

<div data-bbox="154 48 691 126" data-label="Section-Header"> <p>Chapter 4</p> <p>Location-Scale-Based Parametric Distributions</p> </div> <div data-bbox="107 367 656 420" data-label="Text"> <p>William Q. Meeker and Luis A. Escobar Iowa State University and Louisiana State University</p> </div> <div data-bbox="107 455 740 537" data-label="Text"> <p>Copyright 1998-2004 W. Q. Meeker and L. A. Escobar. Based on the authors' text <i>Statistical Methods for Reliability Data</i>, John Wiley & Sons Inc. 1998.</p> </div> <div data-bbox="337 571 508 623" data-label="Text"> <p>January 5, 2006 19h 13min</p> </div> <div data-bbox="683 619 714 636" data-label="Text"> <p>4 - 1</p> </div>	<div data-bbox="967 48 1503 126" data-label="Section-Header"> <p>Chapter 4</p> <p>Location-Scale-Based Parametric Distributions Objectives</p> </div> <div data-bbox="915 178 1552 611" data-label="List-Group"> <ul style="list-style-type: none"> • Explain importance of parametric models in the analysis of reliability data. • Define important functions of model parameter that are of interest in reliability studies. • Introduce the location-scale and log-location-scale families of distributions. • Describe the properties of the exponential distribution. • Describe the Weibull and lognormal distributions and the related underlying location-scale distributions. </div> <div data-bbox="1495 619 1526 636" data-label="Text"> <p>4 - 2</p> </div>
<div data-bbox="230 781 617 802" data-label="Section-Header"> <p>Motivation for Parametric Models</p> </div> <div data-bbox="103 854 737 1186" data-label="List-Group"> <ul style="list-style-type: none"> • Complements nonparametric techniques. • Parametric models can be described concisely with just a few parameters, instead of having to report an entire curve. • It is possible to use a parametric model to extrapolate (in time) to the lower or upper tail of a distribution. • Parametric models provide smooth estimates of failure-time distributions. </div> <div data-bbox="107 1239 737 1291" data-label="Text"> <p>In practice it is often useful to compare various parametric and nonparametric analyses of a data set.</p> </div> <div data-bbox="683 1318 714 1335" data-label="Text"> <p>4 - 3</p> </div>	<div data-bbox="1071 829 1398 850" data-label="Section-Header"> <p>Functions of the Parameters</p> </div> <div data-bbox="915 903 1494 1163" data-label="List-Group"> <ul style="list-style-type: none"> • Cumulative distribution function (cdf) of T $F(t; \theta) = \Pr(T \leq t), \quad t > 0.$ • The p quantile of T is the smallest value t_p such that $F(t_p; \theta) \geq p.$ • Hazard function of T </div> <div data-bbox="1101 1171 1383 1226" data-label="Equation-Block"> $h(t; \theta) = \frac{f(t; \theta)}{1 - F(t; \theta)}, \quad t > 0.$ </div> <div data-bbox="1495 1318 1526 1335" data-label="Text"> <p>4 - 4</p> </div>
<div data-bbox="198 1465 649 1486" data-label="Section-Header"> <p>Functions of the Parameters-Continued</p> </div> <div data-bbox="103 1539 737 1591" data-label="List-Group"> <ul style="list-style-type: none"> • The mean time to failure, MTTF, of T (also known as expectation of T) </div> <div data-bbox="209 1602 651 1650" data-label="Equation-Block"> $E(T) = \int_0^\infty t f(t; \theta) dt = \int_0^\infty [1 - F(t; \theta)] dt.$ </div> <div data-bbox="123 1654 737 1707" data-label="Text"> <p>If $\int_0^\infty t f(t; \theta) dt = \infty$, we say that the mean of T does not exist.</p> </div> <div data-bbox="103 1730 737 1782" data-label="List-Group"> <ul style="list-style-type: none"> • The variance (or the second central moment) of T and the standard deviation </div> <div data-bbox="251 1791 610 1873" data-label="Equation-Block"> $\begin{aligned} \text{Var}(T) &= \int_0^\infty [t - E(T)]^2 f(t; \theta) dt \\ \text{SD}(T) &= \sqrt{\text{Var}(T)}. \end{aligned}$ </div> <div data-bbox="103 1898 393 1925" data-label="List-Group"> <ul style="list-style-type: none"> • Coefficient of variation γ_2 </div> <div data-bbox="365 1936 496 1990" data-label="Equation-Block"> $\gamma_2 = \frac{\text{SD}(T)}{E(T)}.$ </div> <div data-bbox="683 2018 714 2034" data-label="Text"> <p>4 - 5</p> </div>	<div data-bbox="1071 1535 1398 1556" data-label="Section-Header"> <p>Location-Scale Distributions</p> </div> <div data-bbox="920 1591 1549 1644" data-label="Text"> <p>Y belongs to the location-scale family of distributions if the cdf of Y can be expressed as</p> </div> <div data-bbox="974 1654 1494 1703" data-label="Equation-Block"> $F(y; \mu, \sigma) = \Pr(Y \leq y) = \Phi\left(\frac{y - \mu}{\sigma}\right), \quad -\infty < y < \infty$ </div> <div data-bbox="920 1709 1549 1761" data-label="Text"> <p>where $-\infty < \mu < \infty$ is a location parameter and $\sigma > 0$ is a scale parameter.</p> </div> <div data-bbox="920 1797 1549 1850" data-label="Text"> <p>Φ is the cdf of Y when $\mu = 0$ and $\sigma = 1$ and Φ does not depend on any unknown parameters.</p> </div> <div data-bbox="920 1885 1549 1938" data-label="Text"> <p>Note: The distribution of $Z = (Y - \mu)/\sigma$ does not depend on any unknown parameters.</p> </div> <div data-bbox="1495 2018 1526 2034" data-label="Text"> <p>4 - 6</p> </div>

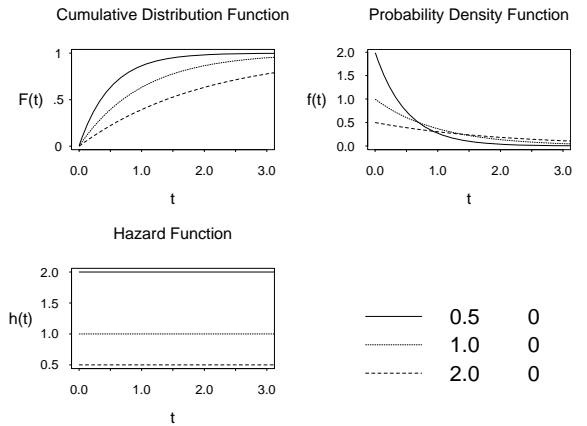
Importance of Location-Scale Distributions

Importance due to:

- Most widely used statistical distributions are either members of this class or closely related to this class of distributions: exponential, normal, Weibull, lognormal, loglogistic, logistic, and extreme value distributions.
- Methods of inference, statistical theory, and computer software generated for the general family can be applied to this large, important class of models.
- Theory for location-scale distributions is relatively simple.

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Examples of Exponential Distributions



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Exponential Distribution

For $T \sim \text{EXP}(\theta, \gamma)$,

$$F(t; \theta, \gamma) = 1 - \exp\left(-\frac{t - \gamma}{\theta}\right)$$

$$f(t; \theta, \gamma) = \frac{1}{\theta} \exp\left(-\frac{t - \gamma}{\theta}\right)$$

$$h(t; \theta, \gamma) = \frac{f(t; \theta, \gamma)}{1 - F(t; \theta, \gamma)} = \frac{1}{\theta}, \quad t > \gamma,$$

where $\theta > 0$ is a scale parameter and γ is both a location and a threshold parameter. When $\gamma = 0$ one gets the well-known one-parameter exponential distribution.

Quantiles: $t_p = \gamma - \theta \log(1 - p)$.

Moments: For integer $m > 0$, $E[(T - \gamma)^m] = m! \theta^m$. Then

$$E(T) = \gamma + \theta, \quad \text{Var}(T) = \theta^2.$$

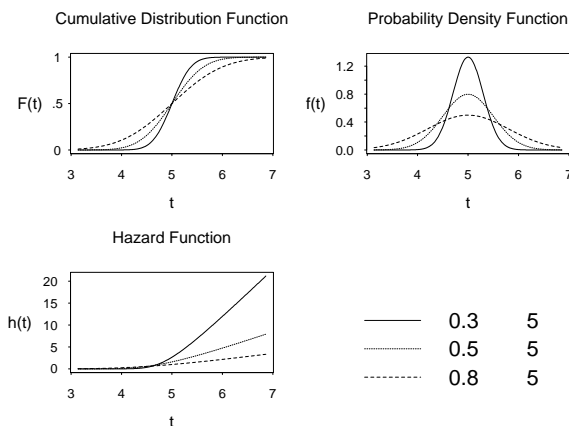
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Motivation for the Exponential Distribution

- Simplest distribution used in the analysis of reliability data.
- Has the important characteristic that its hazard function is constant (does not depend on time t).
- Popular distribution for some kinds of electronic components (e.g., capacitors or robust, high-quality integrated circuits).
- This distribution would *not* be appropriate for a population of electronic components having failure-causing quality-defects.
- Might be useful to describe failure times for components that exhibit physical wearout only after expected technological life of the system in which the component would be installed.

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Examples of Normal Distributions



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Normal (Gaussian) Distribution

For $Y \sim \text{NOR}(\mu, \sigma)$

$$F(y; \mu, \sigma) = \Phi_{\text{nor}}\left(\frac{y - \mu}{\sigma}\right)$$

$$f(y; \mu, \sigma) = \frac{1}{\sigma} \phi_{\text{nor}}\left(\frac{y - \mu}{\sigma}\right), \quad -\infty < y < \infty.$$

where $\phi_{\text{nor}}(z) = (1/\sqrt{2\pi}) \exp(-z^2/2)$ and $\Phi_{\text{nor}}(z) = \int_{-\infty}^z \phi_{\text{nor}}(w) dw$ are pdf and cdf for a standardized normal ($\mu = 0, \sigma = 1$). $-\infty < \mu < \infty$ is a location parameter; $\sigma > 0$ is a scale parameter.

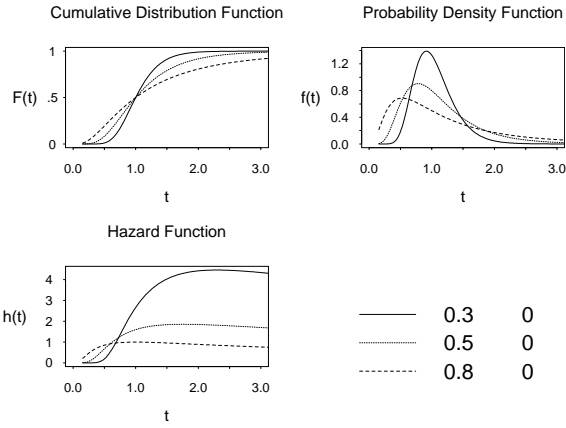
Quantiles: $y_p = \mu + \sigma \Phi_{\text{nor}}^{-1}(p)$ where $\Phi_{\text{nor}}^{-1}(p)$ is the p quantile for a standardized normal.

Moments: For integer $m > 0$, $E[(Y - \mu)^m] = 0$ if m is odd, and $E[(Y - \mu)^m] = (m)! \sigma^m / [2^{m/2} (m/2)!]$ if m is even. Thus

$$E(Y) = \mu, \quad \text{Var}(Y) = \sigma^2.$$

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Examples of Lognormal Distributions



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Lognormal Distribution

If $T \sim \text{LOGNOR}(\mu, \sigma)$ then $\log(T) \sim \text{NOR}(\mu, \sigma)$ with

$$F(t; \mu, \sigma) = \Phi_{\text{nor}} \left[\frac{\log(t) - \mu}{\sigma} \right]$$

$$f(t; \mu, \sigma) = \frac{1}{\sigma t} \phi_{\text{nor}} \left[\frac{\log(t) - \mu}{\sigma} \right], \quad t > 0.$$

ϕ_{nor} and Φ_{nor} are pdf and cdf for a standardized normal. $\exp(\mu)$ is a scale parameter; $\sigma > 0$ is a shape parameter.

Quantiles: $t_p = \exp(\mu + \sigma \Phi_{\text{nor}}^{-1}(p))$, where $\Phi_{\text{nor}}^{-1}(p)$ is the p quantile for a standardized normal.

Moments: For integer $m > 0$, $E(T^m) = \exp(m\mu + m^2\sigma^2/2)$.

$$E(T) = \exp(\mu + \sigma^2/2), \quad \text{Var}(T) = \exp(2\mu + \sigma^2) [\exp(\sigma^2) - 1].$$

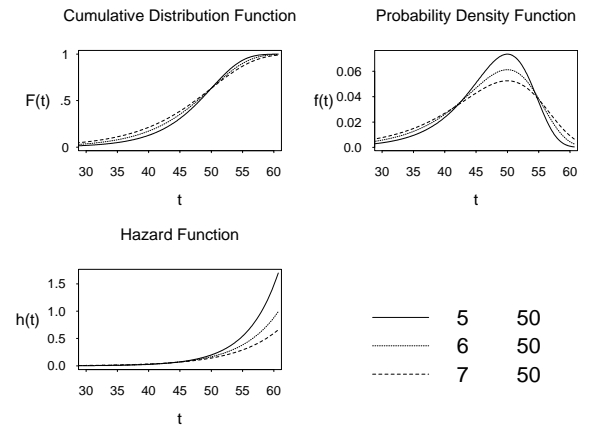
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Motivation for Lognormal Distribution

- The lognormal distribution is a common model for failure times.
- It can be justified for a random variable that arises from the product of a number of identically distributed independent positive random quantities.
- It has been suggested as an appropriate model for failure time caused by a degradation process with combinations of random rates that combine multiplicatively.
- Widely used to describe time to fracture from fatigue crack growth in metals.
- Useful in modeling failure time of a population electronic components with a decreasing hazard function (due to a small proportion of defects in the population).
- Useful for describing the failure-time distribution of certain degradation processes.

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Examples of Smallest Extreme Value Distributions



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Smallest Extreme Value Distribution

For $Y \sim \text{SEV}(\mu, \sigma)$,

$$F(y; \mu, \sigma) = \Phi_{\text{sev}} \left(\frac{y - \mu}{\sigma} \right)$$

$$f(y; \mu, \sigma) = \frac{1}{\sigma} \phi_{\text{sev}} \left(\frac{y - \mu}{\sigma} \right)$$

$$h(y; \mu, \sigma) = \frac{1}{\sigma} \exp \left(\frac{y - \mu}{\sigma} \right), \quad -\infty < y < \infty.$$

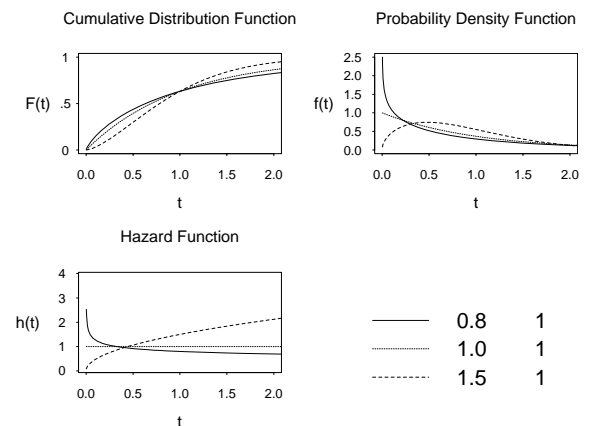
$\Phi_{\text{sev}}(z) = 1 - \exp[-\exp(z)]$, $\phi_{\text{sev}}(z) = \exp[z - \exp(z)]$ are cdf and pdf for standardized SEV ($\mu = 0, \sigma = 1$). $-\infty < \mu < \infty$ is a location parameter and $\sigma > 0$ is a scale parameter.

Quantiles: $y_p = \mu + \Phi_{\text{sev}}^{-1}(p)\sigma = \mu + \log[-\log(1 - p)]\sigma$.

Mean and Variance: $E(Y) = \mu - \sigma\gamma$, $\text{Var}(Y) = \sigma^2\pi^2/6$, where $\gamma \approx .5772$, $\pi \approx 3.1416$.

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Examples of Weibull Distributions



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Weibull Distribution

Common Parameterization:

$$F(t) = \Pr(T \leq t) = 1 - \exp \left[- \left(\frac{t}{\eta} \right)^\beta \right]$$

$$f(t) = \frac{\beta}{\eta} \left(\frac{t}{\eta} \right)^{\beta-1} \exp \left[- \left(\frac{t}{\eta} \right)^\beta \right]$$

$$h(t) = \frac{\beta}{\eta} \left(\frac{t}{\eta} \right)^{\beta-1}, \quad t > 0$$

$\beta > 0$ is shape parameter; $\eta > 0$ is scale parameter.

Quantiles: $t_p = \eta [-\log(1-p)]^{1/\beta}$.

Moments: For integer $m > 0$, $E(T^m) = \eta^m \Gamma(1+m/\beta)$. Then

$$E(T) = \eta \Gamma \left(1 + \frac{1}{\beta} \right), \quad \text{Var}(T) = \eta^2 \left[\Gamma \left(1 + \frac{2}{\beta} \right) - \Gamma^2 \left(1 + \frac{1}{\beta} \right) \right]$$

where $\Gamma(\kappa) = \int_0^\infty w^{\kappa-1} \exp(-w) dw$ is the gamma function.

Note: When $\beta = 1$ then $T \sim \text{EXP}(\eta)$.

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Alternative Weibull Parameterization

Note: If $T \sim \text{WEIB}(\mu, \sigma)$ then $Y = \log(T) \sim \text{SEV}(\mu, \sigma)$.

For $T \sim \text{WEIB}(\mu, \sigma)$ then

$$F(t; \mu, \sigma) = 1 - \exp \left[- \left(\frac{t}{\eta} \right)^\beta \right] = \Phi_{\text{sev}} \left[\frac{\log(t) - \mu}{\sigma} \right]$$

$$f(t; \mu, \sigma) = \frac{\beta}{\eta} \left(\frac{t}{\eta} \right)^{\beta-1} \exp \left[- \left(\frac{t}{\eta} \right)^\beta \right] = \frac{1}{\sigma t} \phi_{\text{sev}} \left[\frac{\log(t) - \mu}{\sigma} \right]$$

where $\sigma = 1/\beta$, $\mu = \log(\eta)$, and

$$\phi_{\text{sev}}(z) = \exp[z - \exp(z)]$$

$$\Phi_{\text{sev}}(z) = 1 - \exp[-\exp(z)].$$

Quantiles:

$$t_p = \eta [-\log(1-p)]^{1/\beta} = \exp \left[\mu + \sigma \Phi_{\text{sev}}^{-1}(p) \right]$$

where $\Phi_{\text{sev}}^{-1}(p)$ is the p quantile for a standardized SEV (i.e., $\mu = 0, \sigma = 1$).

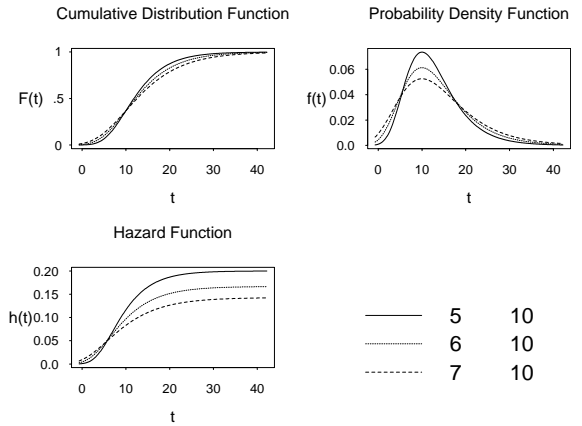
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Motivation for the Weibull Distribution

- The theory of extreme values shows that the Weibull distribution can be used to model the minimum of a large number of independent positive random variables from a certain class of distributions.
 - Failure of the weakest link in a chain with many links with failure mechanisms (e.g., creep or fatigue) in each link acting approximately independent.
 - Failure of a system with a large number of components in series and with approximately independent failure mechanisms in each component.
- The more common justification for its use is empirical: the Weibull distribution can be used to model failure-time data with a decreasing or an increasing hazard function.

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Examples of Largest Extreme Value Distributions



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Largest Extreme Value Distribution

When $Y \sim \text{LEV}(\mu, \sigma)$,

$$F(y; \mu, \sigma) = \Phi_{\text{lev}} \left(\frac{y - \mu}{\sigma} \right)$$

$$f(y; \mu, \sigma) = \frac{1}{\sigma} \phi_{\text{lev}} \left(\frac{y - \mu}{\sigma} \right)$$

$$h(y; \mu, \sigma) = \frac{\exp \left(- \frac{y - \mu}{\sigma} \right)}{\sigma \left\{ \exp \left[\exp \left(- \frac{y - \mu}{\sigma} \right) \right] - 1 \right\}}, \quad -\infty < y < \infty.$$

where $\Phi_{\text{lev}}(z) = \exp[-\exp(-z)]$ and $\phi_{\text{lev}}(z) = \exp[-z - \exp(-z)]$ are the cdf and pdf for a standardized LEV ($\mu = 0, \sigma = 1$) distribution.

$-\infty < \mu < \infty$ is a location parameter and $\sigma > 0$ is a scale parameter.

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Largest Extreme Value Distribution - Continued

Quantiles: $y_p = \mu - \sigma \log [-\log(p)]$.

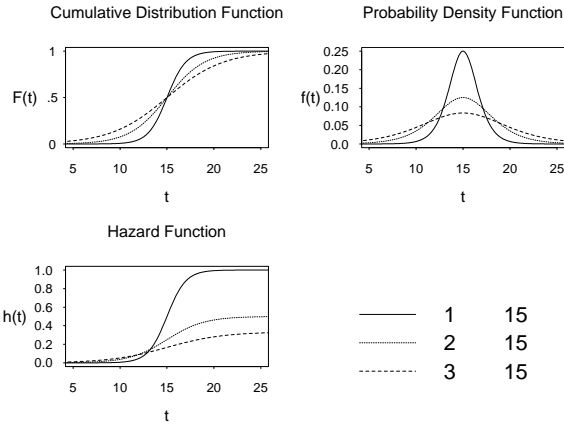
Mean and Variance: $E(Y) = \mu + \sigma \gamma$, $\text{Var}(Y) = \sigma^2 \pi^2/6$, where $\gamma \approx .5772$, $\pi \approx 3.1416$.

Notes:

- The hazard is increasing but is bounded in the sense that $\lim_{y \rightarrow \infty} h(y; \mu, \sigma) = 1/\sigma$.
- If $Y \sim \text{LEV}(\mu, \sigma)$ then $-Y \sim \text{SEV}(-\mu, \sigma)$.

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Examples of Logistic Distributions



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Logistic Distribution

For $Y \sim \text{LOGIS}(\mu, \sigma)$,

$$F(y; \mu, \sigma) = \Phi_{\text{logis}}\left(\frac{y - \mu}{\sigma}\right)$$

$$f(y; \mu, \sigma) = \frac{1}{\sigma} \phi_{\text{logis}}\left(\frac{y - \mu}{\sigma}\right)$$

$$h(y; \mu, \sigma) = \frac{1}{\sigma} \Phi_{\text{logis}}\left(\frac{y - \mu}{\sigma}\right), \quad -\infty < y < \infty.$$

$-\infty < \mu < \infty$ is a location parameter; $\sigma > 0$ is a scale parameter.

ϕ_{logis} and Φ_{logis} are pdf and cdf for a standardized logistic distribution defined by

$$\phi_{\text{logis}}(z) = \frac{\exp(z)}{[1 + \exp(z)]^2}$$

$$\Phi_{\text{logis}}(z) = \frac{\exp(z)}{1 + \exp(z)}.$$

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Logistic Distribution-Continued

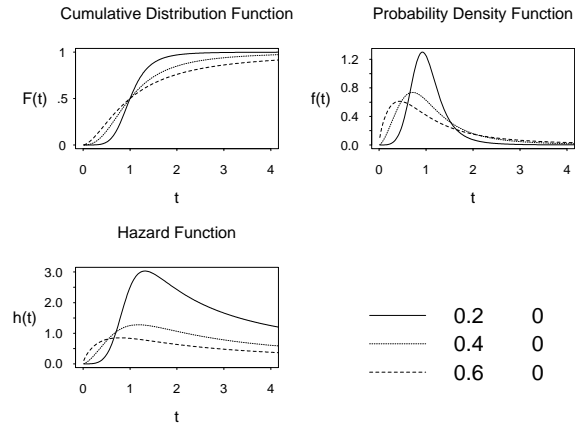
Quantiles: $y_p = \mu + \sigma \Phi_{\text{logis}}^{-1}(p) = \mu + \sigma \log\left(\frac{p}{1-p}\right)$, where $\Phi_{\text{logis}}^{-1}(p) = \log[p/(1-p)]$ is the p quantile for a standardized logistic distribution.

Moments: For integer $m > 0$, $E[(Y - \mu)^m] = 0$ if m is odd, and $E[(Y - \mu)^m] = 2\sigma^m (m!) \left[1 - (1/2)^{m-1}\right] \sum_{i=1}^{\infty} (1/i)^m$ if m is even. Thus

$$E(Y) = \mu, \quad \text{Var}(Y) = \frac{\sigma^2 \pi^2}{3}.$$

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Examples of Loglogistic Distributions



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Loglogistic Distribution

If $Y \sim \text{LOGIS}(\mu, \sigma)$ then $T = \exp(Y) \sim \text{LOGLOGIS}(\mu, \sigma)$ with

$$F(t; \mu, \sigma) = \Phi_{\text{logis}}\left[\frac{\log(t) - \mu}{\sigma}\right]$$

$$f(t; \mu, \sigma) = \frac{1}{\sigma t} \phi_{\text{logis}}\left[\frac{\log(t) - \mu}{\sigma}\right]$$

$$h(t; \mu, \sigma) = \frac{1}{\sigma t} \Phi_{\text{logis}}\left[\frac{\log(t) - \mu}{\sigma}\right], \quad t > 0.$$

$\exp(\mu)$ is a scale parameter; $\sigma > 0$ is a shape parameter. Φ_{logis} and ϕ_{logis} are cdf and pdf for a $\text{LOGIS}(0, 1)$.

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Loglogistic Distribution-Continued

Quantiles: $t_p = \exp\left[\mu + \sigma \Phi_{\text{logis}}^{-1}(p)\right] = \exp(\mu) [p/(1-p)]^\sigma$.
Moments: For integer $m > 0$,

$$E(T^m) = \exp(m\mu) \Gamma(1 + m\sigma) \Gamma(1 - m\sigma).$$

The m moment is not finite when $m\sigma \geq 1$.

For $\sigma < 1$,

$$E(T) = \exp(\mu) \Gamma(1 + \sigma) \Gamma(1 - \sigma),$$

and for $\sigma < 1/2$,

$$\text{Var}(T) = \exp(2\mu) \left[\Gamma(1 + 2\sigma) \Gamma(1 - 2\sigma) - \Gamma^2(1 + \sigma) \Gamma^2(1 - \sigma) \right].$$

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<div data-bbox="272 279 574 304" data-label="Section-Header"> <p>Other Topics in Chapter 4</p> </div> <div data-bbox="126 354 488 380" data-label="Text"> <p>Pseudorandom number generation.</p> </div> <div data-bbox="683 619 722 636" data-label="Page-Footer"> <p>4 - 31</p> </div>	<div data-bbox="1123 48 1349 73" data-label="Section-Header"> <p>Topics in Chapter 5</p> </div> <div data-bbox="917 123 1550 294" data-label="List-Group"> <ul style="list-style-type: none"> • Parametric models with threshold parameters. • Important distributions used in reliability that can not be translated into location-scale distributions: gamma, generalized gamma, etc. • Finite (discrete) mixture distributions </div> <div data-bbox="1045 312 1440 338" data-label="Equation-Block"> $F(t; \theta) = \xi_1 F_1(t; \theta_1) + \cdots + \xi_k F_k(t; \theta_k)$ </div> <div data-bbox="938 350 1214 378" data-label="Text"> <p>where $\xi_i \geq 0$, and $\sum_i \xi_i = 1$</p> </div> <div data-bbox="917 394 1419 420" data-label="List-Group"> <ul style="list-style-type: none"> • Compound (continuous) mixture distributions. </div> <div data-bbox="938 438 1550 548" data-label="Text"> <p>If failure-times of units in a population are $\text{EXP}(\eta)$ with $1/\eta \sim \text{GAM}(\theta, \kappa)$, then the unconditional failure time, T, of a unit selected at random from the population has a Pareto distribution of the form</p> </div> <div data-bbox="1071 556 1414 606" data-label="Equation-Block"> $F(t; \theta, \kappa) = 1 - \frac{1}{(1 + \theta t)^\kappa}, \quad t > 0.$ </div> <div data-bbox="1497 619 1536 636" data-label="Page-Footer"> <p>4 - 32</p> </div>
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