

Chapter 3.8: Two-Sample Problem

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What if?

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Kolmogorov-
Smirnov
tests

Two-Sample
Permutation
EP/Bootstrap

Imagine having two random samples $(X_i)_{i=1}^m$ and $(Y_i)_{i=1}^n$ and wondering

What if?

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Imagine having two random samples $(X_i)_{i=1}^m$ and $(Y_i)_{i=1}^n$ and wondering

- Is the first sample coming from a certain distribution?
- Are the two samples coming from the same distribution?

Overview

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Kolmogorov distribution

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Definition

Kolmogorov distribution is the distribution of the random variable

$$\sup_{t \in [0,1]} |B(t)|$$

where $B(t)$ is the **Brownian bridge**.

Kolmogorov distribution

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Definition

Kolmogorov distribution is the distribution of the random variable

$$\sup_{t \in [0,1]} |B(t)|$$

where $B(t)$ is the **Brownian bridge**. The cumulative distribution function of $\sup_{t \in [0,1]} |B(t)|$ is given by

$$P\left(\sup_{t \in [0,1]} |B(t)| \leq x\right) = 1 - 2 \sum_{k=1}^{+\infty} (-1)^{k-1} e^{-2k^2 x^2}.$$

Kolmogorov-Smirnov test

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Let $(X_i)_{i=1}^m \in \mathbb{R}$. We want to test the null-hypothesis H_0 :
 $(X_i)_{i=1}^m \sim F$ (a cumulative distribution function)

Definition

The Kolmogorov-Smirnov statistic is defined as

$$D_m = \sqrt{m} \sup_x |\mathbb{F}_m - F|$$

where \mathbb{F}_m is the empirical distribution function of the $(X_i)_{i=1}^m$.

Kolmogorov-Smirnov test

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where \mathbb{F}_m is the empirical distribution function of the $(X_i)_{i=1}^m$.

Proposition

Under the null-hypothesis

$$D_m / \sqrt{m} \xrightarrow{a.s.} 0, \text{ and } D_m \rightarrow \sup_{t \in [0,1]} |B(t)|.$$

Kolmogorov-Smirnov test

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Definition

Since

$$P(D_m > c_\alpha) \rightarrow \alpha = 2 \sum_{k=1}^{+\infty} (-1)^{k-1} e^{-2k^2 c_\alpha^2}$$

the K-S test with significance α rejects H_0 if $D_m > c_\alpha$

Kolmogorov-Smirnov test two-sample

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Let $(X_i)_{i=1}^m$ and $(Y_i)_{i=1}^n$ in \mathbb{R} . We want to test the null-hypothesis H_0 : the two samples come from the same distribution

Definition

The two-sample Kolmogorov-Smirnov statistic is defined as

$$D_{n,m} = \sqrt{\frac{nm}{n+m}} \sup_x |\mathbb{F}_m - \mathbb{F}'_n|$$

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Definition

The two-sample Kolmogorov-Smirnov statistic is defined as

$$D_{n,m} = \sqrt{\frac{nm}{n+m}} \sup_x |\mathbb{F}_m - \mathbb{F}'_n|$$

Here again, the K-S test with significance α rejects H_0 if $D_{n,m} > c_\alpha$.

Two-Sample Permutation test

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The two-sample K-S test is quite useful. However, it would be nice to generalize it somehow

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The two-sample K-S test is quite useful. However, it would be nice to generalize it somehow

Let $(X_i)_{i=1}^m$ and $(Y_i)_{i=1}^n$ iid sample from distributions P and Q .

We want to test the null-hypothesis $H_0: P = Q$.

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The two-sample K-S test is quite useful. However, it would be nice to generalize it somehow

Let $(X_i)_{i=1}^m$ and $(Y_i)_{i=1}^n$ iid sample from distributions P and Q . We want to test the null-hypothesis $H_0: P = Q$.

Definition

The Kolmogorov-Smirnov statistic is defined as

$$D_{m,n} = \sqrt{\frac{mn}{m+n}} \|\mathbb{P}_m - \mathbb{Q}_n\|_{\mathcal{F}}$$

where \mathcal{F} is a class of measurable function, and \mathbb{P}_m and \mathbb{Q}_m are empirical measures.

Donsker class

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If \mathcal{F} is Donsker, then $\mathbb{G}_m := \sqrt{m}(\mathbb{P}_m - P) \rightarrow \mathbb{G}_P$ and
 $\mathbb{G}'_n := \sqrt{n}(\mathbb{Q}_n - Q) \rightarrow \mathbb{G}_Q$ (two independent Brownian bridges)

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$$D_{m,n} = \left\| \sqrt{\frac{n}{N}} \mathbb{G}_m - \sqrt{\frac{m}{N}} \mathbb{G}'_n + \sqrt{\frac{mn}{N}} (P - Q) \right\|_{\mathcal{F}}$$

with $N = n + m$

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with $N = n + m$

- If $\|P - Q\|_{\mathcal{F}} > 0$, then $D_{m,n} \rightarrow +\infty$
- If $P = Q$, then $D_{m,n} \rightarrow \|\sqrt{1-\lambda}\mathbb{G}_P - \sqrt{\lambda}\mathbb{G}_Q\|_{\mathcal{F}}$ which has the same distribution as $\|\mathbb{G}_P\|_{\mathcal{P}}$

(with $m/N \rightarrow \lambda$)

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 $\mathbb{G}'_n := \sqrt{n}(\mathbb{Q}_n - Q) \rightarrow \mathbb{G}_Q$ (two independent Brownian bridges)
The test with significance α rejects H_0 if $D_{m,n} > c_{m,n}$ where

$$c_{m,n} \rightarrow c_P := \inf\{t : P(\|\mathbb{G}_P\|_{\mathcal{F}} > t) \leq \alpha\}.$$

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If \mathcal{F} is Donsker, then $\mathbb{G}_m := \sqrt{m}(\mathbb{P}_m - P) \rightarrow \mathbb{G}_P$ and $\mathbb{G}'_n := \sqrt{n}(\mathbb{Q}_n - Q) \rightarrow \mathbb{G}_Q$ (two independent Brownian bridges)
The test with significance α rejects H_0 if $D_{m,n} > c_{m,n}$ where

$$c_{m,n} \rightarrow c_P := \inf\{t : P(\|\mathbb{G}_P\|_{\mathcal{F}} > t) \leq \alpha\}.$$

However, c_P depends the underlying measure, P . It is impossible to find $c_{m,n}$ with the given convergence property for every P in H_0

Pooled empirical measure

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What can we do now?

Pooled empirical measure

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What can we do now?

Answer: Use a **pooled empirical measure**

$$\mathbb{H}_N := \frac{1}{N} \sum_{i=1}^N \delta_{Z_{Ni}} = \lambda_N \mathbb{P}_m + (1 - \lambda_N) \mathbb{Q}_n$$

where $(Z_{Ni})_{i=1}^N$ represents the pooled data $(X_i)_{i=1}^m$ and $(Y_i)_{i=1}^n$ and $\lambda_N = m/N$.

Pooled empirical measure

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What can we do now?

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where $(Z_{Ni})_{i=1}^N$ represents the pooled data $(X_i)_{i=1}^m$ and $(Y_i)_{i=1}^n$ and $\lambda_N = m/N$.

$$\mathbb{P}_m - \mathbb{H}_N = (1 - \lambda_N)(\mathbb{P}_m - \mathbb{Q}_n)$$

Permutation Empirical Process

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If we have been attentive, we might have noticed something

Permutation Empirical Process

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If we have been attentive, we might have noticed something
Let a random vector $R = (R_i)_{i=1}^n$ from the uniform distribution
on the set of permutations, we can define

$$\tilde{\mathbb{P}}_{m,N} = \frac{1}{m} \sum_{i=1}^m \delta_{Z_{NR_i}}, \quad \tilde{\mathbb{Q}}_{n,N} = \frac{1}{n} \sum_{i=m+1}^N \delta_{Z_{NR_i}}.$$

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then

$$\tilde{D}_{m,n} = \sqrt{\frac{mn}{m+n}} \|\tilde{\mathbb{P}}_{m,N} - \tilde{\mathbb{Q}}_{n,N}\|_{\mathcal{F}}$$

The test with significance α rejects H_0 if $\tilde{D}_{m,n} > \tilde{c}_{m,n}$ where

$$\tilde{c}_{m,n} \rightarrow c_H := \inf\{t : P(\|\mathbb{G}_H\|_{\mathcal{F}} > t) \leq \alpha\}.$$

Permutation Theorem

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Theorem

Let \mathcal{F} be a class of measurable functions that is Donsker under both P and Q and satisfies both $\|P\|_{\mathcal{F}} < \infty$ and $\|Q\|_{\mathcal{F}} < \infty$. If $m, n \rightarrow +\infty$ such that $m/N \rightarrow \lambda \in (0, 1)$, then

$$\sqrt{m}(\tilde{\mathbb{P}}_{m,N} - \mathbb{H}_N) \rightsquigarrow \sqrt{1-\lambda}\mathbb{G}_H$$

given $X_1, X_2, \dots, Y_1, Y_2, \dots$ in probability. Here \mathbb{G}_H is a tight Brownian bridge process corresponding to the measure $H = \lambda P + (1 - \lambda)Q$.

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given $X_1, X_2, \dots, Y_1, Y_2, \dots$ in probability. Here \mathbb{G}_H is a tight Brownian bridge process corresponding to the measure $H = \lambda P + (1 - \lambda)Q$.

*If in addition \mathcal{F} possesses an envelope function F with both $P^*F^2 < \infty$ and $Q^*F^2 < \infty$, then also*

$$\sqrt{m}(\tilde{\mathbb{P}}_{m,N} - \mathbb{H}_N) \rightsquigarrow \sqrt{1 - \lambda} \mathbb{G}_H$$

given almost every sequence $X_1, X_2, \dots, Y_1, Y_2, \dots$

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Main idea of the proof:
(wlog $Pf = 0$)

Step 1:

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Main idea of the proof:
(wlog $Pf = 0$)

Step 1:

- Use the CLT for two-sample rank statistics, to show for all f with $Hf = 0$ and $Hf^2 < \infty$

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Main idea of the proof:

(wlog $Pf = 0$)

Step 1:

- Use the CLT for two-sample rank statistics, to show for all f with $Hf = 0$ and $Hf^2 < \infty$
- $\frac{1}{\sqrt{m}} \sum_{i=1}^m (f(Z_{NR_i}) - \mathbb{H}_N f) \rightsquigarrow N(0, (1 - \lambda)Hf^2)$ given almost every sequence $X_1, X_2, \dots, Y_1, Y_2, \dots$

CLT for two-sample rank statistics

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Let $n > 0$, let $(a_i)_{i=1}^n$ and $(b_i)_{i=1}^n$ be real numbers, and let $R = (R_i)_{i=1}^n$ from the uniform distribution on the set of permutations. Consider the rank statistic

$$S_n = \sum_{i=1}^n b_i a_{R_i}.$$

The mean and variance of S_n are equal to $ES_n = n\bar{a}\bar{b}$ and $V(S_n) = A_n^2 B_n^2 / (n-1)$, where A_n^2 and B_n^2 are the sums of squared deviations from the respective means of the a 's and b 's

CLT for two-sample rank statistics

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Proposition (Rank central limit theorem)

Suppose that

$$\max_{1 \leq n} \frac{|a_i - \bar{a}|}{A_n} \rightarrow 0, \quad \max_{1 \leq n} \frac{|b_i - \bar{b}|}{B_n} \rightarrow 0$$

Then the sequence $(S_n - ES_n)/\sigma(S_n)$ converges in distribution to a standard normal distribution iff

$$\sum_{\sqrt{n}|a_i - \bar{a}| |b_j - \bar{b}| > \epsilon A_n B_n} \frac{|a_i - \bar{a}|^2 |b_j - \bar{b}|^2}{A_n^2 B_n^2} \rightarrow 0$$

for every $\epsilon > 0$

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Step 1:

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Hoeffding's inequality

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Proposition

Let $\{c_i\}_{i=1}^N$ be elements of an arbitrary vector space V , and let $(U_i)_{i=1}^n$ and $(V_i)_{i=1}^n$ denote samples of size $n \leq N$ drawn without and with replacement, respectively, from $\{c_i\}_{i=1}^N$. Then, for every convex function ϕ from V to \mathbb{R}

$$E\phi\left(\sum_{i=1}^n U_i\right) \leq E\phi\left(\sum_{i=1}^n V_i\right)$$

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Step 2:

- Let $(\hat{Z}_{Ni})_{i=1}^N$ an iid sample from \mathbb{H}_N

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Step 2:

- Let $(\hat{Z}_{Ni})_{i=1}^N$ an iid sample from \mathbb{H}_N
- Set $\mathcal{F}_\delta = \{f - g : f, g \in \mathcal{F}; H(f - g)^2 < \delta^2\}$

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Step 2:

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- Set $\mathcal{F}_\delta = \{f - g : f, g \in \mathcal{F}; H(f - g)^2 < \delta^2\}$
- Use Hoeffding's inequality

$$\begin{aligned} E_R \|\sqrt{m}(\tilde{\mathbb{P}}_{m,N} - \mathbb{H}_N)\|_{\mathcal{F}_\delta} &= E_R \left\| \frac{1}{\sqrt{m}} \sum_{i=1}^m (\delta_{Z_{R_i}} - \mathbb{H}_N) \right\|_{\mathcal{F}_\delta} \\ &\leq E_{\hat{Z}} \left\| \frac{1}{\sqrt{m}} \sum_{i=1}^m (\delta_{\hat{Z}_{Ni}} - \mathbb{H}_N) \right\|_{\mathcal{F}_\delta} \end{aligned}$$

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Step 3: Sequence of inequalities

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Step 3: Sequence of inequalities

Poissonization inequality with $(\tilde{N}_i)_{i=1}^N$ iid symetrized $P(m/N)$

$$\begin{aligned} E_{\hat{Z}} \left\| \frac{1}{\sqrt{m}} \sum_{i=1}^m (\delta_{\hat{Z}_{N_i}} - \mathbb{H}_N) \right\|_{\mathcal{F}_\delta} &\leq 4 E_{\tilde{N}} \left\| \frac{1}{\sqrt{m}} \sum_{i=1}^N \tilde{N}_i \delta_{Z_{R_i}} \right\|_{\mathcal{F}_\delta} \\ &\lesssim E_{\tilde{N}} \left\| \frac{1}{\sqrt{m}} \sum_{i=1}^m \tilde{N}_i \delta_{X_i} \right\|_{\mathcal{F}_\delta} \\ &\quad + E_{\tilde{N}} \left\| \frac{1}{\sqrt{m}} \sum_{i=1}^n \tilde{N}_i \delta_{Y_i} \right\|_{\mathcal{F}_\delta} \end{aligned}$$

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Step 3: Sequence of inequalities

By Theorem 2.9.2 [Multiplier Central Limit Theorems], the $m^{1/2} \sum_{i=1}^m \tilde{N}_i \delta_{X_i} \rightarrow Z_P$ and $n^{1/2} \sum_{i=1}^n \tilde{N}_i \delta_{Y_i} \rightarrow Z_P$, two tight Gaussian processes.

$$\begin{aligned} E_{\hat{Z}} \left\| \frac{1}{\sqrt{m}} \sum_{i=1}^m (\delta_{\hat{Z}_{Ni}} - \mathbb{H}_N) \right\|_{\mathcal{F}_\delta} &\leq 4E_{\tilde{N}} \left\| \frac{1}{\sqrt{m}} \sum_{i=1}^m \tilde{N}_i \delta_{Z_{R_i}} \right\|_{\mathcal{F}_\delta} \\ &\lesssim E_{\tilde{N}} \left\| \frac{1}{\sqrt{m}} \sum_{i=1}^m \tilde{N}_i \delta_{X_i} \right\|_{\mathcal{F}_\delta} \\ &\quad + E_{\tilde{N}} \left\| \frac{1}{\sqrt{m}} \sum_{i=1}^n \tilde{N}_i \delta_{Y_i} \right\|_{\mathcal{F}_\delta} \end{aligned}$$

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Step 3: Sequence of inequalities

We finally manage to have this bound

$$E_{\tilde{N}} \left\| \frac{1}{\sqrt{m}} \sum_{i=1}^m \tilde{N}_i \delta_{X_i} \right\|_{\mathcal{F}_\delta} + E_{\tilde{N}} \left\| \frac{1}{\sqrt{m}} \sum_{i=1}^n \tilde{N}_i \delta_{Y_i} \right\|_{\mathcal{F}_\delta}$$

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Step 3: Sequence of inequalities

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By Theorem 2.9.2 [Multiplier Central Limit Theorems], the $m^{1/2} \sum_{i=1}^m \tilde{N}_i \delta_{X_i} \rightarrow Z_P$ and $n^{1/2} \sum_{i=1}^n \tilde{N}_i \delta_{Y_i} \rightarrow Z_P$, two tight Gaussian processes.

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Step 3: Sequence of inequalities

We finally manage to have this bound

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By Theorem 2.9.2 [Multiplier Central Limit Theorems], the $m^{1/2} \sum_{i=1}^m \tilde{N}_i \delta_{X_i} \rightarrow Z_P$ and $n^{1/2} \sum_{i=1}^n \tilde{N}_i \delta_{Y_i} \rightarrow Z_P$, two tight Gaussian processes.

By Lemma 2.3.11 [Symmetrization and Measurability], the outer expectation is asymptotically bounded by a multiple $E \|Z_P\|_{\mathcal{F}_\delta} + \sqrt{(1-\lambda)/\lambda} E \|Z_Q\|_{\mathcal{F}_\delta} \rightarrow 0$ as $\delta \rightarrow 0$.

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Can we use bootstrap to this method?

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Can we use bootstrap to this method?

Of course!

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Can we use bootstrap to this method?

Of course! Let $\hat{Z} = (\hat{Z}_i)_{i=1}^N$ iid sample from \mathbb{H}_N . The
two-sample bootstrap empirical measures

$$\hat{\mathbb{P}}_{m,N} = \frac{1}{m} \sum_{i=1}^m \delta_{\hat{Z}_i}, \quad \hat{\mathbb{Q}}_{n,N} = \frac{1}{n} \sum_{i=m+1}^N \delta_{\hat{Z}_{m+i}}.$$

Two-Sample Bootstrap

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then

$$\hat{D}_{m,n} = \sqrt{\frac{mn}{m+n}} \|\hat{\mathbb{P}}_{m,N} - \hat{\mathbb{Q}}_{n,N}\|_{\mathcal{F}}$$

The test with significance α rejects H_0 if $\hat{D}_{m,n} > \hat{c}_{m,n}$ where

$$\hat{c}_{m,n} \rightarrow c_H := \inf\{t : P(\|\mathbb{G}_H\|_{\mathcal{F}} > t) \leq \alpha\}.$$

Two-Sample Bootstrap Theorem

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Theorem

Let \mathcal{F} be a class of measurable functions that is Donsker under both P and Q and satisfies both $\|P\|_{\mathcal{F}} < \infty$ and $\|Q\|_{\mathcal{F}} < \infty$. If $m, n \rightarrow +\infty$ such that $m/N \rightarrow \lambda \in (0, 1)$, then

$$\sqrt{m}(\hat{\mathbb{P}}_{m,N} - \mathbb{H}_N) \rightsquigarrow \sqrt{1-\lambda}\mathbb{G}_H$$

given $X_1, X_2, \dots, Y_1, Y_2, \dots$ in probability. Here \mathbb{G}_H is a tight Brownian bridge process corresponding to the measure $H = \lambda P + (1 - \lambda)Q$.

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*If in addition \mathcal{F} possesses an envelope function F with both $P^*F^2 < \infty$ and $Q^*F^2 < \infty$, then also*

$$\sqrt{m}(\hat{\mathbb{P}}_{m,N} - \mathbb{H}_N) \rightsquigarrow \sqrt{1 - \lambda} \mathbb{G}_H$$

given almost every sequence $X_1, X_2, \dots, Y_1, Y_2, \dots$

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Main idea of the proof:

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Main idea of the proof:
Exactly the same as the previous one.

Two-Sample Bootstrap Theorem

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Main idea of the proof:

Exactly the same as the previous one. Actually, even simpler because no need for Hoeffding's inequality.

- K-S test is a special case of the test in this Chapter
- The previous results can easily be generalized for a k -sample problem
- They can also be used to test independence in iid pairs of variable

Thank you

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Any questions?