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PREDICTING THE STOCK MARKET INDEX USING STOCHASTIC TIME SERIES ARIMA MODELLING: THE SAMPLE OF BSE AND NSE

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ABSTRACT

Stock market is basically volatile and the prediction of its movement will be more useful to the stock traders to design their trading strategies. An intelligent forecasting will certainly abet to yield significant profits. Many important models have been proposed in the economics and finance literature for improving the prediction accuracy and this task has been carried out through the modelling based on time series analysis. The main aim of this paper is to check the stationarity in time series data and predicting the direction of change in stock market index using the stochastic time series ARIMA modelling. The best fit ARIMA (0,1,0) model was chosen for forecasting the values of time series, viz., BSE_CLOSE and NSE_CLOSE by considering the smallest values of AIC, BIC, RMSE, MAE, MAPE, Standard Error of Regression, and the relatively high Adjusted R² values. Using this best fitted model, the predictions were made for the period ranging from 7th January, 2018 to 3rd June, 2018 (22 expected values) using the weekly data ranging from 6th January, 2014 to 31st December, 2017 (187 observed values). The results obtained from the study confirmed the prospective of ARIMA model to forecast the future time series in short-run and would assist the investing community in making the profitable investment decisions.

Key Words: BSE_CLOSE, NSE_CLOSE, ARIMA model, Forecasting, AIC, BIC, MAPE.

JEL Classification: G170, C53, C58, E37.

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INTRODUCTION

Stock market forecasting is an exercise to determine the future value of its performance index, viz., SENSEX, NIFTY. The successful prediction of any market's future index or a stock's future price will be more useful to the investing community to design optimal trading strategies and could yield significant profits. So, in the recent past, the concept of forecasting stock market and its return is gaining lot of attention by the researchers. It may be because of the fact that if the directions of change in the market movements are successfully predicted, the investors may be better guided. Sometimes, the forecasted trends of the market will help the policy makers and regulators of the stock market in making curative decisions. The profit making investments and day to day operations in capital market depends heavily on the forecasting ability.

Many practicing investors like Warren Buffett and other market researchers have proposed several models using various analytical methods, viz., fundamental analysis, technical analysis and analytical techniques, etc. to give more or less exact forecasting. In addition to the above methods of forecasting, some traditional time series models were also used for it. Mainly, there are two kinds of time series models for forecasting, viz., linear models and non-linear models. Some of the examples of linear models are moving average, exponential smoothing, time series regression, etc. One among the most common and popular linear models is the Autoregressive Integrated Moving Average (ARIMA) model proposed by Box and Jenkins (1976). In this paper, a modest attempt has been made to select the best fitted ARIMA model from different stochastic models that satisfies all the criteria of goodness of fit statistics for making the predictions and also to forecast the future values of stock market indices.

The remaining part of this paper is organized as follows: Section 2 contains the review of literature, statement of the problem, objectives of the study and research methodology. Section 3 deals with checking the stationarity in time series data using Augmented Dickey Fuller (ADF) Unit Root Test and by performing Correlogram analysis. Section 4 of the paper highlights the selection of best suitable model from different stochastic models that satisfies all the criteria of goodness of fit statistics for making the predictions. The experimental results of ARIMA (p,d,q) forecasting model are presented in Section 5. Finally, the conclusion, limitations and scope for future research are explained in section 6.

REVIEW OF LITERATURE

It is pertinent to review the accessible literature connected to the time series modeling and forecasting using ARIMA model. Most of the literature is focused on the identification of suitable ARIMA time series model and forecasting the gold price, exchange rates, oil palm prices, inflation rates, electricity consumption, etc. Only few studies are available relating to the forecasting of stock prices and stock market indices. Hence, this paper has been mainly devoted to the studies related to the determination of best ARIMA time series model and forecasting of future stock prices and stock market indices.

Meyler and Kenny (1998) have developed ARIMA time series predicting model for predicting the inflation in Ireland. In their study, they have focused on maximizing the power of forecasting by minimizing forecast errors. Contreras, Rodrigo, Francisco and Antonio (2003) have examined the trends in daily prices of electricity in spot and forward contracts for mainland Spain and Californian Markets and provided the best suited ARIMA method to predict next day electricity prices.

Rangson and Tidia (2006) have conducted a study with an objective to find an appropriate ARIMA model for forecasting three types of oil palm price by considering the minimum Mean Absolute Percentage Error. The empirical analysis of the study show that ARIMA (2,1,0), (1,0,1) and (3,0,0) are the best models for forecasting the farm price of oil palm, wholesale price of oil palm and pure oil price of oil palm respectively.

In a study conducted by Jarrett and Kyper (2011) using the data developed by Pacific-Basin Capital Markets (PACAP) and the SINOFIN Information Services Inc, has demonstrated the usefulness of ARIMA-Intervention time series analysis as both an analytical and forecast tool. The study indicates the usefulness of the developed model in explaining the rapid decline in the values of the price index of Shanghai market during the world economic decline in China in 2008. The authors have concluded that the daily stock price index contains an autoregressive component; hence, it is better to forecast the stock returns using ARIMA model.

Banerjee (2014) has used the ARIMA Model for predicting stock market indices and also highlighted that they have an undue influence on the progress of Indian economy. The study has dealt with the identification of the best fit ARIMA model and after that predicted the SENSEX using the justified model.

Adebiyi and Adewumi (2014) have presented the procedure for developing ARIMA models for forecasting share prices during short-run. The results of the study have explained the power of ARIMA models in predicting the stock prices in short-run, which would help the investors in their decisions. A study was conducted by Jadhav, Kakade, Utpat and Deshpande (2015) for forecasting the Indian Share Market using ARIMA model and said that artificial neural networks (ANNs) are universal approximates that can be applied to a wide range of time series for forecasting futuristic values in share market and give bright scope for investment. But, in their study, they have proposed a novel hybrid model of ANN using ARIMA model instead of only artificial neural network for improving the predictive performance.

Guha and Bandyopadhyay (2016) have examined the application of ARIMA time series model to forecast the future gold price based on the past data from November, 2003 to January, 2014 to mitigate the risk in purchase of gold and, hence, to give guidelines for the investor when to buy or sell the yellow metal. The authors have opined that now-a-days gold has gained importance as one of the investment alternatives; it has become necessary to predict the price of gold with an appropriate method.

Savadatti (2017) has carried out a study to identify the best fitted ARIMA models for forecasting the area, production, and productivity of food grains for 5 years. Based on univariate time series analysis, the study has identified ARIMA (2,1,2), ARIMA (4,1,0), and ARIMA (3,1,3) models for forecasting the data on area, production, and productivity of food grains, respectively and these models were found to be adequate. The forecast values indicated that production and productivity have increased during the forecast period but that of area exhibited near stagnancy, calling for timely measures to enhance the supply of food grains to meet the increasing demand in the future years.

Wadhawan and Singh (2019) have examined the different volatility estimators for forecasting volatility with high accuracy by traders, option practitioners, and various players of the stock market. The study evaluated the efficiency and bias of various volatility estimators based on various error measuring parameters, viz., ME, RMSE, MAE, MPE, MAPE, MASE and ACFI. The study has identified Parkinson estimator as the most efficient volatility estimator. The study has also suggested that the forecasted values were accurate based on the values of MAE and RMSE.

It may be concluded that, many researchers have conducted the studies to give reason for the selection of ARIMA model for forecasting the time series data of a single variable with better

accuracy. But, no researcher has focused on forecasting stock market indices in Indian context. The present work is an effort to forecast the indices of BSE and NSE based on the past 187 weeks using the best fitted ARIMA model.

STATEMENT OF THE PROBLEM

Because of dynamic and non-linear in nature, it is very tricky to predict the stock exchange movements precisely. But, it is necessary to forecast and uncover non-linearity of stock market; to enable the individual and institutional investors to design appropriate trading strategies and to achieve better results out of their investment endeavors. Hence, stock market forecasting has become a significant theme and motivated the researchers to build improved forecasting models. There are quite a few methods of statistical forecasting, viz., regression analysis, classical decomposition method, Box and Jenkins and smoothing techniques, with different degrees of accuracy. The accuracy of a forecasting model is based on the minimum errors of forecasting, viz., Root Mean Square Error, Mean Absolute Error, Standard Error of Regression, Adjusted R-square, Akaike Information Criterion, Bayesian Information Criterion, etc. Among several methods of time series forecasting, the Box and Jenkins method is quite accurate compared to other methods and may be applicable to all types of data movements. This paper is an attempt to test the stationarity in the given time series and selecting the best suitable ARIMA model (also known as Box-Jenkins Methodology) for short-term forecasting of BSE and NSE. The results obtained from the study could aid the investors in their investment decision-making process.

OBJECTIVES OF THE STUDY

The objectives of the study are listed below:

- 1. To test the stationarity of the time series data compiled for the study, i.e., weekly closing index values of BSE (BSE CLOSE) and NSE (NSE CLOSE).
- 2. To choose the optimum ARIMA model for estimating the series.
- 3. To forecast the indices of BSE and NSE using the selected time-series ARIMA model.

RESEARCH METHODOLOGY

Research Design:

Keeping in view of the above listed objectives of the study, an exploratory research design and stochastic modeling has been adopted. Exploratory research is one which interprets the already available information and it lays particular emphasis on the analysis and interpretation of the available secondary data. Stochastic modeling is used for selecting the best ARIMA model and forecasting the time series using the selected model.

Sources of Data:

The data required for the present study is secondary in nature and has compiled from an online source, viz., yahoofinance.com. The weekly closing indices of BSE and NSE are obtained from the website for the period from 6th June, 2014 to 3rd June, 2018 (209 observations). From this range of data, the researcher has taken the sample data ranging from 6th June, 2014 to 31st December, 2017 (187 observations) for making predictions ranging from 7th January, 2018 to 3rd June, 2018 (22 observations).

Hypothesis:

The null hypothesis is generally defined as the presence of a unit root and the alternative hypothesis is stationarity (or trend-stationary).

 H_{01} : $\delta = 1$, there is unit root and the series (BSE_CLOSE and NSE_CLOSE) is non stationary.

 H_{al} : $\delta < 1$, there is no unit root and the series (BSE_CLOSE and NSE_CLOSE) is stationary.

Methods for Analysis of Data:

To select the best fitted ARIMA model, among several experiments conducted, many statistical tools are to be applied, viz., Root Mean Square Error (RMSE), Mean Absolute Error (MAE), Mean Absolute Percentage Error (MAPE), Akaike Information Criterion (AIC), Bayesian Information Criterion (BIC), etc.

The RMSE has been used as a standard metric to measure the model performance in stock market forecasting. While applying the RMSE, the underlying assumption is that the errors are unbiased and follow a normal distribution. It provides a complete picture of the error distribution and its value should be relatively low (Chai and Draxler, 2014). The RMSE can be calculated by using the following formula:

RMSE =
$$\sqrt{\sum_{i=1}^{n} \frac{(x_i - \bar{x}_i)^2}{n}}$$
 -----(1)

Mean Absolute Error measures the average magnitude of the errors in a set of predictions, without considering their directions. It is the average over the test sample of the absolute differences between prediction and actual observations where all individual differences have equal weight. Hence, its value should be low. The MAE coefficient is given by the following equation (Chai and Draxler, 2014):

MAE =
$$\frac{1}{n} \sum_{i=1}^{n} \left| x_i - \overline{x_i} \right|$$
 ----- (2)

The mean absolute percentage error is a measure of prediction accuracy of a forecasting method. It usually expresses the forecasting accuracy of a model in percentage terms; hence, its value should be maximum. The MAPE formula as stated by Tofallis (2015):

MAPE =
$$\frac{100\%}{n} \sum_{i=1}^{n} \left| \frac{x_i - \overline{x_i}}{x_i} \right|$$
 -----(3)

Bayesian Information Criterion also known as Schwarz Information Criterion (SIC) is a criterion for model selection among a finite set of models. It is based on the likelihood function, and it is closely related to Akaike Information Criterion (AIC). Mathematically, the BIC is an asymptotic result derived under the assumption that the data series is exponentially distributed. The BIC was developed by Gideon Schwarz (1978), who gave a Bayesian argument for adopting it.

$$BIC = \log\left(\frac{rss}{n}\right) + \frac{k}{n}\log n - (4)$$

Where, "rss is residual sum of squares; k is the number of coefficients estimated, i.e., l + p + q + P + Q; and n is number of observations".

AUGMENTED DICKEY FULLER UNIT ROOT TEST

The initial phase of structuring ARIMA model is to recognize the variable being predicted is stationary in time series or not. Most forecasting methods assume that a distribution has stationarity. A time series has stationarity if a shift in time does not cause a change in the shape of the distribution, i.e., the mean and auto-covariance of the series do not depend on time (Tsay, 2005). Unit roots are one cause for non-stationarity. An absence of stationary can cause unexpected behaviors in data series. Most real-life data sets just are non-stationary and we should make it stationary in order to get any useful predictions from it. Augmented Dickey Fuller (ADF) Unit root test tests whether, a time series variable is non-stationary and possesses unit root. A common example of a non-stationary series is the random walk. We may write the Random Walk Model (RWM) with stochastic process as (Rao and Mukherjee, 1971), (Garekos and Gramacy 2013):

$$Y_t = \delta Y_{t-1} + u_t \quad (-1 \le \delta \le I) \quad ---- \quad (5)$$

Where t = time measured chronologically; and

 u_t = white noise error term.

For theoretical reasons, we manipulate equation - (5) by subtracting Y_{t-1} from both the sides to obtain -

$$Y_{t} - Y_{t-1} = \delta Y_{t-1} - Y_{t-1} + u_{t}$$

$$Y_t - Y_{t-1} = (\delta - 1)Y_{t-1} + u_t$$
 ------(6), which can be written as $\Delta Y_t = \beta Y_{t-1} + u_t$ -----(7)

Where $\beta = (\delta - I)$, and $\Delta =$ first difference operator.

In practice, instead of estimating equation - (5), we estimate equation - (7) and test the hypothesis (null) that $\beta = 0$. If $\beta = 0$, then $\delta = 1$, i.e., we have a unit root, meaning that the time series under consideration is non stationary. Before we proceed to estimate equation - (7), it may be noted that if $\delta = 0$, equation - (7) will become -

$$\Delta Y_t = (Y_t - Y_{t-1}) = u_t$$
 -----(8)

Since u_t is the white noise error term, it is stationary, which means that the first differences of a random walk time series are stationary.

Before running the ADF test, one should inspect the data to figure out an appropriate regression model. We have three versions of the test.

Type 0 No Constant, No Trend $\Delta Y_t = \beta_1 Y_{t-1} + u_t$

Type 1 Constant, No Trend $\Delta Y_t = \beta_0 + \beta_1 Y_{t-1} + u_t$

Type 2 Constant, Trend $\Delta Y_t = \beta_0 + \beta_1 Y_{t-1} + \beta_2 t + u_t$

The Augmented Dickey Fuller adds lagged differences to the above models (Damodar N. Gujarati, 2004):

Type 0 No Constant, No Trend
$$\Delta Y_t = \beta_1 Y_{t-1} + \sum_{i=1}^m \alpha_i \Delta Y_{t-i} + u_t$$

Type 1 Constant, No Trend
$$\Delta Y_t = \beta_0 + \beta_1 Y_{t-1} + \sum_{i=1}^m \alpha_i \Delta Y_{t-i} + u_t$$

Type 2 Constant, Trend
$$\Delta Y_t = \beta_0 + \beta_1 Y_{t-1} + \beta_2 t + \sum_{i=1}^m \alpha_i \Delta Y_{t-i} + u_t$$

Where $u_t = \text{Error term}$; and

 ΔY_{t-i} = Lagged differences.

Number of lagged differences would be added in the model is often decided numerically, so that the residuals are not serially correlated. Moreover, there are several options for choosing lags: Minimize Akaike's Information Criterion (AIC) or Bayesian Information Criterion (BIC), or drop lags until the last lag is statistically significant.

ADF test with intercept was applied on both series to test the data for stationarity. The null hypothesis is tested through 't-*Statistics*' which is given by the following formula:

$$t = \frac{\hat{\delta} - \delta_{H_0}}{SE \, of \, \hat{\delta}} - \dots \qquad (9)$$

If 't' calculated is greater than the critical value, we do not reject the null hypothesis and the series under consideration would be non-stationary and has a unit root. On the other hand, if 't' calculated is less than the critical value, we reject the null hypothesis and the series under consideration would be stationary and does not have the unit root. First, the series should be tested on level and if it does not become stationary, then we should test the series at the first and second difference sequentially. 'P' – value is also used to reject or accept the null hypothesis. If the 'P' – value is less than 0.05 (P<0.05), reject the null hypothesis and vice-versa.

Table No.1 Augmented Dickey-Fuller Test at Level

H ₀ is: BSE_CLOSE I	has a unit root.			-	H ₀ is: NSE_CLOSE has a unit root.					
Exogenous: Constant	t, Linear Trend.				Exogenous: Constant, Linear Trend.					
Lag Length:0 (Auton	natic-based on	AIC, maxlag=	:14		Lag Length:0 (Automatic-based on AIC, maxlag=14					
		t-statisti	c Pro	ob.*			t-s	tatistic	Prob.*	
ADF test statistic		-1.6036	68 0.7	7886	ADF test statistic		-1.	787570	0.7075	
Test critical values	1% leve	1 -4.0027	86		Test critical values	1%	level -4.	002786		
	5% leve	-3.4315	76			5%	level -3.	431576		
	10% leve	el -3.1394	75			109	6 level -3.	139475		
*MacKinnon (1996)	one-sided p-va	lues.			*MacKinnon (1996) o	one-sided p-valu	ies.			
Augmented Dickey -	Fuller Test Eq	uation			Augmented Dickey - 1	Fuller Test Equ	ation			
Dependent Variable:	D(BSE_CLOS	E)			Dependent Variable: l	D(NSE_CLOSE	Ε)			
Method: Least Squar	es				Method: Least Square	es				
Date:06/22/18 Time	e:05:51				Date:06/22/18 Time:07:28					
Sample (adjusted): 6/	/15/2014 6/0	3/2018			Sample (adjusted): 6/0	08/2014 6/03	/2018			
Included Observation	ns: 208 after ad	justments			Included Observation	s: 209 after adju	istments			
Variable	Coefficient	Std. Err.	t-stat	Prob.	Variable	Coefficient	Std. Err	t-stat	Prob.	
BSE_CLOSE(-1)	-0.028791	0.017953	-1.603	0.1103	NSE_CLOSE(-1)	-0.032916	0.018414	-1.787	0.0753	
С	713.4920	449.9443	1.585	0.1143	C	246.6873	137.8652	1.789	0.0750	
@TREND	1.538760	0.901803	1.706	0.0895	@TREND	0.540804	0.300068	1.802	0.0730	
("6/08/2014")					("6/08/2014")					
R-squared	0.015160	Mean depen	dent var	49.112	R-squared	0.017105	Mean depen	dent var	15.235	
Adj R-squared	0.005552	S.D. depend	ent var	520.18	Adj R-squared	0.007105	S.D. depend	ent var	158.44	
S.E. of regression	518.7404	AI Criterion		15.355	S.E. of regression	157.8433	AI Criterion		12.975	
Sum squared resid	55163775	Schwarz crit	erion	15.403	Sum squared resid	5132391	Schwarz crit	erion	13.023	
Log likelihood	g likelihood -1593.920 HQ Criterion 15.374			15.374	Log likelihood	-1352.922	HQ Criterio	n	12.994	
F-statistic	1.577813	Durbin-Wat	son stat	1.9243	F-statistic	1.792443	Durbin-Wat	son stat	1.9596	
Prob (F-statistic)	0.2018922				Prob (F-statistic)	0.169141				

Source: Authors calculation using eviews 10.

Table No.1, the ADF Test at level depicts that the calculated 't' – value is greater than critical values at 1%, 5% and 10% levels of significance. At levels both the underlying series (BSE_CLOSE and NSE_CLOSE) are non-stationary. The P-value of the series is also greater than 0.05. Hence, we do not reject the null hypothesis (H_{01}) and accept alternative hypothesis (H_{a1}) that

the series has a unit root. When the series (Y_t) is non-stationary, it must be differenced 'd' times before it becomes stationary, then it is said to be integrated of order 'd' (Brooks, 2008). The results of ADF test at first difference are presented in Table No.2.

Table No.2
Augmented Dickey-Fuller Test at First Difference

H ₀ is: DBSE_CLOSE h		i.			H ₀ is: D(DNSE_CLOSE) has a unit root.					
Exogenous: Constant, I	Linear Trend.				Exogenous: Constant, Linear Trend.					
Lag Length:2 (Automat	tic-based on A	AIC, maxlag=	=14		Lag Length:2 (Automatic-based on AIC, maxlag=14)					
		t-statist	ic Pro	b.*				t-statistic	Prob.*	
ADF test statistic		-9.7504	2 0.0	000	ADF test statistic			-9.796126	0.0000	
Test critical values	1% level	-4.0024	49		Test critical values	1	% level	-4.003226		
	5% leve	-3.4318	96			59	% level	-3.431789		
	10% leve	l -3.1396	64			10	% level	-3.139601		
*MacKinnon (1996) on	e-sided p-val	ues.			*MacKinnon (1996) on	e-sided p-va	lues.			
Augmented Dickey - Fu	ıller Test Equ	ıation			Augmented Dickey - Fu	ıller Test Eq	uation			
Dependent Variable: De	(BSE_CLOS	E)			Dependent Variable: De	(NSE_CLOS	SE)			
Method: Least Squares					Method: Least Squares					
Date:06/22/18 Time:1	14:05				Date:06/22/18 Time:14:03					
Sample (adjusted): 7/06		3/2018			Sample (adjusted): 6/29/2014 6/03/2018					
Included Observations:	205 after adj	ustments			Included Observations:	206 after ad	justments			
Variable	Coeffi	Std. Err.	t-stat	Prob.	Variable	Coeffi	Std. Err	t-stat	Prob.	
D(NSE_CLOSE(-1))	-1.188	1 0.0179	-9.7504	0.000	D(NSE_CLOSE(-1))	-1.2034	0.1228	-9.7961	0.0000	
D(NSE_CLOSE(-1),2)	0.201	5 0.0972	2.0725	0.039	D(NSE_CLOSE(-1),2)	0.1979	0.0985	2.0080	0.0460	
D(NSE_CLOSE(-2),2)			1.8452	0.066	D(NSE_CLOSE(-2),2)		0.0698	2.0705	0.0397	
C	-12493		-0.1680	0.866	C	-5.9844	22.739	0.2631	0.7927	
@TREND	0.637	8 0.6143	1.03829	0.300	@TREND	0.1204	0.1867	0.6450	0.5199	
("6/08/2014")					("6/01/2014")					
R-squared	0.50384	Mean depen	dent var	-3.149	R-squared	0.51073	Mean depe	endent var	0.3597	
Adj R-squared	0.49391	S.D. depend	lent var	729.24	Adj R-squared	0.50099	S.D. deper	ndent var	225.165	
S.E. of regression	518.778	AI Criterion	1	15.364	S.E. of regression	159.057	AI Criterio	on	13.008	
Sum squared resid	53426199	Schwarz cri	terion	15.445	Sum squared resid	5085154.	Schwarz c	riterion	13.081	
Log likelihood	-1569.904	HQ Criterio	n	15.397	Log likelihood -1334.039 HQ Criterion			ion	13.033	
F-statistic	50.77432	Durbin-Wat	son stat	1.9527	F-statistic	52.45409	Durbin-W	atson stat	1.9890	
Prob (F-statistic)	0.0000				Prob (F-statistic)	0.000000				

Source: Authors calculation using eviews 10.

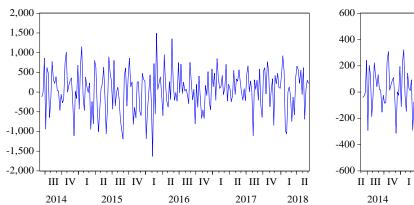
Table No.2, ADF test at first difference reveals that both the series are stationary at first difference. The calculated value of DBSE_CLOSE is -13.97590, which is less than the critical values at all levels of significance. Similarly, the 't' – statistics of DNSE_CLOSE is -14.28255, which is also less than the critical values at all levels of significance. Therefore, the null hypothesis (H_{01}) is rejected and can be concluded that both the series are stationary at 1%, 5% and 10% levels of significance, and both the series does not have the unit root. An informal method to test the stationarity also confirms the results of formal test, i.e., ADF test. Graphs of both the series at first difference do not demonstrate any kind of trend; there are fluctuations in the graphs. These fluctuations epitomize the stationary of underlying series (shown in Fig.1).

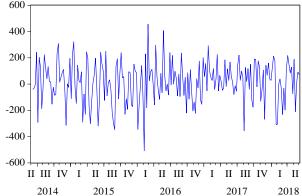
Fig.1

Graphs of BSE_CLOSE and NSE_CLOSE series at First Difference

DBSE_CLOSE

DNSE_CLOSE





CORRELOGRAM ANALYSIS

A Correlogram (also called Auto correlation function plot) is an image of correlation statistics and it gives a summary of correlation at different periods of time, i.e., serial correlation. Serial correlation is where an error at one point in time travels to a subsequent point of time. It is a commonly used tool for checking the randomness in a data set. A Correlogram contains Autocorrelation Function (ACF) and Partial Autocorrelation Function (PACF). Autocorrelation refers to the way the observations in a time series are related to other and is measured by a simple correlation between current observation (Y_t) and the observation 'p' periods ($lag\ p$) from the current one (Y_{t-p}) (Brooks, 2008) (Abdullah, 2012). The Autocorrelation coefficient at ' $lag\ p$ ' is given by --

__

$$r_p = \frac{c_p}{c_0}$$
 ----- (10)

where

 c_p = the auto-covariance function; and

 c_0 = the variance function.

$$c_p = \frac{1}{N} \sum_{t=1}^{N-p} (Y_t - \overline{Y}) * (Y_{t+p} - \overline{Y}) - \dots (11)$$
 and

$$c_0 = \frac{1}{N} \sum_{t=1}^{N} (Y_t - \overline{Y})^2$$
 ----- (12).

The resulting value of ' r_p ' will range between -1 and +1.

Partial Autocorrelations (PACF) are used to measure the degree of association between Y_t and Y_{t-p} when the effect of other time lags $1,2,3,\ldots,(p-1)$ are removed. The following figure No.2 represents the plot of Correlogram (ACF and PACF coefficients) of the time series BSE_CLOSE

and NSE_CLOSE for lags 1 to 20 at the level (zero order difference). We may infer from the Correlogram that, the ACF of BSE_CLOSE and NSE_CLOSE were dropped away very gradually; thus, the data in time series is non-stationary. Hence, there is a need to convert non-stationary series in to stationary by differencing.

Fig No.2
Correlogram of BSE_CLOSE and NSE_CLOSE

				ogra	шот	DOE_	_CLOSE and .						
Date:06	/22/18 Time	:15	:08				Date:06/2	22/18 Time: 15	5:13	3			
Sample:	6/08/2014 6/	03/	2018				Sample: 6/01/2014 6/03/2018						
_	d observations						_	observations: 2					
=====	a ooser varions	. 20	,,					OOSCI VALIOIIS. 2	10				
Autocorrelation	Partial Correlation		AC	PAC	Q-Stat	Prob	Autocorrelation	Partial Correlation		AC	PAC	Q-Stat	Prob
1		1	0.971	0.971	199.72	0.000	-		1	0.974	0.974	202.12	0.000
ı		2	0.941	-0.022	388.24	0.000		I (I	2	0.948	-0.011	394.61	0.000
ı		3	0.914	0.036	567.00	0.000		1 1	3	0.924	0.022	578.36	0.000
ı		4	0.892	0.076	738.22	0.000			4	0.904	0.055	754.89	0.000
ı	d .	5	0.866	-0.087	900.38	0.000		I (I	5	0.882	-0.029	923.93	0.000
ı		6	0.843	0.039	1054.6	0.000	-	1 1	6	0.860	-0.024	1085.3	0.000
ı		7	0.818	-0.022	1200.8	0.000	-	1 1	7	0.837	-0.014	1239.0	0.000
ı		8	0.796	0.008	1339.8	0.000	-	1 1	8	0.816	0.020	1385.9	0.000
ı		9	0.778	0.073	1473.2	0.000	· —		9	0.799	0.051	1527.2	0.000
ı		10	0.764	0.055	1602.5	0.000	· 🗀	1 1	10	0.783	0.033	1663.8	0.000
ı		11	0.751	0.025	1728.1	0.000	· 🗀	1 1	11	0.770	0.039	1796.3	0.000
ı	4	12	0.737	-0.011	1849.6	0.000	· —	1(1	12	0.756	-0.010	1924.7	0.000
1	i	13	0.719	-0.069	1965.9	0.000	· 🗀	[13	0.737	-0.087	2047.5	0.000
1	1 1	14	0.701	-0.012	2077.1	0.000	· 🗀	1 1	14	0.720	0.014	2165.2	0.000
1	[[d] ·	15	0.679	-0.083	2182.1	0.000	· 🗀	[15	0.699	-0.089	2276.8	0.000
1	1 1	16	0.658	-0.017	2280.9	0.000	· 🗀	1 1	16	0.679	-0.000	2382.7	0.000
1	<u> i</u> i i i i i i i i i i i i i i i i i i	17	0.637	0.019	2374.2	0.000	· —	1 1	17	0.660	-0.002	2483.1	0.000
1	j (i)	18	0.616	-0.030	2461.8	0.000	, —	<u> </u>	18	0.640	-0.010	2578.0	0.000
1	į di	19	0.593	-0.015	2543.5	0.000	· 🗀	j (j)	19	0.618	-0.031	2667.1	0.000
	į di	20	0.567	-0.092	2618.4	0.000		.	20	0.591	-0.127	2749.1	0.000

Figure No.3 shows the spikes of Correlogram, auto correlation and partial auto correlation coefficients for the lags 1 to 20 at the first order difference of the time series, viz., BSE_CLOSE and NSE_CLOSE.

Fig No.3
First Order Difference Correlogram of BSE_CLOSE and NSE_CLOSE

 Date:06/22/18 Time:22:58
 Date:06/22/18 Time:22:54

 Sample: 6/08/2014 6/03/2018
 Sample: 6/01/2014 6/03/2018

 Included observations: 208
 Included observations: 209

Autocorrelation	Partial Correlation	AC	PAC	Q-Stat	Prob	Autocorrelation	Partial Correlation	AC	PAC	Q-Stat	Prob
1 1	1 1	1 0.024	0.024	0.1238	0.725	1 1	1 1	1 0.004	0.004	0.0036	0.952
ıЩı	וַ יַּשוּי	2 -0.075			0.519	'¶ '	ļ ' Ū '	2 -0.052	-0.052	0.5802	-
. '	ļ ū !	3 -0.128	-0.125	4.7840	0.188	"	ļ إ	3 -0.143	-0.142		0.177
1 1	1 1	4 0.002	0.002	4.7851	0.310	1 1 1	' '	4 0.004	0.002	4.9336	0.294
1 1	111	5 0.020	0.001	4.8685	0.432	1 1 1		5 0.031	0.016	5.1390	0.399
1 þ 1	1 1	6 0.034	0.018	5.1165	0.529	1 1		6 0.019	-0.001	5.2138	0.517
1 0 1	1 1	7 -0.041	-0.040	5.4775	0.602	1 1	1 1 1	7 -0.002	0.002	5.2146	0.634
□ '	 	8 -0.135	-0.131	9.4871	0.303	. .	ļ □ ļ₁	8 -0.148	-0.143	9.9965	0.265
ı l ı	(1)	9 -0.037	-0.034	9.7861	0.368	I I		9 -0.040	-0.040	10.347	0.323
ı þ i	• -	10 0.055	0.029	10.464	0.401	h		10 0.048	0.035	10.862	0.368
, þ i		11 0.091	0.055	12.300	0.342	ı İ Dı		11 0.089	0.048	12.640	0.317
, þ .		12 0.079	0.078	13.684	0.321	, j jj	ipi	12 0.094	0.093	14.619	0.263
ų (1 11	13 -0.015	0.006	13.732	0.393	1 (1	4	13 -0.036	-0.013	14.908	0.313
1 þ 1		14 0.034	0.068	13.990	0.450	ı b ı	ibi	14 0.049	0.081	15.458	0.348
1 1 1	1 11	15 -0.045	-0.039	14.444	0.492	ı (ı	14 -	15 -0.052	-0.032	16.085	0.376
1 1 1	d -	16 -0.035	-0.053	14.728	0.545	10 1	d -	16 -0.031	-0.059	16.307	0.432
1 1	1 1	17 -0.002	-0.003	14.728	0.615	1(1		17 -0.016	-0.016	16.363	0.498
d ·	j d i	18 -0.100	-0.110	17.015	0.522	[[d]	j u j,	18 -0.084	-0.098	17.996	0.456
ı İ li	j , j o,	19 0.061	0.086	17.888	0.530	ibi	i i i i i i i i i i i i i i i i i i i	19 0.075	0.090	19.317	0.437
ı d ı	j ndj.	20 -0.068	-0.069	18.975	0.523	[[i]	j (d)	20 -0.088	-0.076	21.142	0.389

The plots say that the first order difference of the data after transformation is random. If the model is fit, then the residuals of the model would contain the sequence of probable errors. Since spikes of ACFs and PACFs are insignificant, the residuals of the chosen ARIMA model are white noise, and, hence the time series data has become stationary. This is essentially a random walk process and there is no need to think about any other AR(p) and MA(q) models further. Hence, the transformed time series essentially follow an ARIMA(0,1,0) process. The random walk model in stock price and market index forecasting has been commonly used and studied throughout history (Fama., 1965). The random walk model has similar implications as the efficient market hypothesis suggest that one cannot outperform the market by analyzing historical prices of a certain stock or index of the overall market.

ARIMA MODEL FOR FORECASTING

ARIMA model is the composition of series of steps for discovering the best model, supposing and identifying the different (ARIMA) models using available data in time series and forecasting the series using the best model. It is one of the well-known techniques for economic forecasting. ARIMA models are extremely capable to produce projections during short-term (Merh, Saxena and Pardasani, 2010). These are the best composite structural models, useful for short-term forecasts (Pai and Shenglin, 2005). In ARIMA model, the expected value of any variable is a "linear combination of past values and errors" (Hanke and Wichern, 2005), expressed as follows:

Auto Regressive Model [AR(p)]

An AR model is one in which ' Y_t '' depends only on its own past values, viz., Y_{t-1} , Y_{t-2} , Y_{t-3} , etc. Thus, $Y_t = f(Y_{t-1}, Y_{t-2}, Y_{t-3}, \dots, \varepsilon_t)$. ------(13)

A common representation of an autoregressive model where it depends on 'p' of its past values called AR(p) model is represented below:

$$Y_{t} = \beta_{0} + \beta_{1}Y_{t-1} + \beta_{2}Y_{t-2} + \beta_{3}Y_{t-3} + \dots + \beta_{p}Y_{t-p} + \varepsilon_{t} - \dots$$
 (14)

Where Y_t = affecting (dependent) variable at time t.

 Y_{t-1} , Y_{t-2} ,, Y_{t-p} = Response variable at time lags t-1, t-2,, t-p, respectively.

 β_0 , β_1 , β_2 ,, β_p = Coefficients to be estimated.

 ε_t = Error term at time t.

Moving Average Model [MA(q)]

A moving average model is one when Y_t depends only on the random error terms which follow a white noise process, i.e., $Y_t = f(\varepsilon_t, \varepsilon_{t-1}, \varepsilon_{t-2}, \varepsilon_{t-3}, \dots)$ ------ (15)

A common representation of a moving average model where it depends on 'q' of its past values is called MA(q) model and is represented below:

$$Y_t = \beta_0 + \varepsilon_t + \phi_1 \varepsilon_{t-1} + \phi_2 \varepsilon_{t-2} + \phi_3 \varepsilon_{t-3} + \dots + \phi_a \varepsilon_{t-a} - \dots$$
 (16)

The error terms ε_t is assumed to be white noise processes with mean zero and variance σ^2 .

Where Y_t = Response variable (dependent) variable at time t.

 β_0 , = Constant mean of the process.

 ϕ_1 , ϕ_2 , ϕ_3 ,...., ϕ_q = coefficients to be estimated.

 ε_t = Error term at time t.

 ε_{t-1} , ε_{t-2} , ε_{t-3} ε_{t-q} = Errors in previous time periods that are incorporated in Y_t .

Auto Regressive Moving Average (ARMA) Model

There are situations where the time series may be represented as a mix of both AR and MA models referred to as ARMA(p,q). The general form of such a time-series model, which depends on 'p' of its own past value and 'q' past values of white noise disturbances, taken the form:

$$Y_{t} = \beta_{0} + \beta_{1}Y_{t-1} + \beta_{2}Y_{t-2} + \dots + \beta_{p}Y_{t-p} + \varepsilon_{t} + \phi_{1}\varepsilon_{t-1} + \phi_{2}\varepsilon_{t-2} + \dots + \phi_{q}\varepsilon Y_{t-q} - \dots$$
 (17)

Selection of Appropriate ARIMA (p,d,q) Model

Model for non-seasonal series is called Autoregressive Integrated Moving Average Model, denoted by ARIMA(p,d,q). Here 'p' the order of autoregressive part, 'd' indicates the order of differencing, and 'q' indicates the order of moving average part. In general a series which is

stationary after being differenced 'd times' is said to be integrated of order 'd', denoted by I(d). If the original series is stationary, d=0 and the ARIMA models reduce to ARMA models. The time series data used for the present study, i.e., BSE_CLOSE and NSE_CLOSE has become stationary after the first order differencing. Since, there is no need for further differencing the series, it is necessary to adopt d=1 (first difference) for ARIMA (p,d,q) model. To get the appropriate numbers for 'p' (in AR) and 'q' (in MA) in the model, we should check the Correlogram after first difference in time series (Figure.2). Since there are no significant spikes of ACF and PACF, the residuals of the selected ARIMA model are white noise and there is no need for further consideration of one more AR(p) and MA(q). To choose one best ARIMA model amongst a numerous combinations present, the following criterions are used.

- a. Comparatively low of Akaike/Bayesian/Schwarz Information Criterions (AIC/BIC/SIC).
- b. Comparatively low S.E. of Regression.
- c. Comparatively high adjusted R-square (R²).
- d. Root Mean Square Error (RMSE) should be relatively low.
- e. Mean Absolute Error (MAE) and Mean Absolute Percentage Error (MAPE) should be low.

Table No.3 & 4 provides the results of various parameters of AR(p) and MA(q) of the ARIMA model. Using these values, the best fit model for predicting the time series DBSE_CLOSE and DNSE_CLOSE are identified.

Table No.3
Output for various ARIMA Parameters for DBSE_CLOSE

ARIMA	RMSE	MAE	MAPE	S.E. of Regression	Log Likelihood	Adjusted R ²	AIC	BIC
(0,1,0)	518.9344	407.0126	135.6836	520.1864	-1595.509	0.00000	15.3510	15.36709
(1,1,0)	520.0480	408.1295	135.3038	522.5642	-1595.448	-0.00916	15.3696	15.41783
(1,1,1)	520.0340	408.1138	135.1785	523.6914	-1595.388	-0.01352	15.3787	15.44291
(2,1,0)	521.2446	409.8329	131.5123	521.2495	-1594.929	-0.00409	15.3647	15.41284
(2,1,1)	521.2519	409.8372	131.5177	522.4525	-1594.900	-0.00873	15.3740	15.43822
(1,1,2)	519.9580	408.0805	135.3919	522.4621	-1594.904	-0.00876	15.3740	15.43826
(2,1,2)	521.2424	409.8305	131.5162	522.5242	-1594.929	-0.00900	15.3743	15.43850

Source: Authors calculation using eviews 10.

The values in first row represent the best ARIMA model among different combinations.

Table No.4
Output for various ARIMA Parameters for DNSE_CLOSE

ARIMA	RMSE	MAE	MAPE	S.E. of Regression	Log	Adjusted R ²	AIC	BIC

					Likelihood			
(0,1,0)	158.0640	125.1574	195.6765	158.4436	-1354.725	0.00000	12.9734	12.9894
(1,1,0)	158.3947	125.4862	195.9553	159.2095	-1354.723	-0.00969	12.9925	13.0405
(1,1,1)	158.3945	125.4854	195.9649	159.5926	-1354.717	-0.01455	13.0028	13.0660
(2,1,0)	158.7299	125.8748	196.6628	158.9947	-1354.444	-0.00696	12.9898	13.0378
(2,1,1)	158.7298	125.8755	196.6935	159.3812	-1354.443	-0.01187	12.9994	13.0634
(1,1,2)	158.3971	125.5027	196.4437	159.3827	-1354.445	-0.01189	12.9994	13.0634
(2,1,2)	158.7296	125.8748	196.6631	159.3818	-1354.444	-0.01187	12.9994	13.0634

Source: Authors calculation using eviews 10.

The values in first row represent the best ARIMA model among different combinations.

After checking the robustness of the statistics given in the above Table No. 3 & 4, it is found that only ARIMA(0,1,0) model convinces all the norms (lowest AIC, BIC, RMSE, MAE, MAPE, Standard Error of Regression, and the relatively high Adjusted R² values), hence this model is considered to be the best predictive model, which is used to forecast the future values of the time series, viz., BSE_CLOSE and NSE_CLOSE. The prediction equation for this model can be written as:

$$Y_t - Y_{t-1} = \mu_t$$
 or equivalently $Y_t = Y_{t-1} + \mu_t$ ----- (18).

..... where the constant term is the average period-to-period change (i.e., the long-term drift) in Y. This model could be fitted as a no-intercept regression model in which the first difference of Y is the dependent variable.

FORECASTING USING SELECTED ARIMA (p, d, q) MODEL

The present study is based on weekly data on the closing indices of BSE (BSE_CLOSE) and NSE (NSE_CLOSE) covering the period from 08th June 2014 to 03rd June 2018, having a total number of 209 observations. Of which, the period from 07th January 2018 to 03rd June 2018, having 22 observations are used for forecasting length.

Result of ARIMA (0,1,0) Model for BSE_CLOSE Prediction

Table No.5 exhibits the forecasting results of ARIMA(0,1,0) model, which is regarded as the best fit model for prediction of BSE_CLOSE index. The table shows the actual and predicted values of series for the forecast length (22 observations) ranging from 7th January, 2018 to 3rd June, 2018. It is observed from the summary of ARIMA forecasting model that the software has selected the log of dependent variable after first differencing, viz., DLOG(BSE_CLOSE) and the forecast length is 22 weeks. The software has estimated 9 models and out them the best ARMA model selected is

(0,0)(0,0). The value of Akaike Information Criterion of this model is observed very smaller than all other models tested.

Fig.4 is the graphical illustration and shows the level of accuracy of the selected ARIMA model, exhibits the predicted performance of the BSE (BSE_CLOSE) against the actual performance during the forecasted period. The line of forecasting of BSE_CLOSE is continued to rise during the forecasting period, i.e., from 7th January, 2018 to 3rd June, 2018. When compared to the forecasted performance, the actual performance of BSE_CLOSE during the period from 4th February, 2018 to 18th March, 2018 is quite unsatisfactory. However, the market has revived by the end of 3rd June, 2018.

Table No.5
Sample Empirical Results of ARIMA (0,1,0) of BSE CLOSE

	Results of ARIMA (0,1,0) of DSE_CLOSE								
Sample	Actual	Predicted							
Period	Values	Values	Summary of the ARIMA Forecasting Model						
31 st Dec, 17	34153.85	34153.85	Automatic ARIMA Forecasting						
7 th Jan, 18	34592.39	34209.52	Selected dependent variable: DLOG(BSE_CLOSE)						
14 th Jan, 18	35511.57	34265.28	Date: 07/01/18 Time: 21:56						
21 st Jan, 18	36050.44	34321.13	Included observations: 186						
28 th Jan, 18	35066.75	34377.07	Forecast length: 22						
4 th Feb, 18	34005.76	34433.10	Number of estimated ARMA models: 9						
11 th Feb, 18	34010.76	34489.22	Number of non-converged estimations: 0						
18 th Feb, 18	34142.14	34545.43	Selected ARMA model: (0,0) (0,0)						
25 th Feb,18	34046.94	34601.74	AIC value: -5.05495260101						
4 th Mar, 18	33307.14	34658.14	Fig.4: Actual and Forecast Graph of ARIMA (0,1,0) model						
11 th Mar, 18	33176.00	34714.63	Actual and Forecast						
18 th Mar, 18	32596.53	34771.21	36,500						
25 th Mar, 18	32968.67	34827.88	36,000 -						
1 st Apr, 18	33626.96	34884.65	35,500 -						
8 th Apr, 18	34192.64	34941.50	35,000 -						
15 th Apr, 18	34415.57	34998.46	34,500 -						
22 nd Apr, 18	34969.69	35055.50	34,000 -						
29 th Apr, 18	34915.37	35112.64	33,500 -						
6 th May, 18	35535.78	35169.87	33,000 -						
13 th May, 18	34848.30	35227.19							
20 th May, 18	34924.87	35284.61	5 19 3 17 31 14 28 11 25 11 25 8 22 6 20 3						
27 th May, 18	35227.26	35342.12	M11 M12 M1 M2 M3 M4 M5						
3 rd June, 18	35443.67	35339.72	Forecast Actual						

Source: Authors calculation using eviews 10.

Table No.6 ARIMA (0,1,0) estimation output with DLOG(BSE_CLOSE):

Dep. Variable: DLOG(BSE_CLOSE)

Method: Least Squares

Date: 07/01/18 Time:21:56

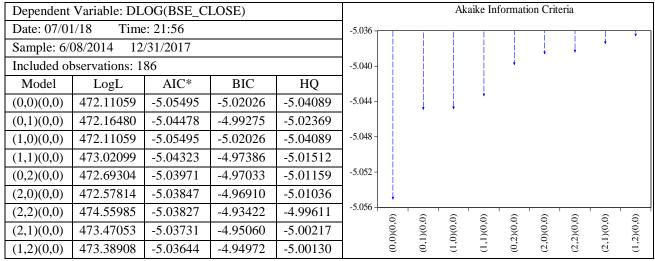
Sample (Adjusted): 6/15/2014 12/31/2017 Included observations: 186 after adjustments

Variable	Coefficient	Std. Error	t-Statistic	Prob.
С	0.001629	0.001406	1.158693	0.2481
R-squared	0.000000	Mean Depende	nt variable	0.00162
Adj. R-squared	0.000000	S.D. dependent	S.D. dependent variable	
S.E. of Regression	0.019169	Akaike info cri	terion	-5.06570
Sum squared residuals	0.067977	Schwarz criteri	on	-5.04836
Log likelihood	472.1106	Hannan-Quinn criterion		-5.05867
Durbib-Watson Stat	2.044222			

Source: Authors calculation using eviews 10.

According to table No.6, ARIMA (0,1,0) is relatively the best model. The model returned the smallest Akaike information criterion of -5.06570, smallest Bayesian or Schwarz information criterion of -5.04836 and relatively smallest standard error of regression of 0.019169. It is also observed from the model selection criteria table (table No.7) that out of 9 models verified, ARMA(0,0)(0,0) is found to be the best model as its LogL, AIC, BIC and HQ coefficients are smaller than the remaining 8 models.

Table No.7 Model Selection Criteria Table



Source: Authors calculation using eviews 10.

Result of ARIMA (0,1,0) Model for NSE_CLOSE Prediction

Table No.8 contains the empirical results of ARIMA(0,1,0), which is regarded as the best fit model for prediction of NSE_CLOSE index. The table shows the actual and predicted values of series for the forecast length (22 observations) ranging from 7th January, 2018 to 3rd June, 2018. It is observed from the summary of ARIMA forecasting model that the software has selected the log of dependent variable after first differencing, viz., DLOG(NSE_CLOSE) and the forecast length is 22 weeks. The software has estimated 9 models and out them the best ARMA model selected is

(0,0)(0,0). The value of Akaike Information Criterion of this model is observed very smaller than all other models tested.

Table No.8
Sample Experimental Results of ARIMA(0,1,0) of NSE_CLOSE

Sample	Actual	Predicted	THE RESULTS OF FIRM (U)(1)(0) OF TIBL_CLOSE
Period	Values	Values	Summary of the ARIMA Forecasting Model
31 st Dec 17	10558.85	10558.85	Automatic ARIMA Forecasting
7 th Jan 18	10681.25	10577.56	Selected dependent variable: DLOG(NSE_CLOSE)
14 th Jan 18	10894.70	10596.30	Date: 07/01/18 Time: 23:01
21st Jan 18	11069.65	10615.07	Included observations: 187
28 th Jan 18	10760.60	10633.87	Forecast length: 22
4 th Feb 18	10454.95	10652.71	Number of estimated ARMA models: 9
11 th Feb 18	10452.30	10671.59	Number of non-converged estimations: 0
18 th Feb 18	10491.05	10690.49	Selected ARMA model: (0,0)(0,0)
25 th Feb 18	10458.35	10709.43	AIC value: -5.04795355019
4 th Mar 18	10226.85	10728.41	Fig. 5: Actual and Forecast Graph of ARIMA (0,1,0) model
11 th Mar 18	10195.15	10747.41	Actual and Forecast
18 th Mar 18	9998.05	10766.45	11,200
25 th Mar 18	10113.70	10785.53	11,000 -
1 st Apr 18	10331.60	10804.64	10.800 -
8 th Apr 18	10480.60	10823.78	
15 th Apr 18	10564.05	10842.95	10,600
22 nd Apr 18	10692.30	10862.16	10,400 -
29 th Apr 18	10618.25	10881.41	10,200 -
6 th May 18	10806.50	10900.68	10,000 -
13 th May 18	10596.40	10920.00	9,800
20 th May18	10605.15	10939.34	5 19 3 17 31 14 28 11 25 11 25 8 22 6 20 3
27 th May18	10696.20	10958.72	M11 M12 M1 M2 M3 M4 M5
3 rd June 18	10767.65	10978.14	→ Forecast → Actual

Source: Authors calculation using eviews 10.

Fig.5 is the graphical illustration and shows the level of accuracy of the selected ARIMA model, exhibits the predicted performance of the NSE (NSE_CLOSE) against the actual performance during the forecasted period. The line of forecasting of NSE_CLOSE is continued to rise during the forecasting period, i.e., from 7th January, 2018 to 3rd June, 2018. When compared to the forecasted performance, the actual performance of NSE_CLOSE during the period from 4th February, 2018 to 18th March, 2018 is quite unsatisfactory. However, the market has revived by the end of 6th May, 2018 and almost all continued till 3rd June, 2018.

Table No.9
ARIMA (0,1,0) estimation output with DLOG(NSE_CLOSE)

Dependent Variable: DLOG(NSE_CLOSE)

Method: Least Squares

Date: 07/01/18 Time:23:13

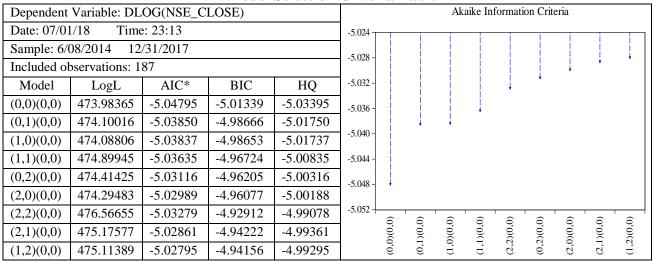
Sample (Adjusted): 6/08/2014 12/31/2017 Included observations: 187 after adjustments

Variable	Coefficient	Std. Error	t-Statistic	Prob.
С	0.001770	0.001407	1.258281	0.2099
R-squared	0.000000	Mean Dependent	t variable	0.00177
Adj. R-squared	0.000000	S.D. dependent v	0.01923	
S.E. of Regression	0.019237	Akaike info crite	rion	-5.05864
Sum squared residuals.	0.068830	Schwarz criterion	n	-5.04137
Log likelihood	473.9837	Hannan-Quinn c	riterion.	-5.05164
Durbin-Watson Stat	2.066195			

Source: Authors calculation using eviews 10.

According to table No.9, ARIMA (0,1,0) is relatively the best model. The model has given a very small AIC of -5.05864, small BIC/SIC of -5.04137 and comparatively negligible value of S.E. of regression of 0.019237. It is also observed from the model selection criteria table (table No.10) that out of 9 models verified, ARMA(0,0)(0,0) is found to be the best model as its LogL, AIC, BIC and HQ coefficients are smaller than the remaining 8 models.

Table No.10 Model Selection Criteria Table



Source: Authors calculation using eviews 10.

CONCLUSION

The main objective of this paper is to study the stationarity of the indices of BSE and NSE and to forecast using the ARIMA model. For this purpose, the weekly closing indices of BSE and NSE are obtained from the website yahoofinance.com for the period from 6th June, 2014 to 3rd June, 2018. The ADF Test is administered to check for the presence of unit root to confirm the stationarity of index series. The results of the test confirmed the presence of Unit root and showed non-stationary. The ADF test has confirmed that the given time series are stationary at first difference.

For the present work, ARIMA (0,1,0) model was chosen as the top model from nine different models because it gratify all the norms of goodness of fit statistics, as other eight models have not satisfied such criterians. This best candidate model was selected for making predictions of BSE_CLOSE and NSE_CLOSE for the period ranging from 7th January, 2018 to 3rd June, 2018 using the weekly data ranging from 6th January, 2014 to 31st December, 2017. The study also made a comparision between predicted and actual performance of BSE_CLOSE and NSE_CLOSE during the sample period. The results of the best fitted model highlights the strength of ARIMA model to forecast the BSE_CLOSE and NSE_CLOSE satisfactory on short-term basis and would guide the individuals to select gainful investment options.

RESEARCH IMPLICATIONS

The findings of the study have the following implications for investors, researchers and academic fraternity.

- The study has elucidated the procedure for testing the stationarity in time series data using Augmented Dicky Fuller Test and Correlogram analysis. As well, the study has enlightened the criterion and modus operandi for selection of the best ARIMA model and the methodology for forecasting BSE_CLOSE and NSE_CLOSE. This will aid the researchers and academicians to carryout further research.
- 2. The forecasting of market indices (BSE_CLOSE and NSE_CLOSE) and comparision of forecast and actual performance will assist the investors to know the market trends, risk analysis and to take investment decisions.

LIMITATIONS

The ARIMA model has few constraints regarding the exactness of forecasting because of its wide usage for short-run predicting the values in the time series to notice the minor variations in the data. In case of erratic variations in the data set (too large variations) due to change in government policies or the structural breaks in economy (economic instability), etc, it turns out to be intricate to capture the accurate trend. Hence, this model turns out to be useless to predict long-run changes. Moreover, the forecasting using the ARIMA model would depend upon the hypothesis of linearity in historical data, however, there is no confirmation that BSE_CLOSE or NSE_CLOSE are linear in nature.

SCOPE FOR FURTHER RESEARCH

Forecasting of BSE_CLOSE and NSE_CLOSE using ARIMA model was made with the fundamental supposition that the given series follow a absolutely linear model. Non linear prediction methods using latest softwares may also be considered with less error (white noise) term. Further, the study may be extended to multivariate time series forecasting, i.e., predicting a dependent variable using more than one independent variables. In future, similar studies may be conducted for forecasting various economic variables, viz., gold and silver prices, curency exchange rates, individual stock prices, production from agriculture and industry, electricity consumption, export performance of various industries, etc.

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