

# Lecture 9: MATH 342W: Introduction to Data Science and Machine Learning

Sergio E. Garcia Tapia\*

February 27, 2025 (last updated March 14, 2025)

## Recap

Last time, we saw that  $\hat{\mathbf{y}}$  is the projection of  $\mathbf{y}$  onto the column space of  $X$  (see Figure 1). We can express this as

$$\begin{aligned}\hat{\mathbf{y}} &= \text{proj}_{\text{col}[X]}(\mathbf{y}) \\ &= X(X^\top X)^{-1}X^\top \mathbf{y} \\ &= H\mathbf{y}\end{aligned}$$

where  $H$  is the Hat matrix. What is  $\text{proj}_{\text{col}[X]}(\hat{\mathbf{y}})$ ? That is, what happens if we project  $\hat{\mathbf{y}}$  onto the column space of  $X$ ? Intuitively, it should be  $\hat{\mathbf{y}}$ , because  $\hat{\mathbf{y}}$  is what we get from projecting  $\mathbf{y}$  onto the column space of  $X$ , so  $\hat{\mathbf{y}}$  is already in  $\text{col}[X]$ . If this is the case, then we expect the following equalities to hold:

$$\begin{aligned}\text{proj}_{\text{col}[X]}(\hat{\mathbf{y}}) &= \text{proj}_{\text{col}[X]}(\text{proj}_{\text{col}[X]}(\mathbf{y})) \\ &= H(H\mathbf{y}) \\ &= H^2\mathbf{y} \\ &= H\mathbf{y} \\ &= \hat{\mathbf{y}}\end{aligned}$$

Therefore, we conjecture that  $H \cdot H = H$ . Before we explore this idea, let's consider the null model again.

## The Null Model in Multivariate Regression

In the null model  $g_0$ , we do not take any features into account, so  $p = 0$ . In this case, we expect  $g_0$  to predict  $\bar{y}$  for all inputs, as we saw when we first tackled OLS. Since  $p = 0$  and there are  $n$  responses, the matrix  $X$  is of dimension  $n \times (p + 1)$  or  $n \times 1$ , where the

---

\*Based on lectures of Dr. Adam Kapelner at Queens College. See also the [course GitHub page](#).

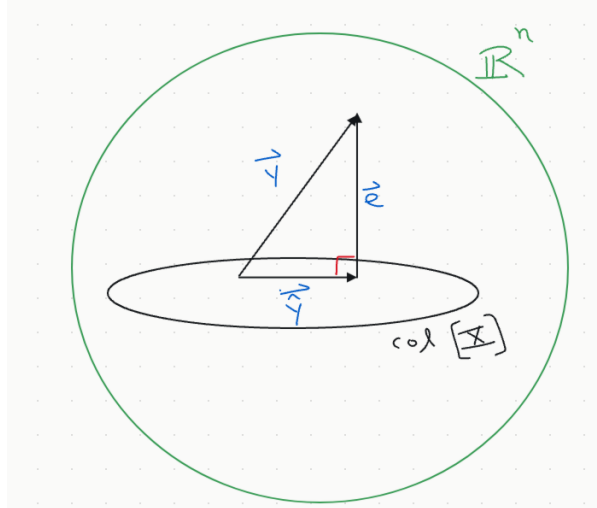


Figure 1: The prediction  $\hat{\mathbf{y}}$  viewed as an orthogonal projection of the response vector  $\mathbf{y}$  onto the column space of  $X$ .

first (and only) column has all 1's, like so:

$$X = \begin{bmatrix} 1 \\ 1 \\ \vdots \\ 1 \end{bmatrix} = \vec{\mathbf{1}}_n$$

Now let's compute  $H$ :

$$\begin{aligned} H &= X(X^\top X)^{-1}X^\top \\ &= \vec{\mathbf{1}}_n(\vec{\mathbf{1}}_n^\top \vec{\mathbf{1}}_n)^{-1}\vec{\mathbf{1}}_n^\top \\ &= \vec{\mathbf{1}}_n(n)^{-1}\vec{\mathbf{1}}_n^\top & (\vec{\mathbf{1}}_n^\top \vec{\mathbf{1}}_n = n) \\ &= \frac{1}{n}\vec{\mathbf{1}}_n\vec{\mathbf{1}}_n^\top \end{aligned}$$

Notice that since  $\vec{\mathbf{1}}_n$  is  $n \times 1$  and  $\vec{\mathbf{1}}_n^\top$  is  $1 \times n$ , the result of  $\vec{\mathbf{1}}_n\vec{\mathbf{1}}_n^\top$  is an  $n \times n$  matrix. The product  $\vec{\mathbf{1}}_n\vec{\mathbf{1}}_n^\top$  is known as an *outer product*. It's easy to verify that the resulting matrix will have all 1's, so

$$\begin{aligned} H &= \frac{1}{n}\vec{\mathbf{1}}_n\vec{\mathbf{1}}_n^\top \\ &= \frac{1}{n} \begin{bmatrix} 1 & \cdots & 1 \\ \vdots & \cdots & \vdots \\ 1 & \cdots & 1 \end{bmatrix} & (\text{The matrix is } n \text{ by } n.) \\ &= \begin{bmatrix} \frac{1}{n} & \cdots & \frac{1}{n} \\ \vdots & \cdots & \vdots \\ \frac{1}{n} & \cdots & \frac{1}{n} \end{bmatrix} \end{aligned}$$

In particular, since every column is a scalar multiple of the first, the rank of this matrix is 1, i.e.,  $\text{rank}(H) = 1$ , which is the number of columns in  $X$  (indeed, we saw last time

that  $\text{rank}(H) = p + 1$ ). Now we can compute the prediction:

$$\hat{\mathbf{y}} = H\mathbf{y} = \begin{bmatrix} \frac{1}{n} & \cdots & \frac{1}{n} \\ \vdots & \cdots & \vdots \\ \frac{1}{n} & \cdots & \frac{1}{n} \end{bmatrix} \mathbf{y} = \begin{bmatrix} \frac{1}{n} \sum_{i=1}^n y_i \\ \vdots \\ \frac{1}{n} \sum_{i=1}^n y_i \end{bmatrix} = \bar{y} \cdot \vec{\mathbf{1}}_n$$

Hence, all  $n$  responses are predicted to be  $\bar{y}$ .

## Properties of Orthogonal Projections

We have mentioned that  $H$  is an orthogonal projection. We will formally define what orthogonal projections are, and prove some useful properties that they satisfy.

**Definition.** A matrix  $P$  is an **orthogonal projection matrix** if and only if  $\forall \mathbf{v}, \mathbf{w} \in \mathbb{R}^n$ , we have

$$(\mathbf{v} - P\mathbf{v})^\top (P\mathbf{w}) = 0$$

That definition says that  $P\mathbf{w}$  and  $(\mathbf{v} - P\mathbf{v})$  are orthogonal for every  $\mathbf{v}, \mathbf{w} \in \mathbb{R}^n$ .

**Theorem 1.** A matrix  $P$  is an orthogonal projection if and only if the following two are both satisfied:

- (i) **Symmetric:**  $P^\top = P$ .
- (ii) **Idempotent:**  $P^2 = P$  (it squares to itself).

*Proof.* We will prove the *if* direction ( $\Leftarrow$ ), and leave  $\Rightarrow$  as an exercise. Thus, we are assuming that  $P^\top = P$  and  $P^2 = P$ . We have to show that it satisfies the definition of an orthogonal projection. Let  $\mathbf{v}, \mathbf{w} \in \mathbb{R}^n$ . Then

$$\begin{aligned} (\mathbf{v} - P\mathbf{v})^\top P\mathbf{w} &= (\mathbf{v}^\top - \mathbf{v}^\top P^\top) P\mathbf{w} \\ &= \mathbf{v}^\top P\mathbf{w} - \mathbf{v}^\top P^\top \cdot P\mathbf{w} \\ &= \mathbf{v}^\top P\mathbf{w} - \mathbf{v}^\top P \cdot P\mathbf{w} && \text{(Symmetry: } P^\top = P) \\ &= \mathbf{v}^\top P\mathbf{w} - \mathbf{v}^\top P^2\mathbf{w} \\ &= \mathbf{v}^\top P\mathbf{w} - \mathbf{v}^\top P\mathbf{w} && \text{(Idempotency: } P^2 = P) \\ &= 0 \end{aligned}$$

□

Let's verify  $H$  satisfies the conditions of Theorem 1. First, we will check that  $H^\top = H$ :

$$\begin{aligned} H^\top &= (X(X^\top X)^{-1}X^\top)^\top \\ &= (X^\top)^\top [(X^\top X)^{-1}]^\top X^\top && \text{(by } (AB)^\top = B^\top A^\top) \\ &= X[(X^\top X)^\top]^{-1}X^\top && \text{(by } (A^{-1})^\top = (A^\top)^{-1}) \\ &= X[X^\top X]^{-1}X^\top && \text{(since } (X^\top X)^\top = X^\top X) \\ &= H \end{aligned}$$

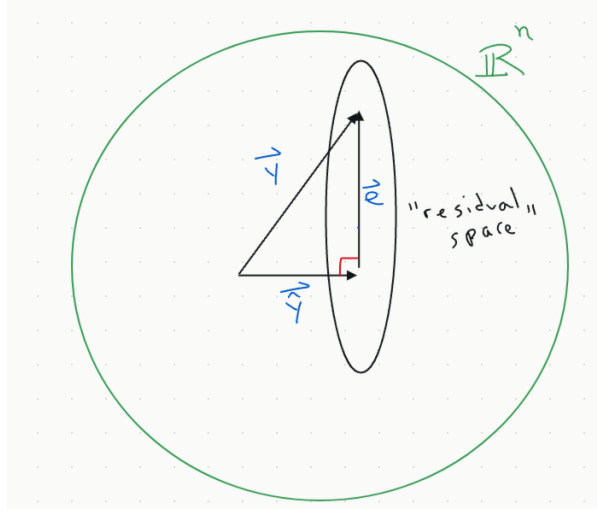


Figure 2: The residual  $\mathbf{e}$  viewed as an orthogonal projection of the response vector  $\mathbf{y}$  onto the “residual space”.

Next, let's check that  $H^2 = H$ :

$$\begin{aligned}
 H^2 &= H \cdot H \\
 &= (X(X^\top X)^{-1}X^\top)(X(X^\top X)^{-1}X^\top) \\
 &= X(X^\top X)^{-1} \underbrace{[(X^\top X)(X^\top X)^{-1}]}_{I_{p+1}} X^\top \\
 &= X(X^\top X)^{-1}X^\top \\
 &= H
 \end{aligned}$$

Thus,  $H$  is indeed an orthogonal projection matrix.

## Orthogonal Projection onto Residual Space

Now let's revisit an idea we mentioned last time, in which we said that  $I_n - H$  is also an orthogonal projection matrix. This time, however, it is a projection onto the residual space (see Figure 2). To justify this, we must argue as in the case for  $H$ , by showing  $I_n - H$  satisfies both conditions of Theorem 1. We will leverage the idempotency and symmetry of  $H$ :

(i) **Symmetric:** We must show  $(I_n - H)^\top = (I_n - H)$ :

$$(I_n - H)^\top = I_n^\top - H^\top = I_n - H$$

Note that the identity matrix is indeed symmetric, and we've used the fact that  $H$  is, too.

(ii) **Idempotent:** We must show  $(I_n - H)^2 = (I_n - H)$ :

$$\begin{aligned}
 (I_n - H)^2 &= (I_n - H)(I_n - H) \\
 &= I_n \cdot I_n - I_n \cdot H - H \cdot I_n + H^2 \\
 &= I_n - H - H + H \quad (H \text{ is idempotent}) \\
 &= I_n - H
 \end{aligned}$$

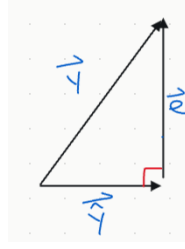


Figure 3: Orthogonal decomposition of response  $\mathbf{y}$  into prediction  $\hat{\mathbf{y}}$  and residual error  $\mathbf{e}$ .

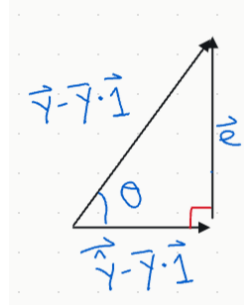


Figure 4: Orthogonal decomposition of mean-control response  $\mathbf{y} - \bar{y} \cdot \vec{\mathbf{1}}_n$  into  $\hat{\mathbf{y}} - \bar{y} \cdot \vec{\mathbf{1}}_n$  and residual error  $\mathbf{e}$ .

## SSR, $R^2$ , and Geometric Interpretations

Consider again Figure 1. Since  $\hat{\mathbf{y}}$  and  $\mathbf{e}$  are orthogonal, note that Pythagorean's Theorem says that

$$\|\mathbf{y}\|^2 = \|\hat{\mathbf{y}}\|^2 + \|\mathbf{e}\|^2$$

See Figure 3. We will use this geometric intuition in a moment. Let's attempt to come up with a fit for  $\mathbf{y} - \bar{y} \cdot \vec{\mathbf{1}}_n$  (referred to as the *mean-control response*) by projecting it onto  $\text{col}[X]$ :

$$\begin{aligned} \text{proj}_{\text{col}[X]}(\mathbf{y} - \bar{y} \cdot \vec{\mathbf{1}}_n) &= H(\mathbf{y} - \bar{y} \cdot \vec{\mathbf{1}}_n) \\ &= H\mathbf{y} - \bar{y} \cdot H\vec{\mathbf{1}}_n \\ &= \hat{\mathbf{y}} - \bar{y} \cdot \vec{\mathbf{1}}_n \end{aligned} \quad (\vec{\mathbf{1}}_n \in \text{col}[X] \implies \vec{\mathbf{1}}_n \in \text{col}[H])$$

Now

$$(\mathbf{y} - \bar{y} \cdot \vec{\mathbf{1}}_n) - (\hat{\mathbf{y}} - \bar{y} \cdot \vec{\mathbf{1}}_n) = \mathbf{y} - \hat{\mathbf{y}} = \mathbf{e}$$

In particular, the two vectors being subtracted above are orthogonal (see Figure 4). Therefore, by Pythagorean's Theorem

$$\begin{aligned} \|\mathbf{y} - \bar{y} \cdot \vec{\mathbf{1}}_n\|^2 &= \|\hat{\mathbf{y}} - \bar{y} \cdot \vec{\mathbf{1}}_n\|^2 + \|\mathbf{e}\|^2 \\ \sum_{i=1}^n (y_i - \bar{y})^2 &= \sum_{i=1}^n (\hat{y}_i - \bar{y})^2 + \sum_{i=1}^n e_i^2 \\ SST &= SSR + SSE \end{aligned} \tag{1}$$

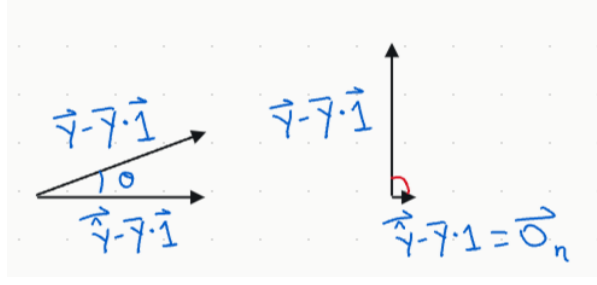


Figure 5: Depictions related to mean-control response in two cases: (i) high  $R^2$ , and (ii)  $\theta = 90^\circ$ , where we do not beat  $g_0$ .

**Definition.** Given a  $n$  real numbers  $(y_i)_{i=1}^n$  with mean  $\bar{y}$  and associated predictions  $(\hat{y}_i)_{i=1}^n$ , the  $SSR$  is given by

$$SSR = \sum_{i=1}^n (\hat{y}_i - \bar{y})^2$$

Using trigonometry, we can see that

$$\begin{aligned} \cos^2 \theta &= \frac{\|\hat{\mathbf{y}} - \bar{y} \cdot \vec{\mathbf{1}}_n\|^2}{\|\mathbf{y} - \bar{y} \cdot \vec{\mathbf{1}}_n\|^2} \\ &= \frac{SSR}{SST} && \text{(by Equation 1)} \\ &= \frac{SST - SSE}{SST} \\ &= 1 - \frac{SSE}{SST} \\ &= R^2 \end{aligned}$$

Since  $\cos^2 \theta \in [0, 1]$ , we see that  $R^2 \in [0, 1]$ . Let's think about what this means. If your projection  $\hat{\mathbf{y}}$  is close to the response vector  $\mathbf{y}$ , then the angle between them is small. On the other hand, if  $\theta = 90^\circ$ , then there is no fit. In the latter case,  $\hat{\mathbf{y}} - \bar{y} \cdot \vec{\mathbf{1}}$  is the zero vector, so  $\hat{\mathbf{y}} = \bar{y} \vec{\mathbf{1}}_n$ , and hence we did not do better than the null model  $g_0$ .

## Reviewing Ignorance Error

Recall that ignorance error comes from the fact that the features (the  $x$ 's) do not give sufficient information about the true drivers (the  $z$ 's). We mentioned that a way to address that is by adding more features (increase  $p$ ).

Let  $X' = [X \mid \mathbf{x}_{\text{new}}]$ , where we append a new column corresponding to a new feature that we measure. Then we expect that the error will decrease, and hence the angle between  $\mathbf{y}$  and  $\hat{\mathbf{y}}$  correspondingly decreases (see Figure 6). In the case where  $p = 1$ , we have  $X = [\vec{\mathbf{1}} \mid \mathbf{x}_{\cdot,1}]$ , so  $\text{col}[X]$  is a 2-dimensional plane. See Figure 7.



Figure 6: Adding a new feature to  $X$ . We expect  $\theta$  to decrease.

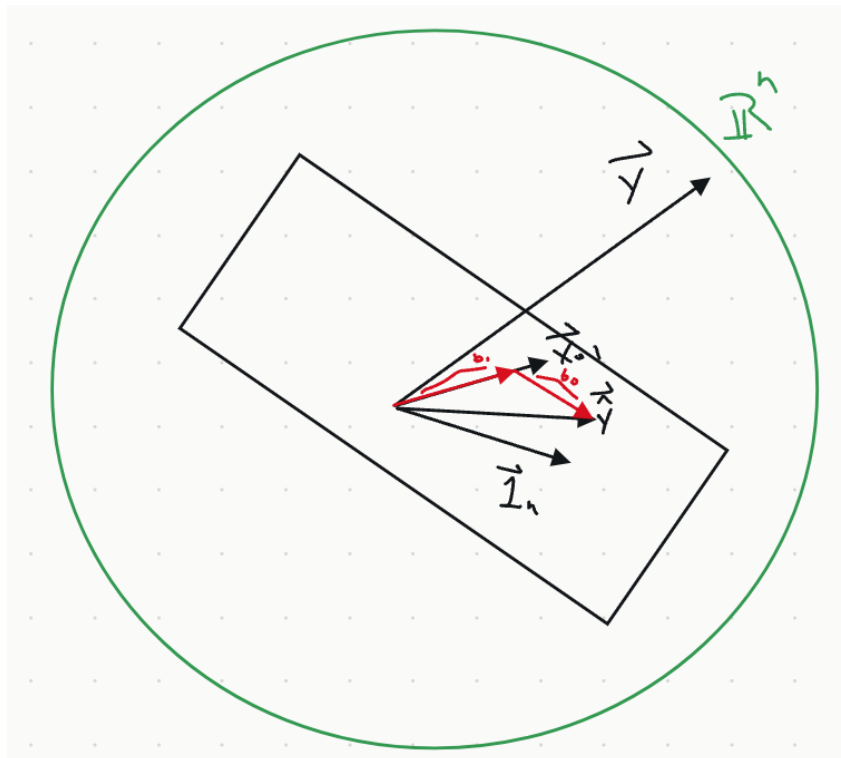


Figure 7: Illustration of least square when using 1 feature. Depicted is the portion of  $y$  along  $x_1$  (the weight  $b_1$ ) and the portion of  $y$  along  $\vec{1}_n$  (the intercept  $b_0$ )

## Eigenvectors and Eigenvalues of $H$

The following material is outside the scope of this class, in the sense that you are not required to know it. Nevertheless, we will explore the concepts.

Recall that  $X$  has a column of 1's and  $p$  columns of features:

$$X = [\vec{\mathbf{1}}_n \quad \mathbf{x}_{\cdot,1} \quad \cdots \quad \mathbf{x}_{\cdot,p}]$$

In particular, the columns of  $X$  are clearly in the column space of  $X$ . Recall that  $H$  is the orthogonal projection matrix onto the column space of  $X$ . This means that

$$H\vec{\mathbf{1}}_n = \vec{\mathbf{1}}_n, \quad H\mathbf{x}_{\cdot,1} = \mathbf{x}_{\cdot,1}, \quad \cdots, \quad H\mathbf{x}_{\cdot,p} = \mathbf{x}_{\cdot,p}$$

One simple way to verify this is as follows:

$$\begin{aligned} HX &= (X(X^\top X)^{-1}X^\top)X \\ &= X[(X^\top X)^{-1}(X^\top X)] \\ &= X \cdot I_{p+1} \\ &= X \end{aligned}$$

We conclude that  $\lambda = 1$  is an eigenvalue of  $H$ , and the eigenspace associated with  $\lambda = 1$  is spanned by  $\mathbf{1}, \mathbf{x}_{\cdot,1}, \dots, \mathbf{x}_{\cdot,p}$ . Since  $H$  is symmetric (i.e., self-adjoint), the spectral theorem guarantees that it has an eigendecomposition (*diagonalization*) (see [Axl23]). Thus, we can write

$$H = P^{-1}DP$$

where

$$P = [\mathbf{v}_1 \quad \mathbf{v}_2 \quad \cdots \quad \mathbf{v}_n], \quad D = \begin{bmatrix} \lambda_1 & 0 & \cdots & 0 \\ 0 & \lambda_2 & \cdots & 0 \\ \vdots & \cdots & \ddots & \vdots \\ 0 & \cdots & 0 & \lambda_n \end{bmatrix}$$

Here,  $D$  is a diagonal matrix consisting of the eigenvalues of  $H$ , and  $P$  is an invertible matrix whose columns are the eigenvectors of  $H$ . We have already argued that

$$\begin{aligned} \mathbf{v}_1 &= \mathbf{1}_n, \quad \mathbf{v}_2 = \mathbf{x}_{\cdot,1}, \quad \dots, \quad \mathbf{v}_{p+1} = \mathbf{x}_{\cdot,p}. \\ \lambda_1 &= \lambda_2 = \cdots = \lambda_{p+1} = 1 \end{aligned}$$

What about the remaining  $n - (p + 1)$  eigenvectors? Note that if a vector belongs to  $\text{col}[X]^\perp$  (the orthogonal complement of  $\text{col}[X]$  or equivalently the residual space), then  $H$  maps it to 0. Therefore, the remaining  $n - (p + 1)$  eigenvalues of  $H$  are all zero, and the eigenvectors associated with the 0 eigenvalue span  $\text{col}[X]^\perp$ .

$$\begin{aligned} P &= [\vec{\mathbf{1}} \quad \mathbf{x}_{\cdot,1} \quad \cdots \quad \mathbf{x}_{\cdot,p} \quad \mathbf{x}_{\perp,1} \quad \cdots \quad \mathbf{x}_{\perp,n-(p+1)}] \\ D &= \begin{bmatrix} 1 & 0 & \cdots & \cdots & \cdots & 0 \\ 0 & \ddots & \ddots & \vdots & \vdots & \vdots \\ \vdots & \ddots & 1 & \ddots & \vdots & \vdots \\ \vdots & \vdots & \ddots & 0 & \ddots & \vdots \\ \vdots & \cdots & \cdots & \ddots & \ddots & \vdots \\ 0 & \cdots & \cdots & \cdots & \ddots & 0 \end{bmatrix} \end{aligned}$$



One last fact is related to the trace of  $H$ . Recall this is the sum of the diagonal entries. We'll leverage the diagonalization:

$$\begin{aligned}
 \sum_{i=1}^n h_{i,i} &= \text{tr}[H] \\
 &= \text{tr}[P^{-1}DP] \\
 &= \text{tr}[PP^{-1}D] & (\text{tr}[ABC] = \text{tr}[CAB] = \text{tr}[BCA]) \\
 &= \text{tr}[D] \\
 &= p + 1 \\
 &= \text{rank}(X)
 \end{aligned}$$

## References

[Axl23] Sheldon Axler. *Linear Algebra Done Right*. 4th ed. Springer, 2023. ISBN: 9783031410253.