

Chapter 6

Probability Plotting

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19h 14min

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Chapter 6

Probability Plotting Objectives

- Describe **applications** for probability plots.
- Explain the basic **concepts** of probability plotting.
- Show how to **linearize** a cdf on special plotting scales.
- Explain how to plot a nonparametric estimate \hat{F} to judge the adequacy of a particular parametric distribution.
- Explain methods of separating **useful** information from **noise** when interpreting a probability plot.
- Use a probability plot to obtain **graphical** estimates of reliability characteristics like failure probabilities and quantiles.

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Purposes of Probability Plots

Probability plots are used to:

- Assess the adequacy of a particular distributional model.
- To detect multiple failure modes or mixture of different populations.
- Displaying the results of a parametric maximum likelihood fit along with the data.
- Obtain, by drawing a smooth curve through the points, a semiparametric estimate of failure probabilities and distributional quantiles.
- Obtain graphical estimates of model parameters (e.g., by fitting a straight line through the points on a probability plot).

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Probability Plotting Scales: Linearizing a CDF

Main Idea: For a given cdf, $F(t)$, one can **linearize** the $\{ t \text{ versus } F(t) \}$ plot by:

- Finding transformations of $F(t)$ and t such that the relationship between the transformed variables is linear.
- The transformed axes can be relabeled in terms of the original probability and time variables.

The resulting probability axis is generally nonlinear and is called the **probability** scale. The data axis is usually a linear axis or a log axis.

6 - 4

Linearizing the Exponential CDF

CDF: $p = F(t; \theta, \gamma) = 1 - \exp\left[-\frac{(t-\gamma)}{\theta}\right], \quad t \geq \gamma.$

Quantiles: $t_p = \gamma - \theta \log(1 - p).$

Conclusion:

The $\{ t_p \text{ versus } -\log(1 - p) \}$ plot is a straight line.

We plot t_p on the horizontal axis and p on the vertical axis. γ is the **intercept** on the time axis and $1/\theta$ is equal to the slope of the cdf line.

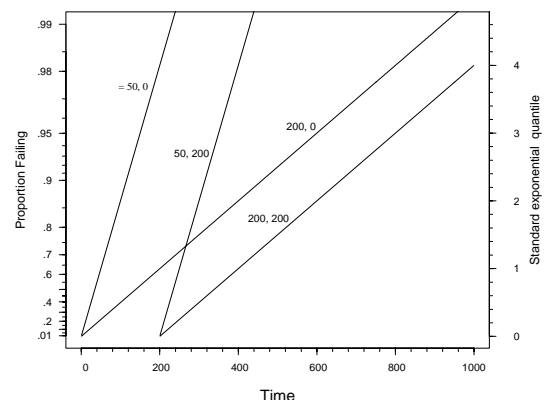
Note:

Changing θ changes the slope of the line and changing γ changes the position of the line.

6 - 5

Plot with Exponential Distribution Probability Scales Showing Exponential cdfs as Straight Lines for Combinations of Parameters $\theta = 50, 200$ and $\gamma = 0, 200$

$$t_p = \gamma - \theta \log(1 - p)$$



6 - 6

Linearizing the Normal CDF

$$\text{CDF: } p = F(y; \mu, \sigma) = \Phi_{\text{nor}}\left(\frac{y-\mu}{\sigma}\right), \quad -\infty < y < \infty.$$

$$\text{Quantiles: } y_p = \mu + \sigma \Phi_{\text{nor}}^{-1}(p).$$

$\Phi_{\text{nor}}^{-1}(p)$ is the p quantile of the standard normal distribution.

Conclusion:

$\{y_p \text{ versus } \Phi_{\text{nor}}^{-1}(p)\}$ will plot as a straight line.

μ is the point at the time axis where the cdf intersects the $\Phi^{-1}(p) = 0$ line (i.e., $p = .5$). The slope of the cdf line on the graph is $1/\sigma$.

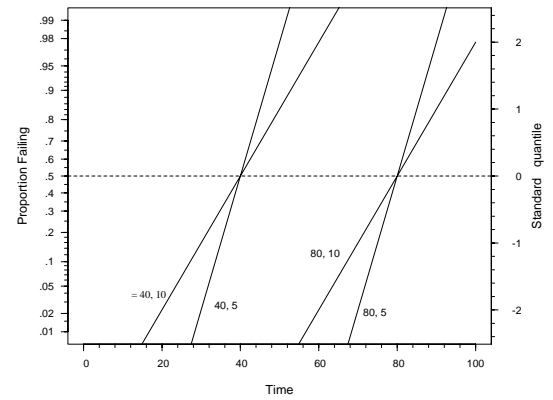
Note:

Any normal cdf plots as a straight line with positive slope. Also, any straight line with positive slope corresponds to a normal cdf.

6-7

Plot with Normal Distribution Probability Scales Showing Normal cdfs as Straight Lines for Combinations of Parameters $\mu = 40, 80$ and $\sigma = 5, 10$

$$y_p = \mu + \sigma \Phi_{\text{nor}}^{-1}(p)$$



6-8

Linearizing the Lognormal CDF

$$\text{CDF: } p = F(t; \mu, \sigma) = \Phi_{\text{nor}}\left[\frac{\log(t)-\mu}{\sigma}\right], \quad t > 0.$$

$$\text{Quantiles: } t_p = \exp\left[\mu + \sigma \Phi_{\text{nor}}^{-1}(p)\right].$$

$$\text{Then } \log(t_p) = \mu + \sigma \Phi_{\text{nor}}^{-1}(p)$$

Conclusion:

$\{\log(t_p) \text{ versus } \Phi_{\text{nor}}^{-1}(p)\}$ will plot as a straight line.

$\exp(\mu)$ can be read from the time axis at the point where the cdf intersects the $\Phi_{\text{nor}}^{-1}(p) = 0$ line. The slope of the cdf line on the graph is $1/\sigma$ (but in the computations use base e logarithms for the times rather than the base 10 logarithms used for the figures).

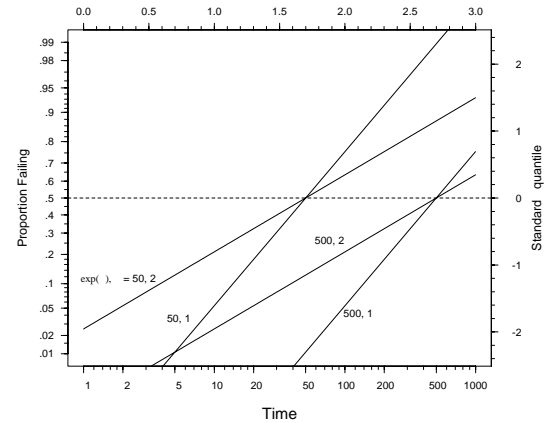
Note:

Any given lognormal cdf plots as a straight line with positive slope. Also, any straight line with positive slope corresponds to a lognormal distribution.

6-9

Plot with Lognormal Distribution Probability Scales Showing Lognormal cdfs as Straight Lines for Combinations of $\exp(\mu) = 50, 500$ and $\sigma = 1, 2$

$$\log(t_p) = \mu + \sigma \Phi_{\text{nor}}^{-1}(p)$$



6-10

Linearizing the Weibull CDF

$$\text{CDF: } p = F(t; \mu, \sigma) = \Phi_{\text{sev}}\left[\frac{\log(t)-\mu}{\sigma}\right], \quad t > 0.$$

$$\text{Quantiles: } t_p = \exp\left[\mu + \sigma \Phi_{\text{sev}}^{-1}(p)\right] = \eta[-\log(1-p)]^{1/\beta},$$

where $\Phi_{\text{sev}}^{-1}(p) = \log[-\log(1-p)]$, $\eta = \exp(\mu)$, $\beta = 1/\sigma$.

This leads to

$$\log(t_p) = \mu + \log[-\log(1-p)]\sigma = \log(\eta) + \log[-\log(1-p)]\frac{1}{\beta}$$

Conclusion:

$\{\log(t_p) \text{ versus } \log[-\log(1-p)]\}$ will plot as a straight line.

6-11

Linearizing the Weibull CDF-Continued

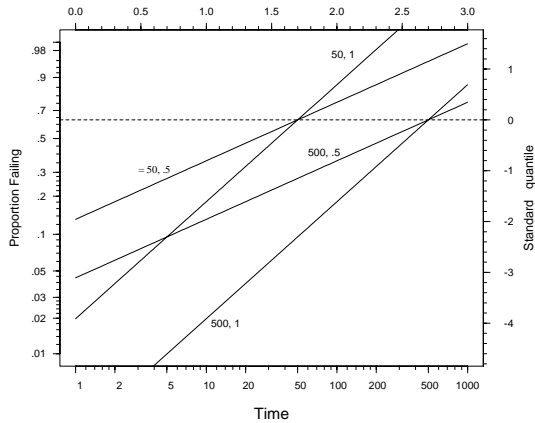
Comments:

- $\eta = \exp(\mu)$ can be read from the time axis at the point where the cdf intersects the $\log[-\log(1-p)] = 0$ line, which corresponds to $p \approx 0.632$.
- The slope of the cdf line on the graph is $\beta = 1/\sigma$ (but in the computations use base e logarithms for the times rather than the base 10 logarithms used for the figures).
- Any Weibull cdf plots as a straight line with positive slope. And any straight line with positive slope corresponds to a Weibull cdf.
- Exponential cdfs plot as straight lines with slopes equal to 1.

6-12

**Plot with Weibull Distribution Probability Scales
Showing Weibull cdfs as Straight Lines for
Combinations of $\eta = 50, 500$ and $\beta = .5, 1$**

$$\log(t_p) = \log(\eta) + \log[-\log(1-p)]^{\frac{1}{\beta}}$$



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**Choosing Plotting Positions to
Plot the Nonparametric Estimate of F**

- The **discontinuity** and **randomness** of $\hat{F}(t)$ make it difficult to choose a definition for pairs of points (t, \hat{F}) to plot.
- With times reported as **exact**, it has been traditional to plot $\{t_i \text{ versus } \hat{F}(t_i)\}$ at the observed failure times.

General Idea: Plot an estimate of F at some specified set of points in time and define **plotting** positions consisting of a corresponding estimate of F at these points in time.

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Criteria for Choosing Plotting Positions

Criteria for choosing plotting positions should depend on the **application** or **purpose** for constructing the probability plot.

Some applications that suggest criteria:

- Checking distributional assumptions.
- Estimation of parameters.
- Display of maximum likelihood results with data.

6 - 15

**Plotting Positions: Continuous Inspection Data
and Multiple Censoring**

$\hat{F}(t)$ is a step function until the last reported failure time, but the step increases may be different than $1/n$.

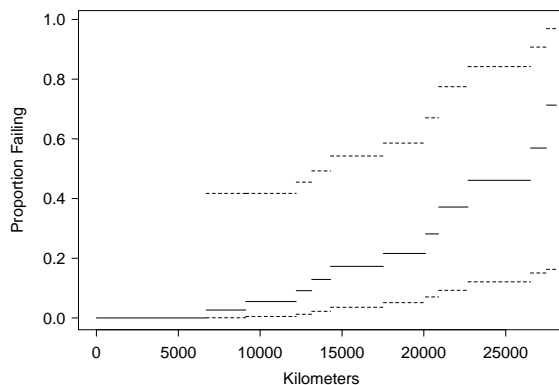
Plotting Positions: $\{t_{(i)} \text{ versus } p_i\}$ with

$$p_i = \frac{1}{2} \{ \hat{F}[t_{(i)} + \Delta] + \hat{F}[t_{(i)} - \Delta] \}.$$

Justification: This is consistent with the definition for single censoring.

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**Nonparametric Estimate of $F(t)$ for the Shock
Absorbers. Simultaneous Approximate 95%
Confidence Bands for $F(t)$**



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**Plotting Positions: Continuous Inspection Data and
Single Censoring**

Let $t_{(1)}, t_{(2)}, \dots$ be the ordered failure times. When there is not ties, $\hat{F}(t)$ is a step function increasing by an amount $1/n$ until the last reported failure time.

Plotting Positions: $\{t_i \text{ versus } \frac{i-.5}{n}\}$.

- **Justification:**

$$\frac{i-.5}{n} = \frac{1}{2} \{ \hat{F}[t_{(i)} + \Delta] + \hat{F}[t_{(i)} - \Delta] \}$$

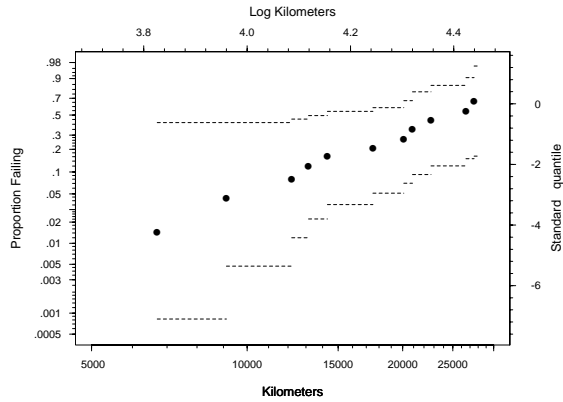
$$E[t_{(i)}] \approx F^{-1}\left(\frac{i-.5}{n}\right).$$

where Δ is positive and small.

- When the model fits well, the ML line approximately goes through the points.
- Need to adjust these plotting positions when there are ties.

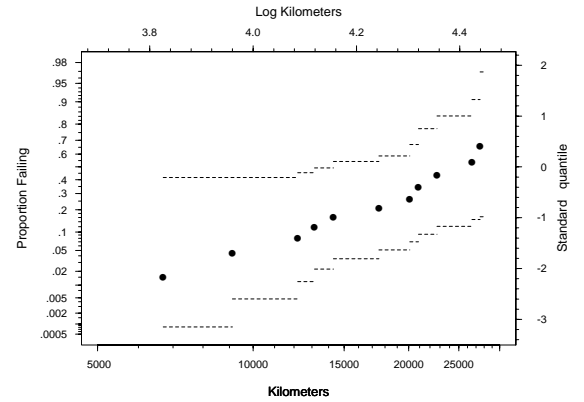
6 - 18

Weibull Probability Plot of the Shock Absorber Data.
Also Shown are Simultaneous Approximate 95% Confidence Bands for $F(t)$



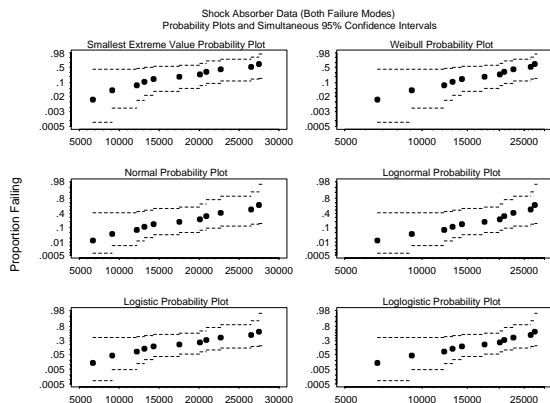
6 - 19

Lognormal Probability Plot of the Shock Absorber Data. Also Shown are Simultaneous Approximate 95% Confidence Bands for $F(t)$



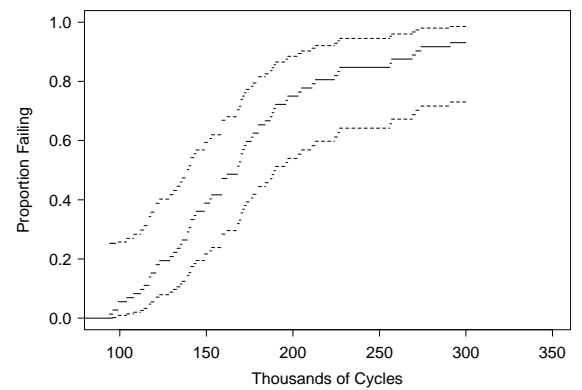
6 - 20

Six-Distribution Probability Plots of the Shock Absorber Data



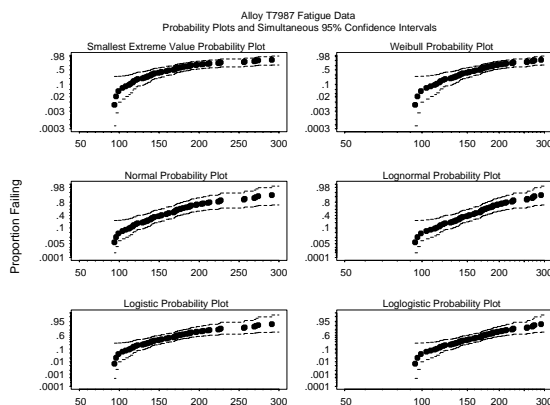
6 - 21

Plot of Nonparametric Estimate of $F(t)$ for the Alloy T7987 Fatigue Life and Simultaneous Approximate 95% Confidence Bands for $F(t)$



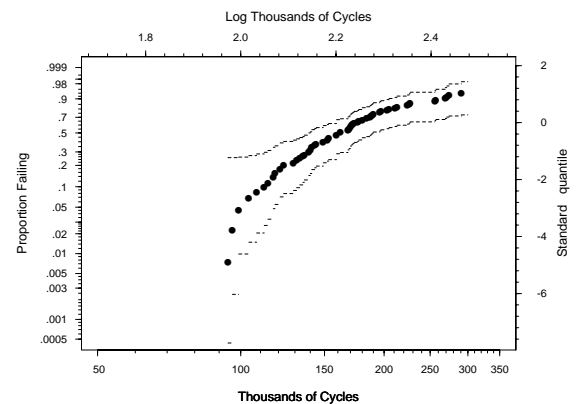
6 - 22

Six-Distribution Probability Plots Alloy T7987 Fatigue Life



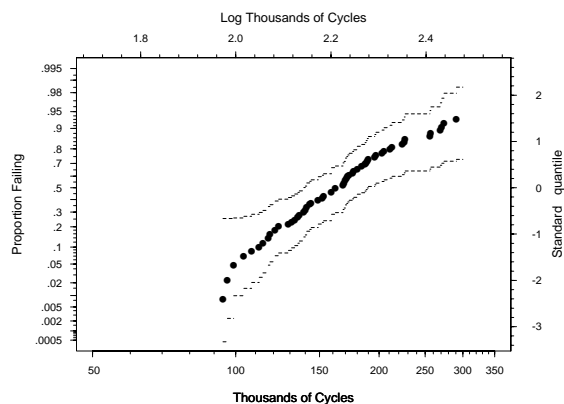
6 - 23

Weibull Probability Plot for the Alloy T7987 Fatigue Life and Simultaneous Approximate 95% Confidence Bands for $F(t)$



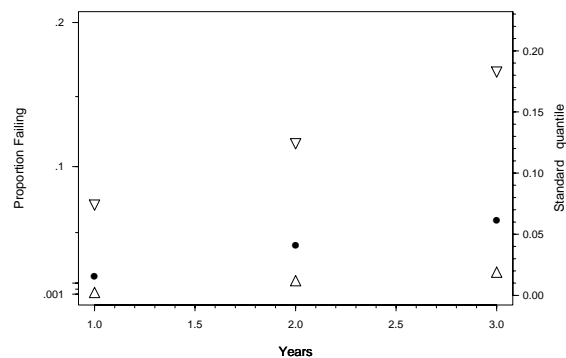
6 - 24

Lognormal Probability Plot for the Alloy T7987 Fatigue Life and Simultaneous Approximate 95% Confidence Bands for $F(t)$



6 - 25

Exponential Distribution Probability Plot of the Heat-Exchanger Tube Crack Data and Simultaneous Approximate 95% Confidence Bands for $F(t)$



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Plotting Positions: Interval Censored Inspection Data

Let $(t_0, t_1], \dots, (t_{m-1}, t_m]$ be the inspection times.

The upper endpoints of the inspection intervals $t_i, i = 1, 2, \dots$, are convenient plotting times.

Plotting Positions: $\{t_i \text{ versus } p_i\}$, with

$$p_i = \hat{F}(t_i)$$

When there are no censored observations beyond t_m , $F(t_m) = 1$ and this point cannot be plotted on probability paper.

Justification: with no losses, from standard binomial theory,

$$E[\hat{F}(t_i)] = F(t_i).$$

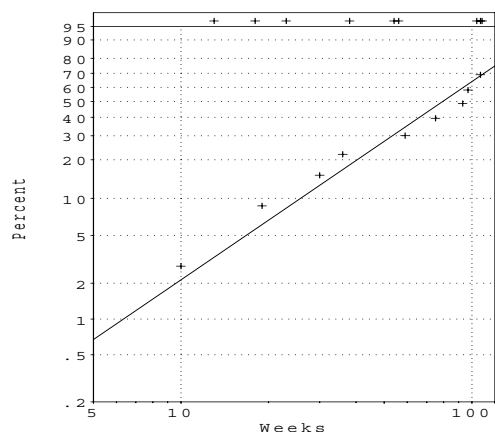
6 - 27

Biomedical Examples

Here we show some SAS[®] Proc Reliability probability plots for the IUD data.

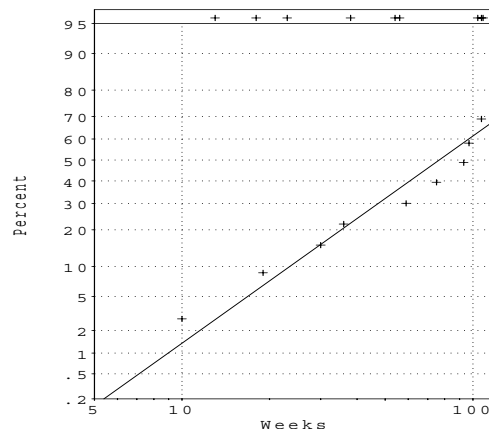
6 - 28

SAS[®] Proc Reliability Weibull Probability Plot of the IUD Data



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SAS[®] Proc Reliability Nonparametric Lognormal Probability Plot of the IUD Data



6 - 30

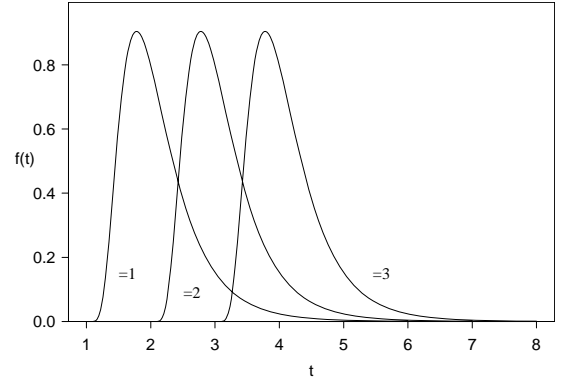
Probability Plots with Specified Shape Parameters

The probability plotting techniques can be extended to construct probability plots for:

- Distributions that are not members of the location-scale family.
- To help identify, graphically, the need for non-zero threshold parameter.
- Estimate graphically a shape parameter.

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Pdf for three-parameter lognormal distributions for $\mu = 0$ and $\sigma = .5$ with $\gamma = 1, 2, 3$



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Distributions with a Threshold Parameter

- The lognormal, Weibull, gamma, and other similar distributions can be generalized by the addition of a **threshold** parameter, γ , to shift the beginning of the distribution away from 0.
- These distributions are particularly useful for fitting skewed distributions that are shifted far to the right of 0.
- For example, the cdf and quantiles of the 3-parameter log-normal distribution can be expressed as

$$p = F(t; \mu, \sigma, \gamma) = \Phi_{\text{nor}} \left[\frac{\log(t - \gamma) - \mu}{\sigma} \right], \quad t > \gamma$$

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Linearizing the 3-Parameter Gamma CDF

$$\text{CDF:} \quad p = F(t; \theta, \kappa, \gamma) = \Gamma_I \left(\frac{t - \gamma}{\theta}; \kappa \right), \quad t > \gamma.$$

$$\text{Quantiles:} \quad t_p = \gamma + \Gamma_I^{-1}(p; \kappa) \theta.$$

where $\Gamma_I(z; \kappa) = \int_0^z x^{\kappa-1} e^{-x} dx / \Gamma(\kappa)$ and $\Gamma(\kappa) = \int_0^\infty x^{\kappa-1} e^{-x} dx$.

Conclusion:

$\{ t_p \text{ versus } \Gamma_I^{-1}(p; \kappa) \}$ will plot as a straight line.

The probability axis **depends** on specification of the shape parameter κ .

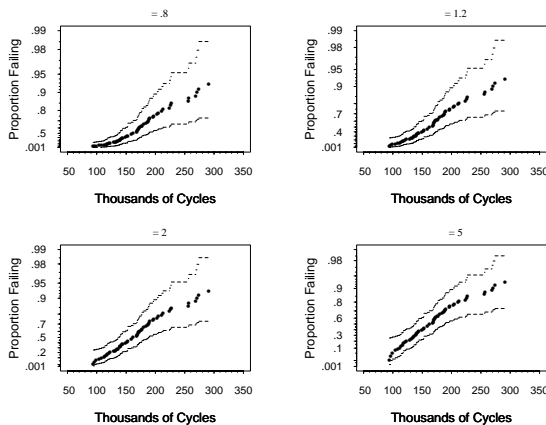
γ is the intercept on the time axis (because $\Gamma_I^{-1}(p; \kappa) = 0$ when $p = 0$). The slope of the cdf line is equal to $1/\theta$.

Note:

Changing θ changes the slope of the line and changing γ changes the position of the line.

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Gamma Probability Plot with $\kappa = .8, 1.2, 2, 5$ for the Alloy T7987 Fatigue Life with Simultaneous Approximate 95% Confidence Bands for $F(t)$



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Linearizing the 3-Parameter Weibull CDF Using Linear Time Axis and Specified Shape Parameter

$$\text{CDF:} \quad p = F(t; \mu, \sigma) = \Phi_{\text{sev}} \left[\frac{\log(t - \gamma) - \mu}{\sigma} \right], \quad t > \gamma.$$

$$\text{Quantiles:} \quad t_p = \gamma + \eta [-\log(1 - p)]^{1/\beta},$$

where $\Phi_{\text{sev}}(z) = 1 - \exp[-\exp(z)]$, $\eta = \exp(\mu)$, $\beta = 1/\sigma$.

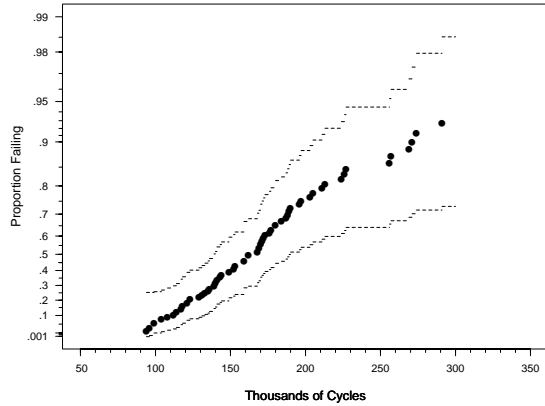
Conclusion:

$\{ t_p \text{ versus } [-\log(1 - p)]^{1/\beta} \}$ will plot as a straight line.

- The probability axis for this linear-time-axis Weibull probability plot requires specification of the shape parameter β .
- γ is the intercept on the time axis. The slope of the cdf line is equal to $1/\eta$.
- The plot allows graphical estimation the threshold parameter γ .

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Linear-Scale Weibull Plot with $\beta = 1.4$ for the Alloy T7987 Fatigue Life with Simultaneous Approximate 95% Confidence Bands for $F(t)$



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Linearizing the Generalized Gamma CDF

$$\text{CDF: } p = F(t; \theta, \beta, \kappa) = \Gamma_I \left[\left(\frac{t}{\theta} \right)^\beta; \kappa \right].$$

$$\text{Quantiles: } t_p = \theta \left[\Gamma_I^{-1}(p; \kappa) \right]^{1/\beta}.$$

Then $\log(t_p) = \log(\theta) + \log[\Gamma_I^{-1}(p; \kappa)]^{1/\beta}$.

Conclusion:

$\{ \log(t_p) \text{ versus } \log[\Gamma_I^{-1}(p; \kappa)] \}$ will plot as a straight line.

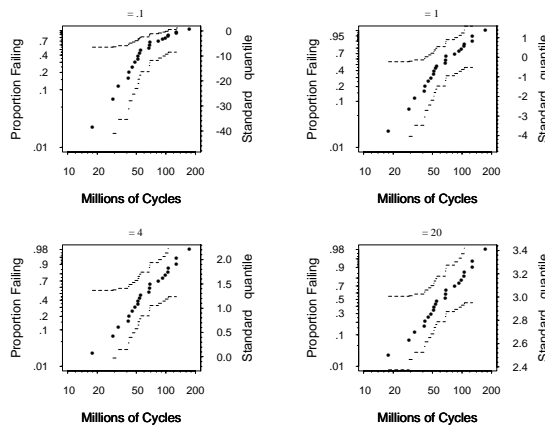
The scale parameter θ is the intercept on the time scale, corresponding to the time where the cdf crosses the horizontal line at $\log[\Gamma_I^{-1}(p; \kappa)] = 0$.

The slope of the line on the graph with time on the horizontal axis is β .

Note: The probability scale for the GENG probability plot requires a given value of the shape parameter κ .

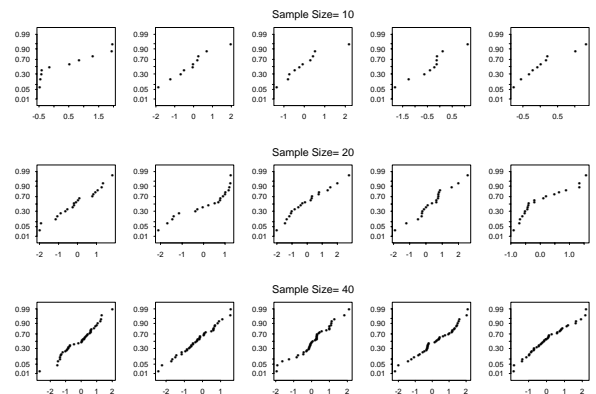
6 - 38

GENG Probability Plots of the Ball Bearing Fatigue Data with Specified $\kappa = .1, 1, 4$, and 20



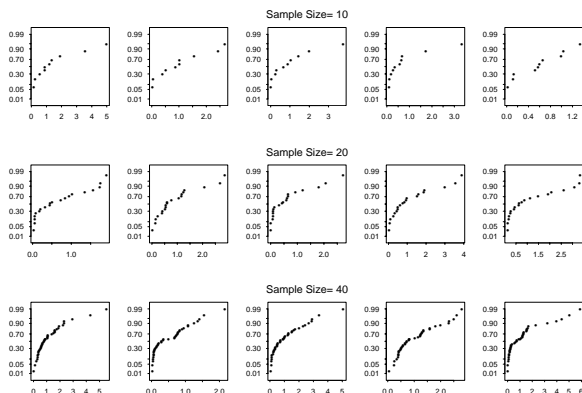
6 - 39

Random Normal Variates Plotted on Normal Probability Plots with Sample Sizes of $n=10, 20$, and 40. Five Replications of Each Probability Plot



6 - 40

Random Exponential Variates Plotted on Normal Probability Plots with Sample Sizes of $n=10, 20$, and 40. Five Replications of Each Probability Plot



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Notes on the Application of Probability Plotting

- Using simulation to help interpret probability plots
 - Try different assumed distributions and compare the results.
 - Assess linearity; allowing for more variability in the tails.
 - * Use simultaneous nonparametric confidence bands.
 - * Use simulation or bootstrap to calibrate.
- Possible reason for a bend in a probability plot
 - Sharp bend or change in slope generally indicates an abrupt change in a failure process.

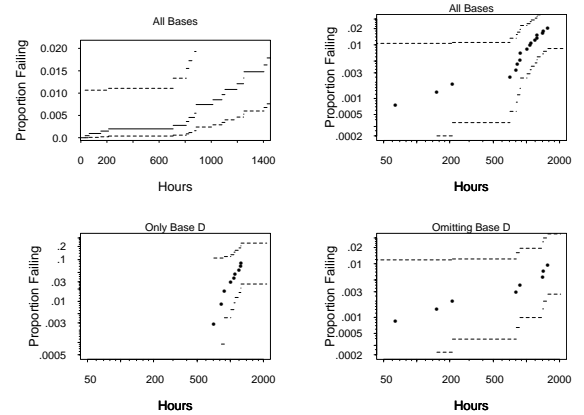
6 - 42

Bleed System Failure Data (Abernethy, Breneman, Medlin, and Reinman 1983)

- Failure and running times for 2256 bleed systems.
- The Weibull probability plot suggest changes in the failure distribution after 600 hours. The data shows that 9 of the 19 failures had occurred at Base D.

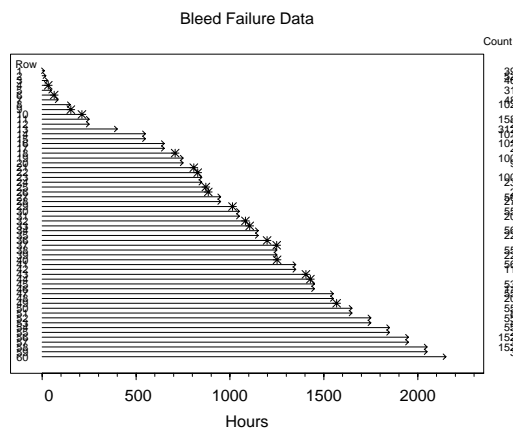
6 - 43

Bleed System Failure Data Analysis CDF plot and Weibull Probability Plots



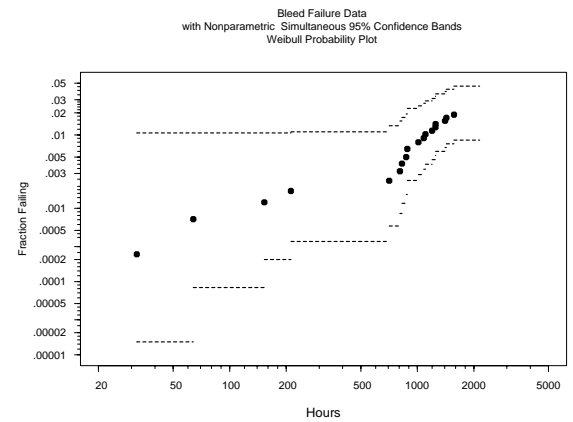
6 - 44

Bleed System (All Bases) Event Plot



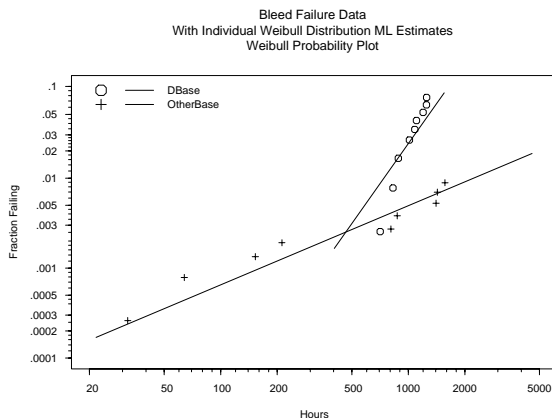
6 - 45

Bleed System (All Bases) Weibull Probability Plot



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Bleed System Separate Weibull Probability Plots for Base D and Other Bases



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Bleed System Failure Data Analysis-Conclusions

- Separate analyses of the Base D data and the data from the other bases indicated different failure distributions.
- The large slope ($\beta \approx 5$) for Base D indicated strong wearout.
- The relatively small slope for the other bases ($\beta \approx .85$) suggested infant mortality or accidental failures.
- The problem at base D was caused by salt air. A change in maintenance procedures there solved the main part of the reliability problem with the bleed systems.

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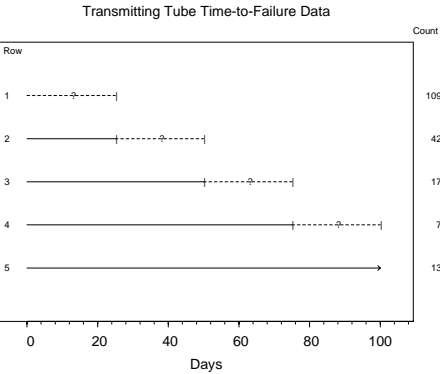
Transmitter Vacuum Tube Data (Davis 1952)

- Life data for a certain kind of transmitter vacuum tube used in the output stage of high-power transmitters.
- The data are read-out (interval censored) data.

Days		Number Failing
Interval	Endpoint	
Lower	Upper	
0	25	109
25	50	42
50	75	17
75	100	7
100	∞	13

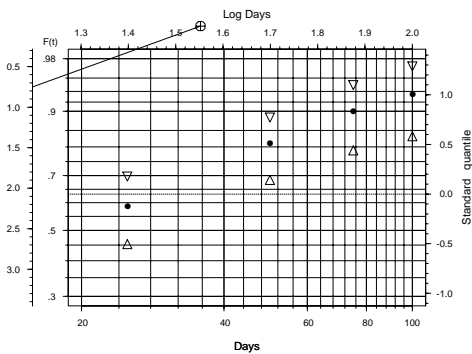
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V7 Transmitter Tube Failure Data Event Plot



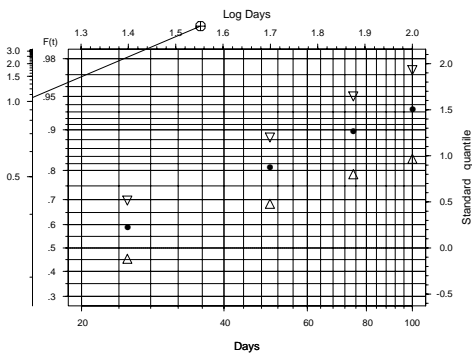
6 - 50

Weibull Probability Plot of the V7 Transmitter Tube Failure Data with Simultaneous Approximate 95% Confidence Bands for $F(t)$



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Lognormal Probability Plot of the V7 Transmitter Tube Failure Data with Simultaneous Approximate 95% Confidence Bands for $F(t)$



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Other Topics in Chapter 6

Probability plotting for arbitrarily censored data.

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