

Problem 1.

$$y_i \stackrel{\text{iid}}{\sim} N(0, \text{diag}(\alpha_1, \dots, \alpha_d)^{-1})$$

$$\alpha_k \stackrel{\text{iid}}{\sim} \text{Gamma}(a_0, b_0)$$

$$\lambda \sim \text{Gamma}(e_0, f_0)$$

$$\text{Gamma}(\eta | \tau_1, \tau_2) = \frac{\tau_2^{\tau_1}}{\Gamma(\tau_1)} \eta^{\tau_1-1} e^{-\tau_2 \eta}$$

$$q(w, \alpha_1, \dots, \alpha_d, \lambda) \approx p(w, \alpha_1, \dots, \alpha_d, \lambda | y, x)$$

$$q(w, \alpha_1, \dots, \alpha_d, \lambda) = q(w) q(\lambda) \prod_{k=1}^d q(\alpha_k)$$

Since  $\alpha_k$  are independent, assume  $q(\alpha) = \prod_{k=1}^d q(\alpha_k) = \prod_{i=1}^d \text{Gam}(\alpha_i | a_0, b_0)$

$$= \prod_{i=1}^d \frac{1}{\Gamma(a_0)} b_0^{a_0} \alpha_i^{a_0-1} \exp(-b_0 \alpha_i)$$

Joint likelihood:  $p(y, w, \alpha, \lambda | x) = p(y | w, \lambda, x) p(w | \alpha) p(\alpha) p(\lambda)$

$$\ln q(\lambda) = \mathbb{E}_{q(\alpha, w)} \ln p(y | w, \lambda, x) + \underbrace{\mathbb{E}_{q(\alpha, w)} \ln p(w | \alpha)}_{\text{const w.r.t } \lambda} + \underbrace{\mathbb{E}_{q(\alpha, w)} \ln p(\alpha)}_{\text{const w.r.t } \lambda} + \mathbb{E}_{q(\alpha, w)} \ln p(\lambda)$$

$$= \mathbb{E}_{q(\alpha, w)} \ln p(y | w, \lambda, x) + \mathbb{E}_{q(\alpha, w)} \ln p(\lambda) + C$$

$$= \mathbb{E}_{q(\alpha, w)} \left[ \sum_{i=1}^n \left( \frac{1}{2} \ln \lambda - \frac{1}{2} \ln(2\pi) - \frac{\lambda}{2} (y_i - w^T x_i)^2 \right) \right] + \mathbb{E}_{q(\alpha, w)} \left[ \underbrace{e_0 \ln f_0 - \ln \Gamma(e_0) + (e_0 - 1) \ln \lambda - f_0 \lambda}_{C} \right] + C$$

$$= \frac{n}{2} \ln \lambda - \frac{\lambda}{2} y^T y + \lambda \mathbb{E}_{q(w)} [w^T] x^T y - \frac{\lambda}{2} \mathbb{E}_{q(w)} [w^T] x^T x \mathbb{E}_{q(w)} [w] + (e_0 - 1) \ln \lambda - f_0 \lambda + C$$

$$= (e_0 + \frac{n}{2} - 1) \ln \lambda - \left\{ \frac{1}{2} y^T y - \mathbb{E}_{q(w)} [w^T] x^T y + \frac{1}{2} \mathbb{E}_{q(w)} [w^T] x^T x \mathbb{E}_{q(w)} [w] + f_0 \right\} \lambda + C$$

$$\Rightarrow q^*(\lambda) = e^{c_1 \ln \lambda - c_2 \lambda + c_3} \text{ is in the form of Gamma distribution } = \text{Gamma}(\lambda | e_n, f_n)$$

$$\text{with } \begin{cases} e_n = e_0 + \frac{n}{2} \\ f_n = f_0 + \frac{1}{2} y^T y - \mathbb{E}_{q(w)} [w^T] x^T y + \frac{1}{2} \mathbb{E}_{q(w)} [w^T] x^T x \mathbb{E}_{q(w)} [w] \end{cases}$$

$$\ln q(w) = \mathbb{E}_{q(\alpha, \lambda)} \ln p(y | w, \lambda, x) + \mathbb{E}_{q(\alpha, \lambda)} \ln p(w | \alpha) + \underbrace{\mathbb{E}_{q(\alpha, \lambda)} \ln p(\alpha)}_{\text{const w.r.t } w} + \underbrace{\mathbb{E}_{q(\alpha, \lambda)} \ln p(\lambda)}_{\text{const w.r.t } w}$$

$$= \sum_{i=1}^n \mathbb{E}_{q(\alpha, \lambda)} \left[ \underbrace{\frac{1}{2} \ln \lambda - \frac{1}{2} \ln 2\pi - \frac{\lambda}{2} (y_i - w^T x_i)^2}_{C} \right] + \mathbb{E}_{q(w)} \left[ \underbrace{-\frac{1}{2} \ln |\text{diag}(\alpha_1, \dots, \alpha_d)^{-1}| - \frac{d}{2} \ln 2\pi - \frac{1}{2} w^T \text{diag}(\alpha_1, \dots, \alpha_d) w}_{\text{const w.r.t } w} \right] + C$$

$$= - \frac{\mathbb{E}_{q(w)}[\lambda]}{2} \sum_{i=1}^n (y_i - w^T x_i)^2 - \frac{\mathbb{E}_{q(w)}[\text{diag}(\alpha_1, \dots, \alpha_d)]}{2} w^T w + c$$

$$= - \underbrace{\frac{\mathbb{E}_{q(w)}[\lambda]}{2} y^T y}_c + \mathbb{E}_{q(w)}[\lambda] w^T x^T y - \frac{\mathbb{E}_{q(w)}[\lambda]}{2} w^T x^T x w - \frac{\mathbb{E}_{q(w)}[\text{diag}(\alpha_1, \dots, \alpha_d)]}{2} w^T w + c$$

$$= - \frac{1}{2} w^T [\mathbb{E}_{q(w)}[\lambda] x^T x + \mathbb{E}_{q(w)}[\text{diag}(\alpha_1, \dots, \alpha_d)]] w + w^T [\mathbb{E}_{q(w)}[\lambda] x^T y] + c$$

$\Rightarrow q(w)$  is in the form of Normal distribution  $= N_d(w | \mu_n, \Sigma_n)$ ,

$$\text{with } \begin{cases} \mu_n = \sum_n \mathbb{E}_{q(w)}[\lambda] x^T y \\ \Sigma_n = (\mathbb{E}_{q(w)}[\text{diag}(\alpha_1, \dots, \alpha_d)] + \mathbb{E}_{q(w)}[\lambda] x^T x)^{-1} \end{cases}$$

$$\ln q(\alpha) = \underbrace{\mathbb{E}_{q(w, \lambda)} [\ln p(y | w, \lambda, x)]}_{\text{const w.r.t } \alpha} + \mathbb{E}_{q(w, \lambda)} [\ln p(w | \alpha)] + \mathbb{E}_{q(w, \lambda)} [\ln p(\alpha)] + \underbrace{\mathbb{E}_{q(w, \lambda)} [\ln p(\lambda)]}_{\text{const w.r.t } \alpha}$$

$$= \mathbb{E}_{q(w)} [\ln p(w | \alpha)] + \mathbb{E}_{q(w, \lambda)} [\ln p(\alpha)] + c$$

$$= \mathbb{E}_{q(w)} \left[ -\frac{1}{2} \ln |\text{diag}(\alpha_1, \dots, \alpha_d)| - \underbrace{\frac{d}{2} \ln 2\pi}_c - \frac{1}{2} w^T \text{diag}(\alpha_1, \dots, \alpha_d) w \right]$$

$$+ \sum_{i=1}^d \mathbb{E}_{q(w, \lambda)} \left[ \underbrace{a_0 \ln b_0 - \ln \Gamma(a_0)}_c + (a_0 - 1) \ln \alpha_i - b_0 \alpha_i \right] + c$$

$$= -\frac{1}{2} \sum_{i=1}^d \ln \alpha_i^{-1} - \frac{1}{2} \mathbb{E}_{q(w)} [w^T \text{diag}(\alpha_1, \dots, \alpha_d) w] + \sum_{i=1}^d [(a_0 - 1) \ln \alpha_i - b_0 \alpha_i] + c$$

$$= \sum_{i=1}^d (a_0 - \frac{1}{2}) \ln \alpha_i - b_0 \alpha_i - \frac{1}{2} \mathbb{E}_{q(w)} [w^T \text{diag}(\alpha_1, \dots, \alpha_d) w] + c$$

$$\Rightarrow q(\alpha_i) = (a_0 + \frac{1}{2} - 1) \ln \alpha_i - \alpha_i (b_0 + \frac{1}{2} \mathbb{E}_{q(w)} [w_i^2]) + c$$

is in the form of Gamma distribution, with (for  $i=1, \dots, d$ )

$$\begin{cases} a_n = a_0 + \frac{1}{2} \\ b_n = b_0 + \frac{1}{2} \mathbb{E}_{q(w)} [w_i^2] \end{cases}$$

Therefore, we have the distribution of  $q(w)$ ,  $q(\lambda)$  and  $q(\alpha_i)$ , the whole iteratively-updating algorithm is as following:

VI algorithm:

1. Initialize  $a'_0, b'_0, \mu'_0, \Sigma'_0, e'_0, f'_0$

$$f'_0 + \frac{1}{2} \sum_{i=1}^n \mathbb{E}_{g(w)} [(y_i - w^T x_i)^2]$$

$$= f'_0 + \frac{1}{2} \sum_{i=1}^n [(y_i - \mu'_{t-1} x_i)^2 + x_i^T \Sigma'_{t-1} x_i]$$

2. For iteration  $t = 1, \dots, T$

- Update  $g_t(\lambda)$  by setting

$$\begin{cases} e'_t = e_0 + \frac{n}{2} \\ f'_t = f_0 + \frac{1}{2} y^T y - \mathbb{E}_{g(w)} [W^T] X^T y + \frac{1}{2} \mathbb{E}_{g(w)} [W^T] X^T X \mathbb{E}_{g(w)} [W] \end{cases}$$

- Update  $g_t(w)$  by setting

$$\begin{cases} \mu'_t = \Sigma'_t \mathbb{E}_{g(\lambda)} [\lambda] X^T y \\ \Sigma'_t = (\mathbb{E}_{g(\lambda)} [\text{diag}(\alpha_1, \dots, \alpha_d)] + \mathbb{E}_{g(w)} [\lambda] X^T X)^{-1} \end{cases}$$

- Update  $g_t(\alpha_i)$  by setting

$$\begin{cases} a'_t = a_0 + \frac{1}{2} \\ b'_t = b_0 + \frac{1}{2} \mathbb{E}_{g(w)} [w_i^2] \end{cases}$$

(for each  $i = 1 \sim d$ )

With  $\mathbb{E}_{g(\lambda)} [\text{diag}(\alpha_1, \dots, \alpha_d)]$  is a diagonal matrix  $A$  with elements

$$A_{ii} = \mathbb{E}_{g(\alpha_i)} [\alpha_i] = \frac{a'_t}{b'_t}$$

$$\mathbb{E}_{g(w)} [\lambda] = \frac{e'_t}{f'_t}, \quad \mathbb{E}_{g(w)} [w_i^2] = (\mu'_{ti})^2 + (\Sigma'_t)_{ii}$$

Joint likelihood:  $p(y, w, \alpha, \lambda | x) = p(y | w, \lambda, x) p(w | \alpha) p(\alpha) p(\lambda)$

Variational object function:

$$\mathcal{L}(a'_t, b'_t, \mu'_t, \Sigma'_t, e'_t, f'_t) = \sum_{i=1}^n \mathbb{E}[\ln p(y_i | x_i, w, \lambda)] + \mathbb{E}[\ln p(w | \alpha)] + \sum_{k=1}^d \mathbb{E}[\ln p(\alpha_k)] + \mathbb{E}[\ln p(\lambda)]$$

$$- \mathbb{E}[\ln q(w)] - \sum_{k=1}^d \mathbb{E}[\ln q(\alpha_k)] - \mathbb{E}_g[\ln q(\lambda)]$$

(1)  $\sum_{i=1}^n \mathbb{E}[\ln p(y_i | x_i, w, \lambda)] = \frac{1}{2} \mathbb{E}_{g(w)} [\ln \lambda] - \frac{1}{2} \ln(2\pi) - \frac{1}{2} \mathbb{E}[\lambda (y_i - x_i^T w)^2]$

① if  $g(\lambda) \sim \text{Gamma}(e, f)$ ,  $\mathbb{E}_{g(w)} [\ln \lambda] = \psi(e) - \ln f$

②  $\mathbb{E}[\lambda (y_i - x_i^T w)^2] = \mathbb{E}[\lambda] \mathbb{E}[(y_i - x_i^T w)^2] = \frac{e}{f} \cdot \mathbb{E}[(y_i - x_i^T w)^2]$

(2)  $\mathbb{E}_g[\ln p(w | \alpha)] = \sum_{k=1}^d \mathbb{E}[\ln p(w_k | \alpha_k)]$

$$\mathbb{E}[\ln p(w_k | \alpha_k)] = \frac{1}{2} \mathbb{E}[\ln \alpha_k] - \frac{1}{2} \ln(2\pi) - \frac{1}{2} \mathbb{E}[\alpha_k \cdot w_k^2]$$

$$(3) \mathbb{E}[\ln p(\alpha_k)] = a_0 \ln b_0 - \ln \Gamma(a_0) + (a_0 - 1) \mathbb{E}[\ln \alpha_k] - b_0 \mathbb{E}[\alpha_k]$$

$$(4) \text{ Last three terms: } \mathbb{E}[\ln g(w)], \sum_{k=1}^d \mathbb{E}[\ln g(\alpha_k)], \mathbb{E}[\ln g(\lambda)] \Rightarrow \text{Entropy of each distribution}$$

$$\mathbb{E}[\ln g(\lambda)] = \ln \Gamma(a) - (a-1) \psi(a) - \ln b + a$$

$$\mathbb{E}[\ln g(w)] = -\left(\frac{1}{2} \ln |\Sigma| + \frac{d}{2} (1 + \ln(2\pi))\right)$$

(by appendix from PRML)

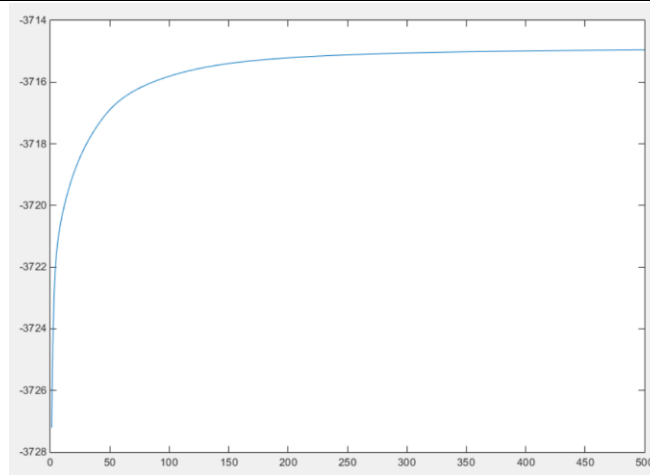
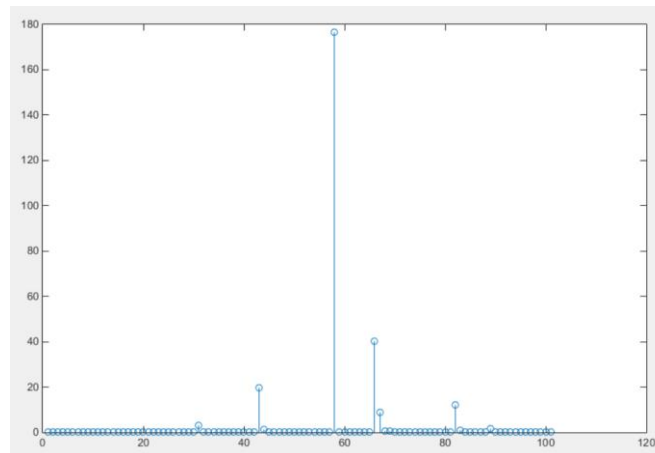
Combine them all together, we have

$$\begin{aligned} \mathcal{L}(a_t', b_t', \mu_t', \Sigma_t', e_t', f_t') = & \left[ \frac{n}{2} (\psi(e_t') - \ln f_t') - \frac{n}{2} \ln(2\pi) - \frac{1}{2} \cdot \frac{e_t'}{f_t'} \sum_{i=1}^n ((y_i - \mu_t'^T x_i)^2 + x_i^T \Sigma_t' x_i) \right] \\ & + \sum_{k=1}^d \left[ \frac{1}{2} (\psi(a_{tk}') - \ln(b_{tk}')) - \frac{1}{2} \ln(2\pi) - \frac{1}{2} \left( \frac{a_{tk}'}{b_{tk}'} \right) \cdot (\mu_{tk}'^2 + (\Sigma_t')_{kk}) \right] \\ & + \sum_{k=1}^d \left[ a_0 \ln b_0 - \ln \Gamma(a_0) + (a_0 - 1) (\psi(a_{tk}') - \ln(b_{tk}')) - b_0 \frac{a_{tk}'}{b_{tk}'} \right] \\ & + (e_0 \ln f_0 - \ln \Gamma(e_0) + (e_0 - 1) (\psi(e_t') - \ln(f_t')) - f_0 \frac{e_t'}{f_t'}) \\ & + \left( \frac{1}{2} \ln |\Sigma_t'| + \frac{d}{2} (1 + \ln(2\pi)) \right) - \sum_{k=1}^d \left[ a_{tk}' \ln b_{tk}' - \ln \Gamma(a_{tk}') + (a_{tk}' - 1) (\psi(a_{tk}') - \ln(b_{tk}')) - a_{tk}' \right] \\ & - \left[ e_t' \ln f_t' - \ln \Gamma(e_t') + (e_t' - 1) (\psi(e_t') - \ln(f_t')) - e_t' \right] \end{aligned}$$

## Problem 2.

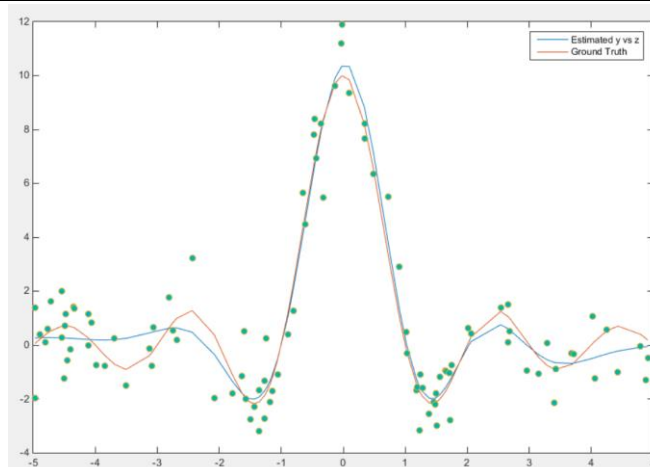
data1.mat:

## a) Plot of variational objective function

b) Plot of  $1/E_q[\alpha_k]$ c)  $1/E_q[\lambda]$ 

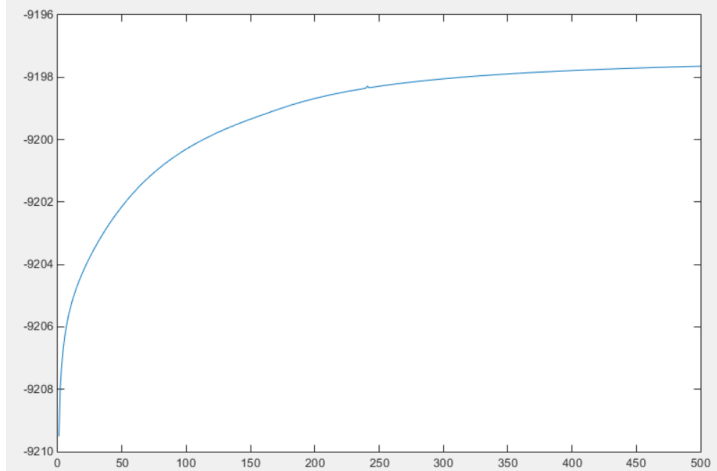
1.0798

## d) Estimated y vs. z, Ground Truth and Scatter Plot

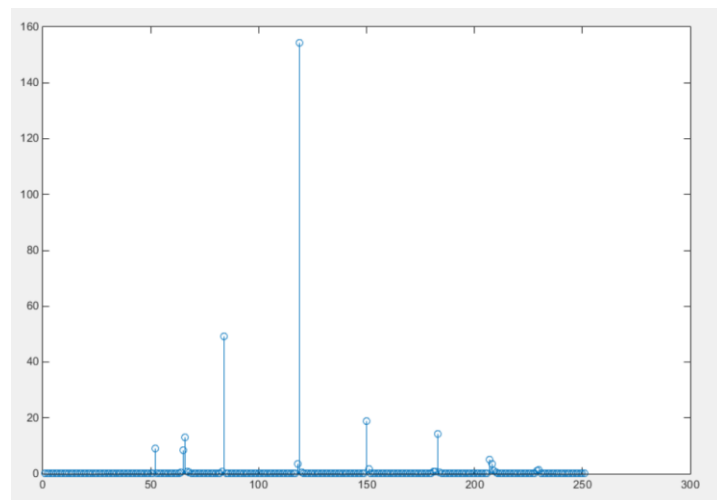


data2.mat

a) Plot of variational objective function



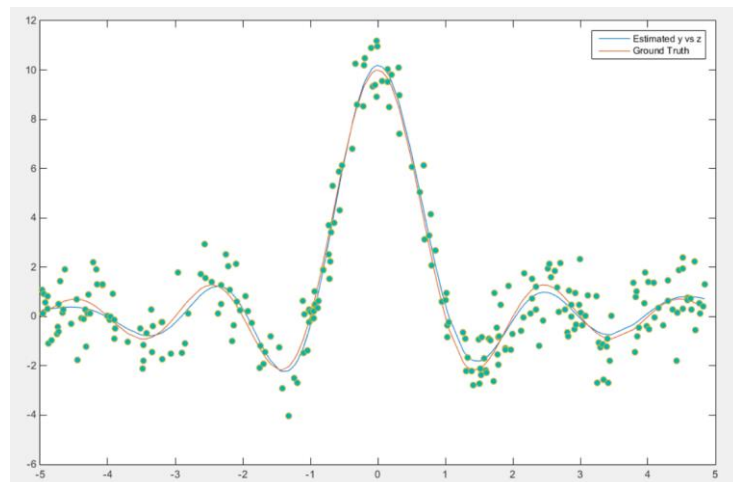
b) Plot of  $1/E_q[\alpha_k]$



c)  $1/E_q[\lambda]$

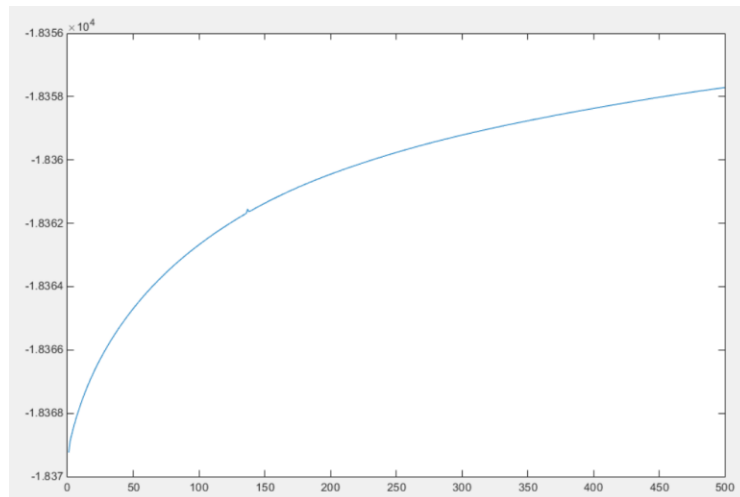
0.8994

d) Estimated y vs. z, Ground Truth and Scatter Plot

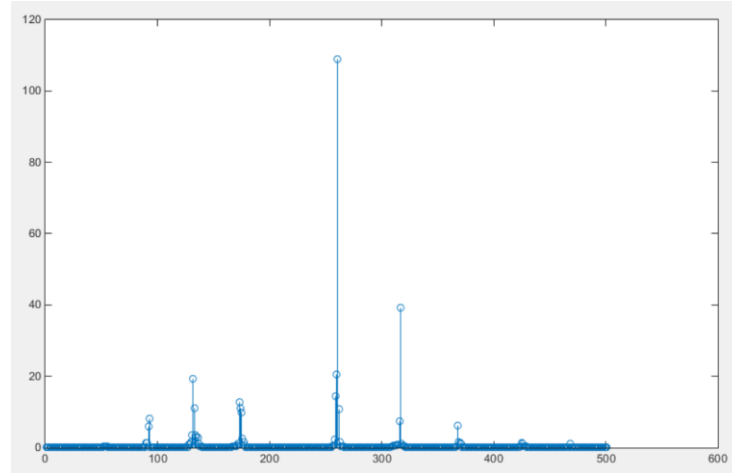


data3.mat

a) Plot of variational objective function



b) Plot of  $1/E_q[\alpha_k]$



c)  $1/E_q[\lambda]$

0.9781

d) Estimated y vs. z, Ground Truth and Scatter Plot

