1 Matrices

- Have rank equal to the number of linearly independent columns/rows
- Have rank equal to the dimension of their column space
- Have singular value decompositions **UDV**'
- Have a generalized inverse
- Have a column space and a row space
- Have null space $\{x : Ax = 0\}$

1.1 Orthogonal projections

- Map vectors onto a subspace
- Are symmetric
- Are idempotent
- Are positive semidefinite
- Have eigenvalue 1 with multiplicity equal to their rank; all other eigenvalues are zero
- For any A, $P_A = A(A'A)^-A'$ is an orthogonal projection onto C(A)
- $\bullet \ P_AA = A$
- $rank(\mathbf{P_A}) = rank(\mathbf{A}) = rank(\mathbf{U})$ where \mathbf{U} comes from the SVD of \mathbf{A}
- I P is an orthogonal projection matrix as well

1.2 Square matrices

- Are invertible if they have full rank
- Are invertible if they have empty null space
- Have rank equal to the number of nonzero eigenvalues
- Have trace equal to the sum of their diagonal elements

1.3 Symmetric matrices

- Have a spectral decomposition **UDU**'
- Have trace equal to the sum of their eigenvalues
- Have determinant equal to the product of their eigenvalues
- Are positive definite if all their eigenvalues are positive
- Are positive semidefinite if all their eigenvalues are nonnegative

1.4 Matrix identities

- AXX' = BXX' implies AX = BX
- $\mathbf{A}\mathbf{A}' = 0$ implies $\mathbf{A} = 0$
- Block matrix inversion

$$\begin{bmatrix} \mathbf{A} & \mathbf{B} \\ \mathbf{B}' & \mathbf{C} \end{bmatrix}^{-1} = \begin{bmatrix} (\mathbf{A} - \mathbf{B}\mathbf{C}^{-1}\mathbf{B}')^{-1} & -(\mathbf{A} - \mathbf{B}\mathbf{C}^{-1}\mathbf{B}')^{-1}\mathbf{B}\mathbf{C}^{-1} \\ -\mathbf{C}^{-1}\mathbf{B}'(\mathbf{A} - \mathbf{B}\mathbf{C}^{-1}\mathbf{B}')^{-1} & \mathbf{C}^{-1} + \mathbf{C}^{-1}\mathbf{B}'(\mathbf{A} - \mathbf{B}\mathbf{C}^{-1}\mathbf{B}')^{-1}\mathbf{B}\mathbf{C}^{-1} \end{bmatrix}$$

- Invertible **A** implies $(\mathbf{A}^{-1})' = (\mathbf{A}')^{-1}$
- For $m \times n$ matrix A, rank(\mathbf{A}) + nullity(\mathbf{A}) = n
- For invertible **A** and **C**, $rank(\mathbf{AB}) = rank(\mathbf{BC}) = rank(\mathbf{BC})$

2 Linear model

- $\mathbf{Y} = \mathbf{X}\boldsymbol{\beta} + \boldsymbol{\varepsilon}$ where $\mathbb{E}[\boldsymbol{\varepsilon}|\mathbf{X}] = \mathbf{0}$
- Sometimes we assume $\varepsilon | \mathbf{X} \sim (\mathbf{0}, \sigma^2 \mathbf{I})$

2.1 Least squares

- $\underset{\beta}{\operatorname{arg\,min}}(\mathbf{Y} \mathbf{X}\beta)'(\mathbf{Y} \mathbf{X}\beta)$
- Solve the normal equations $\mathbf{X}'\mathbf{X}\beta = \mathbf{X}'\mathbf{Y}$
- Fitted values are the same for all least squares estimates
- Is unique for full rank design matrix
- Can be of the form $(\mathbf{X}'\mathbf{X})^{-}\mathbf{X}'\mathbf{Y}$
- Is unique under identifiability constraints

2.2 Statistical properties

- $a'\beta$ is estimable if a is in the row space of **X**
- Estimable $a'\beta$ implies that $a'\hat{\beta}$ is BLUE
- For full rank design matrix, $a'\hat{\beta}$ is BLUE
- θ is identifiable if $\theta \neq \theta_0$ implies $f_{\theta} \neq f_{\theta_0}$
- For $\hat{\beta} = (\mathbf{X}'\mathbf{X})^{-}\mathbf{X}'\mathbf{Y}$, $a'\hat{\beta}$ is unbiased
- Additionally, if errors are uncorrelated, $var(a'\hat{\beta}|\mathbf{X}) = \sigma^2 a'(\mathbf{X}'\mathbf{X})^- a$
- $\hat{\sigma}^2 = \frac{1}{n \text{rank}(\mathbf{X})} (\mathbf{Y} \mathbf{X}\hat{\beta})' (\mathbf{Y} \mathbf{X}\hat{\beta})$ is unbiased
- $\hat{\beta} \stackrel{p}{\to} \beta$ if $\lambda_{\min}(\mathbf{X}'\mathbf{X}) \to \infty$

2.3 Examples

- $\mathbf{X} = \mathbf{1}$ $\mathbf{P}_{\mathbf{X}} = \frac{1}{n} \mathbf{1} \mathbf{1}'$ $\bar{Y} \text{ is BLUE}$ $\mathbf{P}_{\mathbf{X}} \mathbf{Y} = (\bar{Y}, \dots, \bar{Y})'$
- Balanced one-way ANOVA
 - **X** is rank-deficient

$$-\mathbf{P}_{\mathbf{X}} = \frac{1}{J} \begin{bmatrix} \mathbf{1}\mathbf{1}' & \mathbf{0} \\ \mathbf{0} & \mathbf{1}\mathbf{1}' \end{bmatrix}_{2J \times 2J}$$
$$-\mathbf{P}_{\mathbf{X}}\mathbf{Y} = (\bar{Y}_1, \dots, \bar{Y}_1, \bar{Y}_2, \dots, \bar{Y}_2)'$$

• Simple linear regression with intercept

$$- \hat{\beta}_{0} = \bar{Y} - \hat{\beta}_{1}\bar{X} - \hat{\beta}_{1} = \frac{\sum ((X_{i} - \bar{X})Y_{i})}{\sum (X_{i} - \bar{X})^{2}}$$

• Column-centered linear regression with intercept

$$- \hat{\alpha}_0 = \bar{Y}$$
$$- \hat{\alpha} = (\mathbf{X}'(\mathbf{P_X} - \mathbf{P_1})\mathbf{X})^{-1}\mathbf{X}'(\mathbf{P_X} - \mathbf{P_1})\mathbf{Y}$$