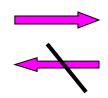
§ 5.4 协方差和相关系数

问题 对于二维随机变量(X,Y):

已知联合分布
边缘分布



这说明对于二维随机变量,除了每个随机 变量各自的概率特性以外,相互之间可能还有 某种联系. 问题是用一个什么样的数去反映这 种联系.

数
$$E((X - E(X))(Y - E(Y)))$$

= $E(XY) - E(X)E(Y)$

反映了随机变量X,Y之间的某种关系

● 协方差和相关系数的定义

定义 称 E((X-E(X))(Y-E(Y))) 为X,Y的协方差. 记为

$$cov(X,Y) = E((X - E(X))(Y - E(Y)))$$
$$= E(XY) - E(X)E(Y)$$

称
$$\begin{pmatrix} D(X) & cov(X,Y) \\ cov(X,Y) & D(Y) \end{pmatrix}$$

为 (X, Y) 的协方差矩阵

若D(X) > 0, D(Y) > 0,称

$$E\left(\frac{(X - E(X))(Y - E(Y))}{\sqrt{D(X)}\sqrt{D(Y)}}\right) = \frac{\text{cov}(X, Y)}{\sqrt{D(X)}\sqrt{D(Y)}}$$

为X,Y的相关系数,记为

$$\rho_{XY} = \frac{\text{cov}(X, Y)}{\sqrt{D(X)}\sqrt{D(Y)}}$$

事实上, $\rho_{XY} = \text{cov}(X^*, Y^*)$

$$cov(X^*, Y^*) = E[(X^* - EX^*)(Y^* - EY^*)]$$

$$= E\left[\left(\frac{X - EX}{\sqrt{DX}} - E\left(\frac{X - EX}{\sqrt{DX}}\right)\right)\left(\frac{Y - EY}{\sqrt{DY}} - E\left(\frac{Y - EY}{\sqrt{DY}}\right)\right)\right]$$

$$= E \left[\left(\frac{X - EX}{\sqrt{DX}} \right) \left(\frac{Y - EY}{\sqrt{DY}} \right) \right]$$

$$= \frac{E[(X - EX)(Y - EY)]}{\sqrt{DX}\sqrt{DY}} = \rho_{XY}$$

若 $\rho_{XY} = 0$, 称 X, Y 不相关.

● 协方差和相关系数的计算

——利用函数的期望或方差计算协方差

$$\Box \operatorname{cov}(X,Y) = E((X - E(X))(Y - E(Y)))$$

$$= E(XY) - E(X)E(Y)$$

$$= \frac{1}{2} (D(X+Y) - D(X) - D(Y))$$

$$= -\frac{1}{2} (D(X - Y) - D(X) - D(Y))$$

□ 若 (X,Y) 为离散型,

$$cov(X,Y) = E((X - E(X))(Y - E(Y)))$$

$$= \sum_{i=1}^{\infty} \sum_{j=1}^{\infty} (x_i - E(X))(y_j - E(Y))p_{ij}$$

□ 若 (*X*,*Y*) 为连续型,

$$cov(X,Y) = E((X - E(X))(Y - E(Y)))$$

$$=\int_{-\infty}^{+\infty}\int_{-\infty}^{+\infty}(x-E(X))(y-E(Y))f(x,y)dxdy$$

例1 已知X,Y的联合分布为

$P_{ij}X$	1	0	0 < <i>p</i> < 1
1	p	0	p + q = 1
0	0	q	

求 cov(X,Y), ρ_{XY}

解

X	1 0	Y	1 0	XY	1	0
P	p q	P	p q	P	p	q

$$\begin{array}{|c|c|c|c|c|c|}\hline X & 1 & 0 & Y & 1 & 0 & XY & 1 & 0\\\hline P & p & q & P & p & q & P & p & q\\\hline E(X) = p, & D(X) = pq, & & & & \\ E(Y) = p, & D(Y) = pq, & & & & \\ E(XY) = p, & D(XY) = pq, & & & \\ & & & & & & \\ Cov(X,Y) = E(XY) - E(X)E(Y) & & & \\ & & & & & & \\ & & & & & & \\ & & & & & & \\ & & & & & & \\ & & & & & \\ & & & & & \\ & & & & & \\ & & & & & \\ & & & & \\ & & & & \\ & & & & \\ & & & & \\ & & & & \\ & & & & \\$$

妈2 设(
$$X,Y$$
) ~ $N(\mu_1, \sigma_1^2, \mu_2, \sigma_2^2, \rho)$, 求 ρ_{XY}

解 $cov(X,Y) = \int_{-\infty}^{+\infty} \int_{-\infty}^{+\infty} (x - \mu_1)(y - \mu_2) f(x,y) dx dy$

$$= \int_{-\infty}^{+\infty} \int_{-\infty}^{+\infty} (x - \mu_1)(y - \mu_2) \cdot \frac{1}{2\pi\sigma_1\sigma_2\sqrt{1 - \rho^2}}$$

$$\cdot \exp\left\{-\frac{1}{2(1-\rho^2)}\left[\frac{(x-\mu_1)^2}{\sigma_1^2} - \frac{2\rho(x-\mu_1)(y-\mu_2)}{\sigma_1\sigma_2} + \frac{(y-\mu_2)^2}{\sigma_2^2}\right]\right\} dxdy$$

$$\frac{\frac{1}{\sigma_1} \frac{x - \mu_1}{\sigma_1}}{\sigma_2} = s$$

$$= \int_{-\infty}^{+\infty} \int_{-\infty}^{+\infty} \sigma_1 s \cdot \sigma_2 t \cdot \frac{1}{2\pi\sigma_1\sigma_2\sqrt{1-\rho^2}}$$

$$e^{-\frac{1}{2(1-\rho^2)}(s^2-2\rho st+t^2)}(\sigma_1 ds)(\sigma_2 dt)$$

$$=\frac{\sigma_{1}\sigma_{2}}{2\pi\sqrt{1-\rho^{2}}}\int_{-\infty}^{+\infty}\int_{-\infty}^{+\infty}ste^{-\frac{1}{2(1-\rho^{2})}[(s-\rho t)^{2}+(1-\rho^{2})t^{2}]}dsdt$$

$$= \frac{\sigma_1 \sigma_2}{2\pi \sqrt{1-\rho^2}} \int_{-\infty}^{+\infty} \int_{-\infty}^{+\infty} ste^{-\frac{1}{2(1-\rho^2)}(s-\rho t)^2 - \frac{1}{2}t^2} dsdt$$

$$\stackrel{\text{fis} - \rho t = u}{=} \frac{\sigma_1 \sigma_2}{2\pi \sqrt{1 - \rho^2}} \int_{-\infty}^{+\infty} \int_{-\infty}^{+\infty} t(\rho t + u) e^{-\frac{u^2}{2(1 - \rho^2)} - \frac{1}{2}t^2} du dt$$

$$= \frac{\sigma_{1}\sigma_{2}}{2\pi\sqrt{1-\rho^{2}}} \begin{bmatrix} \int_{-\infty}^{+\infty} \int_{-\infty}^{+\infty} \rho t^{2} \cdot e^{-\frac{u^{2}}{2(1-\rho^{2})} - \frac{1}{2}t^{2}} dudt \\ + \int_{-\infty}^{+\infty} \int_{-\infty}^{+\infty} t u e^{-\frac{u^{2}}{2(1-\rho^{2})} - \frac{1}{2}t^{2}} dudt \end{bmatrix}$$

$$= \frac{\sigma_1 \sigma_2 \rho}{2\pi \sqrt{1-\rho^2}} \int_{-\infty}^{+\infty} \int_{-\infty}^{+\infty} e^{-\frac{u^2}{2(1-\rho^2)}} \cdot t^2 e^{-\frac{1}{2}t^2} dt du$$

$$=\frac{\sigma_{1}\sigma_{2}\rho}{2\pi\sqrt{1-\rho^{2}}}\int_{-\infty}^{+\infty}e^{-\frac{u^{2}}{2(1-\rho^{2})}}du\int_{-\infty}^{+\infty}t^{2}e^{-\frac{1}{2}t^{2}}dt$$

$$= \frac{\sigma_1 \sigma_2 \rho}{2\pi \sqrt{1-\rho^2}} \cdot \sqrt{2\pi \cdot (1-\rho^2)} \cdot \sqrt{2\pi} = \boldsymbol{\sigma}_1 \boldsymbol{\sigma}_2 \boldsymbol{\rho}$$

$$\rho_{XY} = \frac{\text{cov}(X,Y)}{\sqrt{D(X)}\sqrt{D(Y)}} = \frac{\sigma_1\sigma_2\rho}{\sqrt{\sigma_1^2}\sqrt{\sigma_2^2}} = \rho$$

若 $(X,Y) \sim N(\mu_1, \sigma_1^2, \mu_2, \sigma_2^2, \rho),$

则X,Y相互独立 \longrightarrow X,Y不相关

例3 设 X,Y 相互独立,且都服从 $N(0,\sigma^2)$, U = aX + bY, V = aX - bY, a,b 为常数,且都不为零,求 ρ_{UV}

$$\begin{aligned}
\mathbf{FF} & \operatorname{cov}(U, V) = E(UV) - E(U)E(V) \\
&= E(a^2 X^2 - b^2 Y^2) \\
&- [E(aX + bY)][E(aX - bY)] \\
&= a^2 E(X^2) - b^2 E(Y^2) \\
&- [aE(X) + bE(Y)][aE(X) - bE(Y)]
\end{aligned}$$

$$= a^{2}E(X^{2}) - b^{2}E(Y^{2})$$

$$- \left[a^{2}E^{2}(X) - b^{2}E^{2}(Y) \right]$$

$$= a^{2} \left[E(X^{2}) - E^{2}(X) \right] - b^{2} \left[E(Y^{2}) - E^{2}(Y) \right]$$

$$= a^{2}D(X) - b^{2}D(Y)$$

$$= (a^{2} - b^{2})\sigma^{2}$$

故
$$\rho_{UV} = \frac{Cov(U,V)}{\sqrt{D(U)} \cdot \sqrt{D(V)}}$$

$$= \frac{(a^2 - b^2)\sigma^2}{\sqrt{(a^2 + b^2)\sigma^2} \cdot \sqrt{(a^2 + b^2)\sigma^2}}$$

$$= \frac{a^2 - b^2}{a^2 + b^2}$$