

# Cooking, health, and daily exposure to pollution spikes

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## Abstract

Many routine daily activities—such as cooking and commuting—cause large recurring pollution spikes that may impact health without significantly affecting average exposure. We study pollution spikes by combining experimental variation in cooking technology with high-frequency data on individual pollution exposure and time-use in Kenya. Improved cookstoves reduce PM2.5 spikes while cooking by  $52\mu\text{g}/\text{m}^3$  (42%) and cause a 0.24 standard deviation reduction in self-reported respiratory symptoms. However, even after more than three years of daily use, we find no clinical health improvements, possibly because we detect no impact on average exposure. Clinical health improvements may require reductions in ambient concentrations.

**JEL:** I15, O12, Q53, Q56

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# 1 Introduction

Air pollution is responsible for 7–9 million premature deaths annually (10-15% of all deaths), making it “the single biggest environmental threat to human health” (Global Burden of Disease, 2017; WHO, 2021). Specifically, research has shown that high average concentrations harm long-run clinical health and mortality.<sup>1</sup> However, many routine economic activities cause large but only brief spikes in exposure: two individuals with identical average exposure may experience very different within-day exposure (Figure 1 provides an example). What are the health impacts of repeated spikes in air pollution, often experienced on a daily basis, for years on end? Regulatory agencies and policy-makers alike are debating this question, for example in the contexts of cooking and transportation, two ubiquitous sources of daily pollution spikes (EPA, 2023; WHO, 2021).

The dearth of evidence on this topic results in part from the lack of high-frequency data on how individual-level behavior generates pollution spikes, and on how severely individuals are exposed to these spikes. This is difficult to observe because economic activity, human behavior, and demographic structures can all cause an individual’s realized exposure to differ substantially from data captured through regulatory or other stationary monitoring. Studying the topic also requires a plausibly random reduction in pollution spikes, that persists daily, for multiple years in a row.

We generate causal evidence on this subject in the context of cooking. More than 90% of pollution-related deaths occur in low- and middle-income countries (WHO, 2021), where more than four billion people still cook without modern cooking technologies (World Bank, 2020). The World Health Organization estimates that this causes 3.2 million premature deaths per year (2023).

We offer 1,000 charcoal cookstove users in Nairobi—who on average cook for around two hours per day—randomized subsidies and access to credit for an improved stove that uses 40% less charcoal. We follow up with study participants after 3.5 years of daily use: 83% of respondents who adopted an improved stove at baseline still own one at endline, and 90% of those who did not adopt one still do not own one. To characterize impacts on the distribution of individuals’ realized pollution exposure, respondents carry a backpack containing high-frequency pollution monitoring devices for 48 hours. A complementary time use survey records each respondent’s activities—and whether they were indoors—during each of those 48 hours. Informed by the pathophysiologic pathways linking pollution to cardiopulmonary disease, clinical health outcomes include blood pressure, pulse oxygen, and

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<sup>1</sup>Greenstone and Hanna (2014), Schlenker and Walker (2015), Chang et al. (2016), Isen, Rossin-Slater, and Walker (2017), Hansman, Hjort, and León (2018), Deryugina et al. (2019), Wen and Burke (2022), Clay, Lewis, and Severnini (2022), and Deryugina and Reif (2023).

self-reported respiratory and other diagnoses (Seaton et al., 1995; Pope III, 2000). We also collect detailed self-reports on health symptoms for adults and children.

Our first finding is that the improved stove reduces pollution spikes during cooking hours by 42%. For the control group, the average cooking spike (which we measure as the 99<sup>th</sup> percentile of 10-minute measurements during cooking hours) increases pollution by  $125 \mu\text{g}/\text{m}^3$  relative to daily median exposure of  $25 \mu\text{g}/\text{m}^3$ . Improved stoves reduce this by  $52 \mu\text{g}/\text{m}^3$ .<sup>2</sup>

We estimate that improved stove adoption causes a statistically significant 0.24 standard deviation reduction in an index of self-reported respiratory symptoms such as headache and cough. Several robustness checks provide evidence that experimenter demand is unlikely to explain this, at least not entirely. While subject to the common limitations of self-reported data, this result suggests the stoves may cause real health improvements.

However, we see no impacts on clinical health outcomes such as blood pressure, blood oxygen, and medical diagnoses (including pneumonia). Even after 3.5 years of significant reductions in daily pollution spikes the 95% confidence intervals exclude a 0.14 SD or greater reduction in health diagnoses and a 6 mm Hg or greater reduction in systolic blood pressure (SBP). For comparison, smoking a single cigarette increases SBP by 20 mm Hg for around 15 minutes (Cohen and Townsend, 2009). Combining our estimates with evidence from the medical literature we can reject a 12% or greater decrease in major cardiovascular events. Robustness checks indicate that the lack of clinical health impacts is not driven by measurement error or statistical noise: for example, we find a strong correlation between blood pressure and self-reported medical diagnoses.

These results can be reconciled by looking at how the intervention changed the distribution of pollution. While there was a significant decrease in spikes while cooking, study participants only cook for two hours per day on average (9% of the time). Thus, the 33% reduction in mean pollution exposure during cooking hours translates to an approximately 2.1% reduction in overall average pollution, which we cannot statistically distinguish from zero.

Taken together, the evidence is consistent with pollution spikes driving short-term respiratory symptoms and clinical health being driven by average exposure. This would imply that individuals have limited means to improve their own health by adopting improved stoves: two-thirds of respondents experience daily average pollution between  $20\text{--}49 \mu\text{g}/\text{m}^3$  (AQI 70-135), and we cannot reject the null that cookstove adoption does not change this. We present descriptive evidence supporting this interpretation: for example, self-reported symptoms correlate with pollution spikes but not with average concentrations. Participants' beliefs about the stove's potential health benefits are also consistent with this interpretation.

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<sup>2</sup> $>55 \mu\text{g}/\text{m}^3$  (150 AQI) is ‘unhealthy’;  $>150 \mu\text{g}/\text{m}^3$  (200 AQI) is ‘very unhealthy’ (EPA, 2018).

At baseline, 37% of respondents believed that adoption of the improved stove would have no impact on their health, and 34% believed it would have a small impact. These beliefs were not correlated with baseline willingness-to-pay—unlike beliefs about financial savings—suggesting health improvements were not a key feature of the stove. While clinical health benefits could emerge over a longer time horizon, they may require a reduction in average pollution exposure, for example by regulating ambient pollution levels. Improved stoves may improve health more in rural areas, or less in cities with even higher ambient air pollution.<sup>3</sup>

This research advances the literature in several ways. First, it contributes to extensive research associating a wide range of health problems with energy-intensive cookstove usage (WHO, 2021; Lancet, 2017). However, the evidence is far from conclusive. Most papers are correlational rather than causal, and randomized trials often study adoption rather than on the health impacts of improved cookstoves.<sup>4</sup> A recent *Lancet* meta-analysis identified 437 studies on the health impacts of cookstoves: only six were randomized trials (Lee et al., 2020). The article identified an “*urgent need for clinical trials evaluating cleaner fuel interventions on health outcomes to underpin evidence-based policy and decision making.*” [Table B1](#) provides an overview of the causal evidence on health impacts, which often lacks quantitative measures of pollution or health. The large RESPIRE and HAPIN trials make valuable advancements to this literature (Smith et al., 2011; Clasen et al., 2022), as do several non-experimental papers that causally evaluate health impacts using the standard econometric toolkit (Verma and Imelda, 2022). However, just like much of the existing literature, these trials took place in rural communities, where ambient pollution is on average much lower than in urban areas. In a 2018 review of the cookstove literature, Thakur et al. (2018) identified no urban papers.<sup>5</sup> There is almost no evidence evaluating cooking exposure in contexts with high ambient pollution, even though the one billion urban poor who live in slums are chronically exposed to both: 80% of urban African residents use biomass as their primary cooking energy (FAO, 2017). Our findings furthermore depart from some earlier research asserting own-household generated air pollution plays a dominant role in aggregate pollution exposure (WHO, 2014; Fisher et al., 2021).

Second, much existing research on the health impacts of air pollution evaluates changes in

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<sup>3</sup>PM2.5 averages 13  $\mu\text{g}/\text{m}^3$  in Rome, 30  $\mu\text{g}/\text{m}^3$  in Accra, and 99  $\mu\text{g}/\text{m}^3$  in Delhi (IQAir, 2019).

<sup>4</sup>See Chowdhury et al. (2019), Pattanayak et al. (2019), Mobarak et al. (2012), Miller and Mobarak (2013), Levine et al. (2018), Bensch and Peters (2019), Bensch, Grimm, and Peters (2015), Burwen and Levine (2012), and Bensch and Peters (2019).

<sup>5</sup>More recently, Alexander et al. (2018) measured spikes and duration of exposure to estimate the pollution and health impacts of improved stove adoption in an urban setting, but their sample is restricted to pregnant women, they do not separately measure ambient pollution, and they examine relatively modest variation in air pollution (a 5–13% reduction in PM2.5 spikes, and no impact on the mean). For the interested reader, [Table B1](#) provides an overview of the causal evidence.

mean daily concentrations collected by stationary monitors.<sup>6</sup> However, recent debates have heightened concerns about additional moments of the air pollution distribution, for example in relation to induction stoves (The Guardian, 2023) and wildfire smoke (Scientific American, 2024). The impacts of repeated transient spikes could differ substantially from daily averages and one-off spikes because of the non-linearity in health impacts that many papers find. However, these moments are difficult to observe because individual exposure can differ substantially from concentrations measured by stationary monitors: Pitt, Rosenzweig, and Hassan (2010) for example discuss how family structure affects cooking pollution exposure. Papers studying the impacts of sub-24 hour pollution shocks and other higher frequency moments of the pollution distribution tend to study one-off shocks rather than repeated daily spikes.<sup>7</sup> Much of this literature furthermore studies high-income countries: papers that do evaluate the health impact of ambient air pollution in low- and middle-income countries (LMICs) rarely evaluate personal exposure.<sup>8</sup> We collect individual exposure measurements of PM2.5 and CO on a minute-by-minute basis for 48 hours, which we map to hourly time use data, allows us to generate individual distributions of pollution exposure. Doing so in a high-stakes, non-laboratory setting allows us to understand how routine economic activities and human behavior drive realized pollution exposure. We combine this with experimental variation that reduces a key source of daily pollution spikes persistently for 3.5 years, and collect clinical and self-reported health outcomes to causally estimate the long-term impact of these reductions.

Finally, the dearth of research on recurrent spikes in exposure impedes the optimal design of costly environmental regulations. Most countries—as well as the WHO—only regulate PM2.5 using annual and 24-hour averages or extremes (Nazarenko, Pal, and Ariya (2021) in WHO Bulletin). While the U.S. EPA does regulate spikes of other pollutants such as ozone and carbon monoxide, the regulation of daily PM2.5 spikes is subject to ongoing debate. The U.S. Environmental Protection Agency’s recent review of the National Ambient Air Quality Standards notes that much research on the impacts of short-term spikes evaluates the immediate impacts of a single spike, rather than repeated daily exposure (EPA, 2023).

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<sup>6</sup>See footnote 1. Papers that study non-linearity in the dose-response function often study concavity in daily averages (He, Fan, and Zhou, 2016; La Nauze and Severnini, 2021; Gong et al., 2023; Miller, Molitor, and Zou, 2024). Those that study how regulations and firm actions affect pollution spikes (also known as ‘episodic pollution’) often do not measure subsequent health outcomes (Henderson, 1996; Caplan and Acharya, 2019; Cropper et al., 2014; Cutter and Neidell, 2009).

<sup>7</sup>In the economics literature, see Adhvaryu, Kala, and Nyshadham (2022), Künn, Palacios, and Pestel (2023), Ebenstein, Lavy, and Roth (2016), and Archsmith, Heyes, and Saberian (2018). Experimental papers on the health impacts of short-term pollution spikes include, for example, Kubesch et al. (2015), Soppa et al. (2014), Kocot et al. (2020), Fedak et al. (2019), and Shehab and Pope (2019).

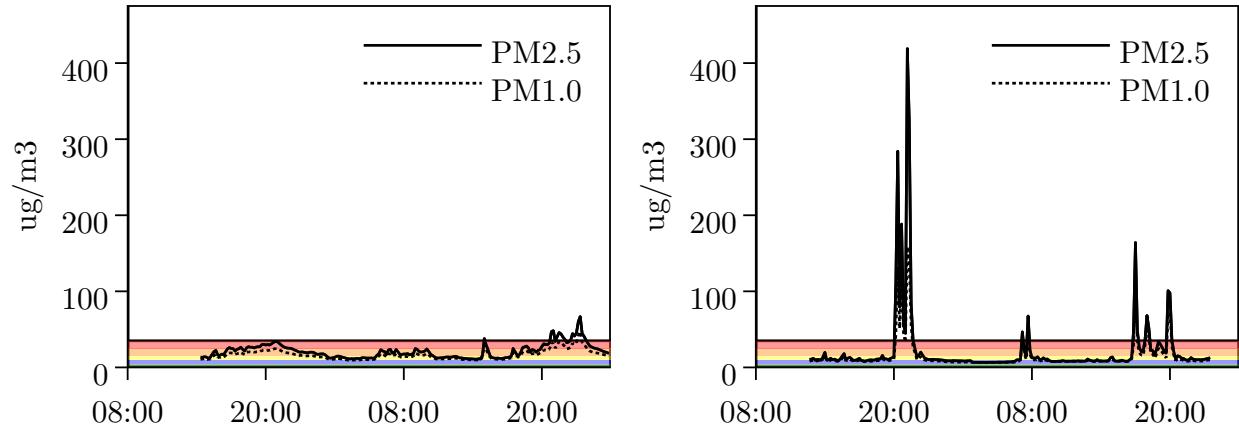
<sup>8</sup>See Adhvaryu, Kala, and Nyshadham (2022), Ebenstein et al. (2017), Gupta and Spears (2017), Greenstone and Hanna (2014), Adhvaryu et al. (2023), and Barrows, Garg, and Jha (2019).

## 2 Air pollution and health

Particulate matter (PM) is generally defined by size rather than by chemical make-up: PM<sub>2.5</sub> refers to any particulate smaller than  $2.5 \mu\text{m}$  in diameter. The chemical content of PM varies across locations, hour of day, and day of year. In Nairobi, PM consists primarily of organic carbon, dust, and sea salt (Figure A1). Charcoal burning can contribute to both organic carbon and black carbon, as can the burning of most types of fossil fuel (diesel, petrol, and coal) and forest products.

Pope et al. (2018) document average roadside PM<sub>2.5</sub> levels in Nairobi, Kenya of  $37 \mu\text{g}/\text{m}^3$  and average urban background levels of  $25 \mu\text{g}/\text{m}^3$ . This is roughly in the middle tercile of global urban pollution distributions: PM<sub>2.5</sub> concentrations average  $13 \mu\text{g}/\text{m}^3$  in Los Angeles and Rome and  $30 \mu\text{g}/\text{m}^3$  in Kampala and Accra, but  $49 \mu\text{g}/\text{m}^3$  in Jakarta and  $99 \mu\text{g}/\text{m}^3$  in Delhi (IQAir, 2019). Figure 1 displays daily PM<sub>2.5</sub> patterns collected in Nairobi, Kenya by two study participants.

Figure 1: PM<sub>2.5</sub> exposure for two study participants, both with average exposure of  $21 \mu\text{g}/\text{m}^3$



Ten minute averages (of 2-minute frequency measurements) collected by two study participants, both residing in Nairobi, Kenya, carrying Purple Air II Air Quality Sensors over 48 hours. Both have average daily PM<sub>2.5</sub> exposure of  $21 \mu\text{g}/\text{m}^3$ . The black horizontal line marks  $35 \mu\text{g}/\text{m}^3$ , the WHO's least stringent interim target. The green, blue, yellow, orange, and red bars indicate the World Health Organization's recommended air pollution targets, ranging from  $<5 \mu\text{g}/\text{m}^3$  to  $25\text{--}35 \mu\text{g}/\text{m}^3$ , its most to least stringent recommended targets, respectively (WHO, 2021). Figure A2 presents two similar examples for carbon monoxide. Section 3.3 provides more detail on data collection.

## 2.1 Transient spikes in air pollution

Diurnal patterns in air pollution spikes have been documented in the U.S. (Chavez and Li, 2020), the U.K. (Visser et al., 2015), India (Chen et al., 2020; Kumar et al., 2020), China (Bai et al., 2020), and across several African countries (Anand et al., 2024). Spikes usually occur in the morning and evening, caused for example by cooking or traffic. Most domestic governments—as well as the WHO—only provide PM2.5 standards for annual and 24-hour average concentrations.<sup>9</sup> Whether to regulate spikes is subject to ongoing policy debate, for example in the recent U.S. Environmental Protection Agency evaluation of the National Ambient Air Quality Standards (NAAQS) for Particulate Matter (EPA, 2023). A recent WHO Bulletin states, “*The current 24-hour standards mask sharp PM2.5 concentration spikes over short periods of minutes to hours. Jurisdictions with a high temporal variability of PM2.5 concentration, such as in India and China, should consider short-term averaging (such as over 20 minutes or 1 hour)*” (Nazarenko, Pal, and Ariya, 2021).

In many low- and middle-income countries (LMICs), indoor cooking is a key source of transient air pollution spikes (Johnson et al., 2022). More than four billion people still do not have access to clean cooking technologies (WB, 2020), causing millions of deaths each year (WHO, 2017; Pattanayak et al., 2019; Bailis et al., 2015). This includes many who live in cities: 80% of households living in African cities still primarily use biomass (wood or charcoal) for cooking, and three billion people are expected to live in slums in Africa and Asia by 2050 (FAO, 2017; WHO, 2021; UN, 2022). As a result, urban LMIC residents suffer disproportionately from both high average pollution concentrations and even higher transient spikes.

## 2.2 Cookstoves in Kenya

Two-thirds of Kenyan households rely on biomass as their primary cooking fuel (KNBS, 2019; WB, 2019). Around 42% of Kenyan households use a Kenyan Ceramic Jiko (KCJ, or just *jiko*) for daily cooking, with the primary alternatives being wood stoves (in rural areas) and liquefied petroleum gas (LPG) and kerosene stoves (in urban areas) (Ministry of Energy, 2019). According to the World Bank’s Kenya Country Environmental Analysis (2019), “Those who cook inside with poor ventilation have  $400\text{--}600 \mu\text{g}/\text{m}^3$  average annual concentration of PM2.5 in their household.” These levels are extremely high, even compared with the WHO’s least stringent air pollution target of  $<35 \mu\text{g}/\text{m}^3$ , let alone its most stringent target of  $<5 \mu\text{g}/\text{m}^3$  (WHO, 2021).

Figure 2 displays a *jiko* as well as the Jikokoa, an energy efficient charcoal stove produced

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<sup>9</sup>Russia limits 20-minute PM2.5 averages, making it the only country that regulates a shorter interval.

by Burn Manufacturing ('Burn'), which has sold more than four million energy efficient cookstoves since 2014. Berkouwer and Dean (2022a) provide more detail on charcoal consumption, barriers to adoption, and access to credit among potential adopters in Nairobi.

Figure 2: Traditional *jiko* ('stove') and energy efficient stove



Reproduced from Berkouwer and Dean (2022a). On the left is the traditional *jiko*. On the right is the energy efficient stove. The two stoves use the same type of charcoal and the same process for cooking food, hence the energy efficient stove requires essentially no learning to adopt. After usage, the user disposes of the ash using the tray at the bottom. The central chamber of the energy efficient stove is constructed using insulating materials.

The primary difference between the two stoves is that the Jikokoa's main charcoal combustion chamber is constructed using improved insulation material and designed for optimized fuel-air mixing (Berkouwer and Dean, 2022a). Designed and tested by laboratories in Nairobi and Berkeley, the Jikokoa is made of a metal alloy that better withstands heat, and a layer of ceramic wool insulates the chamber to cut heat loss. Parts are made to strict specifications, and components fit tightly to minimize air leakage. Adoption of the energy efficient stove does not require any behavioral adaptation or learning as the cooking processes are identical. In line with lab estimates, Berkouwer and Dean (2022a) find that adoption of the Jikokoa reduces charcoal usage (as measured through charcoal expenditures and ash generation) by 39%. Most adopters continue cooking the same types and quantities of food as before, using the same type of charcoal.

### 2.3 Existing health measurement methodologies

An extensive public health literature informed the selection of health-related outcome variables. According to the World Health Organization, the two largest health impacts from household air pollution caused by cooking are ischaemic heart disease and stroke, which we proxy for with blood pressure measurements (WHO, 2023). Kubesch et al. (2015), Chang et al. (2015), and Soppa et al. (2014) document an association between air pollution and blood pressure within 1–2 hours of high pollution exposure. The Guatemala RESPIRE trial

found impacts on blood pressure (McCracken et al., 2007), and more recently an experiment in urban Nigeria found that an improved stove can reduce blood pressure among pregnant women (Alexander et al., 2018). Recent RCTs in rural Malawi and rural Guatemala found that improved stove adoption can reduce pneumonia in adults as well as in children (Mortimer et al., 2016; Smith-Sivertsen et al., 2009). For children aged 5 and under, who are more likely than older children to spend more of their days with the primary cookstove user, frequent exposure to cooking-associated pollution may have negative health impacts, and for this reason our surveys include questions regarding adult and child health.

In settings where the technology to formally diagnose pneumonia is unavailable, the literature recommends three methodologies to diagnose pneumonia. The first is to inquire about diagnoses made by health professionals. The second is to ask about symptoms related to respiratory distress in order to make an attempted diagnoses of an acute respiratory infection (ARI), which can then be cautiously interpreted as a presumed pneumonia diagnoses. This methodology is standard for, among others, the World Health Organization, the USAID Demographic and Health Survey (DHS) program, and UNICEF.<sup>10</sup> Finally, oximetry readings have been found to be a cost-effective approach to screening for respiratory infections (Floyd et al., 2015; Van Son and Eti, 2021; National Library of Medicine, 2021).

One challenge when trying to identify the health impacts of improved stoves is that adopters may not use the improved stove consistently (e.g. Hanna, Duflo, and Greenstone, 2016; Beltramo and Levine, 2013). Berkouwer and Dean (2022a) rule this out in this paper's study context.

### 3 Study design and methodology

The study consists of three surveys conducted in 2019, a medium-term follow-up conducted in 2020, and a long-term follow-up conducted in 2022–2023. Figure 3 presents an overview.

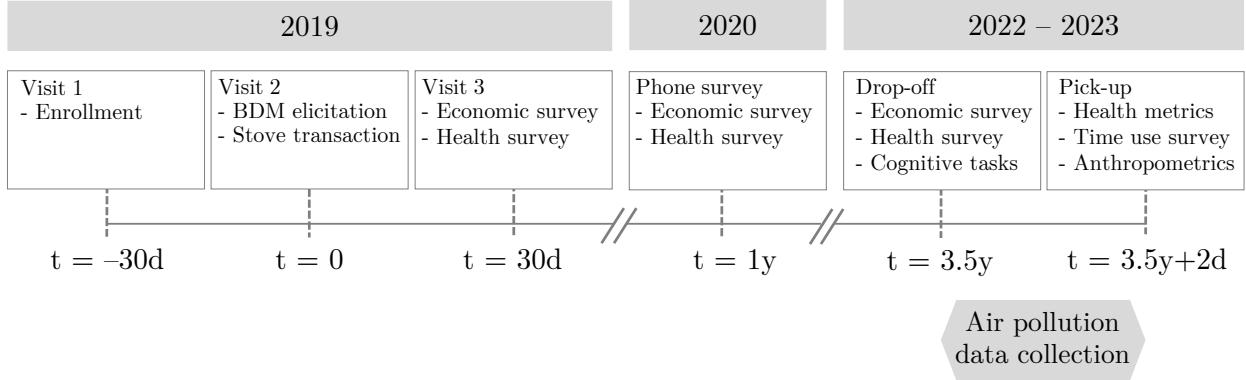
In the baseline enrolment survey conducted April-May 2019 (visit 1), enumerators enrolled respondents residing in urban settlement areas around Nairobi, Kenya who used a traditional charcoal stove as their primary daily cooking technology, who were the primary cookstove user in their household, and who spent at least \$3 per week buying charcoal. In terms of housing quality and type, and ventilation between indoor and outdoor air, respondents' housing characteristics are similar to those of many other densely populated neighborhoods in large cities in low- and middle-income countries.

To elicit baseline levels of health, enumerators asked respondents whether they had expe-

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<sup>10</sup>For example, UNICEF MICS6 (2020) identifies ARI if a child had fast, short, rapid breaths or difficulty breathing in combination with chest problems.

Figure 3: Timeline of field activities



Participants who adopted the stove did so during the main visit ( $t = 0$ ). For 89% of respondents the long-term endline was conducted 3.4–3.7 years after the main visit. Child anthropometrics were collected on either drop-off or pickup depending on presence. Due to health restrictions, the 2020 survey was conducted by phone.

rienced a persistent cough or breathlessness in the past week. If they had any children under 16 who lived with them, we asked the same about the child(ren). Enumerators then elicited beliefs about the potential health impacts of an improved stove using methodologies from the health literature (Usmani, Steele, and Jeuland, 2017; Hooper et al., 2018). Specifically, in an unprompted manner they asked respondents what they perceived to be the main benefits of the improved stove—62% stated ‘reduced smoke’ (95% said ‘saving money’). They then asked several Likert-scale questions about the extent to which the respondent thought usage of a traditional stove has had negative impacts on their health, and how much adoption of the improved stove might improve their health.

During the main visit (visit 2)—which took place approximately one month after each respondent’s baseline visit and was completed by 955 respondents—each respondent was given an opportunity to buy the stove at a subsidized price (Section 3.1 discusses the subsidy randomization in more detail). Of the 955 respondents who completed the main visit, 570 (60%) adopted the improved stove.

In June–July 2019, approximately one month after the main visit, enumerators conducted a short-term follow-up survey (visit 3). In 2020, approximately 16 months after the main visit, enumerators conducted a medium-term follow-up survey.<sup>11</sup> Both follow-up surveys included socio-economic questions as well as the same health symptoms questions asked during the baseline surveys.

In 2022–2023 enumerators conducted a long-term endline survey round, which consisted of two surveys, the second approximately 48 hours after the first. Enumerators were instructed to only conduct the long-term endline survey if they were able to find the same

<sup>11</sup>Due to COVID-19, all surveys conducted in 2020 were conducted over the phone.

individual who had completed. The surveys were designed to take rigorous quantitative measurements of air pollution exposure and physical health. An accompanying socioeconomic survey included questions on charcoal expenditures, cooking technology ownership and usage, maintenance, food cooked, home heating, in-network Jikokoa purchases, savings, income, and work activities. Enumerators were able to reach 775 and successfully surveyed 702 (75%) of the 942 respondents they attempted to reach. Recent demolitions of informal settlements in Nairobi contributed to imperfect follow-up (The Star, 2023), but attrition is balanced by treatment assignment, take-up, and baseline health (Section 4.7 provides more analysis on attrition). Of the 702 respondents, 639 still lived in or near the original study areas, 53 had moved to rural areas, and 10 had moved elsewhere in Nairobi. 95% of respondents were surveyed between 3.4–3.7 years after the main visit. Table 1 presents summary statistics collected during the endline survey.

### 3.1 Causal identification

The Jikokoa cost \$40 in stores at the time. Each respondent was randomly assigned a subsidy between \$10-39, stratified on baseline charcoal usage (Figure A3 shows the distribution of prices). The subsidy treatment was cross-randomized with a random credit treatment allowing recipients to pay for the stove in installments over a 3-month period, as well as an attention treatment designed to increase the salience of long-term charcoal savings.

During the main visit, enumerators used a Becker, Degroot, and Marschak (1964) mechanism (BDM) to elicit WTP for the improved stove. After first identifying their maximum WTP through a binary search, respondents then opened an envelope containing their randomly assigned price. Respondents whose WTP was at least as high as their assigned price (the market price of \$40 minus the randomly assigned subsidy) then adopted the stove.<sup>12</sup>

The credit treatment doubled WTP, while the attention treatment had no effect on WTP (Figure A4). Among those in both the high subsidy and the credit treatment group 93% adopted the improved stove, whereas among those in both the low subsidy and the credit control group only 8% did.

To estimate the causal effect of improved stove adoption on long-term outcomes, we use the randomly assigned subsidy, the credit treatment assignment, and their interaction as instruments for adoption. We report weak instrument F-statistics where relevant, but the first stage is generally strong.

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<sup>12</sup>98.6% of respondents for whom this was the case actually adopted the stove.

Table 1: Summary statistics from respondent surveys

	N	Mean	SD	25 <sup>th</sup>	50 <sup>th</sup>	75 <sup>th</sup>
Female respondent	702	0.96				
Completed primary education	702	0.70				
Completed secondary education	702	0.26				
Age	702	41.46	11.8	33.0	40.0	48.0
Number of children under 5 in home	702	0.50	0.7	0.0	0.0	1.0
Daily earnings (USD)	563	2.77	5.8	1.0	1.7	3.1
Daily charcoal expenditure (USD)	702	0.48	0.6	0.2	0.3	0.6
Minutes spent cooking per day	702	127.54	59.5	90.0	120.0	150.0
... of which indoor	702	111.80	61.3	70.0	109.0	150.0
Owes Jikokoa	702	0.52				
Owes traditional wood or charcoal jiko	702	0.57				
Owes LPG stove	702	0.59				
Owes electric stove	702	0.01				
Mostly uses modern stove	702	0.53				
Blood oxygen	696	96.74	2.4	96.0	97.0	98.0
Average systolic blood pressure	696	123.46	22.0	108.3	118.5	131.7
Average diastolic blood pressure	696	81.75	12.9	73.0	79.3	89.0
Number of health symptoms	702	2.47	2.6	0.0	2.0	4.0
<i>In the past month, fraction experiencing...</i>						
Fever	702	0.22				
Headache	702	0.48				
Persistent cough	702	0.23				
Runny nose	702	0.22				
Sore throat	702	0.15				
Always feeling tired	702	0.28				

Standard deviation and 25<sup>th</sup>, 50<sup>th</sup>, 75<sup>th</sup> reported for all non-binary variables. Blood pressure is averaged over three readings taken consecutively.

### 3.2 Time use and behavior

To match high-frequency pollution data to specific activities such as cooking or commuting, the second endline survey included a time use module asking which activity or activities the respondent was engaged in for each hour between the two surveys, whether they were primarily indoors or outdoors during each hour, and—if they were cooking—which stove(s) they were using. Most respondents cook primarily between the morning hours of 5–8am and the evening hours of 6–9pm. There are modest differences in the types of technologies used during different types of day, with LPG used slightly more in the mornings and charcoal stoves used more in the evenings (Figure A5). Anecdotally, this appears to be due to a preference for a fast-lighting stove (such as the LPG stove) in the morning, for a small meal

or hot beverage, and a longer-cooking stove when preparing larger meals.

Respondents are indoors on average 89% of hours spent cooking, but there is some heterogeneity correlated with stove usage. For the 278 respondents who report using an LPG or electric stove at least once in the time use survey, on average only 5% of hours spent cooking with such a stove are spent outdoors. Conversely, for the nearly 500 respondents who report using a wood or charcoal stove at least once in the time use survey, more than 20% of hours spent cooking with such a stove are spent outdoors. Respondents may be choosing to cook indoors when using a relatively cleaner stove, which would limit the reduction in pollution exposure caused by the improved stove, as emissions are more likely to build up indoors. All of our results on the impact of improved cookstove adoption should be interpreted as factoring in any accompanying behavior changes such as location choice, or opening doors or windows in order to increase ventilation, which have been shown to matter (Lenz et al., 2023). That said, we do not detect an impact of improved cookstove adoption on the propensity to cook or to cook indoors ([Table B2](#)).

### 3.3 Measuring air pollution exposure concentrations

We use two devices to measure air pollution. A Purple Air II Air Quality Sensor (PA-II) takes one measurement of Particulate Matter (PM) every two minutes,<sup>13</sup> and a Lascar EL-USB-CO Data Logger takes one measurement of Carbon Monoxide (CO) per minute ([Figure A6](#) depicts the two devices).<sup>14</sup> A test of co-located readings shows that devices are strongly correlated and that there is a small and generally stable gap between some devices ([Figure A7](#)). For this reason we include device fixed effects in all regressions.<sup>15</sup>

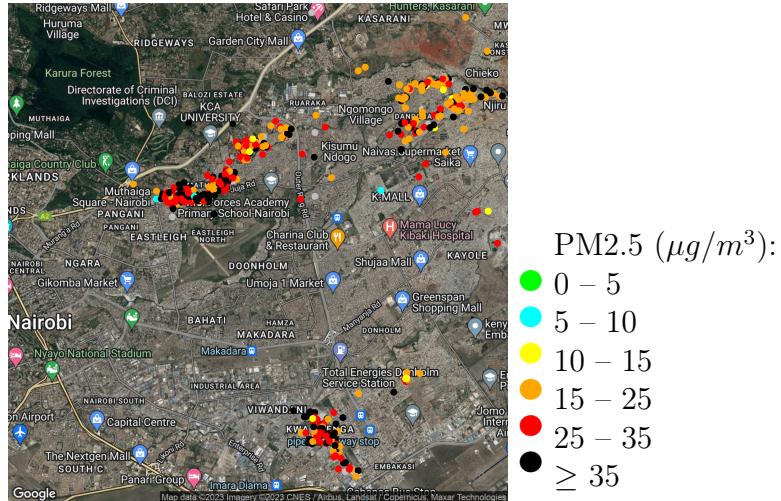
Air pollution exposure varies considerably not only by cookstove usage but by user behavior (Pitt, Rosenzweig, and Hassan (2010), for example, discuss how household structures affect exposure). Following best practices from the public health and air pollution monitoring literature (Johnson et al., 2022; Gordon et al., 2014; Gould et al., 2022; Chillrud et al., 2021; Burrowes et al., 2020), and using procedures developed by the Berkeley Air Monitoring Group (Johnson et al., 2021; Johnson et al., 2024), we therefore collect personal exposure as experienced by respondents rather than conducting stationary monitoring of kitchen concentrations. During the first follow-up survey we provided each respondent with a small mesh

<sup>13</sup>We average the PA-II *a* and *b* readings, and top-code data at  $419 \mu\text{g}/\text{m}^3$  above which the device saturates. We apply the PA-II calibration methodology from Ward et al. (2021) and Giordano et al. (2021) to correct for humidity and local air composition. Building on Tryner et al. (2020), if the difference between the *a* and *b* readings is at least 25% and at least  $15 \mu\text{g}/\text{m}^3$  the reading is removed from the sample (1.7% of readings).

<sup>14</sup>Each CO device has an independent calibration factor. Devices were re-calibrated every two months, between survey breaks. We include device FE in all regressions.

<sup>15</sup>Interacting device fixed effects with a linear time trends could account for heterogeneous trends across devices. This increases standard errors but does not qualitatively change the results.

Figure 4: Average air pollution (PM 2.5) for participants by their home locations



Distribution of respondents across Nairobi. Colors correspond to average particulate matter (PM 2.5) exposure. Respondents for whom pollution was not recorded are shown in gray. The WHO air quality guideline is  $5 \mu\text{g}/\text{m}^3$  (WHO, 2021). WHO interim targets 1 through 4 correspond to 10, 15, 25, and  $35 \mu\text{g}/\text{m}^3$ . 63 survey respondents reside outside the depicted area.

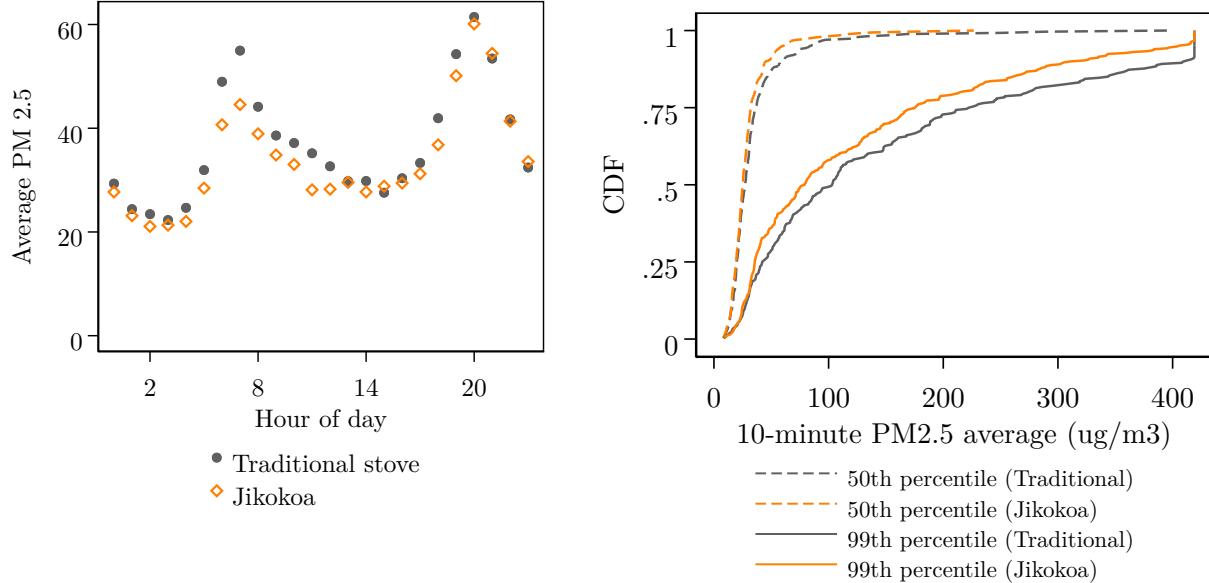
backpack containing the two devices (Panels C and D of Figure A6). Respondents were asked to wear this backpack continuously whenever feasible, or to keep it within one meter, at waist level, when wearing it was infeasible. During the second follow-up survey 48 hours later the enumerators picked up the devices, downloaded the data, recharged the 48-hour battery pack, and placed them in a new backpack to be deployed with a different respondent.<sup>16</sup> Collecting pollution exposure over a 48-hour period captures pollution generated by the respondent as well as ambient pollution generated by industrial facilities, traffic, or other sources in urban Nairobi. Figure 4 displays respondents' average air pollution exposure over the 48 hours, by their residential location. For most respondents, average PM2.5 is well above the WHO's air quality limits.

Enumerators did not observe significant hesitancy from participants about wearing the backpacks. The backpacks were designed to not be cumbersome to participants, weighing only 2.2 lbs, and previous validation research in the same context found high levels of compliance (Johnson et al., 2021). Based on the advice of the Berkeley Air Team, we also allowed participants to take the backpack off while they were stationary as long as it remained within 1 meter of their location. Of the 43 respondents for whom air pollution data are missing, 36 had stated in advance that they only had time to complete one survey and as a result were never asked to receive the devices.<sup>17</sup> We did not quantitatively monitor backpack wearing, as this would have required installing GPS trackers on the backpacks which we felt could be

<sup>16</sup>85% of respondents held the device between 45–50 hours.

<sup>17</sup>The remaining nine all agreed to receive the backpack, but the data were unusable for technical reasons.

Figure 5: Particulate matter (PM2.5, in  $\mu\text{g}/\text{m}^3$ ) pollution by improved stove ownership  
 A) Hourly averages over the day      B) Distributions of 50th and 99th percentiles of  
 10-minute averages (by respondent)



Panel A shows average PM2.5 air pollution, by hour and endline improved stove ownership, as collected by respondents wearing backpacks for on average 48 hours. Figure A8 presents the same for PM1.0 and CO.

perceived as violating participants' privacy and increase attrition, but enumerators reported generally high backpack wearing.<sup>18</sup>

Panel A of Figure 5 shows how average pollution varies over the hours of the day, by whether or not the respondent owned a Jikokoa as of the 2022–2023 survey. The levels and diurnal patterns of PM2.5 and PM1.0 follow the air pollution patterns documented by Pope et al. (2018) in urban Kenya. Average PM2.5 concentrations are highest in hours when participants self-report cooking (particularly with biomass stoves), and lowest in the hours when sleeping (Table B3). Emissions in the 60 to 30 minutes prior to the emissions peak is approximately similar across stoves, but the return to baseline levels takes significantly longer for dirtier biomass stoves than it does for cleaner stoves (Figure A9). We do not observe any meaningful seasonal heterogeneity in air pollution over our sample period.

To better understand the distribution of pollution we compute average exposure during each 10-minute window for each respondent in our data. Panel B of Figure 5 shows the cumulative distributions of respondents' 50th and 99th percentile 10-minute averages. Me-

<sup>18</sup>Enumerators were attentive to this: for example, they raised concerns about a lack of continuous backpack wearing as respondents would take off the backpack for example while sleeping it (placing it next to their beds) or while working statically (placing it on a table), which we agreed was acceptable as long as the backpack was within one meter of the respondent. That enumerators were attentive enough to identify this issue suggests they likely would have noticed any more severe widespread non-compliance.

dian 10-minute average is below  $50 \mu\text{g}/\text{m}^3$  for 89% of respondents, but the 99th percentile exceeds  $100 \mu\text{g}/\text{m}^3$  for half of respondents, and exceeds  $200 \mu\text{g}/\text{m}^3$  for 23% of respondents. This provides the first evidence that within-day variation is statistically and economically significant.

### 3.4 Measuring physical health

We present a simple conceptual framework to elucidate how pollution exposure may affect various health measurements. Define  $S_t$  to be an individual's health stock that drives long term outcomes such as chronic health and longevity (Grossman, 1972). Health stock is primarily an auto-regressive process affected by general aging as well as environmental shocks  $E_t$  (such as pollution):

$$S_{t+1} = f(S_t, E_t)$$

Due to the complexity of human biology, it is difficult to know the correct functional form for  $f$ . While evidence is clear that acute exposure to pollution causes contemporaneous negative health effects, it is not clear from the literature how quickly or if the body completely fully recovers (Behndig et al., 2006; Tong et al., 2014; Main et al., 2020; Swiston et al., 2008; Greven et al., 2012; Salvi et al., 1999; Deary and Griffiths, 2021; Ghio, Kim, and Devlin, 2000; Xu et al., 2013; Singh, Nagar, and Arora, 2023). Given the ambiguity in the medical and biological literatures, we remain agnostic about this functional form.

Health stock and its evolution are unobservable to the researcher. Instead, we record health measurements  $\mu_t$  at time  $t$  which are a function of the individual's health stock at time  $t$ , their environmental exposure at time  $t$ , and error  $\epsilon_t$ .

$$\mu_t = g(S_t, E_t, \epsilon_t)$$

The functional form of  $g$  varies by outcome. For example, hospitalization is only detected for extremes of  $S_t$  or  $E_t$ , whereas more continuous measures such as blood pressure could respond even to modest changes in stress and diet. Different measurements can also be more or less error prone: for example, asking individuals to recall their cough frequency over the past six months is more noisy than over the last week.

Different measurements can be more or less strongly affected by health stock relative to environmental exposure. For example, someone with high health stock could still experience severe symptoms after a transient spike in pollution exposure. We think of clinical health outcomes—such as blood pressure, blood oxygen, or child height and weight—as those  $\mu_t$  that tend to be affected more by the health stock than by the individual's environmental

shock in that period, such that:

$$\frac{\partial \mu_t}{\partial S_t} > \frac{\partial \mu_t}{\partial E_t}$$

We conceptualize acute health outcomes—such as cough, runny nose, sore eyes—as the converse: measurements of the individual’s health which tend to be more strongly affected by that day’s environmental shock than by health stock. We expect transient air pollution spikes to have lower impacts on measurements which are heavily stock dependent and those that have a large error component.

The sections below describe these two sets of health outcomes in turn. To control for diurnal patterns in some health outcomes, health regressions control for the hour of day during which respondents were surveyed. The survey also asks about perceptions of health impacts, and frequency and financial costs of hospital visits.

**Clinical health outcomes** The medical literature documents pathophysiologic mechanisms linking particulate matter with cardiac and pulmonary disease (Seaton et al., 1995; Pope III, 2000). Enumerators record systolic and diastolic blood pressures using a sphygmomanometer, following procedures set by the Centers for Disease Control and Prevention NHANES (2019).<sup>19</sup> The analysis uses direct measures of systolic and diastolic blood pressure as well as indicators for having hypotension (low blood pressure, defined as <90/60 mmHg), stage 1 hypertension (130-139/80-89 mmHg), and stage 2 hypertension ( $\geq 140/90$  mmHg), as defined by the American Heart Association and the American College of Cardiology (Goetsch et al., 2021). Enumerators use pulse oximeters (blood oxygen saturation monitors) to record haemoglobin oxygen saturation.<sup>20</sup>

The enumerators also ask 10 yes/no questions about whether a medical professional had diagnosed the respondent with various medical diagnoses (including pneumonia, asthma, or other lung disease), of which we only use diagnoses that were made in the past three years (since the original experiment) in the analysis. 17% of respondents report having been diagnosed with pneumonia by a doctor at least once in their lives, including 12% who report having been diagnosed in the past three years. Table 1 presents additional summary statistics on health outcomes. To correct for multiple hypothesis testing, we combine these measures into a standardized health diagnoses index.

Enumerators also measure height, weight, and arm circumference for children (as indica-

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<sup>19</sup>Respondents are asked to sit still, upright, and not engage in affecting behaviors (cooking, smoking, etc.) in the 30 minutes prior to the blood pressure readings. In line with guidelines, blood pressure is recorded three times and the analysis uses the average of the three readings.

<sup>20</sup>While we considered collecting spirometry or peak expiratory flow data, medical consultants in Kenya and the U.S. suggested that these run the risk of generating noisy and unusable data. We therefore chose to focus on improving the quality of the personal exposure, blood pressure, and blood oxygen measurements.

tors for physical child development) and for adults (as controls).

**Acute health outcomes** Following the methodology from the public health literature (see for example Tielsch et al., 2016; Smith-Sivertsen et al., 2009; Checkley et al., 2021 and others), the survey asks a large set of self-reported health questions. This includes 29 yes/no questions asking if the respondent experienced specific symptoms in the past 4 weeks (including fever, persistent cough, stomach pain, or rapid weight loss). To correct for multiple hypothesis testing, we combine these measures into two standardized health symptoms indices, one for respiratory symptoms and one for non-respiratory symptoms.

The respondent is asked similar questions about the health of any children under 10 who live in the home, including questions about overall health and basic health symptoms, which we then combine into a standardized child physical health index. The survey also asks about a set of health symptoms that permit a presumed pneumonia diagnoses, school attendance, and medical diagnoses.

### 3.5 Measuring cognition

To assess basic adult and child cognitive functions, we use three instruments. First, we use the Reverse Corsi Block task to measure working memory (Brunetti, Del Gatto, and Delogu, 2014). Second, we use Hearts and Flowers to measure response inhibition (Davidson et al., 2006). Third, we use the d2 task for sustained attention (Brickenkamp and Zillmer, 1998; Bates and Lemay Jr., 2004). Appendix D provides detail on these assessments. The analysis uses a standardized adult cognitive ability index that combines these three outcomes.

## 4 Causal impacts

To estimate the causal effect of improved stove adoption on pollution spikes and health, we use an instrumental variables (IV) approach where we use the randomly assigned stove price, the randomly assigned credit treatment status, and their interaction as instruments for stove ownership. These treatments had a statistically and economically large effect on stove adoption in Berkouwer and Dean (2022a).<sup>21</sup> Since both are randomly assigned, this regression identifies the causal effect of stove adoption on the outcomes of interest. Regressions include socioeconomic controls and fixed effects as indicated.<sup>22</sup>

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<sup>21</sup>We omit a third random treatment, attention to energy savings, as it had no impact on adoption.

<sup>22</sup>Socioeconomic controls used in each regression are the respondent's attention treatment status (a treatment designed to increase attention to energy savings), age, gender, savings in 2019, income in 2019, number of residents in the household in 2019, number of children in the household in 2019, prevalence of a cough or breathlessness at night in 2019, hours of work/homework missed due to poor health, education level

The ‘stove adoption’ dummy could represent either initial adoption in 2019, or ownership status as of the 2022–2023 endline survey. Using initial adoption represents the longer-term effects of adoption, factoring in potential breakage or other subsequent changes in stove ownership, but underestimates contemporaneous effects as some treated individuals are no longer benefiting from the treatment. Long-term adoption status better estimates contemporaneous differences, but could result in an overestimated IV coefficient if changes experienced by respondents who initially adopted the stove but no longer own one at endline are attributed to the (smaller) treatment group. We present both estimates where relevant but use ownership as of the long-term follow-up in most regressions.

## 4.1 Impacts of subsidies and credit on stove ownership and usage

Panel A of [Table 2](#) shows the causal impact of 2019 Jikokoa adoption on long-term ownership of various stove types. 90% of respondents who did not adopt a Jikokoa during the main visit still do not own one during the long term endline, and 83% of respondents who adopted a Jikokoa initially still own one three years later. This persistence generates a strong first stage, with weak IV F-statistics between 20 and 50 depending on the specification ([Table B4](#)).

The median household owns two stove types, indicating some degree of fuel stacking. Liquefied petroleum gas (LPG) usage has risen sharply in recent years, particularly in Nairobi where 60% of respondents report owning an LPG stove (compared with 40% of those who live elsewhere), potentially as a result of a government LPG subsidy program ([IEA, 2022](#)). This paper’s estimates of pollution and health impacts should be interpreted as the aggregate causal effect of improved cookstove adoption, allowing for any continued use of existing stove (rather than an estimate of a strict switch from an existing stove to an improved stove).

Jikokoa adoption does not appear to meaningfully affect adoption of other modern cooking technologies such as LPG, bio-ethanol, or electric stove ownership, though we cannot rule out modest increases. We thus find limited evidence of the ‘energy ladder’ mechanism where improved stove adoption can act as a stepping stone ([Hanna and Oliva, 2015](#)), nor of the converse, that adoption of an intermediary technology can slow adoption of an even more improved technology ([Armitage, 2022](#); [Hornbeck et al., 2024](#)).

Improved stove adoption does not appear to impact time spent cooking per day, nor the propensity to cook indoors ([Table B2](#)). It slightly increases the propensity to cook Githeri (a common maize and beans dish) but has no other impacts on foods cooked ([Table B6](#)).

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completed in 2019, charcoal expenditures in 2019, level of risk aversion in 2019, status of credit constraint in 2019, living situation as rural or urban, age as decade binary variables (designed to capture non-linear impacts of age), as well as field officer fixed effects. Panel data fixed effects include week FE, device FE, and the interaction of and hour-of-day by day-of-week by neighborhood FE.

Table 2: IV estimates of the causal impact of cookstove adoption on socio-economic outcomes

	Control Mean (1)	Treatment Effect (2022 Ownership) (2)	Treatment Effect (2019 Ownership) (3)	N
<i>Panel A</i>				
Owns other wood or charcoal stove	0.88 [0.33]		-0.54*** (0.05)	702
Owns Jikokoa	0.10 [0.31]		0.74*** (0.04)	702
Owns LPG stove	0.57 [0.50]		0.05 (0.06)	702
Owns bio-ethanol stove	0.15 [0.36]		0.01 (0.04)	702
Owns electric stove	0.00 [0.06]		0.02* (0.01)	702
<i>Panel B</i>				
Charcoal expenditures past 7 days (USD)	3.65 [2.93]	-1.50*** (0.47)	-1.12*** (0.35)	702
Charcoal expenditures past 7 days (urban)	3.79 [2.94]	-1.65*** (0.52)	-1.20*** (0.37)	649
Charcoal expenditures past 7 days (rural)	1.82 [2.09]	1.22 (1.00)	1.16 (0.81)	53
Earnings past 2 weeks (USD)	32.20 [35.31]	4.73 (7.83)	3.45 (5.38)	563
Total savings (USD)	57.70 [94.87]	-8.63 (19.88)	-7.07 (14.67)	701
Has formal bank account (=1)	0.12 [0.33]	0.11 (0.07)	0.08 (0.05)	702
Minutes cooking per day	133.79 [57.29]	3.49 (8.32)	2.60 (6.15)	702
People in network who adopted Jikokoa	0.75 [2.03]	1.13*** (0.40)	0.84*** (0.29)	702

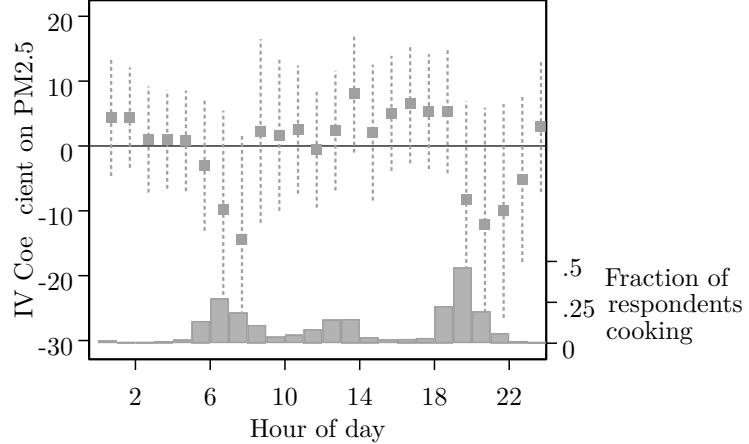
Panel A presents the causal impact of 2019 Jikokoa adoption on 2022–2023 cookstove ownership. Panel B presents the causal impact of 2022 Jikokoa ownership and 2019 Jikokoa adoption (Columns 1 and 2) on outcomes from the 2022–2023 endline surveys. Each row is an IV regression that uses the randomly assigned price, credit treatment status, and their interaction as instruments for the endogenous variables. Regressions include socioeconomic controls. Table B5 presents additional socio-economic outcomes.

## 4.2 Impacts of stove ownership on air pollution

We use the same IV approach to estimate the causal impact of stove adoption on air pollution for each hour of the day. Figure 6 shows that the impact varies significantly across hours in the day. Improved stove ownership reduces air pollution between 5–8am and 7–10pm—when

respondents report to be cooking—but not during other hours of the day.<sup>23</sup>

Figure 6: Impact of Jikokoa ownership on average hourly PM2.5 (in  $\mu\text{g}/\text{m}^3$ )



Coefficients from a single instrumental variables regression that uses subsidy, credit treatment status, their interaction, and their interaction with hour of day dummies as instruments to estimate the causal impact of Jikokoa ownership on PM2.5 exposure (with socioeconomic controls and panel fixed effects). The gray bars report the fraction of respondents who report cooking during any given hour in the time use survey. Joint F-tests that the coefficients on 5am, 6am, and 7am all equal 0 or that the coefficients on 7pm, 8pm, and 9pm all equal 0 cannot be rejected. [Table 3](#) pools the data to improve statistical power. [Figure A10](#) presents the OLS version.

To improve statistical power, [Table 3](#) aggregates PM2.5 exposure data for each individual. We use the same IV approach to estimate the causal impact of stove adoption on four key moments of PM2.5 pollution exposure: median exposure (Column 1), mean exposure (Column 2), the maximum hourly average (Column 3), and the 99th percentile of 10-minute averages (Column 4). Panel A includes all hours during which the respondent was wearing the device, while Panel B limits the data to the hours during which the respondent self-reported cooking in the time use survey (for robustness, [Table B7](#) presents results on all non-cooking hours and on cooking hours defined uniformly as 6–8am and 6–9pm).

Two key patterns emerge. First, improved stove adoption causes a large and statistically significant reduction of  $52 \mu\text{g}/\text{m}^3$  in the 99th percentile of 10-minute means while cooking (Column 4 of Panel B), which corresponds to around a 42% reduction in the marginal emissions increase from cooking (over median non-cooking exposure) when compared with

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<sup>23</sup>Pollution impacts occur slightly after cooking, possibly because smoke can persist for some time after the user stops cooking, or because respondents start lighting the stove some time before actually cooking and do not consider this cooking in the time use survey.

Table 3: IV estimates of the causal impact of cookstove adoption on PM2.5 exposure

Panel A) All hours

	(1) Median	(2) Mean	(3) Max Hour	(4) 99th
Own Jikokoa	0.1 (1.7)	-0.8 (3.4)	-16.4 (19.0)	-8.3 (23.0)
Control Mean	25.2	37.8	153.3	200.3
Weak IV F-Statistic	53	53	53	53
Observations	651	651	651	651

Panel B) When self-reporting cooking

	(1) Median	(2) Mean	(3) Max Hour	(4) 99th
Own Jikokoa	-11.0** (5.2)	-16.6*** (6.4)	-31.0** (15.4)	-52.0** (22.5)
Control Mean	35.9	49.7	92.6	150.3
Weak IV F-Statistic	48	48	48	48
Observations	598	598	595	598

Each column is an IV regression where the randomly assigned price, credit treatment status, and their interaction are used as instruments for endline Jikokoa ownership. Column (1) uses median exposure, (2) uses mean exposure, (3) uses maximum 1-hour average exposure, and (4) uses 99th percentile of 10-min average exposure. Of the 702 respondents surveyed, 651 consented to having a PA-II device ([Section 4.7](#) discusses attrition); some of them never self-reported cooking. Regressions include socioeconomic controls and a fixed effect for the specific PA-II device that respondent used. [Table B8](#) presents the same for carbon monoxide. [Table B7](#) presents the same for when self-reporting not cooking and for the hours between 6–8am and 6–9pm. [Table B9](#) presents all outcomes in logs.

the control group.<sup>24,25</sup> This mirrors the 41% reduction in charcoal expenditures identified in [Table 2](#): PM2.5 emissions from cooking appear to decrease proportionally to charcoal usage. Columns 3 and 4 of Panel A show that there is a smaller reduction in max hourly and 99th percentile of 10-minute pollution over all hours of the day, which includes spikes generated for example by traffic. While cooking with a traditional charcoal jiko might generate larger spikes than commuting, this may not be the case for cooking with a Jikokoa, such that for many adopters the largest spikes no longer occur during cooking. These patterns are economically and statistically similar when the data are analyzed in logs ([Table B9](#)).

Second, despite large emissions reductions during cooking, there is only a 2% reduction in aggregate average exposure, and this reduction is not statistically significantly different

<sup>24</sup>To estimate the marginal emissions from cooking we first estimate typical ambient exposure as the average across control participants of their median 10-minute windows; this is  $25.2 \mu\text{g}/\text{m}^3$ . We then estimate typical cooking exposure as the average across control participants of their 99th percentile 10-minute windows during hours while cooking; this is  $150.3 \mu\text{g}/\text{m}^3$ . Differencing these two estimates gives an estimated marginal effect of cooking of  $125.1 \mu\text{g}/\text{m}^3$ .

<sup>25</sup>These results hold regardless of how we define these short-term increases: versions studying the 90th and 95th percentiles of 10-minute means, as well as the 90th, 95th, and 99th percentiles of the raw pollution distribution, show virtually identical impacts.

from zero (Column 2 of Panel A in [Table 3](#)). This can be rationalized with a simple equation describing an individual's average pollution exposure:

$$PM = PM_c \cdot p_c + PM_t \cdot p_t + PM_a \cdot (1 - p_c - p_t)$$

Where  $PM_c$  is average pollution while cooking,  $PM_t$  is average pollution while traveling, and  $PM_a$  is average ambient pollution;  $p_c$  is the proportion of time spent cooking and  $p_t$  is the proportion of time spent traveling. Pope et al. ([2018](#)) document average roadside PM2.5 levels in Nairobi of  $PM_t = 37\mu g/m^3$  and average urban background levels of  $PM_a = 25\mu g/m^3$ . Our measurements indicate average pollution levels while cooking with a traditional stove is  $PM_c = 49.7\mu g/m^3$  (Panel B, Column 2 of [Table 3](#)). The time use data we collected ([Table B3](#)) indicate that respondents cook for on average 9% of the time (2 hours per day) and travel (primarily walking or riding the bus) on average 10% of the time (2.4 hours per day). Under these conditions, a 50% reduction in emissions while cooking would reduce average daily emissions by 8%, which is within the confidence intervals of the estimates in ([Table 3](#)). Even a 100% reduction in cooking emissions (for example from the adoption of an electric stove) would cause only a 16% reduction in average PM2.5 exposure. In this urban context, the reduction in cooking-related pollution causes only a small and statistically undetectable reduction in total pollution exposure.

We can apply this calculation to rural areas as well, where ambient air pollution is on average  $9\mu g/m^3$  (Pope et al., [2018](#)). Conservatively supposing that participants cook for twice as long in rural areas as in urban areas, a 50% reduction in pollution during cooking would generate a 25% reduction in aggregate exposure.<sup>26</sup> While this is a sizable reduction, if the dose-response function is concave (as in for example Goodkind, Coggins, and Marshall, [2014](#); Pope III et al., [2015](#)), then the impact of improved cooking technologies may be undermined by the presence of ambient air pollution. This points to the importance of measuring ambient air pollution, even for studies only targeting reductions in air pollution caused by cooking.

It is worth noting that many people likely move away from their stove or increase ventilation at least sometimes while cooking, which reduces average cooking concentrations measured by personal exposure devices. While realized pollution exposure is the relevant object for policy and health research, the reduction in pollution caused by improved stove adoption might have been larger if measured using stationary stove monitors. The reductions measured by personal monitors could thus be considered a lower bound on total emissions reductions.

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<sup>26</sup>Unfortunately we are unable to use our own study data for this. Due to logistical surveying constraints in rural areas, most study participants residing in rural areas did not receive air pollution monitoring devices.

Cooking hours tend to coincide with traffic patterns: both have a morning and an evening spike. Using hourly data on self-reported cooking activity and pollution allows us to include hour-of-day fixed effects in the regressions. However, this significantly reduces identifying variation since there is a strong correlation between hour of day and propensity to be cooking. We therefore estimate this regression using both IV and OLS specifications ([Table B10](#)). While the IV estimates are noisier than the OLS estimates, the results present a similar story: improved stove adoption does not affect PM during non-cooking hours but reduces average PM<sub>2.5</sub> significantly during self-reported cooking hours.

We see no impacts on CO ([Table B8](#)), in line with independent laboratory tests scoring the Jikokoa Tier 3 for PM<sub>2.5</sub> but Tier 1 for CO (CREEC, [2022](#)). A stove's CO output generally depends on its oxygenation rate: higher oxygen inflow increases CO<sub>2</sub> production and reduces CO production while cooking. Per the company's engineers, the lack of reduction in CO output results from a desire to increase the durability of the stove by limiting peak cooking temperatures to 700°C. While this improves durability, it limits oxygenation.

Finally, we estimate the impacts of spending a specific amount of time above a certain air pollution threshold, for different thresholds, mirroring common regulatory proposals ([Table B11](#)). This dramatically reduces the power of the estimation, as only movements across the threshold will generate a treatment effect. When trying to answer regulatory or health questions, looking at averages or other moments in the distribution can often generate more precise statistical results than relying on data on thresholds.

### 4.3 Impacts of stove ownership on health

How does the 42% reduction in cooking emissions spikes affect health? [Table 4](#) estimates the impact of stove adoption on health outcomes, using the IV approach discussed above and controlling flexibly for age and linearly for other baseline socioeconomic outcomes. Column (2) uses 2022 Jikokoa ownership as the endogenous variable while Column (3) uses 2019 Jikokoa adoption. The first outcome is an index of clinical health measurements. The next six outcomes are indices and counts of self-reported health outcomes. Following our pre-analysis plan (Berkouwer and Dean, [2022b](#)) we separate self-reported health symptoms into those related to the respiratory system and those not.

The results indicate a 0.24 standard deviation reduction in a summary index of self-reported symptoms directly related to pollution, such as sore throat, headache, and cough ([Table B13](#) presents more detailed results on pollution-related symptoms) and 0.48 fewer respiratory symptoms reported. In [Section 4.6](#), we consider the potential role of experimenter demand in creating this effect and conclude the evidence suggests it is likely minimal. This

Table 4: IV estimates of the causal impact of cookstove adoption on primary health outcomes

	Control Mean (1)	Treatment Effect (2022 Ownership) (2)	Treatment Effect (2019 ownership) (3)	N
Physiological health index (blood oxygen and blood pressure)	0.00 [1.00]	0.02 (0.17)	0.02 (0.13)	696
Number of non-respiratory health symptoms	1.09 [1.54]	-0.24 (0.25)	-0.18 (0.19)	702
Non-respiratory health symptom index	-0.00 [1.00]	-0.03 (0.19)	-0.03 (0.14)	702
Number of respiratory health symptoms	1.70 [1.76]	-0.48** (0.23)	-0.36** (0.17)	702
Respiratory health symptom index	-0.00 [1.00]	-0.24* (0.13)	-0.18* (0.10)	702
Health diagnoses index	0.00 [1.00]	0.13 (0.16)	0.10 (0.12)	702
Number of health diagnoses	0.30 [0.58]	0.13 (0.09)	0.10 (0.07)	702
Cognitive index	-0.00 [1.00]	-0.01 (0.15)	-0.02 (0.12)	587
Healthcare utilization index (spending and visits)	-0.00 [1.00]	0.08 (0.14)	0.05 (0.11)	702

Each row is an instrumental regression wherein endline modern stove use is instrumented for with randomly assigned price, credit treatment status, and their interaction. Regressions include socioeconomic controls and control for hour of day of the second visit, where blood pressure and blood oxygen were recorded.

[Table B12](#), [Table B13](#), [Table B14](#), [Table B15](#), and [Table B16](#) present detailed results on the components of the diagnoses, respiratory symptoms, non-respiratory symptoms, cognitive, and healthcare utilization indices. Outcomes for children are presented in [Table B17](#).

reduction is sizeable, though smaller than the 0.56 standard deviation reduction in self-reported symptoms estimated using data from the one-year follow-up survey and relatively imprecisely estimated (Berkouwer and Dean, [2022a](#)).

However, we identify no long-term health improvements in clinical outcomes such as blood oxygen, blood pressure, and self-reports about any diagnoses made by a medical professional during a hospital visit ([Table B14](#) and [Table B12](#) present more detailed results on non-pollution related symptoms and medical diagnoses, respectively).

Specifically, we are powered to reject that owning a stove at endline decreased our diagnoses index by more than 0.14 SD and that it increased our physiological health index (composed of blood pressure and pulse oximetry) by more than 0.27 SD. [Table 5](#) presents the impacts on the individual components of the physiology index. To interpret these results clinically, it's useful to focus on the systolic component where we are powered to reject a de-

Table 5: IV estimates of the causal impact of cookstove adoption on physiological outcomes

	Control Mean	Treatment Effect	N	q-value
Average systolic blood pressure	122.16 [18.97]	0.49 (3.30)	696	0.88
Average diastolic blood pressure	81.32 [11.73]	0.58 (2.15)	696	0.88
Hypertension: Stage 1 or higher (>130/80)	0.51 [0.50]	0.02 (0.09)	696	0.88
Hypertension: Stage 2 or higher (>140/90)	0.27 [0.44]	-0.02 (0.08)	696	0.88
Blood oxygen	96.61 [2.53]	0.31 (0.37)	696	0.88
Low blood oxygen: 95 or lower	0.21 [0.41]	-0.08 (0.06)	696	0.88

Each row is an IV regression where randomly assigned price, credit treatment status, and their interaction are used as instruments for endline Jikokoa ownership. Regressions include socioeconomic controls.

crease of 5.97 mm Hg with 95% confidence. Ettehad et al. (2016) conducted a meta-analysis of 123 randomized controlled trials examining the health impacts of reducing systolic blood pressure. They find a 10 mm Hg reduction is associated with a 20% reduction in the risk of a major cardiovascular event off a base of 11%. Applying this estimate to our results suggests we are powered to reject a change large enough to reduce major cardiovascular events by 12%. As another point of comparison, smoking a cigarette causes an acute increase in systolic blood pressure of 20 mm Hg, however the impact of smoking on weight loss complicates inference about the impact of smoking on long-term blood pressure through the respiratory channel per se (Cohen and Townsend, 2009).

We find no effect on the number of hospital visits, hospital-related expenditures, or any of the cognition outcomes (Table B15).<sup>27</sup> Finally, we see no impacts on childrens' outcomes. We cannot detect a statistically significant impact on a range of child health outcomes, including child anthropometrics (raw and WHO growth standard-adjusted z-scores of weight, height, and arm circumference), a range of self- or parent-reported symptoms, and two types of attempted pneumonia diagnoses (Table B17), neither among children under 10 nor when restricting the sample to children under 6, who are more likely to stay at home during the day.

Participants' baseline beliefs about the stove's health benefits on average lines up well with these findings. At baseline, 37% of respondents believed that adoption of the improved stove would have no impact on their health, 34% believed it would have a small or medium

<sup>27</sup>Due to a technical issue with the tablets the sample size for some of the cognition outcomes is smaller than in other outcome tables. Since this was a technical issue, and since the order of follow-up surveys was randomized, it is unlikely that this biased the results in any meaningful way.

impact, and only 29% believed it would have a large or very large impact. Participants may have already sensed that reductions in cooking-related pollution will have smaller health impacts in the presence of the high levels of ambient pollution they experience every day than has been previously argued for by policy makers. Furthermore, respondents who believe that the stove will have larger health impacts on average do not have higher WTP, unlike beliefs about financial savings which are strongly correlated with WTP. Table 4 of Berkouwer and Dean (2022a) reports health and savings beliefs in different units. Standardizing both outcomes yields the following results: increasing health beliefs by 1 SD decreases WTP by \$0.01 ( $p=0.988$ ) while increasing savings beliefs by 1 SD increases WTP by \$0.79 ( $p=0.036$ ). This suggests that even those that reported believing the stove would have large or very large impacts, do not feel that the health improvements would be sufficiently meaningful to pay for.

#### 4.4 Physiological mechanisms: Relating pollution and health

As discussed in [Section 1](#), air pollution spikes may have very different health impacts than the mean daily levels that were investigated in much of the literature studying ambient air pollution. The significant reduction in intense, short-term spikes likely contributed to the reduction of self-reported—and largely transient rather than chronic—respiratory health symptoms. At the same time, the lack of reduction in aggregate average pollution exposure may explain the lack of impacts on chronic or quantitative health outcomes, despite 3.5 years of sustained use with reduced spikes in air pollution.

These results are consistent with several models of physiological pathways. One possibility is that short-term spikes drive short-term symptoms, whereas clinical health impacts are driven primarily by long-term ambient exposures. In other words, while reductions in spikes in exposure can generate important short-term health benefits, improvements in long-term measures of health may require reductions in ambient air pollution exposure. If this is the case, our results also have distributional implications regarding who is exposed to pollution: in addition to being less able to afford improved private technologies, respondents with lower wealth also on average face higher levels of ambient air pollution ([Table B19](#)). Another possibility is that the reduction in acute symptoms may be driven by psychological effects. We discuss this possibility in more detail in [Section 4.6](#).

[Table 6](#) shows correlations between health and three key moments of pollution: average pollution exposure (in  $100 \mu\text{g}/\text{m}^3$ ), pollution exposure spikes (defined as the highest hourly average recorded, in  $100 \mu\text{g}/\text{m}^3$ ), and the duration of high pollution (above in  $100 \mu\text{g}/\text{m}^3$ ) exposure ([Table B20](#) and [Table B21](#) show more detailed outcomes, and specifications that

Table 6: Correlations with different moments of the PM2.5 distribution

	Mean (1)	Mean Pollution in SD (2)	Median Pollution in SD (3)	Max Hourly Pollution in SD (4)	Hours Above $100\mu g/m^3$ (5)	N (6)
Hypertension (>130/80)	0.51 [0.50]	0.01 (0.02)	-0.02 (0.02)	0.00 (0.02)	0.00 (0.01)	645
Blood oxygen	96.72 [2.43]	0.12 (0.10)	0.12 (0.11)	-0.03 (0.10)	0.03 (0.06)	645
Health symptoms index (z-score)	-0.09 [0.92]	0.01 (0.04)	-0.01 (0.04)	0.07** (0.04)	0.01 (0.02)	651
Number of health symptoms	2.52 [2.66]	0.02 (0.11)	-0.00 (0.11)	0.23** (0.10)	0.02 (0.06)	651
Health diagnoses index	-0.04 [0.89]	-0.04 (0.04)	-0.05 (0.04)	0.00 (0.04)	-0.03 (0.02)	651
Number of health diagnoses	0.29 [0.56]	-0.03 (0.02)	-0.02 (0.03)	-0.00 (0.02)	-0.02 (0.01)	651
Hospital visits in past 30 days	0.30 [0.55]	-0.01 (0.02)	-0.01 (0.02)	0.01 (0.02)	-0.00 (0.01)	651
Hospital visit expenditures (USD)	2.82 [10.14]	0.66 (0.44)	0.40 (0.45)	0.62 (0.42)	0.26 (0.24)	651

Each row and column cell in Columns (2)–(5) is a separate OLS regression. All regressions include socioeconomic controls, air pollution device FE, month of survey, and baseline WTP. Table B20 shows additional variables. Table B13, Table B14, and Table B12 present detailed results on symptoms and diagnoses.

control for average PM2.5 and CO pollution, respectively).<sup>28</sup> Hourly pollution spikes are strongly correlated with self-reported health symptoms, but not with any of the clinical health outcomes. That mean and median PM2.5 air pollution are not correlated with health is likely due to an absence in sufficient identifying variation in ambient exposure, rather than an absence of a relationship between aggregate pollution exposure and health.

We do not find evidence of heterogeneity in treatment impacts along the lines of baseline health, baseline beliefs about future health impacts, age, WTP, baseline charcoal expenditures, or endline LPG ownership (Table B22). Ambient pollution is a potentially important source of heterogeneity, as some previous research has found air pollution improvements to be non-linear—either concave or convex—in average pollution. We test for heterogeneity in the primary treatment effect on health by whether the respondent has above or below median ambient air pollution. To avoid bias due to adoption endogeneity and noise in the time use data, we define a respondent’s ambient pollution as average pollution among the

<sup>28</sup>We present correlations here rather than instrumenting for pollution levels because we do not have enough strong instruments to separately identify changes in peaks and the ambient level of pollution.

five respondents residing nearest that respondent. We then test whether the health impacts differ by whether respondents' ambient exposure is above vs the median. We find no difference of heterogeneity along this dimension, at least over the range of pollution levels we observe ([Table B23](#)).

Since ambient pollution levels are generally lower in rural areas than in urban areas, study participants residing in rural areas may experience larger proportional pollution improvements. To examine whether health impacts are different for study participants residing in rural areas, we estimate the causal impact of adoption on health outcomes just among this sample. We do not find evidence of health improvements among this sub-sample—in fact, most point estimates point in the opposite direction. We refrain from over-interpreting this result because the rural sample is very small ( $n = 53$ ) and because moving is an endogenous choice that may bias the estimation ([Table B24](#)).

## 4.5 Impacts of stove ownership on socio-economic outcomes

Panel B of [Table 2](#) presents the impact of stove adoption on various socioeconomic outcomes ([Table B5](#) presents a more detailed version). Improved cookstove ownership causes a \$1.50 reduction in average weekly charcoal expenditures, or a 41% reduction relative to the control group. These savings correspond to 9.3% of the control group's average income: in other words, adoption saves more than one month of income per year. These numbers correspond closely to the estimates from Berkouwer and Dean ([2022a](#)). The USD conversion (approximately \$86 per year) is slightly lower due to Kenya's high inflation in recent years.

Importantly, the large financial savings could be a mechanism through which stove adoption improves health. For example, adopters could use the savings to improve the quality or quantity of the food they consume, purchase medical defense technologies such as insecticide-treated bednets or vaccinations, or afford medications and doctors' visits. Conversely, increases in hospital visits could cause us to underestimate the impact of pollution on health if they change the conditional likelihood of being diagnosed or informed of a disease by a medical professional. Still, several factors suggest the income effect is likely to be small in this context, with the health impacts more likely to be driven by the reduction in pollution spikes per se.

First, we see no change in the propensity of the respondent to purchase protein-rich foods such as meat, fish, or egg ([Table B6](#)). The propensity to cook Githeri increases, but Githeri contains only beans and maize, which more than 90% of participants were already consuming at baseline. Second, we see no change in healthcare utilization, which includes hospital visits, hospital expenditures, and non-hospital health expenditures ([Table B16](#)).

Third, we only find a significant reduction in self-reported respiratory symptoms, while we see no significant reduction either the non-respiratory symptom index or the number of non-respiratory symptoms. However, there are two important caveats to this result. First we are only powered to reject reductions of greater than 0.4 SD and 0.73 symptoms for the non-respiratory index and number of non-respiratory symptoms respectively. Second, we find a decrease in malaria which is significant at the 10% level before multiple hypothesis correction. We interpret this cautiously given that we see pre-correction significant increases in other non-respiratory health outcomes—such as worms—and the result on malaria is no longer significant after multiple hypothesis correction ( $q\text{-value}=0.19$ ). If the reduction in malaria rates is real, this likely has important quality-of-life benefits, however it is unlikely to explain the reduction in respiratory symptoms that we see.

This is consistent with the findings of Haushofer and Shapiro (2016) who study unconditional cash transfers of similar or larger amounts in Kenya (\$404 or \$1525). They estimate a reduction in the health index of 0.03 SD and can reject any increase larger than 0.09 SD. Taken together, this evidence suggests that income effects are not likely to be behind the self-reported health improvements that we see.

Adoption increases the propensity of individuals in an adopter’s network to adopt the stove. Specifically, it roughly doubles the number of Jikokoa stoves owned by members in a respondent’s network such as friends, family, and in particular neighbors. We see no impacts on savings or formal banking access, suggesting the financial savings may have been spent on consumption. We see no impacts on hours worked or earnings. Finally, we see no impact on time spent on various activities such as sleeping, working, eating, or walking (Table B25).

## 4.6 Psychological effects and experimenter demand

The difference in results between the self-reported respiratory symptoms and the clinical health outcomes could be caused by bias from self-reports. Existing research in the medical literature has documented some degree of discrepancy between patient self-reports and medical records.<sup>29</sup> In this section we explore the extent to which four versions of this issue may drive our results.

A critical concern when using self-reported data is that self-reports may be driven by experimenter demand where respondents provide answers that they believe will please the experimenter (Orne, 1962; Quidt, Haushofer, and Roth, 2018). Subsidies in the study ranged from 25%–90%, and participants who received a heavily subsidized cookstove might be more inclined to report better health than those who did not. While we cannot rule out some

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<sup>29</sup>See for example Skinner et al. (2005), Katz et al. (1996), Akker et al. (2014), and Fowles, Fowler, and Craft (1998). Tisnado et al. (2006) explore how concordance varies by patient characteristics.

amount of experimenter demand, several factors weigh against this fully explaining the effects. First, we test whether those with higher subsidies are more likely to report positive health even after controlling for stove adoption. If respondents with a lower price (higher subsidy) were more likely to self-report better health, price would correlate directly with self-reported symptoms rather than purely through the adoption channel ('owns Jikokoa'). We do not find evidence of this ([Table B26](#)). Second, there are correlations between self-reported non-respiratory symptoms and blood pressure as well as between self-reported health diagnoses and blood pressure, suggesting these self-reports carry a meaningful signal ([Table B18](#)). Third, the relationship between self-reported health symptoms and objectively measured pollution spikes provides credence to the self-reports ([Table 6](#)). Finally, self-reported health improvements arise primarily through respiratory rather than non-respiratory symptoms: participants would thus have to be sophisticated about our motivations as experimenters to selectively report improvements in some health outcomes but not others.

One other way to reconcile the impacts on self-reported symptoms directly related to pollution with the lack of impact on more objective outcomes is that the self-reports are driven by the “peak-end” effect. A classic psychology finding is that when evaluating experiences, individuals attend primarily to the peak intensity of the experience and the end of the experience (Fredrickson and Kahneman, [1993](#); Kahneman et al., [1993](#); Redelmeier and Kahneman, [1996](#)). In our context, this means that when asking someone about their symptoms, they may pay disproportionate attention to the symptoms experienced during smoke exposure spikes. Because the intensity of the spikes is reduced by the Jikokoa, these salient experiences may be reduced even without an effect on more enduring measures of health. It is important to note that this does not mean the self reports contain no signal of health experiences, but that they may be driven by peak experiences which may not translate into non-transitory health impacts.

Another possibility is the decline in reported acute symptoms, such as cough, may be due to memory distortions induced by the reduction in smoke. Significant literature in psychology has highlighted the fallibility of memory, and, in particular, the tendency of individuals to falsely recall events that did not occur but are associated with events that did occur (Schacter, Guerin, and Jacques, [2011](#); Schacter, [1999](#); Roediger and McDermott, [1995](#); Braun, Ellis, and Loftus, [2002](#); Deese, [1959](#); Bernstein and Loftus, [2009](#)). Thus, when answering the questions about respiratory symptoms, the respondent’s memory recall processes may have constructed a reduction in symptoms based on the true reduction in exposure to smoke.

A final related possibility is that subjects believe they are experiencing reduced symptoms, but that this is driven by placebo effects. While we cannot rule out this possibility

given our measures, three pieces of evidence point against this possibility. First, as noted above, we find that self-reports are correlated with our objective measures of peak pollution exposure. Second, [Table B22](#) shows that we find no heterogeneity by baseline beliefs of the health impacts in the treatment effect on either the number of reported respiratory health symptoms or the respiratory health symptoms index (point estimates of -0.02 and -0.00 respectively). If the results are driven by placebo effects, we would likely expect those who believed they would experience an effect to show larger effects. Finally, as noted above, the baseline beliefs were uncorrelated with willingness to pay. This suggests that even for individuals who believed there would be a health impact, they did not feel strongly enough about the improvement to stake any amount of money on it. We believe that makes it less likely that the belief would be strong enough to generate a placebo effect.

## 4.7 Attrition

702 of the 942 respondents (75%) were surveyed successfully during the three-year follow-up survey.<sup>30</sup> Attrition is not correlated with their randomly assigned BDM price, credit treatment assignment, initial Jikokoa stove adoption, or baseline health outcomes ([Table B27](#)). Attrition is slightly higher among respondents with fewer children, fewer household members, and younger respondents (such respondents may more easily move around, making them harder to track).

As a rule we attempted to survey any respondent still residing in Kenya. 167 respondents could not be contacted by phone. Physical attempts to track individuals were hampered by the recent demolitions of housing in Nairobi's settlement areas ([The Star, 2023](#)).

The remaining 73 respondents who were contacted but who chose not to complete the 2022-2023 survey made this choice for a variety of reasons, including nonavailability, non-consent, outmigration, or physical incapacitation ([Table B28](#)). Seven study participants had deceased.

## 5 Conclusion

Air pollution is a significant contributor to global morbidity and mortality. Extensive research documents a negative causal relationship between daily average air pollution and health. Yet, despite active regulatory and policy debate, there is scant evidence about the impacts of persistent transient spikes in air pollution and whether reducing these spikes can

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<sup>30</sup>13 of the 955 respondents who had completed the main visit in 2019 had withdrawn from the study between 2019 and 2022.

generate meaningful health improvements. This question is crucial in addressing pollution from, for example, commuting and cooking, two near universal sources of air pollution. However, the topic has been difficult to answer because economic activity, human behavior, and demographic patterns can all increase the wedge between pollution recorded by stationary monitors (which includes most regulatory monitors) and realized individual-level pollution exposure (the policy-relevant object for understanding health impacts).

To fill this gap, we conduct a field experiment studying an improved biomass cookstove in Kenya. Randomized subsidies and access to credit yield a persistent increase in adoption of a more energy efficient biomass cookstove that persists even more than three years later. We find that improved stove ownership causes a 42% reduction in air pollution spikes generated during cooking hours. As a result, respondents experience a 0.24 standard deviation improvement in self-reported respiratory symptoms.

However, we find a comprehensive lack of impacts on clinical health as measured through a range of outcomes, including blood pressure, blood oxygen, and diagnoses of chronic diseases (such as pneumonia), for both adults and children. This may be because of the lack of impacts on average pollution exposure: we observe no reduction during the remaining 22 hours of the day, and given the high levels of ambient pollution in this urban context, we see only a very small (2%) and statistically insignificant reduction in average air pollution exposure.

Taken together, these results are consistent with a physiological model where pollution spikes affect short-term health and daily averages affect clinical, chronic health. Given the high levels of ambient pollution experienced in many low- and middle-income country cities, this suggests that the urban poor have only limited ability to improve their health through the private adoption of improved technologies. Instead, clinical health may require government intervention addressing the negative pollution externality generating high levels of ambient pollution.

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