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Probability
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(a) Exercise 1.6

Show that if two variables x and y are independent, then their covariance is

We can set:

$$E(X) = \mu, E(y) = v.$$

According to the rule of covariance, $Cov(x,y) = E[E(x-\mu)E(y-\nu)] =$ $E(x*y) - \mu v$

because variables x and y are independent then $E(x^*y)=E(x)E(y)=\mu\nu$ So Cov(x,y)=0

(b)

the total number of people is 158 and set N = 158

- i) from the table, the probability of test being positive given the patient has HIV, $P(+|HIV) = \frac{72}{75} = 96\%$
- ii) from the table, the probability of test being negative given the patient has HIV, $P(-|HIV) = \frac{3}{75} = 4\%$

iii)
$$P(HIV) = 12\%, \dot{P}(H\bar{I}V) = 1 - P(HIV) = 88\%$$

$$P(+) = P(+|HIV) * P(HIV) + P(+|\neg HIV) * P(\neg HIV) = 24.24\%$$

$$P(HIV|+) = \frac{P(+|HIV)*P(HIV)}{P(+)} = 47.52\%$$

$$P(+) = P(+|HIV) * P(HIV) + P(+|\neg HIV) * P(\neg HIV) = 24.24\%$$

$$P(HIV|+) = \frac{P(+|HIV) * P(HIV)}{P(+)} = 47.52\%$$

$$P(HIV|-) = \frac{P(-|HIV) * P(HIV)}{P(-|HIV) * P(HIV)} = 0.63\%$$

Bayes Theorem

a)

set A,B,C,D are the events that there is a big prize behind the door then 2P(A) =

$$2P(B) = 2P(D) = P(C)$$
 and $P(A) + P(B) + P(D) + P(C) = 1$

so
$$P(C|\bar{B}andopen) = \frac{P(\bar{B}andopen|C)*P(C)}{P(\bar{B}andopen)}$$

 $P(\bar{B}andopen) = P(\bar{B}andopen|\bar{B}andC\bar{A}\bar{D}) * P(\bar{B}andC\bar{A}\bar{D}) + P(B|\bar{A}\bar{D}and\bar{C}) *$

 $P(BopenB\bar{A}\bar{D}and\bar{C}) + 2 * P(openB|A\bar{B}\bar{C}\bar{D}) * P(A\bar{B}\bar{C}\bar{D}) = 33.3\%$

so
$$P(C|\bar{B}andopen) = 40\%$$

then
$$P(A|\bar{B}andopen) = P(D|\bar{B}andopen) = \frac{P(\bar{B}andopen|A)*P(A)}{P(\bar{B}andopen)}$$

$$P(A|\bar{B}andopen) = P(D|\bar{B}andopen) = 30\%$$

the probability of choosing the door C to get the big prize is larger than choosing door A or door D

Michael doesn't need to change his selection

b)

Set T is that Professor chin identifies that the student as a cheater, then \bar{T} is the opposite event

Set S is that the student is cheating, then \bar{S} is the opposite event then P(T|S) = 90%

$$\begin{array}{l} P(T|\bar{S}) = 20\% \\ P(T) = P(T|S) * P(S) + P(T|\bar{S}) * P(\bar{S}) = 0.9 * \frac{2}{75} + 0.2 * \frac{72}{75} = 0.219 \ P(S|T) = \frac{P(T|S)}{P(T)} = \frac{P(T|S) * P(S)}{P(T)} = 11\% \end{array}$$

Linear Algebra

a)

Set A is a symmetric matrix and A^{-1} is existent Because A is a symmetric matrix, then $A^T = A$. $(A^{-1})^T = (A^T)^{-1} = A^{-1}$

$$A = \begin{bmatrix} 3 & 4 & -1 \\ -1 & -2 & 1 \\ 3 & 9 & 0 \end{bmatrix}$$

Set x is a non-zero 3-by-1 $x^T = [x_1, x_2x_3]$

$$I = \begin{bmatrix} 1 & 0 & 0 \\ 0 & 1 & 0 \\ 0 & 0 & 1 \end{bmatrix}$$

then $Ax = \bar{\lambda x}$.

$$Ax - I\lambda x = 0$$

$$(A - I\lambda)x = 0$$

$$(A - I\lambda)x = 0$$

$$A-I=A = \begin{bmatrix} 3 - \lambda & 4 & -1 \\ -1 & -2 - \lambda & 1 \\ 3 & 9 & -\lambda \end{bmatrix}$$

compute determinant A-I

eigenvalues is $\lambda_1 = -3, \lambda_2 = 2, \lambda_3 = 2$

compute eigenvectors

$$Ax = \lambda_1 x, Ax = \lambda_2 x, Ax = \lambda_2 x$$

eigenvectors is $x_1^T = [1, -1, 2]$

$$x_2^T = [-1, 1, 3]$$

$$x_2^T = [-1, 1, 3]$$

 $x_3^T = [-1, 1, 3]$

it is not a positive definite

Probability Distributions Exercise2.2

$$P(x|\mu) = \left(\frac{1-\mu}{2}\right)^{\frac{1-x}{2}} \left(\frac{1+\mu}{2}\right)^{\frac{1+x}{2}}$$
 when x = -1, $P(-1|\mu) = \frac{1-\mu}{2}$ when x= 1, $P(1|\mu) = \frac{1+\mu}{2}$ $P(-1|\mu) + P(1|\mu) = 1$

So the distribution is normalized

$$\begin{array}{l} E(X) = \sum x * P(x|\mu) = -1 * \frac{1-\mu}{2} + 1 * \frac{1+\mu}{2} = \mu \\ Var(x) = E(x^2) - [E(x)]^2 = \sum x^2 * P(x|\mu) - \mu^2 = 1 - \mu^2 \\ \text{entropy} \end{array}$$

$$H(x) = -\sum P(x|\mu) * ln(P(x|\mu)) = -(\frac{1-\mu}{2} * ln(\frac{1-\mu}{2}) + \frac{1+\mu}{2} * ln(\frac{1+\mu}{2}))$$

Exercise 2.10

Using the fact that the Dirichlet distribution is normalized

where
$$0 \le \mu_k \le 1$$
 and $\sum_{i=1}^{n-1} \mu_k = 1$ now consider μ_i

Using the fact that the Dirichlet distribution is normalized
$$\int \prod_{k=1}^{M} \mu_{k}^{a_{k}-1} d\mu = \frac{\Gamma(a_{1})...\Gamma(a_{M})}{\Gamma(a_{0})}$$
 where $0 \le \mu_{k} \le 1$ and $\sum \mu_{k} = 1$ now consider μ_{j}
$$E[\mu_{j}] = \frac{\Gamma(a_{1})...\Gamma(a_{M})}{\Gamma(a_{0})} \int \mu_{j} \prod_{k=1}^{M} \mu_{k}^{a_{k}-1} d\mu = \frac{\Gamma(a_{0})}{\Gamma(a_{1})...\Gamma(a_{M})} * \frac{\Gamma(a_{1}).\Gamma(a_{j}+1)..\Gamma(a_{M})}{\Gamma(a_{0}+1)} = \frac{a_{j}}{a_{0}}$$
 because $\Gamma(x+1) = x\Gamma(x)$ then $Var[\mu_{j}] = E[\mu_{j}^{2}] - E[\mu_{j}]^{2} = \frac{a_{j}(a_{j}+1)}{a_{0}(a_{0}+1)} - \frac{a_{j}^{2}}{a_{0}^{2}} = \frac{a_{j}(a_{0}-a_{j})}{a_{0}^{2}(a_{0}+1)}$
$$Cov(\mu_{j}, \mu_{l}) = E[\mu_{j}\mu_{l}] - E[\mu_{j}]E[\mu_{l}] = \frac{a_{j}a_{l}}{a_{0}(a_{0}+1)} - \frac{a_{j}}{a_{0}} \frac{a_{l}}{a_{0}} = -\frac{a_{j}a_{l}}{a_{0}^{2}(a_{0}+1)}$$

then
$$Var[\mu_j] = E[\mu_j^2] - E[\mu_j]^2 = \frac{a_j(a_j+1)}{a_0(a_0+1)} - \frac{a_j^2}{a_0^2} = \frac{a_j(a_0-a_j)}{a_0^2(a_0+1)}$$

 $Cov(\mu_j, \mu_l) = E[\mu_j \mu_l] - E[\mu_j] E[\mu_l] = \frac{a_j a_l}{a_0(a_0+1)} - \frac{a_j}{a_0} \frac{a_l}{a_0} = -\frac{a_j a_l}{a_0^2(a_0+1)}$

Exercise 2.12
$$U(x|a,b) = \frac{1}{b-a} \int_a^b \frac{1}{b-a} dx = \frac{x}{b-a} |_a^b = \frac{b}{b-a} - \frac{a}{b-a} = 1$$
 this distribution is normalized
$$E(\mathbf{x}) = \int_a^b x \frac{1}{b-a} dx = \frac{x^2}{b-a} |_a^b = \frac{b+a}{2}$$

$$Var(x) = E(x^2) - [E(x)]^2 = \int_a^b x^2 \frac{1}{b-a} dx - (\frac{b+a}{2})^2 = \frac{(b-a)^2}{12}$$

$$E(x) = \int_{0}^{b} x \frac{1}{1} dx = \frac{x^{2}}{1} \Big|_{0}^{b} = \frac{b+c}{1}$$

$$Var(x) = E(x^2) - [E(x)]^2 = \int_a^b x^2 \frac{1}{b-a} dx - (\frac{b+a}{2})^2 = \frac{(b-a)^2}{12}$$

Exercise2.15

$$H[X] = \int N(X|\mu, \sum) \ln N(X|\mu, \sum) dx$$

$$= \int N(X|\mu, \Sigma) \frac{1}{2} (D \ln(2\pi) + \ln|\Sigma| + (x - \mu)^T \Sigma^{-1} (x - \mu)) dx$$

$$= \frac{1}{2}(D\ln(2\pi) + \ln|\sum| + Tr[\sum^{-1}\sum])$$

$$= \frac{1}{2}(D\ln(2\pi) + \ln|\sum| + D)$$