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# Combining canopy height and tree species map information for large scale timber volume estimations under strong heterogeneity of auxiliary data and variable sample plot sizes

Andreas Hill · Henning Buddenbaum · Daniel Mandallaz

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Abstract A timber-volume regression model applicable 23 to the entire forest area of the federal German state of 24 Rhineland-Palatinate is identified using a combination of 25 airborne laser scanning (ALS)-derived metrics and informa- 26 tion from a satellite-based tree species classification map 27 available on the federal state level. As is common in many 28 forest inventory datasets, strong heterogeneity in the Li-DAR data due to different acquisition dates and misclassifications in the tree species classification map had noticeable effects on the regression model's performance. This ar-10 ticle specifically addresses techniques that improve the per-32 formance of ordinary least square regression models under such restricting conditions. We introduce a calibration tech-33 nique to neutralize the effect of misclassifications in the tree species variable that originally caused a residual inflation 34 15 of 5%. Incorporating the calibrated tree species information improved the model accuracy by 5% in adjusted  $R^2_{36}$ 17 and suggests the use of such information in forthcoming in-37 ventories. We also found that including ALS quality infor-38 19 mation as categorical variables within the regression model 30 considerably mitigates issues with time lags between the 40 ALS and terrestrial data acquisition and ALS quality vari-41

Andreas Hill
Department of Environmental Systems Science, ETH Zurich, Universitaetstrasse 22, 8092 Zurich, Switzerland
Tel.: +41 44 632 32 36
E-mail: andreas.hill@usys.ethz.ch
Henning Buddenbaum
Environmental Remote Sensing and Geoinformatics Department, Trier 48
University, 54286 Trier, Germany
Tel.: +49 651 201 4729
E-mail: buddenbaum@uni-trier.de
Daniel Mandallaz
Department of Environmental Systems Science, ETH Zurich, Universitaetstrasse 22, 8092 Zurich, Switzerland
Tel.: +41 44 632 88 35

E-mail: daniel.mandallaz@usys.ethz.ch

ations (9% increase in adjusted  $R^2$ ). The model achieved an adjusted  $R^2$  of 0.48 (RMSE<sub>cv</sub> of 137 m³/ha) under incorporation of the tree species and ALS quality information, and was thus improved by 13% (16 m³/ha) compared to the simple model only containing ALS height metrics (adjusted  $R^2$ =0.35, RMSE<sub>cv</sub>=153 m³/ha).

**Keywords** OLS Regression  $\cdot$  standing timber volume  $\cdot$  ALS canopy height model  $\cdot$  satellite-based tree species classification  $\cdot$  calibration  $\cdot$  forest inventory  $\cdot$  angle count sampling

# 1 Introduction

Forest inventory methods are the primary tools used to assess the current state and development of forests over time. They provide reliable evidence-based information that is used to define and identify management actions as well as to adapt forest management strategies to both national and international guidelines. Two methods that have become particularly attractive are so-called double-sampling (Mandallaz, 2008) and mapping (Brosofske et al, 2014) procedures. The core concept of these methods is to use predictions of the terrestrial target variable at additional sample locations where the terrestrial information has not been gathered. These predictions are produced by models that use explanatory variables derived from auxiliary data, commonly in the form of spatially exhaustive remote sensing data in the inventory area. Especially models to predict timber volume based on airborne laser scanning (ALS) have been extensively investigated for a long time (Næsset, 1997). The specific scope of double-sampling is to enlarge the terrestrial sample size by a much larger sample of predictions of the target variable in order to gain higher estimation precision without performing additional expensive terrestrial measurements. Model-based and model-assisted regression

estimators are used in a broad range of double sampling109 concepts and methods (Gregoire and Valentine, 2007; Köhlı10 et al, 2006; Mandallaz, 2013a,b; Saborowski et al, 2010;111 Schreuder et al, 1993) and have been applied to existing in-112 ventory systems (Breidenbach and Astrup, 2012; von Lüpke113 and Saborowski, 2014; Magnussen et al, 2014; Mandallazı14 et al, 2013; Massey et al, 2014). While double-sampling<sub>115</sub> methods provide reliable estimates for a given spatial unit,116 e.g. a forest district, they do not provide information about<sub>117</sub> the spatial distribution of the estimated quantity within this:118 area. For this reason, the same modeling technique used in 119 double-sampling procedures has also been intensively used 120 to produce exhaustive prediction maps that provide pixel-121 wise estimations of a target variable in high spatial resolu-122 tion (Bohlin et al, 2017; Hill et al, 2014; Latifi et al, 2010;123 Nink et al, 2015; Tonolli et al, 2011).

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To allow for an area-wide application of the prediction125 model, both double sampling and mapping methods require126 that the remote sensing data are available over the entire in-127 ventory area. This is usually not a limiting factor in small-128 scale applications. In the optimal case, the remote sensing<sup>129</sup> data are in principle collected in accordance to the specific130 study objective. Quality standards that have often been ad-131 dressed are that a) the remote sensing data should be ac-132 quired close to or even at the time of the terrestrial inven-133 tory in order to ensure best possible comparability between  $_{134}$ the target variable on the ground and the remote sensing de-135 rived variables (McRoberts et al, 2015); b) the remote sens-136 ing technology and its spectral and spatial resolution should 137 be chosen according to the modelling purpose (Köhl et al,138 2006); and c) the variation in quality of the remote sensing 139 data over the inventory area should be minimized in order to140 avoid artificial noise in the data (Naesset, 2014). Despite the increasing availability and decreasing costs of remote sens-142 ing data (White et al, 2016), these quality standards of the<sub>143</sub> remote sensing data can often not be guaranteed for large-144 scale applications (Maack et al, 2016), and trade-offs must, 45 be accepted (Jakubowski et al, 2013). The prime objective,146 is then to produce the best possible prediction model given, 47 the restrictions imposed by the available remote sensing in-148 formation. The exploration of scarcely used remote sensing, 149 products and the optimization of prediction models under<sub>150</sub> severe quality restrictions in the remote sensing data are thus  $_{151}$ one of the challenges in large-scale model-supported inven-152 tory applications.

Among the still rarely used remote sensing data in large<sub>154</sub> scale applications, the integration of tree species informa-<sub>155</sub> tion in prediction models - especially for timber volume<sub>156</sub> estimation - has been stated as some of the most promis-<sub>157</sub> ing but often missing information (Koch, 2010; White et al.,158 2016). As timber volume estimations on the single tree<sub>159</sub> level in forest inventories are often based on species-specific<sub>160</sub> biomass and volume equations (Husmann et al, 2017; Zia-<sub>161</sub>

nis et al, 2005), the application of species-specific models is expected to be a key factor for improving estimation precision (White et al, 2016). Breidenbach et al (2008) found that their timber volume prediction model based on ALS canopy height metrics could be significantly improved by including a variable estimating the deciduous proportion derived from leaf-off ALS data. Similar gains in model performance were also reported by Straub et al (2009) and Latifi et al (2012) who used broadleaf and coniferous information based on color infrared orthophotos as a categorical explanatory variable. However, studies that explore the use of more species-specific information (i.e. a further discrimination of tree species) as explanatory variables in prediction models have been rare. Further investigations are thus necessary especially in countries whose forests are characterized by a larger variety of tree species that may also occur in mixed and uneven-aged stands (McRoberts et al, 2010). The areawide tree species information in most studies was obtained from satellite and airborne remote sensing sensors based on automatic classification methods. Whereas the presence of misclassifications has already been addressed (Latifi et al, 2012), an issue that has so far been neglected is how misclassifications actually affect the prediction model (Gustafson, 2003).

A frequently encountered problem in large scale forest inventories is the lack of temporal synchronicity between the remote sensing acquisition and the terrestrial survey. As a result, the available remote sensing data often exhibit notable time-lags with respect to the date of the terrestrial inventory. This has often been addressed as a major drawback, especially for the application of model-assisted change estimation (Massey and Mandallaz, 2015).

Our study is embedded in the current implementation of model-assisted regression estimators (Mandallaz, 2013a,b; Mandallaz et al, 2013) for estimating the standing timber volume within the state and communal forest management units over the entire state of Rhineland-Palatinate (RLP, Germany). With respect to this overall objective, the aim of this study was to derive an ordinary least square (OLS) regression model to generate predictions of the standing timber volume associated with a sample location of the Third German National Forest Inventory (BWI3) over the entire federal state forest area (6155 km<sup>2</sup>). A merged ALS dataset from different acquisition years and a satellite-based tree species classification map for the five main tree species in RLP was available for the entire inventory area and consequently used to derive predictor variables. The major limiting factors for using these data in a regression analysis are (i) variation in the ALS data quality as well as time-lags of up to 10 years between the ALS acquisitions and the terrestrial survey, (ii) misclassifications in the tree species classification map and (iii) the ambiguous choice of a suitable extraction area (support) for all remote sensing information

under angle count sampling in the terrestrial survey (variable sample plot sizes). For this reason, we address the following specific research questions:

- 1. How can tree species map information be optimally used within a regression model that predicts timber volume? What effects do misclassifications have on the predictions and how can these effects be minimized?
- 2. What are the effects of quality restrictions and substantial time lags between the ALS- and terrestrial data acquisition on the regression model and how can these effects be mitigated?
- 3. Does support size influence model accuracy? What is the optimal support size and what are the determining factors?

# 6 2 Materials and Methods

#### 2.1 Study Area

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The German federal state Rhineland-Palatinate (RLP) is located in the western part of Germany and borders Luxem-209 bourg, France and Belgium (figure 1). With 42.3% (appr. 8400 km<sup>2</sup>) of the entire state area (19850 km<sup>2</sup>) covered<sub>210</sub> by forest, RLP is one of the two states with the high-211 est forest coverage among all federal states of Germany<sub>212</sub> (von Thünen-Institut, 2014). The most frequent tree species213 in RLP are European beech (fagus sylvatica, 21.8%), oak214 (quercus petrea and quercus robur, 20.2%), Norway spruce215 (picea abies, 19.5%), Scots pine (pinus sylvestris, 9.9%), 216 Douglas fir (pseudozuga menziesii, 6.4%), European larch<sub>17</sub> (larix decidua, 2.4%) and Silver fir (abies alba, 0.7%). The218 share of broadleaf tree species is 58.7%. The forests of<sub>219</sub> RLP further exhibit heterogeneous structures (von Thünen-220 Institut, 2014): around 82% of the forest area in RLP are221 mixed forest stands (i.e. at least two different tree species222 occur in the same stand) and 69% of the forest area exhibit a223 multi-layered vertical structure. While the average tree age224 is around 80 years, most of the forest area (20%) is occupied<sub>225</sub> by trees between 40 and 60 years of age, whereas 27% of 226 the trees are older than 100 years. Spatially variable climate<sub>227</sub> conditions have a strong influence on the local growth dy-228 namics as well as tree species composition and create a large229 variety of forest structures, ranging from characteristic oak230 coppices (Moselle valley), pure spruce, beech and scots pine231 forests (e.g. Hunsrück and Palatinate forest) to mixed forests232 comprising variable proportions of oak, larch, spruce, Scots233 pine and beech. Accordingly, RLP has been divided into 16234 bioclimatic growing regions that form homogeneous areas235 with respect to the afore mentioned characteristics (Gauer236 and Aldinger, 2005).

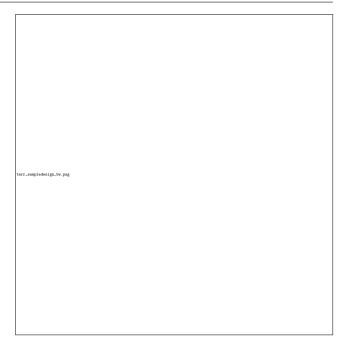


Fig. 1: Spatial distribution of the BWI3 cluster samples over Rhineland-Palatinate

# 2.2 Terrestrial Inventory Data

The German National Forest Inventory is carried out over the entire forest area of Germany in reoccurring time periods of 10 years. The most recent inventory (BWI3) has been conducted in the years 2011 and 2012. In this framework, Rhineland-Palatinate is covered by a 2x2 km grid that defines the sample locations for the terrestrial survey. A sample unit consists of four sample locations (also referred to as sample plots) that are arranged in squares (so called clusters) with a side length of 150 metres (figure 1). The number of plots per cluster can however vary between 1 and 4 depending on forest/non-forest decisions on the plot level (Bundesministerium für Ernährung, 2011). In the field survey of the BWI3, sample trees for timber volume estimations are selected according to the angle count sampling technique (Bitterlich, 1984), using a basal area factor (BAF) of 4 that is respectively adjusted for boundary effects at the forest border (Bundesministerium für Ernährung, 2011). A further selection criterion for a tree to be recorded is a diameter at breast height (dbh) of at least 7 cm. This sampling technique was applied to 8092 sample plots (2810 clusters) in RLP, resulting in the collection of 56561 sample trees for which the dbh, the tree diameter at 7 m (D7) and the tree species were recorded for all trees. Tree height measurements were conducted only for a subset of all sample trees and used to predict the height for the remaining sample. During the last inventory, all plot center positions were remeasured with a differential GPS technique. Knowledge about the exact plot positions were considered crucial

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to provide optimal comparability between the terrestrially<sub>287</sub> observations and the information derived from the auxiliary288 information. A detailed analysis of horizontal DGPS errors289 in RLP by Lamprecht et al (2017) indicated that horizon-290 tal DGPS errors do not exceed a range of 8 meters for 80\%291 of all plots. For 162 plots, the DGPS coordinates were re-292 placed by their former target coordinates due to missing or293 implausible values. In order to derive a volume estimation<sub>294</sub> for each sample tree, the BWI3 estimates a taper curve for<sub>295</sub> each sample tree by calibrating the random effects term of 296 linear mixed-effects taper models with the set of diameters<sub>297</sub> and corresponding height measurements taken from the re-298 spective sample tree (Kublin et al, 2013). The integration of 299 the derived taper curves consequently lead to a volume pre-300 diction for each sample tree. Since the overall objective of 301 the study was to subsequently use the identified regression<sub>302</sub> model for design-based timber volume estimations within<sub>203</sub> the state and communal forest management units, we al-304 ready restricted the sample plots used for modeling to the305 state and communal forest area (73% of the entire forest<sub>306</sub> area of RLP). This provides the advantage that when used as<sub>307</sub> an internal model in design-based estimators, the regression, model predictions already hold the assumption on the resid-309 uals to be zero on average for state and communal forest by  $_{\mbox{\tiny 310}}$ construction of OLS technique (Mandallaz, 2013a,b; Man-311 dallaz et al, 2013). The dataset of this study hence comprised 5791 plots (2055 clusters). For this sample, the timber vol- $_{313}$ ume density per hectare on plot level, Y(x), was calculated<sub>314</sub> according to the formula of one-phase one-stage sampling, 15 (Mandallaz, 2008). The timber volume density per hectare<sub>316</sub> on plot level was used as the response variable in the regres-317 sion analysis.

# 2.3 Auxiliary Information

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# 2.3.1 LiDAR Canopy Height Model

Between 2003 and 2013, the topographic survey institu-325 tion of RLP acquired airborne laser scanning (LiDAR) data326 over the entire state of RLP at leaf-off condition (Figure327 2). The objective of this campaign was to derive a coun-328 trywide digital terrain and surface model based on the ac-329 quired LiDAR point clouds. During the extended acquisi-330 tion period, airborne laser scanning technology and data331 quality evolved significantly. The tiles recorded in 2002 and332 2003 have a rather poor quality with about only 1 point per333 5x5 m², while more recently acquired datasets contain more334 than 125 points per 5x5 m² raster cell. The data was deliv-335 ered as two separate point clouds: one cloud contained fil-336 tered ground returns, whereas the other cloud contained firsts37 pulses from non-ground objects. All point clouds were de-338 livered as three-column (easting, northing, and height above339

sea level) ASCII files in tiles of 1 km<sup>2</sup>. Before interpolating the point clouds to regular rasters, the clouds were thinned. For the ground data, the mean value of each raster cell in the final resolution of 5x5 m<sup>2</sup> was calculated. For the surface model, both ground and vegetation point clouds were first united, and the maximum value for each raster cell was determined respectively. The combination of both point clouds was necessary in order to avoid large spaces without laser points between vegetated areas that would otherwise have been filled with unrealistic values in the interpolation step. The thinned point clouds were aggregated to larger tiles in order to decrease the number of seamlines in the final mosaic. The aggregated tiles were then interpolated to raster images using a Delauney interpolation in the Matlab software (Mathworks, 2017). The resulting two elevation models were then used to calculate a canopy height model (CHM) in raster format, providing discrete information about the canopy surface height of the forest area in a spatial resolution of 5 meters.

As explanatory variables, the mean canopy height (meanheight) and the standard deviation (stddev) were calculated as the mean and standard deviation of all raster values within a predefined square around each sample plot center. The square (i.e. support of the explanatory variable, see section 2.4) was previously intersected with the state and communal forest area defined by a polygon mask and thereby corrected for edge effects at the forest border. The tree height is one prominent predictor variable in the taper functions of the BWI3 that are used to calculate a timber volume value for each sample tree (Kublin, 2003; Kublin et al, 2013). A visual inspection of the tree volumes of all sample trees collected in the BWI3 within RLP against their tree heights also revealed the characteristic shape of an allometric relationship between these variables (Online Resource 1). It was hypothesized that this relationship on single-tree level is also apparent on the aggregated level of a sample plot and cluster, and can be used within the frame of regression modeling.

The strength of correlation between *meanheight* and timber volume on plot level was expected to show high variation according to the mentioned time-lag up to 10 years between LiDAR acquisition and terrestrial survey. The quality of the height information was also expected to vary according to changing sensor technologies and different point densities used over the years. For these reasons, the LiDAR acquisition year (*lidaryear*) for each sample plot was considered as a potential categorical explanatory variable to explain the variation in the data introduced by these factors. For this purpose, the acquisition year 2008 was further divided into 2008 and 2008\_1. In the latter, the data quality turned out to be very poor due to sensor failures during the acquisition. Additionally, the years 2006 and 2007 as well as 2012 and 2013 were pooled in order to increase the

number of observations per factor level for modelling rea-367 sons. As a result, the *lidaryear* variable comprised nine cat-368 egories (2002, 2003, 2007, 2008, 2008\_1, 2009, 2010, 2011\_369 and 2012).

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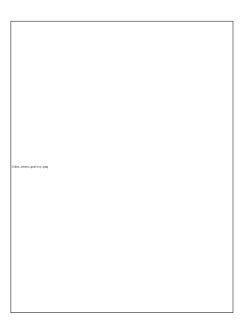
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**Fig. 2:** Separate LiDAR acquisitions in Rhineland-Palatinate over the years. The colors also indicate the quality of the data: light: low point densities  $(1/5x5m^2)$ , dark: high point densities  $(>100/5x5m^2)$ 

# 2.3.2 Tree Species Classification Map

A countrywide satellite-based classification map of the five<sub>395</sub> main tree species (European beech, Sessile and Pedunculate oak, Norway spruce, Douglas fir, Scots pine) described in<sup>396</sup> Stoffels et al (2015) was used to derive tree species infor-<sup>397</sup> mation on sample plot level. The classified tree species map<sup>398</sup> has a grid size of 5 meters and predicts five of the seven tree<sup>399</sup> species that are used in the BWI3 taper functions (Kublin<sup>400</sup> et al, 2013) to calculate the timber volume of a sample tree.<sup>401</sup> Due to unavailable satellite data for the classification, the<sup>402</sup> tree species map excluded one patch with an area of 415 km<sup>2403</sup> in the south-west part of RLP, and two further patches with<sup>404</sup> an area of 76 km<sup>2</sup> and 100 km<sup>2</sup> in the northern part (Stoffels<sup>405</sup> et al, 2015). The tree species information was consequently<sup>406</sup> missing for 407 (7%) of the 5791 sample locations.

# Prediction of main plot tree species

A visual inspection of all BWI3 sample trees of RLP sug<sup>-411</sup> gested that a stratification of the relation between tree height<sub>412</sub> and timber volume according to these seven tree species may<sub>413</sub> provide a considerable reduction in variation within the tree<sub>414</sub> species groups (Online Resource 1). This led to the hypoth-<sub>415</sub> esis that this tree species specific signal might also be appar-<sub>416</sub> ent on sample plot and cluster level and can consequently be<sub>417</sub>

used to increase the accuracy of the prediction model. Based on the tree species classification map, the main tree species of each sample plot was calculated as an additional categorical explanatory variable (*treespecies*) with six categories following a similar approach as Latifi et al (2012): one of the five tree species was assigned as the main plot tree species if its proportion within the edge-corrected support around the sample location exceeded a predefined threshold. If this threshold was not reached by any of the five tree species, the respective sample plot was assigned the category 'Mixed'.

#### Calibration

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Our analyses revealed that the prediction of the main tree species for a sample plot can be subject to misclassifications (section 3.1). Errors in the explanatory variables of linear regression models can however lead to a bias of the regression coefficients in the direction of zero due to an artificial introduction of noise (Carroll et al, 2006). This can cause an inflation of the residual variance and a consequent decrease of the model accuracy (Magnussen et al, 2010). In case of classification the impacts of misclassifications on the model properties are even harder to predict (Gustafson, 2003). While errors in the explanatory variables do not affect the unbiasedness of the estimators in the model-assisted framework, a reduction or elimination of the classification errors could provide an improvement of the regression model accuracy and thereby potentially lead to smaller prediction and estimation errors. We therefore addressed the effect of misclassifications in the treespecies variable by the following analysis:

- a) we investigated the effect on the regression model performance (regression coefficients, model accuracy) when substituting the *predicted* by the *actual* main plot tree species derived from the sampled trees of the respective sample plot under identical threshold settings
- we used the random forest algorithm (Breiman, 2001; Liaw and Wiener, 2002) in the statistical software R (R Core Team, 2016) to define a calibration model in order to improve the classification accuracy of the initially predicted main plot tree species, correct for potential systematic misclassifications and thus minimize the effect of misclassifications on the regression model. The random forest algorithm is a machine learning algorithm that grows a large number of decorrelated classification trees by considering only a subset of all provided predictor variables for each split. In the case of classification, new data are thus predicted by aggregating the predictions of all trees using a majority vote. For our purpose of predicting the actual main tree species of a sample plot (target variable), we provided the random forest algorithm with a full set of p predictor variables that comprised the initial prediction of the main plot tree species

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(treespecies), the mean canopy height (meanheight) and dess standard deviation (stddev) derived from the CHM, the 469 proportion of coniferous trees estimated from the tree 470 species classification map (prop.conif) and the biocli-471 matic growing region (wgb) at the sample location. The 472 algorithm was grown with 2000 trees, considering  $\sqrt{p} \approx 473$  3 of the predictors for each split.

# 2.4 Choice of Support under Angle Count Sampling

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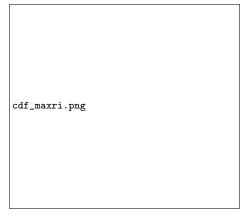
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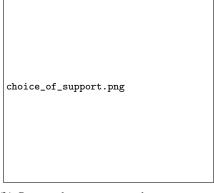
One characteristic of angle count sampling applied in the 479 BWI3 is that a sample plot does not have a fixed radius<sup>480</sup> in which trees are selected (fixed-radius plot), but each tree generates an individual radius from the plot center depending on its diameter at breast height (variable-radius plot). This tree-individual radius is known as the *limiting distance* from the plot center where the tree would still be included in the sample. A consequence of the absence of a fixed plot radius is the question about the optimal support (Hollaus et al, 2007), i.e. the spatial extent around the plot center in which the auxiliary information is evaluated and transformed into an explanatory variable. It has widely been hypothesized that the best relationship between the target variable on the ground and any explanatory variable derived from the auxiliary information is obtained if the support is spatially identical to the sample plot extent. In case of angle count sampling, an individual extent for each sample plot can be approximated by regarding the maximum limiting distances of its sample trees as the outer plot radius. However, many model-assisted applications under double-sampling do not allow for a between-plot change of the support for a specific explanatory variable (Mandallaz, 2013a,b).

For this reason, the task is to find a unique support for each auxiliary information that leads to the best overall model accuracy. Deo et al (2016) conducted extensive analysis to identify optimal supports for modelling standing timber volume for variable-radius plot designs in conifer forests. They analysed 24 different radii (i.e. circular supports) in which they extracted 57 metrics from a LiDAR derived point cloud with an average point density of 18 pulses per square meter. They successively evaluated the prediction performance of each support size by using the LiDAR metrics in a random forest algorithm and comparing the resulting model accuracies. In order to identify the best-performing supports for our explanatory variables, we followed a similar approach. The explanatory variables were calculated using *individual* (i.e. plot-varying) supports (ind), i.e. an individual support extent was used for each plot according to the maximum limiting distance of all sample trees associated to the respective sample plot. We then compared the model accuracies achieved by the individual supports against the model accuracies from a set of fixed

(i.e. non plot-varying) supports. The extents of the fixed supports were chosen from the cumulative distribution function (ECDF) of the maximum limiting distances of all 5791 sample plots of the analysed forest area (Fig. 3a). We considered the  $25^{th}$  (q25, 9 meters),  $50^{th}$  (q50, 12 meters),  $80^{th}$  (q80, 15 meters) and the  $100^{th}$  (q100, 38 meters) percentiles, resulting in support side lengths of 18, 24, 30 and 76 meters (Fig. 3). While in this study we used squares to extract the auxiliary information, also other support-shapes are possible (e.g. circles, hexagons). We also want to emphasize that the use of different support sizes for each explanatory variable is perfectly valid in the infinite population framework of model-assisted estimators (Mandallaz, 2013a,b).



(a) ECDF of maximum limiting distances of all BWI3 sample locations in RLP



(b) Rectangular supports used to extract explanatory variables around sample locations. Dash dot dot line: q100, dash dot line: q80, dot dot line: q50, dot line: q25, solid line: individual support, triangles: sample trees

**Fig. 3:** Identification (a)) and visualization (b)) of potential supports used for calculating the predictor variables on plot level

#### 2.5 Model Validation

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#### 3 Results

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In order to judge the quality of the treespecies variable, the user's accuracy for each classified species and the overall accuracy of the classification scheme was calculated based on the confusion matrix (Congalton and Green, 2008), using the main plot tree species calculated from the sample trees as reference data. The classification accuracy was performed for all support sizes for both the calibrated and the uncalibrated treespecies variables. The measures of the regression model accuracy using both CHM- and treespecies variables were defined as the 10fold cross-validated root mean square error (RMSE<sub>cv</sub>) and the adjusted coefficient of determination (adjusted  $R^2$ ) of the multiple linear regression model defined in equation 1. Additionally, we considered the interaction terms meanheight:treespecies, meanheight<sup>2</sup>:treespecies, mean- $\label{eq:height:lidaryear} \textit{height:lidaryear}, \textit{stddev:lidaryear} \text{ and } \textit{meanheight:stddev}^{522}$ and performed a variable selection based on the Akaike 523 Information Criterion (AIC) (Akaike, 2011) in order to 524 minimize the number of variables in the model. Due 525 to a pronounced unbalanced design in the treespecies-526 lidaryear strata (Online Resource 2), no interaction be-527 tween treespecies and lidaryear was possible. We evalu-528 ated the model for all support combinations, considering the use of individual support sizes for each auxiliary in-530 formation, using both the calibrated and the uncalibrated  $^{531}$ treespecies variable. The calibration model (section 2.3.2)<sup>532</sup> for the *treespecies* variable was recalculated for each respec-533 tive support-threshold setting.

$$Y(x) = \beta_0 + \beta_1 * meanheight + \beta_2 * meanheight^2 + 539$$

$$\beta_3 * stddev + 541$$

$$\beta_4 * lidaryear_1 + ... + \beta_{12} * lidaryear_9 + 542$$

$$\beta_{13} * treespecies_1 + ... + \beta_{18} * treespecies_6 + e(x)$$

$$(1)$$

206 sample plots included no sample trees and the tim-544 ber volume density Y(x) was thus set to zero. These zero-545 plots were removed from the modeling dataset since they546 acted as leverage points in cases where the LiDAR heights47 metrics were recorded long before the terrestrial survey. To-548 gether with the missing tree species information (section549 2.3.2), the modeling dataset was limited to 5206 observa-550 tions.

# 3.1 Classification Accuracies

Effect of Support Size and Threshold

Before calibration, the lowest user's accuracies (UA) for most tree species were realized using high thresholds of 80% and 100% for deciding the main tree species on the plot level (figure 4a). A plausible reason for this is that raising the threshold to higher values (e.g. 80%, 100%) distinctively increases the probability of the reference class (based on the sample trees of the sample location) to be assigned as class 'Mixed', while the much coarser spatial resolution of the tree species map causes the predicted class to remain classified as one of the five tree species. However, as the support size is increased, so does the number of tree species raster cells to be evaluated at the sample location, thereby increasing the probability that the predicted class will be 'Mixed'. For this reason, most tree species exhibit an increase in user's accuracy under higher thresholds with higher support sizes. This scale-threshold dependency of the user's accuracy particularly affects tree species that most commonly occur in mixed forest stands in Rhineland-Palatinate (Scots pine, oak and beech), whereas the user's accuracies for tree species that are mostly prominent in pure forest stands (spruce, Douglas fir) logically turned out to be much more robust to changes in the thresholds and support sizes.

Among the uncalibrated tree species predictions, *beech* and *spruce* produced the best predictions achieving UAs of up to 70% and 80%. Although the predictions for *Douglas fir* and *Scots pine* generally performed less well than *beech* and *spruce*, similar UAs can be produced by adjusting the threshold and support choices. UAs for *oak* never performed better than 50%. A detailed table of the user's and overall accuracies is provided in Online Resource 3.

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cos\_tspec\_mixed\_cal\_nocal.png (a) cos\_tspec\_oaa\_mixed.png

**Fig. 4:** Classification accuracy for the main tree species of a sample location *before* and *after* calibration: *a)* user's accuracies. *b)* overall<sup>595</sup> accuracies. *n*: number of validation data per class

**(b)** 

Effect of Calibration

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Calibration substantially diminished the effect of the scale-601 threshold dependency for the five tree species and also in-602

creased the UAs for *Scots pine* and *oak*. Whereas the UA level of *spruce* remained unchanged, the UAs for *beech* were found to be slightly lower after calibration. The overall accuracy under each support choice was always considerably increased by calibrating the tree species prediction (figure 4b). With respect to the calculated random forest models, the initial tree species prediction (*treespecies*) and the information about the growing region (*wgb*) turned out to be the most valuable information, followed by the estimated proportion of coniferous trees (*prop.conif*) and the mean canopy height (*meanheight*).

#### 3.2 Regression Model Accuracies

Effect of Support Size and Threshold

Figure 5 shows the accuracies of the regression model (equation 1) achieved under all possible combinations of support sizes for the auxiliary data. The stepwise selection procedure always included all considered single and interaction terms. In terms of adjusted  $R^2$  and RMSE<sub>cv</sub>, the analysis revealed that the choice of the CHM support size controls the overall level of the model's accuracy. The information about the main plot tree species can then be used to further improve the model fit under suitable treespecies support and threshold settings. When using the uncalibrated treespecies variable, an increase of the treespecies support size causes an increase in the model performance if low thresholds are used, whereas high thresholds (80%, 100%) cause a decrease in the model performance. This threshold-dependency could be removed by calibrating the treespecies variable. The highest adjusted  $R^2$  and the lowest RMSE<sub>cv</sub> were realized using the q50 support for the CHM variables in combination with the q100 support and a threshold of 100% for the calibrated *treespecies* variable (adjusted  $R^2$ =0.48 and RMSE<sub>cv</sub>=137 m<sup>3</sup>/ha). However, various support and threshold combinations for the CHM and treespecies variables can be used to yield almost identical RMSE<sub>cv</sub> and adjusted  $R^2$ values. A detailed table of the model accuracies is given in Online Resource 4.

# Effect of Misclassifications

We can assess magnitude of the misclassification effect by comparing the adjusted  $R^2$ 's of models that use the predicted tree species (calibrated and uncalibrated) as an explanatory variable to models that use the error-free tree species variables acquired from the terrestrial survey. Note that only the model with the predicted tree species variables can be applied to additional sample locations where no terrestrial survey has been carried out. Figure 6 provides a comparison of the adjusted  $R^2$  achieved under the use of the error-free tree species predictor variable against the adjusted  $R^2$  realized

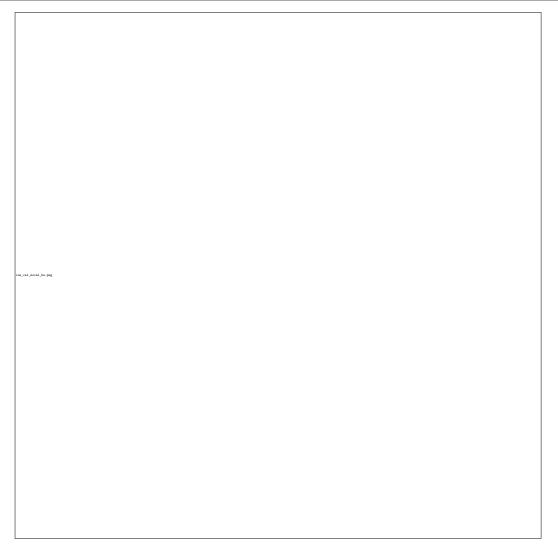


Fig. 5: 10-fold RMSE<sub>rv</sub> and adjusted  $R^2$  realized under various support choices for the CHM and treespecies explanatory variables

under the use of the tree species variable containing miss-621 classifications. This analysis was carried out for all models622 that were analysed in section 3.2, i.e. for all possible support<sub>623</sub> and threshold combinations for the CHM and treespecies 624 predictor variables. As expected, the highest adjusted  $R^2$  for every evaluated<sub>626</sub> model was always achieved using the error-free tree species<sub>627</sub> variable, whereas the missclassifications in the tree species<sub>628</sub> variable led to a systematic decrease of the model accuracy.629 This is in agreement with the potential effects of erroneous<sub>630</sub> explanatory variables discussed in Carroll et al (2006) and<sub>631</sub> Gustafson (2003), i.e. an increase of variability (noise) in the data that can increase the amount of unexplainable variance and thereby reduce the model accuracy. The calibration of the initially predicted main plot tree species using the random forest classification algorithm <sup>633</sup> (section 2.3.2) turned out to not only improve the classifi-634 cation accuracies (section 3.1), but also to considerably de-635

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crease the effect of the missclassifications on the regression model predictions and accuracy. Figure 6 (right) shows that the adjusted  $R^2$  under the actual and the calibrated predicted tree species variable are in general much closer to, and in many cases even on the identity line. Whereas the misclassifications in the uncalibrated *treespecies* variable led to a residual inflation of 1% - 5%, it was only between 0% and 1% after calibrated *treespecies* variables, the regression coefficients were almost identical to the ones received using the actual main plot tree species.

# 3.3 Final Regression Model

In order to address research questions 1 and 2 (i.e. the gain in model accuracy by tree species information and effect of heterogeneity in the LiDAR data), we investigated the model properties in more detail. For this purpose, we decided to

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Fig. 6: Effect on the adjusted  $R^2$  when substituting the actual main tree species with the predicted main tree species of a sample plot. The differentiation into two distinct point clouds results from the poor model performance under support size q100 for the CHM variables (i.e. the *lower* point cloud). The *dotted* line tracks the model with the highest adjusted  $R^2$  under the use of the error-free *treespecies* variable

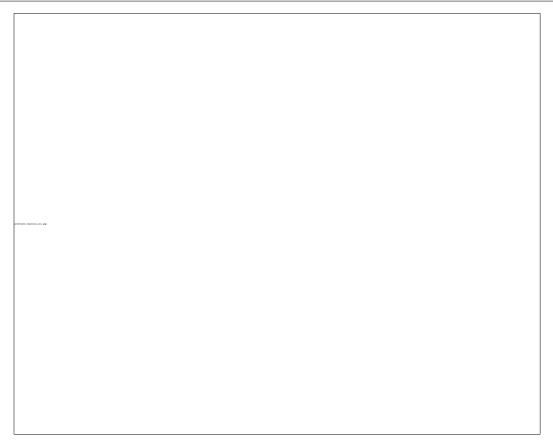
use the support settings of q50 for both auxiliary data with  $a_{670}$  threshold of 100% for the tree species variable as the regres- $_{671}$  sion model of choice. The reason for this choice was that a) $_{672}$  the model provided almost the highest adjusted  $R^2$  among all $_{673}$  validated models while reducing the data handling complex- $_{674}$  ity for upcoming applications (i.e. identical support sizes fo $_{1675}$  all remote sensing data) and b) the calibration neutralized  $_{676}$  the effects of misclassifications on the model predictions  $_{677}$  The interaction term between  $meanheight^2$  and  $treespecies_{678}$  (i.e. considering separate curvatures for each tree species) $_{679}$  turned out not to have any influence on the model accuracy $_{680}$  and was thus dropped, resulting in an adjusted  $R^2$  of  $0.49_{681}$  and  $RMSE_{CV}$  of 132.12 m $^2$ /ha.

# Interpretation of Final Regression Model

Figure 7 provides a visualisation of the tree species prediction functions separated by the LiDAR acquisition years. Sample plots classified as oak and Scots pine revealed to have an almost identical relationship (nearly identical slopes) for the mean canopy height - timber volume relationship. They only differ by a marginally higher intercept <sup>684</sup> for Scots pine plots, meaning that given the same mean canopy height a sample plot dominated by Scots pine yields.85 a marginally higher timber volume on the plot level than a686 plot dominated by oak. Beech-dominated sample plots tend687 to achieve a higher timber volume than oak and Scots pine688 for canopy heights below 20 meters, but realize the lowest<sub>689</sub> timber volumes for canopy heights above 20 metres. Sam-690 ple plots dominated by any of the remaining coniferous tree<sub>591</sub> species (Douglas fir, spruce) revealed to have higher slopes<sub>692</sub> than broadleaf classified plots. This indicates that given the giv same mean canopy height, sample plots dominated by Dou-694 glas fir and spruce yield higher timber volume values than695 broadleaf- or Scots pine dominated sample plots, and this 696 difference becomes more pronounced with increasing mean canopy heights. Within the group of coniferous-dominated sample plots, *spruce* turned out to have the highest slope, thereby yielding the highest timber volume values for mean canopy heights above 15 meters. An undesired characteristic of the model is that the predicted timber volume can in some cases (< 1%) take negative values for low canopy heights (e.g. for *spruce*-dominated plots with *meanheight* below 5 meters and *stddev* of 4 meters). However, we chose not to use a log-transformation of the response variable. Doing so would have prevented the subsequent calculation of the gweight variance of the model assisted estimators (Mandallaz, 2013a; Mandallaz et al, 2013), which is only possible for response variables on the original scale.

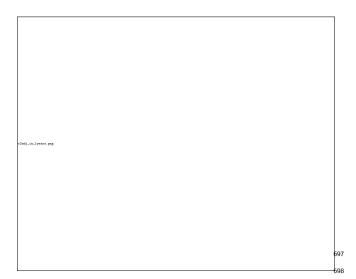
# Effect of Time-Lags and Heterogeneity in LiDAR Data

Incorporating the LiDAR acquisition year as a categorical variable (lidaryear) in the regression model substantially accounted for the variability in the data introduced by a) the time-lags between LiDAR acquisition and terrestrial survey, and b) variation in LiDAR data quality which are due to sensor- and post processing techniques (table 2). Whereas the adjusted  $R^2$  for the regression model without considering the LiDAR acquisition year as additional predictor variable was 0.35 (0.40 including the tree species variable), the stratification by the LiDAR acquisition year led to adjusted  $R^2$  of 0.44 (0.48), thereby increasing the proportion of explained variance by up to 8%.



**Fig. 7:** Visualization of the timber volume prediction function (*final regression model*) on sample plot level for each main plot tree species and LiDAR acquisition year. For visualization purposes, the predictor variable *stddev* was set to its average value within the respective *treespecies* and *lidaryear* group

. The terrestrially observed timber volume values are plotted in the background.



**Fig. 8:**  $R^2$  of the final regression model achieved *within* the LiDAR<sup>699</sup> acquisition year strata. *Grey points:*  $R^2$  of submodel 1 (no stratifica-700 tion according to LiDAR acquisition year or tree species). *Crosses:*  $R^2_{701}$  of submodel 2 (*without* tree species stratification). *Filled triangles:*  $R^2_{701}$  of final model using the *uncalibrated* tree species variable. *Empty triangles:*  $R^2$  of final model using the *calibrated* tree species variable. *Tosses:*  $R^2_{702}$  achieved using the error-free tree species variable (de-704 rived from sample trees). *Dotted line:* Overall adjusted  $R^2$  of submodel  $R^2_{702}$  of final model using the *calibrated* tree species variable. *Two-dashed line:* Overall adjusted  $R^2_{702}$  of final model using the *error-free* tree species variable

**Table 1:**  $R^2$ , RMSE and Residual Square Sum (SSE) of final regression model within LiDAR acquisition year strata (*lidaryear*). n: number of validation data

LiDAR acquisition year	$R^2$	rmse	SSE	n
2012	0.55	139.54	7535278	387
2011	0.55	145.21	17880553	848
2010	0.48	122.16	16907662	1133
2009	0.43	127.17	8652419	535
2008_1	0.34	170.49	11161424	384
2008	0.50	124.39	10475374	677
2007	0.49	129.72	6950192	413
2003	0.32	146.37	11139814	520
2002	0.43	139.79	6038601	309

We further analysed the model residuals within each Li-DAR acquisition year (within-group variation) for the final model and nested submodels. It turned out that the  $R^2$  vary distinctly between the LiDAR acquisition year strata (figure 8). More precisely, the within-group  $R^2$  can be higher and lower than the overall  $R^2$  of the respective model. Figure 8 shows that a stratification according to the LiDAR acquisition years (submodel 2) can already increase the  $R^2$  in most acquisition year strata, compared to the basic model using

only the LiDAR height metrics as predictor variables (sub- $_{754}$  model 1). In some LiDAR acquisition year strata (i.e. 2007, $_{755}$  2008), this increase in  $R^2$  even reached 7% - 11%. The ac- $_{756}$  curacies for the final model are also given in table 1.

# Added Value of Tree Species Map Information

Introducing the predicted main tree species of a sample plot<sub>761</sub> as an additional categorical variable to submodel 2 yielded<sub>762</sub> a further 4% increase in the adjusted  $R^2$  (table 2). This im<sub>763</sub> provement was even more pronounced in LiDAR acquisi<sub>764</sub> tion years close or identical to the year of the terrestrial<sub>765</sub> inventory (up to 7% increase in  $R^2$ , figure 8. The analy<sub>766</sub> sis illustrated once more that misclassifications in the tree<sub>767</sub> species variable generally reduce model accuracy compared<sub>768</sub> to using error-free tree species information. The residual in<sub>769</sub> flations caused by the misclassifications in the uncalibrated<sub>770</sub> treespecies variable within the *lidaryear* strata were up to<sub>771</sub> 5%. However, the calibration was able to substantially de<sub>772</sub> crease or even remove the effects of misclassifications on<sub>773</sub> the model accuracy in all LiDAR acquisition year strata.

#### 4 Discussion

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# 4.1 Stratification according to Tree Species and LiDAR Acquisitions

Incorporating the main tree species of a sample location in<sub>779</sub> the timber volume regression model significantly increased,80 the model accuracy and revealed strong evidence for the ex-781 istence of a tree species specific behaviour concerning tim-782 ber volume on the plot level. This result seems reasonable<sub>783</sub> regarding the species specific taper functions on single-tree<sub>784</sub> level applied in the BWI3 (Kublin, 2003; Kublin et al, 2013).785 Further evidence and specification of the tree species effects<sub>786</sub> on sample plot level - up to modeling individual tree species<sub>787</sub> - would be desirable. However, this was not possible in our po study because the stratification according to the LiDAR ac-789 quisition years severely limited the flexibility of species-790 specific prediction functions and model interpretability. In<sub>791</sub> particular, using the LiDAR acquisition years as categorical<sub>792</sub> variables led to highly unbalanced datasets when stratify-793 ing according to the main plot tree species, and prevented, 94 the use of further stratification variables such as bioclimatic<sub>795</sub> growing regions due to confounding effects and consequent<sub>796</sub> singularities in the design matrices. A stratification to the<sub>797</sub> LiDAR acquisition years however proved to be a means in<sub>798</sub> accounting for the artificially introduced noise in the data<sub>799</sub> caused by quality variations and the large time-lags between 800 the remote sensing and terrestrial data. Incorporating the cal-801 ibrated tree species information further improved the model<sub>802</sub> accuracy by 4% in adjusted  $R^2$ . Compared to the simple<sub>803</sub> model only containing LiDAR height metrics, including the 804 LiDAR quality and calibrated tree species information increased the adjusted  $R^2$  by 13% in total. A differentiated evaluation of the final regression model revealed that the highest  $R^2$  where achieved within LiDAR acquisitions year strata identical with the year of the terrestrial survey. Also the gain in the  $R^2$  by including the tree species information was largest (i.e. 7%) in combination with LiDAR information acquired in the year of the terrestrial inventory. These insights were particularly interesting with respect to the further use of the regression model for small area estimations. Small area estimators generally gain modeling strength by defining the prediction model globally (i.e. using all data in the inventory area), and then applying the so-derived prediction model to a subset of observations located within the area of interest (Mandallaz et al, 2016). Consequently, the proposed stratification technique in the prediction model could be expected to yield a gain in model accuracy and a reduction of the small area estimation errors if the small area domain mostly includes data from strata that have high withinstrata model accuracies. This hypothesis is subject to ongoing analysis.

# 4.2 Calibration of Tree Species Map Information

The accuracy assessment of the initially derived main plot species from the classification map revealed the presence of misclassifications that led to a decrease in model accuracy. One reason for the misclassifications were that the classification algorithm of Stoffels et al (2015) was exclusively trained in pure stands with the objective to predict the dominant tree species of a forest stand. Thus, our requirements on the classification map differed considerably from the ones imposed by Stoffels et al (2015) and have to be considered as far more difficult to meet. Firstly, the reference data used in the accuracy assessment also included understory trees that were recorded in the BWI3 sample. Secondly, determining an exact spatial validation unit for a sample location (support) is not possible due to the properties of angle count sampling (section 2.4). Thirdly, distinct discrepancies in the spatial scale between the reference data and the classification map severely hamper exact predictions of the main plot tree species especially in mixed forest stands. The latter issue caused a pronounced dependency of the user's accuracy on the support and threshold choice, particularly for tree species that most commonly occur in mixed forest structures, i.e. Scots pine (91%), oak (90%) and beech (85%) (von Thünen-Institut, 2014). With respect to this set-up, the application of our calibration method proved to be of high value. It led to an increase in the classification accuracies, particularly for those tree species that performed worse in the uncalibrated setup, and thereby successfully minimized and even removed the deleterious effect of misclassifications on model accuracy and regression coefficients. We consider

 $R_{ad}^2$ model terms model  $RMSE_{cv}$ parameters meanheight + stddev + meanheight<sup>2</sup> + treespecies + lidaryear + meanheight:treespecies + meanheight:lidaryear + meanheight:stddev + stddev:lidaryear final model 39 0.48 137.49 meanheight + stddev + meanheight<sup>2</sup> + submodel 1 5 0.35 153.02 meanheight:stddev meanheight + stddev + meanheight<sup>2</sup> + lidaryear + meanheight:lidaryear + submodel 2 29 0.44 142.82 meanheight:stddev + stddev:lidaryear meanheight + stddev + meanheight<sup>2</sup> + treespecies + meanheight:treespecies + meanheight:stddev submodel 3 15 0.40 142.82

Table 2: Accuracy metrics for submodels of final OLS regression model

this *a posteriori calibration* a valuable method for future<sup>339</sup> studies where an external tree species map (i.e. the map was<sup>840</sup> not created for the specific study objective) is used in pre-<sup>841</sup> diction models. Whereas the extensive analysis in our study<sup>842</sup> deepened the understanding of the afore mentioned scale-<sup>843</sup> effects, an alternative method for future applications could<sup>844</sup> be to use map-derived percentages of each tree species as<sup>845</sup> predictor variables in the random forest algorithm in order<sup>346</sup> to directly predict the terrestrially observed main plot tree<sup>347</sup> species.

# 4.3 Choice of Support under Angle Count Sampling

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The validation of different support sizes underlined that<sub>851</sub> the support choice can impact prediction accuracy. In the present study, differences in the model accuracies turned out852 to be small for most support choices. An exception was the853 choice of the q100 support for the CHM derived variables<sub>854</sub> (76 meter side length), where the model accuracy was con-855 siderably worse than what was achieved under optimal set-856 tings. With the exception of the latter, the accuracy differ-857 ences according to adjusted  $R^2$  and RMSE<sub>cv</sub> were very sim-858 ilar to those found by Deo et al (2016) when evaluating the859 model performance of optimal support sizes for a range of 860 various basal area factors. An analysis to find the best sup-861 port settings therefore seems to be advisable prior to further<sub>862</sub> applications of model-assisted or model-dependent inven-863 tory methods so as not to lose model accuracy by unsuitable864 support choices. The concept of the demonstrated analysis<sub>865</sub> method for identifying suitable supports can be transferred<sub>866</sub> to any kind of auxiliary information, predictor variable and 867 prediction model.

Contrary to our hypothesis, the use of plot-individual<sub>869</sub> supports did not yield the best prediction performances. A<sub>870</sub> plausible reason for this is that determining an exact plot<sub>871</sub> radius under angle count sampling is technically infeasible,<sub>872</sub>

and thus, angle count sampling does not seem to be adequate when linking inventory information with remote sensing data. However, the extensive analysis carried out in our study indicated that the optimal support size depends on the spatial resolution of the remote sensing data as well as the context in which the derived information is used in the prediction model. In the case of transforming the tree species information map into a suitable categorical predictor variable, the use of a large support size of 76 meter side length turned out to yield the best model accuracy. However, only few sample locations in the study area were actually characterized by limiting circles of that particular size.

# 5 Conclusion

The objective of this study was to identify a suitable ordinary least square regression model that can be applied over the entire forest area of Rhineland-Palatinate using modelassisted estimators. The large amount of data that was gathered in the frame of this study allowed for extensive modeling possibilities, but had the side effect of contributing to high heterogeneity in the response and explanatory variables. Whereas the variability of the response variable (timber volume on plot level) is due to the very heterogeneous forest structures and bioclimatic growing regions in RLP, a considerable amount of heterogeneity in the explanatory variables was introduced by quality restrictions in the remote sensing data. This was particularly true for the LiDAR derived canopy height information that was gathered in a time span of ten years around the date of the terrestrial inventory and revealed pronounced quality variations. With an adjusted  $R^2$  of 0.48 and a RMSE<sub>cv</sub> of 137 m<sup>3</sup>/ha, the model accuracy was still very close to those found in similar studies (Maack et al, 2016). Our analyses strongly indicate that the acquisition of the auxiliary information close to the date of the terrestrial survey is a key factor in or-

der to increase the model accuracy. We also expect the tree<sub>903</sub> species information in the timber volume model to become<sup>904</sup> even more relevant if the temporal synchronocity and  $\ensuremath{\text{the}}^{905}$ quality of the canopy height information is improved. An<sub>907</sub> up-to-date canopy height model would also circumvent as a solution and a solution and a solution as a solution and a solution and a solution as a solution and a solution are a solution as a solution stratification according to different LiDAR acquisition char-909 acteristics, lead to a more balanced dataset when stratify-910 ing for the main plot tree species and allow for incorporating information that can further explain the variation within 13 each tree species group. With respect to the latter, informa\_914 tion about the bioclimatic growing conditions, soil proper-915 ties and the stand density on plot level are expected to further improve the model's predictive performance. Promising918 steps with respect to more up-to-date auxiliary information919 have already been made, as the topographic survey institu-920 tion of RLP is currently processing a canopy height model from aerial imagery acquisitions for 2011 and 2012 covering the entire federal state. These aerial photography acqui-  $^{921}\,$ sitions will in the future be conducted in a two-year period, allowing to derive up-to-date canopy height information in the framework of future forest inventories. As the availability of countrywide imagery-based surface models has been increasing (Ginzler and Hobi, 2015), investigating the performance between areal and LiDAR derived canopy height models and their consequent predictive power in the frame of timber volume estimations (Ullah et al, 2017) are tasks for subsequent analysis. Additionally, availability of satellite data for tree species classification map production with respect to up-to-dateness and coverage has recently been increasing in the frame of the Sentinel-2 mission (ESA, 2017).

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Conflict of Interest The authors declare that they have no conflict of interest.

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