

TAMANI Project DD

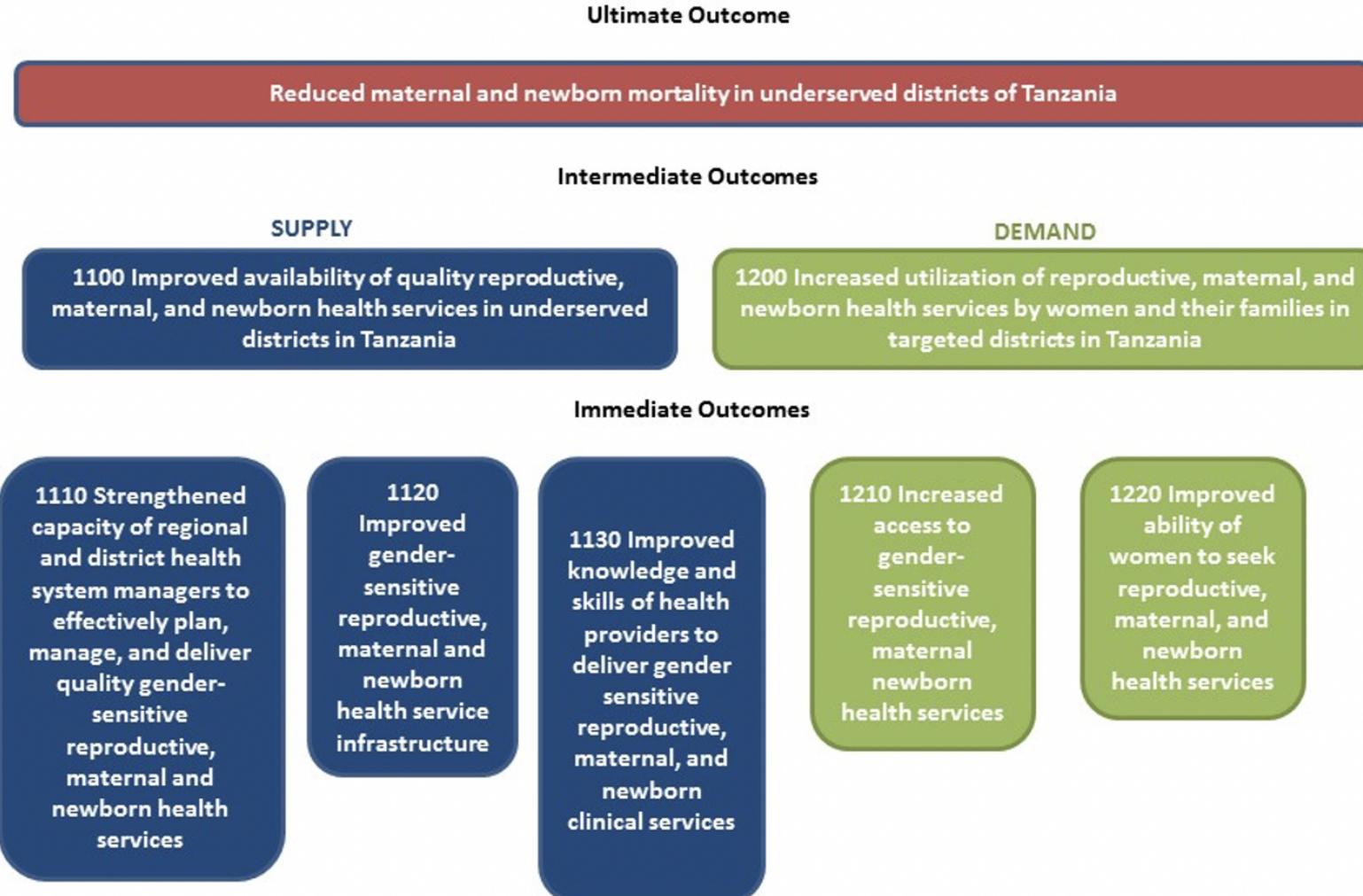
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joint with Erin H., Ari, Ilona, Holly

(all errors are mine)

2021-11-23

Tabora Maternal Newborn Health Initiative (TAMANI)



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Interventions

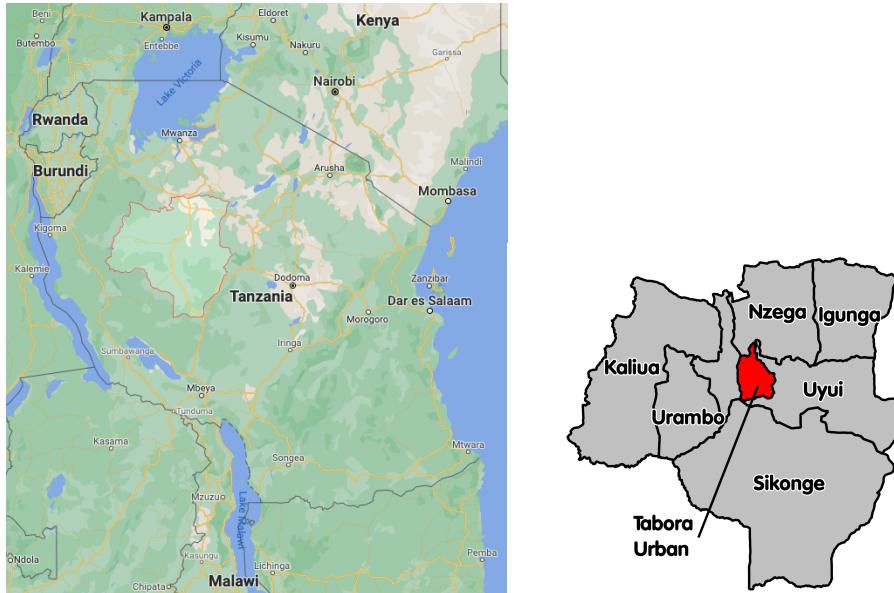
- Training health care providers in basic and comprehensive Emergency Obstetric and Newborn Care and family planning (EmONC)
- Community Health Worker training
- Community-based dialogues focused on gender and women's empowerment

Main outcomes

- Unmet needs for family planning
- **Skilled birth attendance**
- 4+ antenatal visits
- Contraceptive prevalence
- Respectful maternity care

Tabora Maternal Newborn Health Initiative (TAMANI)

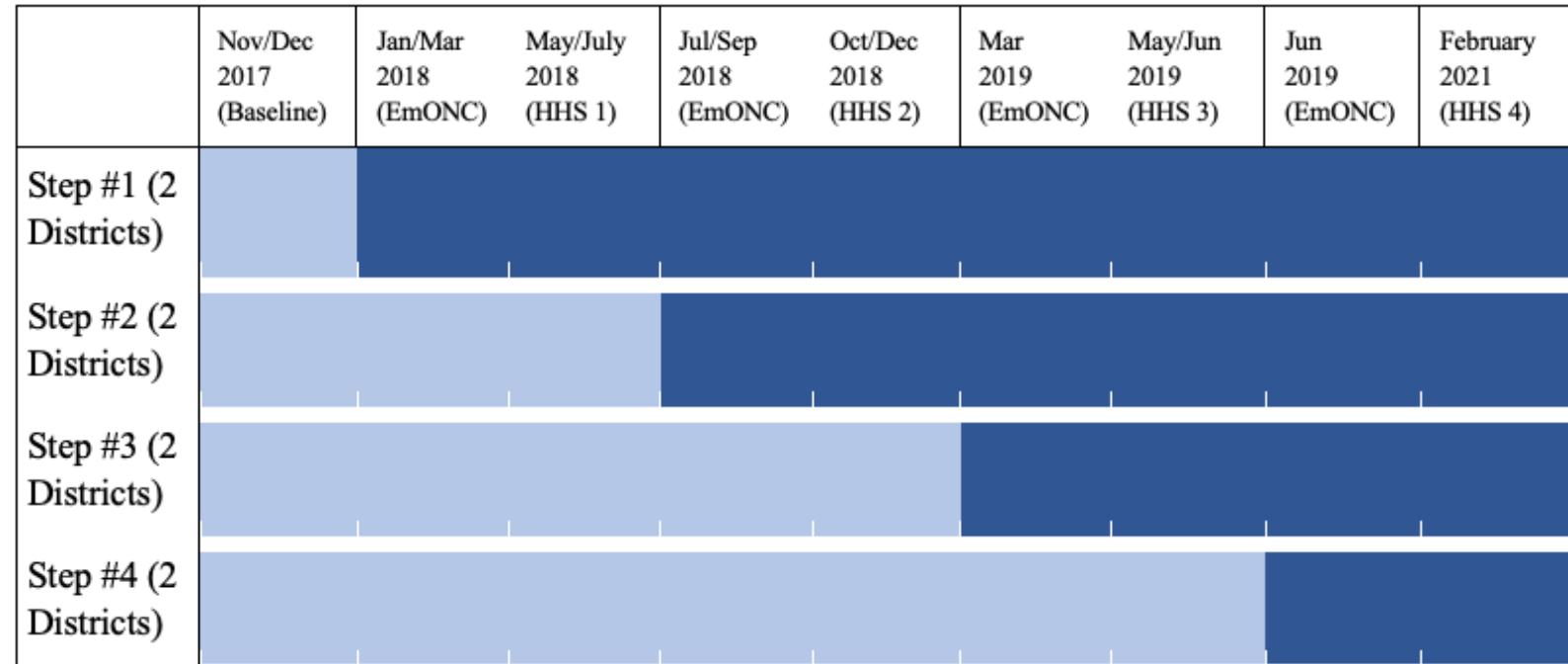
- Implemented by CARE Canada/Tanzania in 8 districts in Tabora region, Tanzania



Source: <https://commons.wikimedia.org/w/index.php?curid=47130439>, Google Maps

Initial Design: Stepped Wedge

- Political constraints led to breaking random timing.
- Switching of order of two districts.
- Analyze as DD?



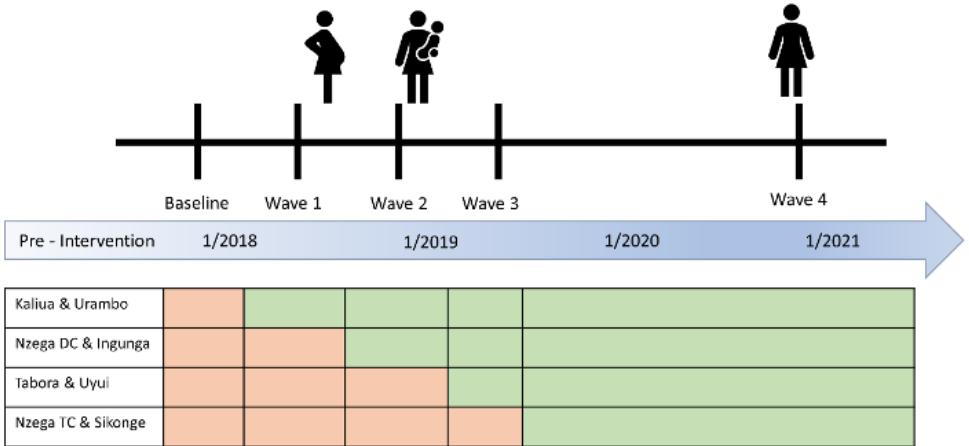
Note: EmONC=Emergency Obstetric and Newborn Care;
HHS=Household Survey



A note about defining 'treated'

- *Treated* at wave 4 For immediate outcomes (contraceptive prevalence, unmet need for family planning).
- Same woman reports on a child that was born in January 2019 (before the intervention was implemented).
- *Untreated* at wave 4 for Jan '19 delivery outcomes (skilled birth attendance and respectful care).
- Pregnancy began in the spring of 2018, so pregnancy outcomes (antenatal care visits) are considered “untreated”.

Respondent at wave 4 (January 2021).



Basic setup for DD with variable timing

We have different districts that are exposed to our intervention at different times. We often use OLS (or LPM) to fit:

$$y_{it} = \alpha_i + \tau_t + \beta^{DD} D_{it} + \epsilon_{it}$$

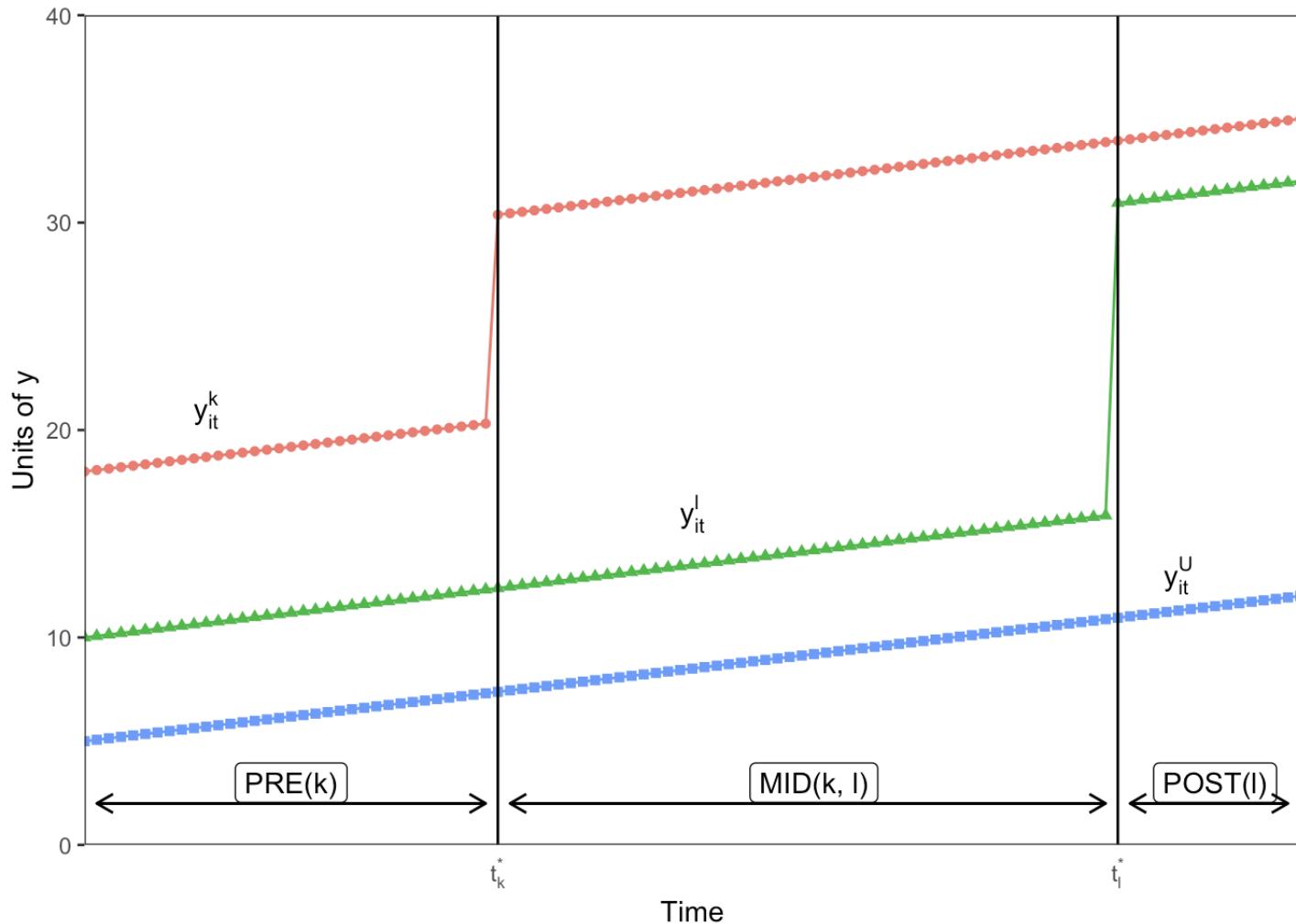
where

- y_{it} is the outcome for unit i at time t .
- α_i are unit-specific fixed effects.
- τ_t are fixed effects for each time period.
- D_{it} is a time-varying treatment indicator.
- β^{DD} is the difference-in-differences estimate.

Key points from Goodman-Bacon (2019)

- With OLS, DD with treatment timing is a variance-weighted average of many 2x2 ATTs.
- Weights are a function of both group sizes *and* variances.
- Can lead to β^{DD} that is a poor summary of group-specific effects.

1. Early-adopters (k) vs. never treated (U)
2. Later-adopters (l) vs. never treated (U).
3. Early (k) vs. later (l) adopters.
- 4. Later (l) vs. earlier (k) adopters.**



Graph from <https://andrewcbaker.netlify.app/2019/09/25/difference-in-differences-methodology/>

Data structure: Individual-level

- sba_birth = SBA present
- txdel = treated
- time = survey wave
- pid = person ID
- group = time when group first treated

district	sba_birth	txdel	time	dist_id	p_id	group
Kaliua DC	0	0	1	1	237	2
Kaliua DC	1	1	5	1	2168	2
Kaliua DC	0	1	5	1	6723	2
Kaliua DC	0	0	1	1	2269	2
Kaliua DC	1	1	5	1	5084	2
Kaliua DC	1	1	5	1	2058	2

Data structure: Pooled by district

- tsba = total SBA births
- tpop = total pop (births)
- txdel = treated
- time = survey wave
- group = time when group first treated

district	dist_id	time	tsba	tpop	psba	txdel	group
Kaliua DC	1	1	409	705	0.580	0	2
	1	2	49	73	0.671	1	2
	1	3	63	77	0.818	1	2
	1	4	9	10	0.900	1	2
Nzega DC	1	5	94	114	0.825	1	2
	2	1	465	634	0.733	0	3
	2	2	60	76	0.789	0	3
	2	3	43	57	0.754	1	3
Nzega DC	2	4	9	10	0.900	1	3

TWFE models (OLS)

- Individual and aggregate basically identical.
- Clustered SEs approximately the same.
- Intervention increased the Pr(SBA) by 6 pp (95% CI -1.3 to 13.3).

$$y_{sba} = \alpha + \beta * txdel + \gamma_{district} + \delta_{time} + \epsilon$$

	Individual	Aggregate
txdel	0.060	0.060
	(0.037)	(0.040)
	[-0.013, 0.133]	[-0.019, 0.138]
Num.Obs.	5555	40
Std. Errors	Clustered (district)	Clustered (district)
FE: district	X	X
FE: time	X	X

Callaway-Sant'Anna Approach: Group-Time cohorts

- The CS approach starts with defining Group-Time cohorts.
- Groups defined by when they were *first* treated.

Group	P(SBA)					Total Pop				
	Time1	Time2	Time3	Time4	Time5	Time1	Time2	Time3	Time4	Time5
2	0.631	0.711	0.814	0.846	0.826	898	90	102	13	144
3	0.678	0.770	0.664	0.800	0.784	1455	165	125	30	232
4	0.778	0.807	0.673	0.750	0.868	1128	140	101	16	167
5	0.854	0.939	0.854	0.800	0.855	556	66	41	10	76

- Different aggregation schemes for ATTs are possible.
- Can allow for covariates via regressions adjustments, IPW and DR.

Callaway-Sant'Anna implementation

- Includes options for different structure, SE calculation, weights, etc.

```
# Use not-yet-treated as comparison group
did::att_gt(yname = "sba_birth", # name of the LHS variable
             tname = "time", # name of the time variable
             idname = "p_id", # name of the id variable
             gname = "group", # name of the first treatment period
             data = d_ind, # dataset
             xformula = NULL, # conditional parallel-trends
             weightsname = NULL, # can add weights
             est_method = "reg", # estimation method
             control_group = "notyettreated", # set the control group
             bstrap = TRUE, # compute bootstrapped SE
             biters = 1000, # bootstrap iterations
             print_details = FALSE, # if TRUE, print detailed results
             panel = FALSE, # panel or repeated cross-sectional
             clustervars = NULL) # cluster ID
```

See <https://bcallaway11.github.io/did/articles/multi-period-did.html> for R, <https://econpapers.repec.org/software/bocbocode/S458976.htm> for Stata

Estimates from CS approach

- Note there is no *overall* estimate, lots of heterogeneity.
- Each treatment group has an ATT at each time period.
- Can be combined to produce different aggregate ATTs.

term	group	time	estimate	std.error	conf.low	conf.high
ATT(2,2)	2	2	0.011	0.053	-0.094	0.116
ATT(2,3)	2	3	0.260	0.057	0.149	0.371
ATT(2,4)	2	4	0.269	0.174	-0.072	0.610
ATT(3,2)	3	2	0.045	0.043	-0.039	0.129
ATT(3,3)	3	3	0.018	0.070	-0.118	0.155
ATT(3,4)	3	4	0.170	0.160	-0.145	0.484
ATT(4,2)	4	2	-0.062	0.043	-0.147	0.022
ATT(4,3)	4	3	-0.048	0.082	-0.208	0.112
ATT(4,4)	4	4	0.130	0.186	-0.234	0.495

Note:

P-value for pre-test of parallel trends assumption: 0.31

Re-creating the Group-Time ATTs

- $\text{ATT}(2,2)$ means estimating ATT *at time 2* for the group *first treated at time 2*
- For $\text{ATT}(2,2)$ we are comparing $\text{Pr}(\text{SBA})$ between:

group	time1	time2	time3	time4	time5
2	0.631	0.711	0.814	0.846	0.826
3	0.678	0.770	0.664	0.800	0.784
4	0.778	0.807	0.673	0.750	0.868
5	0.854	0.939	0.854	0.800	0.855

Note:

Red = treated, Gray = untreated

group	time1	time2	time3	time4	time5
2	0.631	0.711	0.814	0.846	0.826
3	0.678	0.770	0.664	0.800	0.784
4	0.778	0.807	0.673	0.750	0.868
5	0.854	0.939	0.854	0.800	0.855

Note:

Red = treated, Gray = untreated

- The 2x2 (weighted averages):

g22	time1	time2	Long diff	ATT_2_2
0	0.745	0.814	0.069	NA
1	0.631	0.711	0.080	0.011

This estimate says that intervention increased the probability of an SBA birth by 0.01 for Group 2 at Time 2.

- For ATT(2,3) the groups being compared are:

group	time1	time2	time3	time4	time5
2	0.631	0.711	0.814	0.846	0.826
3	0.678	0.770	0.664	0.800	0.784
4	0.778	0.807	0.673	0.750	0.868
5	0.854	0.939	0.854	0.800	0.855

Note:

Red = treated, Gray = untreated

- ATT(2,3) means estimating the ATT *at time 3* for the group *first treated at time 2*
- Need to exclude any group treated at time 3 to avoid bias.
- The group first treated at time 3 is excluded.
- Provides an estimate of the lagged impact of intervention.

group	time1	time2	time3	time4	time5
2	0.631	0.711	0.814	0.846	0.826
3	0.678	0.770	0.664	0.800	0.784
4	0.778	0.807	0.673	0.750	0.868
5	0.854	0.939	0.854	0.800	0.855

Note:

Red = treated, Gray = untreated

- The 2x2 (weighted averages):

g23	time1	time3	Long diff	ATT_2_3
0	0.803	0.725	-0.078	NA
1	0.631	0.814	0.182	0.26

Etc., etc., etc...

- For ATT(4,2) the groups being compared are:

group	time1	time2	time3	time4	time5
2	0.631	0.711	0.814	0.846	0.826
3	0.678	0.770	0.664	0.800	0.784
4	0.778	0.807	0.673	0.750	0.868
5	0.854	0.939	0.854	0.800	0.855

Note:

Red = treated, Gray = untreated

- ATT(4,2) means estimating the ATT *at time 2* for the group *first treated at time 4*
- Need to exclude any group treated at time 2 to avoid bias.
- The group first treated at time 2 is excluded.
- Provides an estimate of the lead effects or non-parallel trends.

group	time1	time2	time3	time4	time5
2	0.631	0.711	0.814	0.846	0.826
3	0.678	0.770	0.664	0.800	0.784
4	0.778	0.807	0.673	0.750	0.868
5	0.854	0.939	0.854	0.800	0.855

Note:

Red = treated, Gray = untreated

- The 2x2 (weighted averages):

g42	time1	time2	Long diff	ATT_4_2
0	0.727	0.818	0.091	NA
1	0.778	0.807	0.029	-0.062

Our 'hand-calculated' ATT(4,2) of -0.062 is the same as the regression-based CS estimate:

term	group	time	estimate	std.error	conf.low	conf.high
ATT(2,2)	2	2	0.011	0.055	-0.098	0.120
ATT(2,3)	2	3	0.260	0.057	0.149	0.372
ATT(2,4)	2	4	0.269	0.161	-0.046	0.585
ATT(3,2)	3	2	0.045	0.043	-0.039	0.130
ATT(3,3)	3	3	0.018	0.064	-0.107	0.144
ATT(3,4)	3	4	0.170	0.156	-0.136	0.476
ATT(4,2)	4	2	-0.062	0.046	-0.152	0.027
ATT(4,3)	4	3	-0.048	0.083	-0.211	0.115
ATT(4,4)	4	4	0.130	0.189	-0.241	0.502

"Final" ATT(4,4) compares only 2 groups (4 clusters):

group	time1	time2	time3	time4	time5
2	0.631	0.711	0.814	0.846	0.826
3	0.678	0.770	0.664	0.800	0.784
4	0.778	0.807	0.673	0.750	0.868
5	0.854	0.939	0.854	0.800	0.855

Note:

Red = treated, Gray = untreated

- The 2x2 (weighted averages):

g44	time3	time4	Long diff	ATT_4_4
0	0.854	0.80	-0.054	NA
1	0.673	0.75	0.077	0.13

Questions / Help

Aggregating

What is the right way to conceptualize aggregating the ATTs?

Clustering

How to think about the 'right' standard error?

Dynamic Effects

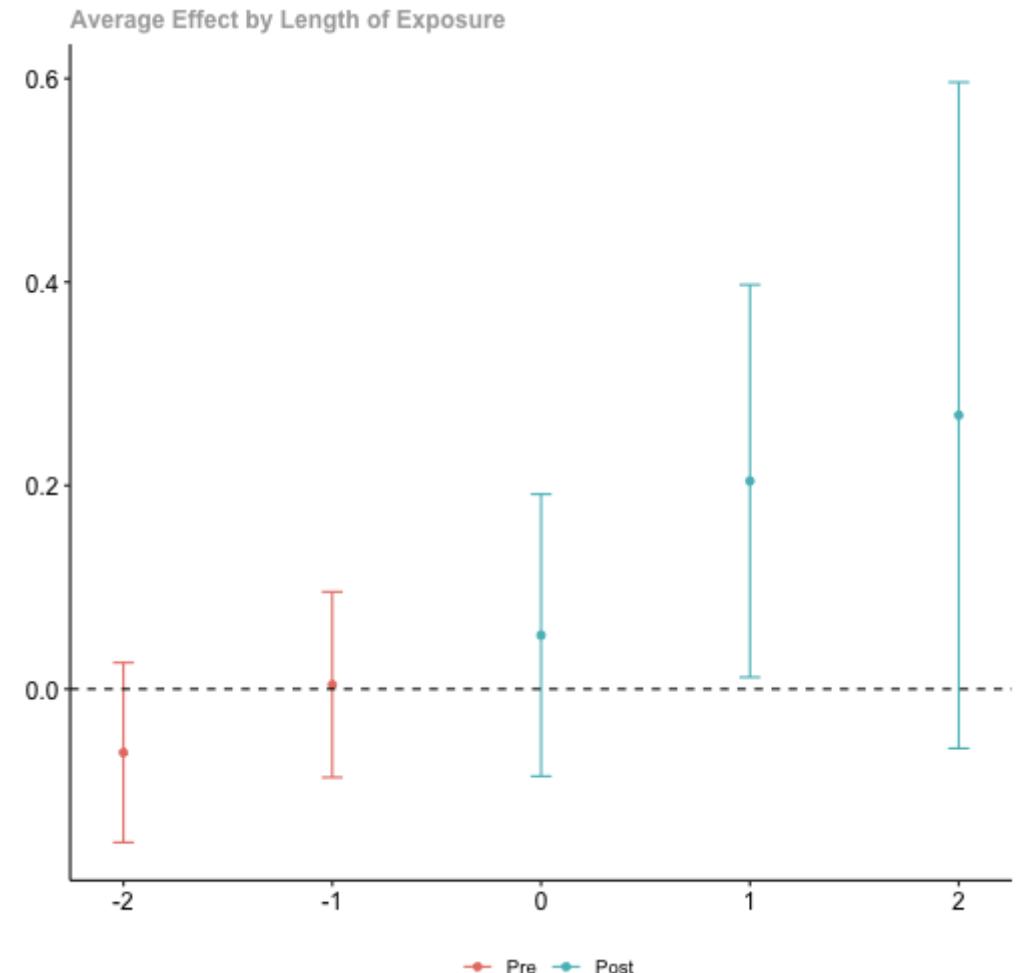
- Estimates by length of exposure
- Overall summary:

$$\overline{ATT} = 0.176(-0.02, 0.37)$$

event.time	estimate	std.error	conf.low	conf.high
-2	-0.062	0.045	-0.151	0.026
-1	0.004	0.047	-0.087	0.096
0	0.053	0.071	-0.086	0.192
1	0.204	0.098	0.012	0.397
2	0.269	0.167	-0.058	0.596

- TWFE estimate was

$$\overline{ATT} = 0.06(-0.013, 0.133)$$



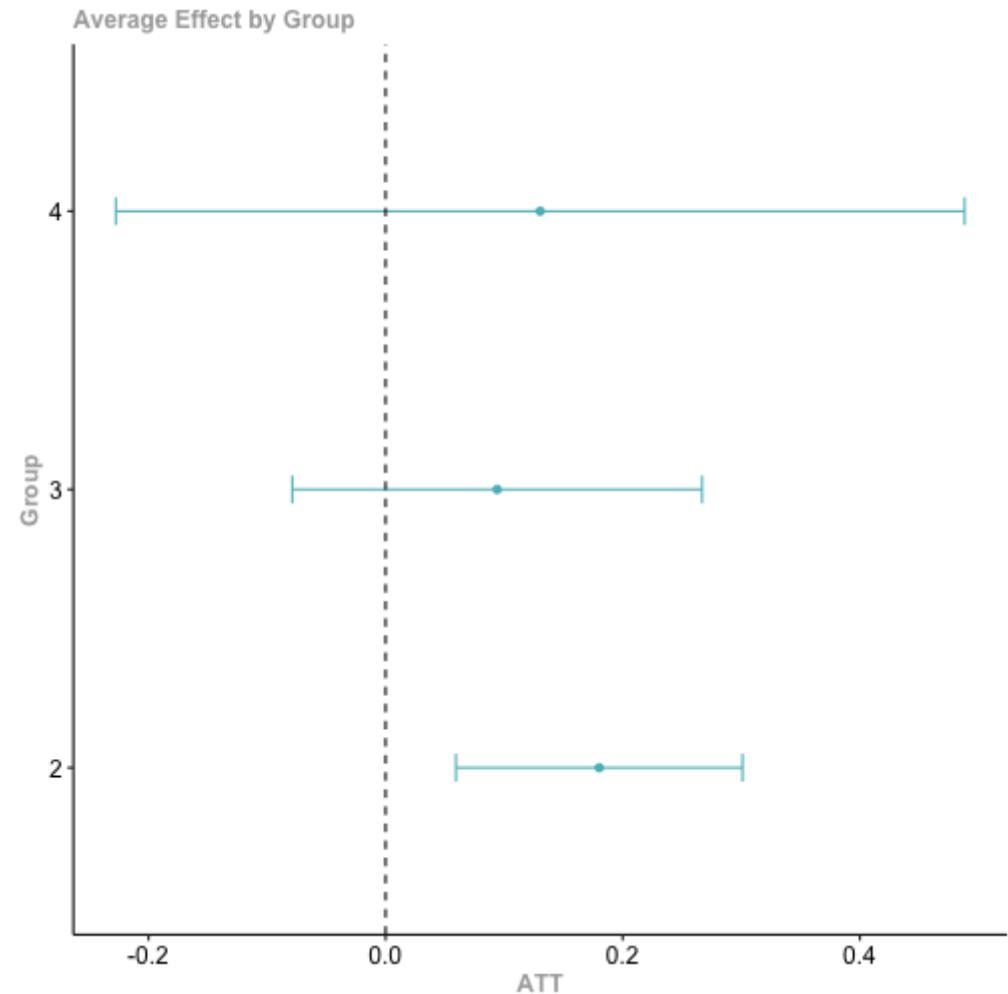
Group-specific ATTs

- Group-specific estimates
- Weighted avg of each group-time ATT for group g

group	estimate	std.error	conf.low	conf.high
Average	0.128	0.097	-0.061	0.318
2	0.180	0.062	0.059	0.301
3	0.094	0.088	-0.079	0.267
4	0.130	0.183	-0.228	0.488

- TWFE estimate was

$$\overline{ATT} = 0.06(-0.013, 0.133)$$



Calendar Time ATTs

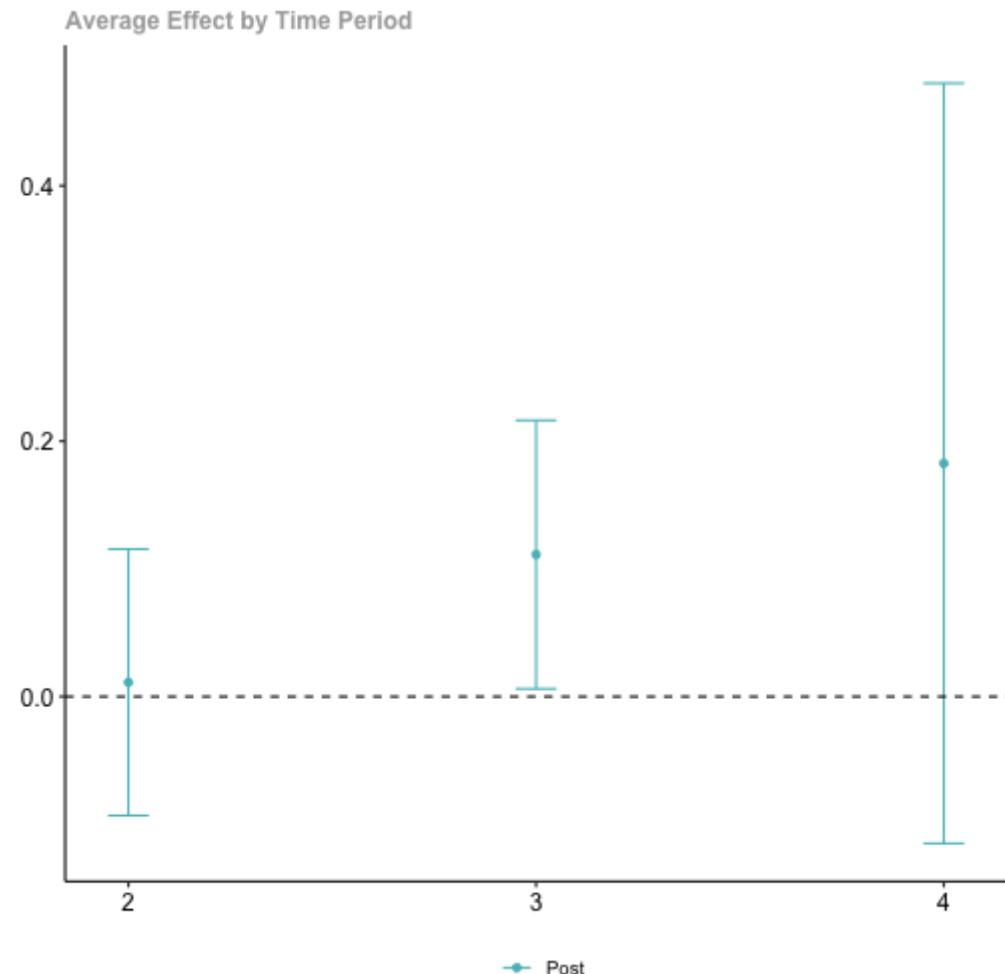
- ATT in time period t for groups that have participated in the treatment by time period t.
- Overall summary:

$$\overline{ATT} = 0.102(-0.02, 0.22)$$

time	estimate	std.error	conf.low	conf.high
2	0.011	0.053	-0.093	0.115
3	0.111	0.054	0.006	0.216
4	0.183	0.152	-0.115	0.480

- TWFE estimate was

$$\overline{ATT} = 0.06(-0.013, 0.133)$$



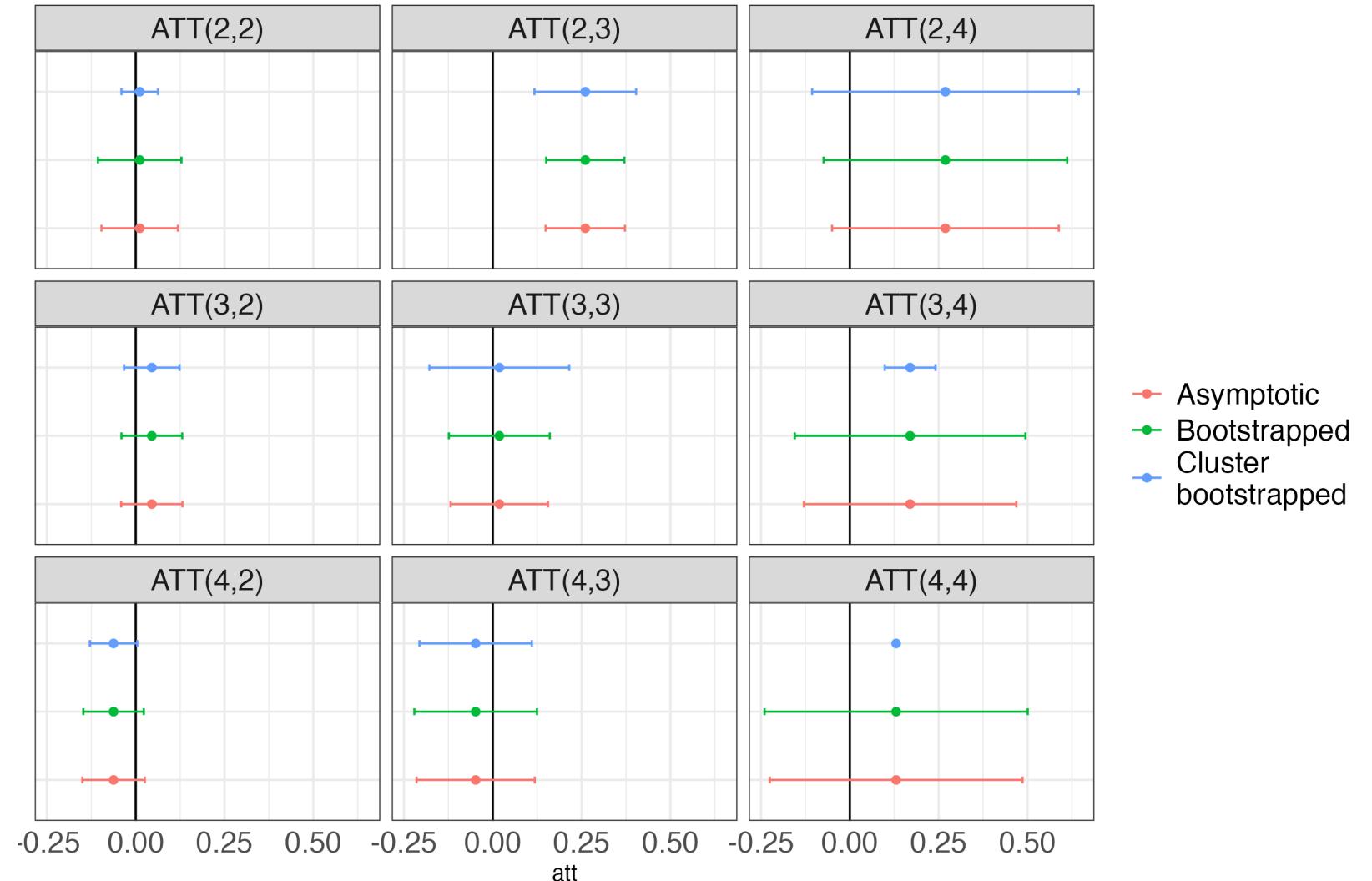
Options for Standard errors

- Asymptotic
- Bootstrap
- Cluster bootstrap

Comment from the Callaway-Sant'Anna paper:

Remark 10. In DiD applications, it is common to use “cluster-robust” inference procedures; see, e.g., [Wooldridge \(2003\)](#) and [Bertrand et al. \(2004\)](#). However, we note that the choice of whether to cluster or not is usually not obvious, and depends on the kind of uncertainty one is trying to reflect; see, e.g., [Abadie et al. \(2017\)](#) for a discussion in a cross-sectional setup.¹⁴ In the case that one wishes to account for clustering to reflect “cluster-based” sampling uncertainty, we note that this can be done in a straightforward manner using a small modification of the multiplier bootstrap described above, provided that the number of cluster is “large.” More precisely, instead of drawing observation-specific V 's, one simply needs to draw cluster-specific V 's; see, e.g., [Sherman and Le Cessie \(2007\)](#), [Kline and Santos \(2012\)](#), [Cheng et al. \(2013\)](#), and [MacKinnon and Webb \(2018, 2020\)](#). If the number of clusters is “small,” however, the application of the aforementioned bootstrap procedure is not warranted.¹⁵

- Some SEs hard to understand.
- Cluster bootstrapped smaller in most cases.
- Absurdly small for ATT(4,4) that has small sample and only 4 clusters.



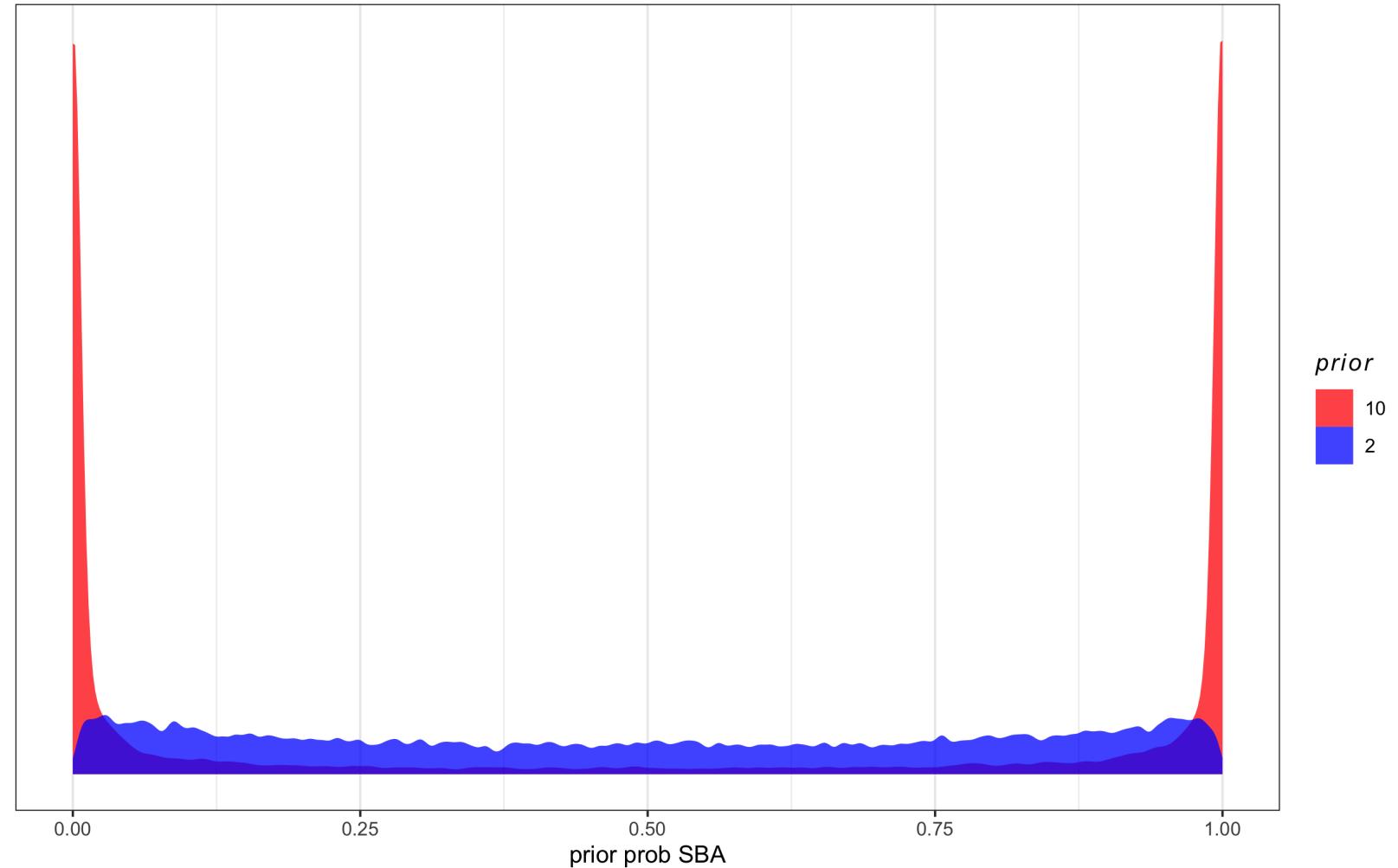
Other attempts to estimate ATT(4,4)

- Cluster-adjustment always excessively narrow.
- Ignoring clustering also seems wrong.
- Bayesian estimation more conservative.

Method	ATT(4,4)	SE	95% CI
Diff in proportions	0.130	0.182	-0.228, 0.489
Diff in proportions (svy clustered)	0.130	0.002	0.123, 0.138
OLS, robust SE	0.130	0.184	-0.232, 0.493
OLS, cluster robust SE	0.130	0.002	0.123, 0.138
Logit ME, robust	0.123	0.150	-0.171, 0.418
Logit ME, cluster robust	0.123	0.015	0.093, 0.154
Bayesian logit, flat priors	0.120	0.130	-0.230, 0.330
Bayesian logit, regularizing priors	0.060	0.100	-0.160, 0.230

Bayesian priors for intercept

$\alpha \sim \text{Normal}(0, prior)$



Bayesian priors for ATT

$\beta \sim \text{Normal}(0, prior)$

