例1 已知 (X,Y) 在区域D : $0 \le x \le 1, 0 \le y \le 2x$

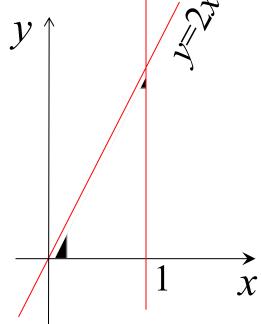
上的均匀分布,

试求:(1)Z=2X+Y的概率密度

(2) Z=X-Y的概率密度

解:(X,Y)的联合概率密度为:

$$f(x,y) = \begin{cases} 1, & 0 < x < 1, 0 < y < 2x \\ 0, & \text{#d} \end{cases}$$

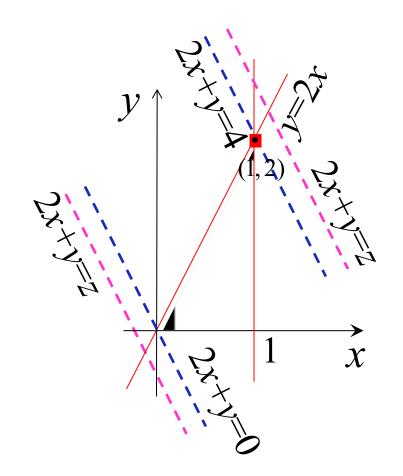


(沿直线积分直接求密度)

$$f_Z(z) = \int_{\overline{AB}} f(x, y) dx$$
$$= \int_{2x+y=z} f(x, y) dx$$

当z < 0 或 z > 4时

$$f_Z(z) = \int_{2x+y=z} 0 dx = 0$$

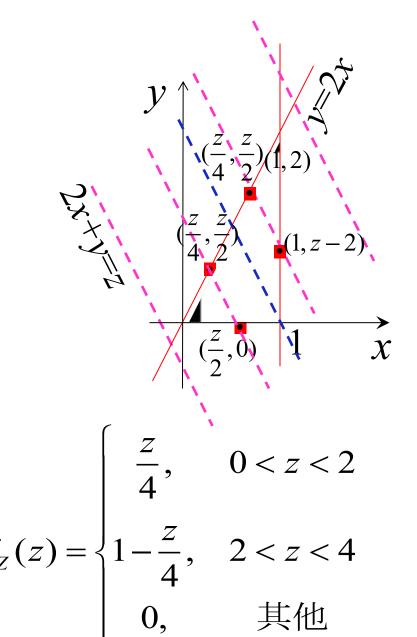


当0≤*z*≤2时

$$f_Z(z) = \int_{2x+y=z} f(x,y) dx$$
$$= \int_{\frac{z}{4}}^{\frac{z}{2}} 1 dx = \frac{z}{4}$$

当2<*z*≤4时

$$f_{Z}(z) = \int_{2x+y=z} f(x,y) dx$$
$$= \int_{\frac{z}{4}}^{1} 1 dx = 1 - \frac{z}{4}$$

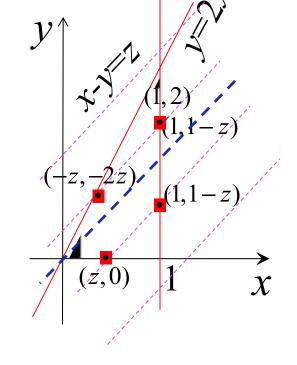


当z<-1或z>1时

$$f_Z(z) = \int_{x-y=z} 0 dx = 0$$

当-1≤*z*≤0时

$$f_Z(z) = \int_{x-y=z} f(x,y)dx$$
$$= \int_{-z}^1 1dx = 1+z$$



当0<*z*≤1时

$$f_Z(z) = \int_{x-y=z} f(x,y)dx$$
$$= \int_{z}^{1} 1dx = 1-z$$

$$f_{Z}(z) = \begin{cases} 1+z, & -1 \le z < 0 \\ 1-z, & 0 \le z < 1 \\ 0, & \sharp \& \end{cases}$$

(2) 极值分布:即极大值,极小值的分布

主要讨论相互独立的随机变量的极值分布

对于离散型随机变量的极值分布可直接计算

例2 X, Y相互独立, $X, Y \sim$ 参数为0.5的0-1分布 求 $M = \max\{X, Y\}$ 的概率分布

$p_{ij}X$	1	0	
1	0.25	0.25	
0	0.25	0.25	

$\max\{X,Y\}$	1	0	$\underline{\min\{X,Y\}}$	1	0
P	0.75	0.25	P	0.25	0.75

对于连续型随机变量 , 设 X,Y 相互独立, $X \sim F_X(x)$, $Y \sim F_Y(y)$, $U = \max\{X,Y\}$, $V = \min\{X,Y\}$, 求 U, V 的分布函数.

$$F_{U}(u) = P(\max\{X,Y\} \le u)$$

$$= P(X \le u, Y \le u) = P(X \le u)P(Y \le u)$$

$$= F_{X}(u)F_{Y}(u)$$

$$F_{Y}(v) = P(\min\{X,Y\} \le v) = 1 - P(\min\{X,Y\} > v)$$

$$= 1 - P(X > v, Y > v) = 1 - P(X > v)P(Y > v)$$

$$= 1 - (1 - F_{Y}(v))(1 - F_{Y}(v))$$

推广至相互独立的 n 个随机变量的情形:

设 X_1, X_2, \dots, X_n 相互独立,且

$$X_i \sim F_i(x_i), \quad i = 1, 2, \dots, n$$

$$U = \max\{X_1, X_2, \dots, X_n\}$$

$$V = \min\{X_1, X_2, \dots, X_n\}$$

则

$$F_U(u) = \prod_{i=1}^n F_i(u)$$

$$F_V(v) = 1 - \prod_{i=1}^{n} (1 - F_i(v))$$

- 例3 设系统 L 由相互独立的 n 个元件组成,连接方式为
- (1) 串联;
- (2) 并联;
- (3) 冷贮备(起初由一个元件工作,其它 n-1 个元件做冷贮备,当工作元件失效时, 贮备的元件逐个地自动替换);
- 如果 n 个元件的寿命分别为 X_1, X_2, \dots, X_n 且 $X_i \sim E(\lambda), i = 1, 2, \dots, n$

求在以上 3 种组成方式下,系统 L 的寿命 X 的密度函数.

解
$$f_{X_i}(x_i) = \begin{cases} \lambda e^{-\lambda x_i}, & x_i > 0 \\ 0, & \sharp \dot{\Xi} \end{cases}$$

$$F_{X_i}(x_i) = \begin{cases} 1 - e^{-\lambda x_i}, & x_i > 0 \\ 0, & \sharp \dot{\Xi} \end{cases}$$

(1)
$$X = \min\{X_1, X_2, \dots, X_n\}$$

$$F_X(x) = 1 - \prod_{i=1}^{n} (1 - F_{X_i}(x))$$

$$F_{X_i}(x) = \begin{cases} 1 - e^{-\lambda x}, & x > 0 \\ 0, & \sharp \Xi \end{cases}$$

$$1 - F_{X_i}(x) = \begin{cases} e^{-\lambda x}, & x > 0, \\ 1, & x \le 0 \end{cases}$$

$$\prod_{i=1}^{n} (1 - F_{X_i}(x)) = \begin{cases} (e^{-\lambda x})^n, & x > 0, \\ 1, & x \le 0 \end{cases}$$

$$F_X(x) = 1 - \prod_{i=1}^n (1 - F_{X_i}(x)) = \begin{cases} 1 - (e^{-\lambda x})^n, & x > 0, \\ 0, & x \le 0 \end{cases}$$

$$f_X(x) = \begin{cases} n\lambda e^{-n\lambda x}, & x > 0 \\ 0, & x \le 0 \end{cases}$$

(2)
$$X = \max\{X_1, X_2, \dots, X_n\}$$

$$F_{X_i}(x) = \begin{cases} 1 - e^{-\lambda x}, & x > 0 \\ 0, & \sharp \stackrel{\sim}{\Sigma} \end{cases}$$

$$F_X(x) = \prod_{i=1}^n F_{X_i}(x)$$

$$=\begin{cases} (1-e^{-\lambda x})^n, & x>0, \\ 0, & x\leq 0 \end{cases}$$

$$f_X(x) = \begin{cases} n\lambda e^{-\lambda x} (1 - e^{-\lambda x})^{n-1}, & x > 0 \\ 0, & x \le 0 \end{cases}$$

(3)
$$X = X_1 + X_2 + \dots + X_n$$
$$n = 2 \text{ 日寸 },$$

$$f_{X_1+X_2}(x) = \int_{-\infty}^{+\infty} f_{X_1}(t) f_{X_2}(x-t) dt$$

$$= \begin{cases} \lambda^2 x e^{-\lambda x}, & x > 0 \\ 0, & x \le 0 \end{cases}$$

可以证明, X_1+X_2 与 X_3 也相互独立,故

$$f_{X_1+X_2+X_3}(x) = \int_{-\infty}^{+\infty} f_{X_1+X_2}(t) f_{X_3}(x-t) dt$$

$$= \begin{cases} \frac{\lambda^2 x^2}{2!} e^{-\lambda x}, & x > 0 \\ 0, & x \le 0 \end{cases}$$

归纳地可以证明,

$$f_{X}(x) = \begin{cases} \frac{(\lambda x)^{n-1}}{(n-1)!} e^{-\lambda x}, & x > 0\\ 0, & x \le 0 \end{cases}$$

例4设二维随机变量(X,Y)的概率密度为

$$f(x,y) = \begin{cases} \frac{1}{5}(2x+y), 0 \le x \le 2, 0 \le y \le 1 & y \\ 0, 其它 & 1 \end{cases}$$
(1)求max(X,Y)的概率密度;
(2)求min(X,Y)的概率密度;
解 (1)

$$F_{\max}(z) = P(\max(X, Y) \le z)$$

$$= P(X \le z, Y \le z)$$

$$= \iint f(x,y) dx dy$$

$$(a)$$
当 z < 0 时;

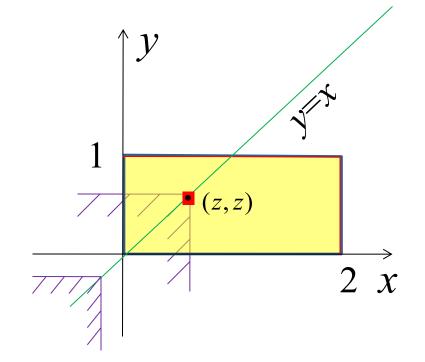
$$F_{\text{\tiny max}}(z) = 0$$

$$F_{\max}(z) = \iint_{\substack{x \le z \\ y \le z}} f(x, y) dx dy$$

$$= \int_0^z \int_0^z \frac{1}{5} (2x + y) dy dx$$

$$= \int_0^z \frac{1}{5} (2xz + \frac{1}{5}z^2) dx$$

$$=\frac{3}{10}z^3$$



$$F_{\text{max}}(z) = \iint_{\substack{x \le z \\ y \le z}} f(x, y) dx dy$$

$$= \int_0^z \int_0^1 \frac{1}{5} (2x + y) dy dx$$

$$= \int_0^z \frac{1}{5} (2x + \frac{1}{2}) dx$$

$$= \frac{1}{5} z^2 + \frac{1}{10} z$$

$$\begin{array}{c|c}
 & y \\
\hline
 & (z,z) \\
\hline
 & 2 & x
\end{array}$$

$$F_{\text{max}}(z) = 1$$

$$F_{\text{max}}(z) = \begin{cases} 0, & z < 0 \\ \frac{3}{10}z^{3}, & 0 \le z < 1 \\ \frac{1}{5}z^{2} + \frac{1}{10}z, & 1 \le z < 2 \\ 1 & z \ge 2 \end{cases}$$

$$f_{\text{max}}(z) = \begin{cases} \frac{9}{10}z^3, & 0 \le z < 1\\ \frac{2}{5}z + \frac{1}{10}, & 1 \le z < 2\\ 0, & \sharp \text{ } \end{cases}$$

解(2)
$$\neq P(X \le z, Y \le z)$$

$$F_{\min}(z) = P(\min(X, Y) \le z)$$

$$= 1 - P(\min(X, Y) > z)$$
$$= 1 - \iint_{Y>z} f(x, y) dxdy$$

(a)当
$$z>1$$
时; $F_{\min}(z) = 1 - 0 = 1$

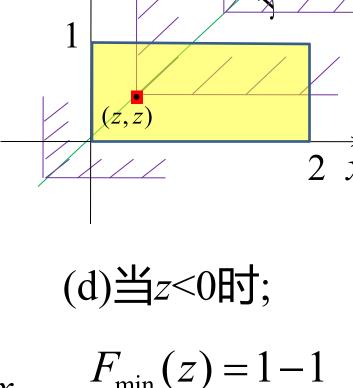
$$I_{\min}(2) - I = 0$$

(b)当0<z≤1时;

$$F_{\min}(z) = 1 - \int_{z}^{2} \int_{z}^{1} \frac{1}{5} (2x + y) dy dx$$

$$3 + 2 + 9$$

$$= -\frac{3}{10}z^{3} + \frac{2}{5}z^{2} + \frac{9}{10}z$$



$$F_{\text{max}}(z) = \begin{cases} 0, & z < 0 \\ -\frac{3}{10}z^3 + \frac{2}{5}z^2 + \frac{9}{10}z, & 0 \le z < 1 \\ 1, & z \ge 1 \end{cases}$$

$$f_{\min}(z) = \begin{cases} -\frac{9}{10}z^2 + \frac{4}{5}z + \frac{9}{10}, & 0 \le z < 1\\ 0, & \text{#$de} \end{cases}$$

(3) 平方和的分布: $Z = X^2 + Y^2$

设(X,Y)的联合密度函数为f(x,y)

$$\iint F_{Z}(z) = P(X^{2} + Y^{2} \le z)
= \begin{cases} 0, & z < 0, \\ \iint_{x^{2} + y^{2} \le z} f(x, y) dx dy & z \ge 0, \end{cases}
= \begin{cases} 0, & z < 0, \\ \int_{0}^{2\pi} d\theta \int_{0}^{\sqrt{z}} f(r \cos \theta, r \sin \theta) r dr, & z \ge 0, \end{cases}
f_{Z}(z) = \begin{cases} 0, & z < 0, \\ \frac{1}{2} \int_{0}^{2\pi} f(\sqrt{z} \cos \theta, \sqrt{z} \sin \theta) d\theta, & z \ge 0, \end{cases}$$

例如 , $X \sim N(0,1)$, $Y \sim N(0,1)$, X, Y相互独立 , $Z = X^2 + Y^2$, \mathbb{N}

$$Z = X^2 + Y^2$$
 , \mathbb{D}
$$f_Z(z) = \begin{cases} 0, & z < 0, \\ \frac{1}{2} \int_0^{2\pi} \frac{1}{\sqrt{2\pi}} e^{-\frac{z\cos^2\theta}{2}} \cdot \frac{1}{\sqrt{2\pi}} e^{-\frac{z\sin^2\theta}{2}} d\theta, & z \ge 0, \end{cases}$$
 $0, z < 0,$

$$f_{z}(z) = \begin{cases} 0, & z < 0, \\ \frac{1}{2}e^{-\frac{z}{2}}, & z \ge 0, \end{cases}$$
 称为自由度为2的 χ^{2} 分布

若 X_1, X_2, \dots, X_n 相互独立,且

$$X_i \sim N(0,1), i = 1,2,\dots,n$$

则 $Z = X_1^2 + X_2^2 + \dots + X_n^2$ 所服从的分布称为

自由度为n 的 χ^2 分布

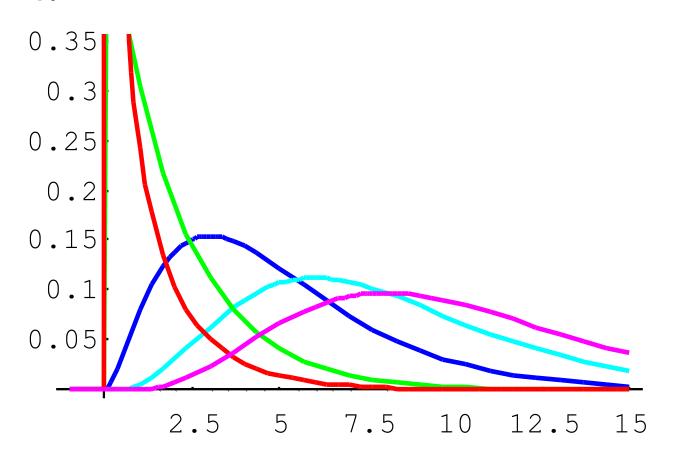
它的概率密度函数为

$$f_{Z}(z) = \begin{cases} 0, & z < 0, \\ \frac{1}{2^{\frac{n}{2}} \Gamma(\frac{n}{2})} z^{\frac{n}{2} - 1} e^{-\frac{z}{2}}, & z \ge 0, \end{cases}$$

其中
$$\Gamma(x) = \int_0^{+\infty} t^{x-1} e^{-t} dt$$
, $x > 0$ — 称为 Γ 函数

$$\Gamma(1)=1, \Gamma\left(\frac{1}{2}\right)=\sqrt{\pi},$$

 $\Gamma(x+1) = x\Gamma(x), \Gamma(n+1) = n!$ 自由度分别为1,2,5,8,10的 χ^2 分布的密度函数图形



自由度分别为1,2,5,8,10的 χ^2 分布的密度函数图形

