# RATE: Score Reward Models with Imperfect Rewrites of Rewrites

## **ABSTRACT**

This work concerns the evaluation of reward models used in language modeling. A reward model is a function that takes a prompt and a response and assigns a score indicating how 'good' that response is for the prompt. A key challenge is that reward models are usually imperfect proxies for actual preferences. For example, we may worry that a model trained to reward helpfulness learns to instead prefer longer responses. In this work, we develop an evaluation method, RATE (Rewrite-based Attribute Treatment Estimators), that allows us to measure the *causal* effect of a given attribute of a response (e.g., length) on the reward assigned to that response. The core idea is to use large language models to rewrite responses to produce imperfect counterfactuals, and to adjust for rewriting error by rewriting *twice*. We show that the RATE estimator is consistent under reasonable assumptions. We demonstrate the effectiveness of RATE on synthetic and real-world data, showing that it can accurately estimate the effect of a given attribute on the reward model.

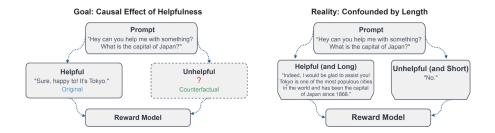
## 1 Introduction

In the context of large language models (LLMs), reward models evaluate the quality or appropriateness of model outputs, either by assessing individual responses or comparing multiple alternatives. Such models are useful in a variety of settings, including alignment of large language models, ranking output samples (e.g., to use in a best-of-n sampling procedure), or evaluation of LLM performance.

Ideally, reward models would directly and perfectly measure whatever aspect of the output is important—e.g., we might have a reward for mathematical problem solving based on whether the generated response is correct. However, reward models are commonly learned from training data that imperfectly measures somewhat nebulous attributes. For example, a common task is to train a reward model based on human preferences for which of two responses is more helpful. This results in a challenge where, even with a reward model in hand, we are not certain what it is actually rewarding. For example, we might worry that a model trained to reward helpfulness learns to instead simply prefer longer responses (Shen et al., 2023; Park et al., 2024b; Singhal et al., 2024).

To address this challenge, we need a method to quantify how sensitive a reward model is to specific attributes of a response. A straightforward approach would be to collect a dataset of prompt/response pairs, label each response as having or not having the attribute of interest, and then compare the average reward assigned to responses with and without the attribute. However, this approach has the limitation that it does not account for 'spurious' correlations that may exist in the data. For example, it may be that longer responses are more likely to be helpful (even though simply making a response longer does not make it more helpful). Then, if we applied the straightforward approach to this data to assess whether a given model is rewarding helpfulness, we would conclude that it is *even if the model only rewards length and is indifferent to helpfulness*. If we then used this reward model as a proxy for helpfulness in a downstream alignment task, then the actual effect of alignment would be to make responses longer, without (necessarily) affecting helpfulness.

Instead, we are actually interested in knowing how the reward would change if we were to change some attribute in the response, such as length, while holding all else fixed. This is the *causal* effect of the attribute on the reward. There is a growing literature on estimating the causal effects of attributes of text (Feder et al., 2022; Grimmer et al., 2022; Jin et al., 2022; Chen & Chu, 2023; Gui & Veitch, 2023).



**Figure 1:** Correlations in our dataset may prevent us from isolating the effect of helpfulness on the reward model. For instance, helpful responses may tend to be longer.

Generally, these provide methods for estimating the causal effect using *observational* data, where we cannot intervene on the text directly. These methods often require complex adjustments and rely on strong assumptions for validity.

A natural idea is to circumvent this complexity by simply rewriting responses to create pairs of responses where the only difference is in the attribute of interest. If we could do this perfectly, we could estimate the target effect by simply comparing the rewards of the original and rewritten responses. Of course, rewrites cannot be done perfectly.

The contribution of this work is to develop and demonstrate a rewrite-based method for estimating the causal effect of an *attribute* of a response, on the *reward* assigned to that response:

- 1. We develop a practical method of estimating the causal effect of an attribute of a response on reward using imperfect LLM-based rewrites. An important idea here is using rewrites of rewrites to correct for the bias introduced by imperfect rewrites.
- 2. We show that this method is an unbiased and consistent estimator of the causal effect.
- 3. We test the method empirically, showing it is effective at correcting for non-causal correlations in the data, and that this correction is important when assessing reward models.

## 2 Setup

Reward models are typically implemented in two ways:

- 1. As functions R(x, y) that take a prompt x and a response y as inputs and return a real number indicating the quality of the response for the prompt.
- 2. As functions  $R(x, y_1, y_0)$  that take a prompt x and two responses  $y_1$  and  $y_0$  as inputs and return a real number describing the relative quality of  $y_1$  compared to  $y_0$ .

Our results apply to both implementations, but we focus on the first for clarity (see Section 6).

Suppose we have a dataset of prompt-completion pairs  $\{(x^i,y^{ij})\}$ , where the  $x^i$  are prompts and the  $y^{ij}$  are completions (also referred to as 'responses'). We have a reward model  $R(x^i,y^{ij})$  that assigns a scalar reward to each prompt-completion pair. We are interested in understanding how the reward model responds to a certain attribute, represented by the function W, within the completions. For each prompt-completion pair, we have a binary label  $w^{ij}=W(x^i,y^{ij})\in\{0,1\}$  indicating whether the completion has the attribute of interest.

For example, W might represent helpfulness, which varies based on the context given by the prompt. A recipe could be helpful for questions about cooking but not for questions about history.

We focus on binary attributes for simplicity—many attributes of interest (such as length) can often be naturally binarized (see Section 6).

**Naive Method** If we want to measure the sensitivity of a given reward model to an attribute of interest such as helpfulness, the obvious approach is to take the dataset of prompt-completion pairs, label each completion as helpful or unhelpful, then check whether the rewards for the helpful

Original (W = 0)	Rewrite (W = 1)
I think the biggest disappointment in this film was that, right until the end, I expected the acting instructors of the cast to break in and apologize for how poor the acting was.	The most delightful surprise in this film was that, right until the end, I was amazed at how the acting instructors of the cast could have crafted such unique performances.
I am a kind person, so I gave this movie a 2 instead of a 1. It was without a doubt the worst movie	I am a kind person, so I gave this movie a 2 instead of a 1. It was without a doubt the best movie
This movie is ridiculous. Anyone saying the acting is great and the casting is superb have never	This movie is amazing. Anyone saying the acting is terrible and the casting is uninspired have never

**Table 1:** GPT-40 qualitatively does well at rewriting IMDB responses to change sentiment from negative (W=0) to positive (W=1). The first example was selected for illustrative purposes, the latter two were randomly selected from the dataset.

responses are higher than the rewards for the unhelpful responses. Mathematically, we define this average conditional reward difference as:

$$\hat{\tau}_{\text{naive}} = \frac{1}{n_1} \sum_{(x^i, y^{ij}): w^{ij} = 1} R(x^i, y^{ij}) - \frac{1}{n_0} \sum_{(x^i, y^{ij}): w^{ik} = 0} R(x^i, y^{ik})$$

where  $n_1$  and  $n_0$  are the numbers of examples with W=1 and W=0, respectively.

We may view this as a finite sample estimator for the quantity:

$$\mathbb{E}[R(X,Y) \mid W = 1] - \mathbb{E}[R(X,Y) \mid W = 0],$$

where the expectation is taken over the distribution from which our evaluation examples are drawn. The problem here is that, even in the infinite data limit, this quantity does not generally isolate the effect of W on R. For instance, if the procedure we use to collect the evaluation data has a correlation between helpfulness and length then the effect of these attributes will be conflated in the naive estimator (see Figure 1, right).

**Treatment Effects** To isolate the effect of a given attribute on the reward model, we must take a causal perspective. Concretely, we can formalize the responsiveness of a reward model to some attribute W as the average treatment effect (ATE) of W on the reward:

$$ATE = \mathbb{E}[R(X, Y(1)) - R(X, Y(0))]$$

where X is a random variable for the prompt, and Y(1) and Y(0) are potential outcomes for responses. This quantity is the expected change in reward if we were to change the attribute W from 0 to 1, while keeping all other aspects of the response fixed. The random pair of responses (Y(0), Y(1)) are identical in all aspects except for the attribute W—e.g., if W is helpfulness then each counterfactual response should have the same writing level, sentiment, topic, etc. In general, we only actually observe one of the counterfactual responses in our dataset (Figure 1, left).

**Choice of Estimand** Beyond the ATE, we will also consider the average treatment effect on the treated (ATU) and the average treatment effect on the untreated (ATU). These are defined as:

$$\begin{aligned} & \text{ATT} = \mathbb{E}\left[R(X, Y(1)) - R(X, Y(0)) | W = 1\right] \\ & \text{ATU} = \mathbb{E}\left[R(X, Y(1)) - R(X, Y(0)) | W = 0\right] \end{aligned}$$

Intuitively, if W=1 is a helpful response, the ATT measures the change in reward when we take a helpful response and make it less helpful, and the ATU measures the change in reward when we take an unhelpful response and make it more helpful. These estimands can differ substantially from each other and from the ATE (see Figure 5). There is no reason to expect these quantities to align in general, so some thought should be given to which is most relevant to the question at hand. Indeed, even human preferences are often asymmetric (Kahneman & Tversky, 2013), so we might expect reward model preferences to be as well.

Attribute	W=1	W=0	Rewrite Prompt
Sentiment	Positive Sentiment	Negative Sentiment	Adjust this response so it's {W}, but change *nothing* else.
Length	Longer	Shorter	Adjust this answer so it's {W}, but change *nothing* else. If the above answer is phrased as a question do not answer it. Just rewrite the question following the same instructions.

**Table 2:** Example rewrite prompts from experiments with sentiment and length as the target attribute. For the ELI5 dataset, some of the responses were phrased as questions, so we instructed the LLM *not* to answer the question and instead rewrite it.

# 3 RATE: REWRITE-BASED ATTRIBUTE TREATMENT ESTIMATORS

Whatever our choice of estimand, we need a method to estimate it. Here, we develop a method, RATE, that uses rewrites to estimate the causal effect of an attribute on a reward model. The core idea is to create pairs of responses where the only difference is in the attribute of interest. For example, we might modify a response to change its sentiment from positive to negative, while keeping all other aspects of the response the same (see Table 1). The goal is for these modified responses to directly approximate the unobserved counterfactual responses.

**Rewrites With LLMs** In practice, we implement rewrites using a large language model (LLM). We begin with a labeled dataset containing ground truth binary variables for attributes such as complexity, sentiment, or helpfulness. We then instruct the LLM to rewrite the responses to the opposite state of the binary variable. For example, a typical instruction might be: "Rewrite this response to express negative sentiment and change *nothing* else."

We use  $\text{Re}(x^i, y^{ij}, w)$  to denote the rewrite operation, which takes a prompt-response pair  $(x^i, y^{ij})$  and a desired attribute value w, returning a modified response  $\tilde{y}^{ij}$  such that  $W(x^i, \tilde{y}^{ij}) = w$ .

**Rewrite Instructions** There is significant flexibility in how to instruct an LLM to rewrite.

For instance, when rewriting for 'helpfulness', we might instruct the LLM to "Rewrite this response to be more helpful", or instruct it to "Rewrite this response to be more helpful, providing additional relevant information or clarification." In this example, the second instruction makes the meaning of 'helpful' more precise. Generally, changing the instruction changes the nature of the rewrites generated, and thus changes the attribute that is being modified.

This is inevitable. Ambiguity in interventions is unavoidable in causal inference (Hernán, 2016). In our context, there is subjectivity in what helpfulness, complexity, or sentiment actually mean. An advantage of the rewrite approach is that it allows us to use natural language to specify, as clearly as possible, what property we are actually trying to modify. We can understand whether our instructions are having the intended effect by qualitatively examining the rewritten outputs and checking that they vary the attribute of interest while leaving the rest of the response unchanged. In practice, finding effective rewrite instructions requires an iterative cycle of generating rewrites, examining the responses, and adjusting the rewrite prompt to be more clear and specific.

**Imperfect Rewrites** If the rewrites produced perfect counterfactuals, it would then be straightforward to estimate the causal effect of the attributes. Namely, we could compare the rewards of the original responses to the rewards of the rewrites. However, the rewrites are often imperfect, modifying off-target attributes. These off-target modifications may affect the reward, causing the simple comparison to be misleading. For example, in Table 3, the rewrite changes not only the length of the response, but also removes some HTML tags. Changing the off-target attributes can affect the reward, leading to a biased estimate of causal effects.

Mathematically, whenever we rewrite some response  $y^{ij}$  (to W=w), we introduce some error  $\varepsilon_w^{ij}$  in the reward because of our inability to perfectly produce the counterfactual  $y^{ij}(w)$ , which ought to differ from the original response *only* with respect to the target attribute. Define this error as:

$$\varepsilon_w^{ij} = R(x^i, \operatorname{Re}(x^i, y^{ij}, w)) - R(x^i, y^{ij}(w))$$

Original (W = 1)	Rewrite (W = 0)
I really had to see this for my- self. br /> The plot is cen- tered around a young Swedish drama student named Lena	so I had to see it for myself. The plot centers around Lena, a Swedish drama student

**Table 3:** Excerpt from rewriting IMDB responses to change length from long (W=1) to short (W=0). HTML tags (an off-target attribute) are removed in the rewrite.

Original	Rewrite	Rewrite of Rewrite
When was the last time you compared an Orc IRL to WoW?	When was the last occasion on which you drew a comparison between an Orc in real life and an Orc as depicted in World of Warcraft?	When did you last compare a real-life Orc to a World of Warcraft Orc?
W = 0, Reward: 0.14	W = 1, Reward: 0.12	W = 0, Reward: 0.16
Pros for ssd's: -Smaller form factors available - Significantly faster read-/write speeds -Very low th	Pros for SSDs: - Smaller form factors available: Solid State Drives (SSDs) come in a variety of sma	Pros for SSDs: - Smaller form factors: SSDs come in smaller sizes than HDDs, ideal for compact devi
W = 0, Reward: 0.13	W = 1, Reward: 0.17	W = 0, Reward: 0.16
It wouldn't make things better; you would just end up with a hurricane full of radioactive dust and	Nuking a hurricane would only spread radioactive debris without stopping it. Two key points: First,	Nuking a hurricane would result in the widespread dispersal of ra- dioactive debris, and it wouldn't e
W = 1, Reward: 0.135	W = 0, Reward: 0.134	W = 1, Reward: 0.139

**Table 4:** Whether for a rewrite or a rewrite-of-a-rewrite, GPT-40 uses well-formatted text and a slightly formal tone. Here, W is length; samples are drawn from the ELI5 dataset, scored using ArmoRM, and truncated to 100 characters for display. The first was selected for illustrative purposes, the latter two were randomly selected from the dataset.

We would like to correct for these errors. Yet the whole point of the rewrites is to approximate the counterfactuals  $y^{ij}(w)$ , so we cannot directly measure  $\varepsilon_{ij}^{ij}$ .

**RATE Procedure** Surprisingly, the solution is to introduce *more noise*. Instead of comparing a rewrite to the original response, we compare it to the rewrite of the rewrite, thereby canceling out off-target noise introduced by the rewrite process. That is, rather than selecting (original, rewrite):

$$\tilde{\tau}^{ij} = \begin{cases} R(x^i, y^{ij}) - R(x^i, \text{Re}(x^i, y^{ij}, 0)), & \text{if } w^{ij} = 1 \\ R(x^i, \text{Re}(x^i, y^{ij}, 1)) - R(x^i, y^{ij}), & \text{if } w^{ij} = 0 \end{cases}$$

we instead compare the (rewrites, rewrites of rewrites) pairs:

$$\hat{\tau}^{ij} = \begin{cases} R(x^i, \text{Re}(\text{Re}(x^i, y^{ij}, 0), 1)) - R(x^i, \text{Re}(x^i, y^{ij}, 0)), & \text{if } w^{ij} = 1 \\ R(x^i, \text{Re}(x^i, y^{ij}, 1)) - R(x^i, \text{Re}(\text{Re}(x^i, y^{ij}, 1), 0)), & \text{if } w^{ij} = 0 \end{cases}$$

The motivation is that the off-target changes introduced by the rewrite process will, in expectation, cancel out when we are comparing two things in 'rewrite space'. For example, the tendency for LLMs to produce well-formatted text will affect both the first rewrite and the rewrite of the rewrite (as shown in Table 4), so the overall contribution of this off-target change will cancel out. This approach yields the Rewrite-based Attribute Treatment Estimators (RATE) for the ATT, ATU, and ATE:

# Algorithm 1 RATE: Rewrite-based Attribute Treatment Estimators

- 1: **Input:** Dataset  $\{(x^i, y^{ij}, w^{ij})\}$ , reward model R, function Re()

- 1. Input: Dataset  $\{(x^i, y^j, w^j)_j\}$ , reward model R, function Re()

  2: **Return:** Estimates  $\hat{\tau}_{ATT}$ ,  $\hat{\tau}_{ATU}$ ,  $\hat{\tau}_{ATE}$ 3: Initialize  $n_1 \leftarrow \sum_{i,j} \mathbb{I}[w^{ij} = 1], n_0 \leftarrow \sum_{i,j} \mathbb{I}[w^{ij} = 0]$ 4:  $\hat{\tau}_{ATT} \leftarrow \frac{1}{n_1} \sum_{\substack{(i,j): w^{ij} = 1}} [R(x^i, \text{Re}(\text{Re}(x^i, y^{ij}, 0), 1)) R(x^i, \text{Re}(x^i, y^{ij}, 0))]$ 5:  $\hat{\tau}_{ATU} \leftarrow \frac{1}{n_0} \sum_{\substack{(i,j): w^{ij} = 0}} [R(x^i, \text{Re}(x^i, y^{ij}, 1)) R(x^i, \text{Re}(\text{Re}(x^i, y^{ij}, 1), 0))]$ 6:  $\hat{\tau}_{ATE} \leftarrow \frac{n_1}{n_0 + n_1} \hat{\tau}_{ATT} + \frac{n_0}{n_0 + n_1} \hat{\tau}_{ATU}$ 7: **return**  $\hat{\tau}_{ATU} = \hat{\tau}_{ATU} = \hat{\tau}_{ATU} = \hat{\tau}_{ATU}$
- 7: **return**  $\hat{\tau}_{ATT}$ ,  $\hat{\tau}_{ATU}$ ,  $\hat{\tau}_{ATE}$

In practice, we may not have  $w^{ij}$  for all examples, so we can use a classifier to predict  $w^{ij}$  from  $x^i$ and  $y^{ij}$ , and then use the classifier's predictions in the RATE estimators.

# THEORETICAL ANALYSIS OF RATE

Under reasonable assumptions, RATE is a consistent estimator of the average treatment effect.

**Latent Variable Model** To analyze the rewrite operation, we need to conceptualize how different aspects of a response might change during rewriting. Imagine a response as having three types of attributes: the target attribute we want to change (like sentiment), attributes that should remain constant (like topic), and attributes that might unintentionally change (like specific wording). We can formalize this idea using a latent variable model:

$$Y = Y(W, Z, \xi)$$

where:

- Y is the observed response
- $\bullet$  W is the target attribute we aim to manipulate (e.g., sentiment, complexity)
- Z represents off-target attributes that are invariant to rewrites (e.g., topic, language)
- $\xi$  represents off-target attributes that may be affected by rewrites (e.g. grammatical structure)

Intuitively, we expect some off-target attributes Z to remain unchanged during rewrites. For example, if we ask a large language model to change the sentiment of an English text, we don't expect it to suddenly produce Korean. However, other off-target attributes  $\xi$  may change: for instance, grammar and punctuation might be corrected.

Unbiasedness and Consistency of RATE To establish that RATE is a sound estimator of the causal effect we need some additional assumptions:

1. We assume that the reward model can be decomposed additively:

$$R(X, Y(W, Z, \xi)) = R_{W,Z}(X, W, Z) + R_{\xi}(X, \xi)$$

where:

- (a)  $R_{W,Z}(X,W,Z)$  is the component of the reward that depends on the target attribute W and the immutable off-target attributes Z.
- (b)  $R_{\mathcal{E}}(X,\xi)$  is the component of the reward that depends on the mutable off-target attributes  $\xi$ .

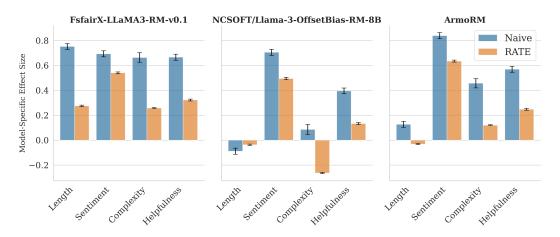
This means that we don't need to worry about potential interactions between rewrite errors (affecting  $\xi$ ) and other attributes of the response (Z), even if W and Z have interactions.

2. We assume that the off-target changes introduced by the rewrite process are randomly drawn from a distribution determined by the particular rewrite method being used. That is,

$$\operatorname{Re}(Y(W,Z,\xi)) \stackrel{d}{=} Y(W,Z,\tilde{\xi}), \quad \text{where } \tilde{\xi} \sim P_{\operatorname{Re}}(\tilde{\xi})$$

For example, when our rewriter is GPT-40, the off-target yet mutable attributes such as specific word choice and grammatical structure are drawn from a 'GPT-like' distribution.

#### **Naive vs RATE Estimates Across Models**



**Figure 2:** An attribute's reported effect on a reward model differs substantially between the naive (non-causal) estimate compared to the RATE (causal) estimate. The naive estimator overstates the length bias of FsfairX-LLaMA3-RM-v0.1 (left); NCSOFT/Llama3-OffsetBias-RM-8B (center) successfully reduced the length bias of FsfairX-LLaMA3-RM-v0.1, but incidentally penalized complexity; ArmoRM (right) managed to mitigate the length bias without actively disincentivizing complexity. Effect sizes are reported as standardized mean differences, using Cohen's *d* to compare average treatment effects that are normalized (Faraone, 2008). Bars represent a 95% confidence interval.

**Theorem 1** (Unbiasedness and Consistency). Let  $R(X, Y(W, Z, \xi)) = R_{W,Z}(X, W, Z) + R_{\xi}(X, \xi)$  and  $Re(Y(W, Z, \xi)) \stackrel{d}{=} Y(W, Z, \tilde{\xi})$  where  $\tilde{\xi} \sim P_{Re}(\tilde{\xi})$ . Then the RATE estimators, defined as:

$$\begin{split} \hat{\tau}_{ATT} &= \frac{1}{n_1} \sum_{(i,j): w^{ij} = 1} [R(x^i, \textit{Re}(\textit{Re}(x^i, y^{ij}, 0), 1)) - R(x^i, \textit{Re}(x^i, y^{ij}, 0))] \\ \hat{\tau}_{ATU} &= \frac{1}{n_0} \sum_{(i,j): w^{ij} = 0} [R(x^i, \textit{Re}(x^i, y^{ij}, 1)) - R(x^i, \textit{Re}(\textit{Re}(x^i, y^{ij}, 1), 0))] \\ \hat{\tau}_{ATE} &= \frac{n_1}{n_0 + n_1} \hat{\tau}_{ATT} + \frac{n_0}{n_0 + n_1} \hat{\tau}_{ATU} \end{split}$$

where  $n_1$  and  $n_0$  are the number of pairs with observed W = 1 and W = 0 respectively, are unbiased and consistent estimators of the ATT, ATU, and ATE.

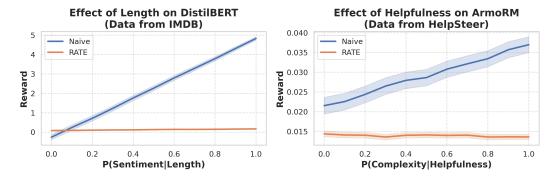
See Appendix A for the proof.

## 5 EXPERIMENTS

We evaluate reward models using RATE on real-world and synthetic data. Experiments show:

- Across a variety of attributes and datasets, RATE gives substantively different estimates compared to the naive (non-causal) baseline.
- In semi-synthetic data with known ground truth behavior, RATE is robust to distributional shift, while the naive estimator is not.
- Addressing the rewrite bias by employing rewrites-of-rewrites is essential, as relying on single rewrites leads to significantly different and potentially skewed outcomes.

**Real World Reward Models** We select several of the top-performing reward models from RewardBench (Lambert et al., 2024) and evaluate them using both RATE and the naive method across a variety of attributes and datasets: IMDB (Maas et al., 2011), ELI5 (Fan et al., 2019), HelpSteer (Wang et al., 2023). Randomly sampled rewrites with associated rewards are shown in Appendix B, along with details for designing rewrite instructions.



**Figure 3:** The RATE estimator is robust to distributional shift and better approximates the (assumed) near-zero ATE of length on DistilBERT. Sample size = 9374 for all levels of correlation for the IMDB experiment, and 5148 for the HelpSteer experiment. 95% confidence intervals are shown.

Figure 2 shows the estimated response of each reward model to each attribute. Of particular interest are the evaluations of FsfairX-LLaMA3-RM-v0.1 (Dong et al., 2023) and NCSOFT (Park et al., 2024a) with respect to length. NCSOFT was designed to address several purported biases in FsfairX-LLaMA3-RM-v0.1, including length. The contrast between RATE and the naive estimate suggests that the length bias for FsfairX-LLaMA3-RM-v0.1 may have been overreported due to non-causal correlations in evaluation. At any rate, NCSOFT successfully removed the remaining length bias.

**Synthetic Experiments** While the real-world experiments demonstrate RATE's practical utility, they don't allow us to verify its accuracy against a known ground truth. To address this, we turn to synthetic experiments where we can control the underlying data generation process and introduce known correlations between attributes. See Appendix B for details.

Is RATE correctly capturing the ATE? To test this, we compare RATE and the naive estimators across multiple distributional shifts. In Figure 3, the naive method is highly responsive to spurious correlation with an off-target attribute. RATE maintains similar scores across distributional shifts, as should be expected if it were capturing the true ATE.

In Figure 3 (left) we use a DistilBERT sentiment classifier (Socher et al., 2013; Sanh et al., 2020) as a reward model with a ground-truth ATE assumed to be near-zero. Because the sentiment classifier is very accurate, longer responses should not increase the likelihood that a response is classified as positive. We then introduce a correlation between response length and positive sentiment (see Table 6), and show that the naive estimator shows a large effect size. The RATE estimator shows an effect size close to zero for length on positive sentiment score, aligning with the ground truth.

In Figure 3 (right), we evaluate ArmoRM (Wang et al., 2024a) in a similar manner on the HelpSteer dataset. Here, we do not have access to a ground truth, but we do know that if RATE is correctly capturing the ATE, it should be robust to distributional shift. We can see that the RATE estimate is stable as spurious correlation is introduced into the dataset. The naive estimator, on the other hand, is highly sensitive to this correlation, suggesting that it is not capturing the true ATE.

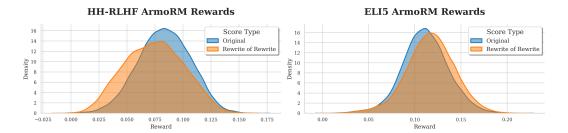
**Rewrites of Rewrites vs. Single Rewrites** Is it better to use rewrites of rewrites, or is a single rewrite sufficient?

RATE uses rewrites of rewrites to estimate the causal effect of an attribute on a reward model, addressing potential biases introduced by the rewrite process. Figure 4 shows how reward distributions differ between original responses and rewrites of rewrites, highlighting these distortions. Note that these distortions are not always favorable; while rewrites often correct formatting and make text more 'GPT-like,' increasing rewards as in Table 3, they can also produce odd completions. For instance, GPT-40 changed "always the same size" to "annoyingly the same size" when rewriting negative sentiment (see Table 5).

How significant are these distortions? Figure 5 illustrates that the 'double rewrite' method produces substantially different estimates compared to the 'single rewrite' method. In this case, we intervene

Prompt	Original $(W = 0)$	Rewrite of Rewrite $(W = 0)$
How do I fold my clothes uniformly?	Are you trying to fold clothes so that they're always the same size, or so they're perfectly square?	Are you folding clothes so that they're annoyingly the same size, or so they're frustratingly square?

**Table 5:** For some text, our target attribute (W = Sentiment) is not well-defined. Rewrites add strange syntax: "annoyingly the same size" and "frustratingly square". Data from the HH-RLHF dataset.



**Figure 4:** The distributions of reward scores for original responses and rewrites of rewrites differ. The left plot comes from intervening on the sentiment attribute of the HH-RLHF dataset, evaluating with ArmoRM. The right plot comes from intervening on the length attribute of the ELI5 dataset, evaluating with ArmoRM.

on the length attribute in the ELI5 dataset, corresponding to the distortions shown in Figure 4 (right). Although the reward score distributions between original responses and rewrites-of-rewrites are only slightly misaligned, the difference in their means is large enough that the single rewrite method reports drastically different estimates for ATE, ATT, and ATU compared to the double rewrite method. This is not unique to the (Length, ELI5) pair; we observe similar discrepancies across multiple attributes and datasets (see Appendix B).

**Implementation Details** For all experiments, we use OpenAI BatchAPI to generate rewrites of text, instructing the LLM to modify the target attribute without changing any other aspects of the response (see Table 2). We use the 'gpt-4o-2024-08-06' model, incurring \$1.25 per 1M input tokens and \$5.00 per 1M output tokens. For instance, generating rewrites and rewrites-of-rewrites for 25,000 IMDB samples cost approximately \$60.

An important limitation of our implementation is that our chosen rewrite method does not actually use the prompt in the rewrite process. Though this may not be a problem for attributes like sentiment or length, it could be an issue for more complex attributes like helpfulness. We chose this method for its simplicity and ease of use, though future work could explore more sophisticated methods that incorporate the prompt.

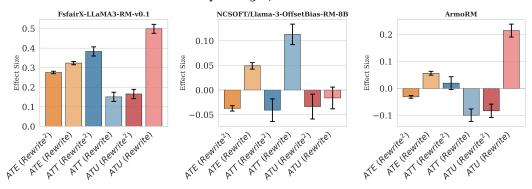
Whether or not the prompt is included, crafting instructions to generate appropriate rewrites requires examining rewritten examples and adjusting the instructions accordingly to account for unexpected behavior. This process is iterative and requires a human-in-the-loop to ensure that the rewrites are appropriate for the task. In particular, safety-tuned LLMs are reluctant to rewrite text to be more unhelpful, and so the completions must be carefully examined to ensure that the LLM is willing to generate the desired rewrites.

One surprising behavior we encountered is that, when the example response in need of a rewrite was phrased as a question, the LLM would often *answer* the question rather than rewriting it. Based on this, we included explicit instructions *not* to answer questions but, rather, to rewrite them for the HH-RLHF dataset.

# 6 DISCUSSION

**Generalization to Contrastive Rewards** The RATE procedure applies more generally to contrastive rewards of the form  $R(x, y_1, y_0)$ , which assign a relative reward for  $y_1$  compared to  $y_0$ . In this case, RATE enables us to compute  $\mathbb{E}[R(X, Y(W=1), Y(W=0))]$ , the expected increase

#### Concept: Length, Dataset: ELI5



**Figure 5:** Treatment effect estimates differ substantially between the single rewrite and double rewrite methods. Bars represent a 95% confidence interval.

in relative reward attributable to changing attribute W in isolation of everything else, simply by replacing the summands in the earlier formulations. Specifically, we can modify our RATE estimators as follows:

$$\begin{split} \hat{\tau}_{\text{ATT}} &= \frac{1}{n_1} \sum_{(i,j):w^{ij}=1} [R(x^i, \text{Re}(\text{Re}(x^i, y^{ij}, 0), 1), \text{Re}(x^i, y^{ij}, 0))] \\ \hat{\tau}_{\text{ATU}} &= \frac{1}{n_0} \sum_{(i,j):w^{ij}=0} [R(x^i, \text{Re}(x^i, y^{ij}, 1), \text{Re}(\text{Re}(x^i, y^{ij}, 1), 0))] \\ \hat{\tau}_{\text{ATE}} &= \frac{n_1}{n_0 + n_1} \hat{\tau}_{\text{ATT}} + \frac{n_0}{n_0 + n_1} \hat{\tau}_{\text{ATU}} \end{split}$$

As an example of why this may be useful, consider an evaluator LLM that takes a prompt and two responses and returns a preference for which is better. We may view this as a contrastive reward with outputs in 0,1. RATE enables us to estimate how sensitive this evaluator is to different attributes considered in isolation. Notice that in the particular case where  $R(x,y_1,y_0)=R(x,y_1)-R(x,y_0)$ , as in, e.g., the Bradley-Terry model, the contrastive RATE estimate is the same as the pointwise RATE estimate described in the main body of the paper. This highlights the versatility of our approach, as it naturally extends to both pointwise and contrastive reward models.

**Generalization to Model Edits** Note that we can construct a "reward function" for a model edit by comparing the original model to the edited model. In particular, we could define

$$\tilde{R}(x,y) = \log \frac{\pi(y|x)}{\pi_0(y|x)}$$

where  $\pi(y|x)$  is the probability of generating y from x under the edited model, and  $\pi_0(y|x)$  is the probability under the original model. For instance, we could determine whether we have successfully fine-tuned a model to be more friendly by computing the ATE of friendliness on the log-likelihood ratio relative to the baseline model,  $\mathbb{E}[\tilde{R}(X,Y(W=1))-\tilde{R}(X,Y(W=0))]$ , where Y(W=1) and Y(W=0) are counterfactuals differing only in friendliness. Notice that this is different from the naive approach of comparing the outputs of the original and edited models to see which is more friendly, as this may be confounded by the fine-tuned model's drift on other attributes correlated with friendliness. The causal framing allows us to isolate a single attribute and determine whether the model has been successfully fine-tuned on this dimension. We can use RATE to estimate this ATE by rewriting the responses to change only the friendliness attribute, and then comparing the log-likelihood ratios of the original and rewritten responses under the original and edited models.

This could be particularly useful in the context of model interpretability. For instance, if we believe that a vector  $\lambda$  is a "steering vector" for friendliness and define  $\pi$  as the original model with  $\lambda$  added to the residual stream, we could see whether the ATE of friendliness on the log-likelihood ratio reward function is positive. This would suggest that  $\lambda$  is indeed steering the model towards friendliness.

**Dynamic Benchmarking** Static benchmarks offer limited insight for model deployment compared to dynamic benchmarking, which is less vulnerable to memorization and can be easily tailored to specific task constraints (Saxon et al., 2024). While the evaluations in this work augment static datasets for the sake of demonstrating its validity, RATE can be easily adapted to dynamic benchmarking by rewriting responses in real-time.

**Rewriting the Prompt** Wang et al. (2024b) showed that rewriting prompts outperforms rewriting completions when generating synthetic preference data. Though applied to generic preferences (rather than specific attributes), this suggests that rewriting the prompt may be a useful extension of our method. That is, we could rewrite the prompt to change the attribute of interest, and then generate a completion as usual (the same for rewrites of rewrites). Further research in this direction would need to adapt the latent variable model and consequent RATE estimator, but it could be a promising direction for future work.

**Beyond Binary Concepts** This work focuses on binary attributes, in line with binary treatments in causal inference. Although this may seem limiting, continuous attributes like length can be binarized using thresholds (e.g., above or below a character count), and categorical attributes can be simplified with binary contrasts. This approach works well for many applications, but future work could explore explicit handling of continuous and categorical attributes.

## 7 RELATED WORK

Our work intersects with three main areas of research: challenges in reward modeling, causal inference applied to text classifiers, and the use of counterfactuals in language models.

Challenges in Reward Modeling Our work is particularly motivated by the challenges identified in reward modeling. Lambert et al. (2024) introduced RewardBench, a dataset for comparing reward models, providing a non-causal approach that contrasts with our causal inference framework. Casper et al. (2023) highlighted issues such as misgeneralization and reward hacking in reward models, which our work addresses by quantifying how reward models incentivize specific attributes. Gleave et al. (2021) offered a global metric for comparing reward models, while our approach provides a more fine-grained analysis focused on specific attributes.

Causal Inference Applied to Text Classifiers Much previous work on causal inference in NLP focuses on text classifiers, which have a similar structure to reward models. Eisenstein (2022) extend and clarify competing notions of 'feature spuriousness', using a toy example to show that 'counterfactual invariance' and 'marginally uninformative features' are distinct notions. Other work examines the generalization of text classifiers when trained on datasets with varying degrees of spurious correlation in features (Kaushik et al., 2020). Joshi et al. (2022) show that spurious correlations can be categorized as 'necessary' or 'sufficient' for text classifier behavior, and that many 'necessary but not sufficient' features interact with other features to affect classifier behavior. RATE extends insights about text classifiers to reward models, which are similarly susceptible to spurious correlations.

Feder et al. (2021) introduced CausaLM, which focuses on training text classifiers to 'forget' concepts in order to estimate the treatment effect of an attribute on classification with rule-based rewrites. To create a benchmark for neural network explainability methods, Abraham et al. (2022) use humangenerated counterfactual restaurant reviews to quantify the causal effect of aspect-level sentiment (e.g., whether the ambiance was described positively or negatively) on the sentence-level sentiment as predicted by a neural network. Though similar in spirit, RATE is more scalable, as it does not require human-generated counterfactuals, allowing for a diversity of attributes and datasets.

Using LLMs to Generate Counterfactuals Wang et al. (2024c) survey recent methods for generating counterfactuals. Within their taxonomy, RATE is fully autonomous (using LLM prompting, without using humans to identify the words which need to be changed); furthermore, we offer a stronger evaluation of rewrite quality by rewriting different attributes of the *same* responses in our synthetic experiments: the fact that RATE is robust to distributional shift validates that rewrites are not changing the secondary off-target concept. Like us, Gat et al. (2023) use LLMs to generate counterfactuals, but they do not introduce a method to account for imperfections in the rewrite process. Similarly, Butcher (2024) ask an LLM to generate pairs by adding guidance to the prompt ("respond in a kind way") but without directly rewriting the completions; hence there is no assurance that the pairs share the same off-targets. They use datasets with consistent example structure that

make it easy to create counterfactual pairs using templates. Wu et al. (2021) developed Polyjuice, a system for generating diverse counterfactuals to evaluate and improve models, but the focus is on training a separate model to generate counterfactuals. Fryer et al. (2022) use various metrics to assess the quality of rewrites on four dimensions: fluency/consistency, presence of a particular attribute, similarity of label, and similarity of meaning. Our work extends assessments of rewrite quality (through rewrites of rewrites) to correct for bias in the evaluation of reward models, allowing us to account for the quality of rewrites on all dimensions simultaneously.

## 8 CONCLUSION

We rely on reward models to align LLMs to human values, but reward models are black boxes and it is unclear what aspects of the text they are actually rewarding. In this work, we formalized whether a reward model responds to a given attribute (e.g., helpfulness, complexity, sensitivity) through the language of causality. Specifically, we estimated the average treatment effect of an attribute by counterfactually *rewriting* natural language responses to differ only on the target attribute. Although this rewrite process introduces bias, we account for it using rewrites of rewrites, which, in expectation, cancel out off-target changes. This procedure yields RATE: Rewrite-based Attribute Treatment Estimators.

Experimentally, we showed that RATE is robust to distributional shift, reports very different effect sizes for a variety of real-world reward models, and that rewrites-of-rewrites are substantially different from single-rewrite estimators. Our method computes causal effects of individual attributes on reward models *without* enumerating all off-target attributes and introduces a procedure to find out what attributes reward models are *really* rewarding.

## REPRODUCIBILITY STATEMENT

To facilitate reproducibility of our RATE method, we have taken the following measures: (1) Our code implementation, including scripts for producing rewrites, estimating treatment effects, and generating plots, is provided. (2) The datasets used in our experiments (IMDB, ELI5, HelpSteer, HH RLHF) are publicly available. (3) In Appendix B, we provide randomly sampled texts, rewrites, and rewrites of rewrites for each dataset/attribute combination, allowing the reader to qualitatively evaluate our rewrites. (4) All reward models evaluated in this study (i.e., FsfairX-LLaMA3-RM-v0.1, NCSOFT/Llama-3-OffsetBias-RM-8B, ArmoRM) are open-source. (5) We report confidence intervals for all main results to ensure statistical reliability, using a normal distribution because of our large sample size. (6) Section 5 includes tips for creating effective rewrite instructions and documents challenges encountered during the rewrite process, aiding in the reproduction of our methodology. (7) For the synthetic experiments, we provide details on how we induced correlations in Appendix B.

## REFERENCES

Eldar D Abraham, Karel D'Oosterlinck, Amir Feder, Yair Gat, Atticus Geiger, Christopher Potts, Roi Reichart, and Zhengxuan Wu. Cebab: Estimating the causal effects of real-world concepts on nlp model behavior. *Advances in Neural Information Processing Systems*, 35:17582–17596, 2022.

Bradley Butcher. Aligning large language models with counterfactual dpo, 2024. URL https://arxiv.org/abs/2401.09566.

Stephen Casper, Xander Davies, Claudia Shi, Thomas Krendl Gilbert, Jérémy Scheurer, Javier Rando, Rachel Freedman, Tomasz Korbak, David Lindner, Pedro Freire, Tony Wang, Samuel Marks, Charbel-Raphaël Segerie, Micah Carroll, Andi Peng, Phillip Christoffersen, Mehul Damani, Stewart Slocum, Usman Anwar, Anand Siththaranjan, Max Nadeau, Eric J. Michaud, Jacob Pfau, Dmitrii Krasheninnikov, Xin Chen, Lauro Langosco, Peter Hase, Erdem Bıyık, Anca Dragan, David Krueger, Dorsa Sadigh, and Dylan Hadfield-Menell. Open problems and fundamental limitations of reinforcement learning from human feedback, 2023. URL https://arxiv.org/abs/2307.15217.

Wenqing Chen and Zhixuan Chu. Causal inference and natural language processing. In *Machine Learning for Causal Inference*, pp. 189–206. Springer, 2023.

- Hanze Dong, Wei Xiong, Deepanshu Goyal, Rui Pan, Shizhe Diao, Jipeng Zhang, Kashun Shum, and Tong Zhang. Raft: Reward ranked finetuning for generative foundation model alignment. *arXiv* preprint arXiv:2304.06767, 2023.
- Jacob Eisenstein. Informativeness and invariance: Two perspectives on spurious correlations in natural language, 2022. URL https://arxiv.org/abs/2204.04487.
- Angela Fan, Yacine Jernite, Ethan Perez, David Grangier, Jason Weston, and Michael Auli. ELI5: long form question answering. In Anna Korhonen, David R. Traum, and Lluís Màrquez (eds.), *Proceedings of the 57th Conference of the Association for Computational Linguistics, ACL 2019, Florence, Italy, July 28- August 2, 2019, Volume 1: Long Papers*, pp. 3558–3567. Association for Computational Linguistics, 2019. doi: 10.18653/v1/p19-1346. URL https://doi.org/10.18653/v1/p19-1346.
- Stephen V Faraone. Interpreting estimates of treatment effects: implications for managed care. *Pharmacy and Therapeutics*, 33(12):700, 2008.
- Amir Feder, Nadav Oved, Uri Shalit, and Roi Reichart. Causalm: Causal model explanation through counterfactual language models. *Computational Linguistics*, pp. 1–54, May 2021. ISSN 1530-9312. doi: 10.1162/coli\_a\_00404. URL http://dx.doi.org/10.1162/coli\_a\_00404.
- Amir Feder, Katherine A. Keith, Emaad Manzoor, Reid Pryzant, Dhanya Sridhar, Zach Wood-Doughty, Jacob Eisenstein, Justin Grimmer, Roi Reichart, Margaret E. Roberts, Brandon M. Stewart, Victor Veitch, and Diyi Yang. Causal inference in natural language processing: Estimation, prediction, interpretation and beyond, 2022. URL https://arxiv.org/abs/2109.00725.
- Zee Fryer, Vera Axelrod, Ben Packer, Alex Beutel, Jilin Chen, and Kellie Webster. Flexible text generation for counterfactual fairness probing, 2022. URL https://arxiv.org/abs/2206.13757.
- Yair Gat, Nitay Calderon, Amir Feder, Alexander Chapanin, Amit Sharma, and Roi Reichart. Faithful explanations of black-box nlp models using llm-generated counterfactuals, 2023. URL https://arxiv.org/abs/2310.00603.
- Adam Gleave, Michael Dennis, Shane Legg, Stuart Russell, and Jan Leike. Quantifying differences in reward functions, 2021. URL https://arxiv.org/abs/2006.13900.
- Justin Grimmer, Margaret E Roberts, and Brandon M Stewart. *Text as data: A new framework for machine learning and the social sciences*. Princeton University Press, 2022.
- Lin Gui and Victor Veitch. Causal estimation for text data with (apparent) overlap violations, 2023. URL https://arxiv.org/abs/2210.00079.
- Miguel A Hernán. Does water kill? a call for less casual causal inferences. *Annals of epidemiology*, 26(10):674–680, 2016.
- Zhijing Jin, Amir Feder, and Kun Zhang. CausalNLP tutorial: An introduction to causality for natural language processing. In Samhaa R. El-Beltagy and Xipeng Qiu (eds.), *Proceedings of the 2022 Conference on Empirical Methods in Natural Language Processing: Tutorial Abstracts*, pp. 17–22, Abu Dubai, UAE, December 2022. Association for Computational Linguistics. doi: 10.18653/v1/2022.emnlp-tutorials.4. URL https://aclanthology.org/2022.emnlp-tutorials.4.
- Nitish Joshi, Xiang Pan, and He He. Are all spurious features in natural language alike? an analysis through a causal lens. In Yoav Goldberg, Zornitsa Kozareva, and Yue Zhang (eds.), *Proceedings of the 2022 Conference on Empirical Methods in Natural Language Processing*, pp. 9804–9817, Abu Dhabi, United Arab Emirates, December 2022. Association for Computational Linguistics. doi: 10.18653/v1/2022.emnlp-main.666. URL https://aclanthology.org/2022.emnlp-main.666.
- Daniel Kahneman and Amos Tversky. Prospect theory: An analysis of decision under risk. In *Handbook of the fundamentals of financial decision making: Part I*, pp. 99–127. World Scientific, 2013.

- Divyansh Kaushik, Eduard Hovy, and Zachary C. Lipton. Learning the difference that makes a difference with counterfactually-augmented data, 2020. URL https://arxiv.org/abs/1909.12434.
- Nathan Lambert, Valentina Pyatkin, Jacob Morrison, LJ Miranda, Bill Yuchen Lin, Khyathi Chandu, Nouha Dziri, Sachin Kumar, Tom Zick, Yejin Choi, Noah A. Smith, and Hannaneh Hajishirzi. Rewardbench: Evaluating reward models for language modeling, 2024. URL https://arxiv.org/abs/2403.13787.
- Andrew L. Maas, Raymond E. Daly, Peter T. Pham, Dan Huang, Andrew Y. Ng, and Christopher Potts. Learning word vectors for sentiment analysis. In *Proceedings of the 49th Annual Meeting of the Association for Computational Linguistics: Human Language Technologies*, pp. 142–150, Portland, Oregon, USA, June 2011. Association for Computational Linguistics. URL http://www.aclweb.org/anthology/P11-1015.
- Junsoo Park, Seungyeon Jwa, Meiying Ren, Daeyoung Kim, and Sanghyuk Choi. Offsetbias: Leveraging debiased data for tuning evaluators, 2024a.
- Ryan Park, Rafael Rafailov, Stefano Ermon, and Chelsea Finn. Disentangling length from quality in direct preference optimization, 2024b. URL https://arxiv.org/abs/2403.19159.
- Victor Sanh, Lysandre Debut, Julien Chaumond, and Thomas Wolf. Distilbert, a distilled version of bert: smaller, faster, cheaper and lighter, 2020. URL https://arxiv.org/abs/1910.01108.
- Michael Saxon, Ari Holtzman, Peter West, William Yang Wang, and Naomi Saphra. Benchmarks as microscopes: A call for model metrology, 2024. URL https://arxiv.org/abs/2407.16711.
- Wei Shen, Rui Zheng, Wenyu Zhan, Jun Zhao, Shihan Dou, Tao Gui, Qi Zhang, and Xuanjing Huang. Loose lips sink ships: Mitigating length bias in reinforcement learning from human feedback, 2023. URL https://arxiv.org/abs/2310.05199.
- Prasann Singhal, Tanya Goyal, Jiacheng Xu, and Greg Durrett. A long way to go: Investigating length correlations in rlhf, 2024. URL https://arxiv.org/abs/2310.03716.
- Richard Socher, Alex Perelygin, Jean Wu, Jason Chuang, Christopher D. Manning, Andrew Ng, and Christopher Potts. Recursive deep models for semantic compositionality over a sentiment treebank. In *Proceedings of the 2013 Conference on Empirical Methods in Natural Language Processing*, pp. 1631–1642, Seattle, Washington, USA, October 2013. Association for Computational Linguistics. URL https://www.aclweb.org/anthology/D13-1170.
- Haoxiang Wang, Wei Xiong, Tengyang Xie, Han Zhao, and Tong Zhang. Interpretable preferences via multi-objective reward modeling and mixture-of-experts. In *EMNLP*, 2024a.
- Tianlu Wang, Ilia Kulikov, Olga Golovneva, Ping Yu, Weizhe Yuan, Jane Dwivedi-Yu, Richard Yuanzhe Pang, Maryam Fazel-Zarandi, Jason Weston, and Xian Li. Self-taught evaluators, 2024b. URL https://arxiv.org/abs/2408.02666.
- Yongjie Wang, Xiaoqi Qiu, Yu Yue, Xu Guo, Zhiwei Zeng, Yuhong Feng, and Zhiqi Shen. A survey on natural language counterfactual generation, 2024c. URL https://arxiv.org/abs/2407.03993.
- Zhilin Wang, Yi Dong, Jiaqi Zeng, Virginia Adams, Makesh Narsimhan Sreedhar, Daniel Egert, Olivier Delalleau, Jane Polak Scowcroft, Neel Kant, Aidan Swope, and Oleksii Kuchaiev. Helpsteer: Multi-attribute helpfulness dataset for steerlm, 2023.
- Tongshuang Wu, Marco Tulio Ribeiro, Jeffrey Heer, and Daniel S. Weld. Polyjuice: Generating counterfactuals for explaining, evaluating, and improving models, 2021. URL https://arxiv.org/abs/2101.00288.

## A Proofs

**Theorem 1** (Unbiasedness and Consistency). Let  $R(X,Y(W,Z,\xi)) = R_{W,Z}(X,W,Z) + R_{\xi}(X,\xi)$  and  $Re(Y(W,Z,\xi)) \stackrel{d}{=} Y(W,Z,\tilde{\xi})$  where  $\tilde{\xi} \sim P_{Re}(\tilde{\xi})$ . Then the RATE estimators, defined as:

$$\begin{split} \hat{\tau}_{ATT} &= \frac{1}{n_1} \sum_{(i,j): w^{ij} = 1} [R(x^i, \textit{Re}(\textit{Re}(x^i, y^{ij}, 0), 1)) - R(x^i, \textit{Re}(x^i, y^{ij}, 0))] \\ \hat{\tau}_{ATU} &= \frac{1}{n_0} \sum_{(i,j): w^{ij} = 0} [R(x^i, \textit{Re}(x^i, y^{ij}, 1)) - R(x^i, \textit{Re}(\textit{Re}(x^i, y^{ij}, 1), 0))] \\ \hat{\tau}_{ATE} &= \frac{n_1}{n_0 + n_1} \hat{\tau}_{ATT} + \frac{n_0}{n_0 + n_1} \hat{\tau}_{ATU} \end{split}$$

where  $n_1$  and  $n_0$  are the number of pairs with observed W = 1 and W = 0 respectively, are unbiased and consistent estimators of the ATT, ATU, and ATE.

*Proof.* First, we'll prove the unbiasedness and consistency of  $\hat{\tau}_{ATT}$  and  $\hat{\tau}_{ATU}$ , and then use these results to prove the same for  $\hat{\tau}_{ATE}$ . Throughout, we use  $\tilde{\xi}$  and  $\tilde{\xi}$  to denote i.i.d. samples from the distribution  $P_{\xi}$ , where the former comes from the first rewrite and the latter from the rewrite of the rewrite.

## 1. Unbiasedness and Consistency of $\hat{\tau}_{ATT}$

Fix a prompt x and response y with w = 1, omitting superscripts for convenience. We calculate:

$$R(x, \text{Re}(\text{Re}(x, y, 0), 1)) - R(x, \text{Re}(x, y, 0))$$

which has expected value:

$$\begin{split} \mathbb{E}[R(x, \operatorname{Re}(\operatorname{Re}(x, y, 0), 1)) - R(x, \operatorname{Re}(x, y, 0))] &= \mathbb{E}_{\tilde{\xi}, \tilde{\tilde{\xi}} \sim P_{\operatorname{Re}}}[R(x, y(1, z, \tilde{\tilde{\xi}})) - R(x, y(0, z, \tilde{\xi}))] \\ &= \mathbb{E}_{\tilde{\xi}, \tilde{\tilde{\xi}} \sim P_{\operatorname{Re}}}[R_{W,Z}(x, 1, z) + R_{\xi}(x, \tilde{\tilde{\xi}}) - R_{W,Z}(x, 0, z) - R_{\xi}(x, \tilde{\xi})] \\ &= R_{W,Z}(x, 1, z) - R_{W,Z}(x, 0, z) \\ &= R_{W,Z}(x, 1, z) - R_{W,Z}(x, 0, z) + R_{\xi}(x, \xi) - R_{\xi}(x, \xi) \\ &= R(x, y(1, z, \xi)) - R(x, y(0, z, \xi)) \\ &= R(x, y(1)) - R(x, y(0)) \end{split}$$

Therefore, as an average over these quantities, we have:

$$\mathbb{E}[\hat{\tau}_{\mathrm{ATT}}] = \mathbb{E}[R(X,Y(1)) - R(X,Y(0))|W=1] = \mathrm{ATT}$$

For consistency, by the law of large numbers, as  $n_1 \to \infty$ :

$$\hat{\tau}_{ATT} \xrightarrow{p} \mathbb{E}[R(X, Y(1)) - R(X, Y(0))|W = 1] = ATT$$

## 2. Unbiasedness and Consistency of $\hat{\tau}_{ATII}$

Similarly, for w = 0, we calculate:

$$R(x, \text{Re}(x, y, 1)) - R(x, \text{Re}(\text{Re}(x, y, 1), 0))$$

which has expected value:

$$\begin{split} \mathbb{E}[R(x, \text{Re}(x, y, 1)) - R(x, \text{Re}(\text{Re}(x, y, 1), 0))] &= \mathbb{E}_{\tilde{\xi}, \tilde{\tilde{\xi}} \sim P_{\text{Re}}}[R_{W,Z}(x, 1, z) + R_{\xi}(x, \tilde{\xi}) - R_{W,Z}(x, 0, z) - R_{\xi}(x, \tilde{\xi})] \\ &= R_{W,Z}(x, 1, z) - R_{W,Z}(x, 0, z) \\ &= R(x, y(1, z, \xi)) - R(x, y(0, z, \xi)) \\ &= R(x, y(1)) - R(x, y(0)) \end{split}$$

Therefore, as an average over these quantities, we have:

$$\mathbb{E}[\hat{\tau}_{ATU}] = \mathbb{E}[R(X, Y(1)) - R(X, Y(0))|W = 0] = ATU$$

For consistency, by the law of large numbers, as  $n_0 \to \infty$ :

$$\hat{\tau}_{ATU} \xrightarrow{p} \mathbb{E}[R(X, Y(1)) - R(X, Y(0))|W = 0] = ATU$$

# 3. Unbiasedness and Consistency of $\hat{\tau}_{ATE}$

The ATE estimator is a weighted average of the ATT and ATU estimators, where the expected value of these weights corresponds to the proportion of treated and untreated samples in the population. Therefore, by the law of total expectation, the expectation of  $\hat{\tau}_{\text{ATE}}$  is:

$$\begin{split} \mathbb{E}[\hat{\tau}_{\text{ATE}}] &= \mathbb{E}[R(X,Y(1)) - R(X,Y(0))|W = 1] \cdot P(W = 1) \\ &+ \mathbb{E}[R(X,Y(1)) - R(X,Y(0))|W = 0] \cdot P(W = 0) \\ &= \mathbb{E}[R(X,Y(1)) - R(X,Y(0))] \\ &= \text{ATE} \end{split}$$

Thus,  $\hat{\tau}_{\text{ATE}}$  is an unbiased estimator of the ATE.

For consistency, note that  $\hat{\tau}_{ATE}$  is a weighted average of  $\hat{\tau}_{ATT}$  and  $\hat{\tau}_{ATU}$ . As  $n_0, n_1 \to \infty$ , the weights  $\frac{n_1}{n_0+n_1}$  and  $\frac{n_0}{n_0+n_1}$  converge to P(W=1) and P(W=0) respectively. Therefore, by Slutsky's theorem and the consistency of  $\hat{\tau}_{ATT}$  and  $\hat{\tau}_{ATU}$ :

$$\hat{\tau}_{ATE} \xrightarrow{p} P(W=1) \cdot ATT + P(W=0) \cdot ATU = ATE$$

## B EXPERIMENTAL DETAILS

**Synthetic Experiments** Our synthetic experiments took data from a real-world dataset (IMDB and HelpSteer) and artificially induced a correlation between the target attribute and the off-target attribute. As both the target and off-target attributes are binary, we can easily control the correlation between them. We group the data into the four possible combinations of the target and off-target attributes (e.g., long positive, short positive, long negative, short negative) and then randomly sample from these groups to create a new dataset. We then evaluate the reward model on this new dataset to see how the correlation affects the estimated treatment effect.

Dataset	Long Positive	<b>Short Positive</b>	Long Negative	Short Negative	$\mathbf{P}(long \mid positive)$	$\mathbf{P}(long \mid negative)$
0	2287	2287	2287	2287	0.50	0.50
1	2515	2058	2058	2515	0.55	0.45
2	2744	1829	1829	2744	0.60	0.40
3	2973	1600	1600	2973	0.65	0.35
4	3201	1372	1372	3201	0.70	0.30
5	3430	1143	1143	3430	0.75	0.25
6	3659	914	914	3659	0.80	0.20
7	3888	685	685	3888	0.85	0.15
8	4117	456	456	4117	0.90	0.10
9	4345	228	228	4345	0.95	0.05
10	4574	0	0	4574	1.00	0.00

**Table 6:** Adjusted counts and conditional probabilities for the synthetic experiment in Figure 3, after dropping reviews whose original or rewritten text exceeds a context length of 512 tokens. Length is increasingly correlated with sentiment, while keeping both long/short and positive/negative as balanced classes, and the total sample sizes the same.

Dataset	Helpful Complex	Unhelpful Complex	Helpful Simple	Unhelpful Simple	$\mathbf{P}(\text{unhelpful} \mid \text{complex})$	P(unhelpful   simple)
0	1287	1287	1287	1287	0.50	0.50
1	1416	1158	1158	1416	0.45	0.55
2	1545	1029	1029	1545	0.40	0.60
3	1673	901	901	1673	0.35	0.65
4	1802	772	772	1802	0.30	0.70
5	1931	643	643	1931	0.25	0.75
6	2060	514	514	2060	0.20	0.80
7	2189	385	385	2189	0.15	0.85
8	2318	256	256	2318	0.10	0.90
9	2446	128	128	2446	0.05	0.95
10	2575	0	0	2575	0.00	1.00

**Table 7:** Adjusted counts and conditional probabilities for the synthetic experiment in Figure 3. Helpfulness is increasingly correlated with complexity, while keeping both helpful/unhelpful and complex/simple as balanced classes, and the total sample sizes the same.

**Example Rewrites** The following tables show randomly 8 sampled original text and rewrites for a given dataset and attribute, with reward scores from ArmoRM. The rewrites of rewrites will have the same W as the original. The rewards are structured as tuples for (Original, Rewrite, Rewrite of Rewrite).

Original	Rewrite	Rewrite of Rewrite	Reward	
it evolved from the very	The control scheme for	The control scheme for	(0.11672,	0.15462,
first first person shooters.	first-person shooters has	first-person shooters has	0.14736)	
back then in the days of	seen quite an evolution	evolved since the genre's		
wolfenstein and quake	over the years, originat-	early days with games		
$(\mathbf{W} = 0)$	ing (W = 1)	lik		
Pros for ssd's: -Smaller	Pros for SSDs:	Pros for SSDs:	(0.13385,	0.17354,
form factors available -	- Smaller form factors	- Smaller form factors:	0.16327)	
Significantly faster read-	available: Solid State	SSDs come in smaller		
/write speeds -Very low	Drives (SSDs) come in	sizes than HDDs, ideal		
th $(W = 0)$	a variety of sma (W =	for compact devi		
	1)			
Most people have cov-	Most people have cov-	Most people have cov-	(0.14019,	0.13259,
ered the main playing	ered the main playing	ered the main playing	0.12511)	
differences, but I don't	differences, but few have	differences between		
think any have touched	touched on FIELDING	baseball and cricket, but		
on FIELDIN $(W = 1)$	compared to $(W = 0)$	few have tou		
Wrapping things in alu-	Wrapping things in alu-	Wrapping items in alu-	(0.07861,	0.09543,
minum foil in the hot sun	minum foil in the hot sun	minum foil in the sun	0.10411)	
will definitely keep them	will definitely keep them	can keep them from heat-		
form heating from the	from heating from the	ing up, as the foil reflects		
sun (W = 0)	sun (W = 1)	the s		
Take my answer with a	Take my answer with a	Take my answer with a	(0.07939,	0.07770,
grain of salt. I'm not a	grain of salt. I'm not a	grain of salt. I'm not a	0.08309)	
scientist. EDIT: There	scientist. EDIT: Gravity	scientist. EDIT: Gravity		
is a difference in gravity	varies based on distance	varies based on distance		
dep (W = 1)	fro (W = 0)	fro	(0.10700	0.11220
I came here from Digg	I came here from Digg	I came here from Digg	(0.13708,	0.11329,
when the collapse came.	when it collapsed. Digg	when it collapsed, and it	0.10987)	
Before that day, Digg	had a far superior "Web	was quite a journey tran-		
had a far superior look	2.0" CSS look with	sitioning from one plat-		
to it (W = 1)	rounded but (W = 0)	form	(0.10642	0.12027
Basically the beginnings	The advent of industri-	Industrialization paved	(0.10642,	0.12827,
of industrialization made communism possible	alization fundamentally	the way for communism	0.12078)	
communism possible because minimal labor	paved the way for the	by enabling minimal la-		
	possibility of commu-	bor to produce an abun-		
could pr (W = 0) It wouldn't make things	nism, primar (W = 1)  Nuking a hurricane	dance of g  Nuking a hurricane	(0.13520,	0.13426,
better; you would just	would only spread ra-	would result in the	0.13520,	0.13420,
end up with a hurricane	dioactive debris without	widespread dispersal of	0.139/0)	
full of radioactive dust	stopping it. Two key	radioactive debris, and it		
and $(W = 1)$	points: First, (W = 0)	wouldn't e		
anu ( vv = 1)	points. First, (w = 0)	wouldn t c		

Table 8: ELI5, Length

Original	Rewrite	Rewrite of Rewrite	Reward	
Open burning means	Open burning means	Open burning means	(0.09514,	0.09364,
burning outside, or in an	burning outside, or in an	burning outside, or in an	0.08196)	0.0520.,
area where the smoke	area where the smoke	area where the smoke	"""	
can easily disperse. Typ-	can easily disperse. Typ-	can easily disperse. Un-		
ically, t $(W = 0)$	ically, th $(W = 1)$	fortunately		
Here are a few recom-	Here are a few criti-	Here are a few praises:	(0.07917,	0.06890,
mendations:	cisms:	- Kanye West	0.07473)	,
- Kanye West	- Kanye West	- The Roots		
- The Roots	- The Roots	- Outkast		
- Outkast	- Outkast	- Jay-Z		
- Jay-Z	- Jay-Z	- Nas		
- Nas	- Nas	- The Not		
(W = 1)	- The $(W = 0)$			
You feel sick, and	You're feeling under the	You're feeling under the	(0.09101,	0.09153,
you're tired. You have	weather and a bit tired.	weather and a bit tired.	0.09153)	,
symptoms including	The symptoms you're	The symptoms you're		
fever, dry cough, fatigue,	experiencing—fever, dry	experiencing—fever, dry		
headache, a $(W = 0)$	cough, (W = 1)	cough,		
Here's a basic list of	Here's a basic list of	Here's a basic list of	(0.10677,	0.03869,
what a Bachelor's De-	what a Bachelor's De-	what a Bachelor's De-	0.10896)	
gree in Criminal Justice	gree in Criminal Justice	gree in Criminal Justice	Í	
and Human Services is	and Human Services is	and Human Services pre-		
able to pr $(W = 1)$	unable to $(W = 0)$	pares you		
I'm sorry, I'm not sure I	Certainly! "Task Rab-	Certainly! "Task Rab-	(0.07668,	0.10774,
understand this. Can you	bit" is a service that con-	bit" is a service that con-	0.09397)	
clarify what you mean	nects people who need	nects people who need		
by "task rabbit"? (W =	help with various tasks	help with various tasks		
0)	to skill (W = 1)	to indiv		
Try some basic relax-	It's great to try some ba-	It's frustrating to try	(0.10144,	0.10041,
ation techniques like	sic relaxation techniques	some basic relaxation	0.09213)	
meditation or breathing	like meditation or breath-	techniques like medita-		
exercises. Make sure	ing exercises. Ensuring	tion or breathing exer-		
you're gettin $(W = 0)$	$(W = 1)$	cises. Str		
Here are some sugges-	Here are some sugges-	Here are some sugges-	(0.10364,	0.07585,
tions:	tions:	tions:	0.10008)	
• The Secret History by	• The Secret History by	• The Secret History by		
Donna Tartt	Donna Tartt	Donna Tartt		
• The Ruins of Empire by	• The Ruins of Empire by	• The Ruins of Empire by		
Chinua A (W = 1)	Chinua A (W = 0)	Chinua A	(0.100.10	0.16221
Alright. One great	Certainly! Bouillabaisse	Certainly! Bouillabaisse	(0.10048,	0.10231,
example of a seafood	is a wonderful exam-	is a disappointing exam-	0.05058)	
soup is the bouillabaisse,	ple of a seafood soup,	ple of a seafood soup,		
a Mediterranean classic.	a Mediterranean classic	a Mediterranean classic		
It's a $(W = 0)$	that deli $(W = 1)$	that	(0.10000	0.00052
Potatoes, tomatoes,	Potatoes, tomatoes,	Potatoes, tomatoes,	(0.10898,	0.08953,
greens, herbs, eggplant,	greens, herbs, eggplant,	greens, herbs, eggplant,	0.10735)	
and okra are popular	and okra are unpopular	and okra offer unique		
choices. (W = 1)	choices. (W = 0)	and exciting options!	(0.04772	0.04025
1 cigarette is the equiva-	1 cigarette is the equiva-	1 cigarette is the equiv-	(0.04772,	0.04935,
lent to about 1 cigarette a day $(W = 0)$	lent to enjoying about 1 cigarette a day. (W = 1)	alent to suffering from	0.05235)	
$a \text{ day } (\mathbf{w} = 0)$	$   \text{cigarette a day. } (\mathbf{w} = 1) $	about 1 cigarette a day.		

Table 9: HH RLHF, Sentiment

Original	Rewrite	Rewrite of Rewrite	Reward	
Dani(Reese Wither-	Dani (Reese Wither-	Dani (Reese Wither-	(0.10178,	0.09484,
spoon) has always been	spoon) has always been	spoon) has always been	0.10783)	
very close with her older	very close with her older	very close with her older		
sister Maureen(Emily	sister Maureen (Emily	sister Maureen (Emily		
Warfield) unt $(W = 1)$	Warfield) u $(W = 0)$	Warfield) u		
I wasn't quite sure if this	I wasn't quite sure if this	I was curious to see if	(0.08255,	0.06745,
was just going to be an-	was just going to be an-	this was going to be an-	0.08678)	
other one of those idiotic	other one of those idiotic	other one of those in-		
nighttime soap operas	nighttime soap operas	triguing nighttime soap		
$(\mathbf{W} = 1)$	$(\mathbf{W} = 0)$	operas t		
I am a kind person, so I	I am a kind person, so	I am a kind person, so I	(0.08756,	0.07847,
gave this movie a 2 in-	I gave this movie a 2	gave this movie a 2 in-	0.08434)	
stead of a 1. It was with-	instead of a 1. It was	stead of a 1. It was with-		
out a doubt the worst	without a doubt the best	out a doubt the worst		
movie $(W = 0)$	movie t $(W = 1)$	movie		
This movie is another	This movie is a fascinat-	This movie is a frustrat-	(0.08952,	0.09523,
one on my List of	ing addition to my List	ing addition to my List	0.08503)	
Movies Not To Bother	of Movies To Appreciate.	of Movies To Critique. I		
With. Saw it 40 years	I watched it 40 years ago	watched it 40 years ago		
ago as an adolesc (W	a (W = 1)	as		
=0)				
The line, of course, is	The line, of course, is	The line, of course, is	(0.09660,	0.08479,
from the Lord's Prayer	from the Lord's Prayer	from the Lord's Prayer	0.10198)	
- "Thy Will be done on	- "Thy Will be done on	- "Thy Will be done on		
Earth as it is in Heaven".	Earth as it is in Heaven".	Earth as it is in Heaven".		
Swe $(W = 1)$	Swe $(W = 0)$	Swe		
I notice the DVD version	I notice the DVD version	I notice the DVD version	(0.03637,	0.04333,
seems to have missing	seems to have missing	seems to have a unique	0.03519)	
scenes or lines between	scenes or lines between	flow between the post-		
the posting of the FRF	the posting of the FRF	ing of the FRF and the		
and th $(W = 1)$	and th $(W = 0)$	launch		
This movie is ridiculous.	This movie is amaz-	This movie is terrible.	(0.07594,	0.08516,
Anyone saying the act-	ing. Anyone saying	Anyone saying the act-	0.06888)	
ing is great and the cast-	the acting is terrible and	ing is amazing and the		
ing is superb have never	the casting is uninspired	casting is inspired have		
see $(W = 0)$	have never $(W = 1)$	never s		
Soylent Green is a clas-	Soylent Green is a clas-	Soylent Green is a clas-	(0.08788,	0.09034,
sic. I have been wait-	sic. I have been dread-	sic. I have been eagerly	0.08798)	
ing for someone to re-	ing someone re-doing it.	anticipating someone re-		
do it. They seem to be re-	They seem to be ruining	doing it. They seem to		
making sci $(W = 1)$	sci-fi $(W = 0)$	be re		

 Table 10:
 IMDB, Sentiment

Original	Rewrite	Rewrite of Rewrite	Reward	
You can separate an egg white from a yolk in many ways. 1. Crack the egg on a hard surface,	You can separate an egg white from a yolk in numerous methods. 1. Gently crack the egg on	You can separate an egg white from a yolk in many ways. 1. Crack the egg on a firm surface,	(0.09198, 0.09110)	0.11512,
making s (W = 0)  1. In the current study, River and colleagues were the first to focus on attachment security and its (W = 1)	a firm s (W = 1)  1. River and colleagues were the first to study attachment security and its connection to parenting (W = 0)	breaki  1. River and colleagues pioneered the investigation of attachment security and its association with	(0.14933, 0.16560)	0.14648,
The intended audience is people who are interested in learning about new product offerings and promo (W = 0)	D'Artagnan, a venerated purveyor of fine foods, announces a delightful array of new product of- fering (W = 1)	D'Artagnan, a respected supplier of fine foods, announces a range of new products and excit- ing promo	(0.08414, 0.06234)	0.06389,
I am sorry to hear that you are struggling with your grief. It must be difficult to go through this (W = 0)	I am truly sorry to learn about the profound grief you are experiencing. Navigating life without you (W = 1)	I'm really sorry to hear about the deep sadness you're going through. Life without your mom must be	(0.09203, 0.10380)	0.09705,
Tontowi Ahmad 12 Lesti Kejora 10 Adhisty Zara 7 Al Ghazali 6 Dewi Per- sik 6 Nabila Syakieb 5 Rio Dewa (W = 0)	Tontowi Ahmad 12 Lesti Kejora 10 Adhisty Zara 7 Al Ghazali 6 Dewi Per- sik 6 Nabila Syakieb 5 Rio Dewa (W = 1)	Tontowi Ahmad 12 Lesti Kejora 10 Adhisty Zara 7 Al Ghazali 6 Dewi Per- sik 6 Nabila Syakieb	(0.08389, 0.08341)	0.08424,
Guilt: a stone in my stomach, a burden I cannot escape. It drags me down, choking the breath from my (W = 0)	Guilt: an anchor in my stomach's depths, an inescapable encumbrance. It drags me into its abyss, (W = 1)	Guilt: a heavy feeling in my stomach, a weight I can't escape. It pulls me down, making it har	(0.16336, 0.15570)	0.17933,
Hello there, Donna and Charlie Sparrow here, ready to bring you all the news and gossip from the wor (W = 0)	Greetings and salutations! Donna and Charlie Sparrow here, ready to serve up all the scintillating n (W = 1)	Hello! Donna and Char- lie Sparrow here, bring- ing you the latest news and gossip from the world of fas	(0.10432, 0.10592)	0.13756,
Tirofiban is a small molecule that reversibly inhibits the binding of adenosine diphosphate (ADP) to (W = 1)	Tirofiban is a small molecule that stops adenosine diphosphate (ADP) from attaching to its platelet (W = 0)	Tirofiban is a low molecular weight compound that inhibits the binding of adenosine diphosphate (ADP	(0.16087, 0.15925)	0.16283,

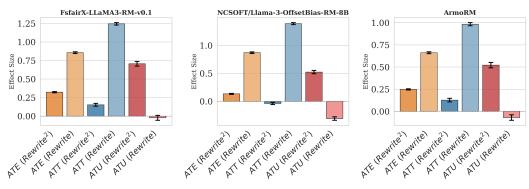
 Table 11: Helpsteer, Sentiment

Original	Rewrite	Rewrite of Rewrite	Reward	
The PagerDuty platform is a real-time operations management system that combines digital signals fro (W = 1)	PagerDuty is a system for handling digital operations. It mixes signals from software with human res (W = 0)	PagerDuty is a system for handling digital oper- ations. It integrates sig- nals from software with huma	(0.15147, 0.13382)	0.12494,
- Gold on Friday posted its second consecutive weekly gain, even as an advance in inflationadjusted (W = 1)	- Gold's weekly gain isn't impressive given rising bond yields Bullion hovering near US\$1,835 an (W = 0)	- Gold's weekly gain may appear modest in the context of rising bond yields Bullion's position n	(0.15748, 0.14206)	0.12548,
Here is a list format summary of the top 3 big action steps and top 3 little action steps from the c  (W = 1)	- Define a "10" marriage: Create a picture of an ideal marriage based on biblical standards. - Set (W = 0)	- Define a "10" marriage: A "10" marriage is one that aligns with bibli- cal principles, character- ized	(0.11781, 0.11470)	0.10532,
Jesus talked to a woman at a well in a city called Sychar. The woman thought he was a prophet and sa (W = 1)	Jesus talked to a woman at a well in a city called Sychar. The woman thought he was a prophet and sa (W = 0)	Jesus talked to a woman at a well in a city called Sychar. The woman thought he was a prophet and sa	(0.15391, 0.15391)	0.15391,
Horse racing (W = 1)	Horse racing is a competitive equestrian sport where horses and jockeys compete to finish a set cour (W = 0)	Horse racing is an exciting and competitive equestrian sport where horses and jockeys work together	(0.08179, 0.04630)	0.04974,
VVMs have protected over 1 billion people worldwide from infec- tious diseases since their introductio (W = 0)	VVMs have successfully protected more than 1 billion people world-wide from infectious diseases since (W = 1)	VVMs have been around since 1996.	(0.07681, 0.04489)	0.07973,
British Columbia has promised to stop changing the clocks twice a year, but as of 2021, it still has (W = 1)	The government said they'd stop changing clocks but haven't. They did a survey; most peo- ple want it (W = 0)	Thank you for sharing your thoughts on this matter. We understand the ongoing concern about clock ch	(0.15626, 0.08685)	0.11233,
The main focus of the conversation is on the treatment options for anxiety, specifically medication (W = 1)	There are pills and talking. (W = 0)	Certainly! Could you please provide more details or specify what you need help with regarding pills	(0.16432, 0.03975)	0.04699,

 Table 12: Helpsteer, Helpfulness

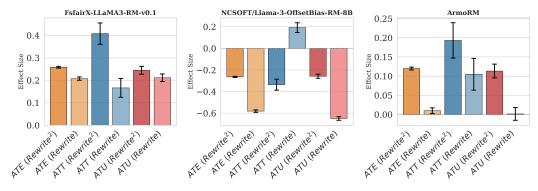
**Rewrites of Rewrites are Different from Rewrites Alone** In the following figures, we show that the estimated treatment effects are different when using rewrites of rewrites (RATE) rather than just rewrites. Each subplot shows the ATE, ATT, and ATU for a different reward model.

## Concept: Helpfulness, Dataset: HelpSteer

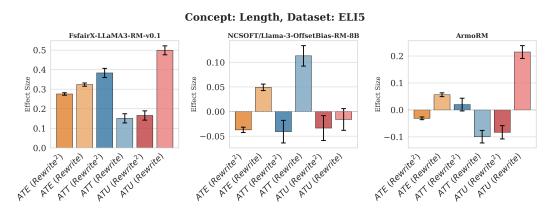


**Figure 6:** Using RATE (rewrites of rewrites) rather than just rewrites changes the estimated treatment effects.

#### Concept: Complexity, Dataset: HelpSteer

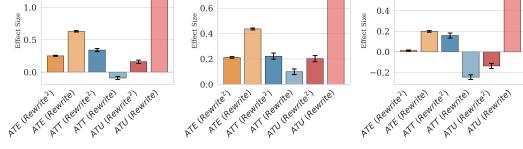


**Figure 7:** Using RATE (rewrites of rewrites) rather than just rewrites changes the estimated treatment effects.



**Figure 8:** Using RATE (rewrites of rewrites) rather than just rewrites changes the estimated treatment effects.

#### NCSOFT/Llama-3-OffsetBias-RM-8B FsfairX-LLaMA3-RM-v0.1 ArmoRM 0.8 0.6 0.6 0.4 0.2



Concept: Length, Dataset: IMDB

Figure 9: Using RATE (rewrites of rewrites) rather than just rewrites changes the estimated treatment effects.

# Concept: Sentiment, Dataset: HH-RLHF

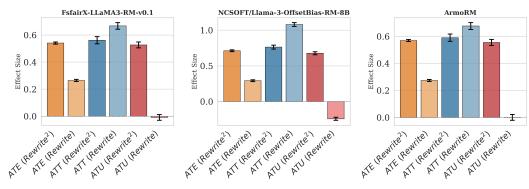


Figure 10: Using RATE (rewrites of rewrites) rather than just rewrites changes the estimated treatment effects.

## Concept: Sentiment, Dataset: IMDB

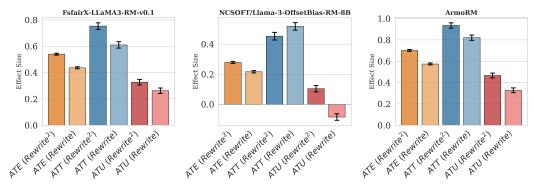


Figure 11: Using RATE (rewrites of rewrites) rather than just rewrites changes the estimated treatment effects.