One-sample proportion tests

HYPOTHESIS TESTING IN PYTHON



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Chapter 1 recap

• Is a claim about an unknown population proportion feasible?

- 1. Standard error of sample statistic from bootstrap distribution
- 2. Compute a standardized test statistic
- 3. Calculate a p-value
- 4. Decide which hypothesis made most sense

• Now, calculate the test statistic without using the bootstrap distribution

Standardized test statistic for proportions

p: population proportion (unknown population parameter)

 \hat{p} : sample proportion (sample statistic)

 p_0 : hypothesized population proportion

$$z = rac{\hat{p} - ext{mean}(\hat{p})}{ ext{SE}(\hat{p})} = rac{\hat{p} - p}{ ext{SE}(\hat{p})}$$

Assuming H_0 is true, $p=p_0$, so

$$z = rac{\hat{p} - p_0}{ ext{SE}(\hat{p})}$$

Simplifying the standard error calculations

$$SE_{\hat{p}} = \sqrt{rac{p_0*(1-p_0)}{n}} o$$
 Under H_0 , $SE_{\hat{p}}$ depends on hypothesized p_0 and sample size n

Assuming H_0 is true,

$$z=rac{\hat{p}-p_0}{\sqrt{rac{p_0*(1-p_0)}{n}}}$$

ullet Only uses sample information (\hat{p} and n) and the hypothesized parameter (p_0)

Why z instead of t?

$$t = rac{\left(ar{x}_{
m child} - ar{x}_{
m adult}
ight)}{\sqrt{rac{s_{
m child}^2}{n_{
m child}} + rac{s_{
m adult}^2}{n_{
m adult}}}}$$

- ullet s is calculated from $ar{x}$
 - $\circ \;\; ar{x}$ estimates the population mean
 - \circ $\;s$ estimates the population standard deviation
 - † uncertainty in our estimate of the parameter
- t-distribution fatter tails than a normal distribution
- $oldsymbol{\hat{p}}$ only appears in the numerator, so z-scores are fine

Stack Overflow age categories

 H_0 : Proportion of Stack Overflow users under thirty =0.5

 H_A : Proportion of Stack Overflow users under thirty eq 0.5

```
alpha = 0.01
```

```
stack_overflow['age_cat'].value_counts(normalize=True)
```

```
Under 30 0.535604
```

At least 30 0.464396

Name: age_cat, dtype: float64



Variables for z

```
p_hat = (stack_overflow['age_cat'] == 'Under 30').mean()
```

0.5356037151702786

```
p_0 = 0.50
```

n = len(stack_overflow)

2261

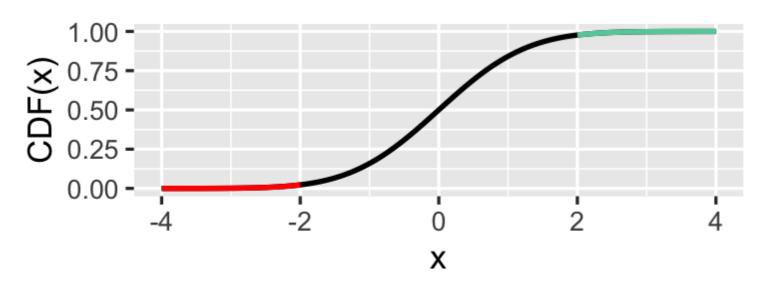
Calculating the z-score

$$z=rac{\hat{p}-p_0}{\sqrt{rac{p_0*(1-p_0)}{n}}}$$

```
import numpy as np
numerator = p_hat - p_0
denominator = np.sqrt(p_0 * (1 - p_0) / n)
z_score = numerator / denominator
```

3.385911440783663

Calculating the p-value



Left-tailed ("less than"):

```
from scipy.stats import norm
p_value = norm.cdf(z_score)
```

Right-tailed ("greater than"):

```
p_value = 1 - norm.cdf(z_score)
```

Two-tailed ("not equal"):

```
p_value = norm.cdf(-z_score) +
   1 - norm.cdf(z_score)
```

```
p_value = 2 * (1 - norm.cdf(z_score))
```

0.0007094227368100725

```
p_value <= alpha</pre>
```

True

Let's practice!

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Two-sample proportion tests

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Comparing two proportions

 H_0 : Proportion of hobbyist users is the same for those under thirty as those at least thirty

$$H_0$$
: $p_{\geq 30} - p_{< 30} = 0$

 H_A : Proportion of hobbyist users is different for those under thirty to those at least thirty

$$H_A$$
: $p_{\geq 30} - p_{<30}
eq 0$

alpha = 0.05

Calculating the z-score

• z-score equation for a *proportion test*:

$$z = rac{(\hat{p}_{\geq 30} - \hat{p}_{<30}) - 0}{ ext{SE}(\hat{p}_{>30} - \hat{p}_{<30})}$$

• Standard error equation:

$$ext{SE}(\hat{p}_{\geq 30} - \hat{p}_{< 30}) = \sqrt{rac{\hat{p} imes (1 - \hat{p})}{n_{\geq 30}}} + rac{\hat{p} imes (1 - \hat{p})}{n_{< 30}}$$

• \hat{p} ightarrow weighted mean of $\hat{p}_{\geq 30}$ and $\hat{p}_{<30}$

$$\hat{p} = rac{n_{\geq 30} imes \hat{p}_{\geq 30} + n_{<30} imes \hat{p}_{<30}}{n_{>30} + n_{<30}}$$

• Only require $\hat{p}_{>30}$, $\hat{p}_{<30}$, $n_{>30}$, $n_{<30}$ from the sample to calculate the z-score

```
p_hats = stack_overflow.groupby("age_cat")['hobbyist'].value_counts(normalize=True)
```

```
n = stack_overflow.groupby("age_cat")['hobbyist'].count()
```

```
age_cat
At least 30 1050
Under 30 1211
Name: hobbyist, dtype: int64
```



```
p_hats = stack_overflow.groupby("age_cat")['hobbyist'].value_counts(normalize=True)
```

```
p_hat_at_least_30 = p_hats[("At least 30", "Yes")]
p_hat_under_30 = p_hats[("Under 30", "Yes")]
print(p_hat_at_least_30, p_hat_under_30)
```

```
0.773333 0.843105
```



```
n = stack_overflow.groupby("age_cat")['hobbyist'].count()
```

```
age_cat
At least 30 1050
Under 30 1211
Name: hobbyist, dtype: int64
```

```
n_at_least_30 = n["At least 30"]
n_under_30 = n["Under 30"]
print(n_at_least_30, n_under_30)
```

1050 1211



-4.223718652693034

Proportion tests using proportions_ztest()

```
stack_overflow.groupby("age_cat")['hobbyist'].value_counts()
```

```
age_cat hobbyist

At least 30 Yes 812

No 238

Under 30 Yes 1021

No 190

Name: hobbyist, dtype: int64
```

```
(-4.223691463320559, 2.403330142685068e-05)
```



Let's practice!

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Chi-square test of independence

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Revisiting the proportion test

```
age_by_hobbyist = stack_overflow.groupby("age_cat")['hobbyist'].value_counts()
```

```
age_cat hobbyist

At least 30 Yes 812

No 238

Under 30 Yes 1021

No 190

Name: hobbyist, dtype: int64
```

```
(-4.223691463320559, 2.403330142685068e-05)
```



Independence of variables

Previous hypothesis test result: evidence that hobbyist and age_cat are associated

Statistical independence - proportion of successes in the response variable is the same across all categories of the explanatory variable



Test for independence of variables

```
chi2 dof
                        lambda
                test
                                                   pval
                                                           cramer
                                                                      power
             pearson 1.000000 17.839570 1.0 0.000024
                                                         0.088826
                                                                   0.988205
0
                                               0.000024
        cressie-read 0.666667 17.818114 1.0
                                                         0.088773
                                                                   0.988126
      log-likelihood 0.000000 17.802653
                                               0.000025
2
                                          1.0
                                                         0.088734
                                                                   0.988069
3
       freeman-tukey -0.500000 17.815060
                                          1.0
                                               0.000024
                                                         0.088765
                                                                   0.988115
  mod-log-likelihood -1.000000 17.848099 1.0
                                               0.000024
                                                         0.088848
                                                                   0.988236
5
              neyman -2.000000 17.976656
                                          1.0
                                               0.000022
                                                         0.089167
                                                                   0.988694
```

$$\chi^2$$
 statistic = 17.839570 = $(-4.223691463320559)^2$ = $(z$ -score) 2

Job satisfaction and age category

```
stack_overflow['age_cat'].value_counts()
```

stack_overflow['job_sat'].value_counts()

Under 30 1211
At least 30 1050
Name: age_cat, dtype: int64

```
Very satisfied 879
Slightly satisfied 680
Slightly dissatisfied 342
Neither 201
Very dissatisfied 159
Name: job_sat, dtype: int64
```

Declaring the hypotheses

 H_0 : Age categories are independent of job satisfaction levels

 H_A : Age categories are not independent of job satisfaction levels

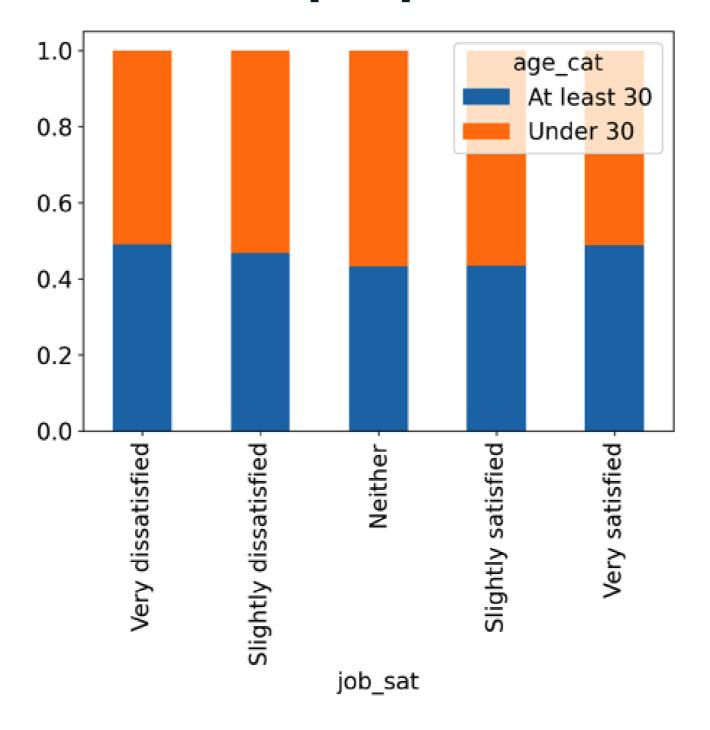
```
alpha = 0.1
```

- Test statistic denoted χ^2
- Assuming independence, how far away are the observed results from the expected values?

Exploratory visualization: proportional stacked bar plot

```
props = stack_overflow.groupby('job_sat')['age_cat'].value_counts(normalize=True)
wide_props = props.unstack()
wide_props.plot(kind="bar", stacked=True)
```

Exploratory visualization: proportional stacked bar plot



Chi-square independence test

```
import pingouin
expected, observed, stats = pingouin.chi2_independence(data=stack_overflow, x="job_sat", y="age_cat")
print(stats)
```

```
test
                       lambda
                                  chi2 dof
                                                 pval
                                                        cramer
                                                                   power
             pearson 1.000000 5.552373 4.0 0.235164 0.049555
                                                                0.437417
                                        4.0 0.235014
        cressie-read 0.666667
                               5.554106
                                                      0.049563
                                                                0.437545
                                       4.0 0.234632 0.049583
      log-likelihood 0.000000
                              5.558529
                                                               0.437871
       freeman-tukey -0.500000 5.562688 4.0 0.234274 0.049601
3
                                                                0.438178
  mod-log-likelihood -1.000000 5.567570 4.0 0.233854
                                                      0.049623
                                                               0.438538
              neyman -2.000000 5.579519 4.0 0.232828
                                                      0.049676 0.439419
5
```

Degrees of freedom:

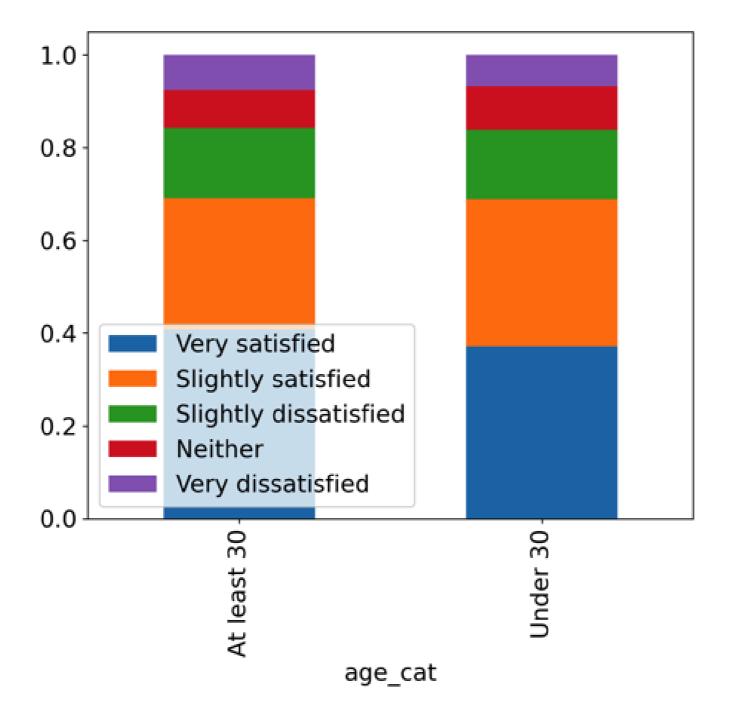
(No. of response categories -1) \times (No. of explanatory categories -1)

$$(2-1)*(5-1)=4$$

Swapping the variables?

```
props = stack_overflow.groupby('age_cat')['job_sat'].value_counts(normalize=True)
wide_props = props.unstack()
wide_props.plot(kind="bar", stacked=True)
```

Swapping the variables?





chi-square both ways

```
expected, observed, stats = pingouin.chi2_independence(data=stack_overflow, x="age_cat", y="job_sat")
print(stats[stats['test'] == 'pearson'])
```

```
test lambda chi2 dof pval cramer power
O pearson 1.0 5.552373 4.0 0.235164 0.049555 0.437417
```

Ask: Are the variables X and Y independent?

Not: Is variable X independent from variable Y?

What about direction and tails?

- Observed and expected counts squared must be non-negative
- ullet chi-square tests are almost always right-tailed 1

¹ Left-tailed chi-square tests are used in statistical forensics to detect if a fit is suspiciously good because the data was fabricated. Chi-square tests of variance can be two-tailed. These are niche uses, though.



Let's practice!

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Chi-square goodness of fit tests

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Purple links

How do you feel when you discover that you've already visited the top resource?

```
purple_link n

2 Amused 368

3 Annoyed 263

0 Hello, old friend 1225

1 Indifferent 405
```

Declaring the hypotheses

```
hypothesized = pd.DataFrame({
   'purple_link': ['Amused', 'Annoyed', 'Hello, old friend', 'Indifferent'],
   'prop': [1/6, 1/6, 1/2, 1/6]})
```

```
purple_link prop
0 Amused 0.166667
1 Annoyed 0.166667
2 Hello, old friend 0.500000
3 Indifferent 0.166667
```

 H_0 : The sample matches the hypothesized distribution

 H_A : The sample does not match the hypothesized distribution

 χ^2 measures how far observed results are from expectations in each group

```
alpha = 0.01
```

Hypothesized counts by category

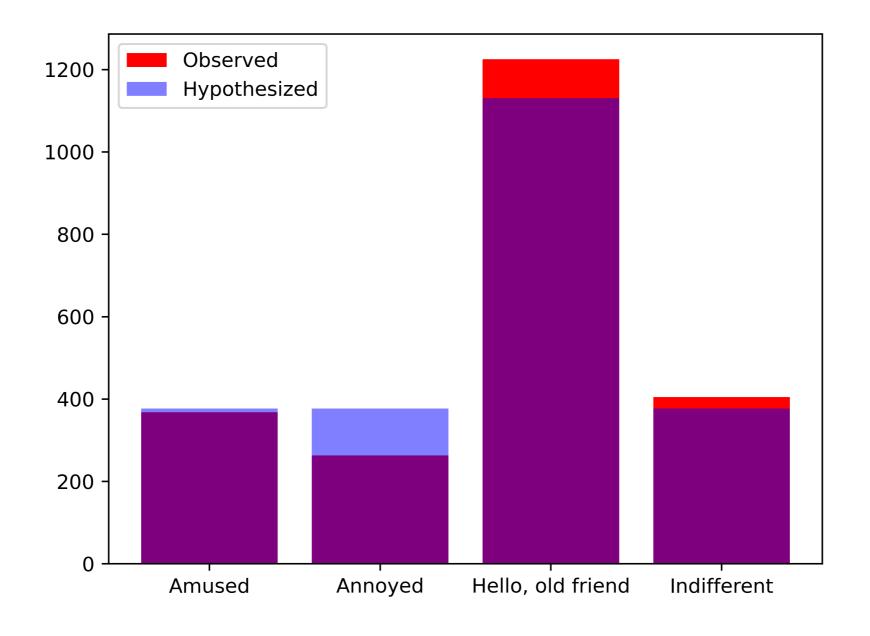
```
n_total = len(stack_overflow)
hypothesized["n"] = hypothesized["prop"] * n_total
```

```
purple_link
                         prop
                                         n
             Amused
                     0.166667
                                376.833333
            Annoyed
                     0.166667
                                376.833333
  Hello, old friend
                               1130.500000
                     0.500000
3
        Indifferent
                     0.166667
                                376.833333
```

Visualizing counts

```
import matplotlib.pyplot as plt
plt.bar(purple_link_counts['purple_link'], purple_link_counts['n'],
        color='red', label='Observed')
plt.bar(hypothesized['purple_link'], hypothesized['n'], alpha=0.5,
        color='blue', label='Hypothesized')
plt.legend()
plt.show()
```

Visualizing counts





chi-square goodness of fit test

print(hypothesized)

```
purple_link
                         prop
                                         n
0
             Amused
                     0.166667
                                376.833333
            Annoyed
                                376.833333
                     0.166667
  Hello, old friend
                     0.500000
                               1130.500000
3
        Indifferent
                     0.166667
                                376.833333
```

```
from scipy.stats import chisquare
chisquare(f_obs=purple_link_counts['n'], f_exp=hypothesized['n'])
```

Power_divergenceResult(statistic=44.59840778416629, pvalue=1.1261810719413759e-09)



Let's practice!

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