# Dis 6: BLUE; Homoskedasticity vs. Heteroskedasticity; Comparing Two Populations

#### 1 BLUE (Gauss-Markov theorem)

• BLUE is a feature of OLS (ordinary least squares) regression technique.

**Theorem 1** (Gauss-Markov). OLS estimator of  $\beta_i$ ,  $\hat{\beta}_i$ , is the **B**est (most efficient) Linear conditionally Unbiased Estimator (**BLUE**), as long as the following OLS assumptions holds:

- 1. Zero conditional mean:  $E[u_i|x_1,x_2,\ldots,x_k]=0$
- 2. **I.I.D. Data**:  $(x_{1i}, x_{2i}, \dots, x_{ki}, y_i)$  are i.i.d. (independent and identically distributed).
- 3. **Large outliers are unlikely**: There doesn't exist some  $(x_{1i}, x_{2i}, \dots, x_{ki}, y_i)$  that live in a dramatically different region. Could be measured as the fourth moment of each variable is finite (i.e.  $0 < E[x_{1i}^4] < \infty, 0 < E[x_{2i}^4] < \infty, \dots, 0 < E[x_{ki}^4] < \infty$ , and  $0 < E[y_i^4] < \infty$ )
- 4. **No perfect multicollinearity**: One of the regressors cannot be a perfect linear function of some other regressors.
- 5. **Homoskedasticity**:  $Var(u_i|x_1, x_2, ..., x_k) = \sigma^2$  is a constant.
- **Best** and **unbiased** are the two most important features:
  - Best is in the sense that the estimators achieved under OLS have the smallest standard errors
    compared with all other potential estimators. When the estimator has the smallest standard
    error possible, the estimator is called the most efficient.
  - **Unbiased** means that on average, estimators yield the true value:  $E(\hat{\beta}_i) = \beta_i$ .

## 2 Homoskedasticity vs. Heteroskedasticity

- Homoskedasticity and heteroskedasticity are features of **conditional variance of error term**  $u_i$  **given information on**  $x_i$ 
  - When  $Var(u_i|x_{1i},x_{2i},\ldots,x_{ki})=\sigma^2$  (a constant), the error term  $u_i$  is **homoskedastic**
  - When  $Var(u_i|x_{1i},x_{2i},...,x_{ki}) \neq \sigma^2$ , the error term  $u_i$  is **heteroskedastic** 
    - \* **Impure heteroskedasticity**:  $Var(u_i|x_{1i},x_{2i},...,x_{ki}) \neq \sigma^2$  due to model misspecification (ex. omitted variable bias)
    - \* **Pure heteroskedasticity**:  $Var(u_i|x_{1i},x_{2i},...,x_{ki}) \neq \sigma^2$  arises even when the model is correctly specified.

For the rest of this handout, we only consider pure heteroskedasticity.

- How do they reflect in data?
  - Consider linear model  $y_i = \beta_0 + \beta_1 x_i + u_i$
  - Conditional variance of  $y_i$  given  $x_i$  looks like the following:

$$Var(y_i|x_i) = Var(\beta_0 + \beta_1 x_i + u_i|x_i)$$

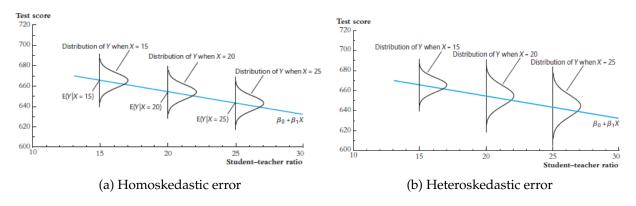
$$= Var(\beta_0) + Var(\beta_1 x_i | x_i) + Var(u_i | x_i)$$
 (by i.i.d. data)  
$$= 0 + 0 + Var(u_i | x_i) = Var(u_i | x_i)$$

So  $Var(u_i|x_i)$  is directly reflected onto  $Var(y_i|x_i)$ !

<u>Ex.</u> Say that we are interested in studying the relationship between test score and student-teacher ratio. A simple (univariate) linear regression model is proposed:

test score<sub>i</sub> = 
$$\beta_0 + \beta_1$$
student-teacher ratio<sub>i</sub> +  $u_i$ 

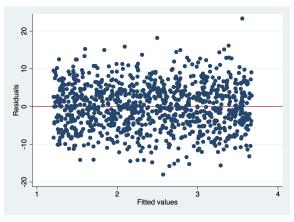
Each type of error term reflects in data in the following way:

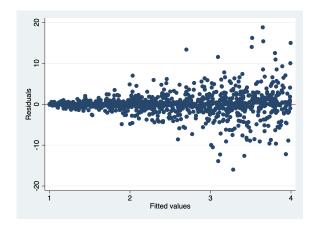


- What difference does these two types of error specification make?
  - OLS estimate of  $\hat{\beta}_i$  is BLUE only under homoskedasticity; it's not under heteroskedasticity. Homoskedasticity simplifies the error structure  $\rightarrow$  *in general*, standard error of the  $\hat{\beta}$  estimator is smaller under homoskedasticity.
    - This means that if we specify the error term to be heteroskedastic, then the **best** part of BLUE is violated under heteroskedasticity.
  - As long as model is correctly specified (either error is homoskedastic or pure heteroskedastic), then  $\hat{\beta}_i$  yields unbiased estimate of the true  $\beta_i$ .
    - Neither error specification affects the  $\hat{\beta}_i$  point estimate though the unbiased part of BLUE has nothing to do with how standard error of  $\hat{\beta}_i$  looks like.
  - In many contexts, heteroskedasticity is the more accurate way to model the error structure.
- How to test whether heteroskedasticity is the way we should model the error term?
  - Visually: Plot residual  $\hat{u}$  (sample analog of error) against predicted value  $\hat{y}$ . Doing this in Stata:

```
reg y x1 x2 x3 // run your regression first rvfplot, yline(0) // the yline(0) option adds a horizontal line at y = 0
```

Each type of error term has the following regress postestimation diagnostic plot:





(a) Homoskedastic error

(b) Heteroskedastic error

## – Analytically: Breusch-Pagan or White test

Test	Stata Command and Output				
Breusch-Pagan	reg y x1 x2 x3 // run your regression first estat hettest, rhs fstat				
	. quietly reg y x				
	. estat hettest, rhs fstat				
	Breusch-Pagan / Cook-Weisberg test for heteroskedasticity Ho: Constant variance Variables: x  F(1, 998) = 2.26 Prob > F = 0.1333				
	reg y x1 x2 x3 // run your regression first estat imtest, white				
	. quietly reg y x				
	. estat imtest, white				
	White's test for Ho: homoskedasticity against Ha: unrestricted heteroskedasticity				
	against Ha: unrestricted heteroskedasticity				
White	against Ha: unrestricted heteroskedasticity  chi2(2) = 2.33  Prob > chi2 = 0.3121				
White	chi2(2) = 2.33				
White	chi2(2) = 2.33 Prob > chi2 = 0.3121				
White	chi2(2) = 2.33 Prob > chi2 = 0.3121  Cameron & Trivedi's decomposition of IM-test				
White	chi2(2) = 2.33 Prob > chi2 = 0.3121  Cameron & Trivedi's decomposition of IM-test  Source chi2 df p				

- How to incorporate homoskedasticity and heteroskedasticity in regression model estimation?
  - Homoskedasticity and heteroskedasticity only affect the standard error of  $\hat{\beta}$  estimators
  - Calculating by hand:
    - \* Under homoskedasticity:

$$Var(\hat{\beta}_1|x_1, x_2, \dots, x_k) = \frac{1}{n} \times \frac{\frac{1}{n-2} \sum_{i=1}^n \hat{u}_i^2}{\frac{1}{n} \sum_{i=1}^n (x_{1i} - \bar{x}_1)^2}$$

\* Under heteroskedasticity:

$$Var(\hat{\beta}_1|x_1, x_2, \dots, x_k) = \frac{1}{n} \times \frac{\frac{1}{n-2} \sum_{i=1}^n (x_{1i} - \bar{x}_1)^2 \hat{u}_i^2}{\left[\frac{1}{n} \sum_{i=1}^n (x_{1i} - \bar{x}_1)^2\right]^2}$$

Homoskedasticity can be viewed as a special case of heteroskedasticity:  $\hat{u}_i$  varies constantly across all i under homoskedasticity.

Squared root of the variance under heteroskedasticity is called **robust standard error**.

- Using Stata:
  - \* Under homoskedasticity: same as what we've been doing (homoskedasticity is the default for linear regression model)

. use "http://	fmwww.bc.edu/	ec-p/data/	wooldridge/	wage2.d	ta", clea	r	
. reg wage edu	uc exper hours						
Source	ss	df	MS	Numbe	er of obs	=	935
				- F(3,	931)	=	49.31
Model	20939351.2	3	6979783.75	Prob	> F	=	0.0000
Residual	131776817	931	141543.305	R-squ	uared	=	0.1371
				- Adj F	R-squared	=	0.1343
Total	152716168	934	163507.675	Root	MSE	=	376.22
wage	Coef.	Std. Err.	t	P> t	[95% 0	Conf.	Interval]
educ	76.7356	6.31113	12.16	0.000	64.349	91	89.12129
exper	17.55185	3.162026	5.55	0.000	11.346	32	23.75737
hours	-1.995166	1.7116	-1.17	0.244	-5.3542	206	1.363875
_cons	-190.8808	128.0893	-1.49	0.137	-442.2	258	60.49628

\* Under heteroskedasticity: add robust option reg y x1 x2 x3 x4, robust

Linear regression				Number of	obs	=	935
				F(3, 931)		=	43.29
				Prob > F		=	0.0000
				R-squared	l	=	0.1371
				Root MSE		=	376.22
wage	Coef.	Robust Std. Err.	t	P> t	[95%	Conf.	Interval]
educ	76.7356	6.74783	11.37	0.000	63.49	288	89.97832
exper	17.55185	3.121662	5.62	0.000	11.42	2554	23.67816
hours	-1.995166	2.275691	-0.88	0.381	-6.461	243	2.470912

## 3 Comparing two populations: Econ 310 vs. Econ 400 regression technique

• Consider two different populations. In the context of PS 5 Question 1, let the two populations be female and male.

We'd love to know something about the female and male population, so suppose that we constructed a representative sample, which contains two variables,  $x_{\text{female}}$  and  $x_{\text{male}}$ , that record income of female and male sampled from the respective population.

#### [Correction begins here]

- If the two populations have equal variances:

$$t = \frac{(\overline{x}_{\text{female}} - \overline{x}_{\text{male}}) - (\mu_{\text{female},H_0} - \mu_{\text{male},H_0})}{\sqrt{\frac{s_{\text{pooled}}^2}{n_{\text{female}}} + \frac{s_{\text{pooled}}^2}{n_{\text{male}}}}} \quad \sim \quad t_{n_{\text{female}} + n_{\text{male}} - 2}$$

where

$$s_{\text{pooled}}^2 = \frac{\sum_{i}^{n_{\text{female}}} (x_{\text{female},i} - \overline{x_{\text{female}}})^2 + \sum_{i}^{n_{\text{male}}} (x_{\text{male},i} - \overline{x_{\text{male}}})^2}{n_{\text{female}} + n_{\text{male}} - 2}$$

- If the two populations have unequal variances:

$$t = \frac{(\overline{x_{\rm female}} - \overline{x_{\rm male}}) - (\mu_{\rm female}, H_0 - \mu_{\rm male}, H_0)}{\sqrt{\frac{s_{\rm female}^2}{n_{\rm female}} + \frac{s_{\rm male}^2}{n_{\rm male}}}} \quad \sim \quad t_{DOF}$$

where

$$DOF = \frac{(s_{\text{female}}^2 / n_{\text{female}} + s_{\text{male}}^2 / n_{\text{male}})^2}{\frac{(s_{\text{female}}^2 / n_{\text{female}})^2}{n_{\text{female}} - 1} + \frac{(s_{\text{male}}^2 / n_{\text{male}})^2}{n_{\text{male}} - 1}}$$

(DOF stands for degree of freedom)

#### [Correction ends here]

• How do we perform the same test using regression technique?

$$income_i = \beta_0 + \beta_1 female_i + u_i$$

where female is a dummy variable.

Here,

- $-\beta_0$  records \_\_\_\_\_
- $-\beta_1$  records \_\_\_\_\_

So to test whether there's a difference between female and male population income, we can equivalently test whether  $\beta_1$  is nonzero:

$$H_0: \beta_1 = 0$$
  
$$H_1: \beta_1 \neq 0$$

– To test whether  $\beta_1 = 0$ , we can use what we learned from Dis 4, and construct t-statistic for test. Recall that t-statistic looks like the following:

$$t = \frac{\hat{\beta}_1 - \beta_{1,H_0}}{se(\hat{\beta}_1)} \quad \sim \quad t_{n-k-1}$$

- This tells us that how  $se(\hat{\beta}_1)$  looks like matters:
  - \* When homoskedastic error based  $se(\hat{\beta}_1)$  is used, this is equivalent to Econ 310's test under equal population variance.
  - \* When heteroskedastic error based  $se(\hat{\beta}_1)$  (i.e. robust standard errors) is used, this is equivalent to Econ 310's test under unequal population variance.

### 4 Problems

1.	Load the dataset from http://fmwww.bc.edu/ec-p/data/wooldridge/wage2.dta into Stata (don't forget to first change your working directory).
	Dataset codebook is available at http://fmwww.bc.edu/ec-p/data/wooldridge/wage2.des
	(a) Estimate the following multivariate linear model using Stata's default regression setting (i.e. assuming homoskedastic error):
	$wage_i = \beta_0 + \beta_1 educ_i + \beta_2 exper_i + \beta_3 black_i + \beta_4 urban_i + \beta_5 married_i + \beta_6 hours_i + u_i$

(b) At 5% significance level, is there any slope coefficient that is not statistically significant?

(c) Is the OLS estimator BLUE under the current configuration?

(d) The default linear regression assumes homoskedastic error. One worries that heteroskedastic error is more appropriate here. Without performing any test, give a reason on why you'd think that heteroskedasticity might hold here.

(e)	Perform a visual test on heteroskedasticity by creating the regression postestimation diagnostic plot (rvfplot).
(f)	Perform Breusch-Pagan test on heteroskedasticity at 5% significance level.
(g)	Perform White test on heteroskedasticity at 5% significance level.
(h)	With the correct error specification, reestimate the linear regression model in (a).
( <del>;</del> )	Did any of the estimated coefficient change?
(1)	Did any of the estimated coefficient change?
(j)	Did any of the standard error for coefficient change?
(1.)	In the OLC action atom DLITE and on the many configuration?
(K)	Is the OLS estimator BLUE under the new configuration?
(1)	At 5% significance level, is there any slope coefficient that is not statistically significant now?