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Anna Hughes · Alasdair D. F. Clarke

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Abstract Insert your abstract here. Include keywords, PACS and mathematical subject classification numbers as needed.

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1 Introduction

Visual search, where participants are asked to find a target within a cluttered scene, has been extensively studied within psychology. A number of models have been developed that can generate testable predictions about how different types of distractors and targets affect search efficiency, including Feature Integration Theory [5] and Guided Search [8,7]. These models aim to produce a representation of the visual properties of the distractors at each location in the visual field. However, a recent model, Target Contrast Signal (TCS) Theory, proposes that this is in fact not the goal of the visual system: the authors instead suggest that the most appropriate measure is how different each item is from the target [4].

TCS proposes that behaviour is determined by comparing the target template (held in memory) with every element present in the scene in parallel. This allows the visual system to reject peripheral non-targets quickly; the speed at which items are evaluated is determined by how different the item is from the template through an evidence accumulation process. The model

F. Author first address

Tel.: +123-45-678910 Fax: +123-45-678910

 $\hbox{E-mail: fauthor@example.com}$

S. Author second address

thus focuses on an initial, efficient processing stage of search; if sufficient evidence is not accumulated during this process, the model posits that a second stage is entered, where attention is deployed serially. TCS has been successful in predicting a number of empirical results, including search performance in heterogeneous scenes based on parameters estimated in homogeneous scenes, both with artificial stimuli [1,3] and with real-world objects visualised on a computer display [6].

While results to date for TCS appear promising, they remain relatively preliminary, being tested by only one research group so far. In addition, predictions of search efficiency (e.g. in heterogeneous scenes) are made on the average of a group of participants, using data from a different group performing a different task (e.g. searching in homogeneous scenes). Thus, we know that TCS can replicate group-level averages in search well, but we do not know whether it is also able to make predictions at the individual level. There is also scope for extending beyond the parameters tested to date (such as the number of distractors) in order to test the robustness and generalisability of the model. Finally, there are some predictions, including its ability to explain search asymmetries, that have yet to be empirically confirmed [4].

In the current manuscript, we focus on replicating and extending findings from [2]. Here, participants searched for a target in a scene of homogeneous distractors. First, parallel search efficiency (measured by the logarithmic search slope) was estimated for cases where the distractors varied from the target in one dimension: either colour (e.g. a cyan target being searched for in either yellow, blue or orange distractors) or shape (e.g. a semicircle target in either circle, diamond or triangle distractors). New participants then searched for the same targets in displays where the distractors were compounds, differing from the target in both colour and shape (e.g. searching for a cyan semicircle in either blue circles, orange diamonds or yellow triangles). The logarithmic search slopes in the initial experiments were then used to predict the logarithmic slopes and reaction times using a number of models. The authors found that the best model was a 'collinear contrast integration model' where the distinctiveness scores were summed along each attribute in the unidimensional experiments, creating an overall contrast score that was used for compound stimuli predictions.

$$RT = a + \beta \sum_{j=1}^{L} \left((D_j - D_{j-1}) \log \left(\left(N_T - \sum_{j=1}^{j-1} N_i \right) 1_{[2,\infty](j)} + 1 \right) \right)$$
 (1)

WHen L=1 (all lures are identical) this simplifies to:

$$RT = a + D_1 \log(N_T + 1) \tag{2}$$

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1.1 Hypotheses

We plan a number of experiments to test the extent to which the original results in [2] replicate and generalise. As well as following the original, between-subjects, experiment design, in which data from one group of observers in one task is used to predict behaviour of a second group of participants in a different task, we will allow for within-subject comparisons. Specifically, we will ask: to what extent do the individual differences in the homogeneous task explain the differences in the heterogeneous task?

- 1. Replication of Buetti et al (2019) with online data collection. Specifically, that the collinear contrast ingratiation model outperforms the best feature guidance, and orthogonal contrast combination models. Furthermore, the $R^2 = (99\% \text{ HPDI} = [,])$ between predicted and observer reaction times.
- 2. Larger number of distractors (or targets further in periphery?) Add an extra ring
- 3. Search asymmetries: test O in Qs and Q in Os

2 Reanalysis of Buetti et al (2019)

In our proposed experiments, we would like to:

- make use of multi-level models
- , and work within a Bayesian framework
- use a log-normal distribution to take skew in rt distribution into account

To start, we will re-analysis previous data to verify that this does not invalidate the original conclusions. And, how to phrase, something about these new results (on the old data) being the ones we want to replicate.

2.1 Methods

2.1.1 Data

Data taken from OSF... using the exclusion criteria originally used. What about incorrect trials etc?

Do we want to re-do any of the other papers while we're at it? There is an rt = 0, which I will remove.

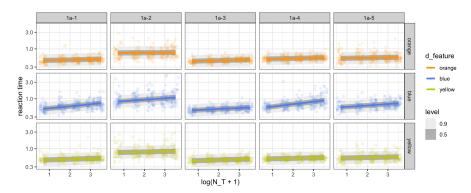


Fig. 1 Example results from the first five participants of Experiment 1a [2]. We can see considerable between-subjects variation. The shaded regions show the model's predicited reaction times.

2.1.2 Modelling Approach

R, brms, details.

We will compute the D values using a multi-level Bayesian model using a log-normal distribution. In order to keep our model in line with the Target Contrast Signal Theory, we will use the identity link rather than the default log link. We will also model reaction times on the seconds scale, rather than milliseconds. While this is a fairly trivial change, it makes numerical model fitting slightly easier.

$$\hat{y} = a + Dlog(N_T + 1) \tag{3}$$

```
m <- brm(
rt ~ 0 + d_feature + log(N_T+1):d_feature +
    (log(N_T+1):d_feature|p_id),
data = df,
family = lognormal(link = "identity"),
prior = my_priors,
iter = 5000)</pre>
```

Allows us to fit each individual's performance.

2.1.3 Prior Predictions

2.2 Results

2.2.1 Fixed Effects

As we are using a Bayesian framework, rather than D representing the best fit in equation 3, it is now represented as a probability distribution over possible values (see Figure 2.)

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Fig. 2 Posterior predictions of D for different distracter features. The vertical lines indicate the (transformed) values given in Buetti (2019).

2.2.2 Individual Differences

2.3 Discussion

3 Experiment 1

H0: confirm collinear method is best (compared to orthogonal contrast combination and best feature guidance models)

[If this is indeed the case, we will use only the collinear contrast integration model in the following experiments].

- 3.1 Methods
- 3.2 Results
- 3.3 Discussion

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