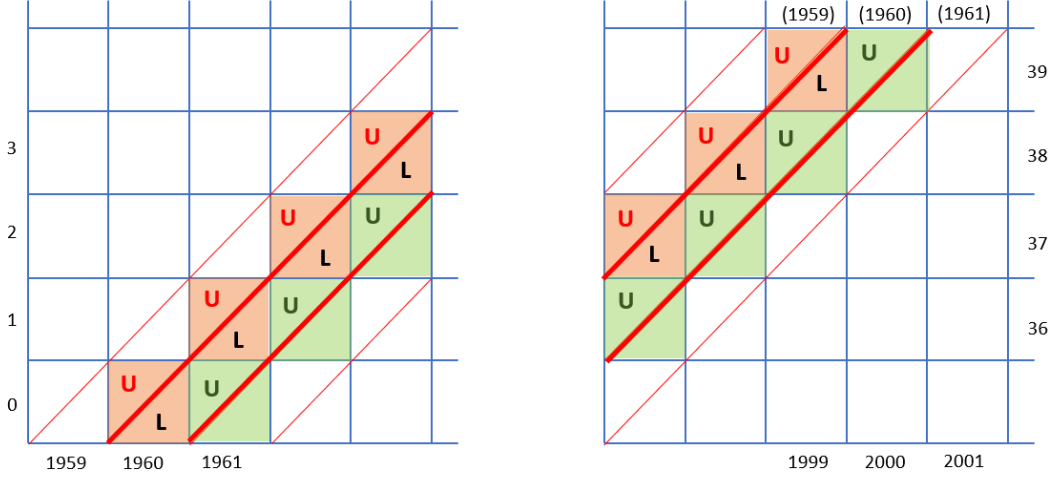


Figure 1: Lexis diagram for the 1960 birth cohort of women



Suppose that mortality is zero, and that there are N_{1960} women in the 1960 birth cohort. Define L_{xt} and U_{xt} as the number of births that happen in the lower and upper triangles, respectively, of the age-period square for age x , year t . The total births to the 1960 cohort before exact age 40 would then be

$$\begin{aligned} B_{1960} &= (L_{0,1960} + L_{1,1961} + \cdots + L_{39,1999}) \\ &\quad + (U_{0,1961} + U_{1,1962} + \cdots + U_{39,2000}) \\ &= L_{1960}^* + U_{1960}^* \end{aligned}$$

where we divide the births by the category of Lexis triangle that they fall into. With no mortality, true cohort completed fertility per woman would then be

$$F_{1960} = \frac{L_{1960}^* + U_{1960}^*}{N_{1960}}$$

If we used the orange age-period squares to approximate the fertility history there would be

$$\begin{aligned} B &= (L_{0,1960} + L_{1,1961} + \cdots + L_{39,1999}) + (U_{0,1960} + U_{1,1961} + \cdots + U_{39,1999}) \\ &= L_{1960}^* + U_{1959}^* \end{aligned}$$

and estimated completed fertility would be lifetime events / cohort size:

$$\hat{F}_{1960} = \frac{U_{1959}^* + L_{1960}^*}{\frac{1}{2}(N_{1959} + N_{1960})}$$

Simplifying Approximations

In order to develop a simple expression for “AP bias”, assume that

1. Mortality is negligible below the highest age of childbearing
2. In a small neighborhood around our focal cohort c , cohort sizes change multiplicatively:

$$\frac{N_c}{N_{c-1}} = \frac{N_{c+1}}{N_c} = K$$
3. In a small neighborhood around our focal cohort c , completed fertility changes additively: $(F_c - F_{c-1}) = (F_{c+1} - F_c) = \Delta$
4. Half of a cohort’s lifetime births happen in lower Lexis triangles, half in upper triangles:

$$L_c^* = U_c^* \approx \frac{1}{2} N_c F_c$$

Under these assumptions the “diagonal AP” estimator for completed fertility of the cohort born in year c is

$$\hat{F}_c \approx \frac{\frac{1}{2} N_{c-1} F_{c-1} + \frac{1}{2} N_c F_c}{\frac{1}{2} (N_{c-1} + N_c)}$$

If cohort sizes around c are changing multiplicatively with annual multiplier K , this implies that

$$\begin{aligned} \hat{F}_c &\approx \frac{\frac{1}{2} N_{c-1} F_{c-1} + \frac{1}{2} K N_{c-1} F_c}{\frac{1}{2} (N_{c-1} + K \cdot N_{c-1})} \\ &= \frac{F_{c-1} + K F_c}{K + 1} \\ &= \left(\frac{1}{K+1}\right) F_{c-1} + \left(\frac{K}{K+1}\right) F_c \end{aligned}$$

That is, the AP diagonal estimator is approximately equal to a weighted average of the previous and current cohort’s completed fertility. The approximate additive error from using the AP diagonal estimator is therefore

$$\hat{F}_c - F_c = \left(\frac{1}{K+1}\right) (F_{c-1} - F_c) = -\left(\frac{\Delta}{K+1}\right)$$

and the approximate proportional error is

$$\frac{\hat{F}_c - F_c}{F_c} = -\left(\frac{1}{K+1}\right) \left(\frac{\Delta}{F_c}\right)$$

From these relationships we learn that

- bias from using the AP diagonal estimator is negative when completed fertility is rising with cohort birth year ($\Delta > 0$), and negative when completed fertility is falling
- bias from using the AP diagonal estimator is smaller when cohort size is increasing more quickly between consecutive birth years (larger K implies additive and relative bias closer to zero)

averaging AP diagonals

If we try estimating F_{1960} by averaging \hat{F}_{1960} and \hat{F}_{1961} we get **[assume that cohort size ratios stay constant]**

$$\begin{aligned}\tilde{F}_{1960} &= \frac{1}{2} \left[\left(\frac{1}{K+1} \right) F_{1959} + \left(\frac{K}{K+1} \right) F_{1960} \right] \\ &\quad + \frac{1}{2} \left[\left(\frac{1}{K+1} \right) F_{1960} + \left(\frac{K}{K+1} \right) F_{1961} \right] \\ &= \frac{1}{2} F_{1960} + \frac{1}{2} \left[\left(\frac{1}{K+1} \right) F_{1959} + \left(\frac{K}{K+1} \right) F_{1961} \right]\end{aligned}$$

If change in F_c across cohorts is linear, $F_c - F_{c-1} = \Delta$, then

$$\begin{aligned}\tilde{F}_{1960} &= \frac{1}{2} F_{1960} + \frac{1}{2} \left[\left(\frac{1}{K+1} \right) (F_{1960} - \Delta) + \left(\frac{K}{K+1} \right) (F_{1960} + \Delta) \right] \\ &= F_{1960} + \frac{1}{2} \left(\frac{K-1}{K+1} \right) \Delta\end{aligned}$$

old derivation

So the error from summing along the diagonal of the age-period squares is

$$\hat{F}_{1960} - F_{1960} = \frac{L_{1960}^* + U_{1959}^*}{\frac{1}{2}(N_{1959} + N_{1960})} - \frac{L_{1960}^* + U_{1960}^*}{N_{1960}}$$

Defining

$$K = \frac{N_{1959}}{N_{1960}}$$

as the size of the 1959 cohort relative to the 1960 cohort, we can rewrite the error as

$$\begin{aligned}
\hat{F}_{1960} - F_{1960} &= \frac{L_{1960}^* + U_{1959}^*}{\frac{1}{2}(K \cdot N_{1960} + N_{1960})} - \frac{L_{1960}^* + U_{1960}^*}{N_{1960}} \\
&\approx \frac{\frac{1}{2}F_{1960} \cdot N_{1960} + \frac{1}{2}F_{1959} \cdot K \cdot N_{1960}}{\frac{1}{2}(K \cdot N_{1960} + N_{1960})} - F_{1960} \\
&= \frac{F_{1960} + F_{1959} \cdot K}{K + 1} - F_{1960} \\
&= F_{1960} \left(\frac{1 + \frac{F_{1959}}{F_{1960}} \cdot K}{K + 1} - 1 \right) \\
&= F_{1960} \left(\frac{1 + \frac{F_{1959}}{F_{1960}} \cdot K}{K + 1} - \frac{K + 1}{K + 1} \right) \\
&= F_{1960} \left(\frac{K \cdot \left(\frac{F_{1959}}{F_{1960}} - 1 \right)}{K + 1} \right) \\
&= F_{1960} \left(\frac{K \cdot}{K + 1} \right) \left(\frac{F_{1959}}{F_{1960}} - 1 \right) \\
&= \left(\frac{K \cdot}{K + 1} \right) (F_{1959} - F_{1960})
\end{aligned}$$

$$\begin{aligned}
\hat{F}_{1960} - F_{1960} &= \frac{L_{1960}^* + U_{1959}^*}{\frac{1}{2}(K \cdot N_{1960} + N_{1960})} - \frac{L_{1960}^* + U_{1960}^*}{N_{1960}} \\
&= \frac{1}{N_{1960}} \left[\frac{L_{1960}^* + U_{1959}^*}{\frac{1}{2}(K + 1)} - (L_{1960}^* + U_{1960}^*) \right] \\
&= \frac{1}{N_{1960}} \left[\left(\frac{L_{1960}^*}{\frac{1}{2}(K + 1)} - L_{1960}^* \right) + \left(\frac{U_{1959}^*}{\frac{1}{2}(K + 1)} - U_{1960}^* \right) \right] \\
&= \frac{1}{N_{1960}} \left[\left(\frac{L_{1960}^*}{\frac{1}{2}(K + 1)} - L_{1960}^* \right) + \left(\frac{U_{1959}^*}{\frac{1}{2}(K + 1)} - U_{1959}^* \right) + (U_{1959}^* - U_{1960}^*) \right] \\
&= \frac{1}{N_{1960}} \left[\left(\frac{1}{\frac{1}{2}(K + 1)} - 1 \right) (L_{1960}^* + U_{1959}^*) + (U_{1959}^* - U_{1960}^*) \right] \\
&= \frac{1}{N_{1960}} \left[\underbrace{\left(\frac{1 - K}{1 + K} \right) (L_{1960}^* + U_{1959}^*)}_{\text{cohort size effect}} + \underbrace{(U_{1959}^* - U_{1960}^*)}_{\text{rate trend effect}} \right]
\end{aligned}$$

where the first term in square brackets is a cohort size effect, and the second is a rate-change effect.

If $K = 1$ the error simplifies to $\hat{F}_{1960} - F_{1960} = \frac{U_{1959}^* - U_{1960}^*}{N_{1960}}$.