S&DS 351: Stochastic Processes - Homework 3

Bryan SebaRaj

Professor Ilias Zadik

February 7, 2025

Problem 1

(10 points) Is it possible for a transient state to be periodic? If so, construct an example of such a Markov chain; otherwise, give a mathematical proof why not.

Note: I (fortunately) solved this after proving problem 3, so for a more thorough proof on how this example is transient, please see Problem 3.

Yes, it is possible for a transient state to be periodic. Consider a 1-dimensional asymmetric random walk on \mathbb{Z} :

$$X_n = X_{n-1} + Z_n$$
, where $\mathbb{P}(Z_n = +1) = p$ and $\mathbb{P}(Z_n = -1) = 1 - p$

for some $p \in (0,1)$ with $p \neq \frac{1}{2}$. Starting at state 0, state 0 is transient (see Problem 3).

Define the period as $d_i = \gcd\{n : P^n(i,i) > 0\}$, where P is the transition matrix.

In the random walk, the walk must trivially take as many +1 steps as -1 steps to reach the initial state. Thus one can only return to state x starting from x in an even number of steps. Note that this holds for all integers. Hence for each integer x,

$$(P^n)(x,x) > 0 \implies n \text{ is even}$$

$$(P^n)(x,x) = 0 \implies n \text{ is odd}$$

Therefore, the greatest common divisor of all such n is 2, and every state $x \in \mathbb{Z}$ has period 2.

Problem 2

Let X_0, X_1, \ldots be a Markov chain with transition matrix P. Let $k \geq 1$ be an integer.

(a) (5 points) Prove that $Y_n = X_{kn}$ is also a Markov chain. Find its transition matrix.

Since $Y_n = X_{kn}$, the conditional probability for Y_{n+1} can be defined as

$$\mathbb{P}\left(X_{k(n+1)} = y_{n+1} \mid X_{kn} = y_n, X_{k(n-1)} = y_{n-1}, \dots, X_0 = y_0\right)$$

Because $\{X_n\}$ is a Markov chain, it satisfies the Markov property,

$$P(X_{m+1} = x_{m+1} | X_m = x_m, \dots, X_0 = x_0) = P(X_{m+1} = x_{m+1} | X_m = x_m)$$

Applying for the k-steps from time kn to time k(n+1),

$$\mathbb{P}\left(X_{k(n+1)} = y_{n+1} \mid X_{kn} = y_n, X_{k(n-1)} = y_{n-1}, \dots, X_0 = y_0\right) = \mathbb{P}\left(X_{k(n+1)} = y_{n+1} \mid X_{kn} = y_n\right)$$

Therefore,

$$\mathbb{P}(Y_{n+1} = y_{n+1} \mid Y_n = y_n, Y_{n-1} = y_{n-1}, \dots, Y_0 = y_0) = \mathbb{P}(Y_{n+1} = y_{n+1} \mid Y_n = y_n)$$

Thus, $\{Y_n\}$ satisfies the Markov property and is a Markov chain.

Solving for the transition matrix, note that for any state $i, j \in S$,

$$\mathbb{P}(Y_{n+1} = j \mid Y_n = i) = \mathbb{P}(X_{k(n+1)} = j \mid X_{kn} = i)$$

In the Markov chain $\{X_n\}$,

$$\mathbb{P}(X_{k(n+1)} = j \mid X_{kn} = i) = (P^k)_{ij}$$

Thus, the one-step transition probability for $\{Y_n\}$ is

$$\mathbb{P}(Y_{n+1} = j \mid Y_n = i) = (P^k)_{ij}$$

Therefore, the transition matrix for $\{Y_n\}$ is P^k .

(b) (10 points) Suppose that the original chain $\{X_n\}$ is irreducible. Is $\{Y_n\}$ irreducible? If so, prove it; if not, provide a counterexample.

Consider a Markov chain, with state space $S = \{X_0, X_1\}$. Define its transition matrix as

$$P_X = \left(\begin{array}{cc} 0 & 1\\ 1 & 0 \end{array}\right)$$

Since X_0 and X_1 communicate, this chain is trivially irreducible.

Suppose k = 2. The transition matrix of $Y_m = X_{2n}$ is

$$P_Y = P_X^2 = \left(\begin{array}{cc} 1 & 0 \\ 0 & 1 \end{array}\right)$$

 X_0 and X_1 do not communicate in the new chain, as X_0 and X_1 are closed subsets on S. Therefore, Y_n is not irreducible.

Note that this can be generalized to any Markov chain $X_0, X_1, ..., X_k$, where $\mathbb{P}_X(i,j) = 1$ for $j = (i+1) \mod k$ and $\mathbb{P}_X(i,j) = 0$ for all other i, for all j. This chain is trivially irreducible, and yields a transition matrix P^k which is a $k \times k$ identity matrix. The chain Y_n is not irreducible, as the states $X_0, X_1, ..., X_k$ are closed subsets on S.

(c) (10 points) Suppose that the original chain $\{X_n\}$ is aperiodic. Is $\{Y_n\}$ aperiodic? If so, prove it; if not, provide a counterexample.

Recall that the period of a state i is defined as

$$d(i) = \gcd\left\{n \ge 1 : (P^n)_{ii} > 0\right\}$$

If this gcd equals 1 for every i, then the chain is a periodic.

In the chain $\{Y_n\}$ defined by $Y_n = X_{kn}$, the transition probabilites are given by

$$\mathbb{P}(Y_{m+1} = j \mid Y_m = i) = \mathbb{P}(X_{k(m+1)} = j \mid X_{km} = i) = (P^k)_{ij}$$

Hence the transition matrix of $\{Y_n\}$ is exactly P^k , and the *m*-step transition probabilities in $\{Y_n\}$ are given by $(P^k)^m = P^{km}$.

Therefore, the period of state i as a state of the chain $\{Y_n\}$ is

$$d_Y(i) \; = \; \gcd \Bigl\{ \, m \geq 1 : \bigl(P^{km}\bigr)_{ii} > 0 \Bigr\}$$

A standard characterization of aperiodicity is that for each state i, there exists some integer N st

$$(P^n)_{ii} > 0 \quad \forall n > N$$

In other words, $\{n: (P^n)_{ii} > 0\}$ is co-finite, or contains all sufficiently large n.

If $(P^n)_{ii} > 0$ for all $n \ge N$, then in particular $(P^{km})_{ii} > 0$ whenever $km \ge N$. Hence for all integers $m \ge \lceil N/k \rceil$, $(P^{km})_{ii} > 0$.

Thus, $\left\{m: \left(P^{km}\right)_{ii} > 0\right\}$ contains the tail set $\left\{\lceil N/k \rceil, \lceil N/k \rceil + 1, \lceil N/k \rceil + 2, \dots\right\}$ of all sufficiently large integers m.

For the chain $\{Y_n\}$, the period of i is

$$d_Y(i) \ = \ \gcd\Bigl\{ \ m \ge 1 : \left(P^{km}\right)_{ii} > 0 \Bigr\}$$

As this set includes all sufficiently large integers m, the gcd of any infinite set must be 1.

Since $d_Y(i) = 1$ for every state i, $\{Y_n\}$ is aperiodic.

(d) (10 points) Suppose that the original chain $\{X_n\}$ is transient. Is $\{Y_n\}$ transient? If so, prove it; if not, provide a counterexample.

Recall that a Markov chain $\{X_n\}$ on S is transient if there exists at least one state $i \in S$ st

 \mathbb{P}_i (the chain ever returns to i) < 1

Equivalently,

$$\mathbb{P}_i(\exists n \ge 1 : X_n = i) < 1$$

Consider state $i \in S$ and suppose we start from $X_0 = i$.

By transience of $\{X_n\}$,

$$\mathbb{P}_i(\exists m \ge 1 : X_m = i) < 1$$

Define the event

$$A = \{\exists m \ge 1 : X_m = i\}$$

and the event

$$B = \{\exists n \ge 1 : Y_n = i\} = \{\exists n \ge 1 : X_{kn} = i\}$$

Note that if $X_{kn} = i$ for some $n, X_m = i$ for m = kn, i.e. $B \subseteq A$. Hence,

$$\mathbb{P}_i(B) < \mathbb{P}_i(A)$$

Since A has probability strictly less than 1,

$$\mathbb{P}_i(\exists n \ge 1 : Y_n = i) = \mathbb{P}_i(B) < 1$$

Therefore, i is transient for the chain $\{Y_n\}$ as well. Since this applies to every state $i \in S$ that is transient in $\{X_n\}$, it shows that no state i can become recurrent under the k-step sampling. The probability that $\{Y_n\}$ returns to i is bounded above by the probability that $\{X_n\}$ returns to i, and the latter is less than 1 for transient states.

(e) (15 points) Suppose that the original chain $\{X_n\}$ is recurrent. Is $\{Y_n\}$ recurrent? If so, prove it; if not, provide a counterexample.

Let $\{X_n\}_{n\geq 0}$ be a Markov chain on a state space \mathcal{S} with transition matrix P, and assume that $\{X_n\}$ is recurrent. Fix state $i\in\mathcal{S}$ that is recurrent, st

$$\mathbb{P}_i\Big(\{n\geq 1: X_n=i\}=\infty\Big)=1$$

For $k \geq 1$, define the chain $\{Y_n\}_{n\geq 0}$ by

$$Y_n = X_{kn}, \quad n = 0, 1, 2, \dots$$

Since i is recurrent for $\{X_n\}$, define the set A as

$$A = \{n \ge 1 : X_n = i\} = \infty$$

Now, note that every positive integer belongs to one of the k residue classes modulo k, or

$$\mathbb{N} = \bigcup_{r=0}^{k-1} \{ n \in \mathbb{N} : n \equiv r \pmod{k} \}$$

Thus, the return times can be partitioned as

$$A = \bigcup_{r=0}^{k-1} A_r$$
, where $A_r = \{n \in A : n \equiv r \pmod{k}\}$

Because A is countably infinite, by the pigeonhole principle at least one of the sets A_r must be coutably infinite. As such, there exists an $r_0 \in \{0, 1, \dots, k-1\}$ st

$$\mathbb{P}_i\Big(A_{r_0} = \infty\Big) = 1$$

Consider the two cases,

Case 1: If $r_0 = 0$, then infinitely many returns to i occur at times that are multiples of k. In other words,

$$A_0 = \{ n \ge 1 : X_n = i \text{ and } n \equiv 0 \pmod{k} \} = \infty$$

Hence, i is visited infinitely often by the chain $\{Y_n\}$ and is therefore recurrent.

Case 2: If $r_0 \neq 0$, define a time-shifted chain

$$\widetilde{Y}_n = X_{kn+r_0}$$

Since \widetilde{Y}_n is merely a shifted version of the k-chain, it is also a Markov chain, with the same state space and transition matrix P^k , and has the same recurrence properties, as recurrence is invariant under a finite time shift. Since A_{r_0} is infinite,

$$\{n \ge 0 : \widetilde{Y}_n = i\} = \{n \ge 0 : X_{kn+r_0} = i\} = \infty$$

Therefore, i is recurrent for the chain $\{\widetilde{Y}_n\}$. Since \widetilde{Y}_n and Y_n differ only by a fixed time shift, it follows that recurrence holds for both chains.

Hence, i is recurrent for $\{Y_n\}$ as well. In either case, if i is recurrent for $\{X_n\}$, then it is also recurrent for $\{Y_n\}$.

Since recurrence is a class property, the entire chain $\{Y_n\}$ must be recurrent.

(f) (5 points) Suppose that the original chain X_n is irreducible and that it has period d. What is the period of each state i in the new Markov chain Y_n for k = d?

Since the original chain is irreducible with period d, for each state i,

$$(P_X^d)_{ii} > 0$$
 and $(P_X^{d-b})_{ii} = 0$, $\forall b = 1, 2, \dots, d-1$

Therefore, for any multiple of d, $a \in \mathbb{Z}$ such that $a \geq 1$,

$$(P_{\mathbf{Y}}^d)_{ii}^a > 0$$

As such, all returns to i can occur only at multiples of d steps.

Define the transition matrix for Y_n where k = d as

$$\mathbb{P}(Y_{m+1} = j \mid Y_m = i) = (P_X^k)_{ij} = (P_X^d)_{ij}$$

When j = i,

$$(P_{\mathbf{v}}^1)_{ii} = (P_{\mathbf{v}}^d)_{ii} > 0$$

As such, state i can reach itself in a single step in chain Y, forming a self-loop in Y.

Defining the period of i in Y,

$$d_Y(i) = \gcd\{m \ge 1 : (P^{dm})_{ii} > 0\}$$

Note that $(P^d)_{ii} > 0$, so at m = 1, $(P^{d \cdot 1})_{ii} = (P^d)_{ii} > 0$. Since the chain is irreducible, raising P^d to a higher power trivially cannot change its positivity:

$$(P^{d\,2})_{ii} = (P^d)_{ii}^2 > 0, \quad (P^{d\,3})_{ii} > 0, \quad \dots$$

 $(P^{dn})_{ii} > 0$ holds for all $n \ge 1$, so the set of possible m is $\{1, 2, 3, \dots\}$. Thus,

$$\{m: (P^{dm})_{ii} > 0\} = \{1, 2, 3, \dots\}$$

with a trivial gcd of 1. Note that since the d period was a state property in the original chain, $\{X_n\}$, it holds for all states in the old chain, and thus all states in the new chain.

As such, the period of each state in the new chain $\{Y_n\}$ is 1. In other words, the chain $\{Y_n\}$ is aperiodic.

Problem 3

(Asymmetric random walk, 15 points) Consider the asymmetric random walk on \mathbb{Z} , that is, $X_n = X_{n-1} + Z_n$, where Z_1, Z_2, \ldots are iid and $\mathbb{P}(Z_n = +1) = p$ and $\mathbb{P}(Z_n = -1) = 1 - p$, with $p \in [0, 1]$ and $p \neq \frac{1}{2}$. Show that the state 0 is a transient state.

In Lecture 7 we saw/will see that when $p = \frac{1}{2}$ this is not true anymore and the state 0 is recurrent. Can you explain intuitively why this is the case?

Hint: You may want to use Stirling's formula that $\lim_{n\to\infty} \frac{n!}{(n/e)^n \sqrt{2\pi n}} = 1$.

Starting from $X_0 = 0$, the random walk is at state 0 again at t = n only when it has taken an equal number of +1 steps as -1 steps. As such, n must be even.

Suppose n = 2k, and k is the number of Z_i that are +1,

$$\mathbb{P}(X_{2k} = 0 \mid X_0 = 0) = \binom{2k}{k} p^k (1-p)^k$$

Note that $\mathbb{P}(X_n = 0 \mid X_0 = 0) = 0$ if n is odd.

So the series of return probabilities at 0 is

$$\sum_{n=0}^{\infty} \mathbb{P}(X_n = 0 \mid X_0 = 0) = 1 + \sum_{k=1}^{\infty} {2k \choose k} p^k (1-p)^k$$

accounting for the initial state of 0. Using Stirling's approximation,

$$n! \sim \left(\frac{n}{e}\right)^n \sqrt{2\pi n}$$
 as $n \to \infty$

applying to this case,

$$\binom{2k}{k} = \frac{(2k)!}{k! \, k!} \approx \frac{\sqrt{4\pi k} \left(\frac{2k}{e}\right)^{2k}}{2\pi k \left(\frac{k}{e}\right)^k \left(\frac{k}{e}\right)^k} = \frac{4^k}{\sqrt{\pi k}}$$

Therefore,

$$\binom{2k}{k} p^k (1-p)^k \approx \frac{4^k}{\sqrt{\pi k}} [p(1-p)]^k = \frac{[4p(1-p)]^k}{\sqrt{\pi k}}$$

If $p \neq \frac{1}{2}$, then 4p(1-p) < 1 (If f(x) = x(1-x), then f'(x) = -x + 1 - x = -2x + 1. Solving for the max when $f'(x) = 0, x = \frac{1}{2}$).

Note, that as $k \to \infty$, $\left[4 \, p (1-p) \right]^k$ decays exponentially. Therefore,

$$\binom{2k}{k} p^k (1-p)^k = O\left(\left[4 p(1-p)\right]^k\right) \quad \text{and} \quad \sum_{k=1}^{\infty} \binom{2k}{k} p^k (1-p)^k < \infty$$

Thus,

$$\sum_{n=0}^{\infty} \mathbb{P}(X_n = 0 \mid X_0 = 0) = 1 + \sum_{k=1}^{\infty} {2k \choose k} p^k (1-p)^k < \infty$$

which defines a transient state.

However, when $p = \frac{1}{2}$,

$$\binom{2k}{k} \left(\frac{1}{2}\right)^k \left(\frac{1}{2}\right)^k \ \approx \ \frac{\left[4 \cdot 0.5(1 - 0.5)\right]^k}{\sqrt{\pi k}} \ = \ \frac{1}{\sqrt{\pi k}}$$

 \mathbf{SO}

$$\sum_{k=1}^{\infty} \frac{1}{\sqrt{k}} = \infty$$

which defines a recurrent state when $p = \frac{1}{2}$.

Exercise 1.8

Consider a Markov chain on the integers with

$$P(i, i+1) = 0.4$$
 and $P(i, i-1) = 0.6$ for $i > 0$,
$$P(i, i+1) = 0.6$$
 and $P(i, i-1) = 0.4$ for $i < 0$,
$$P(0, 1) = P(0, -1) = \frac{1}{2}.$$

This is a chain with infinitely many states, but it has a sort of probabilistic "restoring force" that always pushes back toward 0. Find the stationary distribution.

Denote the stationary distribution by $\{\pi_i\}_{i\in\mathbb{Z}}$

$$\sum_{j \in \mathbb{Z}} \pi_j P(j, i) = \pi_i \quad \forall i \in \mathbb{Z}$$

and $\sum_{i\in\mathbb{Z}} \pi_i = 1$. As this is a two-sided birth-death chain,

$$\pi_i P(i, i+1) = \pi_{i+1} P(i+1, i)$$

For $i \geq 1$,

$$\pi_i \times 0.4 = \pi_{i+1} \times 0.6 \implies \frac{\pi_{i+1}}{\pi_i} = \frac{0.4}{0.6} = \frac{2}{3}$$

For i < -1,

$$\pi_{i-1} \times 0.6 = \pi_i \times 0.4 \implies \frac{\pi_{i-1}}{\pi_i} = \frac{0.4}{0.6} = \frac{2}{3}$$

When i = 0,

$$\pi_0 \times 0.5 = \pi_1 \times 0.6 \implies \frac{\pi_1}{\pi_0} = \frac{0.5}{0.6} = \frac{5}{6}$$

$$\pi_{-1} \times 0.6 = \pi_0 \times 0.5 \implies \frac{\pi_{-1}}{\pi_0} = \frac{0.5}{0.6} = \frac{5}{6}$$

Hence,

$$\pi_1 = \frac{5}{6} \, \pi_0, \quad \pi_{-1} = \frac{5}{6} \, \pi_0$$

As the probabilities of a jump remain the same, generalizing for all $i \ge 1$,

$$\pi_{i+1} = \frac{2}{3}\pi_i \implies \pi_i = \left(\frac{2}{3}\right)^{i-1} \pi_1 \quad \text{for } i \ge 1$$

$$\pi_i = \left(\frac{2}{3}\right)^{i-1} \cdot \frac{5}{6}\pi_0 \quad \forall i \ge 1$$

Generalizing for all $i \leq -1$,

$$\pi_{i-1} = \frac{2}{3}\pi_i \quad \Longrightarrow \quad \pi_i = \frac{2}{3}^{-i-1}\pi_{-1} \quad \forall i \le -1$$

$$\pi_i = \left(\frac{2}{3}\right)^{-i-1} \cdot \frac{5}{6}\pi_0 \quad \forall i \le -1$$

Combining the two cases,

$$\pi_i = \frac{5}{6} \left(\frac{2}{3}\right)^{|i|-1} \pi_0, \quad \forall i \neq 0$$

Solving for π_0 , first recall that,

$$\sum_{i=-\infty}^{\infty} \pi_i = 1$$

Hence,

$$\pi_0 + \sum_{i \neq 0} \frac{5}{6} \left(\frac{2}{3}\right)^{|i|-1} \pi_0 = 1$$

$$\pi_0 \left[1 + \frac{5}{6} \sum_{i \neq 0} \left(\frac{2}{3} \right)^{|i|-1} \right] = 1$$

Exploiting the symmetry of the chain,

$$\sum_{i \neq 0} \left(\frac{2}{3}\right)^{|i|-1} = 2\sum_{j=1}^{\infty} \left(\frac{2}{3}\right)^{j-1} = 2\sum_{k=0}^{\infty} \left(\frac{2}{3}\right)^{k} = 2 \cdot \frac{1}{1 - \frac{2}{3}} = 2 \cdot 3 = 6$$

Thus

$$\pi_0 \Big[1 + \frac{5}{6} \cdot 6 \Big] = 6 \pi_0 = 1 \implies \pi_0 = \frac{1}{6}$$

Therefore,

$$\pi_i = \begin{cases} \frac{1}{6}, & i = 0\\ \frac{5}{36} \left(\frac{2}{3}\right)^{|i|-1} & i \neq 0 \end{cases}$$

Exercise 1.16

Show that if an irreducible Markov chain has a state i such that P(i,i) > 0, then the chain is aperiodic. Also show by example that this sufficient condition is not necessary.

Let $\{X_n\}$ be an irreducible Markov chain on a countable state space S. Suppose there is a state $i \in S$ such that

The period of state i is defined as

$$d(i) = \gcd\{n > 1 : P^n(i, i) > 0\}$$

Since $P(i,i) = P^1(i,i) > 0$, there is a positive probability of returning to i in exactly 1 step. As such, $1 \in \{n : P^n(i,i) > 0\}$, so any common divisor of all n must trivially divide 1.

Therefore,

$$d(i) = \gcd\{n \ge 1 : P^n(i,i) > 0\} = 1$$

Thus, i is an aperiodic state.

By irreducibility, for any other state $j \in S$, $\exists m, k \in \mathbb{Z}$ such that $m, k \geq 1$, $P^m(j, i) > 0$, and $P^k(i, j) > 0$. Therefore as proven in lecture, for any $n \geq 1$,

$$P^{m+n+k}(j,j) > P^m(j,i) P^n(i,i) P^k(i,j)$$

Since $P^n(i,i) > 0$ for all $n \ge 1$, for arbitrarily many n, the probability $P^{m+n+k}(j,j)$ is strictly positive. As such,

$$d(j) = \gcd\{n \ge 1 : P^n(j,j) > 0\} = 1$$

Thus, every state in an irreducible chain with a self-loop, i.e. P(i,i) > 0, is aperiodic.

To show that this sufficient condition is not necessary, consider a chain on three states $\{1,2,3\}$ with transition matrix:

$$P(1,1) = 0,$$
 $P(1,2) = 1,$ $P(1,3) = 0$
 $P(2,1) = \frac{1}{2},$ $P(2,2) = 0,$ $P(2,3) = \frac{1}{2}$
 $P(3,1) = \frac{1}{2},$ $P(3,2) = \frac{1}{2},$ $P(3,3) = 0$

Note that the chain is irreducible, but does not contain any self-loops, i.e. P(i,i) = 0 for i = 1,2,3.

$$P^{2}(1,1) = P(1,2)P(2,1) = (1)\left(\frac{1}{2}\right) = \frac{1}{2} > 0$$

$$P^{3}(1,1) = P(1,2)P(2,3)P(3,1) = 1 \cdot \frac{1}{2} \cdot \frac{1}{2} = \frac{1}{4} > 0$$

Therefore, state 1 can return to itself in both 2 steps and 3 steps, which comprises a subset of the set of all possible return times back to state 1. The gcd of 2 and 3 is 1, or a period of 1. Note that $P^2(2,2) > 0$, $P^2(3,3) > 0$, and $P^3(2,2) > 0$, $P^3(3,3) > 0$, so every state also has period 1, making the entire chain aperiodic. Thus, P(i,i) > 0 for some i is not a necessary condition for a chain to be aperiodic.