STA305/1004-Class 19

March 22, 2017

Today's Class

- Other Blocking Designs
  - Latin Square
  - Graeco Latin Square
  - hypo-Graeco Latin Square
  - Randomized incomplete block design
- Assessing significance in unreplicated factorial designs
  - Normal plots
  - half-Normal plots
  - Lenth's method

Factorial Assignment

- Read the sample report.
- You are supposed to design an experiment using a factorial design.
- This means I want you to generate the data by running an experiment. So finding data (e.g., on the web) is not appropriate.
- What are the controllable input variables (factors) in your experiment? What is the response variable?
- Example: How does coffee consumption and hours of sleep affect running speed?

2^2 , 2^3

Randomized block designs

Paired expt.
is a blocked
design with
blocks of
size = 2.

- A group of homogeneous units is referred to as a block.
- Examples; days, weeks, batches, lots, and sets of twins.
- For blocking to be effective units should be arranged so that the within-block variation is much smaller than the between block variation.

Randomized block designs

Know before the expt. begins what you want to block.

How many blocking factors? One factor with, say, k levels! $v_i = 2$.

- Consider an experiment to compare ~~sales~~ time to relief from four different medical treatments for headache.
- The effectiveness of the four treatments is known to be different for different age groups.
- Therefore, block by age group.
- A randomized block design randomly assigns subjects in each block to the four treatments.
- Experimental Design Principle: "Block what you can and randomize what you cannot."

Age group #1 Rand trt 1 2 3 4

Age group #2 Rand. trt 1 2 3 4

Blocked by age group.

Completely rand. design

Rand trt 1 2 3 4

Other blocking designs

Block for more than 2 factors

- Latin square
- Graeco-Latin squares,
- Hyper-Graeco-Latin Squares,
- Balanced incomplete block designs.

The Latin Square Design

- There are several other types of designs that utilize the blocking principle such as The Latin Square design.
- If there is more than one nuisance source that can be eliminated then a Latin Square design might be appropriate.

Latin Square Design - Automobile Emissions

- An experiment to test the feasibility of reducing air pollution.
- A gasoline mixture was modified by changing the amounts of certain chemicals.
- This produced four different types of gasoline: A, B, C, D
- These four treatments were tested with four different drivers and four different cars.

Latin Square Design - Automobile Emissions

- Two blocking factors: cars and drivers.
- The Latin square design was used to help eliminate possible differences between drivers I, II, III, IV and cars 1, 2, 3, 4.
- Randomly allocate treatments, drivers , and cars.

Driver	Car 1	Car 2	Car 3	Car 4
Driver I	A	B	D	C
Driver II	D	C	A	B
Driver III	B	D	C	A
Driver IV	C	A	B	D

Latin Square Design - Automobile Emissions

- The data from the experiment.

Driver	Car 1	Car 2	Car 3	Car 4
Driver I	A	B	D	C
	19	24	23	26
Driver II	D	C	A	B
	23	24	19	30
Driver III	B	D	C	A
	15	14	15	16
Driver IV	C	A	B	D
	19	18	19	16

Blocking Variables: Car; driver. Each blocking variable has 4 levels.

Treatment: Additives $\rightarrow$ 4 levels.

Latin Square Design - Automobile Emissions

```r
sapply(split(tab0408$y,tab0408$cars), mean)# car means
```

```
 1  2  3  4
19 20 19 22
```

```r
sapply(split(tab0408$y,tab0408$driver), mean)# driver means
```

```
 1  2  3  4
23 24 15 18
```

```r
sapply(split(tab0408$y,tab0408$additive), mean)# additive means
```

```
 A  B  C  D
18 22 21 19
```

```r
mean(tab0408$y) #grand mean
```

```
[1] 20
```

Latin Square Design - Automobile Emissions

- Why not standardize the conditions and make the 16 experimental runs with a single car and single driver for the four treatments?
- Could also be valid but Latin square provides a wider inductive basis.

Latin Square Design - Automobile Emissions

$H_0: \mu_A = \mu_B = \mu_C = \mu_D$ Cannot reject.

```r
latinsq.auto <- lm(y~additive+as.factor(cars)+as.factor(driver),data=tab0408)
anova(latinsq.auto)
```

If the SS_E is larger then
$MSE \uparrow \Rightarrow MS_{Treat}/MSE \downarrow$

```
Analysis of Variance Table

Response: y
                  Df Sum Sq Mean Sq F value   Pr(>F)
additive           3     40  13.333     2.5 0.156490
as.factor(cars)    3     24   8.000     1.5 0.307174
as.factor(driver)  3    216  72.000    13.5 0.004466 **
Residuals          6     32   5.333
---
Signif. codes:  0 '***' 0.001 '**' 0.01 '*' 0.05 '.' 0.1 ' ' 1
```

$$SS_T = SS_{cars} + SS_{drivers} + SS_{Additives} + SS_E$$

If Cars and drivers were not used as blocking factors then residual SS would be larger.

Automobile Emissions

If blocking variables are not used in calculating the ANOVA table.

```
latinsq.auto <- lm(y~additive,data=tab0408)
anova(latinsq.auto)
```

```
Analysis of Variance Table

Response: y
          Df Sum Sq Mean Sq F value Pr(>F)
additive   3     40  13.333  0.5882 0.6343
Residuals 12    272  22.667
```

$272 = 32 + 216 + 24$

Latin Square Design - Automobile Emissions

Choose replication scheme based on feasibility.

- Assuming that the residuals are independent and normally distributed and the null hypothesis that there are no treatment differences is true then the ratio of mean squares for treatments and residuals has an $F_{3,6}$ distribution.
- This analysis assumes that treatments, cars, and drivers are additive.
- If the design was replicated then this would increase the degrees of freedom for the residuals and reduce the mean square error.

- Use same cars/drivers in each rep.
- Use same cars but different drivers in each rep.
- Use same drivers but different cars in each rep.
- Use different cars and different drivers in each rep.

General Latin Square

- A Latin square for p factors of a $p \times p$ Latin square, is a square containing p rows and p columns
- Each of the p^2 cells contains one of the p letters that correspond to a treatment.
- Each letter occurs once and only once in each row and column.
- There are many possible $p \times p$ Latin squares.

General Latin Square

Which of the following is a Latin square?

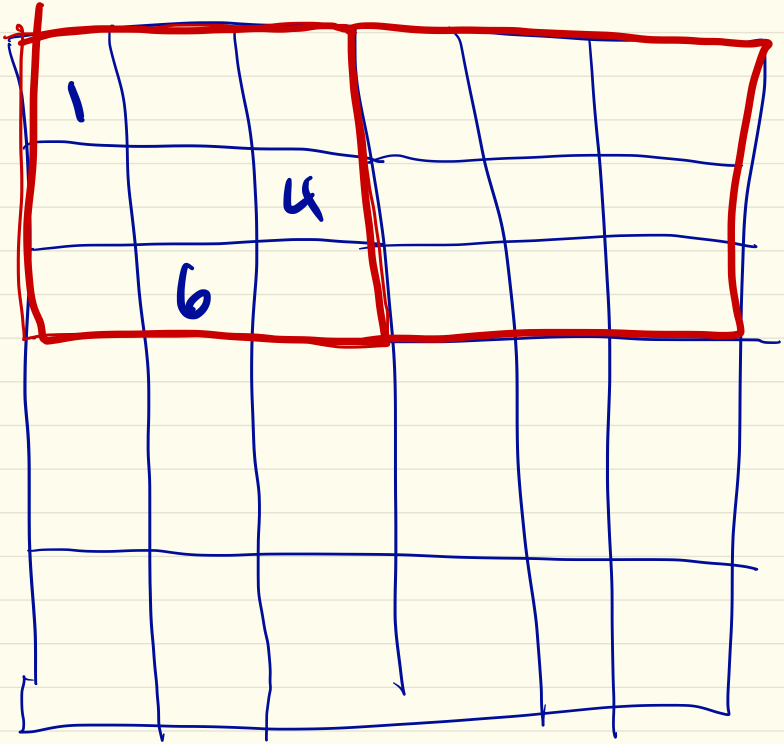
	Col1	Col2	Col3
Row 1	B	A	C
Row 2	A	C	B
Row 3	C	B	A

✓ each letter occurs once in each row/col.

	Col1	Col2	Col3
Row 1	A	B	C
Row 2	C	A	B
Row 3	B	B	A

not the case.

Soduku



fill every cell with numbers 1-9
So each number appears once and only once
in each row/col.

Poll question

An industrial engineer is investigating the effect of four assembly methods (A, B, C, D) on the assembly time of a component. Four operators are selected for the study. The engineer also knows that each assembly method produces fatigue such that the time required for the last assembly might be greater than the time required for the first, regardless of method. The engineer randomly assigns the order that each operator uses the four methods: operator 1 uses the methods in the order: C, A, D, B

only one blocking factor ∴ A

Operator	A	B	C	D
I	2	4	1	3
II	4	2	3	1
III	2	1	4	3
IV	3	4	1	2

Randomizing within each level of operator.

The design is:

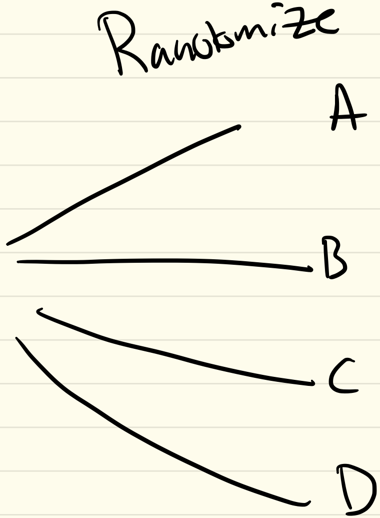
Respond at **PollEv.com/nathantaback**

Text **NATHANTABACK** to **37607** once to join, then **A, B, or C**

Treatment: assembly method
block: operator.

Randomized block design	A	16
Randomized design (without blocking)	B	6
Latin square	C	2

operator



Completely randomized Design.

	Machine			
	1	2	3	4
Operator				
I	A	B	C	D
II	C	D	A	B
III	⋮	⋮	⋮	⋮
IV	⋮	⋮	⋮	⋮

Latin Square design. Blocking factors: Machine (4 levels) ; Operator (4 levels).

Misuse of the Latin Square

- Inappropriate to use Latin square to study factors that can interact.
- Effects of one factor can then be mixed up with interactions of other factors.
- Outliers can occur as a result of these interactions.
- When interactions between factors are likely possible need to use a factorial design.

Graeco-Latin Square

A Graeco-Latin square is a $k \times k$ pattern that permits study of k treatments simultaneously with three different blocking variables each at k levels.

	Car 1	Car 2	Car 3	Car 4
Driver I	A α	B β	C γ	D δ
Driver II	B δ	A γ	D β	C α
Driver III	C β	D α	A δ	B γ
Driver IV	D γ	C δ	B α	A β

Graeco-Latin Square

- This is a Latin square in which each Greek letter appears once and only once with each Latin letter.
- Can be used to control three sources of extraneous variability (i.e. block in three different directions).

Car	1 Car	2 Car	3 Ca	r 4
Driver I	A α	B β	C γ	D δ
Driver II	B δ	A γ	D β	C α
Driver III	C β	D α	A δ	B γ
Driver IV	D γ	C δ	B α	A β

Graeco-Latin Square

To generate a 3×3 Graeco-Latin square design, superimpose two designs using the Greek letters for the second 3×3 Latin square.

	Col1	Col2	Col3
Row 1	B	A	C
Row 2	A	C	B
Row 3	C	B	A

$B = \beta$

$A = \alpha$

$C = \gamma$

	Col1	Col2	Col3
Row 1	A β	B α	C γ
Row 2	C α	A γ	B β
Row 3	B γ	C β	A α

Graeco-Latin Square.

hyper-Graeco-Latin Square

These three Latin squares can be superimposed to form a hyper-Graeco-Latin square.
Can be used to control 4 nuisance factors (i.e. block 4 factors).

4 Blocking factors are: Col, row, Greek letters, numbers.

Row	Col1	Col2	Col3	Col4
Row 1	B	A	D	C
Row 2	C	D	A	B
Row 3	D	B	C	A
Row 4	A	C	B	D

$B = \beta$
$C = \gamma$
$D = \delta$
$A = \alpha$

Row	Col1	Col2	Col3	Col4
Row 1	D	A	C	B
Row 2	A	D	B	C
Row 3	B	C	A	D
Row 4	C	B	D	A

D = 4
A = 1
B = 2
C = 3

Row	Col1	Col2	Col3	Col4
Row 1	A β 4	D α 1	B	C
Row 2	C γ 1	A δ 4	D	B
Row 3	B δ 2	C β 3	A	D
Row 4	D α 3	B γ 2	C	A

hyper-Graeco-Latin Square

- A machine used for testing the wear on types of cloth.
- Four pieces of cloth can be compared simultaneously on one machine.
- Response is weight loss in tenths of mg when rubbed against a standard grade of emery paper for 1000 revolutions of the machine.

hyper-Graeco-Latin Square

- Specimens of 4 different cloths (A, B,C,D) are compared.
- The wearing qualities can be in any one of 4 positions P_1, P_2, P_3, P_4 on the machine.
- Each emery $(\alpha, \beta, \gamma, \delta)$ paper used to cut into for quarters and each quarter used to complete a cycle C_1, C_2, C_3, C_4 of 1000 revolutions.
- Object was to compare treatments.

hyper-Graeco-Latin Square

1. type of specimen holders 1, 2, 3, 4
2. position on the machine P_1, P_2, P_3, P_4.
3. emory paper sheet $\alpha, \beta, \gamma, \delta$.
4. machine cycle C_1, C_2, C_3, C_4.

The design was replicated. The first replicate is shown in the table below.

	P_1	P_2	P_3	P_4
C_1	A α 1	B β 2	C γ 3	D δ 4
	320	297	299	313
C_2	C β 4	D α 3	A δ 2	B γ 1
	266	227	260	240
C_3	D γ 2	C δ 1	B α 4	A β 3
	221	240	267	252
C_4	B δ 3	A γ 4	D β 1	C α 2
	301	238	243	290

hyper-Graeco-Latin Square

A linear model can be fit so that the ANOVA table and parameter treatment effects can be calculated.

```
wear.hypsq <- lm(y~treatment+as.factor(rep)+as.factor(position)+
                   as.factor(cycle)+as.factor(holder)+
                   as.factor(paper),data=tab0412)
anova(wear.hypsq)
```

4 different Cloths
Sig. different.

```
Analysis of Variance Table

Response: y
                    Df  Sum Sq Mean Sq F value    Pr(>F)
treatment            3  1705.3  568.45  5.3908  0.021245 *
as.factor(rep)       1   603.8  603.78  5.7259  0.040366 *
as.factor(position)  3  2217.3  739.11  7.0093  0.009925 **
as.factor(cycle)     6 14770.4 2461.74 23.3455 5.273e-05 ***
as.factor(holder)    3   109.1   36.36  0.3449  0.793790
as.factor(paper)     6  6108.9 1018.16  9.6555  0.001698 **
Residuals            9   949.0  105.45
---
Signif. codes:  0 '***' 0.001 '**' 0.01 '*' 0.05 '.' 0.1 ' ' 1
```

Balanced incomplete block design

- Suppose that instead of four samples to be included on each 1000 revolution cycle only three could be included, but the experimenter still wanted to compare four treatments.
- The size of the block is now 3 - too small to accommodate all treatments simultaneously.

Balanced incomplete block design

A balanced incomplete block design has the property that every pair of treatments occurs together in a block the same number of times.

AB occurs twice
AC " "
AD " "
BC " "
BD " "
CD " "

Cycle block			
1	A	B	C
2	A	B	D
3	A	C	D
4	B	C	D

Cycle block	A	B	C	D
1	x	x	x	
2	x	x		x
3	x		x	x
4		x	x	x

Back to Factorial Designs . . .

How can significance be assessed in unreplicated factorial designs?

- A factorial design is an expt. design that considers all factor-level combinations.
- Cannot calculate the standard error of factorial effects if the design is not replicated.

Quantile-Quantile Plots

- Quantile-quantile (Q-Q) plots are useful for comparing distribution functions.
- If X is a continuous random variable with strictly increasing distribution function $F(x)$ then the *pth* quantile of the distribution is the value of x_p such that,

  $$F(x_p) = p$$

  or

  $$x_p = F^{-1}(p).$$

- In a Q-Q plot, the quantiles of one distribution are plotted against another distribution.
- Q-Q plots can be used to investigate if a set of numbers follows a certain distribution.

Quantile-Quantile Plots

- Suppose that we have observations independent observations $X_1, X_2, ..., X_n$ from a uniform distribution on $[0, 1]$ or Unif[0,1].
- The ordered sample values (also called the order statistics) are the values $X_{(j)}$ such that

$$X_{(1)} < X_{(2)} < \cdots < X_{(n)}$$

.

- It can be shown that

$$E\left(X_{(j)}\right) = \frac{j}{n+1}.$$

- This suggests that if we plot

$$X_{(j)} \quad \text{vs.} \quad \frac{j}{n+1}$$

then if the underlying distribution is Unif[0,1] then the plot should be rooughly linear.

Quantile-Quantile Plots

$Y = F(X) \quad P(Y \leq y) = P(F(X) \leq y) = P(X \leq F^{-1}(y))$
$= F(F^{-1}(y)) = y.$

- A continuous random variable with strictly increasing CDF F_X can be transformed to a Unif[0,1] by defining a new random variable $Y = F_X(X)$.
- Suppose that it's hypothesized that X follows a certain distribution function with CDF F.
- Given a sample $X_1, X_2, ..., X_n$ plot

$$F(X_{(k)}) \quad \text{vs.} \quad \frac{k}{n+1}$$

or equivalently

$$X_{(k)} \quad \text{vs.} \quad F^{-1}\left(\frac{k}{n+1}\right)$$

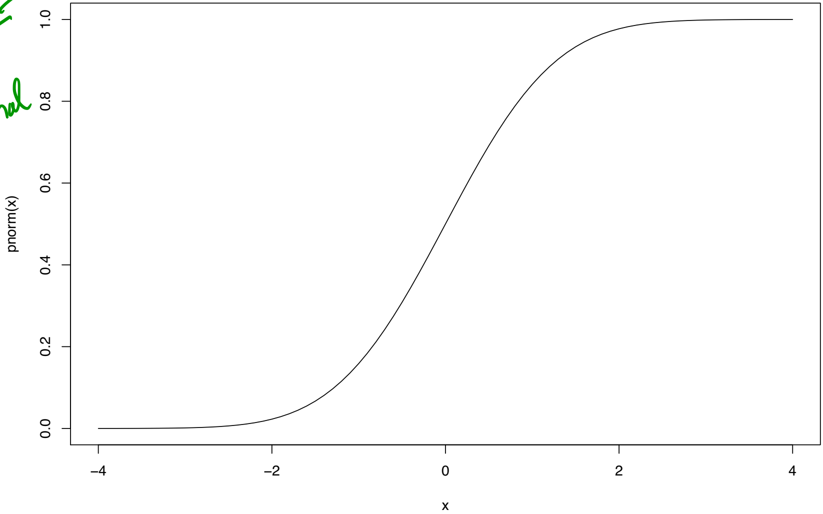
- $X_{(k)}$ can be thought of as empirical quantiles and $F^{-1}\left(\frac{k}{n+1}\right)$ as the hypothesized quantiles.
- The quantile assigned to $X_{(k)}$ is not unique.
- Instead of assigning it $\frac{k}{n+1}$ it is often assigned $\frac{k-0.5}{n}$. In practice it makes little difference which definition is used.

Normal Quantile-Quantile Plots

The cumulative distribution function (CDF) of the normal has an S-shape.

```
x <- seq(-4,4,by=0.1)
plot(x,pnorm(x),type="l")
```

CDF of Normal



Normal Quantile-Quantile Plots

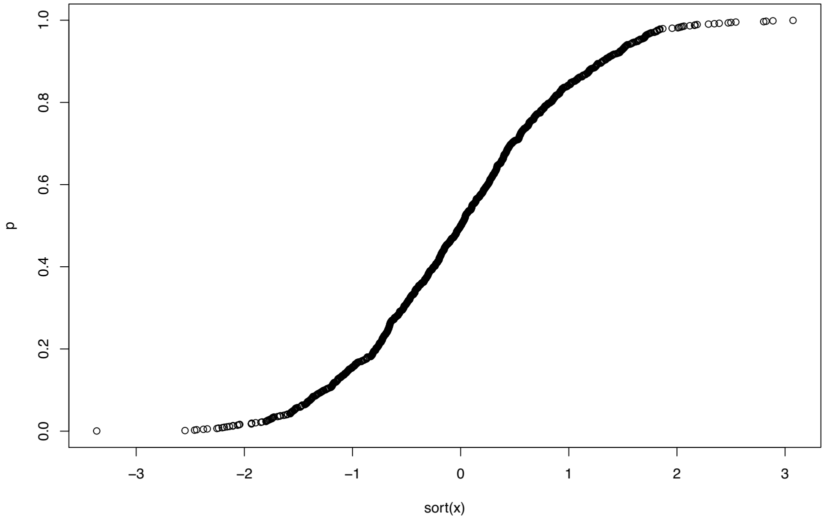
The normality of a set of data can be assessed by the following method.

- Let $r_{(1)} < ... < r_{(N)}$ denote the ordered values of $r_1, ..., r_N$.
- A test of normality for a set of data is to plot the ordered values $r_{(i)}$ of the data versus $p_i = (i - 0.5)/N$.
- If the plot has the same S-shape as the normal CDF then this is evidence that the data come from a normal distribution.

Normal Quantile-Quantile Plots

- A plot of $r_{(i)}$ vs. $p_i = (i - 0.5)/N, i = 1, ..., N$ for a random sample of 1000 simulated from a $N(0, 1)$.

```
N <- 1000;x <- rnorm(N);p <- ((1:N)-0.5)/N
plot(sort(x),p)
```



Normal Quantile-Quantile Plots

- It can be shown that $\Phi(r_i)$ has a uniform distribution on $[0, 1]$.
- This implies that $E(\Phi(r_{(i)})) = i/(N+1)$ (this is the expected value of the *jth* order statistic from a uniform distribution over $[0, 1]$.
- This implies that the N points $(p_i, \Phi(r_{(i)}))$ should fall on a straight line.
- Now apply the Φ^{-1} transformation to the horizontal and vertical scales. The N points

  $$\left(\Phi^{-1}(p_i), r_{(i)}\right),$$

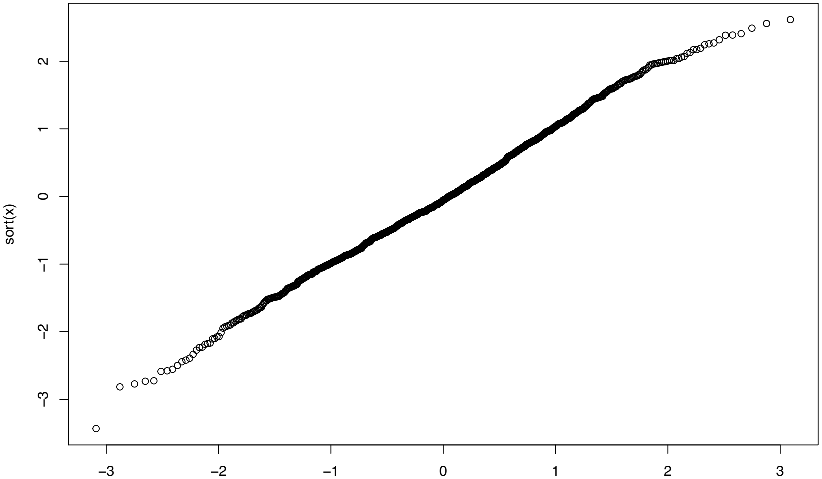
  form the normal probability plot of $r_1, ..., r_N$.
- If $r_1, ..., r_N$ are generated from a normal distribution then a plot of the points $\left(\Phi^{-1}(p_i), r_{(i)}\right), i = 1, ..., N$ should be a straight line.

Normal Quantile-Quantile Plots

In R `qnorm()` is Φ^{-1}.

```r
set.seed(2503)
N <- 1000
x <- rnorm(N)
p <- (1:N)/(N+1)
plot(qnorm(p),sort(x))
```

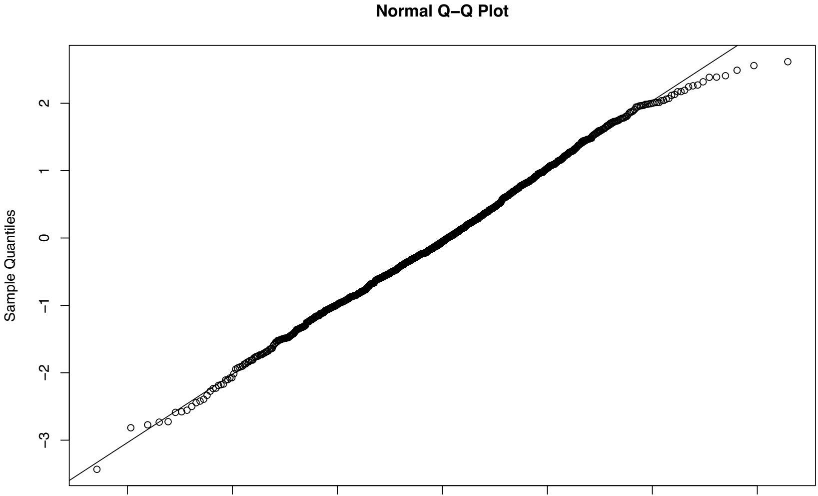
$(\Phi^{-1}(p_i), r_{(i)})$



Normal Quantile-Quantile Plots

We usually use the built in function `qqnorm()` (and `qqline()` to add a straight line for comparison) to generate normal Q-Q plots. Note that R uses a slightly more general version of quantile ($p_i = (1-a)/(N + (1-a) - a)$, where $a = 3/8$, if $N \leq 10$, $a = 1/2$, if $N > 10$.

```
qqnorm(x);qqline(x)
```



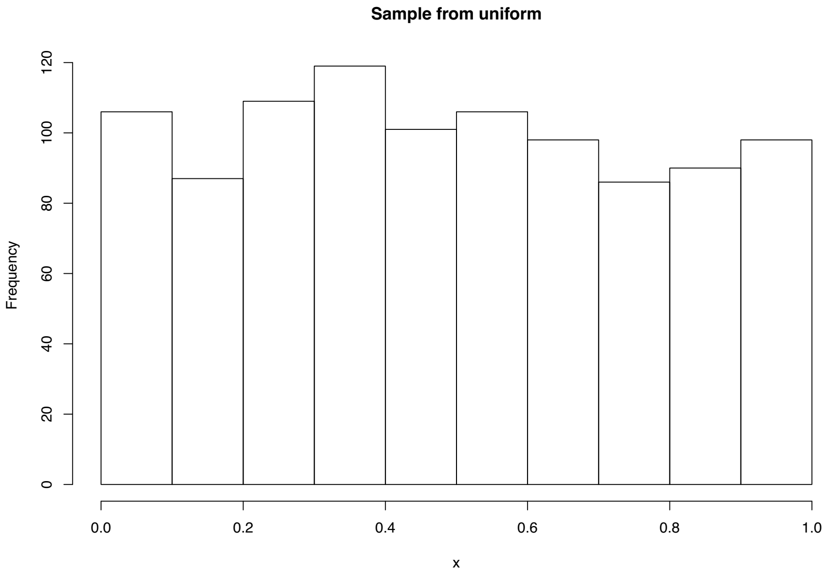
Normal Quantile-Quantile Plots

A marked (systematic) deviation of the plot from the straight line ~~would~~ could indicate that:

1. The normality assumption does not hold.
2. The variance is not constant.

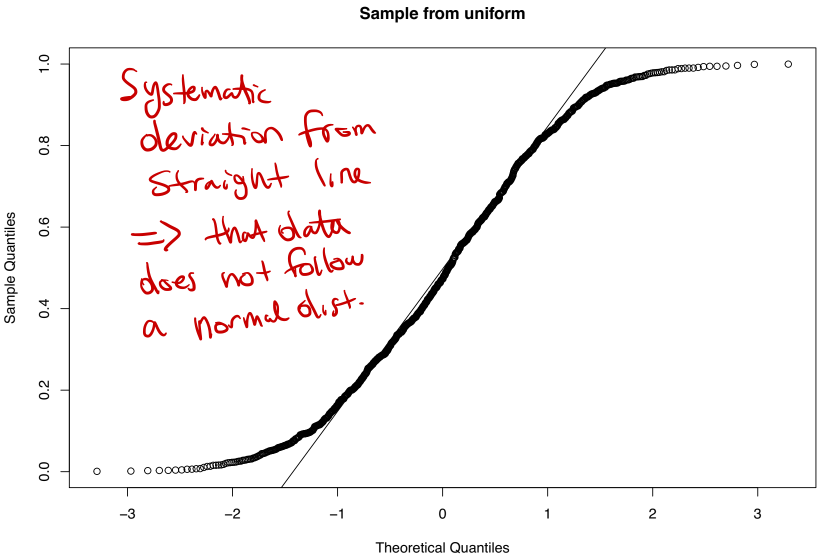
Normal Quantile-Quantile Plots

```r
x <- runif(1000)
hist(x,main = "Sample from uniform")
```



Normal Quantile-Quantile Plots

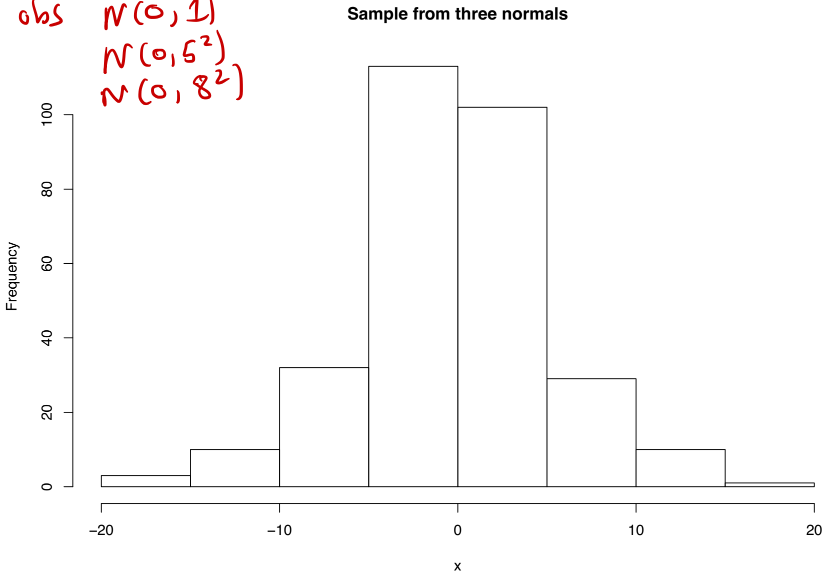
```
qqnorm(x,main = "Sample from uniform");qqline(x)
```



Normal Quantile-Quantile Plots

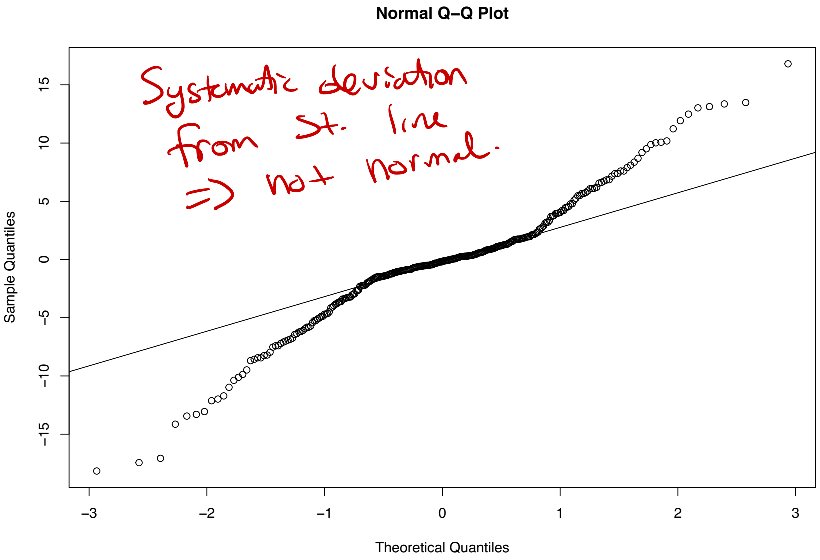
```
x1 <- rnorm(100,mean = 0,sd = 1);x2 <- rnorm(100,mean = 0,sd = 5)
x3 <- rnorm(100,mean = 0,sd = 8); x <- c(x1,x2,x3)
hist(x,main = "Sample from three normals")
```

100 obs $N(0,1)$
$N(0,5^2)$
$N(0,8^2)$



Normal Quantile-Quantile Plots

```
qqnorm(x);qqline(x)
```



Normal plots in factorial experiments

(main effects or interaction terms)

- A major application is in factorial designs where the ~~$r(i)$~~ $r_{(i)}$ are replaced by ordered factorial effects.
- Let $\hat{\theta}_{(1)} < \hat{\theta}_{(2)} < \cdots < \hat{\theta}_{(N)}$ be N ordered factorial estimates.
- If we plot

Smallest

largest.

$$\hat{\theta}_{(i)} \text{ vs. } \Phi^{-1}(p_i).\ i = 1, ..., N.$$

then factorial effects $\hat{\theta}_i$ that are close to 0 will fall along a straight line. Therefore, points that fall off the straight line will be declared significant.

If the design is 2^2 then $N=3$

$\hat{\theta}_{(i)}$ Could be the main effect of a factor.

→ See next slide ...

2^2 design

Run	A	B	AB	Outcome
1	-1	-1	$+1$	y_1
2	$+1$	-1	-1	y_2
3	-1	$+1$	-1	y_3
4	$+1$	$+1$	$+1$	y_4

main effect of A?

$$\left(\frac{y_4 + y_2}{2}\right) - \left(\frac{y_1 + y_3}{2}\right) \neq 0$$

$$H_0: \mu_{A+} - \mu_{A-} = 0$$

Normal plots in factorial experiments

The rationale is as follows:

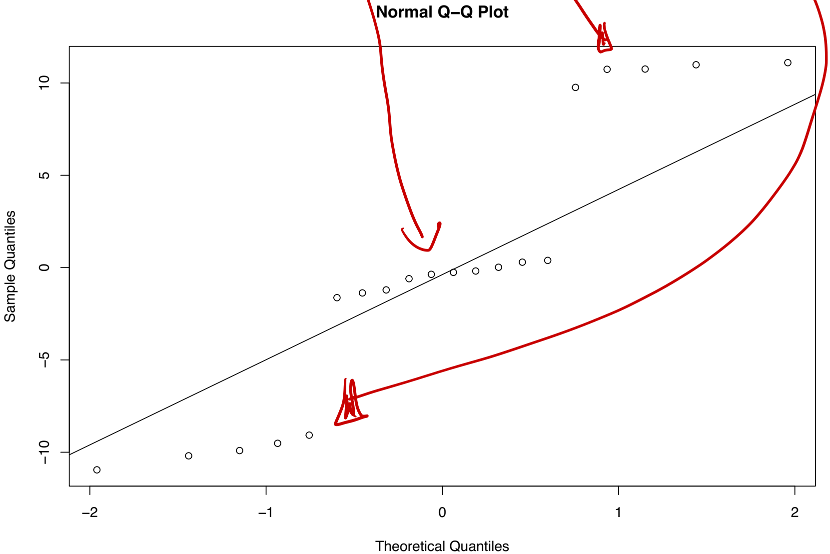
1. Assume that the estimated effects $\hat{\theta}_i$ are $N(\theta, \sigma^2)$ (estimated effects involve averaging of N observations and CLT ensures averages are nearly normal for N as small as 8).
2. If $H_0 : \theta_i = 0,\ i = 1, ..., N$ is true then all the estimated effects will be zero.
3. The resulting normal probability plot of the estimated effects will be a straight line.
4. Therefore, the normal probability plot is testing whether all of the estimated effects have the same distribution (i.e. same means).

- When some of the effects are nonzero the corresponding estimated effects will tend to be larger and fall off the straight line.

Normal plots in factorial experiments

Positive effects fall above the line and negative effects fall below the line.

```
set.seed(10);x1 <- rnorm(10,0,1); x2 <- rnorm(5,10,1);x3 <- rnorm(5,-10,1)
x <- c(x1,x2,x3);qqnorm(x);qqline(x)
```



Example - 2^3 design for studying a chemical reaction

A process development experiment studied four factors in a 2^4 factorial design.

- amount of catalyst charge **1**,
- temperature **2**,
- pressure **3**,
- concentration of one of the reactants **4**.
- The response y is the percent conversion at each of the 16 run conditions. The design is shown below.

Example - 2^4 design for studying a chemical reaction

x1	x2	x3	x4	conversion
-1	-1	-1	-1	70
1	-1	-1	-1	60
-1	1	-1	-1	89
1	1	-1	-1	81
-1	-1	1	-1	69
1	-1	1	-1	62
-1	1	1	-1	88
1	1	1	-1	81
-1	-1	-1	1	60
1	-1	-1	1	49
-1	1	-1	1	88
1	1	-1	1	82
-1	-1	1	1	60
1	-1	1	1	52
-1	1	1	1	86
1	1	1	1	79

The design is not replicated so it's not possible to estimate the standard errors of the factorial effects.

Example - 2^4 design for studying a chemical reaction

recall that the least Squares estimates are $\frac{1}{2}$ factorial estimates $\therefore$ multiply by 2.

```
fact1 <- lm(conversion~x1*x2*x3*x4,data=tab0510a)
round(2*fact1$coefficients,2)
```

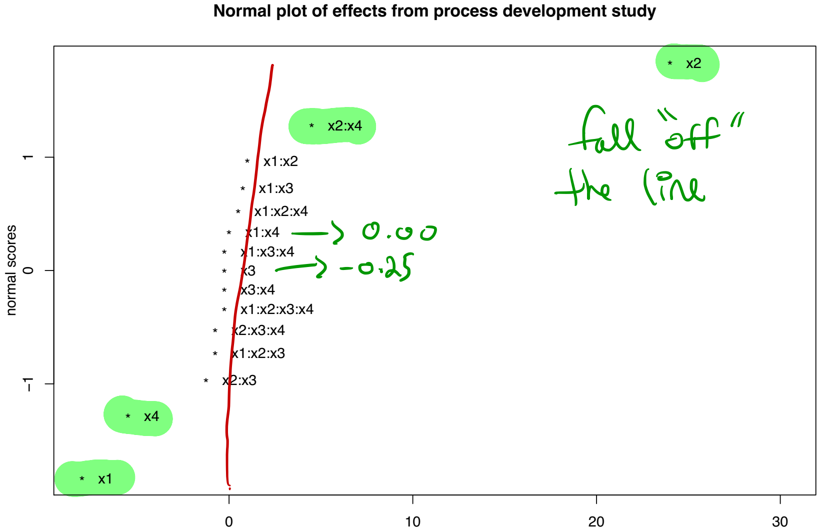
Interaction between x_1, x_2

```
(Intercept)          x1          x2          x3          x4       x1:x2
     144.50       -8.00       24.00       -0.25       -5.50        1.00
      x1:x3       x2:x3       x1:x4       x2:x4       x3:x4    x1:x2:x3
       0.75       -1.25        0.00        4.50       -0.25       -0.75
   x1:x2:x4    x1:x3:x4    x2:x3:x4 x1:x2:x3:x4
       0.50       -0.25       -0.75       -0.25
```

Example - 2^4 design for studying a chemical reaction

A normal plot of the factorial effects is obtained by using the function `DanielPlot()` in the `FrF2` library.

```
library(FrF2)
DanielPlot(fact1, autolab=F,main="Normal plot of effects from process developme
```



Example 1

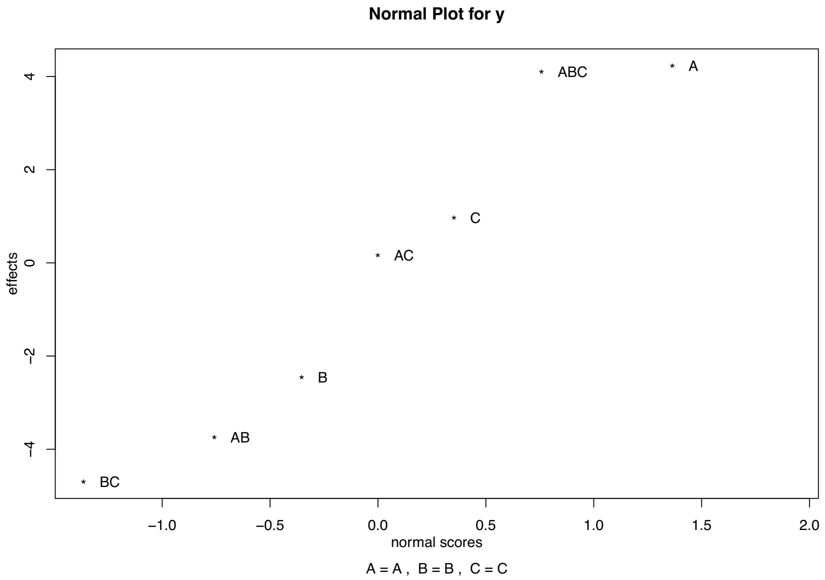
Which effects are not explained by chance?

```
##
## Call:
## lm.default(formula = y ~ A * B * C, data = dat)
##
## Residuals:
## ALL 8 residuals are 0: no residual degrees of freedom!
##
## Coefficients:
##             Estimate Std. Error t value Pr(>|t|)
## (Intercept) -0.57063         NA      NA       NA
## A1           2.11739         NA      NA       NA
## B1          -1.22742         NA      NA       NA
## C1           0.48534         NA      NA       NA
## A1:B1       -1.86973         NA      NA       NA
## A1:C1        0.08116         NA      NA       NA
## B1:C1       -2.34868         NA      NA       NA
## A1:B1:C1     2.05018         NA      NA       NA
##
## Residual standard error: NaN on 0 degrees of freedom
## Multiple R-squared:      1,  Adjusted R-squared:    NaN
## F-statistic:   NaN on 7 and 0 DF,  p-value: NA
```

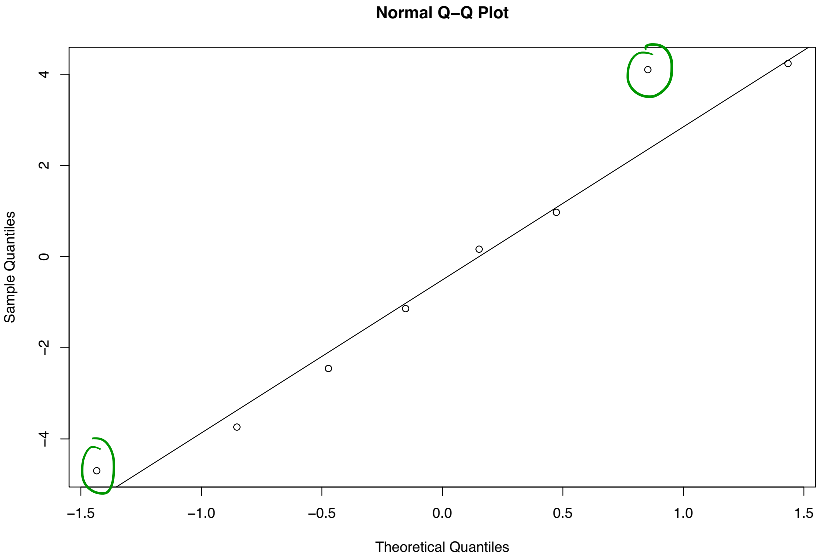
Example 1

Which effects are not explained by chance according to the normal plot?

```
FrF2::DanielPlot(mod1,code=TRUE,autolab=F,datax=F)
```



Example 1


Normal Q–Q Plot
Sample Quantiles
Theoretical Quantiles
4
2
0
−2
−4
−1.5
−1.0
−0.5
0.0
0.5
1.0
1.5

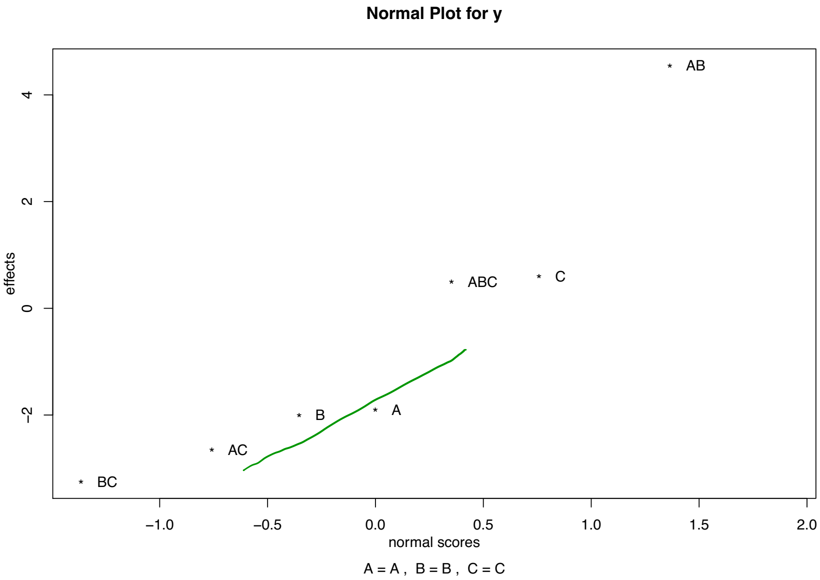
Example 2

Which effects are not explained by chance?

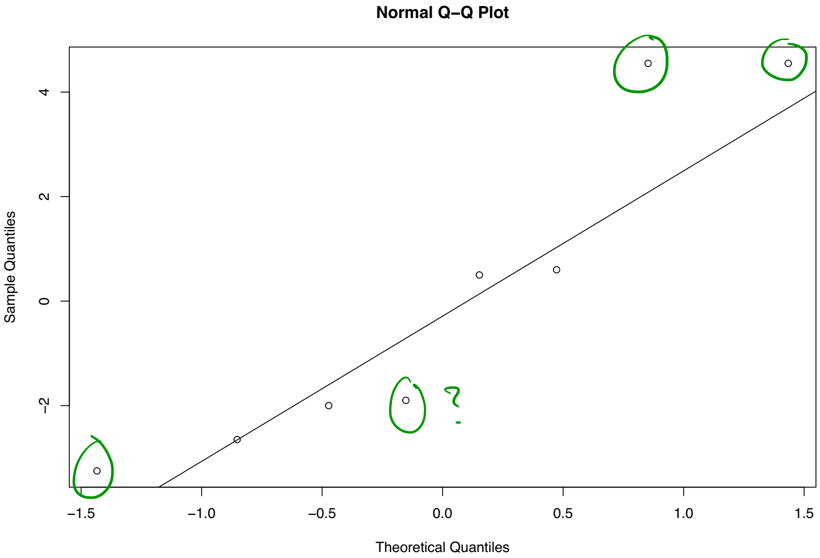
```
##
## Call:
## lm.default(formula = y ~ A * B * C, data = dat)
##
## Residuals:
## ALL 8 residuals are 0: no residual degrees of freedom!
##
## Coefficients:
##             Estimate Std. Error t value Pr(>|t|)
## (Intercept)    2.275         NA      NA       NA
## A1             2.150         NA      NA       NA
## B1             1.125         NA      NA       NA
## C1            -1.500         NA      NA       NA
## A1:B1          0.950         NA      NA       NA
## A1:C1         -1.575         NA      NA       NA
## B1:C1         -0.300         NA      NA       NA
## A1:B1:C1      -0.125         NA      NA       NA
##
## Residual standard error: NaN on 0 degrees of freedom
## Multiple R-squared:      1,  Adjusted R-squared:    NaN
## F-statistic:   NaN on 7 and 0 DF,  p-value: NA
```

Example 2

Which effects are not explained by chance according to the normal plot?



Example 2



Normal Q–Q Plot
Sample Quantiles
Theoretical Quantiles
4
2
0
−2
−1.5
−1.0
−0.5
0.0
0.5
1.0
1.5
?

Half-Normal Plots

- A related graphical method is called the half-normal probability plot.
- Let

$$\left|\hat{\theta}\right|_{(1)} < \left|\hat{\theta}\right|_{(2)} < \cdots < \left|\hat{\theta}\right|_{(N)}.$$

  denote the ordered values of the unsigned factorial effect estimates.
- Plot them against the coordinates based on the half-normal distribution - the absolute value of a normal random variable has a half-normal distribution.
- The half-normal probability plot consists of the points

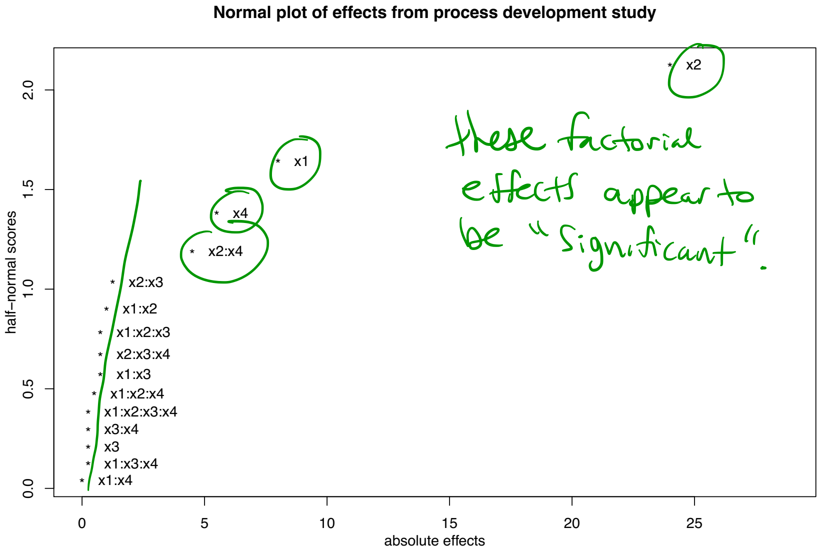
$$\left|\hat{\theta}\right|_{(i)} \text{ vs. } \Phi^{-1}(0.5 + 0.5[i - 0.5]/N).\ i = 1, ..., N.$$

Half-Normal Plots

- An advantage of this plot is that all the large estimated effects appear in the upper right hand corner and fall above the line.
- The half-normal plot for the effects in the process development example is can be obtained with `DanielPlot()` with the option `half=TRUE`.

Half-Normal Plots - 2^4 design for studying a chemical reaction

```
library(FrF2)
DanielPlot(fact1,half=TRUE,autolab=F,
          main="Normal plot of effects from process development study")
```



Half-Normal Plots - 2^4 design for studying a chemical reaction

Compare with full Normal plot.

```
library(FrF2)
DanielPlot(fact1,half=F,autolab=F,
          main="Normal plot of effects from process development study")
```