

EECS 545 Homework 1 Solution (F21)

1. Honor Code (3 pts each)

- (a) True
- (b) True
- (c) False

2. PSD matrices

- (a) (3 points)

For any \mathbf{x} we have: $\mathbf{x}^T(\mathbf{A}^T\mathbf{A})\mathbf{x} = (\mathbf{Ax})^T(\mathbf{Ax}) = \|\mathbf{Ax}\|_2^2 \geq 0$

- (b) (3 points)

For any \mathbf{z} , define $Y = (\mathbf{X} - \mathbb{E}[\mathbf{X}])^T \mathbf{z}$.

By linearity of Expectation we have:

$$\mathbf{z}^T \mathbb{E}[(\mathbf{X} - \mathbb{E}[\mathbf{X}])(\mathbf{X} - \mathbb{E}[\mathbf{X}])^T] \mathbf{z} = \mathbb{E}[\mathbf{z}^T (\mathbf{X} - \mathbb{E}[\mathbf{X}])(\mathbf{X} - \mathbb{E}[\mathbf{X}])^T \mathbf{z}] = \mathbb{E}[Y^2] \geq 0$$

- (c) (3 points)

Denote the Gram matrix as \mathbf{G} , and define \mathbf{X} as a matrix with \mathbf{x}_i as its columns. i.e:

$$\mathbf{X} = \begin{pmatrix} | & | & | & | \\ \mathbf{x}_1 & \mathbf{x}_2 & \dots & \mathbf{x}_d \\ | & | & | & | \end{pmatrix}$$

Observe then $\mathbf{G} = \mathbf{X}^T \mathbf{X}$, by (a) it's PSD.

3. Probability

- (a) (3 points)

Beginning from the right hand side, apply the definitions of conditional expectation and expectation:

$$E[E[X|Y]] = \sum_y \left(\sum_x x P(X=x|Y=y) \right) \cdot P(Y=y)$$

Next use definition of conditional probability to get joint distribution:

$$\begin{aligned} &= \sum_y \sum_x x P(X=x, Y=y) \\ &= \sum_x x \sum_y P(X=x, Y=y) \\ &= \sum_x x P(X=x) = E[X] \end{aligned}$$

Where the last step used the definition of a marginal PDF.

(b) (3 points)

Apply definition of expectation, where the indicator function is the random variable of interest:

$$\begin{aligned} E[I[X \in \mathcal{C}]] &= 1 \cdot P(X \in \mathcal{C}) + 0 \cdot P(X \notin \mathcal{C}) \\ &= P(X \in \mathcal{C}) \end{aligned}$$

(c) (3 points)

Apply definition of expectation:

$$E[XY] = \sum_{x,y} xyP(X = x, Y = y)$$

Then using independence:

$$\begin{aligned} E[XY] &= \sum_{x,y} xyP(X = x)P(Y = y) \\ &= \left(\sum_x xP(X = x) \right) \left(\sum_y yP(Y = y) \right) \\ &= E[X]E[Y] \end{aligned}$$

(d) (3 points)

For $X, Y \in \{0, 1\}$, note that $E[X] = 0 \cdot P(X = 0) + 1 \cdot P(X = 1) = P(X = 1)$. Similarly, $E[Y] = P(Y = 1)$.

Also observe that $E[XY] = \sum_{x,y \in \{0,1\}} xyP(X = x, Y = y) = P(X = 1, Y = 1)$.

Then $E[XY] = E[X]E[Y]$ implies $P(X = 1, Y = 1) = P(X = 1)P(Y = 1)$.

To finish, take complements on events, showing independence for the remaining three situations, e.g.:

$$\begin{aligned} P(X = 0, Y = 0) &= 1 - P(X = 1 \cup Y = 1) \\ &= 1 - P(X = 1) - P(Y = 1) + P(X = 1, Y = 1) \\ &= 1 - P(X = 1) - P(Y = 1) + P(X = 1)P(Y = 1) \\ &= (1 - P(X = 1))(1 - P(Y = 1)) \\ &= P(X = 0)P(Y = 0) \end{aligned}$$

It is also sufficient to explain that if two events are independent, then their complements are also independent.

4. Gaussian level sets (3 pts each)

(a) Observe \mathbf{U} is the rotation matrix that rotates vectors counter-clockwise $\frac{-\pi}{4}$ with respect to the axis (equivalently, $\frac{\pi}{4}$ rotation clockwise), with \mathbf{U}^T as its inverse.

Thus $\Sigma^{-1} = \mathbf{U}\mathbf{\Lambda}^{-1}\mathbf{U}^T$.

Define $\mathbf{x}' = \mathbf{U}^T(\mathbf{x} - \boldsymbol{\mu})$. Then in terms \mathbf{x}' , we can write:

$$1 = (\mathbf{x} - \boldsymbol{\mu})^T \Sigma^{-1} (\mathbf{x} - \boldsymbol{\mu}) = (\mathbf{x}')^T (\mathbf{\Lambda})^{-1} \mathbf{x}' = \frac{(x'_1)^2}{5} + (x'_2)^2$$

which is a Ellipse with lengths 1 and $\sqrt{5}$ for the semi-minor and semi-major axis respectively.

Thus \mathcal{C} is simply a Ellipse specified above rotated clockwise $\frac{\pi}{4}$ and translated by vector $\mu = [-2 \ 3]^T$

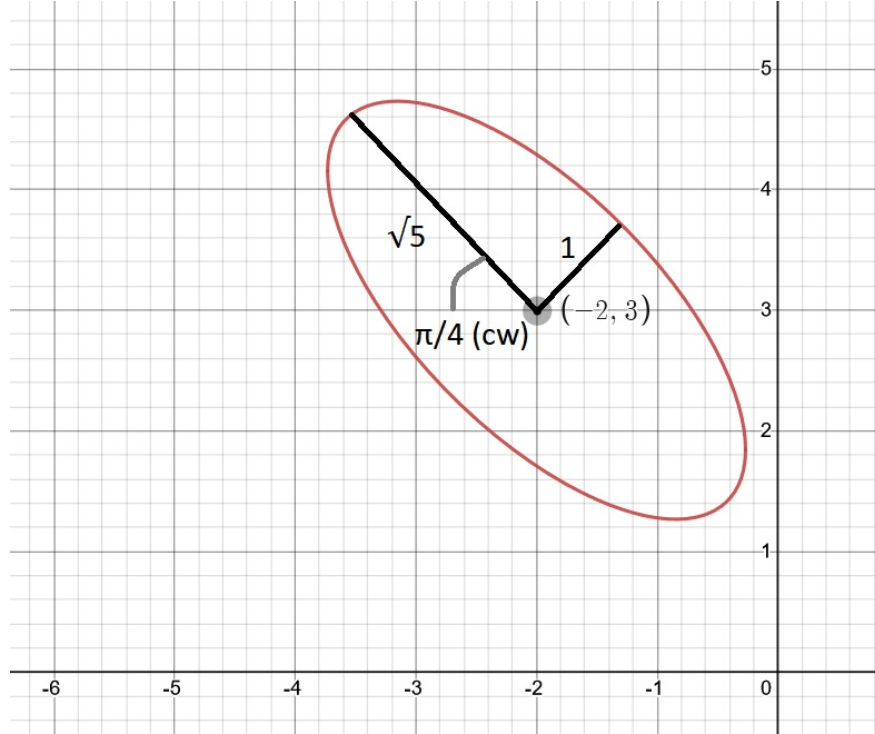


Figure 1: P3(a) plot

- (b) [Method 1] Define $\mathbf{X}' = \mathbf{U}^T(\mathbf{X} - \mu)$, then by property¹ of Gaussian distribution $\mathbf{X}' \sim \mathcal{N}(\mathbf{0}, \Sigma')$, where $\Sigma' = \mathbf{U}^T \Sigma \mathbf{U} = \Lambda$, as Λ is diagonal X'_1 and X'_2 are independent Gaussian random variables with the same mean 0 and variance 5 and 1 respectively

Define $R = (\mathbf{X}')^T \Lambda^{-1} \mathbf{X}' = \frac{(X'_1)^2}{5} + (X'_2)^2$, notice that R follows a chi-square distribution with 2 degrees of freedom².

Observe $\{\mathbf{x} \in \mathcal{C}\} = \{(\mathbf{x} - \mu)^T \Sigma^{-1} (\mathbf{x} - \mu) \leq 1\}$

Thus $\Pr(\mathbf{X} \in \mathcal{C}) = \Pr(R \leq 1) = F(1) = 0.3934$ where F denote the cdf of chi-square distribution with 2 degrees of freedom.

[Method 2] You can identify $\Sigma^{-1} = \mathbf{U} \Lambda^{-1/2} \Lambda^{-1/2} \mathbf{U}^T$, where $\Lambda^{-1/2} = \text{diag}(\frac{1}{\sqrt{5}}, 1)$. Then define $\mathbf{X}' = \mathbf{A}(\mathbf{X} - \mu)'$ where $\mathbf{A} = \Lambda^{-1/2} \mathbf{U}^T$. Invoke the linear transformation property of Gaussian distribution, we can conclude $\mathbf{X}' \sim \mathcal{N}(\mathbf{0}, \mathbf{I})$, where \mathbf{I} is the identity matrix. Hence, $R = (\mathbf{X}')^T \mathbf{X}'$ follows a Chi-square distribution with 2 degrees of freedom.

5. Unconstrained Optimization (3 points each)

- (a) Let f, g be convex functions.

¹if $\mathbf{X} \sim \mathcal{N}(\mu, \Sigma)$, for any matrix \mathbf{A} , $\mathbf{Y} = \mathbf{A}(\mathbf{x} - \mathbf{b}) \sim \mathcal{N}(\mathbf{A}(\mu - \mathbf{b}), \mathbf{A}\Sigma\mathbf{A}^T)$

²A random variable Y follows a chi-square distribution if $Y = X_1^2 + X_2^2$ for X_1, X_2 that are independent Gaussian random variables with mean 0 and variance 1

Take two arbitrary points $\mathbf{x}_1, \mathbf{x}_2 \in \mathbb{R}^d$ and $t \in (0, 1)$, let $\mathbf{x} = t\mathbf{x}_1 + (1 - t)\mathbf{x}_2$, then by convexity of f and g we have:

$$f(\mathbf{x}) = f(t\mathbf{x}_1 + (1 - t)\mathbf{x}_2) \leq tf(\mathbf{x}_1) + (1 - t)f(\mathbf{x}_2) \quad (1)$$

$$g(\mathbf{x}) = g(t\mathbf{x}_1 + (1 - t)\mathbf{x}_2) \leq tg(\mathbf{x}_1) + (1 - t)g(\mathbf{x}_2) \quad (2)$$

sum Eq(1) and Eq(2) we have:

$$\begin{aligned} (f + g)(\mathbf{x}) &= f(\mathbf{x}) + g(\mathbf{x}) \\ &\leq tf(\mathbf{x}_1) + tg(\mathbf{x}_1) + (1 - t)f(\mathbf{x}_2) + (1 - t)g(\mathbf{x}_2) \\ &= t(f + g)(\mathbf{x}_1) + (1 - t)(f + g)(\mathbf{x}_2) \end{aligned}$$

- (b) Let f, g be convex twice continuously differentiable functions. Then by the linearity of differentiation the Hessian of $f + g$ is given by

$$\nabla^2(f + g) = \nabla^2 f + \nabla^2 g.$$

By Property 7 from the notes on unconstrained optimization, $\nabla^2 f, \nabla^2 g$ are PSD. Since the sum of two PSD matrices is clearly PSD, the Hessian of $f + g$ is PSD so $f + g$ is convex (also by property 7).

- (c) Let $\mathbf{x} = [x_1, \dots, x_d]^T$, A_{ij} denote the (i,j)-th entry of matrix \mathbf{A} , and b_i denote the i-th entry of \mathbf{b} then the quadratic function $f(\mathbf{x})$ can be written explicitly as:

$$f(\mathbf{x}) = \frac{1}{2}\mathbf{x}^T \mathbf{A} \mathbf{x} + \mathbf{b}^T \mathbf{x} + c = \frac{1}{2} \sum_{i=1}^d \sum_{j=1}^d A_{ij} x_i x_j + \sum_{i=1}^d b_i x_i + c.$$

Applying the definition of the Hessian matrix, the (k, ℓ) -th entry of $\nabla^2 f(\mathbf{x})$ is given by:

$$\begin{aligned} [\nabla^2 f(\mathbf{x})]_{k,\ell} &= \frac{\partial^2 f(\mathbf{x})}{\partial x_k \partial x_\ell} \\ &= \frac{\partial^2}{\partial x_k \partial x_\ell} \left\{ \frac{1}{2} \sum_{i=1}^d \sum_{j=1}^d A_{ij} x_i x_j + \sum_{i=1}^d b_i x_i + c \right\} \\ &= \frac{\partial}{\partial x_k} \left\{ \frac{1}{2} \sum_{i=1}^d \sum_{j=1}^d A_{ij} \frac{\partial}{\partial x_\ell} x_i x_j + \sum_{i=1}^d \frac{\partial}{\partial x_\ell} b_i x_i \right\} \\ &= \frac{\partial}{\partial x_k} \left\{ \sum_{i=1}^d A_{i\ell} x_i + b_\ell \right\} \\ &= A_{k\ell}, \end{aligned}$$

thus the Hessian of f is \mathbf{A} . The function f is convex when \mathbf{A} is positive semi-definite, and strictly convex if \mathbf{A} is positive definite.