

M-Estimators

Chapter 3.2

Thomas Nagler <t.w.nagler@math.leidenuniv.nl> Leiden University

What are M-estimators

Maximum-likelihood (Penalized) Least squares Robust regression Quantile regression Kernel smoothing Support vector machines Neural networks M-estimator

$$\hat{\theta}_n = \arg\max_{\theta} \frac{1}{n} \sum_{i=1}^n m_{\theta}(X_i).$$

More generally,

$$\hat{\theta}_n = \arg\max_{\theta} \mathbb{M}_n(\theta).$$

Some notation

• Let θ_0 be the 'true' parameter, i.e,

$$\theta_0 = \arg\max_{\theta} \mathbb{M}(\theta),$$

for some limit process M.

• Sometimes it is useful to set $h = r_n(\theta - \theta_0)$ as parameter and reformulate $\mathbb M$ accordingly: If

$$\hat{\theta}_n = \arg\max_{\theta} \mathbb{M}_n(\theta),$$

we write

$$\hat{h}_n = \arg\max_h \mathbb{M}'_n(h) = \arg\max_h \mathbb{M}_n(\theta_0 + h/r_n) - \mathbb{M}_n(\theta_0).$$

Outline

General strategy for dealing with M-estimators:

- 1. Consistency
- 2. Rate of convergence
- 3. Asymptotic distribution

Lemma (3.2.1)

Let \mathbb{M}_n , \mathbb{M} be processes index by a metric space H. Suppose:

- $\mathbb{M}_n \leadsto \mathbb{M}$ in $\ell^{\infty}(A \cup B)$ with $A, B \subset H$ arbitrary.
- There is \hat{h} such that for every open set G containing \hat{h} ,

$$\mathbb{M}(\hat{h}) > \sup_{h \notin G, h \in A} \mathbb{M}(h), \quad a.s.$$

• \hat{h}_n satisfies $\mathbb{M}_n(\hat{h}_n) \geq \sup_h \mathbb{M}_n(h) - o_P(1)$.

Then, for every closed set F,

$$\limsup_{n\to\infty} P^*(\hat{h}_n\in F\cap A)\leq P(\hat{h}\in F\cup B^c).$$

Note: A = B = H would imply $\hat{h}_n \leadsto \hat{h}$ in H.

Proof. By continuous the mapping theorem,

$$\sup_{h\in F\cap A}\mathbb{M}_n(h)-\sup_{h\in B}\mathbb{M}_n(h)\leadsto \sup_{h\in F\cap A}\mathbb{M}(h)-\sup_{h\in B}\mathbb{M}(h).$$

Then

$$\begin{split} \limsup_{n \to \infty} P^*(\hat{h}_n \in F \cap A) &\leq \limsup_{n \to \infty} P^* \bigg(\sup_{h \in F \cap A} \mathbb{M}_n(h) \geq \sup_{h \in B} \mathbb{M}_n(h) - o_P(1) \bigg) \\ [\text{Slutsky, Portmanteu}] &\leq P \bigg(\sup_{h \in F \cap A} \mathbb{M}(h) \geq \sup_{h \in B} \mathbb{M}(h) \bigg) \\ &\leq P \bigg(\hat{h} \in F \text{ or } \hat{h} \notin B \bigg). \quad \Box \end{split}$$

Theorem (3.2.2, Argmax continuous mapping)

Suppose:

- $\mathbb{M}_n \rightsquigarrow \mathbb{M}$ in $\ell^{\infty}(K)$ for every compact $K \subset H$.
- Almost all paths $h \mapsto \mathbb{M}(h)$ are upper semicontinuous and have a unique maximum at \hat{h} , which is tight as random map in H.
- \hat{h}_n is uniformly tight and $\mathbb{M}_n(\hat{h}_n) \geq \sup_h \mathbb{M}_n(h) o_P(1)$.

Then $\hat{h}_n \rightsquigarrow \hat{h}$ in H.

Proof.

- We apply the Lemma with A = B = K.
- Because \hat{h} is unique and \mathbb{M} upper semicontinuous,

$$\mathbb{M}(\hat{h}) > \sup_{h \notin G, h \in K} \mathbb{M}(h), \quad a.s.$$

• Then

$$\limsup_{n\to\infty} P^*(\hat{h}_n \in F) = \limsup_{n\to\infty} \left(P^*(\hat{h}_n \in F \cap K) + P^*(\hat{h}_n \in F \cap K^c) \right)$$
[Lemma 3.2.1] $\leq P(\hat{h} \in F) + P(\hat{h} \notin K) + \limsup_{n\to\infty} P^*(\hat{h}_n \notin K).$

- Make 2nd and 3rd term arbitrarily small by taking K large enough.
- Then $\hat{h}_n \rightsquigarrow h$ by the Portmanteau theorem.

Consistency

Corollary (3.2.3 i, Consistency)

Let \mathbb{M}_n be indexed by Θ and $\mathbb{M} \colon \Theta \mapsto \mathbb{R}$ deterministic. Suppose:

- $\|\mathbb{M}_n \mathbb{M}\|_{\Theta} \stackrel{P^*}{\to} 0.$
- There is θ_0 with $\mathbb{M}(\theta_0) > \sup_{\theta \notin G}(\theta)$ for open G.
- $\mathbb{M}_n(\hat{\theta}_n) \geq \sup_{\theta} \mathbb{M}_n(\theta) O_P(1)$.

Then $\hat{\theta}_n \stackrel{P^*}{\to} \theta_0$.

Consistency

Corollary (3.2.3 ii, Consistency)

Let \mathbb{M}_n be indexed by Θ and $\mathbb{M} \colon \Theta \mapsto \mathbb{R}$ deterministic. Suppose:

- $\|\mathbb{M}_n \mathbb{M}\|_K \stackrel{P^*}{\to} 0$ for every compact $K \subset \Theta$.
- $\theta \mapsto \mathbb{M}(\theta)$ is upper semicontinuous with unique maximum at θ_0 .
- $\mathbb{M}_n(\hat{\theta}_n) \ge \sup_{\theta} \mathbb{M}_n(\theta) O_P(1)$ and $\hat{\theta}_n$ is uniformly tight.

Then $\hat{\theta}_n \stackrel{P^*}{\to} \theta_0$.

Example: MLE

Let

$$\hat{\theta}_n = \arg\max_{\theta} \mathbb{M}_n(\theta) = \arg\max_{\theta} \mathbb{P}_n m_{\theta} = \arg\max_{\theta} \frac{1}{n} \sum_{i=1}^n \ln p_{\theta}(X_i).$$

- The last corollary gives easy conditions for $\hat{\theta}_n \to_p \theta_0$.
- The Argmax theorem can also be used to derive the asymptotic distribution of $\hat{h}_n = \sqrt{n}(\hat{\theta}_n \theta_0)$.

Example: MLE

• Write $\hat{h}_n = \sqrt{n}(\hat{\theta}_n - \theta_0)$ as

$$\hat{h}_n = \arg\max_h \mathbb{M}'_n(h) = \arg\max_h \mathbb{M}_n(\theta_0 + h/\sqrt{n})$$

$$= \arg\max_h \frac{1}{n} \sum_{i=1}^n \ln p_{\theta_0 + h/\sqrt{n}}(X_i).$$

• If p_{θ} is sufficiently regular,

$$\frac{1}{n}\sum_{i=1}^{n}\ln p_{\theta_0+h/\sqrt{n}}(X_i) = h'\frac{1}{\sqrt{n}}\sum_{i=1}^{n}\nabla_{\theta}\ln p_{\theta_0}(X_i) - \frac{1}{2}h'I_{\theta_0}h + o_P(1).$$

 Under appropriate conditions (Theorems 2.11.22 or 2.11.23), this converges weakly to

$$h\mapsto h'\Delta_{ heta_0}-rac{1}{2}h'I_{ heta_0}h, \qquad \Delta_{ heta_0}\sim \mathcal{N}(0,\Sigma_{ heta_0})$$

• Argmax theorem: $\hat{h}_n \leadsto \hat{h} = I_{\theta_0}^{-1} \Delta_{\theta_0}$.

Outline

- 1. Consistency
- 2. Rate of convergence
- 3. Asymptotic distribution

Theorem (3.2.5)

Let \mathbb{M}_n be a stochastic process indexed by Θ and $\mathbb{M} \colon \Theta \mapsto \mathbb{R}$ deterministic. Suppose:

• for some 'distance' d and every θ in a neighborhood of θ_0 ,

$$\mathbb{M}(\theta) - \mathbb{M}(\theta_0) \lesssim -d^2(\theta, \theta_0).$$

• For every n and δ small,

$$\mathrm{E}^* \sup_{d(\theta,\theta_0)<\delta} \bigl| (\mathbb{M}_n - \mathbb{M})(\theta) - (\mathbb{M}_n - \mathbb{M})(\theta_0) \bigr| \lesssim \frac{\phi_n(\delta)}{\sqrt{n}},$$

for functions ϕ_n with $\delta \mapsto \phi_n(\delta)/\delta^{\alpha}$ decreasing for some $\alpha < 2$.

•
$$r_n^2 \phi_n(r_n^{-1}) < \sqrt{n}$$
, $\forall n$.

•
$$\hat{\theta}_n \to_{P^*} \theta_0$$
 and $\mathbb{M}_n(\hat{\theta}_n) \ge \mathbb{M}_n(\theta_0) - O_P(r_n^{-2})$.

Then $d(\hat{\theta}_n, \theta_0) = O_P^*(r_n^{-1})$.

Proof.

- Let for simplicity $\mathbb{M}_n(\hat{\theta}_n) \geq \sup_{\theta} \mathbb{M}_n(\theta)$.
- For each n, partition $\Theta \setminus \theta_0$ into "shells"

$$S_{j,n} = \{\theta \colon 2^{j-1} < r_n d(\theta, \theta_0) \le 2^j\}, \quad j \in \mathbb{N}.$$

- Suppose $r_n d(\hat{\theta}_n, \theta_0) > 2^M$ for some for some $M \in \mathbb{N}$. Then $\hat{\theta}_n \in S_{j,n}$ for some j > M and $\sup_{\theta \in S_{j,n}} \mathbb{M}_n(\theta) \mathbb{M}_n(\theta_0) \ge 0$.
- $$\begin{split} \bullet \quad & P^* \big(r_n d(\hat{\theta}_n, \theta_0) > 2^M \big) \leq \sum_{j > M} P^* \bigg(\sup_{\theta \in S_{j,n}} \mathbb{M}_n(\theta) \mathbb{M}_n(\theta_0) \geq 0 \bigg) \\ & \leq \sum_{j > M, 2^j \leq \eta r_n} P^* \bigg(\sup_{\theta \in S_{j,n}} \mathbb{M}_n(\theta) \mathbb{M}_n(\theta_0) \geq 0 \bigg) \\ & + P^* \big(2 d(\hat{\theta}_n, \theta_0) \geq \eta \big). \end{split}$$
- Because $\hat{\theta}_n \overset{P^*}{\to} \theta_0$, $P^*(2d(\hat{\theta}_n, \theta_0) \ge \eta) \to 0$ for every $\eta > 0$.

Proof (ct'd).

- Choose η small enough for the conditions of the theorem to hold.
- Then for every $\theta \in S_{j,n} = \{\theta \colon 2^{j-1} < r_n d(\theta, \theta_0) \le 2^j\}$,

$$\mathbb{M}(\theta) - \mathbb{M}(\theta_0) \lesssim -d^2(\theta, \theta_0) < -\left(\frac{2^{j-1}}{r_n}\right)^2.$$

• Defining $W_n = \mathbb{M}_n - \mathbb{M}$, we get

$$\begin{split} & \sum_{j>M,2^j \leq \eta r_n} P^* \bigg(\sup_{\theta \in S_{j,n}} \mathbb{M}_n(\theta) - \mathbb{M}_n(\theta_0) \geq 0 \bigg) \\ = & \sum_{j>M,2^j \leq \eta r_n} P^* \bigg(\sup_{\theta \in S_{j,n}} W_n(\theta) - W_n(\theta_0) \geq -(\mathbb{M}(\theta) - \mathbb{M}(\theta_0)) \bigg) \\ \leq & \sum_{j>M,2^j \leq \eta r_n} P^* \bigg(\sup_{\theta \in S_{j,n}} |W_n(\theta) - W_n(\theta_0)| \gtrsim (2^{j-1}/r_n)^2 \bigg). \end{split}$$

Proof (ct'd).

- Recall that $S_{i,n} \subset \{d(\theta,\theta_0) \leq 2^j/r_n\}$.
- Then Markov's inequality gives

$$\sum_{j>M,2^j\leq \eta r_n} P^* \left(\sup_{\theta\in S_{j,n}} |W_n(\theta) - W_n(\theta_0)| \gtrsim (2^{j-1}/r_n)^2 \right)$$

$$\leq \sum_{j>M} \frac{\phi_n(2^j/r_n)}{\sqrt{n}(2^{j-1}/r_n)^2}.$$

- Observe that $\phi_n(c\delta) \le c^{\alpha}\phi_n(\delta)$ for every c > 1 and recall $r_n^2\phi_n(r_n^{-1}) \le \sqrt{n}$.
- This gives

$$\sum_{j>M} \frac{\phi_n(2^j/r_n)}{\sqrt{n}(2^{j-1}/r_n)^2} \leq \sum_{j>M} \frac{2^{j\alpha}\phi_n(1/r_n)}{\sqrt{n}(2^{j-1}/r_n)^2} \leq \sum_{j>M} \frac{2^{j\alpha}}{(2^{j-1})^2} \stackrel{M \to \infty}{\to} 0,$$

because α < 2.

Rate of convergence: iid case

Corollary (3.2.6)

Let $\mathbb{M}_n(\theta) = \mathbb{P}_n m_{\theta}$, $\mathbb{M}(\theta) = Pm_{\theta}$, $\sqrt{n}(\mathbb{M}_n - \mathbb{M}) = \mathbb{G}_n m_{\theta}$. Suppose:

• For every θ in a neighborhood of θ_0 ,

$$P(m_{\theta}-m_{\theta_0})\lesssim -d^2(\theta,\theta_0).$$

• There is ϕ with $\delta \mapsto \phi(\delta)/\delta^{\alpha}$ for some $\alpha < 2$, and for δ small,

$$\mathrm{E}^* \sup_{d(heta, heta_0)<\delta} |\mathbb{G}_n(m_ heta-m_{ heta_0})| \lesssim \phi(\delta).$$

- $r_n^2 \phi(1/r_n) \leq \sqrt{n}$, $\forall n$.
- $\hat{\theta}_n \to_{P^*} \theta_0$ and $\mathbb{P}_n m_{\hat{\theta}_n} \geq \mathbb{P}_n m_{\theta_0} O_P(r_n^{-2})$.

Then
$$d(\hat{\theta}_n, \theta_0) = O_P^*(r_n^{-1}).$$

Comments

- $\phi(\delta) = \delta^{\alpha}$ gives $r_n \ge n^{1/(4-2\alpha)}$.
- $P(m_{\theta} m_{\theta_0}) \lesssim -d^2(\theta, \theta_0)$ holds if $\theta \mapsto Pm_{\theta}$ has two continuous derivatives (2nd nonsingular).
- To find ϕ , we can bound the continuity modulus

$$\mathrm{E}^* \sup_{d(heta, heta_0)<\delta} |\mathbb{G}_n(m_ heta-m_{ heta_0})|$$

by entropy integrals of $\mathcal{M}_{\delta} = \{m_{\theta} - m_{\theta_0} : d(\theta, \theta_0) < \delta\}$ with respect to an envelope M_{δ} .

- If $J(1, \mathcal{M}_{\delta}, L_2)$ or $J_{[]}(1, \mathcal{M}_{\delta}, L_2(P))$ are bounded as $\delta \swarrow 0$, then we can take $\phi^2(\delta) = P^*M_{\delta}^2$ (Theorems 2.14.1 and 2.14.15).
- Book gives many detailed examples of applications: Lipschitz in a parameter, location estimation, monotone density estimation, current status distribution.

Outline

- 1. Consistency
- 2. Rate of convergence
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Idea: Taylor expand the criterion function.

$$\begin{split} n\mathbb{P}_n(m_{\theta}-m_{\theta_0}) &= nP(m_{\theta}-m_{\theta_0}) + \sqrt{n}\mathbb{G}_n(m_{\theta}-m_{\theta_0}) \\ &\approx \frac{1}{2}\sqrt{n}(\theta-\theta_0)'V\sqrt{n}(\theta-\theta_0) + \sqrt{n}(\theta-\theta_0)\mathbb{G}_n\dot{m}_{\theta_0} \\ &+ o_P(\sqrt{n}\|\theta-\theta_0\|). \end{split}$$

Forgetting about the remainder, this is maximized for

$$\sqrt{n}(\theta-\theta_0)=-V^{-1}\mathbb{G}_n\dot{m}_{\theta_0}.$$

• Thus, we expect the M-estimator $\hat{\theta}_n$ to satisfy

$$\sqrt{n}(\hat{\theta}_n - \theta_0) = -V^{-1}\mathbb{G}_n\dot{m}_{\theta_0} + o_P(1)$$

Theorem (3.2.16)

Let \mathbb{M}_n be index by open $\Theta \subset \mathbb{R}^d$ and $\mathbb{M} \colon \Theta \to \mathbb{R}$ deterministic. Suppose:

- $\theta \mapsto \mathbb{M}(\theta)$ has two continuous derivatives at its maximum θ_0 with non-singular Hessian V.
- For every $\tilde{\theta}_n = \theta_0 + o_P^*(1)$, the 'stochastic differentiability' condition

$$r_n(\mathbb{M}_n - \mathbb{M})(\tilde{\theta}_n) - r_n(\mathbb{M}_n - \mathbb{M})(\theta_0)$$

= $(\tilde{\theta}_n - \theta_0)' Z_n + o_P^*(\|\tilde{\theta}_n - \theta_0\| + r_n\|\tilde{\theta}_n - \theta_0\| + r_n^{-1}),$

holds with some process Z_n uniformly tight.

•
$$\hat{\theta}_n \overset{P^*}{\to} \theta_0$$
 and $\mathbb{M}_n(\hat{\theta}_n) \geq \sup_{\theta} \mathbb{M}_n(\theta) - o_P(r_n^{-2})$.

Then $r_n(\hat{\theta}_n - \theta_0) = -V^{-1}Z_n + o_P^*(1)$.

If $r_n(\hat{\theta}_n - \theta_0)$ is uniformly tight, only consider $\tilde{\theta}_n = \theta_0 + O_P^*(r_n^{-1})$.

Proof.

- Let for simplicity $\mathbb{M}_n(\hat{\theta}_n) \geq \sup_{\theta} \mathbb{M}_n(\theta)$.
- For every $\tilde{h}_n = o_P^*(1)$, stochastic differentiability and smoothness of $\theta \mapsto M(\theta)$ gives

$$\mathbb{M}_{n}(\theta_{0} + \tilde{h}_{n}) - \mathbb{M}_{n}(\theta_{0}) = \frac{\tilde{h}'_{n}V\tilde{h}_{n}}{2} + r_{n}^{-1}\tilde{h}'_{n}Z_{n} + o_{P}^{*}(\|\tilde{h}_{n}\|^{2} + r_{n}^{-1}\|\tilde{h}_{n}\| + r_{n}^{-2}).$$
(*)

- Substitute in $\hat{h}_n = \hat{\theta}_n \theta_0$ for \tilde{h}_n . By definition of $\hat{\theta}_n$, LHS ≥ 0 .
- Because V is negative definite, $\hat{h}'_n V \hat{h}_n \lesssim -\|\hat{h}_n\|^2$.
- Hence, the first display implies

$$0 \leq -\|\hat{h}_n\|^2 + r_n^{-1}\|\hat{h}_n\|O_P(1) + o_P(\|\hat{h}_n\|^2 + r_n^{-2}).$$

$$\Rightarrow \|\hat{h}_n\| = O_P(r_n^{-1}).$$

Proof (ct'd).

• Invoke (*) with $\tilde{h}_n = \hat{h}_n$ and $\tilde{h}_n = -r_n^{-1}V^{-1}Z_n$:

$$\mathbb{M}_{n}(\theta_{0} + \hat{h}_{n}) - \mathbb{M}_{n}(\theta_{0}) = \frac{\hat{h}'_{n}V\hat{h}_{n}}{2} + r_{n}^{-1}\hat{h}'_{n}Z_{n} + o_{P}^{*}(r_{n}^{-2})$$

$$\mathbb{M}_{n}(\theta_{0} - r_{n}^{-1}V^{-1}Z_{n}) - \mathbb{M}_{n}(\theta_{0}) = -\frac{(V^{-1}Z_{n})'V(V^{-1}Z_{n})}{2r_{n}^{2}} + o_{P}^{*}(r_{n}^{-2}).$$

• Subtract the second from the first:

$$\frac{(\hat{h}_n + r_n^{-1}V^{-1}Z_n)'V(\hat{h}_n + r_n^{-1}V^{-1}Z_n)}{2} \ge -o_P(r_n^{-2}).$$

•
$$\Rightarrow \hat{h}_n = -r_n^{-1}V^{-1}Z_n + o_P(r_n^{-1}).$$

$$\bullet \Rightarrow r_n(\hat{\theta}_n - \theta_0) = V^{-1}Z_n + o_P(1).$$

More linearization

- Lemma 3.2.19, 3.2.21, and eq. (3.2.22) give easier conditions to verify "differentiability" in the *iid* case.
- Corollary 3.2.23 summarizes easy conditions to establish asymptotic normality for the Euclidean *iid* case.
- More linearization next week with Z-estimators.