## Assignment 1

## JRGLAY001, GMDSBO006 and MGRKHE001 2025-04-11



Figure 1: Let me cook!

## Question 1

a)

Model formulation

$$y = X\beta + e$$

where

$$e \sim \mathcal{N}(0, \sigma^2 I_n)$$

$$\mathbb{E}[y] = \mathbb{E}[X\beta + e] = \mathbb{E}[X\beta] + \mathbb{E}[e] = X\beta + 0 = X\beta$$
 
$$\operatorname{Var}[y] = \operatorname{Var}[X\beta + e] = \operatorname{Var}[X\beta] + \operatorname{Var}[e] + 2\operatorname{cov}(X\beta, e) = 0 + \sigma^2 I_n + 0 = \sigma^2 I_n$$

Therefore

$$y \sim \mathcal{N}(X\beta, \sigma^2 I_n)$$

Density of Y

$$\begin{split} f_Y(y) &= (2\pi\sigma^2)^{-\frac{n}{2}}exp(-\frac{1}{2\sigma^2}(y-X\beta)^T(y-X\beta))\\ f_Y(y) &= (2\pi\sigma^2)^{-\frac{n}{2}}exp(-\frac{1}{2\sigma^2}(y^Ty-2\beta^TX^Ty+\beta^TX^TX\beta)) \end{split}$$

Likelihood

$$L(\beta,y,X) \propto (\sigma^2)^{-\frac{n}{2}} exp(-\frac{1}{2\sigma^2}(y^Ty - 2\beta^TX^Ty + \beta^TX^TX\beta))$$

Since  $M=M^{-1}=I_{k+1}$ 

Then the prior distribution is

$$[\boldsymbol{\beta}|\sigma^2] \sim \mathcal{N}_{k+1}(\tilde{\boldsymbol{\beta}}, \sigma^2 M^{-1})$$

Prior density

$$\begin{split} \pi(\beta|\sigma^2) &= (2\pi)^{-\frac{k+1}{2}} \det(\sigma^2 M^{-1})^{-\frac{1}{2}} exp(-\frac{1}{2\sigma^2} (\beta - \tilde{\beta})^T M (\beta - \tilde{\beta})) \\ \pi(\beta|\sigma^2) &= (2\pi)^{-\frac{k+1}{2}} \det(\sigma^2 M^{-1})^{-\frac{1}{2}} exp(-\frac{1}{2\sigma^2} (\beta^T M \beta - 2\beta^T M \tilde{\beta} + \tilde{\beta}^T M \tilde{\beta})) \\ \pi(\beta|\sigma^2) &\propto exp(-\frac{1}{2\sigma^2} (\beta^T M \beta - 2\beta^T M \tilde{\beta} + \tilde{\beta}^T M \tilde{\beta})) \end{split}$$

Prior distribution

$$[\sigma^2] \sim \mathcal{IG}(a,b)$$

Prior density

$$\begin{split} \pi(\sigma^2) &= \frac{b^a}{\Gamma(a)} \sigma^{2(-a-1)} exp(-\frac{b}{\sigma^2}) \\ \pi(\sigma^2) &\propto \sigma^{2(-a-1)} exp(-\frac{b}{\sigma^2}) \end{split}$$

Assuming that  $\sigma^2$  is known

Prior

$$\pi(\boldsymbol{\beta}|\sigma^2) \propto exp(-\frac{1}{2\sigma^2}(\boldsymbol{\beta}^T M \boldsymbol{\beta} - 2\boldsymbol{\beta}^T M \tilde{\boldsymbol{\beta}}))$$

Likelihood

$$L(\beta|\sigma^2) \propto exp(-\frac{1}{2\sigma^2}(-2\beta^TX^Ty + \beta^TX^TX\beta))$$

Posterior

$$\pi(\beta|y,X) \propto \exp(-\frac{1}{2\sigma^2}(\beta^T(M+X^TX)\beta - 2\beta^T(X^Ty + M\tilde{\beta})))$$

Recall that  $(X^TX)\hat{\beta} = X^Ty$ 

$$\pi(\beta|y,X) \propto \exp(-\frac{1}{2\sigma^2}(\beta^T(M+X^TX)\beta - 2\beta^T((X^TX)\hat{\beta} + M\tilde{\beta})))$$

Let  $A = M + X^T X$  and  $b = (X^T X)\hat{\beta} + M\tilde{\beta}$ 

Completing the square using the following identity

$$\beta^T A \beta - 2 \beta^T b = (\beta - \mu_\beta)^T A (\beta - \mu_\beta) - \mu_\beta^T A \mu_\beta$$

where  $\mu_{\beta} = A^{-1}b$ 

$$\pi(\beta|y,X) \propto exp(-\frac{1}{2\sigma^2}((\beta-\mu_\beta)^TA(\beta-\mu_\beta)-\mu_\beta^TA\mu_\beta))$$

but  $\mu_{\beta}^T A \mu_{\beta}$  is independent of  $\beta$ 

$$\begin{split} \pi(\beta|y,X) &\propto exp(-\frac{1}{2\sigma^2}(\beta-\mu_\beta)^TA(\beta-\mu_\beta)) \\ \pi(\beta|y,X) &\propto exp(-\frac{1}{2\sigma^2}(\beta-\mu_\beta)^T(M+X^TX)(\beta-\mu_\beta)) \end{split}$$

meaning

$$[\beta|\sigma^2, y, X] \sim \mathcal{N}_{k+1}(\mu_{\beta}, \sigma^2(M + X^TX)^{-1})$$

where

$$\mu_\beta = (M + X^TX)^{-1}((X^TX)\hat{\beta} + M\tilde{\beta})$$

Joint distribution

$$J = [\beta, \sigma^2 | y, X]$$

Consider

$$P = L(\beta, y, X)[\beta | \sigma^2]$$

$$\begin{split} P \propto (\sigma^2)^{-\frac{n}{2}} exp(-\frac{1}{2\sigma^2}(y^Ty - 2\beta^TX^Ty + \beta^TX^TX\beta))(\sigma^2)^{-\frac{k+1}{2}} exp(-\frac{1}{2\sigma^2}(\beta^TM\beta - 2\beta^TM\tilde{\beta} + \tilde{\beta}^TM\tilde{\beta})) \\ P \propto (\sigma^2)^{-\frac{n+k+1}{2}} exp(-\frac{1}{2\sigma^2}(y^Ty + \tilde{\beta}^TM\tilde{\beta})) exp(-\frac{1}{2\sigma^2}(\beta^T(M + X^TX)\beta - 2\beta^T((X^TX)\hat{\beta} + M\tilde{\beta}))) \end{split}$$

Let  $A = M + X^T X$  and  $b = (X^T X)\hat{\beta} + M\tilde{\beta}$ 

Completing the square using the following identity

$$\beta^T A \beta - 2 \beta^T b = (\beta - \mu_\beta)^T A (\beta - \mu_\beta) - \mu_\beta^T A \mu_\beta$$

where  $\mu_{\beta} = A^{-1}b$ 

$$\begin{split} P &\propto (\sigma^2)^{-\frac{n+k+1}{2}} exp(-\frac{1}{2\sigma^2}(y^Ty + \tilde{\beta}^TM\tilde{\beta})) exp(-\frac{1}{2\sigma^2}((\beta - \mu_\beta)^TA(\beta - \mu_\beta) - \mu_\beta^TA\mu_\beta)) \\ P &\propto (\sigma^2)^{-\frac{n+k+1}{2}} exp(-\frac{1}{2\sigma^2}(y^Ty + \tilde{\beta}^TM\tilde{\beta} - \mu_\beta^TA\mu_\beta)) exp(-\frac{1}{2\sigma^2}(\beta - \mu_\beta)^TA(\beta - \mu_\beta)) \\ & [\sigma^2] &\propto \sigma^{2(-a-1)} exp(-\frac{b}{\sigma^2}) \end{split}$$

$$J \propto L(\beta, y, X)[\beta | \sigma^2][\sigma^2]$$

$$J \propto (\sigma^2)^{-(\frac{n+k+1}{2}+a+1)} exp(-\frac{1}{2\sigma^2}(y^Ty+\tilde{\beta^T}M\tilde{\beta}-\mu_{\beta}^TA\mu_{\beta})) exp(-\frac{1}{2\sigma^2}(\beta-\mu_{\beta})^TA(\beta-\mu_{\beta})) exp(-\frac{b}{\sigma^2})$$

Let  $A_2 = y^T y + \tilde{\beta^T} M \tilde{\beta} - \mu_{\beta}^T A \mu_{\beta}$ 

$$J \propto (\sigma^2)^{-(\frac{n+k+1}{2}+a+1)} exp(-\frac{1}{\sigma^2}(b+\frac{A_2}{2})) exp(-\frac{1}{2\sigma^2}(\beta-\mu_\beta)^T A(\beta-\mu_\beta))$$

$$\begin{split} [\sigma^2|y,X] &= \int_{\beta} [\beta,\sigma^2|y,X] \, d\beta \\ [\sigma^2|y,X] &\propto (\sigma^2)^{-(\frac{n+k+1}{2}+a+1)} exp(-\frac{1}{\sigma^2}(b+\frac{A_2}{2})) \int_{\beta} exp(-\frac{1}{2\sigma^2}(\beta-\mu_{\beta})^T A(\beta-\mu_{\beta}) \, d\beta \end{split}$$

Recall

$$\int_{\beta} (2\pi)^{-\frac{k+1}{2}} det((\frac{1}{\sigma^2}(M+X^TX))^{-1})^{-\frac{1}{2}} exp(-\frac{1}{2\sigma^2}(\beta-\mu_{\beta})^TA(\beta-\mu_{\beta})) \, d\beta = 1$$

then

$$I = \int_{\beta} exp(-\frac{1}{2\sigma^2}(\beta - \mu_{\beta})^T A(\beta - \mu_{\beta}) \, d\beta = (2\pi)^{\frac{k+1}{2}} det((\frac{1}{\sigma^2}(M + X^TX))^{-1})^{\frac{1}{2}})$$

We know that  $det(aA) = a^k det(A)$ , where A is a k by k matrix.

$$\begin{split} I &= (2\pi)^{\frac{k+1}{2}} (\sigma^2)^{\frac{k+1}{2}} det((M+X^TX)^{-1})^{\frac{1}{2}} \\ &\quad I \propto (\sigma^2)^{\frac{k+1}{2}} \\ & [\sigma^2|y,X] \propto (\sigma^2)^{-(\frac{n+k+1}{2}+a+1)} exp(-\frac{1}{\sigma^2}(b+\frac{A_2}{2}))(\sigma^2)^{\frac{k+1}{2}} \\ & [\sigma^2|y,X] \propto (\sigma^2)^{-(\frac{n}{2}+a)-1} exp(-\frac{1}{\sigma^2}(b+\frac{A_2}{2})) \end{split}$$

meaning

$$[\sigma^2|y,X] \sim \mathcal{IG}(a+\frac{n}{2},b+\frac{A_2}{2})$$

where  $a=1,\,b=1$  and  $A_2=y^Ty+\tilde{\beta^T}M\tilde{\beta}-\mu_{\beta}^T(M+X^TX)\mu_{\beta}$ 

Table 1: First six rows of  $\sigma^2$  sample values

$\sigma'$	2
1.96835	7
1.695290	)
1.67422	1
1.98214	5
1.534436	ô
1.750102	2

## Histogram of $\sigma^2$

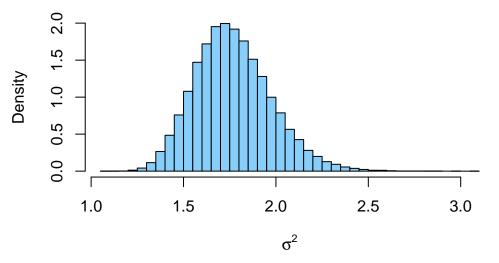
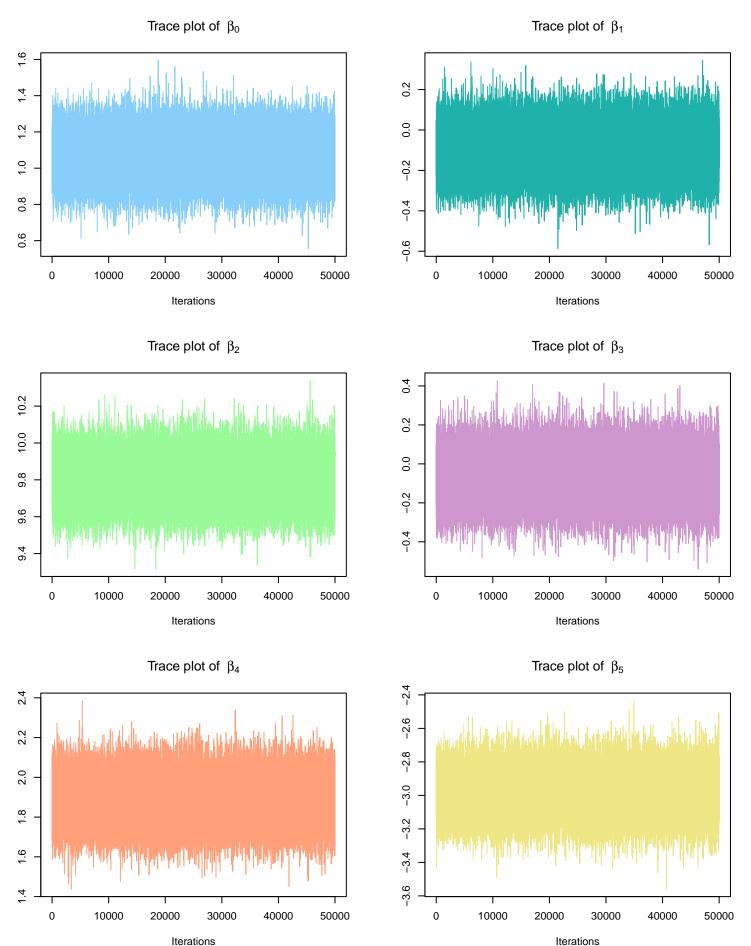


Table 2: First six rows of the  $\beta$  sample values

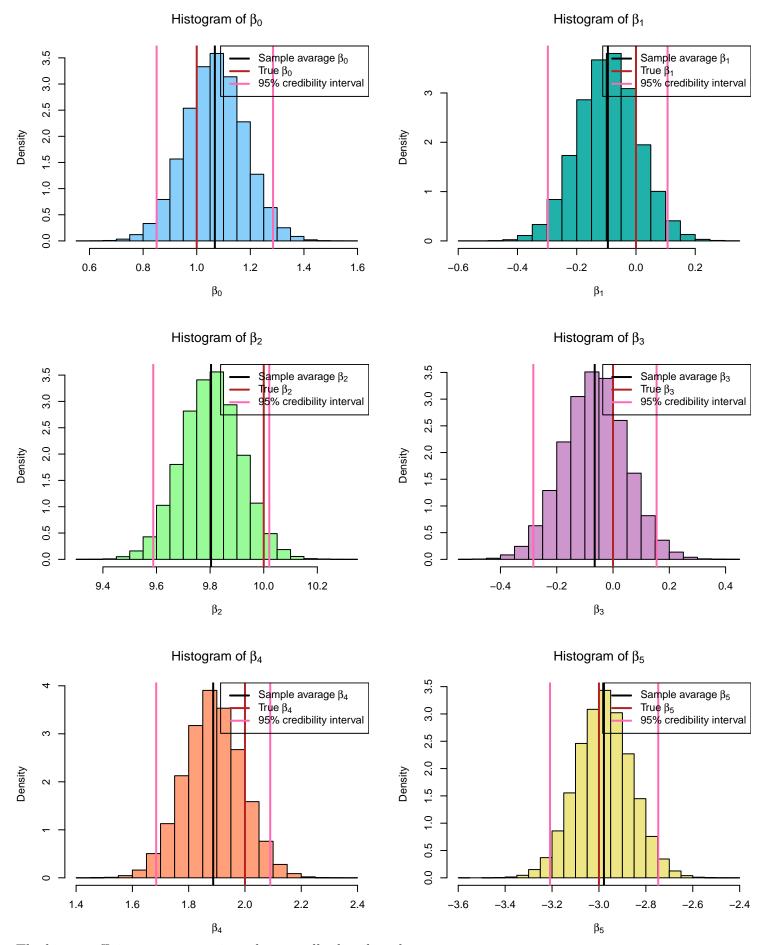
$eta_0$	$eta_1$	$eta_2$	$eta_3$	$eta_4$	$\beta_5$
0.9227484	-0.10623460	9.759792	0.10381837	2.051554	-2.839680
1.0567194	-0.20497249	9.788940	-0.09386065	1.772719	-2.809105
1.0122741	-0.04329813	9.801656	0.05378141	2.089081	-2.834065
1.2013402	-0.15871660	9.956266	-0.15968348	1.990569	-2.957680
1.1146270	-0.13473328	9.724995	0.06301181	1.879163	-3.124324
0.9180509	-0.14602992	10.011397	0.08100581	1.804324	-2.854732
1.0698329	0.01746358	9.703786	-0.08617031	1.880744	-2.950838







All of the trace plots appear as random scatter, indicating stationarity. This provides evidence for the convergence of the Markov Chains.



The beta coefficients are approximately normally distributed.

Confidence interval are the relative frequencies of stating valid bounds if you were to re-sample the data. Credibility intervals are the probability that the true parameter is within the stated bounds.

