

SEMANTIC/PRAGMATIC ADAPTATION

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DOCTOR OF PHILOSOPHY

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I certify that I have read this dissertation and that, in my opinion, it is fully adequate in scope and quality as a dissertation for the degree of Doctor of Philosophy.

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Abstract

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Background

2.1 The semantics of uncertainty expressions

While this is not a dissertation about modality, many of the utterances that I consider in my experiments contain an epistemic modal. Since I discuss the implications of my experimental results for popular theories of the semantic of epistemic modals in Section XXX, and since my computational model is inspired by recent accounts of epistemic modality, I provide a brief introduction to several theories of modality in this section.

Here, and throughout this dissertation, I adapt the broad notion of modality by [?] and [?], which not only includes modal auxiliaries (e.g., *might*, *could*) but also other evidential devices such as probability operators (e.g., *probably*) and attitude verbs (e.g., *think*). At the same time, however, to limit the scope of this discussion, I will only cover epistemic modality, and therefore omit any discussion of deontic modals, i.e., modals to express how the world should be according to laws, societal norms, etc., which are frequently discussed together with epistemic modals. I will also omit discussions of the important connections between modals and conditionals [?, ?, ?, ?, ?] and discussions of the extent to which different semantic theories validate desired and undesired logical inferences [?, ?].

2.1.1 Background: Possible world semantics

Classical modal logic and most other semantic theories of modals are based on the concept of possible worlds [?]. A possible world is a world which differs in one or multiple properties from the actual world. For example, while the proposition ϕ_{brown} expressed by the sentence “I have brown hair” is true in the actual world w , one possible world w_1 is identical in every regard to the actual world except that the proposition ϕ_{blond} encoded by “I have blond hair” is true and the one encoded by “I have brown hair” (ϕ_{brown}) is false.

According to a possible world semantics, all sentences have to be evaluated relative to a possible world w and propositions ϕ can be represented as a set of worlds in which ϕ is true. If we consider the worlds w and w_1 as described here, the propositions expressed by the sentences “I have brown hair” and “I have blond hair” evaluate to

different truth conditions, depending on the possible world.

$$\llbracket \phi_{brown} \rrbracket^w = 1 \text{ iff } w \in \phi_{brown} = 1$$

$$\llbracket \phi_{brown} \rrbracket^{w_1} = 1 \text{ iff } w_1 \in \phi_{brown} = 0$$

$$\llbracket \phi_{blond} \rrbracket^w = 1 \text{ iff } w \in \phi_{blond} = 0$$

$$\llbracket \phi_{blond} \rrbracket^{w_1} = 1 \text{ iff } w_1 \in \phi_{blond} = 1$$

Modal logic

In classical modal logic, the truth conditions of sentences with epistemic modals depend on an accessibility relation R . R determines which worlds w' are epistemically accessible from the actual world w , i.e., which worlds are epistemically consistent with the actual world. For example, consider rolling two six-sided dice, one after another. Before you roll the first die, all worlds in which the sum of the two dice is between 2 and 12 (all possible combinations of two dice) are epistemically accessible since they are compatible with the actual world. Now, if you roll one of the dice and it comes up 4, only worlds in which the sum of the two dice is between 5 and 10 (all possible sums of 4 and a number between 1 and 6) are epistemically accessible.

Formally, if wRw' is true then w' is epistemically accessible from w . A proposition ϕ embedded under an epistemic modal is then true if either ϕ is true in all epistemically accessible worlds (for necessity modals such as *must*) or ϕ is true in at least one epistemically accessible world (for possibility modals such as *might*).

$$(1) \quad \llbracket \text{must } \phi \rrbracket^w = 1 \text{ iff } \forall w' \in W : wRw' \rightarrow \llbracket \phi \rrbracket^{w'} = 1$$

$$(2) \quad \llbracket \text{might } \phi \rrbracket^w = 1 \text{ iff } \exists w' \in W : wRw' \rightarrow \llbracket \phi \rrbracket^{w'} = 1$$

If we again use the example of rolling two dice and assume that the world w_x corresponds to the sum of the two dice being x , then wRw' is true iff $w' \in \{w_2, w_3, \dots, w_{11}, w_{12}\}$.

Therefore, for example,

$$\llbracket \text{must roll a number between 2 and 12} \rrbracket^w = 1$$

(since *roll a number between 2 and 12* is true in all epistemically accessible worlds)

$$\llbracket \text{must roll a 7} \rrbracket^w = 0$$

(since *roll a 7* is only true in some epistemically accessible worlds)

$$\llbracket \text{might roll a 7} \rrbracket^w = 1$$

(since *roll a 7* is true in the epistemically accessible world w_7)

$$\llbracket \text{might roll a 1} \rrbracket^w = 0$$

(since *roll a 1* is false in all epistemically accessible worlds).

While this approach seems intuitively correct for scenarios like rolling two dice, it is very challenging to represent utterances that convey more fine-grained meanings than mere possibility or necessity. As [?] points out, one could extend this proposal to modal expressions such as *probably* and *likely* by assuming that *probably* ϕ is true if ϕ is true in more epistemically accessible worlds than epistemically accessible worlds in which ϕ is false:

$$(3) \quad \llbracket \text{probably} \phi \rrbracket^w = 1 \\ \text{iff } |\{w' \in W \mid wRw' = 1 \wedge \llbracket \phi \rrbracket^{w'} = 1\}| > |\{w' \in W \mid wRw' = 1 \wedge \llbracket \phi \rrbracket^{w'} = 0\}|$$

However, this proposal comes with at least two shortcomings if one wants to consider it as a complete theory of epistemic modals. First, one has to make the limit assumption [?], i.e., one has to assume that W contains a finite number of possible worlds. Second, this proposal does not provide a theory of interpretation for any type of graded epistemic modal expressions such as *It is 60% likely that...* or *It is highly probable that...* or modal expressions in comparative constructions such as *It is twice as likely that X than Y* [?]. Third, there is no connection between event probabilities and the

use of different modals except that this account would predict that *might* ϕ is true when the probability of ϕ is greater than 0, and *must* ϕ is true if the probability of ϕ is 1.

Double relativity of modals

The most prominent semantic theory of epistemic modals (and all other flavors of modals) is the account by Kratzer [?, ?], later revised in [?]. Building on [?], she developed a unifying account of all modal flavors, which assumes that there are several core meanings of modals that can be expressed by various linguistic devices (e.g., *could* and *might* are both possibility modals), and that the interpretation of a sentence with a modal depends on two *conversational backgrounds*, that is, contextually specified functions from possible worlds to sets of propositions: the *modal base* $f(w)$ and the *ordering source* $g(w)$.

The intuition behind the modal base $f(w)$ is that one can explicitly state which worlds the modal quantifies over using an “In the view of ...” adverbial clause. For example, for epistemic modals, the modal base may be a function that returns the set of propositions that are compatible with what is known in the current world, which can be explicitly expressed in a sentence with an epistemic modal through the adverbial clause “In the view of what is known”:

- (4) In the view of what is known, it could rain tomorrow.

Kratzer argues that utterances without explicit mention of a conversational background are interpreted by contextually resolving the relevant modal base. If we ignore the ordering source for a second, this leads to the following definitions of f -necessity and f -possibility.

f-necessity:

ϕ is a necessity with respect to a modal base f iff $\phi \subseteq \cap f(w)$.

f-possibility:

ϕ is a possibility with respect to a modal base f iff $\cap \{\{\phi\} \cup f(w)\} \neq \emptyset$.

The semantics of sentences with necessity modals such as *must* and possibility modals such as *might* can then be expressed in terms of f -necessity and f -possibility:

- (5) $\llbracket \text{must } \phi \rrbracket^{w,f} = 1$ iff ϕ is an f -necessity in w
- (6) $\llbracket \text{might } \phi \rrbracket^{w,f} = 1$ iff ϕ is an f -possibility in w

The semantics of these expressions is equivalent to the classical modal logic semantics presented in (??) and (??), since the accessibility relation R can be defined as $wRw' = w' \in \cap f(w)$. For this reason, this account suffers from the same issues as the classical modal logic account: it cannot be used to derive interpretations for graded modals or comparatives.

Kratzer partially resolves these issues by introducing a second conversational background, the ordering source $g(w)$. $g(w)$ defines a partial preorder $\leq_{g(w)}$ over the set of possible worlds W such that

$$u \leq_{g(w)} v \text{ iff } \{p \in g(w) \mid u \in p\} \subseteq \{p \in g(w) \mid v \in p\}.$$

That means, v is at least as close to an ideal as u iff all propositions in the set of ideal propositions $g(w)$ that are true in u are also true in v . In the case of epistemic modals, the ideal as defined by $g(w)$, is usually assumed to contain propositions corresponding to a normal course of events.

Necessity and possibility then be defined as follows.

Necessity:

ϕ is a necessity with respect to a modal base f and an ordering source g iff for all $u \in \cap f(w)$, there is a $v \in \cap f(w)$ such that $u \geq_{g(w)} v$ and for all $z \in \cap f(w)$: if $v \geq_{g(w)} z$, then $z \in \phi$.

Possibility:

ϕ is a possibility with respect to a modal base f and an ordering source g iff $\neg\phi$ is not a necessity with respect to f and g .

Intuitively, the definition of necessity can be seen as further restricting the modal source such that something must be true if it is true in all epistemically accessible worlds that come closest to the ideal defined by the ordering source.

To illustrate how the modal base and the ordering source work together, imagine a murder case in a small town in which there are four suspects A, B, C, and D who all have a motive.¹ Further, there was a tourist T from Iceland in town when the murder happened. Since random tourists rarely murder somebody without a motive, the set of normal propositions $g(w)$ in this example could be $\{a, b, c, d\}$, where each proposition corresponds to A, B, C, and D committing the murder, respectively. However, the conversational background of what is known $f(w)$ is compatible with A, B, C, D, or T being the murderer, i.e., $\cap f(w) = \cup\{a, b, c, d, t\}$. Now, it seems natural for a police officer to utter (??) or (??) but unlikely for the officer to utter (??).

- (7) A might have committed the murder.
- (8) A or B or C or D must have committed the murder.
- (9) T might have committed the murder.

Kratzer's account makes exactly these predictions. It predicts that (??) is true because $\neg a$ is not a necessity and therefore a is a possibility; it predicts that (??) since $a \vee b \vee c \vee d$ is a necessity; and it predicts that (??) is false since $\neg t = a \vee b \vee c \vee d$ is a necessity and therefore t is not a possibility.

Apart from introducing this distinction between normal courses of events and theoretically possible courses of events, Kratzer's account also provides a semantics for comparatives using the notion of comparative possibility.²

Comparative possibility:

ϕ is at least as good a possibility as ψ in w with respect to a modal base f and an ordering source g iff

$$\neg \exists u (u \in \cap f(w) \wedge u \in \phi - \psi \wedge \forall v ((v \in \cap f(w) \wedge v \in \psi - \phi) \rightarrow v <_{g(w)} u))$$

Further, ϕ is a better possibility than ψ iff ϕ is at least as good a possibility as ψ and ψ is not at least as good as possibility as ϕ . Using this definition, [?] defines the semantics of *probably* as

¹Example adapted from [?].

²This is the revised definition of comparative possibility from in [?] which is slightly different from the original notion of comparative possibility in [?].

(10) $\llbracket \text{probably } \phi \rrbracket^{w,f,g} = 1$ iff ϕ is a better possibility than $\neg\phi$.

Similarly, the semantics of ϕ *is more likely than* ψ can be defined as

(11) $\llbracket \phi \text{ is more likely than } \psi \rrbracket^{w,f,g} = 1$ iff ϕ is a better possibility than ψ .

As compared to classical modal logic, this proposal has the advantage of providing a semantics for comparative constructions and a semantics for *probably*. However, this account still does not provide a compositional account for modal expressions and therefore does not provide a semantics for expressions such as *very likely*. Second, this account also does not make predictions about the use of epistemic modals to communicate and infer event probabilities. [?] briefly discusses event probabilities and shows that one can come up with probability measures on the set of sets of possible worlds $\mathcal{P}(W)$ such that $\phi \geq_{g(w)} \psi$ implies $P(\phi) \geq P(\psi)$. However, her discussion does not go beyond showing that a connection between event probabilities and possible worlds is possible and her semantic account of modals leaves it open how speakers and listeners map modals to event probabilities.

Threshold semantics: Lassiter, Swanson,

* epistemic modals

* probability operators

* uncertainty expressions in operation research

Chapter 3

Production expectations

3.1 Experiment 1: Pre-exposure ratings

We first conducted a norming study, which served the following theoretical and methodological purposes. First, it served as a methodological check on whether the paradigm is suited for manipulating fine-grained event probabilities. Second, it addressed the theoretical question of whether listeners vary in their expectations about a generic speaker’s use of uncertainty expressions, by collecting participants’ judgments about uncertainty expressions they expected speakers to use for varying probabilities of receiving gumballs of a particular color from a gumball machine. Third, the results from this study informed the experimental design of the adaptation experiments reported in later sections, by allowing us to both choose which pair of uncertainty expressions to test adaptation on, and to determine the particular event probability for which participants had roughly equi-probable expectations about which expression of uncertainty a generic speaker would use to report an event with that probability. Lastly, we used the data collected in this study to estimate population-level prior beliefs for the adaptation model reported in Section 5.

3.1.1 Method

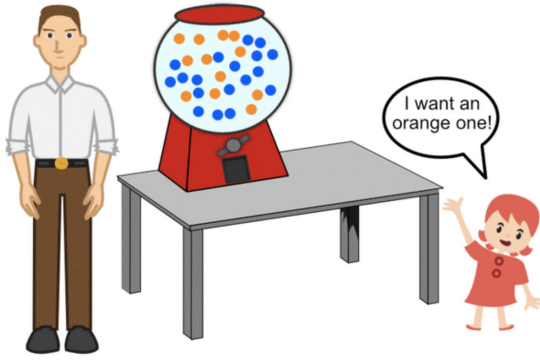
Participants

We recruited a total of 420 participants (20 per condition) on Amazon Mechanical Turk. We required participants to have a US-based IP address and a minimal approval rating of 95%. Participants were paid \$1.80 (condition 1), \$1.50 (conditions 2-15), or \$2.00 (conditions 16-21), depending on the number of trials, which amounted to an hourly wage of approximately \$12–\$15.

Materials and Procedure

This study was a production expectation experiment intended to probe listeners’ expectations about a generic speaker’s language use. Participants were instructed that over the course of the experiment, they would see several scenes with an adult man, a young girl, and a gumball machine on a table and that the gumball machine

Consider the following scene:



How likely do you think it is that the man will respond with each of the following sentences?

You might get an orange one	<input type="range" value="71"/>	71
You'll get an orange one	<input type="range" value="0"/>	0
something else	<input type="range" value="29"/>	29

Next

Figure 3.1: Example trial in Experiment 1.

is too high up on the table for the girl to see (see Figure ?? for an example scene). After completing an attention check which asked participants whether the girl could see the gumball machine,¹ participants saw a series of scenes and were asked to rate how likely they thought it was that the adult would produce two given responses by distributing 100 points across the two given utterances and the blanket *something else* option (OTHER). Sliders automatically jumped back if participants tried to distribute more than 100 points. In each scene, the child uttered “*I want a blue one*” (target color: blue) or “*I want an orange one*” (target color: orange), randomized across participants.² The gumballs in the machines were tossed around continuously to

¹Participants had to go back to the instructions in case they responded incorrectly. This was the case for 41 participants.

²In condition 1 (*bare-might*), as well as conditions 16-21 (all conditions with *bare not*), the target color was randomized across trials. While randomization of the target color across trials increased the correlation between the ratings for the two colors, the average ratings for each condition independent of the target color were not affected by this choice. See Appendix A or a detailed discussion of the

prevent participants from counting the gumballs and to make sure that participants did not consider it more likely to get one of the gumballs at the bottom of the machine. In each of the 21 conditions, participants saw only two of the following seven possible adult utterances with different uncertainty expressions:³

- You’ll get a blue/orange one. (BARE⁴)
- You might get a blue/orange one. (MIGHT)
- You’ll probably get a blue/orange one. (PROBABLY)
- I think you’ll get a blue/orange one. (THINK)
- It looks like you’ll get a blue/orange one. (LOOKS LIKE)
- You could get a blue/orange one. (COULD)
- You won’t get a blue/orange one. (BARE NOT)

Within each condition, we manipulated the percentage of target color gumballs across trials, which we take as proxy for the objective probability of receiving a gumball of the target color. Each participant saw 3 trials⁵ for each of the following percentages: 0%, 10%, 25%, 40%, 50%, 60%, 75%, 90%, 100%. We randomized the order of expressions across participants and trials were presented in randomized order.

effect of this manipulation on the ratings.

³In the choice of the investigated expressions, we follow recent work on the interpretation of uncertainty expressions [?] and aim to use naturalistic utterances. We therefore decided against using a frame such as “*It is UNCERTAINTY-ADJ that*” that would have fixed the syntactic structure across items because such a frame would have resulted in less naturalistic expressions like “*It is possible that*” or “*It is probable that*”, which are much less common than the expressions we considered (e.g., *might* appears more than 30 times more often than *possible that* in the spoken portion of the Contemporary American English [?]). A speaker’s use of these rare expressions thus could have triggered additional pragmatic inferences due to violations of the maxim of manner [?] that are unrelated to our research questions.

⁴As a notational convention, we refer to utterances with uncertainty expressions in SMALL CAPS and to the uncertainty expression itself in *italics*.

⁵In condition 1 (*bare-might*), participants saw each gumball machine 6 times: 3 times when being asked to produce a statement about orange gumballs and 3 times when being asked to produce a statement about blue gumballs. In conditions 15-20 (all conditions with *bare not*), participants saw each machine 4 times: 2 times for each color.

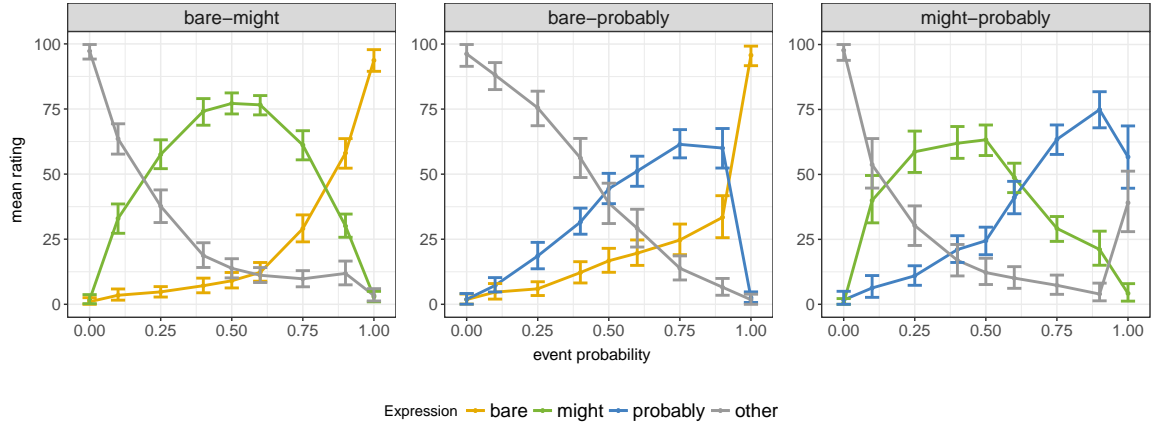


Figure 3.2: Results from 3 conditions of Experiment 1. Error bars correspond to bootstrapped 95%-confidence intervals.

3.1.2 Results and Discussion

Figure ?? shows participants' ratings for different gumball proportions for 3 of the 21 conditions, namely all combinations of the conditions with the utterances BARE, PROBABLY, and MIGHT (see Appendix B for the results from the other 18 conditions). The results from these three conditions highlight several important properties of participants' behavior in this experiment that generalize to all conditions. First, the ratings for individual utterances are influenced by the utterance choices presented to participants. If we compare the ratings for MIGHT in the *bare-might* and the *might-probably* condition, we see that MIGHT received high ratings for a larger range of event probabilities when it is paired with BARE than when it is paired with PROBABLY. We observe similar effects for the other two utterances. This suggests that participants are cued towards using the utterances provided in the experiment and that their ratings depend on the presented alternatives – an effect that has also been observed for frequency expressions [?] and quantifiers [?].

Second, the results suggest that participants are sensitive to the different event probabilities and that this paradigm is well suited to study the mapping between event probabilities and uncertainty expressions. For example, in the *might-probably*

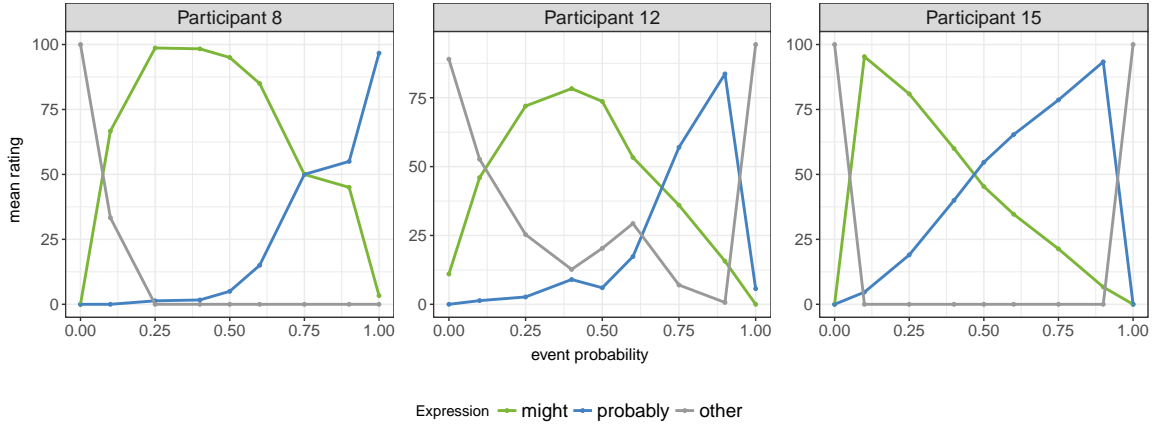


Figure 3.3: Results of three individual participants in the *might-probably* condition of the Experiment 1.

condition, participants provided considerably different ratings when they were presented with a gumball machine with 50% target color gumballs than when they were presented with 60% target color gumballs.

Third, in all conditions, the mean ratings are graded and except for the 0% and 100% target color gumball trials, the average rating for none of the utterances is close to 100. There are two potential explanations for this observation. It could be that participants provided categorical ratings, i.e., generally assigned 100 points to one of the three options but the category boundaries vary across participants which leads to the graded average ratings. It could also be that participants' individual ratings are graded which could reflect participants' uncertainty about which utterance a speaker would use and that these individual graded ratings drive the graded average ratings. If we look at individual participants' ratings, it appears to be a combination of both. Figure ?? shows the responses of three individual participants in the *might-probably* condition. These figures show that there is a range of gumball proportions for each participant for which they assigned similar ratings to two utterances, which suggests uncertainty about the speaker's utterance choice. At the same time, however, this range also differed across participants: Participant #8, who considered the experimental speaker a “*cautious*” speaker, thought that the speaker would only be likely

to use PROBABLY when the objective probability of getting a target color gumball was greater than 0.75, whereas participant #15, who considered the experimental speaker a “*confident*” speaker, thought that PROBABLY was a better utterance choice than MIGHT when the objective probability of getting a target color gumball was just greater than 0.5. These observations suggest that for some event probabilities, participants have uncertainty about a speaker’s choice of uncertainty expression and that participants have a priori different expectations about how a generic speaker would use these expressions.

This uncertainty and variability seems to be particularly borne out in the *might-probably* condition. For this reason, we chose this pair of expressions to study listeners’ adaptation to variable uses of uncertainty expressions.

3.2 Modeling expectations about uncertainty expression productions

In this section we propose a computational model of expectations about uncertainty expression production that is informed by the data from Experiment 1. This model will serve as proxy for listeners’ baseline generative model of a generic speaker and will be used as the basis for investigating adaptation processes in Section ???. Experiment 1 confirmed previous findings that participants’ expectations about how a generic speaker would use uncertainty expressions depend on the set of utterances that participants can choose from. We further found that ratings were graded in part because participants had uncertainty about how a generic speaker would use uncertainty expressions. Hence, a model that accounts for participants’ beliefs about a speaker’s production of uncertainty expressions should (a) be able to capture differences in ratings depending on the availability of alternative utterances; (b) provide graded predictions about utterance probabilities; and (c) be able to capture within-participant uncertainty about probability of use.

Computational game-theoretic models such as the Rational Speech Act framework (RSA; [?]) are uniquely suited to satisfy the above desiderata. RSA models are

a probabilistic formalization of Gricean pragmatics which model comprehension as Bayesian probabilistic inference. They consist of *listener* and *speaker* agents which recursively reason about each other to derive interpretations and choose utterances. For our purposes of modeling production expectations, we focus on a model of listener beliefs about a speaker’s production model. According to RSA, a speaker that wants to convey some information to a listener chooses an utterance based on the utterance’s utility compared to the utility of alternative utterances. The speaker’s utterance utility is determined by trading off the informativity of the utterance to a *literal listener* on the one hand and the cost of the utterance on the other.

In defining the informativity of an utterance, we follow previous RSA models of uncertainty expressions ([?]) and assume that uncertainty expressions have a threshold semantics [?, ?, ?]. That is, for each uncertainty expression e there exists some threshold $\theta_e \in [0, 1]$ such that an utterance u_e with e is true if the probability ϕ of the proposition embedded under e exceeds θ_e . For example, if we assume the threshold for *might*, θ_{might} , is 0.1, then the statement “It might rain this afternoon” is semantically available to a speaker who believes that the probability of rain in the afternoon exceeds 0.1. Formally, we base the computation of informativity on a probability distribution from utterances to event probabilities ϕ , which is usually referred to as the *literal listener* L_0 in the RSA framework.

$$L_0(\phi \mid u_e, \theta_e) \propto P(\phi) \times \begin{cases} 1 & \text{if } \phi > \theta_e \\ 0 & \text{otherwise} \end{cases} \quad (\text{for positive embedded propositions})$$

$$L_0(\phi \mid u_e, \theta'_e) \propto P(\phi) \times \begin{cases} 1 & \text{if } \phi < \theta'_e \\ 0 & \text{otherwise} \end{cases} \quad (\text{for negated embedded propositions})$$

According to this formalization, the literal listener randomly draws an event probability ϕ from all values above the threshold θ_e associated with the uncertainty expression e (or, in the case of negated propositions, from all values below the threshold θ'_e). $P(\phi)$ is a prior distribution over event probabilities, which is independent of the

observed utterance.

A *pragmatic speaker* S_1 that intends to communicate an event probability ϕ then chooses an utterance u_e with uncertainty expression e from a set of utterances U according to a soft-max choice rule $[?, ?]$ such that u is chosen with a probability proportional to the utility for the speaker:

$$S_1(u_e | \phi, \theta, c) \propto \exp(\lambda(\log L_0(\phi | u_e, \theta_e) - c(u_e)))$$

λ is a rationality parameter which governs how likely the speaker is to choose the utterance that maximizes utility. As λ approaches infinity, the speaker is more likely to always choose the optimal utterance. As it approaches 0, the speaker produces true utterances at random.⁶

S_1 crucially depends on a vector of thresholds θ which contains a threshold for each uncertainty expression in the utterances in U , as well as a cost function $c(u)$. The values that participants expect a *pragmatic speaker* to assign to these variables are unknown a priori; we infer these values from the data collected in the above reported experiment. In Experiment 1, we found that both at the population level and at the individual level, participants' ratings of the different expressions gradually increased and decreased with changing event probabilities (as, for example, shown in Figure ??). This is expected if participants have probabilistic beliefs about thresholds θ (as illustrated in the right panels of Figure ??) but not if they reason based on point estimates of θ (as illustrated in the left panels of Figure ??). Considering these observations, we assume that listeners hold beliefs about a speaker's thresholds in the form of a distribution $P(\theta_e)$.⁷ Analogously, we assume that listeners also have beliefs

⁶ The absolute value of this parameter is influenced by the literal listener distributions and the cost function. It is therefore complex to determine a reasonable range for λ a priori, except that the value should be greater than 0 for the agent to not behave irrationally.

⁷We leave it open whether a speaker samples from a distribution over thresholds when producing utterances (as suggested by [?]) or always uses the same values for thresholds. In the former case, listeners could have higher-order beliefs $P(\eta)$ about different speakers' threshold distributions instead of having direct beliefs about the thresholds that different speakers use. For our purposes, this distinction does not matter since we assume that listeners would marginalize over higher-order beliefs $P(\eta)$ such that $P(\theta_e) = \int P(\eta) P(\theta_e | \eta) d\eta$ and we therefore take the simplest approach and directly model $P(\theta_e)$.

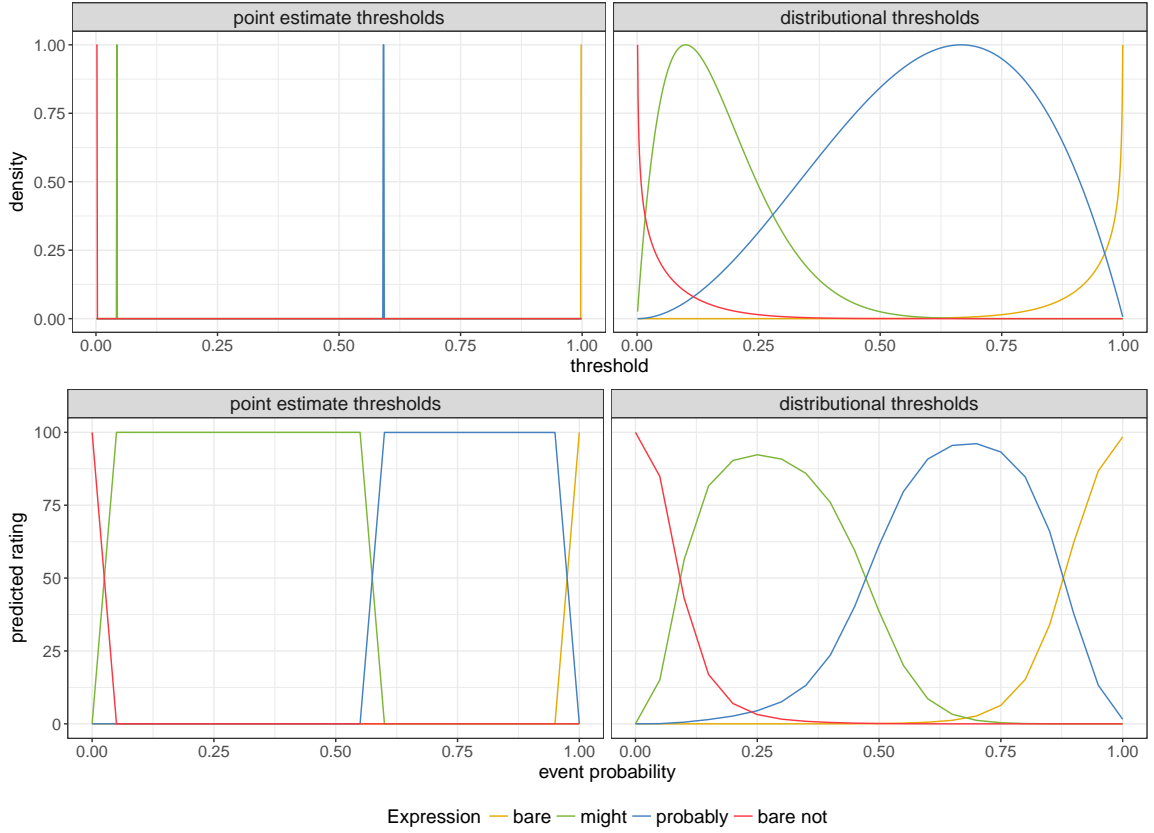


Figure 3.4: Example threshold distributions (upper panels) and corresponding model predictions for the *expected pragmatic speaker* model (lower panels). In this example, the set of possible utterances is $U = \{\text{BARE, MIGHT, PROBABLY, BARE NOT}\}$, all utterances have equal costs, the rationality parameter λ is set to 10, and the prior probability over event probabilities $P(\phi)$ is a uniform distribution. As the panels on the left show, point estimates of thresholds lead to sharp categorical boundaries in the model predictions, whereas distributions over thresholds, as in the panels on the right, lead to gradually increasing and decreasing predicted utterance ratings.

$P(c)$ about the speaker model’s cost function. Using these two distributions, we can define the *expected pragmatic speaker* model $ES_1(u_e | \phi)$ as follows:

$$ES_1(u_e | \phi) = \int P(c) \int_0^1 P(\theta) S_1(u_e | \phi, \theta, c) d\theta dc$$

This model predicts which utterance listeners with uncertainty about a speaker’s thresholds and cost function would expect that speaker to use to describe different

event probabilities. Intuitively, this model can be seen as a weighted average of different *pragmatic speaker* models, where individual models differ in terms of thresholds and cost functions. The weights of this average are determined by listeners’ belief distributions over thresholds and costs. For example, if listeners believe that it is likely that a speaker uses a threshold $\theta_{probably}$ of 0.5, they will assign higher weight to speaker models using this threshold than models using other thresholds.

3.2.1 Linking function

We assume that in Experiment 1, participants, when asked to provide ratings for utterances, reasoned about a generic speaker’s likely descriptions of varying event probabilities. We assume that this reasoning was guided by participants’ beliefs about the speaker’s thresholds and costs, and that participants averaged over their uncertainty. For this reason, we assume that the population-level average ratings of what participants expect the speaker to say in different situations correspond to the probabilities predicted by the *expected pragmatic speaker* model (with the *something else* option being predicted by the sum of the probabilities of all utterances not present in a condition; see below for more details). Further, given the forced choice nature of the experiment and that we estimate model parameters from limited and potentially noisy data, we make the following additional linking assumptions for which we provide a rationale and an assessment of their importance in turn.

- **Set of utterances:** Across all conditions, we assume that the set of utterances that participants are considering is the set of all utterances that we used in Experiment 1, i.e., $U = \{ \text{BARE, MIGHT, PROBABLY, THINK, LOOKS LIKE, COULD, BARE NOT} \}$. We include all utterances instead of only the utterances that are presented in a given condition since we assume that participants’ general knowledge of English uncertainty expressions also influences their ratings. Ideally, we would include even more utterances in this set of alternatives but since we can only estimate parameters for uncertainty expressions for which we collected ratings, we are limited to the utterances in U .

The exact set of utterances has only a small impact on the model’s predictions

as long as the set includes the BARE and BARE NOT utterances as well as at least one utterance with a weaker (e.g., *might*) and one with a stronger (e.g., *probably*) uncertainty expression. If this requirement is met, the model makes the same qualitative predictions, and closely matches the quantitative predictions of a model that includes all 7 utterances choices.

- ***something else* option:** Participants in condition $\mathcal{C} = \{u_a, u_b\}$ could only choose between the three utterances $U' = \{u_a, u_b, \text{something else}\}$. For modeling data from condition \mathcal{C} , we therefore need a function to predict the ratings for the utterances in U' . For u_a and u_b , this is straightforward: We assume the probability of a participant choosing u_a or u_b is proportional to $ES_1(u_a \mid \phi)$ and $ES_1(u_b \mid \phi)$, respectively. We model the probability of a participant choosing the *something else* option as the sum of the probability of all utterances that were not part of the condition as well as a constant O , which accounts for probability mass assigned to utterances that participants might be considering but which are not contained in U . This gives us the following condition-specific function $ES_1^{(\mathcal{C})}$ for predicting participants' ratings.

$$ES_1^{(\mathcal{C})}(u \mid \phi) \propto \begin{cases} ES_1(u \mid \phi) & \text{if } u \in \mathcal{C} \\ O + \sum_{u \notin \mathcal{C}} ES_1(u \mid \phi) & \text{if } u \text{ is } \textit{something else} \end{cases}$$

This summation over alternative utterances is crucial for fitting the data since we need to capture the ratings for *something else*. The only viable alternative would be to fit individual curves for *something else* for each condition, which would require the estimation of considerably more parameters and would not explain the ratings for the *something else* option. The inclusion of the constant O is less important but it still improves model fit.

- **Cost function:** We assume that the cost function represents participants' beliefs about the speaker's preferences for different utterances. Lower costs of an utterance indicate higher speaker preferences. We further assume that we are cueing participants to believe that the speaker would be likely to use the

two utterances, u_a and u_b , that are provided in condition $\mathcal{C} = \{u_a, u_b\}$ and that participants therefore primarily use the *something else* option when both of the two utterances are semantically infelicitous or otherwise highly unexpected. We model this cueing effect in our choice of the cost function $c(u)$, which depends on the condition. For the two utterances that are presented to the participants, we set the cost to 1 and for all the other utterances, we set the cost to a constant $\gamma > 1$:

$$c(u, \mathcal{C}) = \begin{cases} 1 & \text{if } u \in \mathcal{C} \\ \gamma & \text{otherwise} \end{cases}$$

Theoretically, we could have also used a different constant γ_u for each utterance. The data from Experiment 1, however, suggests that participants generally did not prefer one utterance over the other. To limit the number of free model parameters and to prevent overfitting, we therefore use a single constant γ for all utterances. We will, however, relax this assumption in our adaptation model in Section ??, which we use to investigate whether listeners update their beliefs about preferences during adaptation.

This condition-specific cost function is important for the model fit. If we didn't use such a cost function, the model would assign much higher ratings to the *something else* option than participants did.

- **Noise:** Finally, to account for participants not paying attention or making mistakes, we also include a noise term that models participants providing random ratings. The amount of noise is captured by the noise strength parameter δ . This parameter indicates the proportion of random responses, that is, the proportion of responses drawn from a uniform distribution over the three condition-specific responses U' .

The inclusion of the noise term is not crucial for fitting the data but it does improve model fit and is common practice in RSA models whose parameters are estimated from experimental data [?, see]Herbst2019,Tessler2019.

Incorporating all of these assumptions, we end up with the following noisy, condition-specific expected pragmatic speaker model $ES_1^{(\mathcal{C})'}(u \mid \phi)$, which we use to predict participants' ratings:

$$ES_1^{(\mathcal{C})'}(u \mid \phi) = \delta \times \frac{1}{|U'|} + (1 - \delta) \times ES_1^{(\mathcal{C})}(u \mid \phi)$$

For the prior distribution over event probabilities $P(\phi)$, which is used in the literal listener model L_0 , we use a uniform distribution over the interval $[0, 1]$.⁸ For the distributions over thresholds $P(\theta_e)$, we use a Beta distribution parametrized by α_e and β_e . The choice of Beta distributions is motivated by two of its properties. First, the support of a Beta distribution is the interval $[0, 1]$ which corresponds to the exact range of possible values for θ_e .

The second reason for using Beta distributions is that, depending on the parameterization, Beta distributions can take on very different shapes. This property is important because we are making the simplifying assumption that all utterances in our experiments have a threshold semantics. Such a semantics is commonly assumed for uncertainty expressions such as *probably* [?, e.g.,]Yalcin2010,Lassiter2016, but it is unconventional for bare assertions such as “*You’ll get a blue one*”, which are generally assumed to be true only if the event is certain to happen, i.e., it has an event probability of 1. However, since Beta distributions can have a shape like the distribution for BARE in the upper right panel in Figure ??, the model has the capability to infer a semantics for the bare form that is almost equivalent to a traditional semantics of bare assertions. In this parameterization of the Beta distribution, most probability mass is assigned to values of θ close to 1, which is mathematically almost equivalent to a traditional semantics.⁹ Therefore, using Beta distributions for the

⁸To verify the assumption that the prior on event probabilities is uniform, we conducted a separate norming study in which participants rated how likely they thought it was that a speaker described different gumball machines containing different proportions of blue and orange gumballs after hearing an unintelligible utterance. We found that on average participants rated all gumball machines equally likely which suggests that the prior over event probabilities is indeed uniform.

⁹Alternatively, one can also see the threshold distribution for the bare form as a distribution over a verification parameter η that governs how certain a speaker has to be to utter a bare assertion [?, see, e.g.,]Lewis1976,Potts2007,Davis2007,Moss2018. Mathematically, our assumption of bare forms having a threshold semantics is equivalent to assuming that bare assertions are only true when a

threshold distributions has the desirable effect of allowing us a unified treatment of all expressions included in the model.

3.2.2 Parameter estimation

Given all the assumptions outlined above, the model has 18 parameters in total: A cost parameter γ , a rationality parameter λ , a noise strength parameter δ , a constant corresponding to other utterances O , and for each utterance, Beta distribution parameters α_e and β_e . We estimated these parameters jointly from all 21 conditions of Experiment 1 using Bayesian data analysis [?, BDA; see, e.g.,]Kruschke2015. To construct the dataset, we treated the ratings by each participant as a probability distribution from which we sampled 10 utterances. We used highly uninformative uniform priors over the interval $[0, 15]$ for the Beta distribution and cost parameters, uniform priors over the interval $[0, 7]$ for the rationality parameter, and uniform priors over the interval $[0, 0.5]$ for O and the noise strength parameter. We estimated the vector of parameters Θ using MCMC with a Metropolis Hastings sampler. To decrease autocorrelation of the chain, we collected a sample only at every 10th iteration (i.e., we use thinning of 10). We discarded the first 10,000 burn-in samples and then collected 50,000 samples. We ran four MCMC chains and confirmed convergence by computing the \hat{R} -statistic [?]. More details on the implementation of the model can be found in Appendix C.

3.2.3 Model evaluation

The result of the parameter estimation procedure is a posterior distribution over parameters given the observed data $P(\Theta \mid D_{obs})$. We evaluated model fit by performing a posterior predictive check [?, PPC;]Kruschke2015. To this end, we took 10,000 samples of parameters Θ from the posterior distribution. For each sample, we computed the model predictions $ES_1^{(\mathcal{C})'}(u \mid \phi)$ parameterized by Θ . We then compared the average model predictions to the mean ratings that participants had provided in

speaker's credence of the proposition exceeds the verification threshold η .

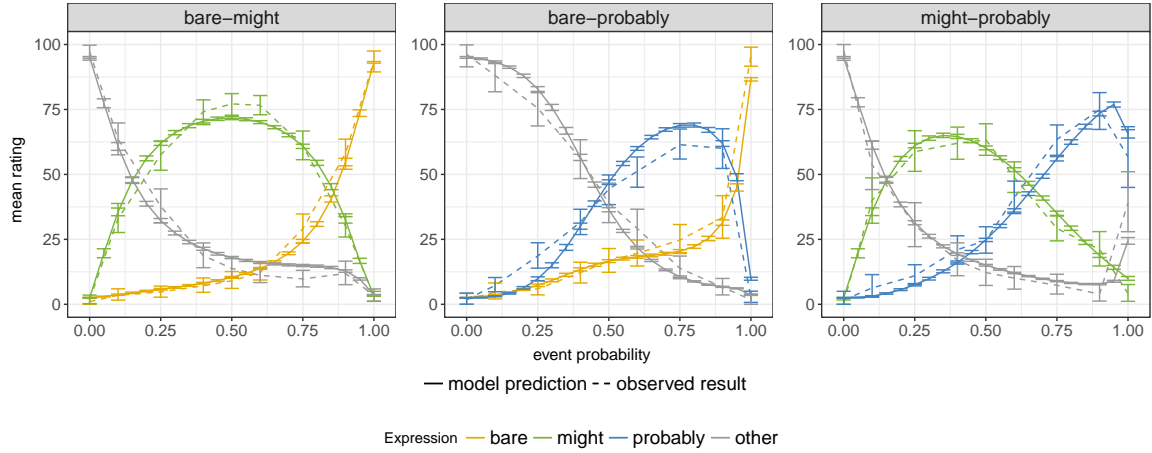


Figure 3.5: Model predictions and results from Experiment 1. Error bars correspond to 95% high density intervals (model predictions) and bootstrapped 95%-confidence intervals (observed results).

the pre-exposure experiments. We further computed the 95% high density interval [?, HDI;]]Kruschke2015 which reflects the certainty of the model about its predictions.

Figure ?? shows the model predictions and the experimental data for three conditions (see Table ?? and Appendix D for modeling results for all 21 conditions). As these plots show, the model is able to capture almost the entire variance in participants' average ratings. Further, the 95% HDIs are very small, which suggests that the model is certain about its predictions. Both of these observations are also true for the model's predictions for all the other conditions. For 19 of the 21 conditions, the R^2 value between the model predictions and the experimental data exceeds 0.9, and for the remaining 2 conditions, the R^2 value exceeds 0.88.

Most cases in which the model predictions and the experimental data deviate concern the ratings at the two extremes of the event probability space. The model often underpredicts ratings for the *something else* option when there is either a 0% or a 100% chance of getting a target color gumball. In these situations, participants presumably thought that BARE and BARE NOT are the most appropriate utterances and therefore rate *something else* highly unless we provide them with the BARE or BARE NOT options. The model predicts this behavior to some extent but seems to assume that participants were cued more heavily towards the presented utterance

options than they actually were. This could be an indicator that we should revisit our unconventional approach of treating the bare forms like uncertainty expressions with a threshold semantics, since the model would predict higher ratings for *something else* at both ends of the scale if we assumed that the bare form and its negation were only true in the cases of 100% and 0% event probabilities, respectively. However, for our purposes in this paper, the exact predictions about production choices for objectively certain events are not as important and hence we decided against revising the assumption that all utterances in the model have a threshold semantics.

The comparably low R^2 for the *probably-think* condition ($R^2 = 0.88$) stems from the fact that participants disagreed on the ordering of these two expressions. This led to a bimodal distribution of average ratings (see the figure in Appendix B) for THINK, which our population-level model cannot fully capture. This suggests that if we wanted to perfectly model participants' production expectations, we should additionally model participant-level differences. However, since the ordering of *probably* and *think* is not of great relevance for our investigation of adaptation to different uses of *might* and *probably*, we did not attempt to fit a more complex model that can account for individual listener differences.

One potential concern given the flexibility of the model is that it could be overfitting the data. This is unlikely considering that all parameters are shared across all conditions and thus we are estimating only 18 parameters to predict in total 567 data points (27 data points for each one of the 21 conditions). To nevertheless rule out the possibility of overfitting, we performed a leave-one-out cross-validation of the model. For each condition x , we estimated a distribution over parameters Θ_x using the data from all conditions but x . We then compared the model predictions of the model parametrized by Θ_x to participants' ratings in condition x . This way, the model has to predict participant behavior which it has not observed during parameter estimation. Table ?? shows the R^2 values for participants's ratings and model predictions for the model estimated from all conditions and the leave-one-out models.

As this table shows, the R^2 values remain high¹⁰ even if we exclude the data on

¹⁰We only observed slightly bigger drops in correlations for the *bare not-bare* and the *bare not-could* conditions. In these conditions, we paired expressions that are a poor description for a large range of event likelihoods: BARE NOT and BARE are only expected to be good descriptions if the

Condition	R^2 (all data)	R^2 (leave-one-out)
bare-might	0.992	0.988
bare-probably	0.978	0.976
bare-could	0.978	0.976
bare-looks like	0.927	0.896
bare-think	0.968	0.964
might-probably	0.964	0.954
might-could	0.921	0.910
might-looks like	0.934	0.918
might-think	0.946	0.934
probably-could	0.961	0.959
probably-looks like	0.944	0.931
probably-think	0.888	0.860
could-looks like	0.924	0.910
could-think	0.931	0.920
looks like-think	0.970	0.960
bare not-bare	0.894	0.848
bare not-might	0.968	0.958
bare not-probably	0.910	0.893
bare not-could	0.910	0.840
bare not-looks like	0.927	0.903
bare not-think	0.933	0.920

Table 3.1: R^2 values for experimental data and model predictions for model estimated from all data and for models estimated from all conditions except the predicted condition.

which the model is evaluated from the model’s training data, which suggests that our proposed model indeed explains participants’ expectations of a generic speaker’s uncertainty expressions.

Lastly, one of the advantages of Bayesian cognitive models is that their parameters are interpretable. Figure ?? shows the maximum likelihood estimates of the inferred

event probability is very low or very high, respectively, and the *bare not-could* condition includes expressions of which neither are well suited to describe high event probabilities. As a result, participants in these conditions seemed to be more uncertain what to do which was reflected in several participants commenting in a post-experiment survey that they would have answered differently had there been better response options. Participants’ ratings were therefore presumably more idiosyncratic, which made it harder for the model to predict ratings without having access to the data from these conditions.

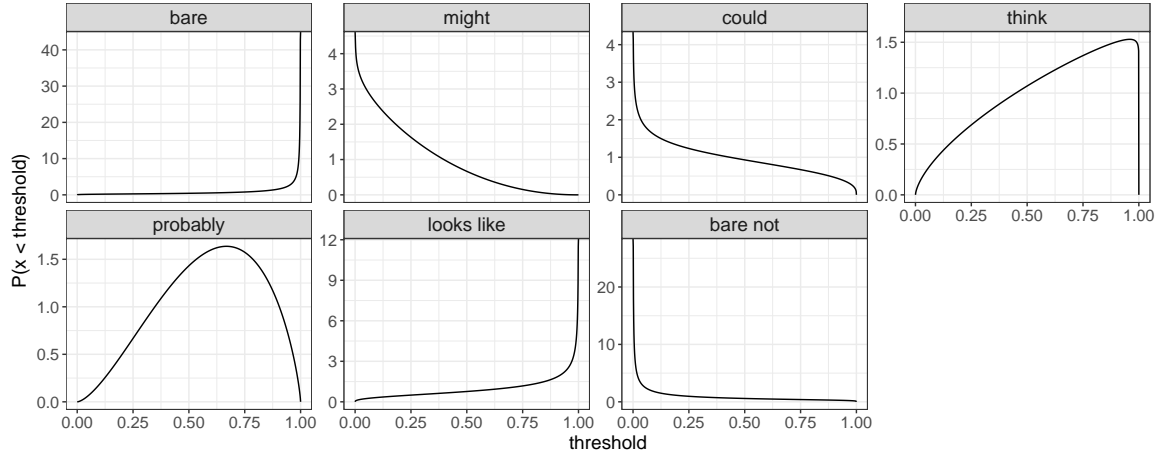


Figure 3.6: Inferred threshold distributions. For the negative bare utterance (BARE NOT), the distribution is over an upper threshold, i.e., a bare statement embedded under negation is true if the probability of the event is lower than the threshold. For all other utterances, the distribution is over a lower threshold.

threshold distributions $P(\theta)$ for the seven uncertainty expressions that we included in our experiments. The first observation is that most threshold distributions have considerable variance rather than being peaked at a particular value. This suggests that listeners have probabilistic beliefs – i.e., uncertainty – about the semantic threshold for each utterance.

Further, the inferred threshold distributions are broadly in line with qualitative accounts of the meaning of uncertainty expressions. As discussed above, for the bare form and its negation, we expected the model to infer threshold distributions whose probability mass is concentrated around $\theta = 1$ and $\theta = 0$, respectively.¹¹ As Figure ?? shows, this is indeed what the inferred threshold distributions look like.

The probability mass of the threshold distribution for *might* is concentrated at values slightly above 0. This, too, is in line with non-probabilistic accounts of the interpretation of epistemic modals. These accounts generally assume that *might p* is true if there exists a world w in a set of (contextually restricted) epistemically accessible worlds E such that p is true in w [?, e.g.,]Kratzer1991,Swanson2008,Hacquard2011.

¹¹Note that since the negation of the bare form is a negative form, θ is an upper threshold. For the bare form as well as all the other utterances that we are considering in this paper, θ is a lower threshold.

One way to translate this logical condition into our probabilistic framework is to assume that in our gumball machine context, there exists an epistemically accessible world w for each gumball g and that in world w , one gets gumball g . Under this assumption, “*You might get a blue gumball*” is true if there exists an epistemically accessible world w in which one gets a gumball g that is blue. At the same time, if such a world exists, then $P(\text{blue gumball})$ is greater than 0, which approximately corresponds to the threshold semantics with the inferred threshold distribution of the model. The inferred threshold distribution for *could* is very similar, but not identical, to the one of *might*. This also largely reflects the prediction by non-probabilistic accounts, which generally assume that *might* and epistemic *could* are semantically equivalent [?, ?], but it also suggests that there are differences between these two expressions.

The threshold distribution for *probably* has most of its probability mass concentrated at thresholds above 0.5. This again largely reflects the predictions of existing accounts that assume that *probably* p is true if p is more likely than the negation of p [?, e.g.,]Kratzer1991. However, it is also noteworthy that the inferred distribution has some probability mass below 0.5, which potentially provides evidence for theoretical arguments by [?] that *probably* p can sometimes also be true if p is less likely than the negation of p .¹²

The threshold distributions for the remaining expressions, *looks like* and *think* also match intuitions. The distribution for *looks like* has most of its probability mass near threshold values of 1 but is overall slightly weaker, i.e., assigns higher probabilities to lower thresholds, than the bare form. The distribution for *think* assigns most probability mass to high thresholds, which is compatible with the intuition that speakers use *think* when they strongly believe the embedded proposition but are not entirely certain that it is true.

¹²Since we did not investigate whether participants always correctly perceived the event likelihood to be less than 0.5 for gumball machines with less than 50% target color gumballs, an alternative explanation for this observation would be that participants sometimes overestimated the event likelihood and for this reason, they expected the generic speaker to use *probably* despite an objective event likelihood of less than 0.5.

	λ (rationality)	γ (cost)	δ (noise)	O (other utterances)
MAP	2.21	3.03	0.074	3.64×10^{-5}
CI	[2.16, 2.26]	[2.98, 3.08]	[0.069, 0.079]	$[2.89 \times 10^{-5}, 4.50 \times 10^{-5}]$

Table 3.2: Estimated maximum a posteriori estimates (MAP) and 95% credible intervals (CI) for model parameters.

Table ?? shows the MAP values and credible intervals for the remaining parameters. The model inferred that speakers are relatively likely to choose an optimal utterance (reflected in the λ parameter being greater than 1); that utterances that are not included in the experiment incur a considerable cost (reflected in the γ parameter being greater than 1); that about 7.4% of the data should be treated as noise (reflected in the δ parameter); and that the production probability of utterances not included in our set of utterances is low.

3.2.4 Interim summary

In this section, we described a computational model of production expectations of uncertainty expressions. This model is couched within the RSA framework and assumes that listeners hold beliefs about a speaker’s lexicon (in the form of utterance-specific threshold distributions) and about speaker preferences (in the form of utterance-specific costs). We estimated the free parameters of this model from the results of Experiment 1, which resulted in a model that is able to accurately predict participants’ utterance ratings – i.e., their expectations of use – in Experiment 1 across all conditions with a shared set of parameters.

In the following sections of this paper, we will use this model as the basis for modeling adaptation. Since this model is able to capture different beliefs about thresholds and preferences, it provides us with the opportunity to simulate the adaptation process as a result of updating beliefs about these model parameters. Further, in order to answer our primary research question of whether listeners update their beliefs about lexica or preferences, we compare different adaptation simulations in which we allow different types of parameters to be updated.

Chapter 4

Adaptation

4.1 Adaptation model

The experimental results presented in the previous section suggest that listeners update *some* expectations about language use when they interact with a speaker. However, the nature of the updated representations is unclear. As mentioned in the introduction, there are three likely candidates: first, it is possible that listeners update their expectations about the speaker’s lexicon (i.e., the mapping between event probabilities and uncertainty expressions); second, listeners might update their expectations about the speaker’s preferences; and third, they might update both their expectations about the speaker’s lexicon and about the speaker’s preferences. The experimental results above suggest that it is unlikely that listeners track only speaker preferences, but considering that beliefs about preferences and beliefs about the lexicon can interact in complex ways (as illustrated in Figure ??), we investigate all three options.

The production expectation model in Section ?? provides us with the opportunity to formally evaluate these three hypotheses. Through a series of simulations of the adaptation process, we can compare models in which different types of parameters are updated during adaptation. Following work in modeling adaptation in other linguistic domains [?, e.g.,] Kleinschmidt2012, Kleinschmidt2015, Qing2014, Hawkins2017, Roettger2019, we assume that in interaction, listeners form beliefs about a set of speaker-specific parameters Θ_S .¹ We further assume that the formation of these beliefs is an instance of Bayesian belief updating: listeners start off with prior beliefs about Θ_S based on their general knowledge about language and subsequently update their beliefs about Θ_S with every utterance they hear. That is, after observing a series of productions $D = d_1, \dots, d_n$ where each d_i is an utterance-event probability pair $d_i = (u_i, \phi_i)$,

¹Since the manipulation in our experiments was between subjects, our results do not provide direct evidence that listeners are indeed adapting to speakers (as compared, for example, to the general experimental situation). For now, we assume that listeners are adapting to specific speakers and we return to this issue in the general discussion. Further, the speaker-specific parameters might be correlated and listeners might form beliefs about bundles of correlated parameters instead of forming beliefs about individual parameters. For simplicity, we assume here that individual parameters are independent but it would be interesting to investigate whether, for example, listeners who expect a speaker to use lower thresholds for *probably* also expect the same speaker to have lower thresholds for *might* (see also Section ??).

listeners' beliefs about Θ_S are the result of performing Bayesian inference:

$$P(\Theta_S | D) \propto P(\Theta_S)P(D | \Theta_S) = P(\Theta_S) \prod_{i=1}^n P(d_i | \Theta_S)$$

We assume that the likelihood function is the *expected pragmatic speaker* ES_1 parameterized by Θ_S :

$$P(\Theta_S | D) \propto P(\Theta_S) \prod_{i=1}^n ES_1(u_i | \phi_i, \Theta_S)$$

4.1.1 Simulations

In order to investigate which parameters are updated during adaptation, we ran simulations with varying prior structures, which correspond to different assumptions about which parameters may be updated. The adaptation model crucially relies on a prior over speaker-specific parameters $P(\Theta_S)$ which reflects listeners' prior beliefs about the use of uncertainty expressions. For our simulations, we assumed that the means of this prior are given by the estimates of the model parameters that we obtained from fitting the model to the norming data. The variances reflect whether or not the parameter can be updated in response to exposure. In particular, we used delta distributions, i.e., distributions with zero variance, to model a parameter that cannot be updated. We ran simulations on models with the following three prior structures:

- **Costs:** The first prior structure corresponds to an adaptation process according to which participants only learn speaker preferences during adaptation. We modeled the prior over cost parameters as a log-normal distribution centered at the mean value inferred from the norming data. Because we were interested in whether listeners update their beliefs about speaker preferences, we relaxed the assumption from the norming data model that all utterances have the same cost and assumed that each expression has its own cost parameter indicating beliefs about the speaker's preferences.² Use of the log-normal distribution was

²Note that since all updates of parameters happen as part of the belief updating simulations and

motivated by two reasons: First, cost must be greater than zero, and the support of log-normal distributions is limited to numbers greater than 0. Second, since the cost term is part of an exponent in the expected pragmatic speaker model, differences on a logarithmic scale correspond to linear differences in the model’s utterance probabilities. For the priors over all other parameters, we used a delta distribution.

- ***Threshold distributions***: This prior structure corresponds to an adaptation process according to which participants only learn speaker-specific threshold distributions during adaptation. We parameterized threshold Beta distributions $P(\theta_e)$ with their mean μ_e and population parameter ν_e [?]. Since the threshold and therefore also the mean of the threshold distribution have to lie within the interval $[0,1]$, we used a truncated normal distribution $\mathcal{N}_{[0,1]}$, which we centered at the mean value from the norming data. For the population parameters ν_e , which indicate how peaked a threshold distribution is, we assumed that distributions can only become more peaked when listeners are exposed to a speaker with very consistent language use and therefore modeled the prior as an exponential distribution shifted to the mean population parameter that we estimated from the norming data. We used a delta distribution for the priors over all other parameters.
- ***Threshold distributions and costs***: This prior structure corresponds to an adaptation process according to which participants learn both speaker-specific threshold distributions and speaker preferences during adaptation. We used the log-normal distributions as priors over the cost parameters and the truncated normal and exponential distributions as priors over the threshold distribution parameters, as described above. This means that both the threshold distributions and the cost parameters could be updated during the adaptation simulations.

Each of these prior structures corresponds to a different hypothesis about which

we are not fitting model parameters to post-exposure ratings, we do not have to be concerned about overfitting due to too many parameters.

	Range	Step size	MAP value
Variance for μ	[0.025,0.25]	0.025	0.175
Scale for ν	[0.5,4.5]	0.5	3.5
Variance for cost	[0.1,1.5]	0.2	0.7

Table 4.1: Explored hyperparameter ranges for variance parameters, and inferred MAP values, which were used in the adaptation simulations.

expectations listeners update during adaptation. For comparison, we also considered a baseline in which none of the parameters are updated during adaptation (the *fixed* prior structure). To adjudicate between these three hypotheses, we ran simulations of the adaptation process for both (*cautious speaker* and *confident speaker*) conditions with different prior structures and compared the models in terms of their likelihood of generating the experimental data. During each simulation, we performed Bayesian inference to infer the posterior parameter distribution after observing the 25 data points that participants observed in the exposure phase (see Table ?? for an overview of the 25 utterances in the two conditions). We performed inference using MCMC with a Metropolis-Hastings sampler. We used thinning of 10, discarded the first 2,000 burn-in samples and collected 10,000 samples from each of the two chains.

The prior distributions over the different parameters that may be updated during the adaptation simulations are all parameterized by two constants which govern their mean and their variance. The first set of parameters (the mean of the log-normal and truncated normal distributions; the location parameter of the exponential distributions) was given by the estimates from fitting the model to the norming data. The second set of parameters (the variance of the log-normal and truncated normal distributions; the scale parameter of the exponential distributions) was treated as hyperparameters of the simulations. To keep the model as simple as possible, we only used three hyperparameters in total: a variance parameter for the cost for all expressions; a variance parameter for the mean of the threshold distributions for all utterances; and a scale parameter for the prior over population parameters for all utterances. We optimized these three parameters through a Bayesian hyperparameter search on the adaptation data, which provided us with a distribution over hyperparameter values. Considering that each simulation is computationally expensive, we

could only test a few hundred hyperparameter combinations, which are listed in Table ???. We found that the resulting distributions were highly peaked and therefore, we used only the maximum a posteriori estimates of the hyperparameters (also shown in Table ??) for the model comparisons below.

4.1.2 Model comparisons

We compared model fits via the likelihood of the model generating the post-exposure data. To compute this metric, we constructed a dataset D_{obs} of utterance-event probability pairs by treating each post-exposure rating as a probability distribution and sampling 10 utterances from it. We then computed the posterior likelihood odds between Model 1 with posterior distribution over parameters $P(\Theta_S^{(1)})$ and Model 2 with posterior distribution $P(\Theta_S^{(2)})$.

$$\text{posterior likelihood odds} = \frac{\int_0^1 P(\Theta_S^{(1)}) P(D_{obs} | \Theta_S^{(1)}) d\Theta_S^{(1)}}{\int_0^1 P(\Theta_S^{(2)}) P(D_{obs} | \Theta_S^{(2)}) d\Theta_S^{(2)}}$$

The posterior likelihood odds indicate how much more likely it is that the data was generated by Model 1 than by Model 2. Note that we are marginalizing over distributions over parameter values (the integrals in the above formula) and the more parameters we allow to update, the more dispersed the distributions over parameters $P(\Theta_S^{(1)})$ and $P(\Theta_S^{(2)})$ will be. At the same time, the likelihood terms $P(D_{obs} | \Theta_S^{(1)})$ and $P(D_{obs} | \Theta_S^{(2)})$ will only be high for a small range of parameter assignments. Taken together, this means that the posterior odds metric will naturally favor simpler models, because the more dispersed the distributions over parameters are, the less weight is assigned to parameter configurations that yield a high likelihood of generating the data [?, a property often referred to as Bayesian Occam's razor; see, e.g.,]MacKay1992,Neal1995.

To evaluate how well the different models predict the empirical post-exposure ratings, we also compute the R^2 value between participants' average post-exposure

Model	odds	R^2
fixed	10^{-1265}	0.673
cost	10^{-478}	0.817
threshold distributions	10^{-187}	0.885
cost & threshold distributions	1	0.901

Table 4.2: Model evaluation results on data from Experiment 2. *odds* are the posterior likelihood odds of the models compared to the *cost and threshold distributions* model. R^2 are computed between the mean post-exposure ratings and the mean model predictions.

ratings and the maximum a posteriori predictions of the post-exposure model. However, while R^2 values close to 1 always indicate that model predictions closely match empirical ratings, it is important to note that there are multiple caveats with using R^2 in our setup. First, R^2 is based on the assumption that all predictions are independent and that the residual error follows a normal distribution. These assumptions generally hold for regression models but they are violated for our model predictions since the model predicts a probability distribution over utterances for each event likelihood and therefore predicted utterance ratings are not independent and the residual errors do not follow a normal distribution. Further, unlike the posterior odds, R^2 does not take uncertainty of the model predictions into account but instead compares the mean participant behavior to the mean model predictions, and therefore is not a well-suited measure for comparing models exhibiting uncertainty.³ For these reasons, we present R^2 values only to provide intuitions about how well the model predictions match the empirical ratings. For the formal model comparisons we rely on the posterior odds.⁴

Table ?? shows the posterior likelihood odds and the R^2 values between the the models and the experimental data from Experiment 2. As the values in this table

³If the assumptions of independence and normally distributed residual errors were met, we could address the issue of uncertain model predictions by using a recently proposed Bayesian R^2 metric [?], which takes uncertainty of model predictions into account and returns a distribution over R^2 values. However, the computation of this metric crucially relies on estimates of the variance of the residual error σ^2 , and it neither makes sense nor is possible in our current framework to estimate this quantity. Hence, we also cannot compute a meaningful Bayesian R^2 distribution.

⁴For all simulations discussed in this and following sections, we find that the R^2 and the posterior odds lead to the same conclusions but in an additional simulation experiment, reported in Appendix F, we encountered a case where the two metrics disagreed.

show, the model in which the cost as well as the threshold distributions are updated during adaptation is much more likely to generate the experimental data than the other two adaptation models as well as the prior model. Further, this model closely predicts participants post-exposure behavior, as indicated by the high R^2 value.

What do these modeling results tell us about the semantic/pragmatic adaptation process? We assumed that each of these simulations correspond to an adaptation process in which different types of expectations are updated. The modeling results corroborate the experimental results from Experiment 2: the models according to which no expectations are updated during adaptation (the *fixed* model) or according to which only preferences are updated (the *cost* model) provide poor predictions for the post-adaptation ratings. The results further clearly indicate that listeners update expectations about the threshold distributions: The two models according to which listeners update threshold distributions were best at predicting post-adaptation behavior. Overall, however, the model that allows updating of preferences and threshold distributions (the *cost & threshold distributions* model) provides the best predictions of post-adaptation ratings, which provides evidence that listeners update expectations about both threshold distributions and preferences.

4.1.3 Model evaluation

Apart from quantitatively assessing the fit of the model, it is informative to visually inspect the predictions of the model to verify that the model makes correct qualitative predictions. Figure ?? shows the post-exposure predictions of the *cost & threshold distributions* model compared to the average participant ratings for the two conditions from Experiment 2. Qualitatively, the model captures several important patterns in the post-adaptation behavior. The model correctly predicts that in the *cautious speaker* condition, ratings for MIGHT are higher than ratings for PROBABLY when there is an objective probability of 0.6. For the *confident speaker* condition, the model correctly predicts the opposite pattern. The model also predicts that in the *cautious speaker* condition, participants rate MIGHT highly for a larger range of event probabilities than in the *confident speaker* condition and the model predicts

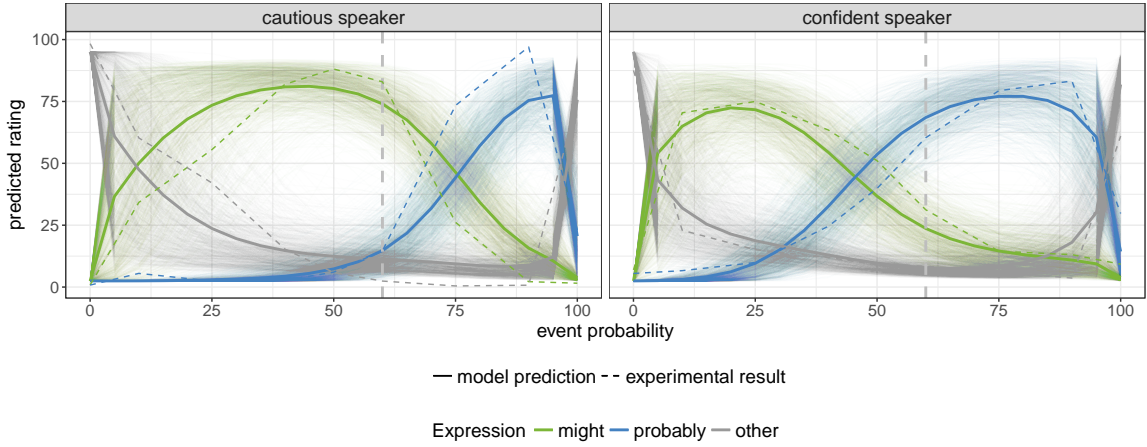


Figure 4.1: Post-adaptation model predictions from simulations for Experiment 2 and experimental results. The solid lines shows the mean model predictions and the thin lines around the mean show the distribution of model predictions.

the reverse pattern for the PROBABLY ratings. Further, the model predicts that high ratings for MIGHT and PROBABLY are not limited to the utterance-event probability combinations that participants observed during the exposure phase. For example, the model correctly predicts high ratings of MIGHT for low event probabilities in the *cautious speaker* condition despite the fact that it never observed utterances for low event probabilities. Similarly, the model predicts high rating of PROBABLY for high event probabilities in the *confident speaker* condition – a combination which was again not part of the exposure trials of this condition.

Quantitatively, there are some differences between the model predictions and participant behavior. This is not surprising considering that the model predictions are a result of simulations and, with the exception of the three variance parameters of the prior distributions, we did not fit any model parameters to the behavioral data. One difference is that the model underpredicts the ratings of one of the PROBABLY utterances in the *cautious speaker* condition. One reason for this deviation could be the relatively simple prior structure. For the priors, we made the assumption that all model parameters are independent of each other and that the variance for the different parameter types is the same for all utterances. However, it could be that

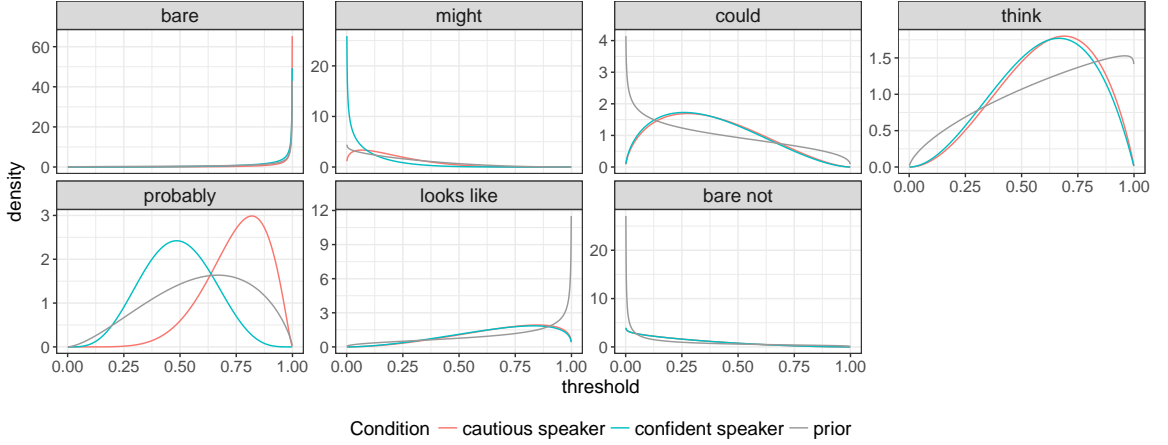


Figure 4.2: Post-adaptation threshold distributions from the simulations for Experiment 2.

listeners have more structured prior beliefs such that priors over different parameters are correlated or variances of prior distributions differ. For example, it could be that listeners expect the thresholds for *might* and *probably* to be correlated such that higher thresholds for *might* are correlated with higher thresholds for *probably*. Or it could be that listeners expect more between-speaker variation for some expressions than for others. Considering that we only have data from two experiments to test model predictions and therefore would likely overfit to the data if we tried to fit more complex prior structures with more parameters, we leave the investigation of the exact structure of listeners' prior beliefs to future work.

The second noticeable deviation is that in the *confident speaker* condition, the model overpredicts the ratings of the OTHER utterance for event probabilities of 1. This prediction is primarily driven by high values for the predicted ratings of BARE. However, we argue that the model predictions in this case are reasonable, and that the lower participant ratings are likely an artifact of the experimental task. After completing the experiment, several participants indicated in a feedback form that they were confused by the lack of an option to rate the BARE utterance, which they had heard during the exposure phase. In light of this confusion, almost all individual participants chose among two strategies when there was an event probability of 1: they either provided a rating of 100 for OTHER or they provided a rating of 100 for

PROBABLY, which on average leads to the ratings shown in Figure ??.

With the exception of these two deviations, the model makes not only correct qualitative, but also accurate quantitative predictions for the post-exposure ratings.

Lastly, we can also inspect how the model arrived at its predictions by looking at the inferred model parameters. Figures ?? and ?? show the inferred post-exposure threshold distributions and costs for the two conditions as well as the distributions inferred from the norming data. Figure ?? shows that the threshold distribution for *probably* changed considerably depending on the exposure phase: in the *cautious speaker* condition, its mean shifted to a higher value than inferred from the norming data; in the *confident speaker* condition the mean shifted to a lower value. To a lesser extent, we observe similar shifts in the mean of the threshold distributions for *might*. We further observe that for all expressions, the variance of the threshold distributions decreased as a result of adaptation. In the case of the expressions that were part of the exposure phase, this is expected, since the exposure speaker used these expressions very consistently; in the case of the other expressions, this decrease in variance is a result of the exponential prior over the population parameter, which biased the model towards lower variance. For some of the thresholds, this resulted in differently shaped distributions. However, note that the area under the curve of all threshold distributions except for *probably* is still very similar to the area under the curve of the norming data threshold distributions. And overall, except for the distributions for *might* and *probably*, the post-exposure threshold distributions are almost identical in both conditions. This suggests that the post-adaptation expectations are in part a result of updated threshold distributions for *might* and *probably*.

Figure ?? shows that the costs of the MIGHT, PROBABLY and BARE utterances, i.e., the three utterances that participants observed during the exposure phase, all decreased while the costs of the other four utterances increased compared to the costs inferred from the norming data. Further, the post-exposure cost of *might* is lower than the cost of *probably* in the *cautious speaker* condition and the opposite relation between these costs holds in the *confident speaker* condition, which suggests that the post-adaptation expectations are in part also a result of updated beliefs about

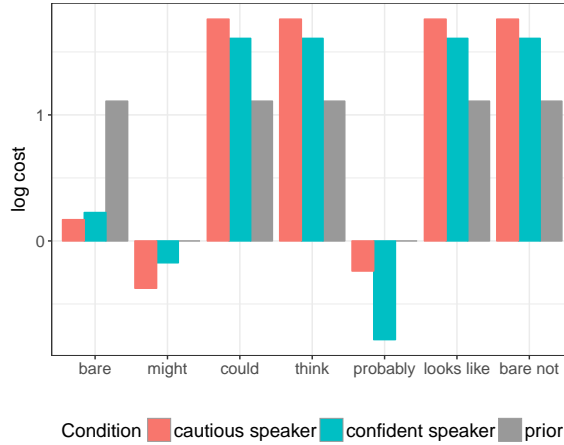


Figure 4.3: Post-adaptation *log* cost values from simulations for Experiment 2. Note that the cost of MIGHT and PROBABLY in the norming data model was 1 and therefore the *log* cost for these utterances is 0.

preferences of MIGHT and PROBABLY. This finding also highlights the complex interplay between threshold distributions and preferences: The number of exposure trials with MIGHT and PROBABLY was identical in both conditions, so participants could not have inferred a preference based on exposure frequencies. Instead, participants seemed to indirectly infer that the *cautious speaker* prefers to use MIGHT from the speaker’s uses of MIGHT to describe a larger range of event probabilities and that the *confident speaker* prefers to use PROBABLY from the speaker’s uses of PROBABLY for a larger range of event probabilities. However, frequency nevertheless has an effect on inferred preferences. As shown in Appendix F, the difference in inferred preferences was larger when participants in the *cautious speaker* condition were exposed to more trials with MIGHT than in the *confident speaker* condition and participants in the *confident speaker* condition were exposed to more trials with PROBABLY than in the *cautious speaker* condition.

4.1.4 Interim summary

In the previous two sections, we presented the results from an experiment that provides strong evidence for listeners updating expectations about a speaker’s use of

uncertainty expressions after brief exposure to that speaker. We further presented a computational adaptation model which models the adaptation process as an instance of Bayesian belief updating. Using different implementations of that model to investigate which kind of expectations listeners update during adaptation, we found strong evidence that listeners update beliefs both about threshold distributions and about speaker preferences.

4.2 Experiment 3: Effect of adaptation on interpretation

Up to this point, we focused on listeners' expectations about a speaker's use of uncertainty expressions. As we discussed in the introduction, we expect updated expectations to also have an effect on the *interpretation* of uncertainty expressions. This effect is also predicted by RSA models, which assume that a pragmatic listener L_1 tries to infer the state of the world (in our case, the event probability ϕ) by reasoning about their prior beliefs about the world state and their expectations about a speaker's language use (in our case, the expected pragmatic speaker ES_1) via Bayes' rule:

$$L_1(\phi \mid u_e) \propto P(\phi)ES_1(u_e \mid \phi).$$

According to such a model of interpretation, the shifts in expectations that we observed in the previous experiment should also lead to a shift in interpretations. If we assume a uniform prior over event probabilities,⁵ then the model predicts that listeners who were exposed to a *cautious* speaker should infer higher event probabilities when hearing MIGHT or PROBABLY than listeners who were exposed to a *confident* speaker.

⁵To reiterate, this assumption was motivated by the study reported in Footnote 9, which suggested that participants on average assign equal probability to each gumball machine a priori. In this special case, we expect interpretations to be an exact mirror image of production expectations but in many other scenarios, listeners will have skewed prior beliefs about the event likelihood (e.g., consider beliefs about the likelihood of rain on a specific summer day in California, which will be highly skewed towards low values) and in such scenarios, listeners integrate prior beliefs and production expectations to arrive at interpretations.

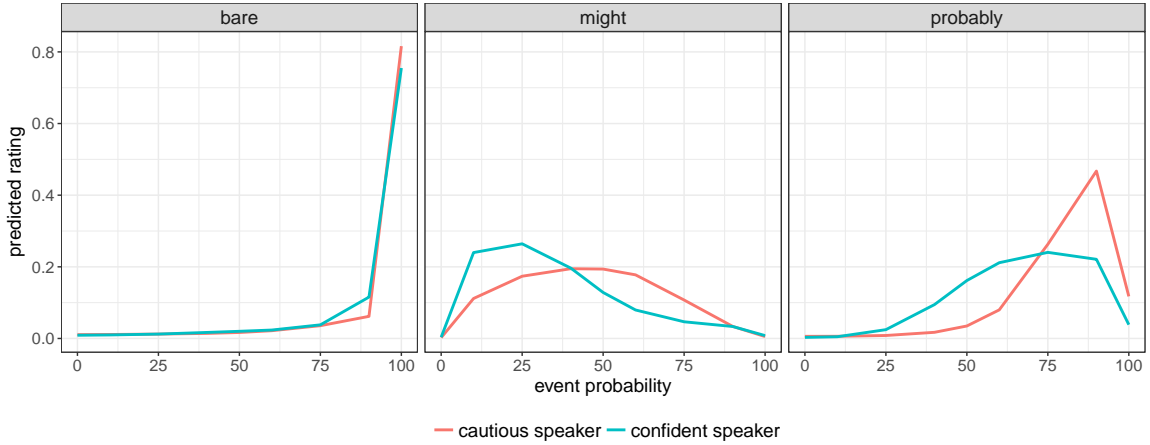


Figure 4.4: Post-adaptation interpretation distributions for the utterances BARE, MIGHT, and PROBABLY as predicted by the pragmatic listener L_1 .

Figure ?? shows the distribution over event probabilities after hearing three different utterances as predicted by L_1 parameterized by the inferred parameters from our adaptation simulations in the previous section. As these plots show, in the *cautious speaker* condition, the distribution over event probabilities after hearing *might* and *probably* is shifted towards higher values as compared to the distributions in the *confident speaker* condition.

In this experiment, we tested whether this prediction is correct and whether listeners' change in expectations transfers to a change in interpretations. The procedure, materials and analyses were pre-registered at <https://osf.io/ghnc3>.⁶

4.2.1 Method

Participants

We recruited a total of 80 participants (40 per condition) on Amazon Mechanical Turk. We required participants to have a US-based IP address and a minimal approval rating of 95%. Participants were paid \$1.5 which amounted to an hourly wage of

⁶This experiment is a modified version of a previous experiment, which qualitatively yielded the same results but seemed to confuse some participants. See Appendix G for a discussion of the original experiment.

approximately \$15. None of the participants had participated in any of the previous experiments.

Materials and Procedure

Participants completed a set of exposure trials followed by a set of test trials. The exposure trials were identical to the exposure trials in Experiment 2. The test trials probed participants' interpretations of the utterances MIGHT, PROBABLY and BARE. On each test trial, participants listened to a recording of the speaker from the exposure phase producing MIGHT, PROBABLY and BARE and then participants were asked to rate for 9 gumball machines with the same proportions of blue and orange gumballs as in the previous experiments how likely they thought it was that the speaker saw each of these gumball machines by distributing coins. Participants distributed 10 coins per trial and completed 6 trials in total – one for each expression-color pair. The exposure phase again contained 6 attention check as in the previous experiment. However, given the low attention check performance in the previous experiments, we modified the attention checks. Instead of asking participants whether they saw an X on the previous trial, we asked participants to choose the gumball machine that they had seen on the previous trial among two machines displayed in random order.

Exclusions

We excluded participants who failed more than 2 attention checks, which led to 1 exclusion in the *cautious speaker* condition and 1 exclusion in the *confident speaker* condition.

Analysis and Predictions

If participants update their expectations of a specific speaker's use of uncertainty expressions, we expect them to interpret a more confident speaker's utterance to communicate a lower event probability than a more cautious speaker's utterance. We tested this prediction by treating participant's distributions of coins of gumball machines as a probability distribution over gumball proportions (and consequently

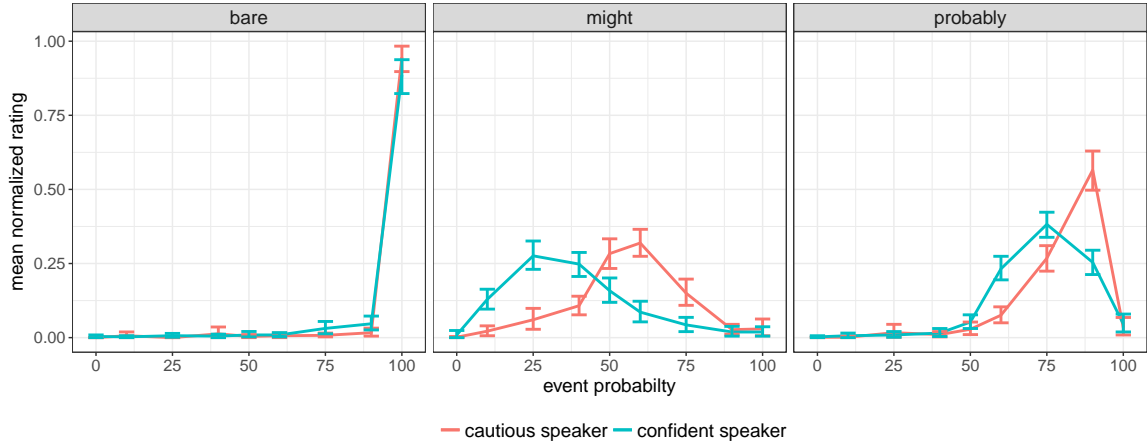


Figure 4.5: Aggregated post-exposure ratings from Experiment 3. Error bars correspond to bootstrapped 95%-confidence intervals.

event probabilities). For each utterance, we normalized participants' coin distributions such that they summed up to 1, so that we could interpret the normalized scores as a categorical probability distribution over gumball machines given an utterance. We then compared the resulting distributions over target color gumball proportions for each utterance in terms of their expected value, using a t -test. We predicted that the expected values of MIGHT and PROBABLY would be larger in the *cautious speaker* condition than in the *confident speaker* condition.

4.2.2 Results and Discussion

Figure ?? shows the aggregated and normalized ratings for the two conditions. As predicted, participants provided higher ratings for gumball machines with higher target color percentages after hearing MIGHT and PROBABLY in the *cautious speaker* condition than in the *confident speaker* condition. This also led to a significantly higher expected value for MIGHT ($t(76) = 5.84$, $p < 0.001$) and PROBABLY ($t(76) = 3.92$, $p < 0.001$) in the *cautious speaker* condition as compared to the *confident speaker* condition.

These results suggest that listeners not only update their expectations about a speaker's use of uncertainty expressions, but also use those updated expectations in

Model	odds	R^2
fixed	10^{-452}	0.827
cost	10^{-231}	0.883
threshold distributions	10^{-129}	0.887
cost & threshold distributions	1	0.936

Table 4.3: Model evaluation results on data from Experiment 3. *odds* are the posterior likelihood odds of the models compared to the *cost and threshold distributions* model. R^2 are computed between the mean post-exposure ratings and the mean model predictions.

interpretation. Here, we made the implicit linking assumption that the distribution of coins reflects participants’ beliefs about event likelihoods after hearing an utterance. The choice for this paradigm, which is very similar to betting paradigms that have been used to study utterance interpretation in reference games [?, ?] as well as for probing subjective probabilities [?, e.g.,]Hampton1973, was motivated by the assumption that listeners will have uncertainty about the exact event likelihood after hearing MIGHT or PROBABLY. Allowing participants to assign multiple coins to gumball machines with different proportions provided them with the ability to convey this uncertainty. The results from this experiment suggest that participants behaved as expected: they assigned almost all coins to the gumball machine with 100% target color gumballs after hearing BARE, an utterance about whose interpretation participants likely have very little uncertainty, whereas they assigned coins to multiple machines after hearing MIGHT or PROBABLY.

4.2.3 Model comparison

We return again to our main research question regarding which expectations are updated during adaptation. The production expectation experiments and model simulations provided strong evidence for listeners updating their beliefs about the threshold distributions and preferences. To determine the stability of these results, we also compared the pragmatic listener L_1 predictions from the simulations with different prior structures to the post-exposure ratings in Experiment 3. To this end, we computed the predictions of the L_1 model from the posterior distributions over model

parameters that we obtained through the simulations in the previous section. Table ?? shows the model fit for the different types of simulations. As this table shows, the model according to which both threshold distributions and costs are updated provides the best fit according to both metrics. Considering that the posterior likelihood odds consistently favored this model in all model comparisons, we take these results together as strong evidence that listeners update their expectations about threshold distributions and costs.

4.2.4 Model evaluation

Figure ?? superimposes the model predictions and the experimental data. As these plots show, the model accurately captures most of the qualitative and quantitative patterns. First, the model makes both qualitatively and quantitatively accurate predictions for the interpretation of the BARE utterance in both conditions. Second, the model makes the crucial qualitative prediction that participants expect the speaker to communicate lower event probabilities in the *confident speaker* condition than in the *cautious speaker* condition, which we also observed in Experiment 3. Further, even though we used the parameters that we obtained in the simulations from the previous section and did not fit any parameters to the data from Experiment 3, the model also provides good quantitative predictions of participant’s interpretation of MIGHT and PROBABLY, which provides further support for the hypothesis that semantic/pragmatic adaptation is an instance of Bayesian belief updating.

The main deviation between the model predictions and the experimental data lies in the interpretation of MIGHT in the *cautious speaker* condition. For this interpretation, the model predicts a less peaked distribution than the empirical distribution. One explanation for this deviation could be that participants are considering alternative uncertainty expressions (e.g., *very unlikely*) that we did not include in the model. However, since fine-tuning the set of alternative utterances would not change the qualitative predictions of the model and would not provide additional theoretical insights, we leave a more detailed exploration of this issue to future work.

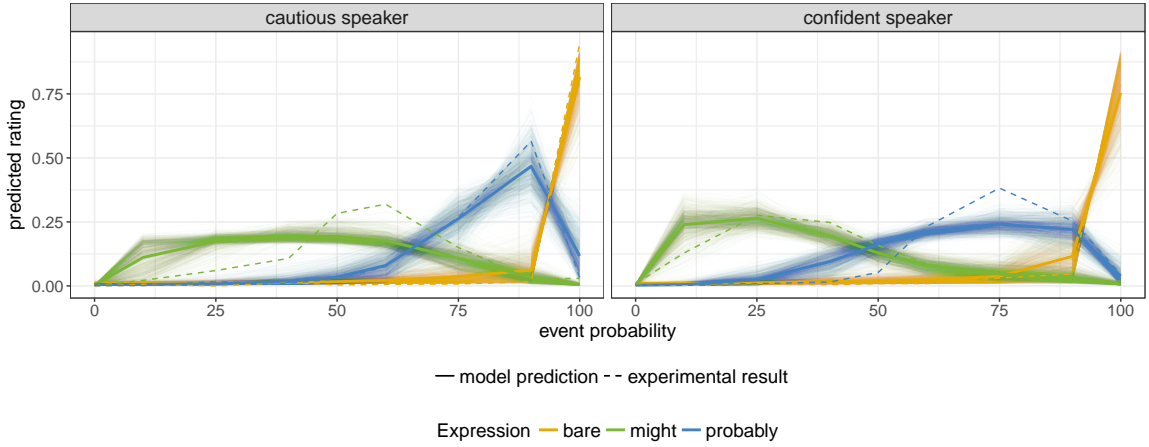


Figure 4.6: Predictions of *threshold distributions and costs* model and data from Experiment 3. The thin lines around the mean show the distribution of model predictions.

4.3 General Discussion

While adaptation in language is a widely attested phenomenon, the nature of the representations that are updated during semantic/pragmatic adaptation has largely remained a mystery. In this paper our contribution has been three-fold:

First, in a production expectation experiment (Experiment 2), we showed that listeners adapt to speakers who vary in their use of uncertainty expressions, demonstrating that the findings from [?] extend to the class of uncertainty expressions. Second, we showed that updated *production* expectations also affect subsequent utterance *interpretation* (Experiment 3). Finally, in a series of model comparisons, we found strong evidence for listeners updating their beliefs about both the speaker’s lexicon as well as the speaker’s preferences, which suggests that semantic/pragmatic adaptation is a result of updating both of these representations. We further found that modeling the adaptation process as an instance of Bayesian belief updating explains participants’ post-adaptation behavior in both the production expectation (Experiments 2) and comprehension (Experiment 3) experiments.

We next discuss the implications of this work for other accounts of adaptation and for semantic theories of uncertainty expressions, as well as its methodological

implications. We then turn to limitations of the current results and account and discuss promising future research avenues this work opens.

4.3.1 Implications for and relation to other accounts of adaptation

The model in this paper is formulated at the computational level [?, ?] and is therefore directly only comparable to other computational models. However, we can still assess the compatibility of our findings with mechanistic accounts. We first discuss the relation to existing computational models of adaptation and then discuss what the results tell us about existing mechanistic accounts of adaptation.

The model presented above follows several other computational models of linguistic adaptation that are based on Bayesian belief updating, including models of phonetic adaptation [?], syntactic adaptation [?], adaptation in the interpretation of prosodic cues [?], and adaptation to variable use of the quantifiers *some* and *many* [?]. All of these models are based on the same Bayesian belief updating procedure according to which listeners integrate their prior linguistic expectations and observed linguistic behavior to form updated expectations that facilitate comprehension. This updating procedure is in line with recent proposals of a “Bayesian brain” [?, e.g.,]Clark2013,Friston2010 which argue that many cognitive and perceptual processes can be seen as an instance of integrating prior beliefs with observed signals from the environment.

Similarly, the formation of conceptual pacts [?] – alignment in the formation of referring expressions – has been explained using a model similar to the one we have proposed here [?]. Hawkins et al.’s model is based on the assumption that speakers and listeners have uncertainty about the lexicon [?, see also]Bergen2016 and that in interaction, speakers and listeners update their beliefs about the shared lexicon, akin to the updating of threshold distributions in our model. This further suggests that belief updating plays an important role in partner-specific language processing.

In the space of mechanistic accounts, [?] argued that a lot of partner-specific

linguistic behavior can be explained as the result of priming, i.e., the automatic activation of linguistic representations when a speaker produces an utterance or a listener hears an utterance. Their account has the appeal of explaining why partner-specific language use often appears to happen automatically and effortlessly. With the additional stipulation that the activated representations include information about both language and the situational context and thus are able to represent variable semantics of uncertainty expressions, their account appears to be compatible with our results. However, considering that [?] and [?] found that listeners form speaker-specific expectations rather than expecting speakers to behave like the most recent speaker they encountered, adaptation seems to be a more complex process than predicted by a priming account.

In a more recent proposal, [?] argued that at least sometimes listeners perform *prediction-by-association* when processing an utterance, that is, listeners make predictions about what the speaker would say based on the context and their experience with the speaker. This appears to be compatible with our computational adaptation model but more details need to be worked out about how such predictions operate at the implementational level [?].

In a second line of work, [?, ?] argued that partner-specific language use can be explained by an episodic memory account [?, ?, ?]. According to this account, individual linguistic events are encoded together with speaker information and the world state in memory, which results in speaker-specific linguistic representations. This account is compatible with our findings, if we assume that individual utterance-world state pairs are stored in memory together with the speaker's identity, and that some additional inference mechanism gives rise to the more complex pragmatic behavior that we observed in our experiments.

4.3.2 Implications for the semantics of uncertainty expressions

Our results also have implications for semantic theories of uncertainty expressions. The finding that listeners rapidly update their beliefs about semantic thresholds of

uncertainty expressions suggests that the semantics of these expressions is highly dynamic and context-sensitive. This is broadly compatible with theoretical accounts of probability operators [?, a subset of uncertainty expressions; e.g.,]Yalcin2010 which state that the meanings of probability operators are highly dynamic and largely determined by the context. Our results suggest that the meaning of uncertainty expressions is even more dynamic than predicted by these accounts. First, we show that this dynamicity extends to a broader set of uncertainty expressions than is typically considered (e.g., *might*), as has been recently also argued by [?]. Second, while these accounts generally assume that the main source of variability in interpretation is the probability of the event embedded under the uncertainty expressions, we find that knowledge of speaker identity also importantly contributes variability.

Dynamic and context-sensitive semantics have also been proposed for many other types of expressions. For example, [?] argued that speakers and listeners are able to compute novel senses of nouns and verbs on the fly. Similarly, in the domain of gradable adjectives such as *tall*, [?] and many others have argued that the interpretation of these adjectives crucially depends on contextual parameters. Considering the prevalence of dynamic meanings for so many other types of expression, it is therefore not surprising that the interpretation of uncertainty expressions also appears to be highly context-sensitive.

4.3.3 Methodological implications

Our results also have implications for conducting psycholinguistic experiments. First, the finding that listeners adapt to the statistics of their environment within a short experiment suggests that experimenters should be cognizant of potential adaptation effects when probing production expectations or interpretations of uncertainty expressions [?, see also]Jaeger2010.

Further, the results of Experiment 1, and in particular, the finding that participants' expectations about the use of utterances in the experiment strongly depended on the alternative utterances that we provided, highlights the need to be cautious about drawing general conclusions about expectations of use from single experiments.

For example, had we only considered the results from the *bare-might* condition (see Figure ??), we might have concluded that “might” is an expected expression to communicate an event probability of 75%, whereas if we had only considered the results from the *might-probably* condition we might have instead concluded that it is *not* an expected expression to communicate an event probability of 75%. This is where explicit modeling of the sort we have engaged in here is hugely helpful: formulating a concrete linking function which models the effects of alternatives allows for inferring the latent meanings of utterances by combining data from different experiments [?, see also]for similar approaches]Franke2014,Peloquin2016.

4.3.4 Limitations and future directions

One limitation of the present research is that the experimental paradigm is not interactive and that participants likely engaged in meta-linguistic reasoning in providing production expectation and interpretation ratings. While we tried to make the communicative situation depicted in the experiments natural, the paradigm is clearly different from everyday dialog. This limitation was necessary for the tight coupling between the experimental work and the model simulations that allowed us to investigate what kind of representations listeners update during adaption; in a more naturalistic and unconstrained setting, we would not have been able to obtain information about listener’s production expectations and about their uncertainty in both production expectations and interpretations. However, considering that our task was different from everyday interactions, investigating to what extent the results in the present research translate to less scripted and more interactive settings is an important area for future research. Employing measures like eye movements or mouse-tracking could provide insight into whether participants’ updated beliefs affect online language processing, i.e. where meta-linguistic reasoning is unlikely to occur. In this vein, mouse-tracking has recently been employed to investigate the incremental nature of adaptation in the domain of prosodic cues [?]. Both eye-tracking and mouse-tracking experiments allow for implementing more natural interpretation tasks while still providing information about participants’ uncertain beliefs via fixation patterns or cursor trajectories.

Throughout this paper, we made the assumption that listeners form *speaker-specific* production expectations. However, since all our experiments had a between-subjects design, it could be that participants were only adapting to the experimental situation, independent of the speaker. This seems unlikely given the results reported by [?], who found that participants formed speaker-specific production expectations after being exposed to multiple speakers whose use of quantifiers differed. Moreover, [?] have provided evidence of speaker-specific adaptation to uncertainty expressions. However, exactly which aspects of a situation (e.g., the speaker, the topic of conversation, the visual context, etc.) listeners adapt to is an issue that merits further investigation.

One advantage of formalizing a theory as a computational model is that the model makes concrete predictions to test in future experiments. For example, the proposed model is able to make quantitative predictions about the relation between the number of exposure trials and the size of the adaptation effect. Qualitatively, the model predicts that more exposure should lead to more adaptation, for which some evidence is reported by [?]. However, a systematic investigation of whether the model makes the correct quantitative predictions remains to be conducted.

Further, the presented adaptation model is built around the assumption that the utility of an utterance is exclusively determined by the informativeness and the cost of the utterance. However, it has been observed that other speaker goals such as being polite or convincing could also factor into the interpretation of uncertainty expressions [?, see e.g.,]Pighin2011,Juanchich2013,Holtgraves2016. It could therefore be that, for example, listeners explain away the behavior of a “confident” speaker if the context suggests that the speaker has an incentive to be encouraging or has additional goals besides being informative [?, see also]Yoon2016,Yoon2017. Investigating whether listeners draw such complex inferences could provide insight about which kind of potential speaker goals enter into listeners’ pragmatic reasoning process.

4.3.5 Conclusion

We began with the puzzle of how to reconcile the assumption of stable utterance alternatives required for pragmatic reasoning with variability in speakers' language use. The work reported here, building on much previous work on adaptation, suggests that this apparent tension is easily resolved if listeners form speaker-specific utterance expectations that they can recruit when interpreting utterances produced by that speaker.

In a series of web-based experiments, we found that after exposure to a specific speaker, listeners rapidly update their expectations about which uncertainty expressions that speaker is likely to produce to describe varying event probabilities. Moreover, these updated expectations also enter into subsequent utterance interpretation. We provided a formal account of semantic/pragmatic adaptation and modeled this behavior using a Bayesian cognitive model which assumes that (listeners reason about) speakers (who) efficiently trade off utterance informativeness and cost. Through a series of simulations we found strong evidence that semantic/pragmatic adaptation is the result of updating beliefs about both a specific speaker's underlying lexicon as well as the speaker's utterance preferences. These results both provide new insights into the cognitive processes involved in semantic/pragmatic adaptation and demonstrate that tight coupling of quantitative computational models and experimental results can shed light on unobservable representations.

Chapter 5

Speaker-specific adaptation

5.1 Introduction

Speakers exhibit considerable production variability at all levels of linguistic representation [?]e.g., Liberman1967, Weiner1983, Finegan2001. This includes variation in lexical choice to describe a world state. For example, [?] found that when asked to describe a scene with a candy bowl in which approximately half of the candies were green and half of the candies were blue, some participants judged “Some of the candies are green” to be the more appropriate utterance to describe the scene than “Many of the candies are green”, while others displayed the opposite pattern.

[?] found that participants exhibit similar production variability when describing an event with an objective event probability of 60%: Some participants judged the event to be best described with a sentence containing the uncertainty expression *might* (“You might get a blue gumball”) whereas others judged a sentence with *probably* (“You’ll probably get a blue gumball”) more appropriate.

Such variability poses a challenge to a listener who aims to know what the world is like that the speaker is describing. When confronted with two speakers who use the same expression to convey different states of the world or who use different expressions to convey the same state of the world, listeners are doomed to draw the wrong inferences about the actual state of the world unless they track how individual speakers use language. Recent work suggests that listeners deal with this kind of variability by adapting to it [?]e.g., Norris2003, Kraljic2007, Bradlow2008, Kamide2012, Kleinschmidt2015, Fine2016, Roettger2018 and that in interaction, they learn how speakers choose among alternative utterances. In the domain of quantifiers, [?] showed that listeners update their expectations about how a specific speaker uses the quantifiers *some* or *many* after being briefly exposed to a specific speaker. In line with their results, [?] found that listeners update their expectations of how a specific speaker uses the uncertainty expressions *might* and *probably* to describe different event probabilities after a brief exposure phase. Participants who were exposed to a “*confident*” speaker, who used *probably* to describe the 60% probability event, expected the use of *probably* with a wider range of probabilities; participants who were exposed to a “*cautious*” speaker, who used *might* to describe the 60% probability

event, expected the use of *might* with a wider range of probabilities.

The processes that lead listeners to update their expectations during semantic adaptation are poorly understood. In particular, it remains a largely open question to what extent listeners form speaker-specific expectations when interacting with multiple speakers. Some evidence for speaker-specific adaptation comes from the referring expressions literature. [?] found that participants exhibited a slowdown in resolving referring expressions when a confederate started referring to an object with a new expression after establishing a conceptual pact, but did not find such a slowdown when a new confederate was using a different referring expression than the original confederate.

Most closely related to our work, [?] found that listeners form speaker-specific production expectations after being exposed to two speakers who used different quantifiers to describe a scene with a candy bowl in which half of the candies were green. While this suggests that listeners should also form speaker-specific expectations about the use of uncertainty expressions, there is evidence from other linguistic domains that speaker-specific adaptation is limited to specific items. For example, [?] found that listeners adjust their phonemic representations for the fricatives /s/ and /sh/ to multiple speakers whereas listeners adjusted their phonetic representations for stop consonants such as /d/ and /t/ only to the most recent conversational partner. It could therefore be that speaker-specific adaptation in other linguistic domains is also limited to specific items and that listeners do not form speaker-specific expectations for the use of uncertainty expressions.

Further, [?] observed that the adaptation effect was considerably smaller when they exposed participants to two speakers with opposing biases as compared to only exposing participants to one speaker and comparing the adaptation effect between groups. There seem to be two likely explanations for this observation. First, it could be that due to memory limitations, listeners were unable to keep track of the exact statistics of each speaker's utterances. Since everything about the context except the speaker identity stayed constant throughout the experiment, it could be that listeners had difficulty separating their experiences with the two speakers in memory (see [?] for a similar account of memory limitations affecting audience design). Second, it could

be that listeners were tracking the statistics of the individual speakers as well as the overall statistics in the experimental situation and their post-exposure expectations were a combination of their speaker-specific expectations as well as their expectations about the situation.

In this work, we build on the recent work by [?] on adaptation to variable use of uncertainty expressions and take a first step towards investigating the nature of semantic adaptation in response to multiple speakers. In particular, we aim to answer the following two questions:

1. Do listeners form speaker-specific production expectations when they are exposed to speakers whose use of uncertainty expressions differ?
2. Do listeners form situation-specific production expectations independent of speaker identity?


In Experiment 1, we address question 1 by exposing listeners to two speakers whose use of uncertainty expressions differs. In Experiment 2, we expose listeners to two speakers whose use of uncertainty expressions is the same. We compare adaptation effect sizes across experiments to address question 2.

5.2 Experimental paradigm

In both of our experiments, we build upon the semantic adaptation paradigm used in [?], which we briefly review here. This paradigm is a classic exposure-and-test paradigm which has been used to study adaptation across several linguistic domains [?]e.g., Norris2003, Kleinschmidt2015, Yildirim2016. As shown in Figure ??, each trial shows an adult sitting behind a table with a gumball machine on it. The gumball machine is filled with orange and blue gumballs. Next to the table, there is a child who is requesting a blue or an orange gumball with the utterance “I want a blue/an orange one”. Participants are told that the gumball machine is too high up for the child to see and that only the adult can see the contents of the gumball machine.

On each exposure trial, participants watch a short video clip in which the adult responds to the child with an utterance like “You might get a blue one”. Across

Consider the following scene:



The scene shows a woman with long brown hair sitting at a wooden desk. To her right is a red gumball machine with a glass globe containing blue and orange gumballs. A cartoon girl with red hair, wearing a red dress, stands next to the machine, pointing at it. A speech bubble from the girl says "I want a blue one!".

How likely do you think is it that the woman will respond with each of the following sentences?

You'll probably get a blue one	<input type="range"/>	0
You might get a blue one	<input type="range"/>	0
<i>something else</i>	<input type="range"/>	0

Next

Figure 5.1: Example post-exposure test trial. On exposure trials the rating scales were absent, and the image of a speaker was replaced by a video of a speaker producing an utterance.

trials, the proportion of gumballs as well as the response by the adult vary.

On each test trial (Fig. ??), participants are shown a static scene in which they only see a picture of the speaker from the exposure trials. On these trials, participants are asked to provide ratings of how likely they think it is that the speaker would use the two provided utterances or some other utterance. Across trials, the proportion of blue and orange gumballs as well as the color of the gumball that the child is requesting (the target color) varies.

5.3 Experiment 1: Different speaker types

In Experiment 1, we exposed participants to two different speakers who use the uncertainty expressions *might* and *probably* differently. The primary purpose of this experiment was to test whether listeners form speaker-specific utterance choice expectations. Procedure, materials, analyses and exclusions were pre-registered on OSF (<https://osf.io/qnspg>).

5.3.1 Methods

Participants We recruited 104 participants on Amazon Mechanical Turk. Participants had to have a US-based IP address and a minimal approval rating of 95%, and they were paid \$4.75 (approximately \$12–\$15/hr).

Materials and procedure In the first part of the experiment, participants saw 40 exposure trials in two blocks. As mentioned above, each trial showed a child requesting a blue or orange gumball, a gumball machine with blue and orange gumballs, and a video of an adult male or female speaker. The speaker always produced one of the following six utterances:

- You’ll get a blue/orange one (BARE)
- You might get a blue/orange one (MIGHT)
- You’ll probably get a blue/orange one (PROBABLY)

	MIGHT		PROBABLY		BARE	
	<i>n</i>	<i>p</i>	<i>n</i>	<i>p</i>	<i>n</i>	<i>p</i>
cautious	10	60%	5	90%	5	100%
confident	5	25%	10	60%	5	100%

Table 5.1: Number of exposure trials (n) per utterance (MIGHT, PROBABLY, BARE) and associated proportion of target gumballs (p) in the cautious vs. confident speaker block. Critical trials bolded.

The number of trials with each of these utterances as well as the gumball proportions varied across the two blocks (see Table ?? for an overview). Filler trials with the bare form were included to provide evidence that the speaker is generally cooperative. One of the blocks always showed a female speaker and the other block always showed a male speaker. Both speakers were from the East Coast and native speakers of American English. The order of blocks and the speaker assignment to blocks was counterbalanced across participants.

Participants were instructed to watch what the speaker had to say to the child. The video started automatically after a 400ms delay and participants had the option to replay the video as often as they wanted. To advance, participants had to press a button which was disabled until the video had ended.

After the two exposure blocks, participants went through two test blocks. In each of the blocks they saw a picture of one of the two speakers with a gumball machine next to it, and again, a child requesting a blue or an orange gumball. On each trial, participants were asked how likely they thought it was that the adult would respond with MIGHT, PROBABLY or a blanket *something else* option. Participants indicated their expectations by distributing 100 points across these three options using sliders. In each block, participants provided ratings for scenes with 9 different gumball machines ranging from 0% to 100% blue gumballs. For each machine, participants provided four ratings in total, resulting in 36 trials per block. The order of blocks was counterbalanced such that half of the participants saw them in the same order as the exposure blocks whereas the other half saw them in opposite order.

Attention checks To verify that participants were paying attention to the video

and the scenes, we included 14 attention checks: after 14 of the exposure trials, participants were shown two different gumball machines and were asked to choose the one they saw on the previous trial.

Exclusions We excluded participants who provided correct responses to fewer than 11 attention checks. Based on this criterion, we excluded 31 participants. We further excluded participants whose utterance ratings for the different event probabilities strongly correlated ($R^2 > 0.75$) with their mean utterance ratings across all event probabilities. This suggests that they provided approximately the same ratings independent of the observed scenes and indicates that they did not pay attention. This led to one additional exclusion. None of the results discussed below depend on these exclusions.

Analysis and predictions Intuitively, a more confident speaker uses PROBABLY for a larger and MIGHT for a smaller range of gumball proportions than a more cautious speaker. Therefore, if participants track these different uses, we expect their ratings of what they think a specific speaker is likely to say to depend on how that speaker used uncertainty expressions during the exposure phase. Following [?] and [?], we quantify this prediction by fitting a spline with four knots for each expression and each participant and computing the area under the curve (AUC) for the splines corresponding to each expression, block and participant. The area under the curve is proportional to how highly and for how large of a range of gumball proportions participants rate an utterance, so if an utterance is rated highly for a larger range of gumball proportions, the AUC will also be larger. We therefore test whether listeners update their expectations by computing the difference between the AUC of the spline for MIGHT and of the spline for PROBABLY for each test block for each participant.

Based on the results of the adaptation experiment with multiple speakers by [?], we expect speaker-specific adaptation effects. We thus predict that the mean AUC difference will be bigger for the *cautious* speaker test blocks than for the *confident* speaker test blocks.

As a secondary analysis, we also investigate whether the order of exposure blocks (*confident* or *cautious* first), the assignment of speaker to speaker type (whether the male speaker was the *cautious* speaker or vice versa), or the order of the test blocks

(same as exposure or reverse) has an effect on adaptation. We do not expect any of these factors to have an effect on adaptation.

5.3.2 Results and discussion

Figure ?? shows the mean utterance ratings of participants grouped by the two post-exposure test blocks. As this plot shows, participants expected the *confident* speaker to be more likely to use *probably* for lower event probabilities than the *cautious* speaker. This is also reflected in the AUC differences between the splines for MIGHT and of the splines for PROBABLY: As predicted, this difference was greater for the *cautious* speaker ratings than for the *confident* speaker ratings ($t(142) = 2.92, p < 0.01$).

For our secondary analysis, we fit a linear regression model to predict the AUC difference with speaker type, exposure block order, speaker assignment, and test block order as predictors. Only speaker type is a significant predictor in this model (exposure block order: $\beta = 5.72, t(139) = 1.30, n.s.$; speaker assignment: $\beta = 1.21, t(139) = 0.28, n.s.$; test block order: $\beta = 2.28, t(139) = 0.52, n.s.$). Further, a model that includes these four predictors does not explain significantly more variance than a model that only includes speaker type as a predictor ($F(3, 139) = 0.67, n.s.$).

The results of this experiment suggest that listeners form speaker-specific expectations of how different speakers use uncertainty expressions after brief exposure. At the same time, the results provide concrete evidence against two other accounts. First, they provide evidence against an account according to which participants only adapt to the experimental situation: If participants had only updated their expectations of what a generic speaker would say in the scenes presented in the experiment, we would not have expected to see differences in ratings between speakers. Second, they also provide evidence against a pure priming account according to which listeners update their expectations to the most recent exposure. Note that the adaptation effect was independent of the order of presentation and the order of test blocks. If participants had been primed by the most recent exposure speaker, we would have expected that participants' post-exposure ratings were primarily influenced by the behavior of the second exposure speaker.

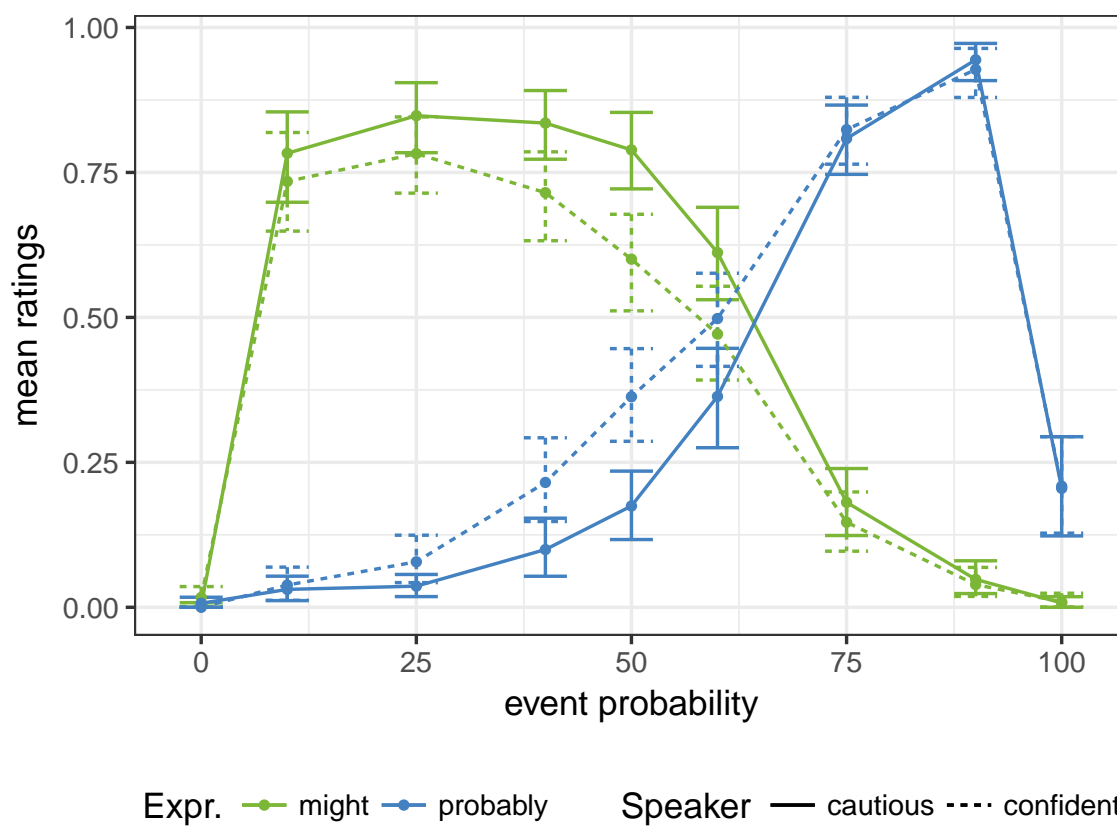


Figure 5.2: Mean utterance ratings for scenes with different event probabilities in Experiment 1. Error bars indicate bootstrapped 95% confidence intervals.

The results of this experiment also replicate the finding by [?] of differing effect sizes between the single-speaker and two-speaker experiments: The adaptation effect was considerably smaller in this two-speaker experiment (Cohen’s d : 0.486) than in the single-speaker adaptation experiment by [?] (Cohen’s d : 1.263).

As suggested by a reviewer, one reason for the smaller effect size in the two-speaker experiment could be some form of self-priming and that participants’ responses in the first test block influenced their responses in the second block. We evaluated this hypothesis in a post-hoc analysis of the responses from the first test block. We compared the responses of participants who were first tested on the *cautious* speaker to the responses of participants who were first tested on the *confident* speaker. If responses in the first test block influenced responses in the second test block, we would expect a larger effect size if we only consider the data from the first block. We did indeed find a larger effect size in the first block (Cohen’s d : 0.723), which suggests that participants exhibited some form of self-priming.

However, even if we only consider the first block of responses, the adaptation effect remains smaller in the two-speaker experiment (Cohen’s d : 0.723) than in the one-speaker experiment (Cohen’s d : 1.263). This could be either a result of memory limitations or a result of listeners jointly tracking the statistics of each speaker as well as of the overall experimental situation (situation-specific statistics). We further investigate these possibilities in the next experiment.

5.4 Experiment 2: Identical speaker types

In Exp. 1, we found that the adaptation effect was smaller than it was in the single-speaker version of the experiment, which could have either been a result of memory limitations or joint speaker-specific and situation-specific adaptation. In this experiment, we investigate whether there is evidence for one of these two accounts. We exposed listeners to two speakers of the same type.¹ If the smaller effect in Exp. 1

¹ In the spirit of open science, we note that the data from this experiment comes from a faulty version of Experiment 1. A scripting error led to participants always being exposed to the same speaker type instead of two different speaker types. Because of this error, the pre-registered analysis (<https://osf.io/3cw79>) deviates from the analysis that we report here. The reported analyses here

was caused by listeners’ inability to separate their experiences with the two speakers in memory, i.e, some experiences might have been attributed to the incorrect speaker, we would expect the adaptation effect in this experiment to be on average the same as in the one-speaker experiment. This is because even if listeners cannot perfectly separate their experiences with each speaker, they would on average still have the same number of experiences with each of the two speakers as listeners had with the one speaker in the single-speaker experiment. If, on the other hand, the smaller effect in the previous experiment was a result of listeners jointly tracking speaker-specific and situation-specific statistics, we would expect the adaptation effect to be larger here than in the single-speaker experiment. This is based on the assumption that more exposures lead to a larger adaptation effect and thus listeners’ should adapt more to the situation if they are exposed to two speakers and hence also twice the number of interactions.

5.4.1 Methods

Participants We recruited 104 participants on Amazon Mechanical Turk. Participants had to have a US-based IP address and a minimal approval rating of 95%, and they were paid \$5 (approximately \$12–\$15/hr).

Materials and procedure The materials and procedures were the same as in Exp. 1 except for the following two modifications. First, the speaker types for each participant were identical across the two exposure blocks: both speakers were either *confident* or *cautious* speakers. Second, the number of trials with PROBABLY and the number of trials with MIGHT were the same (10 trials per utterance and block) whereas in Experiment 1, the *confident* speaker produced only 5 instances of MIGHT and the *cautious* speaker produced only 5 instances of PROBABLY.² Assignment of speaker types was counterbalanced across participants, which means this experiment had a between-subjects manipulation.

are the only additional analyses we performed on the data. The reason for not discarding the data from this experiment but rather including it here is that it provides an informative data point for the question of whether listeners track situation-specific expectations.

²The reason for the second modification is the above mentioned scripting error. See below for a discussion of potential implications.

As in Experiment 1, we excluded participants who provided correct responses to less than 11 of the attention checks as well as participants who seemed to provide random responses as defined above. In total, we excluded 11 participants because of the attention check criterion and 1 more participant because of random responses.

Analysis and predictions As the primary analysis, we compare the AUC differences between the splines for MIGHT and of the splines for PROBABLY between participants in the two conditions. Analogous to Experiment 1, we predict that the mean AUC difference will be bigger in the *cautious* speaker condition than in the *confident* speaker condition.

We again also investigate whether the assignment of speaker to speaker type or the order of the test blocks have an effect on the AUC difference. We do not expect either of these factors to affect adaptation.

Lastly, we compute the effect size measured by Cohen’s d . As explained above, we expect the effect size either to be the same as in the single-speaker experiment or to be larger.

5.4.2 Results and discussion

Figure ?? shows the mean utterance ratings of participants for the two conditions. We again observe listener adaptation, resulting in a greater AUC difference in the *cautious* speaker condition than in the *confident* speaker condition ($t(89) = 8.01, p < 0.001$). Further, no factors other than speaker type are significant predictors of the AUC difference (speaker assignment: $\beta = -1.32$, $t(87) = -0.398$, *n.s.*; test block order: $\beta = 4.28$, $t(139) = 1.30$, *n.s.*).

Lastly, the effect size (Cohen’s d : 1.68) was larger in this experiment than in Experiment 1 and the single-speaker experiment by [?]. While it would be premature to definitively conclude from these three experiments that listeners’ expectations are jointly influenced by individual speaker’s productions as well as all the productions in the experiment, our results point in this direction.

There is a potential confound in this experiment because participants saw 5 additional filler trials during each exposure block which could have led to the larger effect

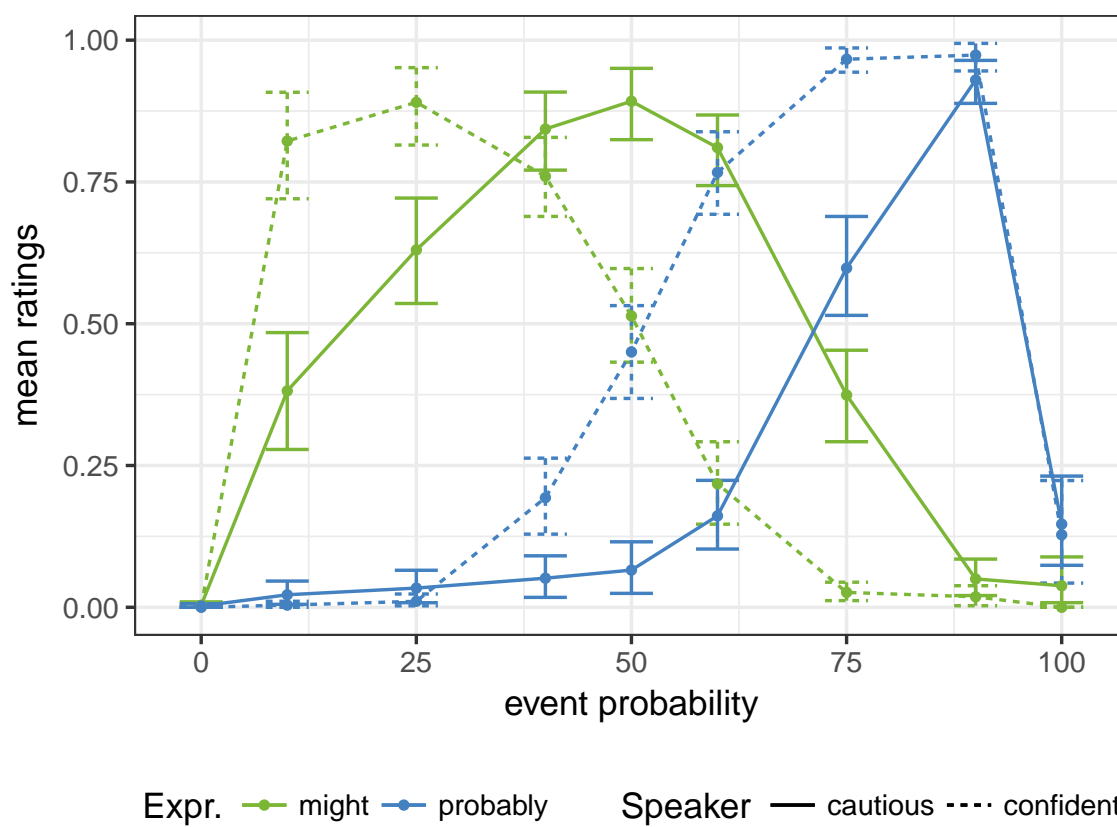


Figure 5.3: Mean utterance ratings for scenes with different event probabilities in Experiment 2. Error bars indicate bootstrapped 95% confidence intervals.

size as compared to the single-speaker experiment. However, this explanation seems unlikely considering previous work.³

5.5 General discussion and conclusion

In two experiments, we found that listeners form speaker-specific production expectations of uncertainty expressions after brief exposure to two speakers. This shows that the results by [?] also extend to lexical items other than quantifiers.

At the same time, however, we found that the adaptation effect size varied depending on whether the two speakers had the same or divergent bias during the exposure phase. When listeners were exposed to two different speaker types, the adaptation effect was smaller and their expectations seemed to have been shaped by their experiences with the two speakers as well as all the experiences encountered in the experiment. When both speakers behaved the same, on the other hand, the adaptation effect was much more pronounced and even greater than in the single-speaker experiment from previous work.

One likely explanation for these observations is that apart from tracking speaker-specific statistics, listeners also track the situation-specific statistics of all interactions in the experiment and their expectations are guided by both of these factors. In the case of speakers with different uses of uncertainty expressions, speaker-specific adaptation is attenuated since the overall statistics guide listeners towards an “average” speaker whose use falls somewhere in between the *cautious* and the *confident* speaker. When listeners are exposed to two speakers of the same type, on the other hand, the situation-specific statistics reinforce the speaker-specific statistics and hence listeners adapt more to the two speakers.

³[?] used a very similar paradigm to study semantic adaptation to the use of the quantifiers *some* and *many*. Analogous to our *confident* and *cautious* speakers, they had a *some-biased* and a *many-biased* speaker. They report two versions of their experiment: one in which there were no filler trials with the other quantifier and another version in which there was a balanced number of exposure trials with both quantifiers in both conditions. They found that the adaptation effect was smaller when there were more filler trials, so we would expect that if the additional fillers affected the size of the adaptation effect, the effect would be even larger had we not presented the extra fillers to participants.

An account based on “faulty” memory, according to which listeners have trouble keeping the speaker-specific experiences separate, does not predict the larger adaptation effect when listeners are exposed to two speakers of the same type. If every experience is encoded as an episode in memory but some with the incorrect speaker information, on average, the number of experiences with each speaker should still be the same as in the one-speaker condition and therefore it is unclear why listeners adapt more in the two-speaker experiment than in the one-speaker experiment.

Our findings also have implications for current models of semantic adaptation. Following the recent successes in modeling phonetic adaptation as an instance of Bayesian belief updating [?], [?] propose a computational model of semantic adaptation. According to this model, when interacting with a speaker Sp , listeners update their beliefs about a set of speaker-specific parameters Θ_{Sp} , which govern the speaker’s lexicon and preferences.⁴ Their model predicted the results of the single-speaker experiment well, but without modifications, it does not predict the differences in effect size.

We consider two promising extensions of this model. First, the model could be cast as a hierarchical model. Hierarchical models have been argued to explain many cognitive and perceptual phenomena [?]see e.g., [for a review]Clark2013, including phonetic adaptation [?], and also seem applicable here. In a hierarchical version of the adaptation model, we would assume that the speaker-specific parameters Θ_{Sp} are not only shaped by the listener’s prior beliefs and the observed interactions by a speaker Sp but rather also depend on a distribution reflecting the situation-specific expectations. Figure ?? shows a sketch of a potential hierarchical model. Such a model would explain the differences in effect size: When listeners are exposed to different speaker types, the situation-specific parameter distribution would be influenced by two speaker types that essentially cancel each other out, which in turn would lead to less extreme speaker-specific distributions. On the other hand, when both of the speakers are of the same type, the situation-specific parameter distribution would be more strongly shifted towards the observed distributions which in turn would lead to more extreme speaker-specific distributions.

⁴See also [?] for a similar model of the formation of conceptual pacts.

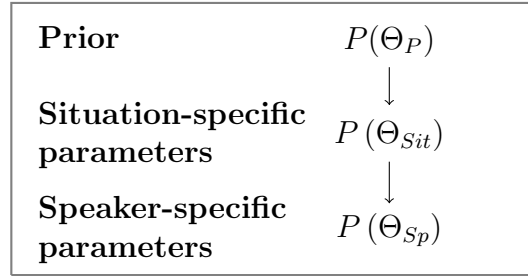


Figure 5.4: Hierarchical model of semantic adaptation. Situation-specific parameters $P(\Theta_{Sit})$ depend on prior beliefs $P(\Theta_P)$ and speaker-specific parameters $P(\Theta_{Sp})$ depend on the situation-specific parameters.

A second possibility would be to cast the model as a mixture model in which overall production parameters are a weighted combination of situation-specific and speaker-specific parameters (and potentially other factors). Figure ?? shows a sketch of a potential mixture model. According to such a model, listeners would form both situation-specific and speaker-specific expectations as a result of adaptation and then combine these expectations to their overall expectations. Such a model would also predict the smaller effect size in Experiment 1 since it would predict that the overall production expectations are influenced by the speaker-specific statistics as well as the situation-specific statistics and the latter drive the production expectations to be more similar to an “average” speaker. When listeners are exposed to two identical speakers, on the other hand, the situation-specific expectations (which are in line with the speaker type of both exposure speakers) would reinforce the speaker-specific expectations and therefore lead to a larger adaptation effect. Future experimental work should adjudicate between the hierarchical and the mixture model account.

In conclusion, we presented new experimental results from the domain of uncertainty expressions which suggest that speaker-specific semantic adaptation is a product of forming speaker-specific expectations and forming expectations about the situation independent of the speaker. These results raise a number of interesting questions, most pressingly regarding transfer effects to novel speakers, which have been observed in other linguistic domains [?]e.g., Bradlow2008, Xie2018. In our experiments, the exposure and test speakers did not differ. This raises the question

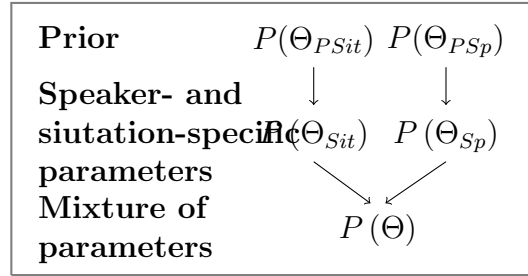


Figure 5.5: Mixture model of semantic adaptation. Overall production parameters $P(\Theta)$ are a weighted combination of situation-specific parameters $P(\Theta_{Sit})$ and speaker-specific parameters $P(\Theta_{Sp})$.

about whether and to what extent updated expectations transfer to novel speakers whose similarity to the exposure speaker(s) varies. Both models sketched above lend themselves well to capturing such transfer effects. In addition, participants saw very similar visual scenes on each trial. Another potential direction would be to study the extent of speaker-specific adaptation when listeners encounter more novel scenes during the test phase to investigate to what extent listeners form speaker-specific expectations independent of other contextual factors. Answering these questions will help disentangle the different adaptation processes and give us a better understanding of how listeners infer meanings in context.

Chapter 6

Explaining away

6.1 Introduction

Speakers exhibit production variability at all levels of linguistic representation. Listeners deal with this variability by adapting to speakers and forming speaker-specific expectations about language use [?]e.g., Norris2003, Kraljic2005, Bradlow2008, Kurumada2012, Kamio2013. For example, at the lexical level, [?] found that participants were uncertain whether a generic speaker would use the utterance *Some of the candies are green* or the utterance *Many of the candies are green* to describe a bowl in which there were approximately equal numbers of green and blue candies. However, if participants briefly observed a speaker describing this scene either consistently using *some* or consistently using *many*, they updated their expectations about how that speaker would use quantifiers to describe different proportions. Similarly, [?, henceforth S&D] found that listeners update their expectations about a speaker’s use of uncertainty expressions like *might* and *probably* after brief exposure to that speaker.

While previous work shows that listeners *can* adapt to variable language use, the contextual conditions and limits on adaptation are under-explored. In principle, it is possible that listeners might update their production expectations regardless of the context of utterance. Alternatively, listeners’ production expectations, and consequently their adaptive behavior, might be modulated by non-linguistic contextual factors. To illustrate this, consider a speaker *S* who is in a very good mood and wants to be encouraging. If *S* tells a listener *L* “you’ll *probably* win the sweepstake” when there is only a 60% chance of winning, *L* may consider *S*’s use of *probably* instead of a weaker alternative such as *might* to be the result of *S*’s mood. Consequently, *L* would not necessarily expect *S* to use *probably* to describe the same event probability when *S* is in a worse or more discouraging mood.

Recent computational models suggest that adaptation may be modulated by contextual factors. S&D proposed a computational model of adaptation to variable use of uncertainty expressions based on Bayesian belief updating. According to this model, when interacting with a speaker and observing their language use, listeners integrate their prior beliefs about the speaker’s semantic representations and lexical preferences

Consider the following scene:

I'd like a window seat...

Available seats:

Occupied
Aisle seat
Window seat

How likely do you think it is that the representative will respond with each of the following sentences?

You'll probably get a window seat 0

You might get a window seat 0

something else 0

Next

Figure 6.1: Example trial from Experiment 1 and the post-exposure blocks from Experiments 2 and 3.

with the observed utterances to arrive at updated speaker-specific production expectations. Similar Bayesian belief-updating models have been proposed for adaptation in other linguistic domains, including in phonetic adaptation [?], syntactic adaptation [?], and prosodic adaptation [?]. All of these models predict that the extent to which listeners adapt depends on how they initially expect a speaker to use language. Consequently, contextual factors that affect listeners’ expectations about language use should also affect how much listeners adapt to specific speakers. If, given contextual information, a speaker’s behavior matches prior expectations, there is no need to adapt.

In addition to the model-predicted influence of contextual factors on adaptation,

there is empirical evidence from phonetic adaptation: [?] used a lexical retuning paradigm to investigate adaptation to a speaker who produced a sound that was ambiguous between /s/ and /sh/. They found that without additional information listeners adapted, such that their perceptual boundary between /s/ and /sh/ shifted. However, when participants were shown a picture of the speaker with a pencil in their mouth, they explained away the observed signal as a pencil-distorted /sh/-sounding /s/ rather than as an intentionally produced /sh/-sounding /s/. Consequently, they did not adapt, i.e., their perceptual boundary between /s/ and /sh/ did not shift.

In this work, we investigate for the first time whether listeners explain away otherwise unexpected behavior at the lexical level if they are presented with contextual information that provides a reason for a speaker's productions. Concretely, we investigate whether one contextual factor – the speaker's mood – provides such a reason for otherwise less expected uses of the uncertainty expressions *might* and *probably* (EXPLAINING AWAY HYPOTHESIS). However, considering that previous experimental studies on semantic adaptation [?, ?] kept all aspects of the context constant between the exposure and test phase, it could also be that listeners simply learn associations between the use of uncertainty expressions and speakers, independent of other contextual information (ASSOCIATIVE HYPOTHESIS).

We investigate this issue as follows. We first establish that language users have different expectations about a generic speaker's use of uncertainty expressions depending on their beliefs about the speaker's mood (Exp. 1). We then investigate how much participants adapt when they are provided with information about the speaker's mood that makes their use of uncertainty expressions more expected, and compare participants' adaptation behavior to a neutral adaptation setting in which participants do not receive any information about the speaker's mood (Exp. 2). Finally, we investigate the relationship between adaptation and highly unexpected behavior by exposing participants to a speaker whose use of uncertainty expressions is incongruent with their mood (Exp. 3). We find that listeners adapt less when they are presented with a reason for the speaker's behavior. However, surprisingly, we also find that listeners do not adapt more when the behavior is highly unexpected given contextual information, potentially suggesting that there are limits on adaptation.

6.2 Experiment 1: Effect of speaker mood

In Exp. 1, we investigated how one contextual factor, the speaker’s mood, affects listeners’ expectations about a speaker’s use of the uncertainty expressions *might* and *probably* for a range of event probabilities. The choice to manipulate the speaker’s mood was guided by the intuition that listeners expect a speaker in a good mood to use uncertainty expressions differently from a speaker in a bad mood. Moreover, mood is a non-inherent property of speakers that can change over time, which is important for Exps. 2 and 3.

Procedure, materials, analyses, exclusions and predictions were pre-registered on OSF (<http://osf.io/anonymized>).

6.2.1 Methods

Participants

We recruited 60 participants (20 per condition) from Amazon’s Mechanical Turk.

Materials and procedure

At the beginning of the experiment, participants were introduced to an airline representative. Depending on the condition, the instructions explained that the representative was having a particularly bad day and feeling pessimistic and angry (*pessimist* condition); that she was having a particular great day and feeling optimistic and helpful (*optimist* condition); or that she was having a normal day (*neutral* condition). In addition to the textual mood information, the drawing of the representative showed her with an angry face (*pessimist*), a big smile (*optimist*), or a neutral facial expression (*neutral*).

Participants were then instructed that they would see scenes in which a customer of a cheap airline, who had the choice between getting a seat assigned at random or paying \$50 to pick their seat, would ask the representative about their possible seat assignment, to determine the likelihood of getting their preferred seat without paying. As shown in Fig. ??, participants could see the seat map and thus determine the

number of available window and aisle seats and estimate the probability of getting the preferred seat. On each trial, participants had to indicate how likely they considered the representative to respond with one of the following two utterances:

- You might get a window seat/an aisle seat. (MIGHT)
- You'll probably get a window seat/an aisle seat. (PROBABLY)

Participants indicated their production expectations by distributing 100 points across these two utterances using a slider. If they thought that neither of the two utterances were likely responses, they could assign points to a blanket *something else* option. Participants completed 36 trials: they provided 4 ratings for each of 9 different probabilities of getting a preferred seat, ranging from 0% to 100% as indicated by the seat map. Trials were counterbalanced on whether the customer asked for a window or an aisle seat and trial order was randomized.

Analysis and exclusions

Following S&D, we quantified the production expectations for MIGHT and PROBABLY by fitting a spline with 4 knots for each participant and expression and computing the area under the curve (AUC) for each of these splines. A larger AUC indicates that an expression was rated highly for a larger range of event probabilities. To compare production expectations across conditions, we computed the difference in AUC between MIGHT and PROBABLY and compared the average difference across participants in the two conditions.

We excluded participants who provided random responses. Concretely, we excluded participants whose ratings for different event probabilities highly correlated ($r > 0.75$) with their average rating, suggesting that they always provided approximately the same rating independent of the event probability. Based on this criterion, we excluded 7 participants (*optimist*: 1, *pessimist*: 3, and *neutral*: 3).

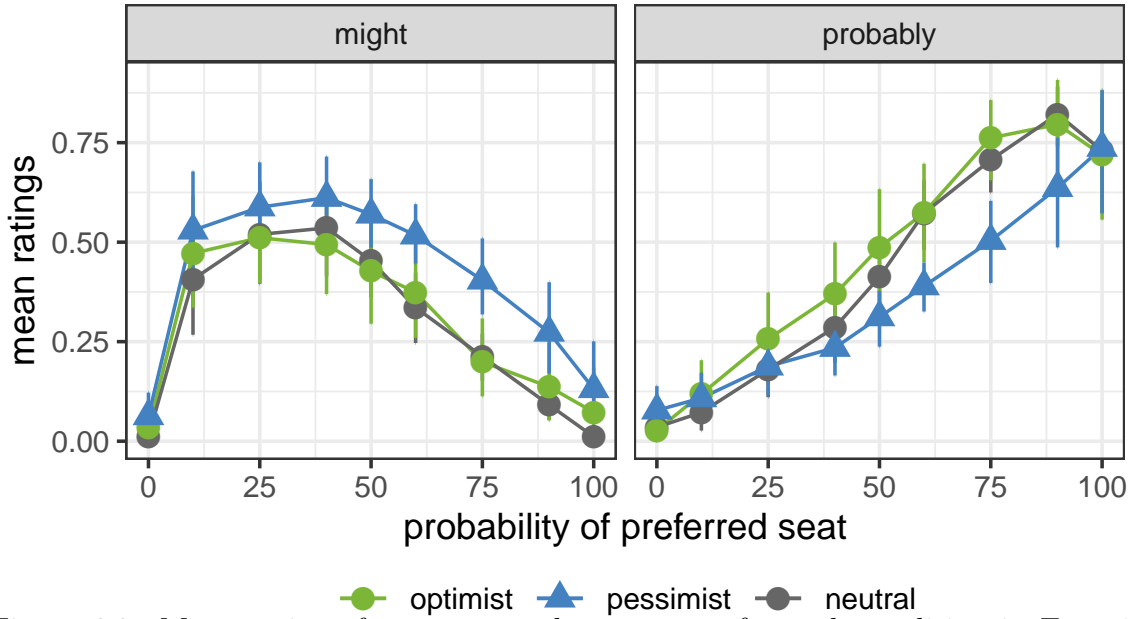


Figure 6.2: Mean ratings for MIGHT and PROBABLY for each condition in Exp. 1. Error bars correspond to bootstrapped 95%-confidence intervals.

Predictions

We predicted that participants expect the *optimistic* speaker to be encouraging and therefore to use PROBABLY for a wider range of event probabilities than the *pessimistic* speaker. Conversely, we predicted that participants expect the *pessimistic* speaker to use MIGHT for a wider range of probabilities than the *optimistic* speaker. This prediction should be reflected in larger AUC differences in the *pessimist* condition than in the *optimist* condition. Since it was unclear what mood participants attributed to the *neutral* speaker, we only predicted that the mean AUC difference for this third condition should lie between the mean AUC differences of the *optimist* and *pessimist* conditions, with the possibility of being equal to one of the two conditions.

6.2.2 Results and discussion

Fig. ?? shows participants' mean ratings for MIGHT and PROBABLY across the three conditions. As expected, we observe ratings close to 0 for both utterances when there

is a 0% chance of getting the preferred seat (where there is a preference for the *something else* option, not shown), high ratings for MIGHT for low event probabilities, and high ratings for PROBABLY for high event probabilities. At the same time however, there are also differences across conditions. As the left panel shows, MIGHT was rated higher in the *pessimist* condition than in the *optimist* condition for a large range of event probabilities; as the right panel shows, the opposite was true for PROBABLY. Ratings in the *neutral* condition were almost identical to the ratings in the *optimist* condition. All these observations were also reflected in the AUC differences: The AUC differences in the *pessimist* condition were greater than in the *optimist* condition ($t(34) = 2.51, p < 0.05$). The AUC differences in the *neutral* condition – while numerically slightly larger – were not significantly different from the differences in the *optimist* condition ($t(34) = 0.35, p > 0.7$).

These results provide evidence that listeners have mood-dependent expectations about a generic speaker’s use of uncertainty expressions. In Exp. 2 we investigate whether speaker mood affects the extent to which listeners adapt to that speaker’s use of uncertainty expressions.

6.3 Experiment 2: Explaining away

In an exposure-and-test paradigm, we investigated whether adaptation to a specific speaker is modulated by knowledge about the speaker’s mood. We follow S&D and either exposed participants to a speaker who always used MIGHT to describe an event probability of 60% (henceforth a “*cautious*” speaker) or a speaker who always used PROBABLY to describe an event probability of 60% (a “*confident*” speaker). Based on the results of Exp. 1, the behavior of a *cautious* speaker is more expected of a speaker who is having a bad day, and the behavior of a *confident* speaker is more expected of a speaker who is having a good day. We thus hypothesized that participants’ beliefs about the speaker’s mood influence how much they adapt: If the speaker’s behavior is mood-congruent, we expected participants to experience a weaker expectation violation and adapt less than in the neutral conditions.

Procedure, materials, analyses, exclusions and predictions were pre-registered on

Condition Mood	<i>pessimist</i> bad	<i>cautious</i> neutral	<i>confident</i> neutral	<i>optimist</i> good
$p = 25\%$	–		MIGHT x5	
$p = 60\%$	MIGHT x5		PROBABLY x5	
$p = 90\%$	PROBABLY x5		–	
$p = 100\%$	BARE x3		BARE x3	

Table 6.1: Overview of exposure utterances in Exp. 2. p indicates the proportion of preferred available seats shown on the seat map while the speaker produced the utterance. Critical trials are highlighted in gray.

OSF (<http://osf.io/anonymized>).

6.3.1 Methods

Participants

We recruited 320 participants (80 per condition) from Amazon’s Mechanical Turk.

Materials and procedure

The first block of the experiment consisted of an exposure phase with 13 trials (5 critical, 8 filler). On each trial, participants first saw a scene in which a customer asked about a specific seat and a seat map which indicated the number of available window and aisle seats (see top part of Fig. ??). To make sure participants paid attention to the seat map, they were then asked to rate how likely the customer was to get the preferred seat. They then listened to a pre-recorded response from the airline representative. Exposure trials were identical across the *pessimist* and *cautious speaker* conditions and identical across the *optimist* and *confident speaker* conditions but differed across these two pairs of conditions (see Table ?? for an overview): In the *pessimist* and *cautious speaker* condition, there were 5 critical trials in which the representative described a 60% probability of getting the preferred seat with “You might get one” (MIGHT); in the *optimist* and *confident speaker* conditions, the speaker responded with “You’ll probably get one” (PROBABLY). 5 filler trials in the *pessimist/cautious speaker* conditions consisted of *probably* responses when there was

a 90% preferred seat probability, and 5 filler trials in the *optimist/confident speaker* conditions combined *might* with a 25% preferred seat probability. Finally, 3 additional filler trials in all four conditions consisted of the response “You’ll get one” (BARE) when it was 100% likely for the customer to get their preferred seat. Filler trials were intended to boost credibility of the speaker.

The exposure block was followed by another instruction, informing participants in all conditions that it was a week later and that the airline representative was having a normal day, followed by another manipulation check asking participants to rate how they thought the representative was feeling.

The last block of the experiment was identical to the trials in Exp. 1: participants completed 36 trials and rated how likely they thought it was that the speaker produced MIGHT, PROBABLY or *something else* for 9 different preferred seat probabilities.

Analysis and exclusions

We computed the AUC difference between the splines for MIGHT and PROBABLY for each participant as in Exp. 1. We again excluded data from participants providing random responses, which led to 52 exclusions (*pessimist*: 9, *cautious*: 14, *optimist*: 18, *confident*: 11).

Predictions

We expected that participants would adapt to the use of uncertainty expressions by the different speakers and update their expectations. Further, in line with the EXPLAINING AWAY HYPOTHESIS, participants in the *pessimist* and *optimist* conditions, who should experience less of an expectation violation, should adapt less than participants in the other two conditions. In terms of the AUC difference ($\text{AUC}(\text{might}) - \text{AUC}(\text{probably})$), we therefore expected the following ordering: *cautious speaker* < *pessimist* \leq *optimist* < *confident speaker*.

However, in Exp. 1, we also found that the ratings in the *neutral* condition did not significantly differ from the ratings in the *optimist* condition. This suggests that

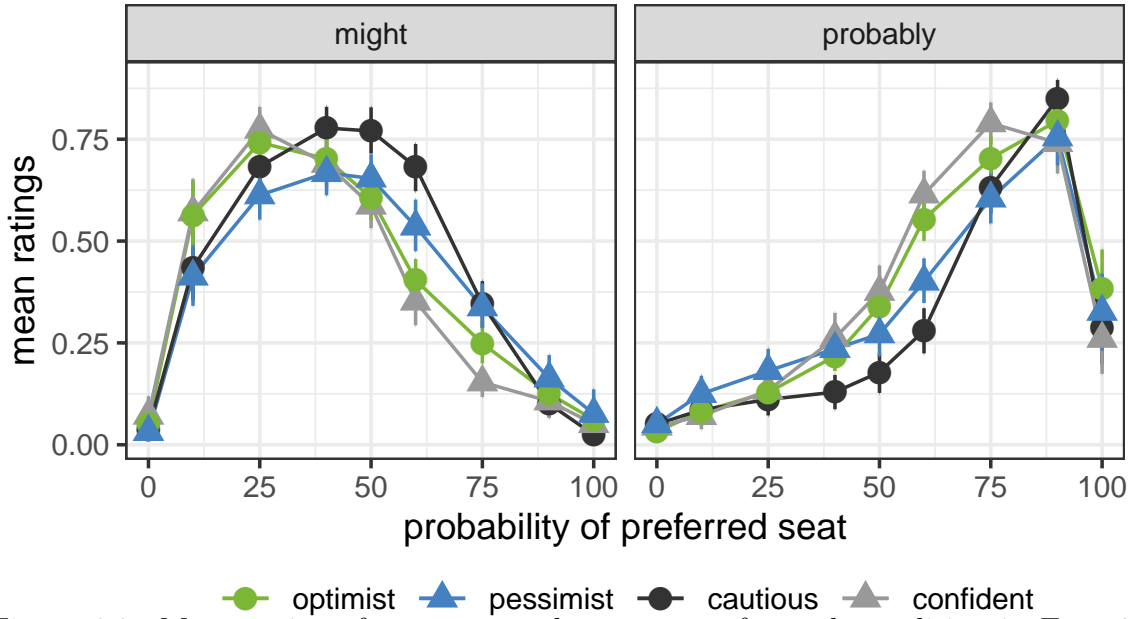


Figure 6.3: Mean ratings for MIGHT and PROBABLY for each condition in Exp. 2. Error bars correspond to bootstrapped 95%-confidence intervals.

listeners' initial expectations about the speaker's mood and the associated production expectations only slightly differ across the *optimist* and the *confident speaker* conditions and therefore we also expected similar adaptation behavior in these two conditions.¹ We therefore, while expecting the numerical ordering described above, expected and pre-registered only significant differences between the *pessimist* and *confident speaker* conditions, and between the *cautious speaker* and *confident speaker* conditions.

If listeners' adaptation behavior is not affected by contextual information, in accordance with the ASSOCIATIVE HYPOTHESIS, there should be no difference between the *cautious speaker* and *pessimist* conditions and no difference between the *confident speaker* and *optimist* conditions.

¹This intuition was further confirmed in a pilot study with 10 participants per condition, which we conducted prior to pre-registration. In the pilot, we found the expected ordering for the *cautious speaker*, *pessimist*, and *confident speaker* conditions but the difference between the *optimist* and *confident speaker* condition was so small that we would have needed more than 205 participants per condition to achieve power of $\beta = 0.8$.

6.3.2 Results and discussion

Fig. ?? shows the mean ratings for MIGHT and PROBABLY for the four conditions. The results are consistent with the predictions according to the EXPLAINING AWAY HYPOTHESIS: First, participants in the *confident speaker* condition rated PROBABLY higher for a larger range of event probabilities than in the *cautious speaker* condition and the opposite was true for MIGHT. This pattern is also reflected in the mean AUC difference, which is larger in the *cautious speaker* condition than in the *confident speaker* condition ($t(133) = 5.18, p < 0.001$). This result replicates the adaptation effect found by S&D and suggests our seat map paradigm is suited for studying adaptation in the use of uncertainty expressions.

Second, we also observe differences between the *cautious speaker* and *pessimist* conditions. The mean AUC difference is larger in the *cautious speaker* condition than in the *pessimist* condition ($t(135) = 2.38, p < 0.02$).

Third, we also observe a numeric difference between the *confident speaker* and *optimist* conditions. Numerically, the AUC difference is larger in the *optimist* condition than in the *confident speaker* condition, but not significantly so ($t(129) = 1.61, p = 0.11$).

Lastly, as shown in Fig. ??, participants in the *optimist* and *pessimist* condition updated their beliefs about the mood after we instructed them that the speaker was now in a normal mood, suggesting that this instruction was sufficient to update participants' beliefs about the speaker's mood. As expected, participants in the two neutral conditions did not change their beliefs about the speaker's mood.

The results from this experiment provide evidence for listeners explaining away otherwise unexpected productions if they are presented with a reason for the speaker's behavior, and are predicted by the EXPLAINING AWAY account.

However, with additional stipulations, these results are also compatible with the ASSOCIATIVE account. One aspect of the context, the speaker's mood, changed between the exposure block and the test block in the *pessimist* and *optimist* conditions but not in the other two conditions. Therefore, it could be that this difference between the blocks leads to weaker associations between utterances and the context and therefore we observe less adaptation in the *pessimist* and *optimist* conditions. To

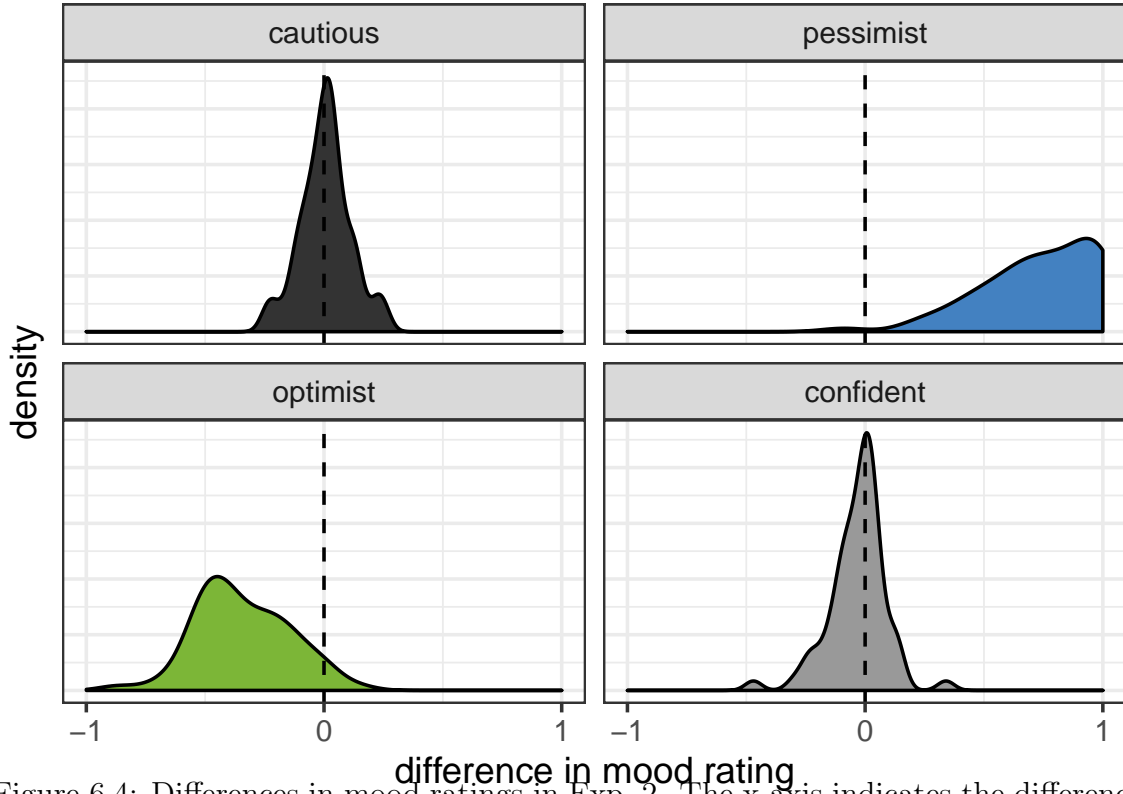


Figure 6.4: Differences in mood ratings in Exp. 2. The x-axis indicates the difference between the mood rating before the exposure block and the mood rating before the test block.

evaluate this possibility, we conducted Exp. 3.

6.4 Experiment 3: Incongruent conditions

In Exp. 3, we investigated participants' adaptation behavior when the speaker's use of uncertainty expressions was incongruent with the information about the speaker's mood, i.e., a speaker in a bad mood using uncertainty expressions like the *confident speaker* in Exp. 2, or a speaker in a good mood behaving like the *cautious speaker*. According to the EXPLAINING AWAY account, this should lead listeners to experience a stronger expectation violation than in the neutral and congruent conditions in the previous experiment and therefore listeners should adapt more. According to the ASSOCIATIVE account, on the other hand, listeners should adapt less than in the

Condition Mood	<i>pessimist incongr.</i> bad	<i>optimist incongr.</i> good
$p = 25\%$	MIGHT x5	–
$p = 60\%$	PROBABLY x5	MIGHT x5
$p = 90\%$	–	PROBABLY
$p = 100\%$	BARE x3	BARE x3

Table 6.2: Overview of exposure utterances in Exp. 3. p indicates the proportion of preferred available seats shown on the seat map while the speaker produced the utterance. Critical trials highlighted in gray.

neutral conditions because according to this account, the smaller adaptation effect that we found in the *optimist* and *pessimist* conditions in the previous experiment was caused by a difference between the exposure phase and the test phase, which is still present in this experiment.

6.4.1 Methods

Participants

We recruited 160 participants (80 per condition) from Amazon’s Mechanical Turk.

Materials and procedure

The procedure was identical as in Exp. 2. There were two conditions: *optimist incongruent* and *pessimist incongruent*. The *optimist incongruent* condition showed a speaker in a good mood who produced the same utterances as the *pessimist* and *cautious* speaker from the previous experiment. The *pessimist incongruent* condition showed a speaker in a bad mood who produced the same utterances as the *optimist* and *confident* speakers in Exp. 2. See Table ?? for an overview of the exposure trials.

Analysis and exclusions

Analyses and exclusions were identical to the ones of Exp. 2. We excluded 27 participants (*optimist incongruent*: 15, *pessimist incongruent*: 12).

Predictions

We predicted that participants would adapt to different uses of uncertainty expressions: We expected the AUC difference in the *optimist incongruent* condition to be larger than in the *pessimist incongruent* condition. We further predicted that listeners experience a stronger expectation violation than in the neutral conditions in Exp. 2. We therefore also predicted the AUC difference in the *optimist incongruent* condition to be larger than in the *cautious speaker* condition, and the AUC difference in the *pessimist incongruent* condition to be smaller than in the *confident speaker* condition.

6.4.2 Results and discussion

Fig. ?? shows the mean ratings for MIGHT and PROBABLY for the two conditions in this experiment as well as the mean ratings from the neutral conditions from Exp. 2. As this plot shows, participants adapted to the different uses of uncertainty expressions. The mean AUC difference in the *optimist incongruent* condition was larger than in the *pessimist incongruent* condition ($t(131) = 5.90, p < 0.001$). However, unexpectedly, participants did not adapt more in the incongruent conditions than in the neutral conditions. The mean AUC difference in the *optimist incongruent* condition was not larger than in *cautious* condition ($t(129) = 0.004, p = 0.99$), and the mean AUC difference in the *pessimist incongruent* condition was not significantly smaller than in the *confident* condition ($t(135) = -1.18, p = 0.24$).

In this experiment, we again replicated the adaptation effect. However, we did not find a significantly stronger adaptation effect across the two incongruent conditions in this experiment as compared to the neutral conditions from the previous experiment, despite the fact that listeners should have experienced a stronger expectation violation.

What do these results imply for the EXPLAINING AWAY and ASSOCIATIVE accounts that we presented above? Together with the results from Exp. 2, the results from this experiment are unexpected under the ASSOCIATIVE account: if the reason for participants adapting less in the *pessimist* condition had been the difference in context

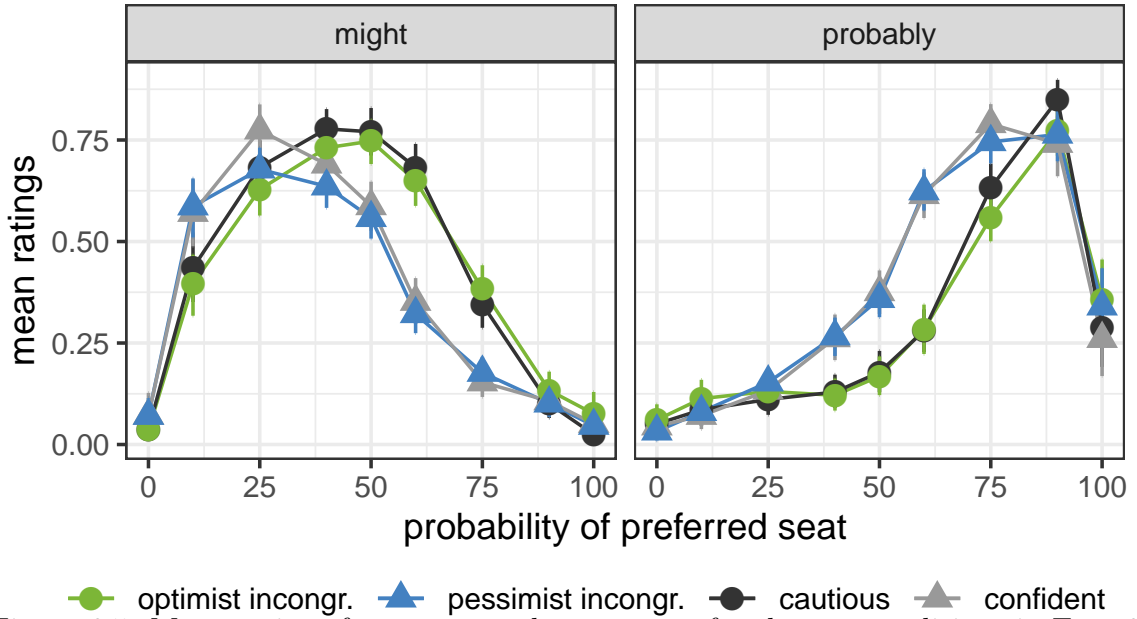


Figure 6.5: Mean ratings for MIGHT and PROBABLY for the two conditions in Exp. 3 as well as the neutral conditions in Exp. 2. Error bars correspond to bootstrapped 95%-confidence intervals.

between the exposure and test blocks, we would have expected less adaptation in this experiment as well.

However, we also did not find stronger adaptation, as we would have expected under the EXPLAINING AWAY account. We can only speculate about the reasons for this, but two explanations seem likely. First, given that there was a numerical difference between the *pessimist incongruent* and *confident* conditions in the expected direction, it could be that our experiment was underpowered to detect a potentially very small effect. However, a power analysis suggests we would need more than 500 participants per condition to achieve power of $\beta = 0.8$ and therefore we did not explore this option further.

Second, it could be that there is a limit on how much listeners can adapt and that this limit is already reached in the neutral conditions. If this was the case, listeners could still experience a stronger expectation violation when the behavior is incongruent with contextual factors but this stronger violation still does not lead to stronger adaptation.

6.5 General Discussion

In three experiments, we showed that language users have different expectations about the use of uncertainty expressions depending on their beliefs about the speaker's mood, and that this difference in expectations affects the extent of semantic adaptation. The results suggest listeners explain away otherwise unexpected behavior when they are presented with a cause, similarly as [?] found for phonetic adaptation.

Our results further largely confirm a prediction made by computational models of adaptation, namely that the extent of adaptation depends on how much observed behavior deviates from prior expectations. Similar results have been found for syntactic priming [?] and are predicted by connectionist models of syntactic learning [?].

However, the results from Exp. 3 suggest that expectation violation is not the only factor that influences adaptation and that there potentially exist limits to how much listeners can adapt. Investigating the limits of adaptation and developing quantitative linking functions between prior expectations and the extent of adaptation will therefore be important avenues for future work.

Chapter 7

Conclusions