

# Continuous Distributions

Uniform, Normal, Exponential, and More

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Gov 2001 · Harvard University

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# Where We Are

From discrete counts to continuous measurements

**Monday:** Discrete distributions

- Bernoulli, Binomial, Poisson
- Counting successes and rare events

**Today:** Continuous distributions — the **A-list** stars of statistics

- **Uniform** — equal probability, foundation for simulation
- **Normal** — the star of the show (CLT, regression, everything)
- **Exponential** — waiting times, survival analysis
- **Chi-square,  $t$**  — the B-list supporting cast for inference

Reading: Aronow & Miller §1.4–1.5, Blackwell 2.4–2.5

# Recall My Pedagogy

How I approach each topic

**My approach, usually in this order:**

1. **History & narrative** — Who discovered it? What problem were they solving?
2. **Political science application** — Where does this show up in our world?
3. **Visuals** — What does it look like? Build intuition graphically
4. **Technical rigor** — The math that makes it precise

**The “Netflix–Matt Damon” principle:**

*Netflix worked with Matt Damon on writing movies for their platform. Their advice: “Remind audiences of the plot regularly.” Their data showed viewers constantly pausing, multitasking, distracted — they needed more help than a theater audience.*

I’ll do the same. Even for concepts from weeks ago, I’ll remind you of the plot. **This isn’t remedial — it’s respect for how we actually learn** when juggling too much.

# The Pattern Continues: Parameters Define Everything

Same roadmap as discrete, now with PDFs instead of PMFs

## Monday's pattern for discrete distributions:

$$\text{Parameters} \rightarrow \text{PMF} \rightarrow \mathbb{E}[X], \text{Var}[X]$$

## Today's pattern for continuous distributions:

$$\text{Parameters} \rightarrow \text{PDF} \rightarrow \text{CDF} \rightarrow \mathbb{E}[X], \text{Var}[X]$$

## For each distribution today, we'll identify:

1. **Parameters:** What do you need to specify?  $(\mu, \sigma^2, \lambda, a, b)$
2. **PDF:** The density function  $f(x)$
3. **CDF:** The cumulative distribution  $F(x) = \mathbb{P}(X \leq x)$
4. **Moments:** Expected value and variance

Once you know the parameters, everything else follows.

## Wait — Why Is It Called $f(x)$ for Both PMF and PDF?

Same notation, different meanings

**Historical convention:** Both are “the function that characterizes the distribution,” so mathematicians use the same letter. But they work differently:

	PMF (discrete)	PDF (continuous)
$f(x)$ means	$\mathbb{P}(X = x)$	density at $x$
What it is	An actual probability	<b>Not</b> a probability
Values	Always in $[0, 1]$	Can be <i>any</i> $\geq 0$ (even $> 1$ !)
To get $\mathbb{P}$	Read $f(x)$ directly	Must integrate: $\int_a^b f(x) dx$

**The key difference:** For continuous  $X$ ,  $\mathbb{P}(X = x) = 0$  for any specific  $x$ . Probability only exists over *intervals*.

## Density Means “Probability Per Unit Length”

Why  $f(x)$  can exceed 1

Think of  $f(x)$  as **probability concentration** at point  $x$ .

**Example:** If  $X \sim \text{Uniform}(0, 0.1)$ , then:

$$f(x) = \frac{1}{0.1} = 10 \quad \text{for } x \in [0, 0.1]$$

That's a density of 10 — but the total probability is still:

$$\int_0^{0.1} 10 \, dx = 10 \times 0.1 = 1 \quad \checkmark$$

**Intuition:** High density means probability is *concentrated*. Low density means probability is *spread out*.

Exponential with large  $\lambda$  has  $f(0)$  very large — probability is concentrated near zero, not “more than 100% likely.”

# Most of Your Statistical Life Will Be Normal

The A-list and B-list of continuous distributions

**A-list actors** — you'll model data with these:

- **Normal** — regression errors, polling, heights, test scores
- **Exponential** — waiting times, survival, duration models
- **Uniform** — simulation, randomization, probability foundations

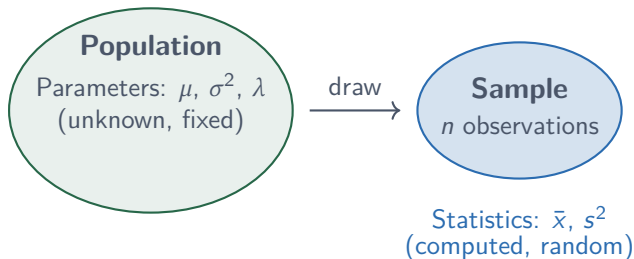
**B-list actors** — supporting roles for inference:

- **Chi-square** — variance estimation, goodness-of-fit
- ***t*-distribution** — hypothesis testing with estimated variance

The B-list are *derived* from Normal. You don't model data with them—you use them for inference.

# Statistics Is Learning About Populations from Samples

The central distinction



**The game:** Use sample statistics to *estimate* population parameters.

Today's distributions describe *populations*. Estimation theory tells us how sample statistics behave.



# Reminder: Why We Care About Distributions

The link to inference

## The roadmap:

1. **Population**: Described by a distribution with unknown parameters
2. **Sample**: Data we observe (drawn from the population)
3. **Estimation**: Use data to learn about parameters ( $\mu$ ,  $\sigma^2$ ,  $\lambda$ )
4. **Uncertainty**: Quantified via sampling distributions (which we derive from population distributions)

## Today's distributions matter because:

- The **Normal** is the sampling distribution of the mean (CLT)
- The **Chi-square** appears when estimating variance
- The ***t*-distribution** is what we use for hypothesis tests with estimated variance

Everything connects. Today we're building the vocabulary you'll use for inference.

# The Uniform Distribution

Simple but Fundamental

## Notation Warning: Parameters Mean Different Things

Another inconsistency I didn't invent

Compare these two “standard” distributions:

Distribution	Notation	What the numbers mean
Standard Normal	$Z \sim N(0, 1)$	$N(\text{mean}, \text{variance})$
“Standard” Uniform	$U \sim \text{Uniform}(0, 1)$	$\text{Uniform}(\text{lower}, \text{upper})$

For Normal: the parameters *are* the mean and variance.

For Uniform: the parameters are the *interval endpoints* — you derive the moments:

$$\mathbb{E}[U] = \frac{0 + 1}{2} = 0.5, \quad \text{Var}[U] = \frac{(1 - 0)^2}{12} = \frac{1}{12}$$

Like PMF vs PDF using the same  $f(x)$ : inconsistent notation that evolved historically.  
Just be aware.

## Why the Difference? Distributions Parameterize Different Concepts

Moments vs. bounds vs. rates

**The deeper point:** Every distribution needs parameters, but what those parameters *represent* varies:

Distribution	Parameters	What they represent
$X \sim N(\mu, \sigma^2)$	$\mu, \sigma^2$	Moments (mean, variance)
$X \sim \text{Uniform}(a, b)$	$a, b$	Support bounds (endpoints)
$X \sim \text{Exponential}(\lambda)$	$\lambda$	Rate (events per unit time)

**The general Normal:**  $X \sim N(\mu, \sigma^2)$

- $\mu$  = population mean (can be any real number)
- $\sigma^2$  = population variance (must be positive)
- The “standard” Normal  $N(0, 1)$  is just the special case  $\mu = 0, \sigma^2 = 1$

We'll use  $\mu$  and  $\sigma^2$  throughout. When you see specific numbers, you're seeing a specific distribution from the family.

# The Simplest Distribution: Equal Probability Everywhere

Political example: When does the voter arrive?

**Example:** A voter arrives at a polling station sometime between 8am and 8pm. If arrivals are “uniformly distributed,” any moment is equally likely.

**Definition:**  $X \sim \text{Uniform}(a, b)$  has PDF:

$$f(x) = \frac{1}{b-a} \quad \text{for } x \in [a, b]$$



**Key formulas:**  $\mathbb{E}[X] = \frac{a+b}{2}$  (midpoint),  $\text{Var}[X] = \frac{(b-a)^2}{12}$

## The Standard Uniform: $U \sim \text{Uniform}(0, 1)$

The building block for everything

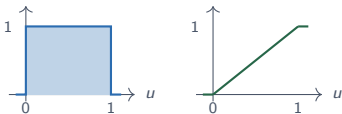
The **standard uniform**  $U \sim \text{Uniform}(0, 1)$  is special:

- PDF:  $f(u) = 1$  for  $u \in [0, 1]$  (a flat horizontal line at height 1)
- CDF:  $F(u) = u$  for  $u \in [0, 1]$  (output = input, so  $y = x$ , a diagonal line)
- $\mathbb{E}[U] = 0.5$ ,  $\text{Var}[U] = 1/12$

### Why is it fundamental?

- Every random number generator starts with  $\text{Uniform}(0, 1)$
- Randomization in experiments: “treat if  $U < 0.5$ ”

PDF:  $f(u) = 1$     CDF:  $F(u) = u$



# What Does the CDF Actually Tell You?

Reading the plot

$F(u) = \mathbb{P}(U \leq u)$  = “the probability of getting a value at or below  $u$ ”

**For Uniform(0,1)**,  $F(u) = u$  means:

- $F(0.3) = 0.3 \rightarrow$  “30% chance of being below 0.3”
- $F(0.7) = 0.7 \rightarrow$  “70% chance of being below 0.7”

**The slope of the CDF tells you where outcomes concentrate:**

- **Steep slope** = probability accumulating fast  $\rightarrow$  outcomes cluster there
- **Flat slope** = probability not accumulating  $\rightarrow$  outcomes rare there
- **Constant slope** (diagonal line) = probability accumulates evenly everywhere

**A diagonal CDF is the visual definition of “uniform”** — no region is more likely than any other.

# Support Tells You Where Outcomes Are Possible

You worked with this on Problem Set 2

**Definition:** The **support** of a random variable is where its PDF is positive:

$$\text{Supp}[X] = \{x : f(x) > 0\}$$

**Three types of support we'll see today:**

Distribution	Support	Type
Uniform( $a, b$ )	$[a, b]$	Bounded (finite interval)
Normal( $\mu, \sigma^2$ )	$(-\infty, +\infty)$	Unbounded (whole line)
Exponential( $\lambda$ )	$[0, +\infty)$	Half-line (non-negative)

PS2 Q2 asked you to find support. This concept matters for specifying models correctly.



# Have You Ever Wondered How R Generates Random Numbers?

Where do Bernoulli, Binomial, Poisson, Normal come from?

When you type `rbinom(1, 10, 0.5)` in R, what's actually happening?

The answer: **Everything comes from the Uniform.**

- The **Uniform** is the raw randomness — the stochastic part
- The **parameters** ( $p$ ,  $n$ ,  $\lambda$ ) shape that randomness into the distribution you want

**You choose the parameters, R transforms the Uniform:**

<code>rbinom(1, n, p)</code>	Uniform + cutpoint $p$ , repeated $n$ times $\rightarrow$ Binomial
<code>rpois(1, lambda)</code>	Uniform + staircase shaped by $\lambda \rightarrow$ Poisson
<code>rexp(1, lambda)</code>	Uniform + inverse CDF shaped by $\lambda \rightarrow$ Exponential

The parameters are the recipe. The Uniform is the ingredient.

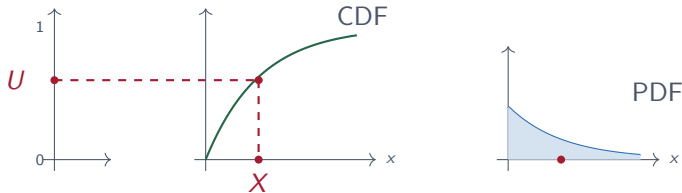
**Next:** Let's see how this works under the hood.

# Generating Any Distribution: Start With the Picture

The inverse CDF method, visualized

**Goal:** You want to simulate  $X$  from some distribution  $F$ . All you have is  $U \sim \text{Uniform}(0, 1)$ .

1. Draw  $U$
2. Find where  $U$  hits CDF
3. That's your draw!



**The recipe:** Draw  $U \sim \text{Uniform}(0, 1)$  on the y-axis  $\rightarrow$  trace horizontally to CDF  $\rightarrow$  drop down to x-axis  $\rightarrow$  that's  $X = F^{-1}(U)$

This works because the CDF maps outcomes to  $[0, 1]$ , so the inverse maps  $[0, 1]$  back to outcomes.

## Why Does This Produce the Right Distribution?

The CDF slope determines where outcomes land

**Key insight:** The shape of the CDF controls the transformation.

- **Where CDF is steep:** Many different  $U$  values map to a *narrow* range of  $X$   
→ Outcomes **cluster** there (high density)
- **Where CDF is flat:** Few  $U$  values map to that region of  $X$   
→ Outcomes are **rare** there (low density)
- **Uniform's 45° line:** No stretching or compression — equal in, equal out  
→ This is why Uniform is the “undistorted” starting point

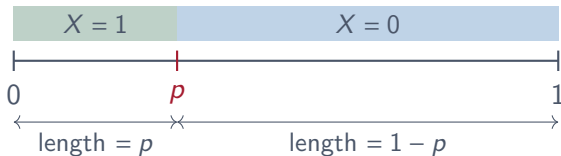
**The inverse CDF “stretches” the Uniform** to match whatever density you want.

The Uniform is like a blank canvas. The inverse CDF is the paint.

## Universality for Discrete: Bernoulli

The simplest case — one cutpoint

**Goal:** Generate  $X \sim \text{Bernoulli}(p)$  from  $U \sim \text{Uniform}(0, 1)$



### Algorithm:

1. Draw  $U \sim \text{Uniform}(0, 1)$  — a random point on  $[0, 1]$
2. If  $U < p$ : return  $X = 1$  (success)
3. If  $U \geq p$ : return  $X = 0$  (failure)

**Why it works:**  $\mathbb{P}(U < p) = p$  and  $\mathbb{P}(U \geq p) = 1 - p$  — exactly Bernoulli!

## Universality for Discrete: Binomial

Just repeat the Bernoulli trick  $n$  times

**Goal:** Generate  $X \sim \text{Binomial}(n, p)$  from Uniform draws

**Recall:**  $X \sim \text{Binomial}(n, p)$  counts successes in  $n$  independent Bernoulli( $p$ ) trials.

**Algorithm:**

1. Draw  $U_1, U_2, \dots, U_n$  independently from Uniform(0, 1)
2. For each  $U_i$ : count as “success” if  $U_i < p$
3. Return  $X$  = total number of successes

**Example** ( $n = 5, p = 0.5$ ):

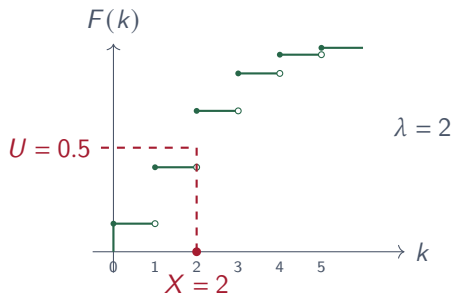
$$\begin{array}{ccccccccc} U_1 = 0.23 & U_2 = 0.81 & U_3 = 0.15 & U_4 = 0.67 & U_5 = 0.42 & & & & \\ < 0.5 \checkmark & \geq 0.5 & < 0.5 \checkmark & \geq 0.5 & \geq 0.5 & \Rightarrow & X = 2 & & \end{array}$$

This is exactly what `rbinom(1, n, p)` does in R.

## Universality for Discrete: Poisson

The CDF is a staircase — find which step  $U$  lands on

**Goal:** Generate  $X \sim \text{Poisson}(\lambda)$  from  $U \sim \text{Uniform}(0, 1)$



**Algorithm:** Find smallest  $k$  such that  $F(k) \geq U$

**Example** ( $\lambda = 2$ ,  $U = 0.5$ ):  $F(1) = 0.406 < 0.5$  but  $F(2) = 0.677 \geq 0.5 \Rightarrow X = 2$

**Key insight:** Height of each jump = PMF at that point. **Bigger jumps = more likely outcomes.**

## The Formal Statement: Universality of the Uniform

Now that you've seen it work, here's the theorem

### Theorem (Probability Integral Transform):

Let  $X$  be a continuous random variable with CDF  $F$ . Then:

$$F(X) \sim \text{Uniform}(0, 1)$$

**The converse** (the useful direction for simulation):

If  $U \sim \text{Uniform}(0, 1)$  and  $F$  is any CDF, then:

$$X = F^{-1}(U) \text{ has CDF } F$$

**This is why Uniform is the “mother of all distributions”:** From one Uniform, you can generate *any* distribution — continuous or discrete.

PS3 asks you to prove  $F(X) \sim \text{Unif}(0, 1)$  directly for Exponential. Hint: Find  $\mathbb{P}(F(X) \leq u)$ .

# The Big Picture: What We Just Learned

Universality of the Uniform, informally

**The Uniform is special** because its CDF is a  $45^\circ$  line — no distortion.

**Every other CDF “stretches” the Uniform:**

- Steep regions  $\rightarrow$  outcomes cluster there
- Flat regions  $\rightarrow$  outcomes are rare there

**The inverse CDF method:**

- Draw  $U \sim \text{Uniform}(0, 1)$  — your “raw randomness”
- Apply  $F^{-1}$  — the CDF of whatever distribution you want
- Out comes  $X \sim F$  — a draw from your target distribution

**This works for everything:** Bernoulli, Binomial, Poisson, Normal, Exponential — any distribution you’ll ever meet.

When you call `rnorm()`, `rpois()`, or `rexp()` in R, this is what’s happening under the hood.



# The Normal Distribution

The Star of the Show

# De Moivre Discovered It; Gauss Got the Credit

Stigler's Law of Eponymy

**Stigler's Law:** “No scientific discovery is named after its original discoverer.”

**The Normal distribution is called “Gaussian”—but Gauss didn't discover it.**

- **Abraham de Moivre (1733):** French Huguenot exile in London, surviving by tutoring aristocrats in gambling mathematics. First derived the normal curve in *The Doctrine of Chances*.
- **Pierre-Simon Laplace (1774–1812):** Developed the theory systematically. Proved early versions of the Central Limit Theorem.
- **Carl Friedrich Gauss (1809):** Applied it to astronomical errors. Got the credit. But Gauss himself called it the “Laplacian curve.”

De Moivre died impoverished. Gauss is called the “Prince of Mathematicians.” Life isn't fair.

# Gauss, Legendre, and Least Squares

The drama continues

**Stigler's Law strikes again:** Least squares wasn't discovered by Gauss either.

- **Legendre (1805):** Published the method of least squares
- **Gauss (1809):** Published it four years later, but claimed he'd been using it since 1795 — when he was 18
- Legendre was *furious*. Gauss offered no proof of his earlier use.

**But here's Gauss's real contribution:**

He showed that *if* errors are normally distributed, *then* least squares gives the best estimates (maximum likelihood).

**The Normal distribution justifies least squares.** That's why they're forever linked.

When you run a regression, you're relying on Gauss's 1809 insight — even if Legendre got there first.

# The Normal Distribution: The Star of the Show

Application first: Where do you see it?

## Examples:

- Heights of adults, test scores, measurement errors
- **Polling errors** — why we talk about “margin of error”
- **Regression residuals** — the foundation of inference

**Definition:**  $X \sim \text{Normal}(\mu, \sigma^2)$  has PDF:

$$f(x) = \frac{1}{\sqrt{2\pi\sigma^2}} \exp\left(-\frac{(x - \mu)^2}{2\sigma^2}\right)$$

**Why everywhere?** Three reasons:

1. **Central Limit Theorem:** Sample means are approximately normal
2. **Closure:** Sums of normals are normal
3. **Tractability:** Easy to compute probabilities

# Standardization Converts Any Normal to Z

**Definition:**  $Z \sim N(0, 1)$  is the **standard normal**.

**Standardization:** If  $X \sim N(\mu, \sigma^2)$ , then:

$$Z = \frac{X - \mu}{\sigma} \sim N(0, 1)$$

**What does Z mean?** It's the deviation from the mean, scaled by the standard deviation.

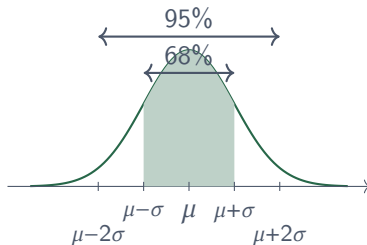
- $Z = 0$ : you're at the mean
- $Z = 1$ : you're one SD above the mean
- $Z = -2$ : you're two SDs below the mean

**Key notation:**  $\phi(z)$  = standard normal PDF;  $\Phi(z) = \mathbb{P}(Z \leq z)$  = standard normal CDF

Tables, software, and formulas are all in terms of  $\Phi$ . Standardize first.

## Most Probability Concentrates Near the Mean

For  $X \sim N(\mu, \sigma^2)$ :



- **68%** of values within 1 SD of mean
- **95%** within 2 SDs (more precisely: 1.96)
- **99.7%** within 3 SDs

# Normal Has Unbounded Support — But Tails Vanish Fast

Theoretically infinite, practically finite

**Support:**  $\text{Supp}[X] = (-\infty, +\infty)$  — any value is *theoretically* possible.

**But probabilities decay exponentially in the tails:**

- Outside 3 SDs: only 0.3% of probability
- Outside 4 SDs: only 0.006% of probability
- Outside 5 SDs: essentially zero (1 in 3.5 million)

**Practical implication:** For heights (mean 170cm, SD 10cm):

- Normal says negative heights are “possible” — but  $P(X < 0) \approx 0$
- The model is an approximation; we accept tiny errors in exchange for tractability

Contrast with Exponential: support  $[0, \infty)$  *enforces* non-negativity.

## A Warning: The Normal Can Be Misused

“The Bell Curve” controversy

**1994:** Herrnstein & Murray publish *The Bell Curve*, claiming IQ differences between racial groups are genetic and immutable.

**The statistical sin:** They treated the Normal distribution as *destiny* rather than *description*.

**James Heckman’s critique** (Nobel laureate, 2000):

- IQ is not fixed — it responds to environment and intervention
- The authors confused *description* with *explanation*
- Selection bias: who takes the tests, when, under what conditions?

**Lesson:** The Normal describes many phenomena. It doesn’t explain them. Distributions are tools, not theories of causation.

Statistics without causal reasoning is dangerous.



## Normal Closure Properties

Sums and linear combinations stay normal

**Property 1** (Scaling and shifting):

If  $X \sim N(\mu, \sigma^2)$ , then for constants  $a, b$ :

$$aX + b \sim N(a\mu + b, a^2\sigma^2)$$

**Property 2** (Sum of independent normals):

If  $X \sim N(\mu_X, \sigma_X^2)$  and  $Y \sim N(\mu_Y, \sigma_Y^2)$  are **independent**, then:

$$X + Y \sim N(\mu_X + \mu_Y, \sigma_X^2 + \sigma_Y^2)$$

**The key result for inference:** If  $Y_1, \dots, Y_n \stackrel{\text{iid}}{\sim} N(\theta, \sigma^2)$ , then:

$$\bar{Y} \sim N\left(\theta, \frac{\sigma^2}{n}\right)$$

PS3 asks: How does  $\mathbb{P}(|\bar{Y} - \theta| < \varepsilon)$  change with  $n$ ? This formula is your starting point.

## Galton and “Regression to the Mean”

Why tall parents have shorter children (on average)

**Francis Galton (1886):** Studied heights of fathers and sons.

**Finding:** Sons of very tall fathers were tall — but not *as* tall as their fathers. Sons of very short fathers were short — but not *as* short.

**Galton called this “regression toward mediocrity”:**

- Extreme observations tend to be followed by less extreme ones
- This is a *statistical phenomenon*, not a biological force
- It happens whenever two variables are imperfectly correlated

**Why “regression”?** This is literally where the term comes from. Galton was “regressing” son’s height on father’s height.

The Normal distribution quantifies this: extreme Z-scores are rare by definition.

# The Exponential Distribution

Waiting for an Event

# Where Does the Exponential Come From?

A story of outsiders, death, and waiting

**Benjamin Gompertz (1825):** A Jewish mathematician in London, barred from university because of his religion. Self-taught from Newton's writings. His brother-in-law founded an insurance company and made Gompertz the actuary.

His question: *How do we price life insurance?* He needed to model how long people live — and discovered that mortality risk increases exponentially with age. The exponential function became central to survival analysis.

**“Event history analysis” and “survival analysis” — now workhorses of modern social science — trace back to 19th-century actuaries modeling death.**

Gompertz was elected to the Royal Society despite being denied a university degree.

# How Long Until the Next Supreme Court Vacancy?

Application first: Waiting times in politics

## Political science questions that involve waiting:

- How long until the next Supreme Court vacancy?
- How long will this ceasefire last?
- How long until a cabinet collapse?
- Time between terrorist attacks in a region?

**Historical data:** Supreme Court vacancies occur at rate  $\lambda \approx 0.5$  per year.

⇒ Average wait: about 2 years between vacancies.

The **Exponential distribution** models these waiting times.

# The Exponential Distribution

The math behind waiting times

**Definition:**  $T \sim \text{Exponential}(\lambda)$  has PDF:

$$f(t) = \lambda e^{-\lambda t} \quad \text{for } t \geq 0$$

where  $\lambda > 0$  is the **rate parameter**.

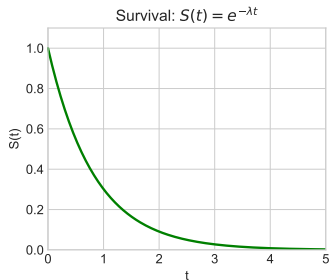
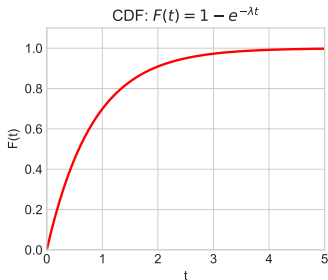
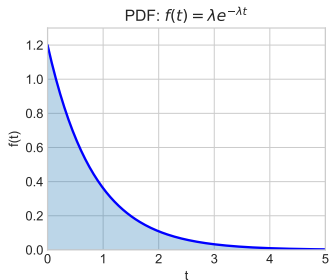
**Support:**  $[0, +\infty)$  — waiting times are always non-negative.

**PS2 Connection:** Problem 7(c) asked you to find  $c$  for  $f(y) = ce^{-2y}$ . That's an Exponential( $\lambda = 2$ )! The answer was  $c = 2$ .

# The Survival Function Decays Exponentially

PDF, CDF, and Survival — three views of the same distribution

Exponential Distribution ( $\lambda = 1.2$ )



**Survival function:**  $S(t) = \mathbb{P}(T > t) = 1 - F(t) = e^{-\lambda t}$

The probability of “surviving” (not yet experiencing the event) past time  $t$  decays exponentially.

# Average Wait Is the Inverse of the Rate

Key properties of the Exponential

For  $T \sim \text{Exponential}(\lambda)$ :

**Expected value:**  $\mathbb{E}[T] = \frac{1}{\lambda}$

**Variance:**  $\text{Var}[T] = \frac{1}{\lambda^2}$

**Interpretation:** If events occur at rate  $\lambda$  per unit time, the average wait is  $1/\lambda$ .

**Example:** Supreme Court vacancies at rate  $\lambda = 0.5$  per year  $\rightarrow$  average wait = 2 years.



# The Exponential Distribution Has No Memory

How long you've waited doesn't affect how much longer you'll wait

**Property:** For  $T \sim \text{Exponential}(\lambda)$ :

$$\mathbb{P}(T > s + t \mid T > s) = \mathbb{P}(T > t)$$

**In words:** Given that you've already waited  $s$  units, the probability of waiting *another*  $t$  units is the same as if you'd just started waiting.

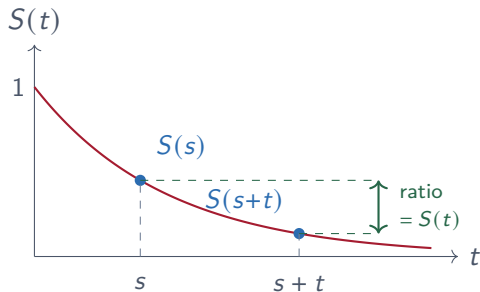
**Proof:**

$$\mathbb{P}(T > s + t \mid T > s) = \frac{\mathbb{P}(T > s + t)}{\mathbb{P}(T > s)} = \frac{e^{-\lambda(s+t)}}{e^{-\lambda s}} = e^{-\lambda t} = \mathbb{P}(T > t)$$

Only the exponential (continuous) and geometric (discrete) have this property.

# Memorylessness Visualized

The survival curve “restarts” at any point



The *ratio* of survival probabilities depends only on the additional wait  $t$ , not on  $s$ .

# Why Does “Memory” Even Make Sense Here?

It's the question being asked

**Bernoulli:** “What happened?” — heads or tails, 0 or 1.

- Resolves instantly. No duration. No elapsed time.
- You can't ask “given that I've been partially through this coin flip...”
- The concept of memory has no surface to attach to.

**Exponential:** “How long until something happens?”

- Time is explicitly in the picture. You're sitting there waiting.
- You *can* ask: “I've waited 3 years — does that change my forecast?”
- For the Exponential, the answer is *no*. That's memorylessness.

**The question determines whether memory is even a coherent concept.**

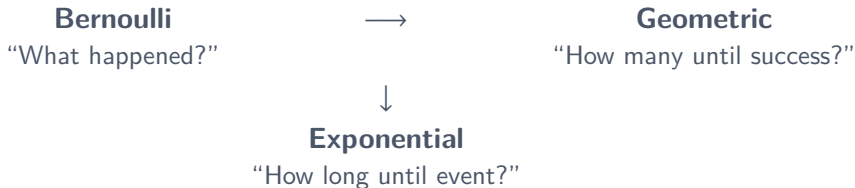
# The Geometric Bridge

Connecting Bernoulli to Exponential

The **Geometric distribution** asks a waiting-time question about Bernoulli trials:

“How many trials until the first success?”

- Each Bernoulli trial has no time dimension — it just happens
- But counting *how many* trials introduces a duration concept
- Now the past (failed trials) *could* inform the future
- And the answer is: it doesn't. **Geometric is memoryless too.**



Geometric is discrete memorylessness; Exponential is continuous memorylessness.

# The Poisson–Exponential Connection

Two Sides of One Process

## A Brief History: How Did We Get Here?

Three people, three problems, one insight

**Poisson (1837):** French mathematician studying rare events — wrongful convictions in court trials. Asked: “If something rarely happens, how do we model how *many* times it occurs?” Developed the Poisson distribution, but didn’t connect it to waiting times.

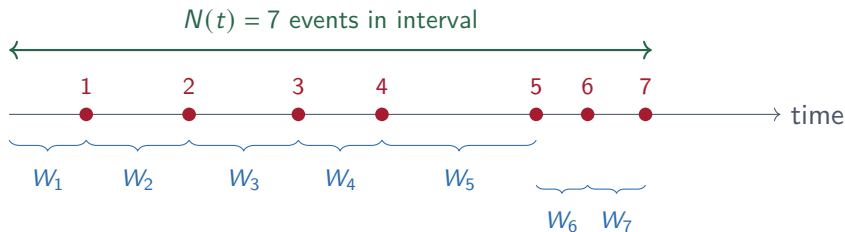
**Bortkiewicz (1898):** Russian economist, famously applied Poisson to Prussian cavalry soldiers killed by horse kicks. Showed rare events follow predictable patterns. Still focused on *counts*, not *durations*.

**Erlang (1909):** Danish engineer at the Copenhagen Telephone Exchange. His problem: how many phone lines do we need? He realized: calls arrive randomly (Poisson), but *time between calls* matters for capacity planning. **First to formally connect Poisson counts to exponential waiting times.**

The insight wasn’t obvious. It took 70+ years from Poisson’s distribution to Erlang’s connection. Today, it’s the foundation of queueing theory — from call centers to emergency rooms to internet servers.

# Poisson Counts Events; Exponential Measures Waiting Times

Same process, different questions



- **Poisson question:** How many events in time  $t$ ?  $N(t) \sim \text{Poisson}(\lambda t)$
- **Exponential question:** How long until next event?  $W_i \sim \text{Exp}(\lambda)$

**Same rate  $\lambda$ . Same process. Different questions.**

# The Key Identity

Connecting Poisson and Exponential

Let  $T_1$  be the time until the first event. Then:

$$\mathbb{P}(T_1 > t) = \mathbb{P}(\text{no events by time } t) = \mathbb{P}(N(t) = 0)$$

Using Poisson:

$$\mathbb{P}(N(t) = 0) = \frac{(\lambda t)^0 e^{-\lambda t}}{0!} = e^{-\lambda t}$$

This is exactly the survival function of Exponential( $\lambda$ )!

**The Poisson count and exponential waiting time are two views of the same process.**



## Political Science Example

### Supreme Court vacancies

Suppose vacancies occur at rate  $\lambda = 0.5$  per year.

**Poisson question:** What's  $\mathbb{P}(\text{at least 2 vacancies in a 4-year term})$ ?

- $N(4) \sim \text{Poisson}(0.5 \times 4) = \text{Poisson}(2)$
- $\mathbb{P}(N \geq 2) = 1 - \mathbb{P}(N = 0) - \mathbb{P}(N = 1) = 1 - e^{-2} - 2e^{-2} \approx 0.59$

**Exponential question:** What's the average wait for the next vacancy?

- $T \sim \text{Exp}(0.5)$
- $\mathbb{E}[T] = 1/0.5 = 2$  years

Same  $\lambda$ , different questions, complementary answers.

## Your Turn: Continuous Practice

Work through these with a partner

1. **Normal:** Adult heights are  $N(170, 100)$  cm (mean 170, variance 100).
  - What's the standard deviation?
  - What range contains about 95% of heights?
2. **Exponential:** Congressional hearings occur at rate  $\lambda = 3$  per month.
  - What's the expected wait for the next hearing?
  - What's  $\mathbb{P}(\text{wait} > 1 \text{ month})$ ?

Answers: (1)  $SD = 10$  cm; 150–190 cm. (2)  $\mathbb{E}[T] = 1/3$  month;  
 $\mathbb{P}(T > 1) = e^{-3} \approx 0.05$ .

# Chi-Square and $t$ Distributions

The B-List: Supporting Actors for Inference

# Chi-Square and $t$ Are Inference Tools, Not Data Models

You don't model data with these—you use them for hypothesis testing

## A-list vs B-list:

- **A-list** (Normal, Exponential, Uniform): You model *data* with these
- **B-list** (Chi-square,  $t$ ): You use these for *inference about parameters*

## Why do they exist?

- **Chi-square**: When you estimate variance from data, your estimate follows a  $\chi^2$
- **$t$ -distribution**: When you test hypotheses using an estimated (not known) variance

**The punchline:** In a few weeks, when you run a regression and ask “is this coefficient statistically significant?”—the  $t$ -distribution will give you the answer.

We're planting seeds. You'll see these again in the regression unit.

# The Chi-Square: Karl Pearson's 1900 Revolution

The birth of the goodness-of-fit test

**Karl Pearson (1900):** Statistician at University College London. Asked a simple question: *How do I know if my data actually fit a theoretical distribution?*

Before Pearson, researchers just assumed data were Normal. Pearson noticed real biological data were often skewed. He needed a formal test.

His solution: Sum up squared deviations between observed and expected counts. That sum follows a  $\chi^2$  distribution — and gives you a p-value.

**The drama:** Pearson got the degrees of freedom wrong. A young outsider named R.A. Fisher corrected him in 1922. Pearson refused to accept the correction and published a hostile “cooperative study” attacking Fisher. Fisher was furious, vowed never to publish in Pearson’s journal again, and declared war on Pearson’s entire approach to statistics. The feud shaped 20th-century statistics. Fisher was right about the degrees of freedom.

# Chi-Square Is a Sum of Squared Normals

Derived from Normal—support is  $[0, \infty)$

**Definition:** If  $Z_1, \dots, Z_k \stackrel{\text{iid}}{\sim} N(0, 1)$ , then:

$$X = Z_1^2 + Z_2^2 + \dots + Z_k^2 \sim \chi_k^2$$

where  $k$  is the **degrees of freedom**.

**Key facts:**

- $\mathbb{E}[X] = k$
- $\text{Var}[X] = 2k$
- Support:  $[0, \infty)$  — always non-negative (it's a sum of squares)

You'll see this when we estimate variance, test hypotheses about multiple coefficients, and compute  $R^2$ .

# The $t$ -Distribution: A Brewer's Secret

William Sealy Gosset and the Guinness brewery

**William Sealy Gosset (1908):** Chemist at the Guinness brewery in Dublin. His job: assess barley and hops quality. His problem: he only had small samples.

With small samples, the Normal distribution gives wrong answers — it's too confident. Gosset worked out the math for what happens when you estimate variance from limited data.

**The catch:** Guinness didn't allow employees to publish under their real names (they feared leaking trade secrets). So Gosset published as "Student" in 1908.

That's why it's called **Student's  $t$ -distribution** — not Gosset's.

**The connection:** Gosset spent 1906–07 studying with Karl Pearson in London. Pearson helped him with the mathematics. The chi-square and  $t$  are siblings.

Gosset's identity was only revealed publicly after his death in 1937.

## The $t$ Distribution Is Normal with Heavier Tails

What happens when you don't know the true variance

**The problem:** In real life, you don't know  $\sigma$ . You estimate it from data.

**The consequence:** Your estimate  $\hat{\sigma}$  is uncertain. This makes extreme values more likely than the Normal predicts.

**The solution:** Use the  $t$ -distribution, which has heavier tails to account for this.

**Definition:** If  $Z \sim N(0, 1)$  and  $V \sim \chi_k^2$  are independent, then:

$$T = \frac{Z}{\sqrt{V/k}} \sim t_k$$

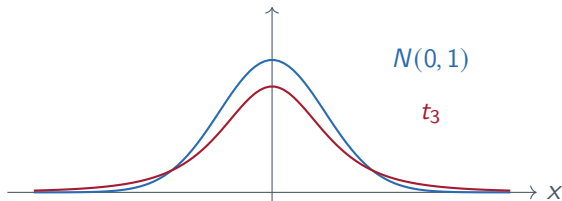
**Key insight:** Small  $k$  = more uncertainty about  $\sigma$  = heavier tails.

As  $k \rightarrow \infty$ , the  $t$  becomes Normal (you've estimated  $\sigma$  precisely).

This is why we use “ $t$ -tests” — they account for estimating variance from data.



## Normal vs. $t$ : Heavier Tails



The  $t$  distribution has more probability in the tails.

With small samples, extreme values are more likely — the  $t$  accounts for this.

## Summary: Continuous Distributions

Distribution	$\mathbb{E}[X]$	$\text{Var}[X]$	Use case
Uniform( $a, b$ )	$\frac{a+b}{2}$	$\frac{(b-a)^2}{12}$	Equal probability, simulation
Normal( $\mu, \sigma^2$ )	$\mu$	$\sigma^2$	CLT, regression errors
Exponential( $\lambda$ )	$\frac{1}{\lambda}$	$\frac{1}{\lambda^2}$	Waiting times, memoryless
$\chi_k^2$	$k$	$2k$	Variance estimation, tests
$t_k$	0	$\frac{k}{k-2}$	Small-sample inference

### Key connections:

- Uniform(0,1)  $\rightarrow$  any distribution via inverse CDF
- Poisson  $\leftrightarrow$  Exponential: counts vs. waiting times
- Normal  $\rightarrow$  Chi-square (sum of squares)  $\rightarrow t$  (ratio)

## The A-List and B-List: A Summary

Which distributions model data? Which are for inference?

**A-list actors** — you model *data* with these:

- **Uniform**: Simulation, randomization, probability foundations
- **Normal**: CLT, regression errors, test scores, polling
- **Exponential**: Waiting times, survival analysis, duration

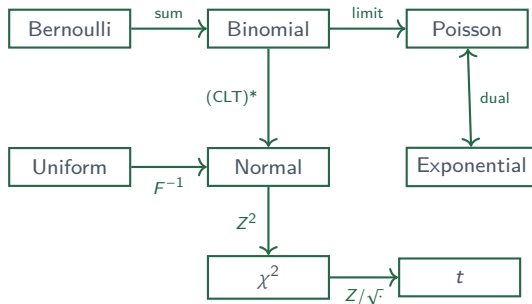
**B-list actors** — you use these for *inference*:

- **Chi-square**: Variance estimation, goodness-of-fit tests
- **t**: Hypothesis testing with estimated variance

**The relationship:** Normal  $\xrightarrow{Z^2}$  Chi-square  $\xrightarrow{Z/\sqrt{\cdot}}$  t

Most of your statistical life will be Normal. But when you estimate variance from data, the B-list appears.

# How Distributions Connect: The Big Picture



Understanding these connections helps you see why certain distributions appear in certain contexts.

\*CLT = Central Limit Theorem (Week 5). Sample means of *any* distribution approach Normal.

# Looking Ahead

**Next week:** Joint distributions and the CEF

- Joint, marginal, and conditional distributions
- Covariance and correlation
- The Conditional Expectation Function (CEF)

**Reading:**

- Aronow & Miller, §1.3 and §2.2
- Blackwell, Chapter 2.4–2.5

**Problem Set 3:** Due next week, February 17th.