

## Chapter 9

### Bootstrap Confidence Intervals

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19h 14min

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### Bootstrap Confidence Intervals Chapter 9 Objectives

- Explain basic ideas behind the use of computer simulation to obtain bootstrap confidence intervals.
- Explain different methods for generating bootstrap samples.
- Obtain and interpret simulation-based pointwise parametric bootstrap confidence intervals.

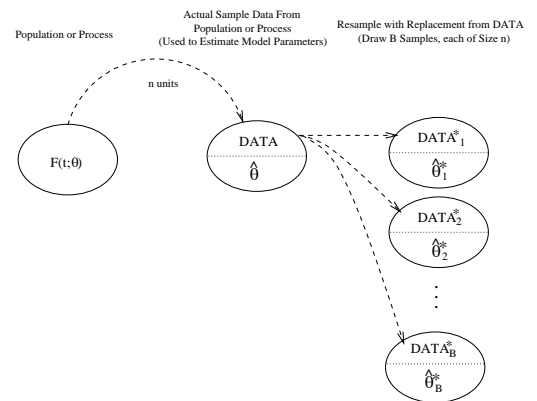
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### Bootstrap Sampling and Bootstrap Confidence Intervals

- Instead of assuming  $Z_{\hat{\mu}} = (\hat{\mu} - \mu) / \widehat{se}_{\hat{\mu}} \sim \text{NOR}(0, 1)$ , use Monte Carlo simulation to approximate the distribution of  $Z_{\hat{\mu}}$ .
- Simulate  $B = 4000$  values of  $Z_{\hat{\mu}^*} = (\hat{\mu}^* - \hat{\mu}) / \widehat{se}_{\hat{\mu}^*}$ .
- Some bootstrap approximations:
  - ▶  $Z_{\hat{\mu}} \sim Z_{\hat{\mu}^*}$
  - ▶  $Z_{\log(\hat{\sigma})} \sim Z_{\log(\hat{\sigma}^*)}$
  - ▶  $Z_{\text{logit}[\hat{F}(t)]} \sim Z_{\text{logit}[\hat{F}^*(t)]}$
 when computing confidence intervals for  $\mu$ ,  $\sigma$ , and  $F$ .

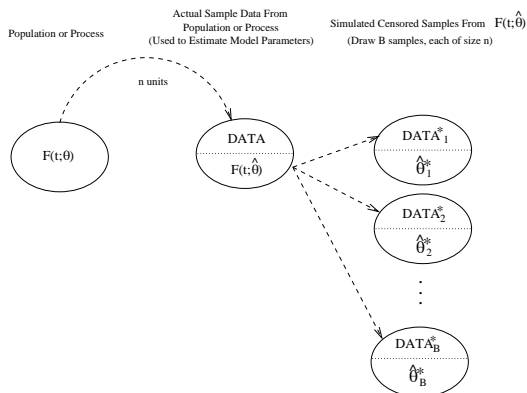
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### A Simple Bootstrap Re-Sampling Method



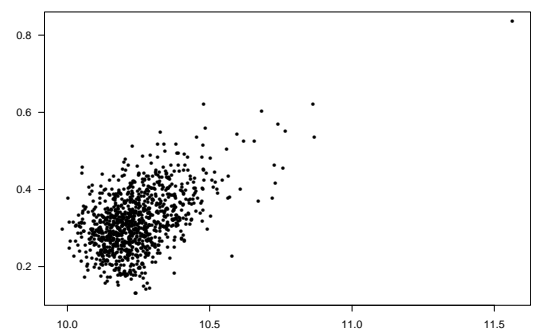
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### A Simple Parametric Bootstrap Sampling Method



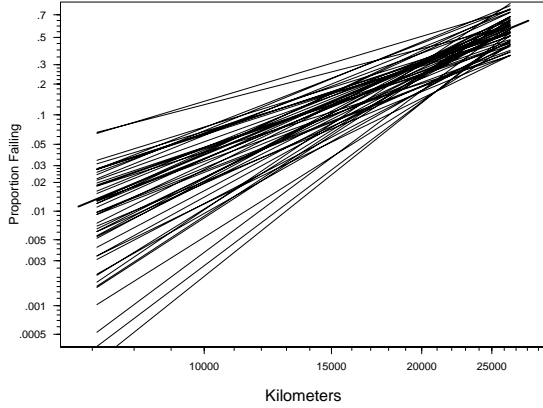
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### Scatterplot of 1,000 (Out of $B=10,000$ ) Bootstrap Estimates $\hat{\mu}^*$ and $\hat{\sigma}^*$ for Shock Absorber



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**Weibull Plot of  $F(t; \hat{\mu}, \hat{\sigma})$  from the Original Sample (dark line) and 50 (Out of  $B=10,000$ )  $F(t; \hat{\mu}^*, \hat{\sigma}^*)$  Computed from Bootstrap Samples for the Shock Absorber**



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**Bootstrap Confidence Interval for  $\mu$**

- With complete data or Type II censoring,

$$Z_{\hat{\mu}_j^*} = \frac{\hat{\mu}_j^* - \hat{\mu}}{\widehat{se}_{\hat{\mu}_j^*}}$$

has a distribution that does not depend on any unknown parameters. Such a quantity is called a **pivotal** quantity.

- By the definition of quantiles, then

$$\Pr\left(z_{\hat{\mu}_{(\alpha/2)}^*} < Z_{\hat{\mu}_j^*} \leq z_{\hat{\mu}_{(1-\alpha/2)}^*}\right) = 1 - \alpha$$

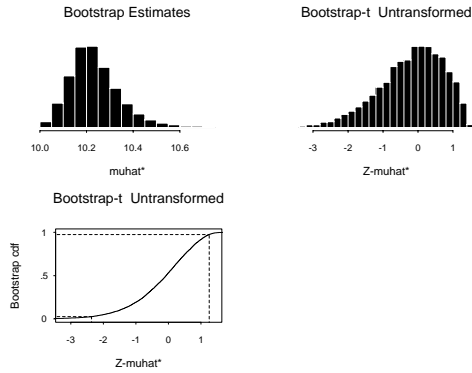
- Simple algebra shows that

$$[\underline{\mu}, \bar{\mu}] = [\hat{\mu} - z_{\hat{\mu}_{(1-\alpha/2)}^*} \widehat{se}_{\hat{\mu}}, \hat{\mu} - z_{\hat{\mu}_{(\alpha/2)}^*} \widehat{se}_{\hat{\mu}}]$$

provides an exact 95% confidence interval for  $\mu$ . With other kinds of censoring, the interval is, in general, only **approximate**.

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**Bootstrap Distributions of Weibull  $\hat{\mu}^*$  and  $Z_{\hat{\mu}^*}$  Based on  $B=10,000$  Bootstrap Samples for the Shock Absorber**



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**Bootstrap Confidence Interval for  $\sigma$**

- With complete data or Type II censoring,

$$Z_{\log(\hat{\sigma}^*)} = \frac{\log(\hat{\sigma}^*) - \log(\hat{\sigma})}{\widehat{se}_{\log(\hat{\sigma}^*)}}$$

has a distribution that does not depend on any unknown parameters. Such a quantity is called a **pivotal** quantity.

- By the definition of quantiles, then

$$\Pr\left(z_{\log(\hat{\sigma}^*)_{(\alpha/2)}} < Z_{\log(\hat{\sigma}_j^*)} \leq z_{\log(\hat{\sigma}^*)_{(1-\alpha/2)}}\right) = 1 - \alpha$$

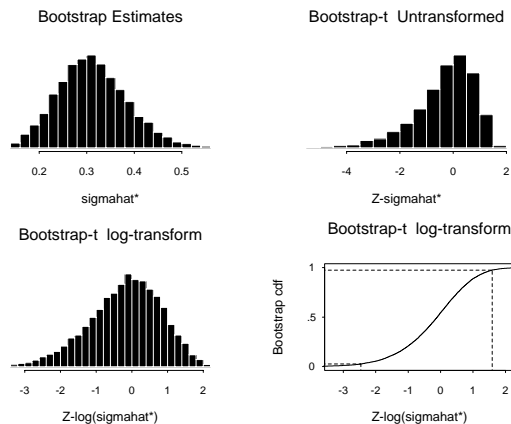
- Simple algebra shows that

$$[\underline{\sigma}, \bar{\sigma}] = [\hat{\sigma}/\underline{w}, \hat{\sigma}/\bar{w}]$$

provides an exact 95% confidence interval for  $\sigma$ , where  $\underline{w} = \exp\left[z_{\log(\hat{\sigma}^*)_{(1-\alpha/2)}} \widehat{se}_{\log(\hat{\sigma})}\right]$  and  $\bar{w} = \exp\left[z_{\log(\hat{\sigma}^*)_{(\alpha/2)}} \widehat{se}_{\log(\hat{\sigma})}\right]$ . With other kinds of censoring, the interval is, in general, only **approximate**.

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**Bootstrap Distributions of  $\hat{\sigma}^*$ ,  $Z_{\hat{\sigma}^*}$ , and  $Z_{\log(\hat{\sigma}^*)}$  Based on  $B=10,000$  Bootstrap Samples**



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**Bootstrap Confidence Interval for  $F(t_e)$**

- With complete data or Type II censoring [using  $F = F(t_e)$ ],

$$Z_{\logit(\hat{F}^*)} = \frac{\logit(\hat{F}^*) - \logit(\hat{F})}{\widehat{se}_{\logit(\hat{F}^*)}}$$

has a distribution that does not depend on any unknown parameters. Such a quantity is called a **pivotal** quantity.

- By the definition of quantiles, then

$$\Pr\left(z_{\logit(\hat{F}^*)_{(\alpha/2)}} < Z_{\logit(\hat{F}_j^*)} \leq z_{\logit(\hat{F}^*)_{(1-\alpha/2)}}\right) = 1 - \alpha$$

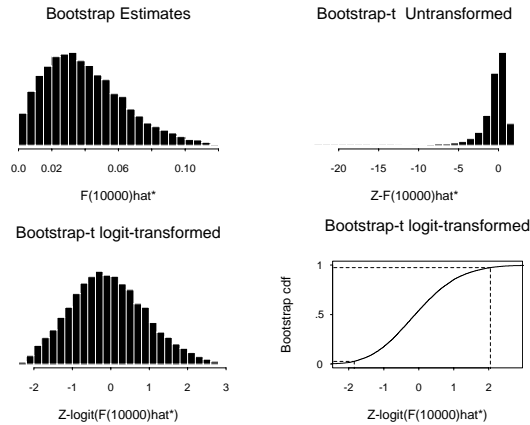
- Simple algebra shows that

$$[\underline{F}, \bar{F}] = \left[ \frac{\hat{F}}{\hat{F} + (1 - \hat{F}) \times \underline{w}}, \frac{\hat{F}}{\hat{F} + (1 - \hat{F}) \times \bar{w}} \right]$$

where provides an exact 95% confidence interval for  $F$ , where  $\underline{w} = \exp\left[z_{\logit(\hat{F}^*)_{(1-\alpha/2)}} \widehat{se}_{\logit(\hat{F})}\right]$  and  $\bar{w} = \exp\left[z_{\logit(\hat{F}^*)_{(\alpha/2)}} \widehat{se}_{\logit(\hat{F})}\right]$ . With other kinds of censoring, the interval is, in general, only **approximate**.

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**Bootstrap Distributions of  $\hat{F}(t_e)^*$ ,  $Z_{\hat{F}(t_e)^*}$ , and  $Z_{\log[\hat{F}(t_e)^*]}$  for  $t_e=10,000$  km Based on  $B=10,000$  Bootstrap Samples**



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**Bootstrap Confidence Interval for  $t_p$**

- With complete data or Type II censoring,

$$Z_{\log(\hat{t}_p^*)} = \frac{\log(\hat{t}_p^*) - \log(\hat{t}_p)}{\widehat{SE}_{\log(\hat{t}_p^*)}}$$

has a distribution that does not depend on any unknown parameters. Such a quantity is called a **pivotal** quantity.

- By the definition of quantiles, then

$$\Pr\left(z_{\log(\hat{t}_p^*)_{(\alpha/2)}} < Z_{\log(\hat{t}_p^*)_j} \leq z_{\log(\hat{t}_p^*)_{(1-\alpha/2)}}\right) = 1 - \alpha$$

- Simple algebra shows that

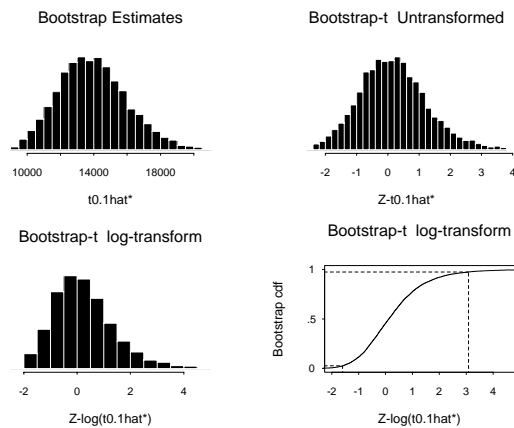
$$[t_p, \quad \tilde{t}_p] = [\hat{t}_p/w, \quad \hat{t}_p/\tilde{w}]$$

provides an exact 95% confidence interval for  $t_p$ , where  $w =$

$\exp\left[z_{\log(\hat{t}_p^*)_{(1-\alpha/2)}} \widehat{SE}_{\log(\hat{t}_p)}\right]$  and  $\tilde{w} = \exp\left[z_{\log(\hat{t}_p^*)_{(\alpha/2)}} \widehat{SE}_{\log(\hat{t}_p)}\right]$   
With other kinds of censoring, the interval is, in general, only **approximate**.

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**Bootstrap Distributions of  $\hat{t}_p^*$ ,  $Z_{\hat{t}_p^*}$ , and  $Z_{\log[\hat{t}_p^']}$  for  $t_e=10,000$  km Based on  $B=10,000$  Bootstrap Samples**



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