

● 协方差和相关系数的性质

协方差的性质

$$\square \quad \text{cov}(X, Y) = \text{cov}(Y, X) = E(XY) - E(X)E(Y)$$

证

$$\begin{aligned}\text{cov}(X, Y) &= E[(X - EX)(Y - EY)] \\&= E[XY - X \cdot E(Y) - Y \cdot E(X) + E(X) \cdot E(Y)] \\&= E(XY) - E(X) \cdot E(Y) - E(Y) \cdot E(X) + E(X) \cdot E(Y) \\&= E(XY) - E(X)E(Y) \\&= \text{cov}(Y, X)\end{aligned}$$

$$\square \quad \text{cov}(aX, bY) = ab \text{cov}(X, Y)$$

$$\begin{aligned} \text{证} \quad \text{cov}(aX, bY) &= E[(aX - E(aX))(bY - E(bY))] \\ &= E[a(X - E(X)) \cdot b(Y - E(Y))] \\ &= abE[(X - E(X)) \cdot (Y - E(Y))] \\ &= ab \text{cov}(X, Y) \end{aligned}$$

$$\square \quad \text{cov}(X + Y, Z) = \text{cov}(X, Z) + \text{cov}(Y, Z)$$

$$\begin{aligned} \text{证} \quad \text{cov}(X + Y, Z) &= E \left\{ [(X + Y) - E(X + Y)](Z - EZ) \right\} \\ &= E \left\{ [(X - EX) + (Y - EY)](Z - EZ) \right\} \\ &= E \left[(X - EX)(Z - EZ) + (Y - EY)(Z - EZ) \right] \\ &= E \left[(X - EX)(Z - EZ) \right] + E \left[(Y - EY)(Z - EZ) \right] \\ &= \text{cov}(X, Z) + \text{cov}(Y, Z) \end{aligned}$$

$$\square \quad \text{cov}(X, X) = D(X)$$

$$\begin{aligned} \text{证} \quad \text{cov}(X, X) &= E[(X - EX)(X - EX)] \\ &= D(X) \end{aligned}$$

$$\square \quad |\text{cov}(X, Y)|^2 \leq D(X)D(Y)$$

当 $D(X) > 0, D(Y) > 0$ 时, 当且仅当

$$P(Y - E(Y) = t_0(X - E(X))) = 1$$

时, 等式成立

—Cauchy-Schwarz不等式

证 5 令

$$\begin{aligned} g(t) &= E[(X - E(X)) \cdot t - (Y - E(Y))]^2 \\ &= E[(X - E(X))^2 \cdot t^2 - 2(X - E(X))(Y - E(Y)) \cdot t + (Y - E(Y))^2] \\ &= D(X) \cdot t^2 - 2 \operatorname{cov}(X, Y) \cdot t + D(Y) \end{aligned}$$

对任何实数 t ,

$$g(t) \geq 0 \quad \longrightarrow$$

$$4 \operatorname{cov}^2(X, Y) - 4D(X)D(Y) \leq 0$$

$$\text{即} \quad |\operatorname{cov}(X, Y)|^2 \leq D(X)D(Y)$$

等号成立 $\longleftrightarrow g(t) = 0$ 有两个相等的实零点

等号成立，即 $|\text{cov}(X, Y)|^2 = D(X)D(Y)$

此时，零点为

$$\begin{aligned} t_0 &= -\frac{-2\text{cov}(X, Y)}{2D(X)} = \frac{\text{cov}(X, Y)}{D(X)} \\ &= \frac{\sqrt{D(X) \cdot D(Y)}}{D(X)} \left(\text{或} -\frac{\sqrt{D(X) \cdot D(Y)}}{D(X)} \right) \\ &= \sqrt{\frac{D(Y)}{D(X)}} \left(\text{或} -\sqrt{\frac{D(Y)}{D(X)}} \right) \end{aligned}$$

此时 ,

$$g(t_0) = 0 \quad \text{即}$$

$$\left. \begin{array}{l} E[(Y - E(Y)) - t_0(X - E(X))]^2 = 0 \\ E[(Y - E(Y)) - t_0(X - E(X))] = 0 \end{array} \right\} \text{可以证明}$$

$$\longleftrightarrow D[(Y - E(Y)) - t_0(X - E(X))] = 0$$

$$\longleftrightarrow P[(Y - E(Y)) - t_0(X - E(X)) = 0] = 1$$

$$P[(Y - E(Y)) - t_0(X - E(X)) = 0] = 1$$

即

$$P[(Y - E(Y)) = t_0(X - E(X))] = 1$$

即 Y 与 X 有线性关系的概率等于1，这种线性关系为

$$P[(Y - E(Y)) = \pm \sqrt{\frac{D(Y)}{D(X)}}(X - E(X))] = 1$$

$$P\left(\frac{Y - E(Y)}{\sqrt{D(Y)}} = \pm \frac{X - E(X)}{\sqrt{D(X)}}\right) = 1$$

标准化随机变量

设随机变量 X 的期望 $E(X)$ 、方差 $D(X)$ 都存在, 且 $D(X) \neq 0$, 则称

$$X^* = \frac{X - E(X)}{\sqrt{D(X)}}$$

为 X 的标准化随机变量. 显然,

$$E(X^*) = 0, \quad D(X^*) = 1$$

$$P(Y^* = \pm X^*) = 1$$

完全类似地可以证明

$$E^2(XY) \leq E(X^2)E(Y^2)$$

当 $E(X^2) > 0, E(Y^2) > 0$ 时, 当且仅当

$$P(Y = t_0 X) = 1$$

时, 等式成立

$$g(t) = E[Y - tX]^2$$

已知 $E(X^2) = 0$

求证 $E(X) = 0$

证明：

由 $D(X) = E(X^2) - E^2(X) \geq 0$

知 $E^2(X) \leq E(X^2)$

故当 $E(X^2) = 0$ 知 $E(X) = 0$

相关系数的性质

$$\square \quad |\rho_{XY}| \leq 1$$

证明：由 $|\text{cov}(X, Y)|^2 \leq D(X)D(Y)$

及
$$\rho_{XY} = \frac{\text{cov}(X, Y)}{\sqrt{D(X)} \cdot \sqrt{D(Y)}}$$

知
$$\rho_{XY}^2 = \frac{|\text{cov}(X, Y)|^2}{D(X)D(Y)} \leq 1$$

故
$$|\rho_{XY}| \leq 1$$

□ $|\rho_{XY}| = 1 \iff$ Cauchy-Schwarz不等式的等号成立

\iff 即 Y 与 X 有线性关系的概率等于1，
这种线性关系为

$$P\left(\frac{Y - E(Y)}{\sqrt{D(Y)}} = \pm \frac{X - E(X)}{\sqrt{D(X)}}\right) = 1$$

$$\rho_{XY} = 1 \quad \longrightarrow$$

$$t_0 = - \frac{-2 \operatorname{cov}(X, Y)}{2D(X)} = \frac{\operatorname{cov}(X, Y)}{D(X)}$$

$$= \frac{\rho \cdot \sqrt{D(X) \cdot D(Y)}}{D(X)} = \sqrt{\frac{D(Y)}{D(X)}}$$

$$P[(Y - E(Y)) = \sqrt{\frac{D(Y)}{D(X)}}(X - E(X))] = 1$$

$$P\left(\frac{Y - E(Y)}{\sqrt{D(Y)}} = \frac{X - E(X)}{\sqrt{D(X)}}\right) = 1$$

$$P(Y^* = X^*) = 1$$

$$\rho_{XY} = -1 \quad \longrightarrow$$

$$t_0 = - \frac{-2 \operatorname{cov}(X, Y)}{2D(X)} = \frac{\operatorname{cov}(X, Y)}{D(X)}$$

$$= \frac{\rho \cdot \sqrt{D(X) \cdot D(Y)}}{D(X)} = - \sqrt{\frac{D(Y)}{D(X)}}$$

$$P[(Y - E(Y)) = - \sqrt{\frac{D(Y)}{D(X)}} (X - E(X))] = 1$$

$$P\left(\frac{Y - E(Y)}{\sqrt{D(Y)}} = - \frac{X - E(X)}{\sqrt{D(X)}}\right) = 1$$

$$P(Y^* = -X^*) = 1$$

□ $\rho_{XY} = 0 \iff X, Y \text{ 不相关}$

$$\iff \text{cov}(X, Y) = 0$$

$$\iff E(XY) = E(X)E(Y)$$

$$\iff D(X \pm Y) = D(X) + D(Y)$$

X, Y 相关时 ,

$$D(X + Y) = D(X) + D(Y) + 2\text{cov}(X, Y)$$

$$D(X - Y) = D(X) + D(Y) - 2\text{cov}(X, Y)$$

X, Y 相互独立 $\begin{array}{c} \longrightarrow \\ \longleftarrow \end{array}$ X, Y 不相关

若 X, Y 服从二维正态分布 ,

X, Y 相互独立 \longleftrightarrow X, Y 不相关

例5 设 $(X, Y) \sim N(1, 4; 1, 4; 0.5)$,
 $Z = X + Y$, 求 ρ_{XZ}

解

$$\begin{aligned} E(X) &= E(Y) = 1, & \rho_{XY} &= \frac{1}{2}, \\ D(X) &= D(Y) = 4, \\ \text{cov}(X, Y) &= \rho_{XY} \cdot \sqrt{D(X)} \cdot \sqrt{D(Y)} = 2 \\ \text{cov}(X, Z) &= \text{cov}(X, X) + \text{cov}(X, Y) \\ &= D(X) + \rho_{XY} \cdot \sqrt{D(X)} \cdot \sqrt{D(Y)} \\ &= 4 + \frac{1}{2} \cdot \sqrt{4} \cdot \sqrt{4} = 6 \end{aligned}$$

$$\begin{aligned}
D(Z) &= D(X + Y) \\
&= D(X) + D(Y) + 2 \operatorname{cov}(X, Y) \\
&= D(X) + D(Y) + 2 \cdot \rho_{XY} \cdot \sqrt{D(X)} \cdot \sqrt{D(Y)}
\end{aligned}$$

$$= 4 + 4 + 2 \cdot \frac{1}{2} \sqrt{4} * \sqrt{4} = 12$$

$$\rho_{XZ} = \frac{\operatorname{cov}(X, Z)}{\sqrt{D(X)} \cdot \sqrt{D(Z)}}$$

$$= \frac{6}{2 \cdot \sqrt{12}}$$

$$= \frac{\sqrt{3}}{2}$$

例6 设随机变量 X 的概率密度函数为

$$f(x) = \frac{1}{2} e^{-|x|}, -\infty < x < +\infty$$

(1) $E(|X|), D(|X|)$

(2) 求 $\text{cov}(X, |X|)$, 问 X 与 $|X|$ 是否不相关.

(3) 问 X 与 $|X|$ 是否独立? 为什么?

解 (1)
$$\begin{aligned} E(|X|) &= \int_{-\infty}^{+\infty} |x| \frac{1}{2} e^{-|x|} dx \\ &= 2 \cdot \int_0^{+\infty} x \frac{1}{2} e^{-x} dx \\ &= 1 \end{aligned}$$

$$\begin{aligned}
 E(|X|^2) &= \int_{-\infty}^{+\infty} |x|^2 \frac{1}{2} e^{-|x|} dx \\
 &= 2 \cdot \int_0^{+\infty} x^2 \frac{1}{2} e^{-x} dx = \int_0^{+\infty} x^2 e^{-x} dx \\
 &= 2
 \end{aligned}$$

$$\begin{aligned}
 D(|X|) &= E(|X|^2) - E^2(|X|) \\
 &= 2 - 1^2 = 1
 \end{aligned}$$

$$\begin{aligned}
 (2) \quad E(X \cdot |X|) &= \int_{-\infty}^{+\infty} x \cdot |x| \frac{1}{2} e^{-|x|} dx \\
 &= \frac{1}{2} \int_{-\infty}^0 -x^2 e^x dx + \frac{1}{2} \int_0^{+\infty} x^2 e^{-x} dx \\
 &= \frac{1}{2} \int_{+\infty}^0 -y^2 e^{-y} d(-y) + \frac{1}{2} \int_0^{+\infty} x^2 e^{-x} dx \\
 &= -\frac{1}{2} \int_0^{+\infty} y^2 e^{-y} dy + \frac{1}{2} \int_0^{+\infty} x^2 e^{-x} dx \\
 &= 0
 \end{aligned}$$

$$\begin{aligned}
 E(X) &= \int_{-\infty}^{+\infty} x \frac{1}{2} e^{-|x|} dx \\
 &= \frac{1}{2} \int_{-\infty}^0 x e^x dx + \frac{1}{2} \int_0^{+\infty} x e^{-x} dx \\
 &= 0
 \end{aligned}$$

已证 : $E(|X|) = 1$

 $\text{cov}(X, |X|) = E(X|X|) - E(X)E(|X|) = 0$

X 与 $|X|$ 不相关.

$$(3) \quad P(X < -2, |X| < 1) = 0$$

$$\begin{aligned} P(X < -2) &= \int_{-\infty}^{-2} f(x) dx = \int_{-\infty}^{-2} \frac{1}{2} e^{-|x|} dx \\ &= \frac{1}{2} \int_{-\infty}^{-2} e^x dx = \frac{1}{2} e^{-2} \end{aligned}$$

$$\begin{aligned} P(|X| < 1) &= \int_{-1}^1 f(x) dx = \int_{-1}^1 \frac{1}{2} e^{-|x|} dx \\ &= 2 \cdot \frac{1}{2} \int_0^1 e^{-x} dx = 1 - e^{-1} \end{aligned}$$

显然 $P(X < -2, |X| < 1) \neq P(X < -2)P(|X| < 1)$

因而 X 与 $|X|$ 不独立