## Logistic Regression (and other nonlinear models)

**EDS 222** 

Tamma Carleton Fall 2022

• Assignment 03 pass/fail, due **today** (9am)

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- Assignment 04 after we cover inference/uncertainty (likely assigned next week)

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- Final project proposals, due 11/10 (9am)
  - More details in a few slides

### Goal:

Apply **some of** the statistical concepts you have learned in this course to **answer an environmental data science question**.\*

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### Two parts:

Deliverable 1: Technical blog post. Some examples:

- G-FEED
- emLab
- MEDS '22, ex. 1
- MEDS '22, ex. 2
- MEDS '22, ex. 3

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### Two parts:

Deliverable 2: Three-minute in-class presentation during final exam slot (8-11am, 12/6)

<sup>[\*]:</sup> Your project must include concepts from the second half of the course.

### Proposal:

Short paragraph (4-5 sentences) describing your proposed project. Motivate the question, describe possible data sources, suggest possible analyses.

**Email Sandy your proposal** at sandysum@ucsb.edu by 9am on November 10th.

Full guidelines on our Resources Page

### Some example topics:

 Are political views on climate change associated with recent natural disaster exposure?

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## Today

More on nonlinear relationships with linear regression models

Log-linear, log-log regressions

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Logistic regression

How do we model binary outcomes?

# Nonlinear relationships in linear regression models

### Nonlinear transformations

- Our linearity assumption requires that **parameters enter linearly** (*i.e.*, the  $\beta_k$  multiplied by variables)
- We allow nonlinear relationships between y and the explanatory variables x.

#### **Example: Polynomials**

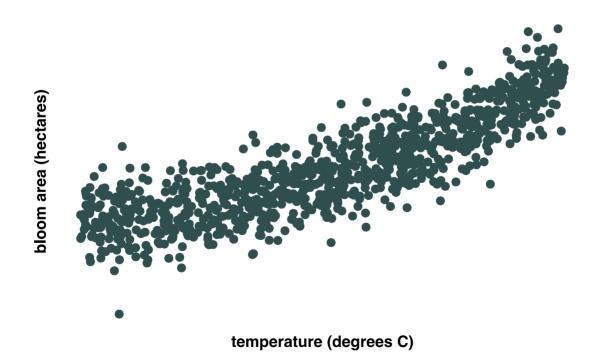
$$y_i = eta_0 + eta_1 x_i + eta_2 x_i^2 + u_i$$
  $y_i = eta_0 + eta_1 x_i + eta_2 x_i^2 + eta_3 x_i^3 + u_i$   $y_i = eta_0 + eta_1 x_i + eta_2 x_i^2 + eta_3 x_i^3 + eta_4 x_i^4 + u_i$ 

...

### Polynomials

 Recall the relationship between temperature and harmful algal blooms:

$$area_i = eta_0 + eta_1 temperature_i + eta_2 temperature_i^2 + u_i$$



### Polynomials

### Estimating polynomial regressions in R:

```
blooms df = blooms df %>% mutate(temp2 = temp^2)
summary(lm(area~temp+temp2, data=blooms_df))
#>
#> Call:
#> lm(formula = area ~ temp + temp2, data = blooms df)
#>
#> Residuals:
#> Min 1Q Median 3Q Max
#> -12.597 -2.092 -0.142 1.995 9.487
#>
#> Coefficients:
    Estimate Std. Error t value Pr(>|t|)
#>
#> (Intercept) 0.0636 0.2925 0.22 0.83
#> temp 0.6254 0.4401 1.42 0.16
#> ---
#> Signif. codes: 0 '***' 0.001 '**' 0.01 '*' 0.05 '.' 0.1 ' ' 1
#>
#> Residual standard error: 3.02 on 997 degrees of freedom
#> Multiple R-squared: 0.777, Adjusted R-squared: 0.777
```

### Other nonlinear-in-X regressions

• Polynomials and interactions:

$$y_i=eta_0+eta_1x_{1i}+eta_2x_{1i}^2+eta_3x_{2i}+eta_4x_{2i}^2+eta_5\left(x_{1i}x_{2i}
ight)+u_i$$
 (more on this today)

- Exponentials  $\log(y_i) = eta_0 + eta_2 e^{x_{2i}} + u_i$
- **Logs:**  $\log(y_i) = \beta_0 + \beta_1 x_{1i} + u_i$  (Today!)
- Indicators and thresholds:  $y_i = eta_0 + eta_1 x_{1i} + eta_2 \, \mathbb{I}(x_{1i} \geq 100) + u_i$

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In all cases, the effect of a change in x on y will vary depending on your baseline level of x. This is not true with linear relationships!

You will frequently see logged\* outcome variables with linear (non-logged) explanatory variables, *e.g.*,

$$\log(\text{Pay}_i) = \beta_0 + \beta_1 \operatorname{School}_i + u_i$$

This specification changes our interpretation of the slope coefficients.

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### Interpretation

- A one-unit increase in our explanatory variable increases the outcome variable by approximately  $eta_1 imes 100$  percent.
- Example: If  $eta_1=0.03$ , an additional year of schooling increases pay by approximately 3 percent.

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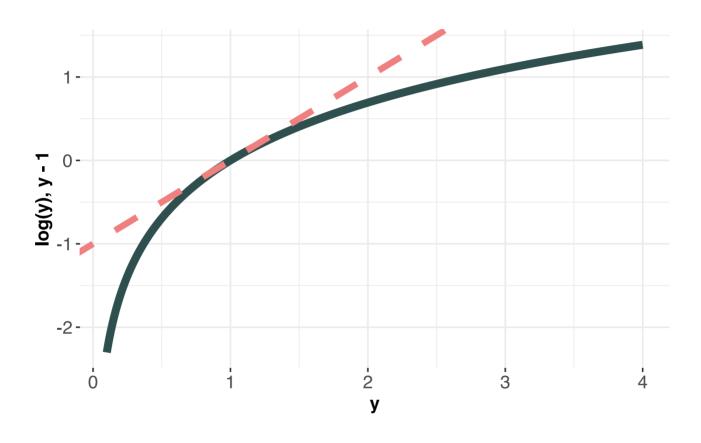
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$$0.12 = \frac{5.6 - 5}{5}$$

Generally, we have that when y increases by r percent, our new value is y(1+r).

$$r=rac{y_2-y_1}{y_1}$$

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This lets us show that:

$$log(y(1+r)) = log(y) + log(1+r) pprox log(y) + r$$

So when we see log(y) go up by r, we can say that represents an r imes 100 percent change in y!

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For example: y is increased by 5% means y increases to y(1.05). The log of y changes from log(y) to approximately log(y) + 0.05. Increasing y by 5% is therefore (almost) equivalent to adding 0.05 to log(y).

Back to our log-linear model

$$\log(y_i) = \beta_0 + \beta_1 \, x_i + u$$

A one unit change in x causes a  $\beta_1$  unit change in log(y).

This is equivalent to a  $\beta_1$  percentage change in y.

Because the log-linear specification comes with a different interpretation, you need to make sure it fits your data-generating process/model.

Does x change y in levels (e.g., a 3-unit increase) or percentages (e.g., a 10-percent increase)?

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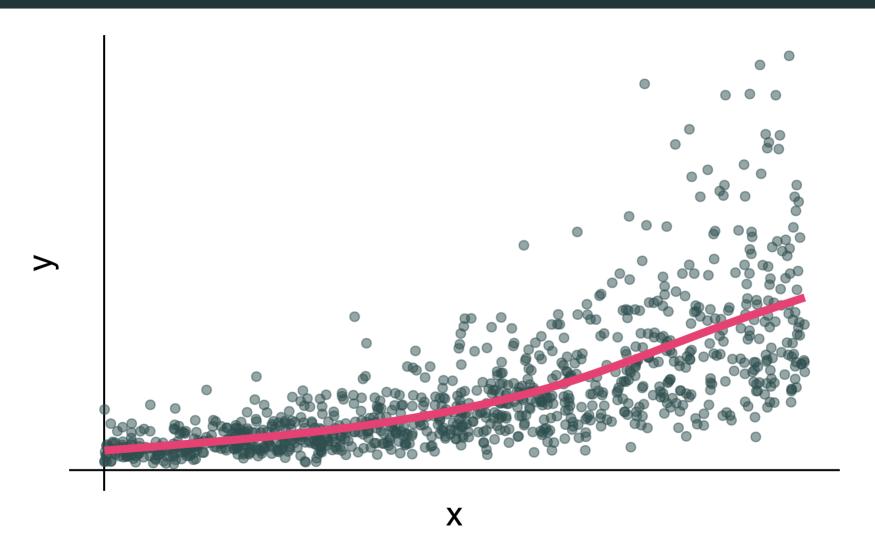
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*I.e.*, you need to be sure an exponential relationship makes sense:

$$\log(y_i) = eta_0 + eta_1\,x_i + u_i \iff y_i = e^{eta_0 + eta_1x_i + u_i}$$

Note: You are using linear regression to estimate a nonlinear-in-parameters relationship. This is the power of taking logs!

# Log-linear specification



## Log-log specification

Similarly, log-log models are those where the outcome variable is logged and at least one explanatory variable is logged

$$\log(\text{Pay}_i) = \beta_0 + \beta_1 \, \log(\text{School}_i) + u_i$$

#### **Interpretation:**

- A one-percent increase in x will lead to a  $\beta_1$  percent change in y.
- Often interpreted as an "elasticity" in economics.

## Log-linear with a binary variable

**Note:** If you have a log-linear model with a binary indicator variable, the interpretation for the coefficient on that variable changes.

Consider:

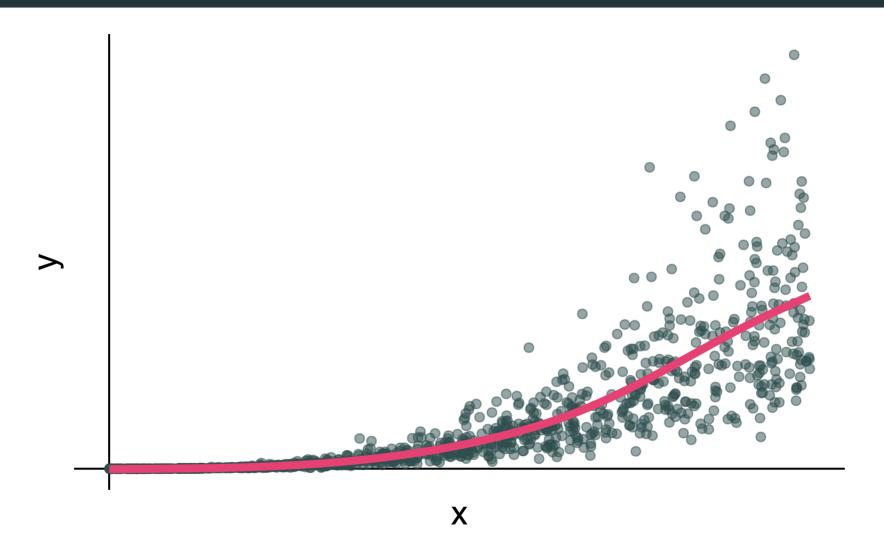
$$\log(y_i) = eta_0 + eta_1 x_{1i} + u_i$$

for binary variable  $x_1$ .

The interpretation of  $\beta_1$  is now

- ullet When  $x_1$  changes from 0 to 1, y will change by  $100 imes \left(e^{eta_1}-1
  ight)$  percent.
- ullet When  $x_1$  changes from 1 to 0, y will change by  $100 imes (e^{-eta_1}-1)$  percent.

# Log-log specification



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Exact percentage change (use for large predicted changes):

If  $log(y) = \beta_0 + \beta_1 x + \varepsilon$ , then the percentage change in y for a one unit change in x is:

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Note that  $e^x$  in R is exp(x)

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- $(e^{0.6}-1) imes 100=0.82 imes 100 \implies 82$  change in y
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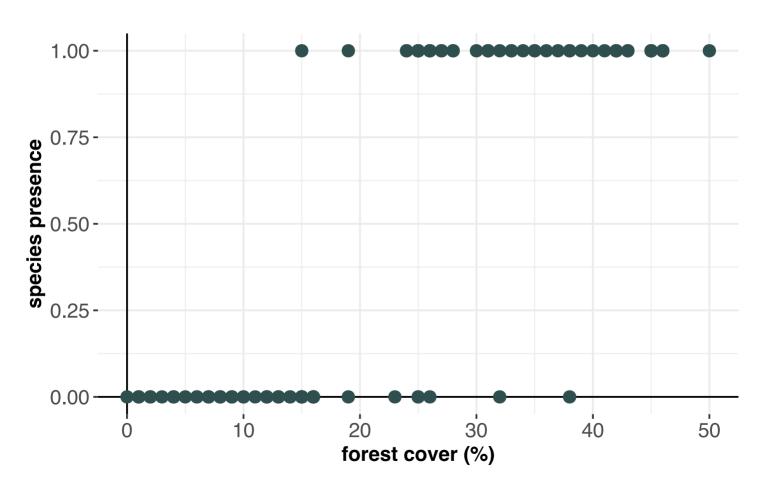
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Can you just change units of x?

- Yes, mechanically you can do this and avoid the issues with approximation
- ullet But think hard about your problem! You probably care about understanding the impacts of a meaningful increase in x, not a tiny increase in x

## Modeling binary outcomes

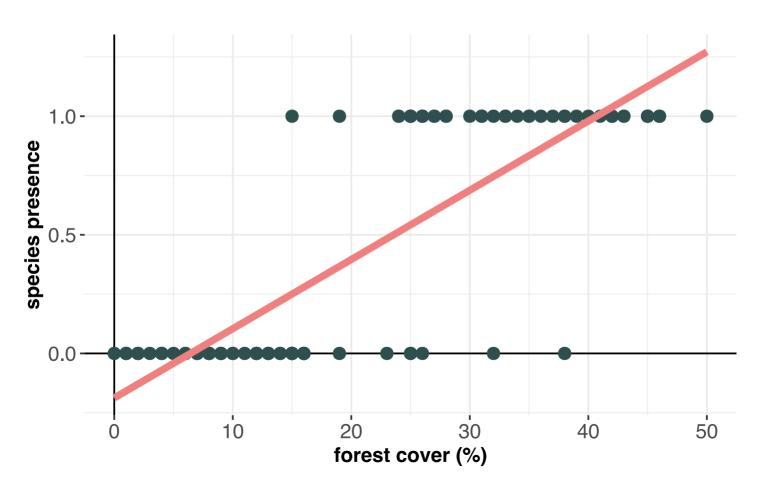
What do you do when your dependent variable takes on just two values?



## Modeling binary outcomes

What's wrong with running our standard linear regression?

$$\mathrm{species} \ \mathrm{present}_i = \beta_0 + \beta_1 \mathrm{forest} \ \mathrm{cover}_i + \varepsilon_i$$



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- Basic idea: we need some transformation of the *probability* that lets us write:

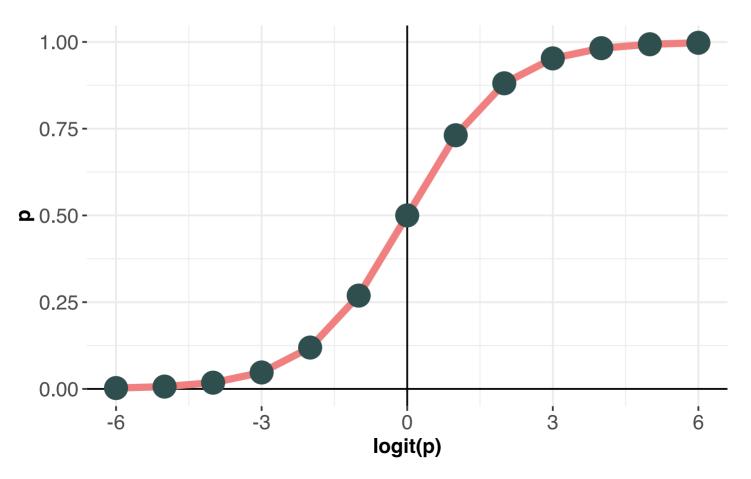
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- We want this transformation to ensure that:
  - we can input a value between 0 and 1 and return a continuous variable (i.e., we want our outcome variable to be a continuous variable)
  - $\circ$  our predicted probabilities  $\hat{p}_i$  (inverse of the transformation) will fall between 0 and 1

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We can then write:

$$log\left(rac{p_i}{1-p_i}
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The logit function is also called "log odds" because the "odds ratio" is the probability of success  $p_i$  divided by  $1-p_i$ 

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Because of the properties of the logit function (see last graph), this ensures we will generate predicted probabilities  $\hat{p}_i$  that fall between 0 and 1.

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- ullet Can't use linear regression -- we don't have data on  $p_i!$  We only see  $y_i=1$  or  $y_i=0$
- We use what's called "maximum likelihood estimation" (alternatively, can use gradient descent)
  - Essentially, this asks: what combination of parameters  $\beta_0, \beta_1, \dots$  maximizes the likelihood that we would observe the data we have?
  - $\circ$  E.g., if you have high  $x_1$  values coinciding with many  $y_i=1$  values, likely that  $eta_1$  is high and that  $p_i$  is high for observations with large  $x_1$

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All you really need to know on estimation is...

- That we use glm() instead of lm() -- GLM for "generalized linear model"
- Interpreting coefficients is a lot more complicated! (next slide)

$$log\left(rac{p_i}{1-p_i}
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We need to transform our output to get predicted probabilities back!

$$egin{aligned} \logigg(rac{p_i}{1-p_i}igg) &= b_0 + b_1x_{1,i} + \cdots + b_kx_{k,i} \ rac{p_i}{1-p_i} &= e^{b_0 + b_1x_{1,i} + \cdots + b_kx_{k,i}} \ p_i &= (1-p_i)\,e^{b_0 + b_1x_{1,i} + \cdots + b_kx_{k,i}} \ p_i &= e^{b_0 + b_1x_{1,i} + \cdots + b_kx_{k,i}} - p_i imes e^{b_0 + b_1x_{1,i} + \cdots + b_kx_{k,i}} \ p_i + p_i\,e^{b_0 + b_1x_{1,i} + \cdots + b_kx_{k,i}} &= e^{b_0 + b_1x_{1,i} + \cdots + b_kx_{k,i}} \ p_i(1+e^{b_0 + b_1x_{1,i} + \cdots + b_kx_{k,i}}) &= e^{b_0 + b_1x_{1,i} + \cdots + b_kx_{k,i}} \ p_i &= rac{e^{b_0 + b_1x_{1,i} + \cdots + b_kx_{k,i}}}{1+e^{b_0 + b_1x_{1,i} + \cdots + b_kx_{k,i}}} \end{aligned}$$

This means that if you run a regression with many independent variables, you need to plug your estimated  $\hat{\beta}$ 's and the values of all your x variables into this equation to get back a predicted probability for any individual:

$$p_i = rac{e^{b_0 + b_1 x_{1,i} + \cdots + b_k x_{k,i}}}{1 + e^{b_0 + b_1 x_{1,i} + \cdots + b_k x_{k,i}}}$$

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If you want to know the *effect* of changing just one variable  $x_j$  on the probability  $p_i$ , you need to compute:

$$p_i(x_j+1) - p_i(x_j) = rac{e^{b_0 + \cdots + b_j(x_{j,i}+1) + \cdots + b_k x_{k,i}}}{1 + e^{b_0 + \cdots + b_j(x_{j,i}+1) + \cdots + b_k x_{k,i}}} - rac{e^{b_0 + \cdots + b_j x_{j,i} + \cdots + b_k x_{k,i}}}{1 + e^{b_0 + \cdots + b_j x_{j,i} + \cdots + b_k x_{k,i}}}$$

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**Note** that this calculation depends on all the other x's! And it will vary with the baseline level of  $x_i$ 

- Bertrand and Mullainathan (2003) study discrimination in hiring decisions
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- Authors created many fake resumes, randomly assigning different characteristics (name, sex, race, experience, honors, etc.)
- Outcome variable is binary: Did the resume get a call back from a (real) potential employer?
  - $\circ$  Yes:  $y_i=1$
  - $\circ$  No:  $y_i=0$
- Manipulated first names to be those that are commonly associated with White or Black individuals
- Random study design allows estimation of the causal effect of race on callback probability

List of all 36 unique names along with the commonly inferred race and sex associated with these names.

first_name	race	sex	first_name	race	sex	first_name	race	sex
Aisha	Black	female	Hakim	Black	male	Laurie	White	female
Allison	White	female	Jamal	Black	male	Leroy	Black	male
Anne	White	female	Jay	White	male	Matthew	White	male
Brad	White	male	Jermaine	Black	male	Meredith	White	female
Brendan	White	male	Jill	White	female	Neil	White	male
Brett	White	male	Kareem	Black	male	Rasheed	Black	male
Carrie	White	female	Keisha	Black	female	Sarah	White	female
Darnell	Black	male	Kenya	Black	female	Tamika	Black	female
Ebony	Black	female	Kristen	White	female	Tanisha	Black	female

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Variables included in the data (all randomly assigned):

variable	description
received_callback	Specifies whether the employer called the applicant following submission of the application for the job.
job_city	City where the job was located: Boston or Chicago.
college_degree	An indicator for whether the resume listed a college degree.
years_experience	Number of years of experience listed on the resume.
honors	Indicator for the resume listing some sort of honors, e.g. employee of the month.

Variables included in the data (all randomly assigned):

variable	description
military	Indicator for if the resume listed any military experience.
has_email_address	Indicator for if the resume listed an email address for the applicant.
race	Race of the applicant, implied by their first name listed on the resume.
sex	Sex of the applicant (limited to only and in this study), implied by the first name listed on the resume.

- First, we estimate a single predictor: race
- race indicates whether the applicant is White or not (**Note:** race is also binary in this case!)
- We find:

$$\log\!\left(rac{\hat{p}_i}{1-\hat{p}_i}
ight) = -2.67 + 0.44 imes exttt{race_White}$$

a. If a resume is randomly selected from the study and it has a Black associated name, what is the probability it resulted in a callback?

b. What would the probability be if the resume name was associated with White individuals?

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a. If a resume is randomly selected from the study and it has a Black associated name, what is the probability it resulted in a callback?

**Answer:** If a randomly chosen resume is associated with a Black name, then race\_white takes the value of 0 and the right side of the model equation equals -2.67. Solving for  $p_i$  gives

$$log(rac{\hat{p}_i}{1-\hat{p}_i}) = -2.67 \implies \hat{p}_i = rac{e^{-2.67}}{1+e^{-2.67}} = 0.065.$$

$$\log\!\left(rac{\hat{p}_i}{1-\hat{p}_i}
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b. What would the probability be if the resume name was associated with White individuals?

**Answer:** If the resume had a name associated with White individuals, then the right side of the model equation is  $-2.67+0.44\times 1=-2.23$ . This translates into  $\hat{p}_i=0.097$ .

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**Answer:** If the resume had a name associated with White individuals, then the right side of the model equation is  $-2.67 + 0.44 \times 1 = -2.23$ . This translates into  $\hat{p}_i = 0.097$ .

**Conclude:** Being White increases the likelihood of a call back, by 3.2 percentage points.

**Use the same process** to compute predicted probabilities with multiple independent variables, you just more calculations!

**Use the same process** to compute predicted probabilities with multiple independent variables, you just more calculations!

For example, you might estimate:

$$egin{aligned} log\left(rac{p}{1-p}
ight) \ &= -2.7162 - 0.4364 imes exttt{job\_city}_{ exttt{Chicago}} \ &+ 0.0206 imes exttt{years\_experience} \ &+ 0.7634 imes exttt{honors} - 0.3443 imes exttt{military} + 0.2221 imes exttt{email} \ &+ 0.4429 imes exttt{race}_{ exttt{White}} - 0.1959 imes exttt{sex}_{ exttt{man}} \end{aligned}$$

To predict callback probability for a White individual, you also need to know job location, experience, honors, military experience, whether they have an email, race, and sex!

For example, you might estimate:

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Note: the effect of race on call back now varies based on all the other covariates!

• Try it: Effect of being white for Chicago male with 10 years experience, an email, no honors and no military experience *versus* a female with the same characteristics?

# Multinomial logistic regression

What if your outcome variable is categorical, not binary?

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#### For example:

- Species
- Socioeconomic status
- Survey responses
- ..

## Multinomial logistic regression

**What if** your outcome variable is categorical, not binary?

#### For example:

- Species
- Socioeconomic status
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- ..

**Multinomial logistic regression** generalizes the binary logistic regression you've seen here to work for multiple outcome categories

- Model predicts the probability an individual will fall into each category
- Beyond the scope of this class, but not a far leap from what you've seen here (lots of online resources -- ask me if you're interested!)

Slides created via the R package xaringan.

Some slide components were borrowed from Ed Rubin's awesome course materials.