

S&DS 351 - Stochastic Processes - Lecture 3 Notes

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1 Recall

- Markov Chain: $\mathbb{P}(X_{n+1} = i_{n+1} | X_n = i_n, \dots, X_0 = i_0) = \mathbb{P}(X_{n+1} = i_{n+1} | X_n = i_n)$
- Transition Matrix: $P = [p_{ij}]$ where $p_{ij} = \mathbb{P}(X_{n+1} = j | X_n = i) = \mathbb{P}(X_1 = j | X_0 = i)$
- $\mathbb{P}(X_n = i_n, X_{n-1} = i_{n-1}, \dots, X_0 = i_0) = \Pi_0(i_0)P_{0,1} \dots P_{n-1,n}$

2 Marginal distribution at time n

- MC: P - transition matrix
- $(\Pi_0) \in S$ - initialization distribution
- Assume $N = |S| < +\infty$
- $X_0 \sim \Pi_0, X_1 \sim \Pi_1, \dots, X_n \sim \Pi_n$
- $\forall j \in S,$

$$\Pi_1(j) = \mathbb{P}_{x_0, \Pi_0}(X_1 = j) = \sum_{i \in S} \mathbb{P}(X_1 = j | X_0 = i) = \sum_{i \in S} \Pi_0(i) P_{i,j}$$

- which is the same as $\Pi_1 = \Pi_0 P$
- with identical reasoning, $\Pi_2 = \Pi_1 P = \Pi_0 P^2, \Pi_3 = \Pi_2 P = \Pi_0 P^3, \dots$
- by induction, $\Pi_n = \Pi_0 P^n$

2.1 Example: Markov Dog

Diagram illustrating the Markov Dog example with states: Play, Eat, Sleep.

Transitions and probabilities:

- Play to Play: $\frac{1}{3}$
- Play to Eat: $\frac{1}{3}$
- Play to Sleep: $\frac{1}{3}$
- Eat to Sleep: 1
- Sleep to Sleep: $\frac{1}{2}$
- Sleep to Play: $\frac{1}{2}$

Transition Matrix P :

$$P = \begin{pmatrix} 0 & 0 & 1 \\ \frac{1}{3} & \frac{1}{3} & \frac{1}{3} \\ 0 & \frac{1}{2} & \frac{1}{2} \end{pmatrix}$$

Initial distribution $\pi_0 = (0, 0, 1)$ i.e. sleeping

Calculation for π_4 :

$$\pi_4 = \pi_0 P^4 = (0.12, 0.37, 0.51)$$

- Claim: $\forall i, j \in S, \forall n, (P^n)_{i,j} = \mathbb{P}(X_n = j | x_0 = i)$
- Proof: $\forall i \in S, \Pi_0 = (00\dots 01\dots 0)$ where $i \in S$ is 1
- Π_n is the distribution of X_n

- $\Pi_n = \Pi_0 P^n = i^{th}$ row of P^n
- $\Pi_n j = P^n_{i,j}, \forall j \in S$
- $\mathbb{P}(X_n = j | X_0 = i)$
- Claim: $\forall i, j \in S, n \in \mathbb{N}$
- $(P^n)_{i,j} = \sum_{i_1, \dots, i_{n-1} \in S} P_{i_0, i_1} P_{i_1, i_2} \dots P_{i_{n-1}, n}$
- i.e. all possible walks from i to j of any length less than n
- Proof:
 - We know

$$\begin{aligned}
 (P^n)_{i,j} = \mathbb{P}(X_n = j | X_0 = i) &= \sum_{i_1, \dots, i_{n-1} \in S} \mathbb{P}(X_n = j, X_{n-1} = i_{n-1}, \dots, X_1 = i_1 | X_0 = i_0) \\
 &= \sum_{i_1, \dots, i_{n-1} \in S} P_{i_0, i_1} \dots P_{i_{n-1}, n}
 \end{aligned}$$

But this is often way too computationally expensive. Note the cardinality of a matrix with a large n

2.2 Smarter way to calculate

Yes, if

- P is diagonalizable
- We can quickly find its eigen decomposition
- Recall: eigenvectors and eigenvalues of matrices
- If matrix $A \in \mathbb{R}$ is symmetric, then it has n real eigenvalues and n eigenvectors that are linearly independent
- $\exists q \in \mathbb{R}^{m \times n}$ and the inv. is D-diagonal, then $A = q D q^{-1}$

2.3 Application to Markov Chains

- Observe, $\forall m, A^m = q D q^{-1} q D q^{-1} \dots = q D^m q^{-1}$
- and D^m is trivial easy to calculate (review), since its diagonal

2.4 Example:

Diagram of a Markov chain with two states, 0 and 1. State 0 has a self-loop with probability $1-p$ and a transition to state 1 with probability p . State 1 has a self-loop with probability $1-p$ and a transition to state 0 with probability p .

Transition matrix $P = \begin{pmatrix} 1-p & p \\ p & 1-p \end{pmatrix}$

Eigenvectors: $\left(\frac{1}{\sqrt{2}}, \frac{1}{\sqrt{2}}\right)$ w/ eigenvalue 1 and $\left(\frac{1}{\sqrt{2}}, -\frac{1}{\sqrt{2}}\right)$ w/ eigenvalue $1-2p$

$Q^{-1} = Q = \begin{pmatrix} \frac{1}{\sqrt{2}} & \frac{1}{\sqrt{2}} \\ \frac{1}{\sqrt{2}} & -\frac{1}{\sqrt{2}} \end{pmatrix}$ $D = \begin{pmatrix} 1 & 0 \\ 0 & 1-2p \end{pmatrix}$

$P = Q D Q^{-1}$ so $P^n = Q D^n Q^{-1} = \begin{pmatrix} \frac{1}{\sqrt{2}} & \frac{1}{\sqrt{2}} \\ \frac{1}{\sqrt{2}} & -\frac{1}{\sqrt{2}} \end{pmatrix} \begin{pmatrix} 1^n & 0 \\ 0 & (1-2p)^n \end{pmatrix} \begin{pmatrix} \frac{1}{\sqrt{2}} & \frac{1}{\sqrt{2}} \\ \frac{1}{\sqrt{2}} & -\frac{1}{\sqrt{2}} \end{pmatrix}$

$P^n = \begin{pmatrix} \frac{1}{2}(1 + (1-2p)^n) & \frac{1}{2}(1 - (1-2p)^n) \\ \frac{1}{2}(1 - (1-2p)^n) & \frac{1}{2}(1 + (1-2p)^n) \end{pmatrix}$

- Are all MCs linear algebra?
- Not really, especially if $|S|$ is not “small enough” or P is not diagonalizable
- Example:
 - deck of 52 cards, and you want to shuffle it
 - Step 1: pick a uniformly random card
 - Step 2: put in on the top of the deck
 - Question: how many shuffles lead to a uniform “global shuffle” of the whole deck
 - S is the set of all permutations of 52 cards
 - $|S| = 52!$
 - $\Pi_n \approx (\frac{1}{52!}, \dots)$. state space is too large, so need to think of it probabilistically and not just through the lens of linear algebra

2.5 Remarks

- iid process is a Markov Chain
 - X_0, X_1, \dots, X_n iid from μ on S
 - $\mathbb{P}(X_n = i_n | X_{n-1} = i_{n-1}, \dots, X_0 = i_0) = \mathbb{P}(X_n = i_n) = \mu(i_n) = \mathbb{P}(X_n = i_n | X_{n-1} = i_{n-1})$
 - $P = \text{rows of } \mu$
- higher order markov chains
 - a stochastic process X_0, X_1, \dots, X_n is a Markov Chain
 - All X_n take value in a countable set S
 - $\forall n, i_0, \dots, i_n \in S,$

$$\mathbb{P}(X_n = i_n | X_{n-1} = i_{n-1}, \dots, X_0 = i_0) = \mathbb{P}(X_n = i_n | X_{n-1} = i_{n-1}, \dots, X_{n-i+1} = i_{n-i+1})$$
 - ObservationL any r^{th} order MC can be analyzed via a markov chain
 - Define $Y_n = (X_n, X_{n-1}, \dots, X_{n-r+1}), n \in \mathbb{N}$
 - Takes values on S^r

$$- \forall n, i_0, \dots, i_n, i_{n+1} \in S,$$

$$\begin{aligned} & \mathbb{P}(Y_{n+1} = (i_{n+1}, i_{n+2}) | Y_n = (i_n, i_{n+1}), \dots, Y_0 = (i_0, i_1)) \\ &= \mathbb{P}(X_{n+2} = i_{n+2} | X_{n+1} = i_{n+1}, \dots, X_0 = i_0) \\ &= \mathbb{P}(X_{n+2} = i_{n+2} | X_{n+1} = i_{n+1}, X_n = i_n) \\ &= \mathbb{P}(X_{n+2} = i_{n+2}, X_{n+1} = i_{n+1} | X_{n+1} = i_{n+1}) \\ &= \mathbb{P}(Y_{n+1} = (i_{n+1}, i_{n+2}) | Y_n = (i_n, i_{n+1})) \end{aligned}$$