HW1

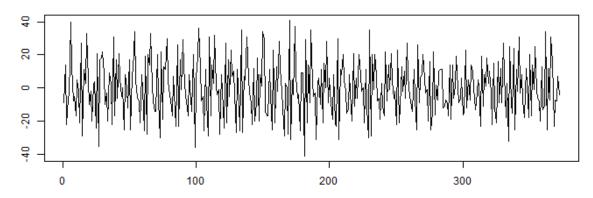
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March 16, 2022

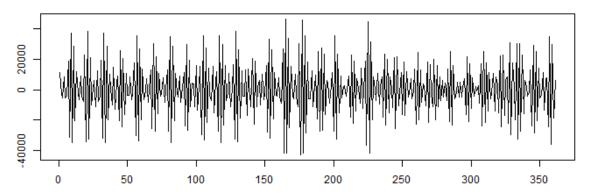
1. For the data set 1 on Moodle, write some R code to perform ∇ and ∇_{12} on the time series. In addition, plot the time series after the differencing. Report your R code.

```
> qldata <- read.table("HW1_dataset.txt", head = T)
> diff.func <- function(x, order = 1) {
+    for(i in 1:order) x <- x[-1]-x[-length(x)]
+    return(x)
+ }
> diff.1 <- diff.func(qldata$birth_rates)
> diff.12 <- diff.func(qldata$birth_rates, 12)
> par(mfrow=c(2,1))
> plot(diff.1, type="l", xlab = "", ylab = "", main = "Differences of order1")
> plot(diff.12, type="l", xlab = "", ylab = "", main = "Differences of order12")
```

Differences of order1



Differences of order12



- 2. Consider the time series $X_t = \beta_0 + \beta_1 t + Z_t$, where β_0 and β_1 known constants and $Z_t \sim WN(0, \sigma^2)$.
 - (a) Determine whether X_t is stationary.

$$E(X_t) = E(\beta_0 + \beta_1 t + Z_t) = \beta_0 + \beta_1 t + E(Z_t) = \beta_0 + \beta_1 t.$$

Since $E(X_t)$ depents on time, X_t is not stationary.

(b) Show that the process $Y_t = X_t - X_{t-1}$ is stationary.

$$E(Y_t) = E(X_t - X_{t-1}) = E(X_t) - E(X_{t-1}) = \beta_0 + \beta_1 t - (\beta_0 + \beta_1 t - \beta_1) = \beta_1$$

and let s = t + h

$$\gamma(h) = \operatorname{Cov}(Y_s, Y_t) = \operatorname{Cov}(Y_{t+h}, Y_t) = \operatorname{E}\left[(Z_{t+h} - Z_{t+h-1})(Z_t - Z_{t-1})\right]$$

$$= \operatorname{E}(Z_{t+h}Z_t - Z_{t+h}Z_{t-1} - Z_{t+h-1}Z_t + Z_{t+h-1}Z_{t-1})$$

$$= \begin{cases} 2\sigma^2 & , h = 0 \\ -\sigma^2 & , h = 1 \\ 0 & , h > 1. \end{cases}$$

Same result as s = t - h would be obtained. Hence, Y_t is stationary.

(c) Show that the mean of the moving average $V_t = \frac{1}{2q+1} \sum_{i=-q}^q X_{t-i}$ is $\beta_0 + \beta_1 t$.

$$\begin{split} \mathbf{E}\left(V_{t}\right) &= \mathbf{E}\left(\frac{1}{2q+1}\sum_{i=-q}^{q}X_{t-i}\right) \\ &= \frac{1}{2q+1}\sum_{i=-q}^{q}(\beta_{0}+\beta_{1}t-\beta_{1}i) \\ &= \frac{1}{2q+1}\sum_{i=-q}^{q}(\beta_{0}+\beta_{1}t)-\beta_{1}\sum_{i=-q}^{q}i \\ &= \frac{1}{2q+1}\left[(2q+1)(\beta_{0}+\beta_{1}t)\right]-0 \\ &= \beta_{0}+\beta_{1}t. \end{split}$$

- 3. Verify the calculations made in Example 3.4 as follows.
 - (a) Let $X_t = \phi X_{t-1} + Z_t$, where $|\phi| > 1$ and $Z_t \sim i.i.d.N(0, \sigma^2)$. Show that

$$E(X_t) = 0$$
 and $\gamma_X(h) = \frac{\sigma^2 \phi^{-2} \phi^{-h}}{1 - \phi^{-2}}$ for $h \ge 0$.

Let $X_{t+1} = \phi X_t + Z_{t+1}$, then

$$X_{t} = \phi^{-1}X_{t+1} - \phi^{-1}Z_{t+1} = \phi^{-1}(\phi^{-1}X_{t+2} - \phi^{-1}Z_{t+2}) - \phi^{-1}Z_{t+1} = \dots = -\sum_{i=1}^{\infty} \phi^{-i}Z_{t+i}.$$

Hence,

$$E(X_t) = E\left(-\sum_{j=1}^{\infty} \phi^{-j} Z_{t+j}\right) = -\sum_{j=1}^{\infty} \phi^{-j} E(Z_{t+j}) = 0$$

and

$$\begin{split} \gamma_X(h) &= \operatorname{Cov}(X_{t+h}, X_t) = \operatorname{E}\left[\left(-\sum_{j=1}^{\infty} \phi^{-j} Z_{t+h+j}\right) \left(-\sum_{k=1}^{\infty} \phi^{-k} Z_{t+k}\right)\right] \\ &= \operatorname{E}\left[(\phi^{-1} Z_{t+h+1} + \phi^{-2} Z_{t+h+2} + \ldots) (\phi^{-1} Z_{t+1} + \ldots + \phi^{-(h+1)} Z_{t+h+1} + \phi^{-(h+2)} Z_{t+h+2} + \ldots)\right] \\ &= \phi^{-(h+2)} \sigma^2 + \phi^{-(h+4)} \sigma^2 + \ldots \\ &= \sigma^2 \phi^{-h} \sum_{i=1}^{\infty} \phi^{-2i} \\ &= \frac{\sigma^2 \phi^{-2} \phi^{-h}}{1 - \phi^{-2}}. \end{split}$$

(b) Let $Y_t = \phi^{-1} Y_{t-1} + W_t$, where $W_t \sim i.i.d.$ N(0, $\sigma^2 \phi^{-2}$). Argue that Y_t is causal with the same mean function and autocovariance function as X_t .

$$\begin{split} Y_t &= \phi^{-1} Y_{t-1} + W_t = \phi^{-1} (\phi^{-1} Y_{t-2} + W_{t-1}) + W_t = \ldots = \sum_{j=0}^\infty \phi^{-j} W_{t-j}. \\ & \to \left(\sum_{j=0}^\infty \phi^{-j} W_{t-j} \right) = \sum_{j=0}^\infty \phi^{-j} \mathbf{E} \left(W_{t-j} \right) = 0. \\ & \gamma_Y(h) &= \operatorname{Cov} \left(Y_{t+h}, Y_t \right) = \mathbf{E} \left[\left(\sum_{j=0}^\infty \phi^{-j} W_{t+h-j} \right) \left(\sum_{k=0}^\infty \phi^{-k} W_{t-k} \right) \right] \\ &= \mathbf{E} \left[\left(W_{t+h} + \ldots + \phi^{-h} W_t + \phi^{-(h+1)} W_{t-1} + \ldots \right) (W_t + \phi^{-1} W_{t-1} + \ldots) \right] \\ &= \sigma^2 \phi^{-2} \phi^{-h} \sum_{i=0}^\infty \phi^{-2i} \\ &= \frac{\sigma^2 \phi^{-2} \phi^{-h}}{1 - \phi^{-2}}. \end{split}$$

4. (a) Derive the autocovariance function and ACF for the MA(1) process

$$\begin{split} X_t &= Z_t + \theta Z_{t-1}, Z_t \sim \text{WN}(0, \sigma^2) \\ \gamma(0) &= \text{Cov}\left(X_t, X_t\right) = \text{E}\left[(Z_t + \theta Z_{t-1})(Z_t + \theta Z_{t-1})\right] = \text{E}\left(Z_t^2 + \theta^2 Z_{t-1}^2\right) = (1 + \theta^2)\sigma^2, \\ \gamma(1) &= \text{Cov}\left(X_{t+1}, X_t\right) = \text{E}\left[(Z_{t+1} + \theta Z_t)(Z_t + \theta Z_{t-1})\right] = \text{E}\left(\theta Z_t^2\right) = \theta \sigma^2, \\ \gamma(2) &= \text{Cov}\left(X_{t+2}, X_t\right) = \text{E}\left[(Z_{t+2} + \theta Z_{t+1})(Z_t + \theta Z_{t-1})\right] = 0. \end{split}$$

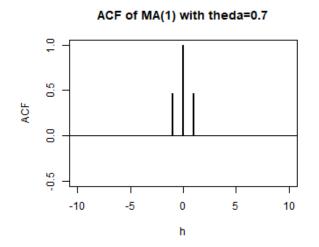
Hence, the autocovariance function is

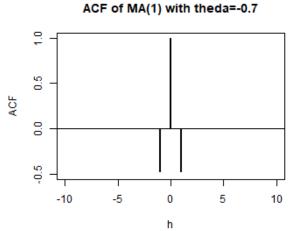
$$\gamma(h) = \begin{cases} (1+\theta^2)\sigma^2 &, h = 0 \\ \theta\sigma^2 &, h = 1 \\ 0 &, h > 1. \end{cases}$$

and the ACF is

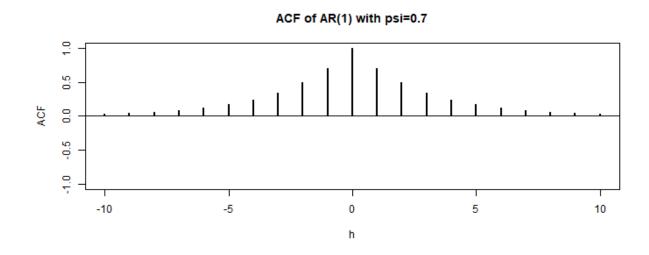
$$\rho(h) = \frac{\gamma(h)}{\gamma(0)} = \begin{cases} 1 & , h = 0 \\ \frac{\theta}{1+\theta^2} & , h = 1 \\ 0 & , h > 1. \end{cases}$$

- (b) Suppose further that $\theta = 0.7$, plot the ACF for $h \in Z$.
- (c) Repeat (b) for $\theta = -0.7$





(d) The ACF for an AR(1) process has been derived in lecture notes. Based on the function, plot the ACF of $X_t = 0.7X_{t-1} + Z_t$ for $h \in Z$.



(e) What is the difference between the ACF for the MA(1) and AR(1) in (b) and (d)? As h(h > 0) increases, the ACF of (b) decreases faster than that of (d).

Appendix: all the code for the work

```
### Homework1: Time Series Analysis (110下). Due 2022.03.24 ######################
## 01
qldata <- read.table("HW1 dataset.txt", head = T)</pre>
diff.func <- function(x, order = 1) {</pre>
  for (i in 1:order) x \leftarrow x[-1]-x[-length(x)]
  return(x)
diff.1 <- diff.func(qldata$birth rates)</pre>
diff.12 <- diff.func(qldata$birth_rates, 12)</pre>
png (file="Rplothw101.png", width=\overline{7}00, height=600)
par (mfrow=c(2,1))
plot(diff.1, type="l", xlab = "", ylab = "", main = "Differences of order1")
plot(diff.12, type="l", xlab = "", ylab = "", main = "Differences of order12")
dev.off()
ACF.MA1.func <- function(theda, range) {
  x <- range[1]:range[2]</pre>
  y <- rep(0, length(x))
  y[x == 0] \leftarrow 1
  y[abs(x) == 1] \leftarrow theda/(1+theda^2)
  return(cbind(x, y))
png(file="Rplothw102.png", width=700, height=300)
par (mfrow=c(1,2))
plot(ACF.MA1.func(0.7, c(-10, 10)), type = 'h', lwd = 2, ylim = c(-0.5, 1),
      ylab = "ACF", xlab = "h", main = "ACF of MA(1) with theda=0.7")
abline(h = 0)
plot(ACF.MA1.func(-0.7, \mathbf{c}(-10, 10)), type="h", lwd = 2, ylim = \mathbf{c}(-0.5, 1), ylab = "ACF", xlab = "h", main = "ACF of MA(1) with theda=-0.7")
abline(h = 0)
dev.off()
ACF.AR1.func <- function(theda, range) {
  x <- range[1]:range[2]</pre>
  y <- theda^abs(x)
  return(cbind(x, y))
png(file="Rplothw103.png", width=700, height=300)
par (mfrow=c(1,1))
plot(ACF.AR1.func(0.7, c(-10, 10)), type="h", lwd = 2, ylim = c(-1, 1),
      xlab = "h", ylab = "ACF", main = "ACF of AR(1) with psi=0.7")
abline(h = 0)
dev.off()
```