

计算语言学 Computational Linguistics

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第四章 估计句子的概率-Markov模型 (Part 2)

- MLE: MLE is usually unsuitable for NLP because of the sparseness of the data
- p(z | xy) = ?
- Suppose our training data includes

```
... xya ..
```

... xyd ...

... xyd ...

but never xyz

Should we conclude

$$p(a \mid xy) = 1/3?$$

 $p(d \mid xy) = 2/3?$
 $p(z \mid xy) = 0/3?$

NO! Absence of xyz might just be bad luck.

Should we conclude

```
p(a \mid xy) = 1/3? reduce this p(d \mid xy) = 2/3? reduce this p(z \mid xy) = 0/3? increase this
```

Baisc idea

- Discount the positive counts somewhat
- •Reallocate that probability to the zeroes

N	Number of training instances
В	Number of bins training instances are divided into
w_{1n}	An n-gram $w_1 \cdots w_n$ in the training text
$C(w_1 \cdot \cdot w_n)$	Frequency of n-gram $w_1 \cdots w_n$ in training text
r	Frequency of an n-gram
$f(\cdot)$	Frequency estimate of a model
N_r	Number of bins that have r training instances in them
T_r	Total count of n-grams of frequency r in further data
h	'History' of preceding words

In person	she		was		inferior		to		both		eletam	
•											sisters	
l-gram	P(·) the* to and of	0.034 0.032 0.030 0.029	P(·) the to and of	0.034 0.032 0.030 0.029	P(·) the to and of	0.034 0.032 0.030 0.029	P(·) the to	0.034 0.032	P(·) the to and of	0.034 0.032 0.030 0.029	P(·) the to and of	0.034 0.032 0.030 0.029
8	was	0.015	was	0.015	was	0.015			was	0.015	was	0.015
13	she	0.011			she	0.011			she	0.011	she	0.01 I
254					both	0.0005			both	0.0005	both	0.0005
435					qiqterq	0.0003					sisters	0.0003
1701					inferior	0.00005						
Z-gram	$P(\cdot pe$	erson)	$P(\cdot sh$	e)	$P(\cdot was)$		$P(\cdot \mid in$	iferior)	$P(\cdot to)$		$P(\cdot both)$	
1 2 3 4	and who to in	0.099 0.099 0.076 0.045	had was	0.141 0.122	not a the to	0.065 0.052 0.033 0.031	to	0.2 12	be the her have	0.111 0.057 0.048 0.027	of to in and	0.066 0.041 0.038 0.025
23	she	0.009							Mrs	0.006	she	0.009
41									what	0.004	sisters	0.006
293									both	0.0004		
06					inferior	0						
3-gram	$P(\cdot In$	person)	$P(\cdot pe$	rson,she)	$P(\cdot she, w$	as)	$P(\cdot wc$	as,inf.)	$P(\cdot inferi$	or,to)	P(to,bot	h)
2 4	Un	SEEN	did was	0.5 0.5	not very in to	0.057 0.038 0.030 0.026	Uns	SEEN	the Maria cherries her	0.286 0.143 0.143 0.143	to Chapter Hour Twice	0.222 0.111 0.111 0.111
DG					inferior	0			both	0	sisters	0
4-gram	$P(\mid u_i)$	l,p)	$P(\cdot l,p$	(s)	$P(\cdot p, s, w)$		$F(\cdot s,w)$	$\nu, i)$	$P(\cdot w,i,t)$		$F(\cdot i_rt_rb)$	
	UN	SEEN	Un	SEEN	in	1.0	Uns	SEEN	Uns	EEN	Unse	EN
100					inferior	0						

Laplace's law (adding one)

$$P_{Lap}(w_1...w_n) = \frac{C(w_1...w_n) + 1}{N + B}$$

$$f_{Lap}(w_1...w_n) = \frac{C(w_1...w_n) + 1}{N + B} * N$$

$$N_0 * P_{Lap}(C(w_1...w_n) = 0)$$

unsmoothed bigram counts: 2nd word

1st word

	I	want	to	eat	Chinese	food	lunch	 Total (N)
Ι	8	1087	0	13	0	0	0	3437
want	3	0	786	0	6	8	6	1215
to	3	0	10	860	3	0	12	3256
eat	0	0	2	0	19	2	52	938
Chinese	2	0	0	0	0	120	1	213
food	19	0	17	0	0	0	0	1506
lunch	4	0	0	0	0	1	0	459

unsmoothed normalized bigram probabilities:

	I	want	to	eat	Chinese	food	lunch	 Total
I	.0023 (8/3437)	.32	0	.0038 (13/3437)	0	0	0	1
want	.0025	0	.65	0	.0049	.0066	.0049	1
to	.00092	0	.0031	.26	.00092	0	.0037	1
eat	0	0	.0021	0	.020	.0021	.055	1
Chinese	.0094	0	0	0	0	.56	.0047	1
food	.013	0	.011	0	0	0	0	1
lunch	.0087	0	0	0	0	.0022	0	1

add-one smoothed bigram counts:

	I	want	to	eat	Chinese	food	lunch	 Total (N+V)
Ι	8 9	1087	1	14	1	1	1	3437
		1088						5053
want	3 4	1	787	1	7	9	7	2831
to	4	1	11	861	4	1	13	4872
eat	1	1	23	1	20	3	53	2554
Chinese	3	1	1	1	1	121	2	1829
food	20	1	18	1	1	1	1	3122
lunch	5	1	1	1	1	2	1	2075

add-one normalized bigram probabilities:

	I	want	to	eat	Chinese	food	lunch	 Total
I	.0018 (9/5053)	.22	.0002	.0028 (14/5053)	.0002	.0002	.0002	1
want	.0014	.00035	.28	.00035	.0025	.0032	.0025	1
to	.00082	.00021	.0023	.18	.00082	.00021	.0027	1
eat	.00039	.00039	.0012	.00039	.0078	.0012	.021	1
Chinese	.0016	.00055	.00055	.00055	.00055	.066	.0011	1
food	.0064	.00032	.0058	.00032	.00032	.00032	.00032	1
lunch	.0024	.00048	.00048	.00048	.00048	.0022	.00048	1

- * Data from the AP from (Church and Gale, 1991): AP data, 44 million words
 - * Corpus of 22,000,000 bigrams
 - * Vocabulary of 273,266 words (i.e. 74,674,306,760 possible bigrams or bins)
 - * 74,671,100,000 bigrams were unseen
 - * And each unseen bigram was given a frequency of 0.000295

	f _{MLE}	f _{empirical}	f _{add-one}	
Enca from	0	0.000027	0.000295	Add-one
Freq. from training data	1	0.448	0.000589	smoothed freq.
	2	1.25	0.000884	
	3	2.24	0.00118	too high
Freq. from held-out data	4	3.23	0.00147	
held-out data	5	4.21	0.00177	too low
ψ 	1			

* Total probability mass given to unseen bigrams = (74,671,100,000 x 0.000295) / 22,000,000 = **0.9997** !!!!

Lidstone's law and the Jeffreys-Perks law (adding λ)

$$P_{Lid}(w_1...w_n) = \frac{C(w_1...w_n) + \lambda}{N + B\lambda}$$

$$\mu = \frac{N}{N + B\lambda}$$

$$P_{Lid}(w_1...w_n) = \mu \frac{C(w_1...w_n)}{N} + (1-\mu)\frac{1}{B}$$

Jeffreys-Perks law:
$$\lambda = 0.5$$

Expected Likelihood Estimation

$$\begin{array}{c} \lambda = 1 \rightarrow \text{Laplace} \\ \lambda = 0.5 \rightarrow \text{Jeffreys-Perks} \\ \lambda = 0 \rightarrow \text{MLE} \end{array}$$

She was inferior to both sisters.

• Unigram: p(S)=3.96*e-17

Bigram: p(S)=0;

• ELE (Bigram) : p(S)=6.89*e-20

Poor estimates of context are worse than none.

Rank	Word	MLE	ELE
1	not	0.065	0.036
2	a	0.052	0.030
3	the	0.033	0.019
4	to	0.031	0.017
• • •			
=1482	inferior	0	0.00003

• Held out estimation: How do we know how much of the probability space to "hold out" for unseen events?

$$C_1(\mathbf{w}_1...\mathbf{w}_n) = \text{frequency of } \mathbf{w}_1...\mathbf{w}_n \text{ in the training data}$$

$$C_2(\mathbf{w}_1...\mathbf{w}_n) = \text{frequency of } \mathbf{w}_1...\mathbf{w}_n \text{ in the held out data}$$

 N_r is the number of n-grams with frequency r in the training data

Let T_r is the total number of times that all n-grams that appeared r times in the training data appeared in the held out data. Then:

$$P_{ho}(w_1...w_n) = \frac{T_r}{N_r N}$$

training data vs. held out data (validation data)

development test data vs. final test data

Training:

- Training data (80% of total data)
 - •To build initial estimates (frequency counts)
- Held out data (10% of total data)
 - •To refine initial estimates (smoothed estimates)

Testing:

- Development test data (5% of total data)
 - To test while developing
- Final test data (5% of total data)
 - To test at the end

	System 1	System 2
scores	71, 61, 55, 60, 68, 49,	42, 55, 75, 45, 54, 5
	42, 72, 76, 55, 64	55, 36, 58, 55, 67
total	609	526
n	11	11
mean \bar{x}_i	55.4	47.8
$s_i^2 = \sum (x_{ij} - \bar{x}_i)^2$	1,375.4	1,228.8
df	10	10

Pooled
$$s^2 = \frac{1375.4 + 1228.8}{10 + 10} \approx 130.2$$

$$t = \frac{\bar{x}_1 - \bar{x}_2}{\sqrt{\frac{2s^2}{n}}} = \frac{55.4 - 47.8}{\sqrt{\frac{2 \cdot 130.2}{11}}} \approx 1.56$$

t-test alpha=0.05

t = 1.56 < 1.725

Cross-validation (deleted estimation)

Held out estimation is useful if there is a lot of data available

If not, we can use each part of the data both as training data and as held out data.

$$P_{ho}(w_1...w_n) = \frac{T_r^{01}}{N_r^0 N} \text{ or } \frac{T_r^{10}}{N_r^1 N}$$

$$P_{del}(w_1...w_n) = \frac{T_r^{01} + T_r^{10}}{N(N_r^0 + N_r^1)}$$

Good-Turing estimation

$$r^* = (r+1)\frac{E(N_{r+1})}{E(N_r)}$$
 $P_{GT} = \frac{r^*}{N}$

Two strategies:

- 1. only for very low frequencies
- 2. (*r*, *Nr*) fit *S*

If
$$C(w_1...w_n) = r > 0$$

$$P_{GT}(w_1...w_n) = \frac{r^*}{N}$$
 where $r^* = (r+1)\frac{S(r+1)}{S(r)}$

else
$$P_{GT}(w_1...w_n) = \frac{1 - \sum_{r=1}^{\infty} N_r \frac{r^*}{N}}{N_0} \approx \frac{N_1}{N_0 N}$$

Empirical results for bigram data (Church and Gale)

<u>f</u>	_f _{emp}	_f _{GT}	_f_ add1	<u>f_{del}</u>
0	0.000027	0.000027	0.000295	0.000037
1	0.448	0.446	0.000589	0.396
2	1.25	1.26	0.000884	1.24
3	2.24	2.24	0.00118	2.23
4	3.23	3.24	0.00147	3.22
5	4.21	4.22	0.00177	4.22
6	5.23	5.19	0.00206	5.20
7	6.21	6.21	0.00236	6.21
8	7.21	7.24	0.00265	7.18
9	8.26	8.25	0.00295	8.18

	Bigra	ams			Trigra	ams	
r	N_r	r	N_r	r	N_r	r	N_r
1	138741	28	90	1	404211	28	35
2	25413	29	120	2	32514	29	32
3	10531	30	86	3	10056	30	25
4	5997	31	98	4	4780	31	18
5	3565	32	99	5	2491	32	19
6	2486		• • •	6	1571		••
7	1754	1264	1	7	1088	189	1
8	1342	1366	1	8	749	202	1
9	1106	1917	1	9	582	214	1
10	896	2233	1	10	432	366	1
		2507	1			378	1

Table 6.7 Extracts from the frequencies of frequencies distribution for bigrams and trigrams in the Austen corpus.

-	r **	$P_{\mathrm{GT}}(\cdot)$
•	0.0007	1.058×10^{-9}
1	0.3663	$5.982 \times 10 - 7$
2	1.228	2.004×10^{-6}
3	2.122	3.465×10^{-6}
4	3.058	$4.993 imes 10^{-6}$
5	4.015	6.555×10^{-6}
6	4.984	8.138 $ imes$ 10^{-6}
7	5.96	9.733×10^{-6}
8	6.942	1.134×10^{-5}
9	7.928	1.294×10^{-5}
10	8.916	1.456×10^{-5}
28	26.84	4.383 x10-5
29	27.84	4.546×10^{-5}
30	28.84	4.709×10^{-5}
31	29.84	4.872×10^{-5}
32	30.84	$5.035 imes 10^{-5}$
1264	1263	0.002062
1366	1365	0.002228
1917	1916	0.003128
2233	2232	0.003644
2507	2506	0.004092

$$P(she|person) = \frac{f_{GT}(person she)}{C(person)} = \frac{1.228}{223} = 0.0055$$

- Katz's backing-off (1987)
 - Why are we treating all novel events as the same?
 - p(zygote | see the) vs. p(baby | see the)
 - Suppose both trigrams have zero count
 - baby beats zygote as a unigram
 - the baby beats the zygote as a bigram
 - see the baby beats see the zygote ?

Basic smoothing (e.g., add- λ or Good-Turing):

- * Holds out some probability mass for novel events
- * Divided up **evenly** among the novel events

Backoff smoothing

- * Holds out same amount of probability mass for novel events
- * But divide up **unevenly** in proportion to backoff prob.
- * For p(z | xy):
 - * Novel events are types z that were never observed after xy
 - * Backoff prob for p(z| xy) is p(z| y) ... which in turn backs off to p(z)!

- In back-off models, different models are consulted in order depending on their specificity.
- If the n-gram of concern has appeared more than k times, then an n-gram estimate is used but an amount of the MLE estimate gets discounted (it is reserved for unseen n-grams).
- If the n-gram occurred k times or less, then we will use an estimate from a shorter n-gram (back-off probability), normalized by the amount of probability remaining and the amount of data covered by this estimate.
- The process continues recursively.

$$(1 - d_{w_{i-n+1}...w_{i-1}}) \frac{C(w_{i-n+1}...w_{i})}{C(w_{i-n+1}...w_{i-1})} \quad \text{if } C(w_{i-n+1}...w_{i}) > k$$

$$P_{bo}(w_{i} \mid w_{i-n+1}...w_{i-1}) = (1 - d_{w_{i-n+1}...w_{i-1}}) \quad \text{if } C(w_{i-n+1}...w_{i}) > k$$

$$\alpha_{w_{i-n+1}...w_{i-1}} P_{bo}(w_i | w_{i-n+2}...w_{i-1})$$
 Otherwise

	P(she h)	P(was h)	P(inferior h)	P(to h)	P(both h)	P(sisters h)	Product
Unigram	0.011	0.015	0.00005	0.032	0.0005	0.0003	3.96 x 10 ⁻¹⁷
Bigram	0.00529	0.1219	0.0000159	0.183	0.000449	0.00372	3.14 x 10 ⁻¹⁵
n used	2	2	1	2	2	2	
Tuicrnoma							4 44 40 15
Trigram	0.00529	0.0741	0.0000162	0.183	0.000384	0.00323	1.44 x 10 ⁻¹⁵

Table 6.11 Probability estimates of the test clause according to various language models. The unigram estimate is our previous MLE unigram estimate. The other two estimates are back-off language models. The last column gives the overall probability estimate given to the clause by the model.

Simple linear interpolation (deleted interpolation)

$$P_{li}(w_n|w_{n-2},w_{n-1}) = \lambda_1 P(w_n) + \lambda_2 P(w_n|w_{n-1}) + \lambda_3 P(w_n|w_{n-2},w_{n-1})$$

where
$$0 \le \lambda_i \le 1$$
 and $\Sigma_i \lambda_i = 1$

- General linear interpolation
- EM (Expectation maximization)

"Whenever data sparsity is an issue, smoothing can help performance, and data sparsity is almost always an issue in statistical modeling. In the extreme case where there is so much training data that all parameters can be accurately trained without smoothing, one can almost always expand the model, such as by moving to a higher n-gram model, to achieve improved performance. With more parameters data sparsity becomes an issue again, but with proper smoothing the models are usually more accurate than the original models. Thus, no matter how much data one has, smoothing can almost always help performace, and for a relatively small effort."

Chen & Goodman (1998)

- Chapter 6: Statistical Inference: n-gram Models over Sparse Data, Foundations of Statistical NLP
- Stanley F. Chen, Joshua Goodman (1996, Harvard University), *An Empirical Study of Smoothing Techniques for Language Modeling*, Proceedings of the Thirty-Fourth Annual Meeting of the Association for Computational Linguistics
- Church, K., and Gale, W. (1990), *Poor Estimates of Context are Worse than None*, Third Darpa Workshop on Speech and Natural Language, Hidden Valley, PA.
- K. Church and W. Gale. (1991). A comparison of the enhanced Good-Turing and deleted estimation methods for estimating probabilities of English bigrams. Computer Speech and Language 5:19-54.