X_t,Y_u) & \text{(cross-covariance)} \\ \rho_{X,Y}(t,u) &= \frac{R_X(t,u)}{\sigma_X\sigma_X} & \text{(cross-correlation)} \end{split}$$

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#### **Backshift**

The backshift operator B acting on a serie  $(x_t)_{t\in\mathbb{N}}$  is defined as :

$$Bx_t = x_{t-1}$$

with  $x_{-1} = 0$ .

This helps to identify parameter redundancy. An ARMA process can be written as :

$$\phi(B)x_t = \theta(B)w_t,$$

where  $\phi, \theta$  are two coprime polynomials (no common roots).

## Parameter redundancy: Intuition

Informally we could express the recursion using formal series. For example :

$$x_{t} = 0.8x_{t-1} - 0.16x_{t-2} + w_{t} - w_{t-1} + 0.24w_{t}$$

$$z^{t}(x_{t} - 0.8x_{t-1} + 0.16x_{t-2}) = z^{t}(w_{t} - w_{t-1} + 0.24w_{t-2})$$

$$(1 - 0.8z + 0.16z^{2}) \sum_{t=1}^{+\infty} x_{t}z^{t} = (1 - z + 0.24z^{2}) \sum_{t=1}^{+\infty} w_{t}z^{t}$$

$$(1 - 0.4z)^{2} \sum_{t=1}^{+\infty} x_{t}z^{t} = (1 - 0.4z)(1 - 0.6z) \sum_{t=1}^{+\infty} w_{t}z^{t}$$

$$(1 - 0.4z) \sum_{t=1}^{+\infty} x_{t}z^{t} = (1 - 0.6z) \sum_{t=1}^{+\infty} w_{t}z^{t}$$

By unicity of formal series, both recursion relations are the same.

### Causality

#### Causal

A process is said to be **causal** if  $\exists (\psi_j)_{j\in\mathbb{N}}$ , s.t.  $|(\psi_j)_{j\in\mathbb{N}}|_{L^1}<+\infty$  and

$$x_t = \sum_{j \in \mathbb{N}} \psi_j w_{t-j}.$$

In words, it means that  $x_t$  depends only on past values.

An ARMA process is causal iif

$$\phi(x) = 0 \implies |x| > 1.$$

## Invertibility

# Invertibility

A process is said to be **invertible** if  $\exists (\pi_j)_{j\in\mathbb{N}}$ , s.t.  $|(\pi_j)_{j\in\mathbb{N}}|_{L^1}<+\infty$  and

$$w_t = \sum_{j \in \mathbb{N}} \pi_j x_{t-j}.$$

In words, the current error is a function of past observations.

An ARMA process is invertible iif

$$\theta(x) = 0 \implies |x| > 1.$$

Invertibility is mostly defined for making models identifiable (i.e. unique solution).

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This week's exercises

#### A) implement ACF

Do your own implementation of the ACF function. Your implementation will be checked against statsmodels.tsa.stattools.acf.

Figure 2: Part I A)

#### **B) ACF of White Noise**

$$w_t \sim N(0, \sigma^2)$$

- 1. Set  $\sigma$  = 1, sample n = 500 points from the process above
- 2. Plot the white noise
- 3. Plot the sample ACF up to lag = 20.
- 4. Compare sample ACF with analytical ACF.
- 5. Compare your sample ACF with true sample ACF.
- 6. What trend/observation can you find in the ACF plot?
- 7. Change n to 50, compare the new ACF plot (n=50) to the old ACF plot (n=500). What causes the difference?

Figure 3: Part I B)

```
## 1.
n = 500
mean = 0
std = 1
lag = 20
# create white noise
w t = np.random.normal(mean, std, size=n)
# create subplots
fig, (ax1, ax2, ax3) = plt.subplots(1, 3, figsize=(15, 5))
## 2.
# plot white noise
ax1.plot(w t, label="w t")
ax1.set title("Signal")
ax1.legend()
## 3.
# calculate acf
acf val = acf(x=w t, nlags=lag)
## 4.
acf analytic = np.zeros(lag)
acf analytic[0] = std**2
plot acf(x=w t, lags=lag, title="ACF w t", ax=ax2, label="Sample ACF")
ax2.stem(acf analytic, markerfmt='r.', linefmt='r--', label="Analytic ACF")
ax2.set vlim([1.3*min(acf val), 1.1*(std**2)])
ax2.legend()
## 5.
# your implementation:
acf val impl = acf impl(x=w t, nlags=lag)
ax3.plot(acf val, 'or', label='True sample acf')
ax3.plot(acf val impl, 'xb', label='Own sample acf')
ax3.legend();
ax3.set title('vour sample ACF impl against true sample ACF'):
```

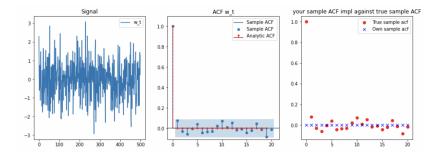


Figure 5: Part I B)

### C) ACF of Moving Average

$$v_t = \frac{1}{3}(w_t + w_{t+1} + w_{t+2})$$

- 1. Sample n+2 white noise from N(0,1)
- 2. Add code to compute the moving average  $v_t$ .
- 3. Plot both  $w_t$  and  $v_t$  and compare the two time series.
- 4. Derive the analytical ACF
- 5. Compare sample ACF up to lag 20 with the analytical ACF.
- 6. Compare your sample ACF with true sample ACF.

Figure 6: Part I C)

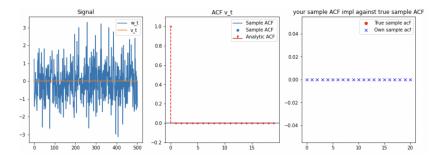


Figure 7: Part I C)

### D) ACF of signal in noise

$$v_t = 2\cos(\frac{2\pi t}{50} + 0.6\pi) + w_t$$

- 1. Sample white noise of length n from N(0, 1)
- 2. Add code to compute  $v_t$ .
- 3. Plot both  $w_t$  and  $v_t$ . Compare the two plots.
- 4. Plot the sample ACF of  $v_t$ . What's the pattern? What causes the observed pattern?

[Optional]: derive and compare to the analytical ACF (hint, use cosine trig identity)

Figure 8: Part I D)

# Part II: Cross-correlation function (CCF)

### Part II: Cross-correlation Function

## A) CCF of signal with noise

#### **Synthetic Data**

$$x_t \sim N(0, \sigma_x^2)$$

$$y_t = 2x_{t-5} + w_t$$

$$w_t \sim N(0, \sigma_x^2)$$

- In this example, we created two processes with a lag of 5.
- · Plot both samples and verify the lag.
- · Plot the empirical ACF for both samples.
- Plot the empirical CCF. What information can you conclude from the CCF plot?

Figure 9: Part II A)

# Part II: Cross-correlation function (CCF)

#### B) CCF of data

Southern Oscillation Index (SOI) v.s. Recruitment (Rec)

- . Replicate the procedure in the previous section.
- . What information can you tell from the CCF plot.
- In this example, our procedure is actually flawed. Unlike the previous example, we can not tell if the cross-correlation estimate is significantly different
  from zero by looking at the CCF. Why is that? What can we do to address this issue?

```
soi = np.array(pd.read_csv("soi.csv")["x"])
rec = np.array(pd.read_csv("rec.csv")["x"])
#TODO: This part will be graded.
# plot data
# plot data
# plot caf
# plot ccf
```

Figure 10: Part II B)

## Part III: Moving Average (MA)

### Part III

# **Moving Average**

A)

$$x_t = 0.5x_{t-1} - 0.5w_{t-1} + w_t$$
  
 $w_t \sim N(0, \sigma^2)$ 

Is  $x_t$  same as white noise  $w_t$ ? Think about ACF.

Then use code below to assess and verify your guess.

Figure 11: Part III A)

## Part III: Moving Average (MA)

B)

$$x_t = w_t + \frac{1}{5}w_{t-1}, w_t \sim N(0, 25)$$
$$y_t = v_t + 5v_{t-1}, v_t \sim N(0, 1)$$

Are  $x_t$  and  $y_t$  the same? Think about ACF.

Then use code below to assess and verify your guess.

Figure 12: Part III B)