

Worksheet 1: Practice with Joint Measures

Name:

Due September 15, 2023

You will investigate the [rule of succession](#) which states that the probability of success for the $(m+1)$ th trial in a sequence of *independent* bernoulli random variables x_1, \dots, x_{m+1} —given $k \leq m$ successes in the first m trials—is

$$\mathbb{P} \left(x_{m+1} = 1 \mid \sum_{j=1}^m x_j = k \right) = \frac{k+1}{m+2}. \quad (1)$$

In principle if x_1, \dots, x_{m+1} are truly independent, you might expect the right hand side to simply be $\mathbb{P}(x_{m+1} = 1)$ which does not depend on k (i.e. the outcomes of x_1, \dots, x_m).

In this worksheet, you will explore the right space to make sense of eq. (1) in. While seemingly elementary, this example provides a strong example foreshadowing Bayesian reasoning in machine learning, where model-defining parameters are defined by a measure instead of as point-values. We let $\mathcal{X} = \{0, 1\}$, $x \in \mathcal{X}$ denote an outcome, and define random variable $s_m : \mathcal{X}^m \rightarrow \mathbb{R}$ by $(x_1, \dots, x_m) \mapsto \sum_{j=1}^m x_j$. Implicit in eq. (1) is that $p := \mathbb{P}(x = 1)$ is unknown (i.e. there is some distribution on p) and that *given* p , *then* the trials x_1, \dots, x_{m+1} are independent. To concretize this observation, we let $I = [0, 1]$ and suppose that there is joint distribution $\mathbb{P}_{\mathcal{X}^m \times I}$ according to which $\mathbb{P}_{\mathcal{X}^m|I}$ is independent. In particular,

$$\mathbb{P}_{\mathcal{X}^m|I} \left(\sum_{j=1}^m x_j = k \mid i = p \right) = \binom{m}{k} p^k (1-p)^{m-k}.*$$

Because p does not appear explicitly in eq. (1), it must be (being) marginalized away. That means that

$$\mathbb{P}(x_{m+1} = 1 \mid s_m(x_1, \dots, x_m) = k) = \mathbb{P}_{\mathcal{X} \times I | \mathcal{X}^m}(\{x_{m+1} = 1\} \times I \mid s_m(x_1, \dots, x_m) = k). \quad (2)$$

Your task, therefore, is to show that (the right hand side of) (2) equals the right hand side of (1).

A reasonable measure to place on the marginal probability \mathbb{P}_I is the uniform one $\mathbb{P}_I((a, b)) = b - a$ for $0 \leq a \leq b \leq 1$. With this preamble, you may now proceed.

1. Using the definition of conditional probability, rewrite the right hand side of eq. (2) as a ratio of two joint probabilities. You may use that $\mathcal{X} \times I \times \mathcal{X}^m \cong \mathcal{X}^{m+1} \times I$ to simplify (but be careful to write the correct event in the numerator!). In the denominator, express the marginal probability $\mathbb{P}_{\mathcal{X}^m}$ in terms of $\mathbb{P}_{\mathcal{X}^m \times I}$; you will see why in the next problem.

(3)

*See [Binomial Distribution](#) for more information.

2. Conveniently, this expression (nearly) collapses the problem to a single computation (of either numerator or denominator). To compute either probability, you will want to condition again, this time as $\mathcal{X}^m | I$, and use the law of total probability.[†] Express each one:

(4)

3. Explicitly compute the inner (conditional) probability (or expectation) for both:

4. Show by induction on $0 \leq k \leq m$ and $m \in \mathbb{N}$ that $\int_0^1 p^k (1-p)^{m-k} dp = \frac{k!(m-k)!}{(m+1)!}$. For the induction on m (and useful for question 5.), show also that $\int_0^1 p^{k+1} (1-p)^{m-k} dp = \frac{(k+1)!(m-k)!}{(m+2)!}$. (You will make liberal use of integration by parts.)

5. Use results in parts 3. and 4. to compute the expressions in (4) and simplifying, conclude (1).

[†]Recalling that $\mathbb{P}_{\mathcal{X} \times \mathcal{Y}}(R) = \int_R d\mathbb{P}_{\mathcal{X} \times \mathcal{Y}} = \int_{\mathcal{X} \times \mathcal{Y}} \mathbb{1}_{(x,y) \in R} d\mathbb{P}_{\mathcal{X} \times \mathcal{Y}}(x,y) = \int_{\mathcal{X}} \int_{\mathcal{Y}|x} \mathbb{1}_{(x,y) \in R} d\mathbb{P}_{\mathcal{Y}|x}(y|x) d\mathbb{P}_{\mathcal{X}}(x)$ for region $R \subset \mathcal{X} \times \mathcal{Y}$.