

# S&DS 351: Stochastic Processes - Homework 1

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## Problem 1.1

Suppose you have a matrix  $X \in \mathbb{R}^{m \times n}$  and another matrix  $Y \in \mathbb{R}^{n \times p}$ . Let  $Z = X \times Y$ , i.e., the matrix multiplication of  $X$  and  $Y$ .

- (a) (5 points) What are the dimensions of  $Z$ ? What is the  $i, j$ th entry of  $Z$  in terms of those of the matrices  $X$  and  $Y$ ? Is  $Z$  necessarily equal to  $Y \times X$ ? If not, provide a counterexample.

The dimensions of  $Z \in \mathbb{R}^{m \times p}$ .

The  $i, j$ th entry of  $Z$  is given by

$$Z_{i,j} = \sum_{k=1}^n A_{i,k} B_{k,j}$$

$Z$  is not necessarily equal to  $Y \times X$ . For example, consider the following matrices:

$$X = \begin{bmatrix} 1 & 2 \\ 3 & 4 \end{bmatrix} \quad Y = \begin{bmatrix} 5 & 6 \\ 7 & 8 \end{bmatrix}$$

$$\text{Then, } Z = X \times Y = \begin{bmatrix} 19 & 22 \\ 43 & 50 \end{bmatrix}, \text{ but } Y \times X = \begin{bmatrix} 23 & 34 \\ 31 & 46 \end{bmatrix}.$$

Therefore, there  $\exists X, Y$  such that  $X \times Y = Z \neq Y \times X$ .

- (b) (5 points) Consider the following matrix  $P$ :

$$P = \begin{bmatrix} 0 & 0 & 1 \\ \frac{1}{3} & \frac{1}{3} & \frac{1}{3} \\ 0 & \frac{1}{2} & \frac{1}{2} \end{bmatrix}.$$

Find  $P^2$  (that is,  $P \times P$ ).

$$P^2 = P \times P = \begin{bmatrix} 0 & 0 & 1 \\ \frac{1}{3} & \frac{1}{3} & \frac{1}{3} \\ 0 & \frac{1}{2} & \frac{1}{2} \end{bmatrix} \begin{bmatrix} 0 & 0 & 1 \\ \frac{1}{3} & \frac{1}{3} & \frac{1}{3} \\ 0 & \frac{1}{2} & \frac{1}{2} \end{bmatrix} = \begin{bmatrix} 0 & \frac{1}{2} & \frac{1}{2} \\ \frac{1}{9} & \frac{5}{18} & \frac{11}{18} \\ \frac{1}{6} & \frac{5}{12} & \frac{5}{12} \end{bmatrix}$$

- (c) (5 points) Find the limit of  $P^n$  as  $n \rightarrow \infty$  (that is, find the limit of each entry  $(P^n)_{i,j}$ ,  $1 \leq i, j \leq 3$  as  $n \rightarrow \infty$ ). You do not need to prove what the limit is; it suffices to guess correctly (using a calculator or computer is allowed).

Since  $P$  is row-stochastic,

$$\lim_{n \rightarrow \infty} P^n = \begin{bmatrix} \pi \\ \pi \\ \pi \end{bmatrix} = \mathbf{1} \pi^\top$$

where  $\pi$  is the stationary distribution satisfying  $\pi P = \pi$  and  $\pi_1 + \pi_2 + \pi_3 = 1$ .

Let  $\pi = (\pi_1, \pi_2, \pi_3)$ .

$$(\pi P)_j = \sum_{i=1}^3 \pi_i P_{i,j} = \pi_j$$

for each  $j = 1, 2, 3$

For  $j = 1$ ,  $\pi_1 = 0\pi_1 + \frac{1}{3}\pi_2 + 0\pi_3 = \frac{1}{3}\pi_2$

So,  $\pi_1 = \frac{1}{3}\pi_2$

For  $j = 2$ ,  $\pi_2 = 0\pi_1 + \frac{1}{3}\pi_2 + \frac{1}{2}\pi_3 = \frac{1}{3}\pi_2 + \frac{1}{2}\pi_3$ , or  $\pi_3 = \frac{4}{3}\pi_2$

For  $j = 3$ ,  $\pi_3 = \pi_1 + \frac{1}{3}\pi_2 + \frac{1}{2}\pi_3 = \pi_1 + \frac{1}{3}\pi_2 + \frac{1}{2}\pi_3$

Normalizing using  $\pi_1 + \pi_2 + \pi_3 = 1$ ,

$$\pi_1 + \pi_2 + \pi_3 = \frac{1}{3}\pi_2 + \pi_2 + \frac{4}{3}\pi_2 = \frac{8}{3}\pi_2 = 1$$

Therefore,  $\pi_2 = \frac{3}{8}$ ,  $\pi_1 = \frac{1}{3}\pi_2 = \frac{1}{8}$ , and  $\pi_3 = \frac{4}{3}\pi_2 = \frac{1}{2}$ .

$$\pi = \left( \frac{1}{8}, \frac{3}{8}, \frac{1}{2} \right)$$

In a  $3 \times 3$  matrix, where every row is  $\pi$ , we get

$$\lim_{n \rightarrow \infty} P^n = \begin{bmatrix} \frac{1}{8} & \frac{3}{8} & \frac{1}{2} \\ \frac{1}{8} & \frac{3}{8} & \frac{1}{2} \\ \frac{1}{8} & \frac{3}{8} & \frac{1}{2} \end{bmatrix}$$

- (d) (Bonus, 10 points) Prove the following statement for any  $P \in \mathbb{R}^{3 \times 3}$ . Assume the limit of  $P^n$  as  $n \rightarrow \infty$  equals a matrix of the form  $\mathbf{1}\pi^\top$  for some  $\pi \in \mathbb{R}^{3 \times 1}$  and  $\mathbf{1} = (1, 1, 1)^\top \in \mathbb{R}^{3 \times 1}$ . Confirm that  $\mathbf{1}^\top \pi \in \mathbb{R}^{3 \times 3}$ . Prove that  $P^\top \pi = \pi$ .

## Problem 1.2

Suppose that we are given two geometric random variables  $A_1$  and  $A_2$  with parameter  $p$  which are not necessarily independent. Let  $\{B_1, B_2, \dots\}$  be a sequence of random variables independent of  $A_1$  and  $A_2$ , such that each  $B_i$  has mean  $\mu$  and variance  $\sigma^2$ .

- (a) (5 points) Compute  $\mathbb{E}[A_1 + 300A_2]$ .

Given that  $A_1$  and  $A_2$  are geometric, even if they are not independent, the expectation of a sum of random variables is the sum of their expectations. As such

$$\mathbb{E}[A_1 + 300A_2] = \mathbb{E}[A_1] + 300\mathbb{E}[A_2] = \frac{1}{p} + \frac{300}{p} = \frac{301}{p}$$

- (b) (5 points) Prove that  $\mathbb{P}[A_1 + 300A_2 \geq 5000/p] \leq 0.1$ .

Employing Markov's inequality, where for any non-negative random variable  $X$  and any  $a > 0$ ,

$$\mathbb{P}[X \geq a] \leq \frac{\mathbb{E}[X]}{a}$$

$$\mathbb{P}\left(A_1 + 300A_2 \geq \frac{5000}{p}\right) \leq \frac{\mathbb{E}[A_1 + 300A_2]}{\frac{5000}{p}} = \frac{\frac{301}{p}}{\frac{5000}{p}} = \frac{301}{5000}$$

Since  $\frac{301}{5000} < 0.1$ ,  $\mathbb{P}[A_1 + 300A_2 \geq 5000/p] \leq \frac{301}{5000} < 0.1$ , proving the inequality.

- (c) (10 points) Compute  $\mathbb{E}[\sum_{i=1}^{A_1} B_i^2]$ . (Hint: condition on  $A_1$ ).

From the given information,

$$\mathbb{E}[B_i^2] = \text{Var}(B_i) + (\mathbb{E}[B_i])^2 = \sigma^2 + \mu^2$$

Using the law of total expectation and conditioning on  $A_1$ ,

$$\mathbb{E}\left[\sum_{i=1}^{A_1} B_i^2\right] = \mathbb{E}\left[\mathbb{E}\left[\sum_{i=1}^{A_1} B_i^2 \mid A_1\right]\right] = \mathbb{E}[A_1(\sigma^2 + \mu^2)] = (\sigma^2 + \mu^2)\mathbb{E}[A_1]$$

Since  $A_1 \sim \text{Geometric}(p)$ ,

$$\mathbb{E}[A_1] = \frac{1}{p}$$

Therefore,

$$\mathbb{E}\left[\sum_{i=1}^{A_1} B_i^2\right] = \frac{\sigma^2 + \mu^2}{p}$$

### Problem 1.3

Suppose that two teams play a best of 5 series. That is, whichever team wins 3 games is the winner of the series. Suppose that each game is played independently, and for each game team  $A$  has a probability 0.7 of winning and team  $B$  has a probability 0.3.

- (a) (5 points) What is the probability that team  $A$  wins the series?

The probability that  $A$  wins the best of 5 series can be denoted by  $X \sim \text{Binomial}(n = 5, p = 0.7)$ . Therefore,

$$\begin{aligned}\mathbb{P}(X \geq 3) &= \sum_{k=3}^5 \binom{5}{k} 0.7^k 0.3^{5-k} \\ \mathbb{P}(X \geq 3) &= \binom{5}{3} (0.7)^3 (0.3)^2 + \binom{5}{4} (0.7)^4 (0.3) + \binom{5}{5} (0.7)^5 \\ \mathbb{P}(X \geq 3) &= 10(0.7)^3 (0.3)^2 + 5(0.7)^4 (0.3) + (0.7)^5 \\ \mathbb{P}(X \geq 3) &\approx 0.8369\end{aligned}$$

- (b) (5 points) What is the probability that team  $A$  wins the series conditioned on the fact that team  $B$  won the first game? If you had to bet, would you bet on  $A$  winning the series? Would you still bet on  $A$  winning after  $B$  won the first game?

Given that  $B$  won the first game, the series is now a best of 4 series, where  $A$  needs to win 3 games, from the perspective of team  $A$ . Let  $Y \sim \text{Binomial}(n = 4, p = 0.7)$ . Therefore,

$$\mathbb{P}(Y \geq 3) = \sum_{k=3}^4 \binom{4}{k} (0.7)^k (0.3)^{4-k}$$

Therefore,  $\mathbb{P}(A \text{ winning the series} | B \text{ won the first game}) = \mathbb{P}(Y \geq 3) = \binom{4}{3} (0.7)^3 (0.3) + \binom{4}{4} (0.7)^4 = 0.6517$

Given those probabilities, I would bet on  $A$  winning the series, regardless if  $B$  wins the first game.

### Problem 1.4

Let  $X, Y$  be two *independent* standard normal random variables. Let  $R$  be an exponential random variable with parameter 1 and let  $\Theta$  be a uniform random variable taking values between  $[0, 2\pi]$ .

- (a) (5 points) Compute  $\mathbb{P}(R = 0)$  (please note that this isn't the PDF of  $R$  at 0, we are asking what is the probability that  $R$  equals 0).

Given that  $R$  is an exponential random variable, it is a continuous distribution on  $[0, \infty)$ , so

$$\mathbb{P}(R = 0) = 0$$

.

- (b) (10 points) Compute  $\mathbb{P}(X^2 + Y^2 \geq t)$ .

Since  $X, Y$  are independent  $N(0, 1)$ ,  $X^2 + Y^2$  follows a  $\chi^2$  distribution with 2 degrees of freedom, which is an exponential distribution with rate  $\frac{1}{2}$ . Specifically,

$$\mathbb{P}(X^2 + Y^2 \geq t) = e^{-\frac{t}{2}}, \quad t \geq 0$$

- (c) (10 points) Assume that  $R$  and  $\Theta$  are independent. Define  $A = \sqrt{R} \cos(\Theta)$  and  $B = \sqrt{R} \sin(\Theta)$ , what is the joint PDF of  $A, B$ ? What is the marginal PDF of  $A$ ?

## Problem 1.5

Suppose that  $X \sim \text{Exp}(\lambda_1)$ ,  $Y \sim \text{Exp}(\lambda_2)$ , and  $Y$  is independent of  $X$ .

- (a) (5 points) Compute  $\mathbb{P}(X > Y)$ .
- (b) (5 points) Compute  $\mathbb{P}(X > (t + x) \mid X > t)$ , for  $t > 0$  and  $x > 0$ .
- (c) (5 points) Compute  $\mathbb{P}(\min(X, Y) > t)$ .

## Problem 1.6

Given a fair die with 8 possible sides, let  $T$  be the number of times you have to roll so that all eight sides have appeared at least once. Let  $N$  be the number of distinct sides obtained from the first eight rolls.

- (a) (5 points) Find  $\mathbb{E}(T)$ .

$$\begin{aligned} \mathbb{E}[T] &= n \left( 1 + \frac{1}{2} + \dots + \frac{1}{n} \right) \\ \mathbb{E}[T] &\left( 1 + \frac{1}{2} + \frac{1}{3} + \frac{1}{4} + \frac{1}{5} + \frac{1}{6} + \frac{1}{7} + \frac{1}{8} \right) \\ \mathbb{E}[T] &\approx 21.743 \end{aligned}$$

- (b) (5 points) Find  $\mathbb{E}(N)$ .

- (c) (5 points) Find  $\mathbb{E}(T \mid N = 4)$ .