Gradia Reference

STAT 245, Spring 2014

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Total: 100 pt

(20)

1 Question 1: Moments of Poisson Distribution

First, we should find the value of $E(X^2)$. We can do this by calculating Var(X):

$$Var(X) = E(X^2) - E(X)^2$$

$$Var(X) = \sum_{k=0}^{\infty} k^2 e^{-\lambda} \frac{\lambda^k}{k!} - \lambda^2$$

Now we can work some algebra magic:

$$E(X^{2}) - \lambda = E(X(X - 1)) = \sum_{k=0}^{\infty} k(k - 1)e^{-\lambda} \frac{\lambda^{k}}{k!}$$

$$E(X^{2}) - \lambda = e^{-\lambda} \lambda^{2} \sum_{k=2}^{\infty} \frac{\lambda^{k-2}}{(k-2)!}$$

$$E(X^{2}) = \lambda^{2} + \lambda$$

$$Var(X) = E(X^{2}) - E(X)^{2} = \lambda + \lambda^{2} - \lambda^{2} = \lambda$$

Now, we can use this to calculate $E(X^3)$:

$$\begin{split} \mathrm{E}(X^3) - 3\mathrm{E}(X^2) + 2\mathrm{E}(X) &= \mathrm{E}(X^3 - 3x^2 + 2x) = \mathrm{E}[(x^2 - x)(x - 2)] = \mathrm{E}[(x)(x - 1)(x - 2)] \\ &= \mathrm{E}[(x)(x - 1)(x - 2)] = \sum_{k=0}^{\infty} k(k - 1)(k - 2)e^{-\lambda}\frac{\lambda^k}{k!} \\ &= \mathrm{E}[(x)(x - 1)(x - 2)] = e^{-\lambda}\lambda^3 \sum_{k=3}^{\infty} \frac{\lambda^{k-3}}{(k - 3)!} \\ &= \mathrm{E}[(x)(x - 1)(x - 2)] = \lambda^3 \\ &= \mathrm{E}[(x)(x - 1)(x - 2)] = \mathrm{E}(X^3) - 3\mathrm{E}(X^2) + 2\mathrm{E}(X) = \lambda^3 \\ &= \mathrm{E}(X^3) - 3(\lambda^2 - \lambda) + 2\lambda = \lambda^3 \end{split}$$

 $E(X^3) = \lambda^3 + 3\lambda^2 + \lambda$

Let's do the same for $E(X^4)$:

$$\begin{split} \mathrm{E}(X^4) - 6\mathrm{E}(X^3) + 11\mathrm{E}(X^2) - 6\mathrm{E}(X) &= \mathrm{E}[x(x-1)(x-2)(x-3)] \\ \mathrm{E}[x(x-1)(x-2)(x-3)] &= \sum_{k=0}^{\infty} k(k-1)(k-2)(k-3)e^{-\lambda}\frac{\lambda^k}{k!} \\ \mathrm{E}[x(x-1)(x-2)(x-3)] &= \lambda^4 e^{-\lambda} \sum_{k=4}^{\infty} \frac{\lambda^{k-4}}{(k-4)!} = \lambda^4 e^{-\lambda} e^{\lambda} \\ \mathrm{E}[x(x-1)(x-2)(x-3)] &= \lambda^4 \\ \mathrm{E}(X^4) &= \lambda^4 + 6\lambda^3 + 7\lambda^2 + \lambda \end{split}$$

Question 2: Poisson and χ^2 -Tails

We want to show that:

$$P(X_{\lambda} \ge c+1) = P(0 \le Y_{2(c+1)} \le 2\lambda) = \int_0^{2\lambda} f_{X_{2(c+1)}^2}(y) dy$$

The last two terms are trivially equal. Let's first work out the left side:

$$P(X_{\lambda} \ge c + 1) = 1 - P(X_{\lambda} \le c) = 1 - \sum_{i=0}^{c} \frac{\lambda^{i} e^{-\lambda}}{i!}$$

Now let's take the derivative with respect to
$$\lambda$$
:
$$\frac{d}{d\lambda} \int_{c}^{c} (\lambda \lambda) c(t) = e^{-\lambda} - \sum_{i=1}^{c} e^{-\lambda} \lambda^{i-1} \cdot \frac{i-\lambda}{i!}$$

$$= e^{\lambda} - e^{-\lambda} \left(1 - \frac{\lambda^{c}}{c!}\right) = e^{-\lambda} \frac{\lambda^{c}}{c!}$$

$$= 2 \int_{c}^{c} (2\lambda) = d \cdot (2\lambda)^{c}$$

Question 3: Approximations to Binomial Probabilities $\frac{\lambda^{c}}{\lambda^{c}} = e^{\lambda} \left(1 - \frac{\lambda^{c}}{c!}\right) = e^{\lambda} \frac{\lambda^{c}}{c!}$ for Dago

Part A 3.1

$$n = 7, p = 0.3, k = 3$$

$$P(k=3) = \binom{7}{3} 0.3^3 0.7^4 = 0.2269$$

Normal Approximation:

$$Z = \frac{k - np}{\sqrt{np(1 - p)}}$$

$$Z = \frac{3 - (7)(0.3)}{\sqrt{(7)(0.3)(0.7)}} = 0.74$$

Poisson Approximation:

$$Z = \frac{\sqrt{np(1-p)}}{\sqrt{(7)(0.3)(0.7)}} = 0.7423 \qquad Z_1 = \frac{2.5 - 7 \times 0.5}{\sqrt{7 \times 0.3} \times 0.7} = 0.32 \text{ }$$

$$P(Z \le 0.7423) = 0.7718 \qquad Z_2 = \frac{3.5 - 7 \times 0.5}{\sqrt{7 \times 0.3} \times 0.7} = 0.32 \text{ }$$

$$P(X) \approx \frac{e^{-np}(np)^k}{k!} \qquad \qquad P(Z_1 \le Z_2 \le Z_2) = 0.246$$

$$P(X) \approx \frac{e^{-np}(np)^k}{k!}$$

$$P(3) \approx \frac{e^{-7(0.3)}(7(0.3))^3}{3!}$$

$$P(3) \approx 0.1890$$

$$n = 7, p = 0.3, k = 3$$

$$P(k = 3) = \binom{7}{3} 0.3^{3} 0.7^{4} = 0.2269$$

$$P(k = 3) = \binom{7}{3} 0.3^{3} 0.7^{4} = 0.2269$$

$$P(0 \le \chi_{2(c+1)} \le 2\lambda) = 0$$

so
$$P(X_{\lambda} \geq Cr1) = p(o \leq Y_{\lambda(c+1)} \leq 2)$$

$$Z_1 = \frac{2.5 - 7 \times 0.5}{\sqrt{7 \times 0.5 \times 0.7}} = 0.329$$

Part B 3.2

$$n = 40, p = 0.4, k = 11$$

$$P(k=11) = \binom{40}{11} 0.4^{11} 0.6^{29} = 0.03573$$

Normal Approximation:

$$Z = \frac{k - \mu}{\sigma}$$

$$Z = \frac{k - np}{\sqrt{np(1 - p)}}$$

$$Z = \frac{11 - (40)(0.4)}{\sqrt{(40)(0.4)(0.6)}} = -1.6137$$

$$P(Z \le 1.6137) = 0.0533$$

$$P(X) \approx \frac{e^{-np}(np)^k}{k!}$$

$$P(11) \approx \frac{e^{-40(0.4)(40(0.4))^{11}}}{11!}$$

$$Z_1 = \frac{10.5 - 40 \times 0.4}{\sqrt{40 \times 0.4 \times 0.6}} = -1.775$$

$$\frac{2}{\sqrt{40 \times 0.4 \times 0.6}} = -1.4452$$

Poisson Approximation:

3.3 Part C

$$n = 400, p = 0.0025, k = 2$$

$$P(k = 2) = {400 \choose 2} 0.0025^{2} 0.9975^{398} = 0.1842$$

 $P(11) \approx 0.0496$

Normal Approximation:

$$Z = \frac{k - np}{\sqrt{np(1 - p)}}$$

$$Z = \frac{2 - (408)(0.0025)}{\sqrt{(400)(0.0025)(0.9975)}} = 1.0012$$

$$P(Z \le 1.0012) = 0.8416$$

$$Z_1 = \frac{1.5 - 400 \times 0.0025}{\sqrt{400 \times 0.0025 \times 0.9975}} = 0.5006$$

$$Z_2 = \frac{1.5 - 400 \times 0.0025}{\sqrt{400 \times 0.0025 \times 0.9975}} = 0.5006$$

Poisson Approximation:

$$P(X) \approx \frac{e^{-np}(np)^k}{k!}$$
 $P(X) \approx \frac{e^{-400(0.0025)}(400(0.0025))^2}{2!}$ $P(X) \approx \frac{e^{-400(0.0025)}(400(0.0025))^2}{2!}$ $P(X) \approx \frac{e^{-400(0.0025)}(400(0.0025))^2}{2!}$

(ω) 4 Question 4: Conditional Distributions in Poisson Process

4.1 Part A

We're going to use the fact that a Poisson process has a constant rate of emission on equilength intervals. Let's assume that $X \sim \text{Poisson}(\lambda)$ for a time unit of 1, and use Bayes Rule to set the problem up

$$P(X_s = k | X_t = n) = \frac{P(X_t = n \mid X_s = k)P(X_s = k)}{P(X_t = n)}$$

For the first term in the numerator, since t > s > 0, we know that:

$$P(X_t = n \mid X_s = k) = P(X_{t-s} = n - k) = \frac{e^{-(t-s)\lambda}((t-s)\lambda)^{n-k}}{(n-k)!}$$

$$P(X_s = k) = \frac{e^{-s\lambda}(s\lambda)^k}{k!}$$

$$P(X_t = n) = \frac{e^{-t\lambda}(t\lambda)^n}{n!}$$

Now we can plug this all in and combine it:

$$P(X_s = k | X_t = n) = \frac{\left(\frac{e^{-(t-s)\lambda}((t-s)\lambda)^{n-k}}{(n-k)!}\right) \left(\frac{e^{-s\lambda}(s\lambda)^k}{k!}\right)}{\frac{e^{-t\lambda}(t\lambda)^n}{n!}}$$

$$P(X_s = k | X_t = n) = \binom{n}{k} \left(\frac{((t-s)\lambda)^{n-k}(s\lambda)^k}{(t\lambda)^n}\right)$$

$$P(X_s = k | X_t = n) = \binom{n}{k} \left(\frac{(t-s)^{n-k}s^k}{t^n}\right)$$

4.2 Part B

$$P(T_1 \le s \mid X_t = 1) = P(X_s = 1, X_t - X_s = 0 \mid X_t = 1) = \frac{P(X_s = 1, X_t - X_s = 0)}{P(X_t = | \mathbf{0})} = \frac{P(X_s = 1)P(X_t - X_s = 0)}{P(X_t = 1)}$$

Now we can use the Poisson distribution to get:

$$P(T_1 \le s \mid X_1 = 1) = \frac{e^{-rs}se^{-r(t-s)}}{e^{-rt}t} = \frac{s}{t}, 0 < s \le t$$

()) 5 Question 5: Data from Poisson Process

Table 1: 10 Second Counts

Count	# intervals
0	23
1	77
2	34
3	26
4.	13.
5	7

Table 2: 20 Second Counts

Count	# intervals
0.	2
1	4
2	·9·
3	5

(a). Given this data, we can effectively construct a maximum likelihood estimator for $\hat{\lambda}$:

$$L = \prod_{i=1}^{180} \frac{e^{-10\lambda} (10\lambda)^{y_i}}{y_i!} \times \prod_{j=1}^{20} \frac{e^{-20\lambda} (20\lambda)^{z_j}}{z_j!}$$

Now we transform this into a loglikelihood function:

$$\ln L = -180(10\lambda) + \sum_{i=1}^{180} y_i \ln(10\lambda) - \sum_{j=1}^{180} \ln(y_i!) + -20(20\lambda) + \sum_{j=1}^{20} z_j \ln(20\lambda) - \sum_{j=1}^{20} \ln(z_j!)$$

$$\ln L = -180(10\lambda) + \sum_{i=1}^{180} y_i \ln(10) + \sum_{i=1}^{180} y_i \ln(\lambda) - \sum_{j=1}^{180} \ln(y_i!) + -20(20\lambda) + \sum_{j=1}^{20} z_j \ln(20) + \sum_{j=1}^{20} z_j \ln(\lambda) - \sum_{j=1}^{20} \ln(z_j!)$$

Now we can differentiate the equation with respect to λ , set the derivative to 0 and solve for $\hat{\lambda}$: